Balanced Allocations: A refined drift theorem

Dimitrios Los¹, Thomas Sauerwald¹, John Sylvester²

¹University of Cambridge, UK, ²University of Liverpool, UK





Outline

Outline

- Balanced allocations (background and some highlights)
- The exponential and hyperbolic cosine potential functions
- The proof of the drift theorem
- The refinement and its applications
- Open problems

Balanced allocations: Background

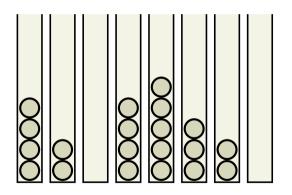
Allocate m tasks (balls) sequentially into n machines (bins).

Allocate m tasks (balls) sequentially into n machines (bins).

Goal: minimise the **maximum load** $\max_{i \in [n]} x_i^m$, where x^t is the load vector after ball t.

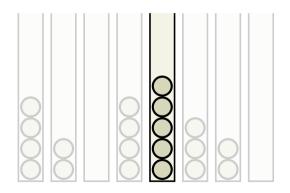
Allocate m tasks (balls) sequentially into n machines (bins).

Goal: minimise the **maximum load** $\max_{i \in [n]} x_i^m$, where x^t is the load vector after ball t.



Allocate m tasks (balls) sequentially into n machines (bins).

Goal: minimise the **maximum load** $\max_{i \in [n]} x_i^m$, where x^t is the load vector after ball t.

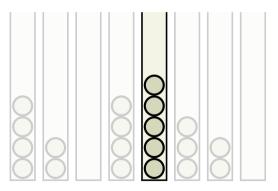


Balanced allocations: Background

Allocate m tasks (balls) sequentially into n machines (bins).

Goal: minimise the **maximum load** $\max_{i \in [n]} x_i^m$, where x^t is the load vector after ball t.

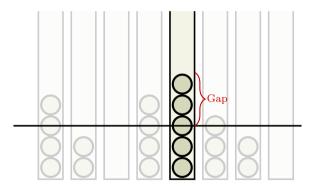
 \Leftrightarrow minimise the **gap**, where $\operatorname{Gap}(m) = \max_{i \in [n]} (x_i^m - m/n)$.



Allocate m tasks (balls) sequentially into n machines (bins).

Goal: minimise the **maximum load** $\max_{i \in [n]} x_i^m$, where x^t is the load vector after ball t.

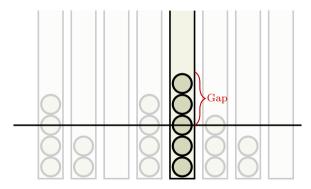
 \Leftrightarrow minimise the gap, where $\operatorname{Gap}(m) = \max_{i \in [n]} (x_i^m - m/n)$.



Allocate m tasks (balls) sequentially into n machines (bins).

Goal: minimise the **maximum load** $\max_{i \in [n]} x_i^m$, where x^t is the load vector after ball t.

 \Leftrightarrow minimise the gap, where $\operatorname{Gap}(m) = \max_{i \in [n]} (x_i^m - m/n)$.



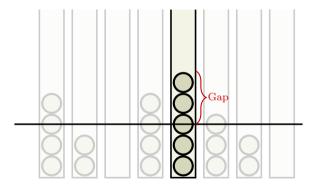
Connections to

Balanced allocations: Background

Allocate m tasks (balls) sequentially into n machines (bins).

Goal: minimise the **maximum load** $\max_{i \in [n]} x_i^m$, where x^t is the load vector after ball t.

 \Leftrightarrow minimise the gap, where $\operatorname{Gap}(m) = \max_{i \in [n]} (x_i^m - m/n)$.

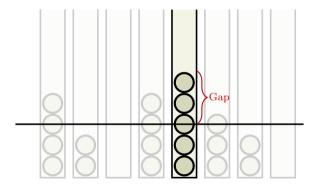


Connections to occupancy problems,

Allocate m tasks (balls) sequentially into n machines (bins).

Goal: minimise the **maximum load** $\max_{i \in [n]} x_i^m$, where x^t is the load vector after ball t.

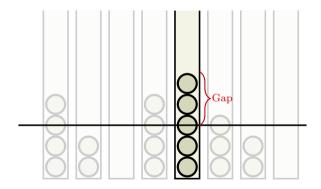
 \Leftrightarrow minimise the gap, where $\operatorname{Gap}(m) = \max_{i \in [n]} (x_i^m - m/n)$.



Connections to occupancy problems, urn processes

Allocate m tasks (balls) sequentially into n machines (bins).

Goal: minimise the **maximum load** $\max_{i \in [n]} x_i^m$, where x^t is the load vector after ball t. \Leftrightarrow minimise the **gap**, where $\operatorname{Gap}(m) = \max_{i \in [n]} (x_i^m - m/n)$.

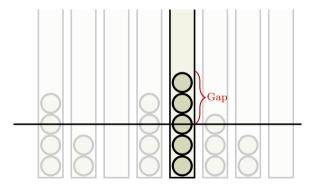


Connections to occupancy problems, urn processes and queuing theory.

Allocate m tasks (balls) sequentially into n machines (bins).

Goal: minimise the **maximum load** $\max_{i \in [n]} x_i^m$, where x^t is the load vector after ball t.

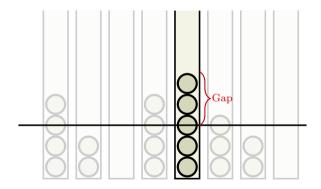
 \Leftrightarrow minimise the gap, where $\operatorname{Gap}(m) = \max_{i \in [n]} (x_i^m - m/n)$.



Connections to occupancy problems, urn processes and queuing theory. Applications in hashing.

Allocate m tasks (balls) sequentially into n machines (bins).

Goal: minimise the **maximum load** $\max_{i \in [n]} x_i^m$, where x^t is the load vector after ball t. \Leftrightarrow minimise the **gap**, where $\operatorname{Gap}(m) = \max_{i \in [n]} (x_i^m - m/n)$.

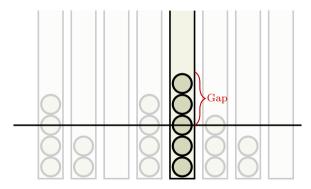


Connections to occupancy problems, urn processes and queuing theory. Applications in hashing, load balancing

Allocate m tasks (balls) sequentially into n machines (bins).

Goal: minimise the **maximum load** $\max_{i \in [n]} x_i^m$, where x^t is the load vector after ball t.

 \Leftrightarrow minimise the gap, where $\operatorname{Gap}(m) = \max_{i \in [n]} (x_i^m - m/n)$.



Connections to occupancy problems, urn processes and queuing theory. Applications in hashing, load balancing and routing.

One-Choice Process:

Iteration: For each $t \geq 0$, sample **one** bin uniformly at random (u.a.r.) and place the ball there.

ONE-CHOICE Process:

Iteration: For each $t \geq 0$, sample **one** bin uniformly at random (u.a.r.) and place the ball there.

■ In the lightly-loaded case (m = n), w.h.p. $Gap(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$ [Gon81].

Balanced allocations: Background

One-Choice Process:

Iteration: For each $t \geq 0$, sample **one** bin uniformly at random (u.a.r.) and place the ball there.

In the lightly-loaded case (m=n), w.h.p. $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$ [Gon81].

Meaning with probability at least $1 - n^{-c}$ for constant c > 0.

One-Choice Process:

Iteration: For each $t \geq 0$, sample **one** bin uniformly at random (u.a.r.) and place the ball there.

- In the lightly-loaded case (m=n), w.h.p. $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$ [Gon81].
- In the heavily-loaded case $(m \gg n)$, w.h.p. $\operatorname{Gap}(m) = \Theta\left(\sqrt{\frac{m}{n} \cdot \log n}\right)$ (e.g. [RS98]).

One-Choice Process:

Iteration: For each $t \geq 0$, sample **one** bin uniformly at random (u.a.r.) and place the ball there.

- In the lightly-loaded case (m = n), w.h.p. $Gap(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$ [Gon81].
- In the heavily-loaded case $(m \gg n)$, w.h.p. $\operatorname{Gap}(m) = \Theta\left(\sqrt{\frac{m}{n} \cdot \log n}\right)$ (e.g. [RS98]).

TWO-CHOICE Process:

Iteration: For each $t \ge 0$, sample two bins independently u.a.r. and place the ball in the least loaded of the two.

One-Choice Process:

Iteration: For each $t \geq 0$, sample **one** bin uniformly at random (u.a.r.) and place the ball there.

- In the lightly-loaded case (m = n), w.h.p. $Gap(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$ [Gon81].
- In the heavily-loaded case $(m \gg n)$, w.h.p. $Gap(m) = \Theta\left(\sqrt{\frac{m}{n} \cdot \log n}\right)$ (e.g. [RS98]).

TWO-CHOICE Process:

Iteration: For each $t \ge 0$, sample **two** bins independently u.a.r. and place the ball in the least loaded of the two.

■ In the lightly-loaded case (m = n), w.h.p. $Gap(n) = log_2 log n + \Theta(1)$ [KLMadH96, ABKU99].

ONE-CHOICE Process:

Iteration: For each $t \geq 0$, sample one bin uniformly at random (u.a.r.) and place the ball there.

- In the lightly-loaded case (m = n), w.h.p. $\operatorname{Gap}(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$ [Gon81].
- In the heavily-loaded case $(m \gg n)$, w.h.p. $\operatorname{Gap}(m) = \Theta\left(\sqrt{\frac{m}{n} \cdot \log n}\right)$ (e.g. [RS98]).

TWO-CHOICE Process:

Iteration: For each $t \geq 0$, sample two bins independently u.a.r. and place the ball in the least loaded of the two.

■ In the lightly-loaded case (m = n), w.h.p. $Gap(n) = log_2 log n + \Theta(1)$ [KLMadH96, ABKU99].

ONE-CHOICE Process:

Iteration: For each $t \geq 0$, sample **one** bin uniformly at random (u.a.r.) and place the ball there.

- In the lightly-loaded case (m = n), w.h.p. $Gap(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$ [Gon81].
- In the heavily-loaded case $(m \gg n)$, w.h.p. $Gap(m) = \Theta\left(\sqrt{\frac{m}{n} \cdot \log n}\right)$ (e.g. [RS98]).

TWO-CHOICE Process:

Iteration: For each $t \ge 0$, sample **two** bins independently u.a.r. and place the ball in the least loaded of the two.

- In the lightly-loaded case (m = n), w.h.p. $Gap(n) = log_2 log n + \Theta(1)$ [KLMadH96, ABKU99].
- In the heavily-loaded case $(m \gg n)$, w.h.p. $Gap(m) = \log_2 \log n + \Theta(1)$ [BCSV06].

One-Choice Process:

Iteration: For each $t \geq 0$, sample **one** bin uniformly at random (u.a.r.) and place the ball there.

- In the lightly-loaded case (m = n), w.h.p. $Gap(n) = \Theta\left(\frac{\log n}{\log \log n}\right)$ [Gon81].
- In the heavily-loaded case $(m \gg n)$, w.h.p. $Gap(m) = \Theta\left(\sqrt{\frac{m}{n} \cdot \log n}\right)$ (e.g. [RS98]).

TWO-CHOICE Process:

Iteration: For each $t \ge 0$, sample **two** bins independently u.a.r. and place the ball in the least loaded of the two.

- In the lightly-loaded case (m = n), w.h.p. $\operatorname{Gap}(n) = \log_2 \log n + \Theta(1)$ [KLMadH96, ABKU99].
- In the heavily-loaded case $(m \gg n)$, w.h.p. $Gap(m) = \log_2 \log n + \Theta(1)$ [BCSV06].

The practical significance of the "power of two choices" was recognised in the 2020 ACM Theory and Practice Award $[ABK^+20]$:

The practical significance of the "power of two choices" was recognised in the 2020 ACM Theory and Practice Award [ABK+20]:

"[...] it is not surprising that the power of two choices that requires only a local decision rather than global coordination has led to a wide range of practical applications. These include i-Google's web index, Akamai's overlay routing network, and highly reliable distributed data storage systems used by Microsoft and Dropbox, which are all based on variants of the power of two choices paradigm."

The practical significance of the "power of two choices" was recognised in the 2020 ACM Theory and Practice Award [ABK+20]:

"[...] it is not surprising that the power of two choices that requires only a local decision rather than global coordination has led to a wide range of practical applications. These include i-Google's web index, Akamai's overlay routing network, and highly reliable distributed data storage systems used by Microsoft and Dropbox, which are all based on variants of the power of two choices paradigm."

The practical significance of the "power of two choices" was recognised in the 2020 ACM Theory and Practice Award [ABK+20]:

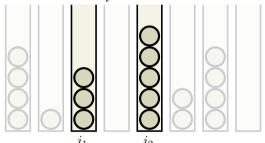
"[...] it is not surprising that the power of two choices that requires only a local decision rather than global coordination has led to a wide range of practical applications. These include i-Google's web index, Akamai's overlay routing network, and highly reliable distributed data storage systems used by Microsoft and Dropbox, which are all based on variants of the power of two choices paradigm."

Question: Why variants and not vanilla Two-Choice?

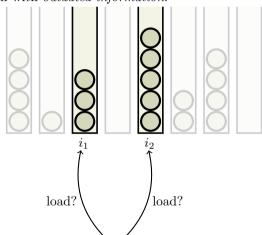
■ Mitzenmacher [Mit00] and Dahlin [Dah00] empirically observed that **Two-Choice** does not perform well with *outdated information*.

Balanced allocations: Background

■ Mitzenmacher [Mit00] and Dahlin [Dah00] empirically observed that Two-Choice does not perform well with *outdated information*.

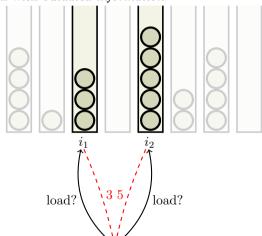


■ Mitzenmacher [Mit00] and Dahlin [Dah00] empirically observed that **Two-Choice** does not perform well with *outdated information*.



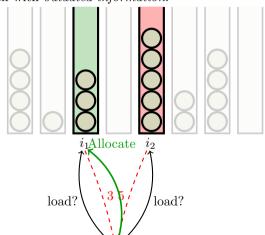
Balanced allocations: Background

■ Mitzenmacher [Mit00] and Dahlin [Dah00] empirically observed that **Two-Choice** does not perform well with *outdated information*.



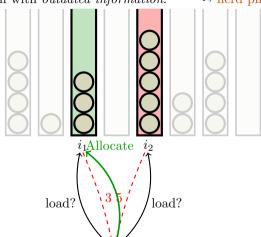
Balanced allocations: Background

■ Mitzenmacher [Mit00] and Dahlin [Dah00] empirically observed that Two-Choice does not perform well with *outdated information*.



■ Mitzenmacher [Mit00] and Dahlin [Dah00] empirically observed that Two-Choice does not perform well with outdated information.

→ herd phenomenon



- Mitzenmacher [Mit00] and Dahlin [Dah00] empirically observed that Two-Choice does not perform well with outdated information.

 → herd phenomenon
- Several *low-latency schedulers* use variants of Two-Choice (Eagle [DDDZ16], Hawk [DDKZ15], Peacock [KG18]).

- Mitzenmacher [Mit00] and Dahlin [Dah00] empirically observed that Two-Choice does not perform well with outdated information.

 → herd phenomenon
- Several *low-latency schedulers* use variants of TWO-CHOICE (Eagle [DDDZ16], Hawk [DDKZ15], Peacock [KG18]). For Sparrow [OWZS13], they remark

- Mitzenmacher [Mit00] and Dahlin [Dah00] empirically observed that Two-Choice does not perform well with outdated information.

 → herd phenomenon
- Several low-latency schedulers use variants of TWO-CHOICE (Eagle [DDDZ16], Hawk [DDKZ15], Peacock [KG18]). For Sparrow [OWZS13], they remark

 The power of two choices suffers from two remaining performance problems: first, server queue length is a poor indicator of wait time, and second, due to messaging delays. multiple schedulers sampling in parallel may experience race conditions.

- Mitzenmacher [Mit00] and Dahlin [Dah00] empirically observed that Two-Choice does not perform well with outdated information.

 → herd phenomenon
- Several low-latency schedulers use variants of TWO-CHOICE (Eagle [DDDZ16], Hawk [DDKZ15], Peacock [KG18]). For Sparrow [OWZS13], they remark

 The power of two choices suffers from two remaining performance problems: first, server queue length is a poor indicator of wait time, and second, due to messaging delays, multiple schedulers sampling in parallel may experience race conditions.
- As noted in [LXK⁺11], *communication* is a shortcoming of TWO-CHOICE in some real-word load balancers:

- Mitzenmacher [Mit00] and Dahlin [Dah00] empirically observed that Two-Choice does not perform well with outdated information.

 → herd phenomenon
- Several low-latency schedulers use variants of TWO-CHOICE (Eagle [DDDZ16], Hawk [DDKZ15], Peacock [KG18]). For Sparrow [OWZS13], they remark

 The power of two choices suffers from two remaining performance problems: first, server queue length is a poor indicator of wait time, and second, due to messaging delays, multiple schedulers sampling in parallel may experience race conditions.
- As noted in [LXK⁺11], *communication* is a shortcoming of Two-Choice in some real-word load balancers:
 - More importantly, the [Two-Choice] algorithm requires communication between dispatchers and processors at the time of job assignment. The communication time is on the critical path, hence contributes to the increase in response time.

- Mitzenmacher [Mit00] and Dahlin [Dah00] empirically observed that Two-Choice does not perform well with outdated information.

 → herd phenomenon
- Several low-latency schedulers use variants of TWO-CHOICE (Eagle [DDDZ16], Hawk [DDKZ15], Peacock [KG18]). For Sparrow [OWZS13], they remark

 The power of two choices suffers from two remaining performance problems: first, server queue length is a poor indicator of wait time, and second, due to messaging delays, multiple schedulers sampling in parallel may experience race conditions.
- As noted in [LXK⁺11], *communication* is a shortcoming of Two-Choice in some real-word load balancers:
 - More importantly, the [Two-Choice] algorithm requires communication between dispatchers and processors at the time of job assignment. The communication time is on the critical path, hence contributes to the increase in response time.
- In the *queuing setting*, Whitt [Whi86] remarks:

- Mitzenmacher [Mit00] and Dahlin [Dah00] empirically observed that Two-Choice does not perform well with outdated information.

 → herd phenomenon
- Several low-latency schedulers use variants of TWO-CHOICE (Eagle [DDDZ16], Hawk [DDKZ15], Peacock [KG18]). For Sparrow [OWZS13], they remark

 The power of two choices suffers from two remaining performance problems: first, server queue length is a poor indicator of wait time, and second, due to messaging delays, multiple schedulers sampling in parallel may experience race conditions.
- As noted in [LXK⁺11], *communication* is a shortcoming of TWO-CHOICE in some real-word load balancers:
 - More importantly, the [Two-Choice] algorithm requires communication between dispatchers and processors at the time of job assignment. The communication time is on the critical path, hence contributes to the increase in response time.
- In the *queuing setting*, Whitt [Whi86] remarks:
 - We have shown that several natural selection rules are not optimal in various situations, but we have not identified any optimal rules. Identifying optimal rules in these situations would obviously be interesting, but appears to be difficult.

An example of a variant of TWO-CHOICE

$(1+\beta)$ -Process:

Parameter: A mixing factor $\beta \in (0, 1]$.

Iteration: For each $t \ge 0$, with probability β allocate one ball via the Two-Choice process, otherwise allocate one ball via the ONE-Choice process.

An example of a variant of Two-Choice

$(1+\beta)$ -Process:

Parameter: A mixing factor $\beta \in (0,1]$.

Iteration: For each $t \ge 0$, with probability β allocate one ball via the Two-Choice process, otherwise allocate one ball via the ONE-CHOICE process.

■ Introduced by Mitzenmacher [Mit96] as a faulty setting for Two-Choice.

An example of a variant of Two-Choice

$(1+\beta)$ -Process:

Parameter: A mixing factor $\beta \in (0,1]$.

Iteration: For each $t \ge 0$, with probability β allocate one ball via the Two-Choice process, otherwise allocate one ball via the ONE-CHOICE process.

- Introduced by Mitzenmacher [Mit96] as a faulty setting for Two-Choice.
- In the heavily-loaded case, [PTW15] proved that the gap is w.h.p. $\mathcal{O}(\frac{\log n}{\beta})$ for any $\beta \in (0, 1]$.

$(1+\beta)$ -Process:

Parameter: A mixing factor $\beta \in (0, 1]$.

Iteration: For each $t \ge 0$, with probability β allocate one ball via the Two-Choice process, otherwise allocate one ball via the ONE-CHOICE process.

- Introduced by Mitzenmacher [Mit96] as a faulty setting for Two-Choice.
- In the heavily-loaded case, [PTW15] proved that the gap is w.h.p. $\mathcal{O}(\frac{\log n}{\beta})$ for any $\beta \in (0, 1]$.
- It has been used to analyze

$(1+\beta)$ -Process:

Parameter: A mixing factor $\beta \in (0,1]$.

Iteration: For each $t \ge 0$, with probability β allocate one ball via the Two-Choice process, otherwise allocate one ball via the ONE-Choice process.

- Introduced by Mitzenmacher [Mit96] as a faulty setting for Two-Choice.
- In the heavily-loaded case, [PTW15] proved that the gap is w.h.p. $\mathcal{O}(\frac{\log n}{\beta})$ for any $\beta \in (0, 1]$.
- It has been used to analyze population protocols [AAG18, AGR21],

$(1+\beta)$ -Process:

Parameter: A mixing factor $\beta \in (0, 1]$.

Iteration: For each $t \ge 0$, with probability β allocate one ball via the Two-Choice process, otherwise allocate one ball via the ONE-CHOICE process.

- Introduced by Mitzenmacher [Mit96] as a faulty setting for Two-Choice.
- In the heavily-loaded case, [PTW15] proved that the gap is w.h.p. $\mathcal{O}(\frac{\log n}{\beta})$ for any $\beta \in (0,1]$.
- It has been used to analyze population protocols [AAG18, AGR21], distributed data structures [ABK+18, AKLN17, Nad21]

$(1+\beta)$ -Process:

Parameter: A mixing factor $\beta \in (0, 1]$.

Iteration: For each $t \ge 0$, with probability β allocate one ball via the Two-Choice process, otherwise allocate one ball via the ONE-CHOICE process.

- Introduced by Mitzenmacher [Mit96] as a faulty setting for Two-Choice.
- In the heavily-loaded case, [PTW15] proved that the gap is w.h.p. $\mathcal{O}(\frac{\log n}{\beta})$ for any $\beta \in (0,1]$.
- It has been used to analyze population protocols [AAG18, AGR21], distributed data structures [ABK+18, AKLN17, Nad21] and online carpooling [GKKS20].

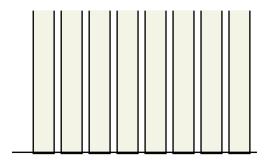
$(1+\beta)$ -Process:

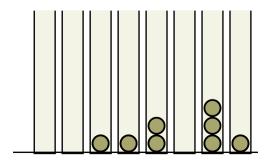
Parameter: A mixing factor $\beta \in (0,1]$.

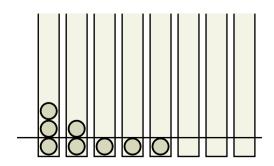
Iteration: For each $t \ge 0$, with probability β allocate one ball via the Two-Choice process, otherwise allocate one ball via the ONE-CHOICE process.

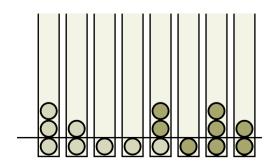
- Introduced by Mitzenmacher [Mit96] as a faulty setting for Two-Choice.
- In the heavily-loaded case, [PTW15] proved that the gap is w.h.p. $\mathcal{O}(\frac{\log n}{\beta})$ for any $\beta \in (0, 1]$.
- It has been used to analyze population protocols [AAG18, AGR21], distributed data structures [ABK+18, AKLN17, Nad21] and online carpooling [GKKS20].

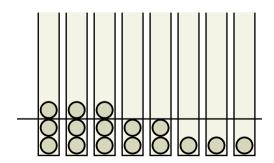
Question: Why choose a $\beta < 1$?

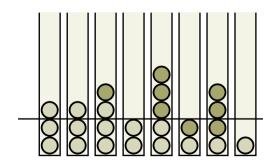


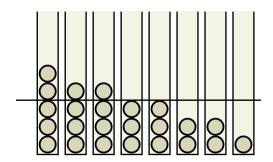


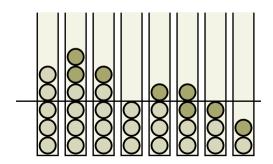


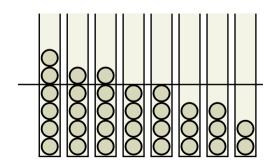


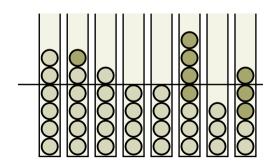


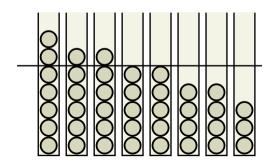




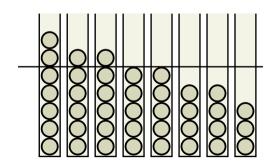




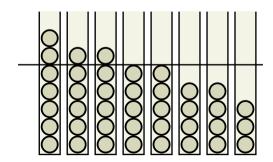




- Berenbrink, Czumaj, Englert, Friedetzky and Nagel [BCE⁺12] studied Two-Choice where balls are allocated in *batches* of size *b* (*b*-BATCHED setting).
- For $b \ge n \log n$, we show that Two-Choice has w.h.p. $\operatorname{Gap}(m) = \Theta(\frac{b}{n})$.

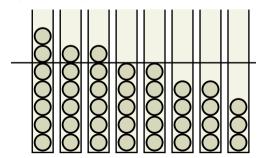


- Berenbrink, Czumaj, Englert, Friedetzky and Nagel [BCE⁺12] studied Two-Choice where balls are allocated in *batches* of size *b* (*b*-BATCHED setting).
- For $b \ge n \log n$, we show that Two-Choice has w.h.p. $\operatorname{Gap}(m) = \Theta(\frac{b}{n})$.
- In constast to the case $b \ll n \log n$, where it is asymptotically optimal.



- Berenbrink, Czumaj, Englert, Friedetzky and Nagel [BCE⁺12] studied Two-Choice where balls are allocated in *batches* of size *b* (*b*-BATCHED setting).
- For $b \ge n \log n$, we show that Two-Choice has w.h.p. $\operatorname{Gap}(m) = \Theta(\frac{b}{n})$.
- In constast to the case $b \ll n \log n$, where it is asymptotically optimal.
- For $b \ge n \log n$, we show that the $(1 + \beta)$ -process has

w.h.p.
$$\operatorname{Gap}(m) = \Theta\left(\sqrt{\frac{b}{n} \cdot \log n}\right)$$
, for $\beta = \Theta(\sqrt{(n/b) \cdot \log n})$.



Probability allocation vectors

Probability allocation vectors

Probability allocation vector p^t , where p_i^t is the prob. of allocating to *i*-th most loaded bin.

Probability allocation vectors

- **Probability allocation vector** p^t , where p_i^t is the prob. of allocating to *i*-th most loaded bin.
- For One-Choice,

$$p_{\text{ONE-CHOICE}} = \left(\frac{1}{n}, \frac{1}{n}, \dots, \frac{1}{n}\right).$$

Probability allocation vectors

- **Probability allocation vector** p^t , where p_i^t is the prob. of allocating to *i*-th most loaded bin.
- For One-Choice,

$$p_{\text{ONE-CHOICE}} = \left(\frac{1}{n}, \frac{1}{n}, \dots, \frac{1}{n}\right).$$

For Two-Choice,

$$p_{\text{TWO-CHOICE}} = \left(\frac{1}{n^2}, \frac{3}{n^2}, \dots, \frac{2i-1}{n^2}, \dots, \frac{2n-2}{n^2}\right).$$

Probability allocation vectors

- **Probability allocation vector** p^t , where p_i^t is the prob. of allocating to *i*-th most loaded bin.
- For One-Choice,

$$p_{ ext{One-Choice}} = \left(\frac{1}{n}, \frac{1}{n}, \dots, \frac{1}{n}\right).$$

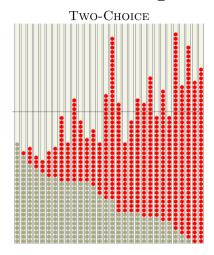
For Two-Choice,

$$p_{\text{TWo-Choice}} = \left(\frac{1}{n^2}, \frac{3}{n^2}, \dots, \frac{2i-1}{n^2}, \dots, \frac{2n-2}{n^2}\right).$$

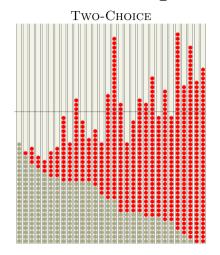
 \blacksquare For $(1+\beta)$ -process,

$$p_{(1+\beta)} = \left(\dots, \beta \cdot \frac{2i-1}{n^2} + (1-\beta) \cdot \frac{1}{n}, \dots\right).$$

An example of a variant of Two-Choice



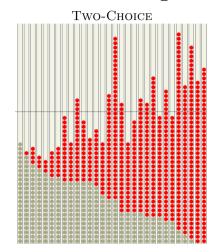
$$p_i = \frac{2i-1}{n^2}$$



$$(1+eta)$$
-Process

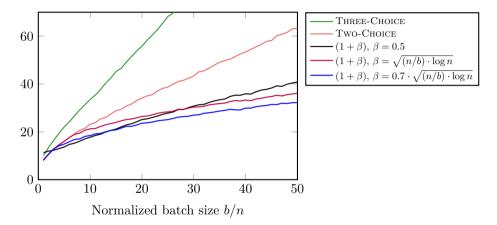
$$p_i = \frac{2i-1}{n^2}$$

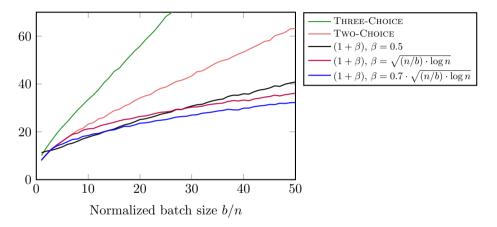
$$p_i = \beta \cdot \frac{2i-1}{n^2} + (1-\beta) \cdot \frac{1}{n}$$



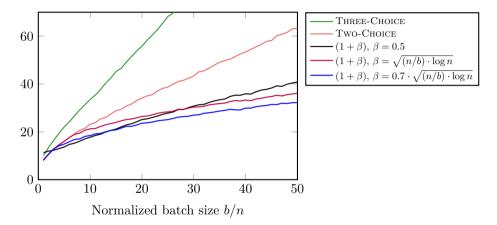
$$p_i = \frac{2i - 1}{n^2}$$

$$p_i = \beta \cdot \frac{2i-1}{n^2} + (1-\beta) \cdot \frac{1}{n}$$

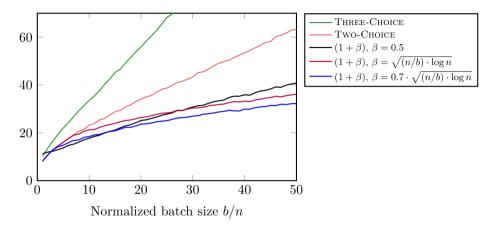




■ The gaps are decreasingly ordered by p_n : $\approx \frac{3}{n}$ (for Three-Choice),



■ The gaps are decreasingly ordered by p_n : $\approx \frac{3}{n}$ (for Three-Choice), $\approx \frac{2}{n}$ (for Two-Choice)



■ The gaps are decreasingly ordered by p_n : $\approx \frac{3}{n}$ (for Three-Choice), $\approx \frac{2}{n}$ (for Two-Choice) and $\approx \frac{1+\beta}{n}$ (for the $(1+\beta)$ -processes).

Potential functions

Techniques for analyzing balanced allocations

Layered induction

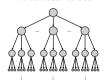


Two-Choice, Memory

Poissonisation

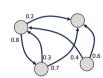
 $X_i \sim \mathsf{Poi}(\frac{m}{n})$ Unweighted, time-independent

Witness trees



 ${\bf Two\text{-}Choice,\ parallel\ allocations}$

Markov chains



Some weights, b-Batched, heterogeneous sampling

Graphical processes



Two-Choice

Potential functions



weights, b-Batched, outdated info, noise graphical, heterogeneous sampling

■ Following [PTW15] we will study ϵ -biased processes that:

- Following [PTW15] we will study ϵ -biased processes that:
 - ightharpoonup Have p is non-decreasing,

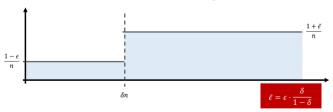
- Following [PTW15] we will study ϵ -biased processes that:
 - \triangleright Have p is non-decreasing,
 - ▶ For some constant $\delta \in (0,1)$, satisfy

$$p_{\delta n} \le \frac{1 - \epsilon}{n}.$$

- Following [PTW15] we will study ϵ -biased processes that:
 - \triangleright Have p is non-decreasing,
 - ▶ For some constant $\delta \in (0,1)$, satisfy

$$p_{\delta n} \le \frac{1 - \epsilon}{n}.$$

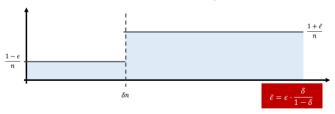
■ **Key Observation 1:** We can assume the following *worst-case* allocation vector



- Following [PTW15] we will study ϵ -biased processes that:
 - ightharpoonup Have p is non-decreasing,
 - ▶ For some constant $\delta \in (0,1)$, satisfy

$$p_{\delta n} \le \frac{1 - \epsilon}{n}.$$

Key Observation 1: We can assume the following worst-case allocation vector



Our main aim will be to derive the w.h.p. $\mathcal{O}((\log n)/\epsilon)$ gap, for any $\epsilon \in (0,1)$.

■ At step $t \ge 0$, the *exponential potential* with smoothing parameter $\alpha > 0$ is defined as

$$\Phi^t := \sum_{i=1}^n e^{\alpha \cdot (x_i^t - t/n)}$$

 \blacksquare At step $t \ge 0$, the *exponential potential* with smoothing parameter $\alpha > 0$ is defined as

$$\Phi^t := \sum_{i=1}^n e^{\alpha \cdot (x_i^t - t/n)} = \sum_{i=1}^n e^{\alpha \cdot y_i^t}.$$

■ At step $t \ge 0$, the *exponential potential* with smoothing parameter $\alpha > 0$ is defined as

$$\Phi^t := \sum_{i=1}^n e^{\alpha \cdot (x_i^t - t/n)} = \sum_{i=1}^n e^{\alpha \cdot y_i^t}.$$

Our goal is to show that $\Phi^t = \mathcal{O}(\text{poly}(n))$,

At step $t \geq 0$, the *exponential potential* with smoothing parameter $\alpha > 0$ is defined as

$$\Phi^t := \sum_{i=1}^n e^{\alpha \cdot (x_i^t - t/n)} = \sum_{i=1}^n e^{\alpha \cdot y_i^t}.$$

Our goal is to show that $\Phi^t = \mathcal{O}(\text{poly}(n))$, as it implies that

$$y_i^t \leq \frac{1}{\alpha} \cdot \log \Phi^t$$

At step $t \geq 0$, the *exponential potential* with smoothing parameter $\alpha > 0$ is defined as

$$\Phi^t := \sum_{i=1}^n e^{\alpha \cdot (x_i^t - t/n)} = \sum_{i=1}^n e^{\alpha \cdot y_i^t}.$$

Our goal is to show that $\Phi^t = \mathcal{O}(\text{poly}(n))$, as it implies that

$$y_i^t \le \frac{1}{\alpha} \cdot \log \Phi^t = \mathcal{O}\left(\frac{1}{\alpha} \cdot \log n\right).$$

At step $t \geq 0$, the *exponential potential* with smoothing parameter $\alpha > 0$ is defined as

$$\Phi^t := \sum_{i=1}^n e^{\alpha \cdot (x_i^t - t/n)} = \sum_{i=1}^n e^{\alpha \cdot y_i^t}.$$

Our goal is to show that $\Phi^t = \mathcal{O}(\text{poly}(n))$, as it implies that

$$y_i^t \le \frac{1}{\alpha} \cdot \log \Phi^t = \mathcal{O}\left(\frac{1}{\alpha} \cdot \log n\right).$$

Even better, if we show that $\Phi^t = \mathcal{O}(n)$,

At step $t \geq 0$, the *exponential potential* with smoothing parameter $\alpha > 0$ is defined as

$$\Phi^t := \sum_{i=1}^n e^{\alpha \cdot (x_i^t - t/n)} = \sum_{i=1}^n e^{\alpha \cdot y_i^t}.$$

Our goal is to show that $\Phi^t = \mathcal{O}(\text{poly}(n))$, as it implies that

$$y_i^t \le \frac{1}{\alpha} \cdot \log \Phi^t = \mathcal{O}\left(\frac{1}{\alpha} \cdot \log n\right).$$

Even better, if we show that $\Phi^t = \mathcal{O}(n)$, then the bins with load $\geq t/n + z$ is at most $\mathcal{O}(n \cdot e^{-\alpha z})$.

At step $t \geq 0$, the *exponential potential* with smoothing parameter $\alpha > 0$ is defined as

$$\Phi^t := \sum_{i=1}^n e^{\alpha \cdot (x_i^t - t/n)} = \sum_{i=1}^n e^{\alpha \cdot y_i^t}.$$

Our goal is to show that $\Phi^t = \mathcal{O}(\text{poly}(n))$, as it implies that

$$y_i^t \le \frac{1}{\alpha} \cdot \log \Phi^t = \mathcal{O}\left(\frac{1}{\alpha} \cdot \log n\right).$$

Even better, if we show that $\Phi^t = \mathcal{O}(n)$, then the bins with load $\geq t/n + z$ is at most $\mathcal{O}(n \cdot e^{-\alpha z})$.

Question: How can we prove this?

■ A drift theorem (or drift inequality) has the form

■ A drift theorem (or drift inequality) has the form

$$\mathbf{E}\left[\Phi^{t+1} \mid \mathfrak{F}^t\right] \leq \Phi^t \cdot \left(1 - \frac{c_1}{n}\right) + c_2,$$

A drift theorem (or drift inequality) has the form

$$\mathbf{E}\left[\Phi^{t+1} \mid \mathfrak{F}^t\right] \leq \Phi^t \cdot \left(1 - \frac{c_1}{n}\right) + c_2,$$

for any possible history of the process \mathfrak{F}^t (or filtration).

■ A drift theorem (or drift inequality) has the form

$$\mathbf{E}\left[\Phi^{t+1} \mid \mathfrak{F}^t\right] \le \Phi^t \cdot \left(1 - \frac{c_1}{n}\right) + c_2,$$

for any possible history of the process \mathfrak{F}^t (or filtration).

■ When Φ^t is large (i.e., $\Phi^t \geq 2 \cdot \frac{c_2}{c_1} \cdot n$), then it drops in expectation by a multiplicative factor,

■ A drift theorem (or drift inequality) has the form

$$\mathbf{E}\left[\Phi^{t+1} \mid \mathfrak{F}^t\right] \le \Phi^t \cdot \left(1 - \frac{c_1}{n}\right) + c_2,$$

for any possible history of the process \mathfrak{F}^t (or filtration).

■ When Φ^t is large (i.e., $\Phi^t \geq 2 \cdot \frac{c_2}{c_1} \cdot n$), then it drops in expectation by a multiplicative factor,

$$\mathbf{E}\left[\Phi^{t+1} \mid \mathfrak{F}^t, \Phi^t \ge 2 \cdot \frac{c_2}{c_1} \cdot n \right] \le \Phi^t \cdot \left(1 - \frac{c_1}{2n}\right),$$

■ A drift theorem (or drift inequality) has the form

$$\mathbf{E}\left[\Phi^{t+1} \mid \mathfrak{F}^t\right] \le \Phi^t \cdot \left(1 - \frac{c_1}{n}\right) + c_2,$$

for any possible history of the process \mathfrak{F}^t (or filtration).

When Φ^t is large (i.e., $\Phi^t \geq 2 \cdot \frac{c_2}{c_1} \cdot n$), then it drops in expectation by a multiplicative factor,

$$\mathbf{E}\left[\Phi^{t+1} \mid \mathfrak{F}^t, \Phi^t \geq 2 \cdot \frac{c_2}{c_1} \cdot n \right] \leq \Phi^t \cdot \left(1 - \frac{c_1}{2n}\right),$$

■ It also implies that $\mathbf{E}\left[\Phi^{t}\right] \leq \frac{c_{2}}{c_{1}} \cdot n$,

■ A drift theorem (or drift inequality) has the form

$$\mathbf{E}\left[\Phi^{t+1} \mid \mathfrak{F}^t\right] \le \Phi^t \cdot \left(1 - \frac{c_1}{n}\right) + c_2,$$

for any possible history of the process \mathfrak{F}^t (or filtration).

When Φ^t is large (i.e., $\Phi^t \geq 2 \cdot \frac{c_2}{c_1} \cdot n$), then it drops in expectation by a multiplicative factor,

$$\mathbf{E}\left[\Phi^{t+1} \mid \mathfrak{F}^t, \Phi^t \geq 2 \cdot \frac{c_2}{c_1} \cdot n \right] \leq \Phi^t \cdot \left(1 - \frac{c_1}{2n}\right),$$

■ It also implies that $\mathbf{E}[\Phi^t] \leq \frac{c_2}{c_1} \cdot n$, since by induction,

$$\mathbf{E}\left[\Phi^{t+1}\right] = \mathbf{E}\left[\mathbf{E}\left[\Phi^{t+1} \mid \Phi^{t}\right]\right]$$

■ A drift theorem (or drift inequality) has the form

$$\mathbf{E}\left[\Phi^{t+1} \mid \mathfrak{F}^t\right] \le \Phi^t \cdot \left(1 - \frac{c_1}{n}\right) + c_2,$$

for any possible history of the process \mathfrak{F}^t (or filtration).

■ When Φ^t is large (i.e., $\Phi^t \geq 2 \cdot \frac{c_2}{c_1} \cdot n$), then it drops in expectation by a multiplicative factor,

$$\mathbf{E}\left[\Phi^{t+1} \mid \mathfrak{F}^t, \Phi^t \geq 2 \cdot \frac{c_2}{c_1} \cdot n \right] \leq \Phi^t \cdot \left(1 - \frac{c_1}{2n}\right),$$

■ It also implies that $\mathbf{E}[\Phi^t] \leq \frac{c_2}{c_1} \cdot n$, since by induction,

$$\mathbf{E}\left[\Phi^{t+1}\right] = \mathbf{E}\left[\mathbf{E}\left[\Phi^{t+1} \mid \Phi^{t}\right]\right] \leq \mathbf{E}\left[\Phi^{t}\right] \cdot \left(1 - \frac{c_{1}}{n}\right) + c_{2}$$

■ A drift theorem (or drift inequality) has the form

$$\mathbf{E}\left[\Phi^{t+1} \mid \mathfrak{F}^t\right] \le \Phi^t \cdot \left(1 - \frac{c_1}{n}\right) + c_2,$$

for any possible history of the process \mathfrak{F}^t (or filtration).

When Φ^t is large (i.e., $\Phi^t \geq 2 \cdot \frac{c_2}{c_1} \cdot n$), then it drops in expectation by a multiplicative factor,

$$\mathbf{E}\left[\Phi^{t+1} \mid \mathfrak{F}^t, \Phi^t \geq 2 \cdot \frac{c_2}{c_1} \cdot n \right] \leq \Phi^t \cdot \left(1 - \frac{c_1}{2n}\right),$$

■ It also implies that $\mathbf{E}[\Phi^t] \leq \frac{c_2}{c_1} \cdot n$, since by induction,

$$\mathbf{E}\left[\Phi^{t+1}\right] = \mathbf{E}\left[\mathbf{E}\left[\Phi^{t+1} \mid \Phi^{t}\right]\right] \le \mathbf{E}\left[\Phi^{t}\right] \cdot \left(1 - \frac{c_{1}}{n}\right) + c_{2} \le \frac{c_{2}}{c_{1}} \cdot n - c_{2} + c_{2}$$

■ A drift theorem (or drift inequality) has the form

$$\mathbf{E}\left[\Phi^{t+1} \mid \mathfrak{F}^t\right] \le \Phi^t \cdot \left(1 - \frac{c_1}{n}\right) + c_2,$$

for any possible history of the process \mathfrak{F}^t (or filtration).

When Φ^t is large (i.e., $\Phi^t \geq 2 \cdot \frac{c_2}{c_1} \cdot n$), then it drops in expectation by a multiplicative factor,

$$\mathbf{E}\left[\Phi^{t+1} \mid \mathfrak{F}^t, \Phi^t \geq 2 \cdot \frac{c_2}{c_1} \cdot n \right] \leq \Phi^t \cdot \left(1 - \frac{c_1}{2n}\right),$$

■ It also implies that $\mathbf{E}[\Phi^t] \leq \frac{c_2}{c_1} \cdot n$, since by induction,

$$\mathbf{E}\left[\Phi^{t+1}\right] = \mathbf{E}\left[\mathbf{E}\left[\Phi^{t+1} \mid \Phi^{t}\right]\right] \leq \mathbf{E}\left[\Phi^{t}\right] \cdot \left(1 - \frac{c_{1}}{n}\right) + c_{2} \leq \frac{c_{2}}{c_{1}} \cdot n - c_{2} + c_{2} = \frac{c_{2}}{c_{1}} \cdot n.$$

■ A drift theorem (or drift inequality) has the form

$$\mathbf{E}\left[\Phi^{t+1} \mid \mathfrak{F}^t\right] \le \Phi^t \cdot \left(1 - \frac{c_1}{n}\right) + c_2,$$

for any possible history of the process \mathfrak{F}^t (or filtration).

When Φ^t is large (i.e., $\Phi^t \geq 2 \cdot \frac{c_2}{c_1} \cdot n$), then it drops in expectation by a multiplicative factor,

$$\mathbf{E}\left[\Phi^{t+1} \mid \mathfrak{F}^t, \Phi^t \geq 2 \cdot \frac{c_2}{c_1} \cdot n \right] \leq \Phi^t \cdot \left(1 - \frac{c_1}{2n}\right),$$

■ It also implies that $\mathbf{E}[\Phi^t] \leq \frac{c_2}{c_1} \cdot n$, since by induction,

$$\mathbf{E}\left[\Phi^{t+1}\right] = \mathbf{E}\left[\mathbf{E}\left[\Phi^{t+1} \mid \Phi^{t}\right]\right] \leq \mathbf{E}\left[\Phi^{t}\right] \cdot \left(1 - \frac{c_{1}}{n}\right) + c_{2} \leq \frac{c_{2}}{c_{1}} \cdot n - c_{2} + c_{2} = \frac{c_{2}}{c_{1}} \cdot n.$$

■ Then, applying Markov's inequality we get that w.h.p. $\mathbf{E} [\Phi^t] = \text{poly}(n)$.

 \blacksquare Let us fix a bin $i \in [n]$.

■ Let us fix a bin $i \in [n]$. Then,

$$\mathbf{E} \left[\Phi_i^{t+1} \mid \Phi_i^t \right] = p_i \cdot e^{\alpha \cdot (y_i^t + 1 - 1/n)} + (1 - p_i) \cdot e^{\alpha \cdot (y_i^t - 1/n)}$$

Let us fix a bin $i \in [n]$. Then,

$$\mathbf{E} \left[\Phi_i^{t+1} \mid \Phi_i^t \right] = p_i \cdot e^{\alpha \cdot (y_i^t + 1 - 1/n)} + (1 - p_i) \cdot e^{\alpha \cdot (y_i^t - 1/n)}$$

$$\leq \Phi_i^t \cdot \left(p_i \cdot \left(1 + \alpha \cdot (1 - 1/n) + \alpha^2 \right) + (1 - p_i) \cdot \left(1 - \alpha/n + \alpha^2/n^2 \right) \right)$$

using the Taylor estimate $e^z \le 1 + z + z^2$ for sufficiently small z.

Let us fix a bin $i \in [n]$. Then,

$$\mathbf{E} \left[\Phi_i^{t+1} \mid \Phi_i^t \right] = p_i \cdot e^{\alpha \cdot (y_i^t + 1 - 1/n)} + (1 - p_i) \cdot e^{\alpha \cdot (y_i^t - 1/n)}$$

$$\leq \Phi_i^t \cdot \left(p_i \cdot \left(1 + \alpha \cdot (1 - 1/n) + \alpha^2 \right) + (1 - p_i) \cdot \left(1 - \alpha/n + \alpha^2/n^2 \right) \right)$$

using the Taylor estimate $e^z \leq 1 + z + z^2$ for sufficiently small z.

Applies also to weights \mathcal{W} with unit expectation and finite MGF, i.e., $e^{\alpha \mathcal{W}} \leq 1 + \alpha + \alpha^2 \cdot S$.

Let us fix a bin $i \in [n]$. Then,

$$\mathbf{E} \left[\Phi_i^{t+1} \mid \Phi_i^t \right] = p_i \cdot e^{\alpha \cdot (y_i^t + 1 - 1/n)} + (1 - p_i) \cdot e^{\alpha \cdot (y_i^t - 1/n)}$$

$$\leq \Phi_i^t \cdot \left(p_i \cdot \left(1 + \alpha \cdot (1 - 1/n) + \alpha^2 \right) + (1 - p_i) \cdot \left(1 - \alpha/n + \alpha^2/n^2 \right) \right)$$

$$= \Phi_i^t \cdot \left(1 + \alpha \cdot \left(p_i - \frac{1}{n} \right) + \alpha^2 \cdot \left(p_i + \frac{1 - p_i}{n} \right) \right),$$

using the Taylor estimate $e^z \le 1 + z + z^2$ for sufficiently small z.

Let us fix a bin $i \in [n]$. Then,

$$\mathbf{E}\left[\Phi_{i}^{t+1} \mid \Phi_{i}^{t}\right] = p_{i} \cdot e^{\alpha \cdot (y_{i}^{t}+1-1/n)} + (1-p_{i}) \cdot e^{\alpha \cdot (y_{i}^{t}-1/n)}$$

$$\leq \Phi_{i}^{t} \cdot \left(p_{i} \cdot \left(1+\alpha \cdot (1-1/n)+\alpha^{2}\right) + (1-p_{i}) \cdot \left(1-\alpha/n+\alpha^{2}/n^{2}\right)\right)$$

$$= \Phi_{i}^{t} \cdot \left(1+\alpha \cdot \left(p_{i}-\frac{1}{n}\right) + \alpha^{2} \cdot \left(p_{i}+\frac{1-p_{i}}{n}\right)\right),$$

using the Taylor estimate $e^z \le 1 + z + z^2$ for sufficiently small z.

For bins with $p_i = \frac{1-\epsilon}{n}$,

Let us fix a bin $i \in [n]$. Then,

$$\mathbf{E} \left[\Phi_i^{t+1} \mid \Phi_i^t \right] = p_i \cdot e^{\alpha \cdot (y_i^t + 1 - 1/n)} + (1 - p_i) \cdot e^{\alpha \cdot (y_i^t - 1/n)}$$

$$\leq \Phi_i^t \cdot \left(p_i \cdot \left(1 + \alpha \cdot (1 - 1/n) + \alpha^2 \right) + (1 - p_i) \cdot \left(1 - \alpha/n + \alpha^2/n^2 \right) \right)$$

$$= \Phi_i^t \cdot \left(1 + \alpha \cdot \left(p_i - \frac{1}{n} \right) + \alpha^2 \cdot \left(p_i + \frac{1 - p_i}{n} \right) \right),$$

using the Taylor estimate $e^z \leq 1 + z + z^2$ for sufficiently small z.

■ For bins with $p_i = \frac{1-\epsilon}{n}$, we have that $(\Delta \Phi_i^{t+1} := \Phi_i^{t+1} - \Phi_i^t)$

$$\mathbf{E}\left[\Delta\Phi_{i}^{t+1}\mid\mathfrak{F}^{t}\right]\leq\Phi_{i}^{t}\cdot\left(-\frac{\alpha\epsilon}{n}+\mathcal{O}\left(\frac{\alpha^{2}}{n}\right)\right)$$

Let us fix a bin $i \in [n]$. Then,

$$\mathbf{E} \left[\Phi_i^{t+1} \mid \Phi_i^t \right] = p_i \cdot e^{\alpha \cdot (y_i^t + 1 - 1/n)} + (1 - p_i) \cdot e^{\alpha \cdot (y_i^t - 1/n)}$$

$$\leq \Phi_i^t \cdot \left(p_i \cdot \left(1 + \alpha \cdot (1 - 1/n) + \alpha^2 \right) + (1 - p_i) \cdot \left(1 - \alpha/n + \alpha^2/n^2 \right) \right)$$

$$= \Phi_i^t \cdot \left(1 + \alpha \cdot \left(p_i - \frac{1}{n} \right) + \alpha^2 \cdot \left(p_i + \frac{1 - p_i}{n} \right) \right),$$

using the Taylor estimate $e^z \leq 1 + z + z^2$ for sufficiently small z.

■ For bins with $p_i = \frac{1-\epsilon}{n}$, we have that $(\Delta \Phi_i^{t+1} := \Phi_i^{t+1} - \Phi_i^t)$

$$\mathbf{E}\left[\Delta\Phi_i^{t+1} \mid \mathfrak{F}^t\right] \leq \Phi_i^t \cdot \left(-\frac{\alpha\epsilon}{n} + \mathcal{O}\left(\frac{\alpha^2}{n}\right)\right) \rightsquigarrow \text{Good bin.}$$

Let us fix a bin $i \in [n]$. Then,

$$\mathbf{E} \left[\Phi_i^{t+1} \mid \Phi_i^t \right] = p_i \cdot e^{\alpha \cdot (y_i^t + 1 - 1/n)} + (1 - p_i) \cdot e^{\alpha \cdot (y_i^t - 1/n)}$$

$$\leq \Phi_i^t \cdot \left(p_i \cdot \left(1 + \alpha \cdot (1 - 1/n) + \alpha^2 \right) + (1 - p_i) \cdot \left(1 - \alpha/n + \alpha^2/n^2 \right) \right)$$

$$= \Phi_i^t \cdot \left(1 + \alpha \cdot \left(p_i - \frac{1}{n} \right) + \alpha^2 \cdot \left(p_i + \frac{1 - p_i}{n} \right) \right),$$

using the Taylor estimate $e^z \leq 1 + z + z^2$ for sufficiently small z.

■ For bins with $p_i = \frac{1-\epsilon}{n}$, we have that $(\Delta \Phi_i^{t+1} := \Phi_i^{t+1} - \Phi_i^t)$

$$\mathbf{E}\left[\Delta\Phi_i^{t+1} \mid \mathfrak{F}^t\right] \leq \Phi_i^t \cdot \left(-\frac{\alpha\epsilon}{n} + \mathcal{O}\left(\frac{\alpha^2}{n}\right)\right) \rightsquigarrow \text{Good bin.}$$

Otherwise, for bins with $p_i = \frac{1+\widetilde{\epsilon}}{n}$,

Let us fix a bin $i \in [n]$. Then,

$$\mathbf{E}\left[\Phi_{i}^{t+1} \mid \Phi_{i}^{t}\right] = p_{i} \cdot e^{\alpha \cdot (y_{i}^{t}+1-1/n)} + (1-p_{i}) \cdot e^{\alpha \cdot (y_{i}^{t}-1/n)}$$

$$\leq \Phi_{i}^{t} \cdot \left(p_{i} \cdot \left(1+\alpha \cdot (1-1/n)+\alpha^{2}\right) + (1-p_{i}) \cdot \left(1-\alpha/n+\alpha^{2}/n^{2}\right)\right)$$

$$= \Phi_{i}^{t} \cdot \left(1+\alpha \cdot \left(p_{i}-\frac{1}{n}\right)+\alpha^{2} \cdot \left(p_{i}+\frac{1-p_{i}}{n}\right)\right),$$

using the Taylor estimate $e^z \leq 1 + z + z^2$ for sufficiently small z.

■ For bins with $p_i = \frac{1-\epsilon}{n}$, we have that $(\Delta \Phi_i^{t+1} := \Phi_i^{t+1} - \Phi_i^t)$

$$\mathbf{E}\left[\Delta\Phi_i^{t+1} \mid \mathfrak{F}^t\right] \leq \Phi_i^t \cdot \left(-\frac{\alpha\epsilon}{n} + \mathcal{O}\left(\frac{\alpha^2}{n}\right)\right) \rightsquigarrow \text{Good bin.}$$

Otherwise, for bins with $p_i = \frac{1+\widetilde{\epsilon}}{n}$,

$$\mathbf{E}\left[\Delta\Phi_i^{t+1} \mid \mathfrak{F}^t\right] \leq \Phi_i^t \cdot \left(+ \frac{\alpha\widetilde{\epsilon}}{n} + \mathcal{O}\left(\frac{\alpha^2}{n}\right) \right)$$

Let us fix a bin $i \in [n]$. Then,

$$\mathbf{E} \left[\Phi_i^{t+1} \mid \Phi_i^t \right] = p_i \cdot e^{\alpha \cdot (y_i^t + 1 - 1/n)} + (1 - p_i) \cdot e^{\alpha \cdot (y_i^t - 1/n)}$$

$$\leq \Phi_i^t \cdot \left(p_i \cdot \left(1 + \alpha \cdot (1 - 1/n) + \alpha^2 \right) + (1 - p_i) \cdot \left(1 - \alpha/n + \alpha^2/n^2 \right) \right)$$

$$= \Phi_i^t \cdot \left(1 + \alpha \cdot \left(p_i - \frac{1}{n} \right) + \alpha^2 \cdot \left(p_i + \frac{1 - p_i}{n} \right) \right),$$

using the Taylor estimate $e^z \leq 1 + z + z^2$ for sufficiently small z.

■ For bins with $p_i = \frac{1-\epsilon}{n}$, we have that $(\Delta \Phi_i^{t+1} := \Phi_i^{t+1} - \Phi_i^t)$

$$\mathbf{E}\left[\Delta\Phi_i^{t+1} \mid \mathfrak{F}^t\right] \leq \Phi_i^t \cdot \left(-\frac{\alpha\epsilon}{n} + \mathcal{O}\left(\frac{\alpha^2}{n}\right)\right) \rightsquigarrow \text{Good bin.}$$

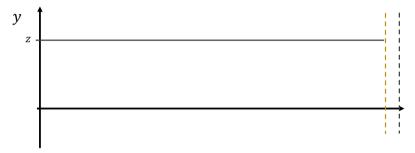
Otherwise, for bins with $p_i = \frac{1+\widetilde{\epsilon}}{n}$,

$$\mathbf{E}\left[\left.\Delta\Phi_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]\leq\Phi_{i}^{t}\cdot\left(+\frac{\alpha\widetilde{\epsilon}}{n}+\mathcal{O}\left(\frac{\alpha^{2}}{n}\right)\right)\rightsquigarrow\operatorname{\mathsf{Bad\ bin}}.$$

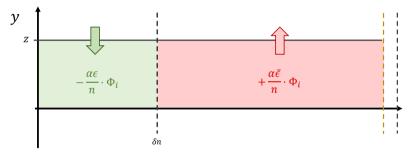
■ There could be too many overloaded bins.

- There could be too many overloaded bins.
- Consider a step $t \geq 0$, where all but one bins have the same load:

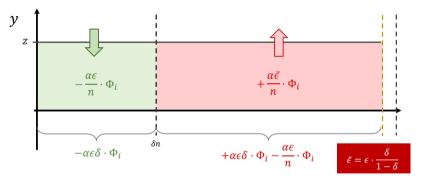
- There could be too many overloaded bins.
- Consider a step $t \geq 0$, where all but one bins have the same load:



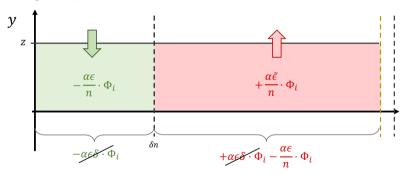
- There could be too many overloaded bins.
- Consider a step $t \geq 0$, where all but one bins have the same load:



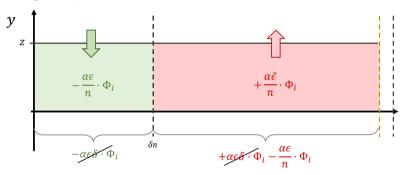
- There could be too many overloaded bins.
- Consider a step $t \geq 0$, where all but one bins have the same load:



- There could be too many overloaded bins.
- Consider a step $t \geq 0$, where all but one bins have the same load:

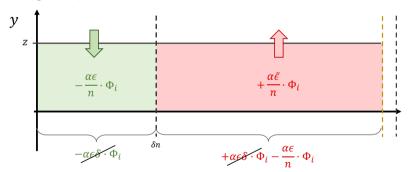


- There could be too many overloaded bins.
- Consider a step $t \geq 0$, where all but one bins have the same load:



We could get only a *very small* decrease.

- There could be too many overloaded bins.
- Consider a step $t \geq 0$, where all but one bins have the same load:



We could get only a very small decrease. \rightsquigarrow Gives $\mathcal{O}(n \log n/\epsilon)$ bound on the gap.

The *hyperbolic cosine potential* [PTW15, Spe77] is defined as

$$\Gamma^t := \Phi^t + \Psi^t := \sum_{i=1}^n e^{\alpha \cdot (x_i^t - t/n)} + e^{-\alpha \cdot (x_i^t - t/n)}$$

The *hyperbolic cosine potential* [PTW15, Spe77] is defined as

$$\Gamma^t := \Phi^t + \Psi^t := \sum_{i=1}^n e^{\alpha \cdot (x_i^t - t/n)} + e^{-\alpha \cdot (x_i^t - t/n)} = \sum_{i=1}^n e^{\alpha \cdot y_i^t} + e^{-\alpha \cdot y_i^t}.$$

The *hyperbolic cosine potential* [PTW15, Spe77] is defined as

$$\Gamma^t := \Phi^t + \Psi^t := \sum_{i=1}^n e^{\alpha \cdot (x_i^t - t/n)} + e^{-\alpha \cdot (x_i^t - t/n)} = \sum_{i=1}^n e^{\alpha \cdot y_i^t} + e^{-\alpha \cdot y_i^t}.$$

Question: Why does the *second term* help?

The *hyperbolic cosine potential* [PTW15, Spe77] is defined as

$$\Gamma^t := \Phi^t + \Psi^t := \sum_{i=1}^n e^{\alpha \cdot (x_i^t - t/n)} + e^{-\alpha \cdot (x_i^t - t/n)} = \sum_{i=1}^n e^{\alpha \cdot y_i^t} + e^{-\alpha \cdot y_i^t}.$$

Question: Why does the *second term* help?



■ For any bin $i \in [n]$, the underload potential satisfies the following drift inequality,

 \blacksquare For any bin $i \in [n]$, the underload potential satisfies the following drift inequality,

$$\mathbf{E}\left[\Psi_i^{t+1} \mid \mathfrak{F}^t\right] \leq \Psi_i^t \cdot \left(1 + \alpha \cdot \left(\frac{1}{n} - p_i\right) + \alpha^2 \cdot \left(p_i + (1 - p_i) \cdot \frac{1}{n^2}\right)\right).$$

For any bin $i \in [n]$, the underload potential satisfies the following drift inequality,

$$\mathbf{E}\left[\Psi_i^{t+1} \mid \mathfrak{F}^t\right] \leq \Psi_i^t \cdot \left(1 + \alpha \cdot \left(\frac{1}{n} - p_i\right) + \alpha^2 \cdot \left(p_i + (1 - p_i) \cdot \frac{1}{n^2}\right)\right).$$

So, the dominant term is the *negative* of that in Φ .

 \blacksquare For any bin $i \in [n]$, the underload potential satisfies the following drift inequality,

$$\mathbf{E}\left[\Psi_i^{t+1} \mid \mathfrak{F}^t\right] \leq \Psi_i^t \cdot \left(1 + \alpha \cdot \left(\frac{1}{n} - p_i\right) + \alpha^2 \cdot \left(p_i + (1 - p_i) \cdot \frac{1}{n^2}\right)\right).$$

- So, the dominant term is the *negative* of that in Φ .
- More specifically, if $p_i = \frac{1+\tilde{\epsilon}}{n}$, then

$$\mathbf{E}\left[\Delta\Psi_{i}^{t+1} \mid \mathfrak{F}^{t}\right] \leq \Psi_{i}^{t} \cdot \left(-\frac{\alpha \widetilde{\epsilon}}{n} + \mathcal{O}\left(\frac{\alpha^{2}}{n}\right)\right) \rightsquigarrow \text{Good bin.}$$

 \blacksquare For any bin $i \in [n]$, the underload potential satisfies the following drift inequality,

$$\mathbf{E}\left[\Psi_i^{t+1} \mid \mathfrak{F}^t\right] \leq \Psi_i^t \cdot \left(1 + \alpha \cdot \left(\frac{1}{n} - p_i\right) + \alpha^2 \cdot \left(p_i + (1 - p_i) \cdot \frac{1}{n^2}\right)\right).$$

- So, the dominant term is the *negative* of that in Φ .
- More specifically, if $p_i = \frac{1+\tilde{\epsilon}}{n}$, then

$$\mathbf{E}\left[\Delta\Psi_i^{t+1} \mid \mathfrak{F}^t\right] \leq \Psi_i^t \cdot \left(-\frac{\alpha\widetilde{\epsilon}}{n} + \mathcal{O}\left(\frac{\alpha^2}{n}\right)\right) \rightsquigarrow \text{Good bin.}$$

Otherwise, if $p_i = \frac{1-\epsilon}{n}$

$$\mathbf{E}\left[\left.\Delta\Psi_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]\leq\Psi_{i}^{t}\cdot\left(+\frac{\alpha\epsilon}{n}+\mathcal{O}\left(\frac{\alpha^{2}}{n}\right)\right)\rightsquigarrow\mathbf{Bad\ bin}.$$

Drift Theorem

Theorem ([PTW15, Section 2])

Consider any process with non-decreasing allocation vector p which is ϵ -biased for some $\epsilon \in (0,1)$ and some constant δ , in the setting with weights sampled from a distribution with finite MGF. Then, for $\Gamma := \Gamma(\alpha)$ with $\alpha := \Theta(\epsilon)$, for any step $t \geq 0$,

$$\mathbf{E}\left[\right. \Delta\Gamma^{t+1} \left. \right| \, \mathfrak{F}^t \left. \right] \leq -\Gamma^t \cdot \frac{\alpha \epsilon}{4n} + \mathrm{poly}(1/\epsilon),$$

and

$$\mathbf{E}\left[\Gamma^{t}\right] \leq n \cdot \operatorname{poly}(1/\epsilon).$$

Refined Drift Theorem

Theorem ([LS22, Corollary 3.2])

Consider any process and a probability vector p being ϵ -biased for some $\epsilon \in (0,1)$ and some constant δ . Further assume that it satisfies for some K > 0 and for any $t \ge 0$,

$$\mathbf{E}\left[\Phi^{t+1} \mid \mathfrak{F}^t\right] \leq \sum_{i=1}^n \Phi_i^t \cdot \left(1 + \left(p_i - \frac{1}{n}\right) \cdot \alpha + K \cdot \frac{\alpha^2}{n}\right),$$

and

$$\mathbf{E}\left[\left|\Psi^{t+1}\right|\left|\mathfrak{F}^{t}\right]\right] \leq \sum_{i=1}^{n} \Psi_{i}^{t} \cdot \left(1 + \left(\frac{1}{n} - p_{i}\right) \cdot \alpha + K \cdot \frac{\alpha^{2}}{n}\right).$$

Then, there exists a constant $c := c(\delta) > 0$, such that for $\alpha \in (0, \min\{1, \frac{\epsilon \delta}{8K}\})$

$$\mathbf{E}\left[\left.\Gamma^{t+1} \mid \mathfrak{F}^t\right.\right] \le \Gamma^t \cdot \left(1 - \frac{\alpha \epsilon \delta}{8n}\right) + c\alpha \epsilon,$$

and

$$\mathbf{E}\left[\Gamma^{t}\right] \leq \frac{8c}{\delta} \cdot n.$$

Refined Drift Theorem

Theorem ([LS22, Corollary 3.2])

Consider any process and a probability vector p being ϵ -biased for some $\epsilon \in (0,1)$ and some constant δ . Further assume that it satisfies for some K > 0 and some R > 0, for any $t \ge 0$,

$$\mathbf{E}\left[\left.\Phi^{t+1}\mid\mathfrak{F}^{t}\right.\right] \leq \sum_{i=1}^{n}\Phi_{i}^{t}\cdot\left(1+\left(p_{i}-\frac{1}{n}\right)\cdot\mathbf{R}\cdot\alpha+K\cdot\mathbf{R}\cdot\frac{\alpha^{2}}{n}\right),$$

and

$$\mathbf{E}\left[\left|\Psi^{t+1}\right|\left|\mathfrak{F}^{t}\right.\right] \leq \sum_{i=1}^{n} \Psi_{i}^{t} \cdot \left(1 + \left(\frac{1}{n} - p_{i}\right) \cdot \mathbf{R} \cdot \alpha + K \cdot \mathbf{R} \cdot \frac{\alpha^{2}}{n}\right).$$

Then, there exists a constant $c := c(\delta) > 0$, such that for $\alpha \in (0, \min\{1, \frac{\epsilon \delta}{8K}\})$

$$\mathbf{E}\left[\left.\Gamma^{t+1}\,\right|\,\mathfrak{F}^t\,\right] \leq \Gamma^t\cdot \frac{\mathbf{R}}{\mathbf{R}}\cdot \left(1-\frac{\alpha\epsilon\delta}{8n}\right) + \frac{\mathbf{R}}{\mathbf{R}}\cdot c\alpha\epsilon,$$

and

$$\mathbf{E}\left[\Gamma^{t}\right] \leq \frac{8c}{\delta} \cdot n.$$

Our goal is to show that:

$$\sum_{i=1}^{n} \Phi_{i}^{t} \cdot \left(\alpha \cdot \left(p_{i} - \frac{1}{n}\right) + K \cdot \frac{\alpha^{2}}{n}\right) + \Psi_{i}^{t} \cdot \left(\alpha \cdot \left(\frac{1}{n} - p_{i}\right) + K \cdot \frac{\alpha^{2}}{n}\right)$$

Our goal is to show that:

$$\sum_{i=1}^{n} \Phi_{i}^{t} \cdot \left(\alpha \cdot \left(p_{i} - \frac{1}{n}\right) + K \cdot \frac{\alpha^{2}}{n}\right) + \Psi_{i}^{t} \cdot \left(\alpha \cdot \left(\frac{1}{n} - p_{i}\right) + K \cdot \frac{\alpha^{2}}{n}\right)$$

$$\leq -\frac{\alpha\epsilon}{4n} \cdot \Gamma^{t} + K \cdot \frac{\alpha^{2}}{n} \cdot \Gamma^{t} + \mathcal{O}(\alpha\epsilon)$$

Our goal is to show that:

$$\sum_{i=1}^{n} \Phi_{i}^{t} \cdot \left(\alpha \cdot \left(p_{i} - \frac{1}{n}\right) + K \cdot \frac{\alpha^{2}}{n}\right) + \Psi_{i}^{t} \cdot \left(\alpha \cdot \left(\frac{1}{n} - p_{i}\right) + K \cdot \frac{\alpha^{2}}{n}\right)$$

$$\leq -\frac{\alpha \epsilon}{4n} \cdot \Gamma^{t} + K \cdot \frac{\alpha^{2}}{n} \cdot \Gamma^{t} + \mathcal{O}(\alpha \epsilon)$$

$$\leq -\frac{\alpha \epsilon}{8n} \cdot \Gamma^{t} + \mathcal{O}(\alpha \epsilon),$$

for any $\alpha \leq \frac{\epsilon}{8K}$. (**Key Observation 2**)

Our goal is to show that:

$$\begin{split} & \sum_{i=1}^{n} \Phi_{i}^{t} \cdot \left(\alpha \cdot \left(p_{i} - \frac{1}{n}\right) + K \cdot \frac{\alpha^{2}}{n}\right) + \Psi_{i}^{t} \cdot \left(\alpha \cdot \left(\frac{1}{n} - p_{i}\right) + K \cdot \frac{\alpha^{2}}{n}\right) \\ & \leq -\frac{\alpha\epsilon}{4n} \cdot \Gamma^{t} + K \cdot \frac{\alpha^{2}}{n} \cdot \Gamma^{t} + \mathcal{O}(\alpha\epsilon) \\ & \leq -\frac{\alpha\epsilon}{8n} \cdot \Gamma^{t} + \mathcal{O}(\alpha\epsilon), \end{split}$$

for any $\alpha \leq \frac{\epsilon}{8K}$. (**Key Observation 2**)

We have the following types of bins

Set	Load	Index	r_i	Dominant Contribution
Good overloaded \mathcal{G}_+	$y_i \ge 0$	$i \le \delta n$	$\frac{1-\epsilon}{n_{\sim}}$	$-\Phi_i \cdot \frac{\alpha \epsilon}{n} + \Psi_i \cdot \frac{\alpha \epsilon}{n}$
Bad overloaded \mathcal{B}_+	$y_i \ge 0$	$i > \delta n$	$\frac{1+\epsilon}{n_{\sim}}$	$+\Phi_i \cdot \frac{\alpha\epsilon}{n} - \Psi_i \cdot \frac{\alpha\epsilon}{n}$
Good underloaded \mathcal{G}_{-}	$y_i < 0$	$i > \delta n$	$\frac{1+\epsilon}{n}$	$+\Phi_i \cdot \frac{\alpha\epsilon}{n} - \Psi_i \cdot \frac{\alpha\epsilon}{n}$
Bad overloaded \mathcal{B}_{-}	$y_i < 0$	$i \le \delta n$	$\frac{1-\epsilon}{n}$	$-\Phi_i \cdot \frac{\alpha \epsilon}{n} + \Psi_i \cdot \frac{\alpha \epsilon}{n}$

Our goal is to show that:

$$\begin{split} & \sum_{i=1}^{n} \Phi_{i}^{t} \cdot \left(\alpha \cdot \left(p_{i} - \frac{1}{n} \right) + K \cdot \frac{\alpha^{2}}{n} \right) + \Psi_{i}^{t} \cdot \left(\alpha \cdot \left(\frac{1}{n} - p_{i} \right) + K \cdot \frac{\alpha^{2}}{n} \right) \\ & \leq -\frac{\alpha \epsilon}{4n} \cdot \Gamma^{t} + K \cdot \frac{\alpha^{2}}{n} \cdot \Gamma^{t} + \mathcal{O}(\alpha \epsilon) \\ & \leq -\frac{\alpha \epsilon}{8n} \cdot \Gamma^{t} + \mathcal{O}(\alpha \epsilon), \end{split}$$

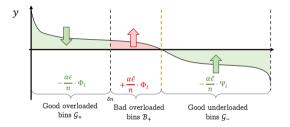
for any $\alpha \leq \frac{\epsilon}{8K}$. (**Key Observation 2**)

We have the following types of bins

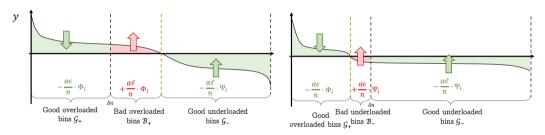
Set	Load	Index	r_i	Dominant Contribution
Good overloaded \mathcal{G}_+	$y_i \ge 0$	$i \le \delta n$	$\frac{1-\epsilon}{n}$	$-\Phi_i \cdot \frac{\alpha \epsilon}{n} + \Psi_i \cdot \frac{\alpha \epsilon}{n}$
Bad overloaded \mathcal{B}_+	$y_i \ge 0$	$i > \delta n$	$\frac{1+\epsilon}{n_{\sim}}$	$+\Phi_i\cdot\frac{\alpha\epsilon}{n}-\Psi_i\cdot\frac{\alpha\epsilon}{n}$
Good underloaded \mathcal{G}_{-}	$y_i < 0$	$i > \delta n$	$\frac{1+\epsilon}{n}$	$+\Phi_i\cdot \frac{lpha\epsilon}{n}-\Psi_i\cdot \frac{lpha\epsilon}{n}$
Bad overloaded \mathcal{B}_{-}	$y_i < 0$	$i \leq \delta n$	$\frac{1-\epsilon}{n}$	$-\Phi_i\cdot rac{lpha\epsilon}{n} + \Psi_i\cdot rac{lpha\epsilon}{n}$

Key Observation 3: For overloaded bins $\Psi_i^t \leq 1$ and for underloaded bins $\Phi_i^t \leq 1$, \rightarrow their contribution is $\mathcal{O}(\alpha \epsilon)$.

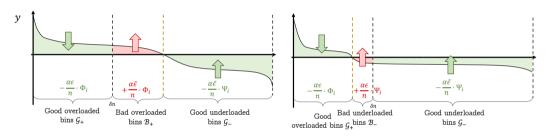
■ There can either be overloaded bad bins



■ There can either be overloaded bad bins or underloaded bad bins.

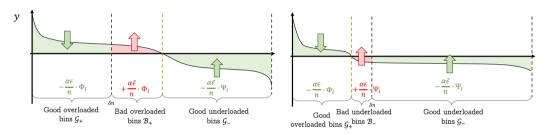


There can either be overloaded bad bins or underloaded bad bins.



Key Observation 4: The second case is symmetric to the first: $\delta' = 1 - \delta$, $\Phi' = \Psi$, $\Psi' = \Phi$ and y' = -y.

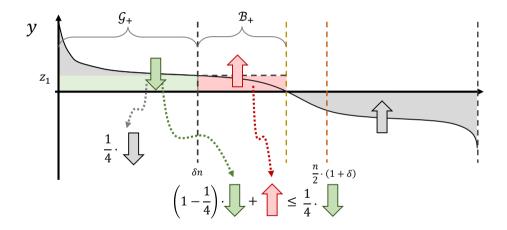
There can either be overloaded bad bins or underloaded bad bins.



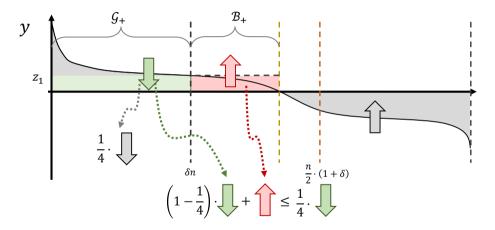
- **Key Observation 4:** The second case is symmetric to the first: $\delta' = 1 \delta$, $\Phi' = \Psi$, $\Psi' = \Phi$ and y' = -y.
- So we only consider Case A.

Case A.1: Not too many overloaded bins

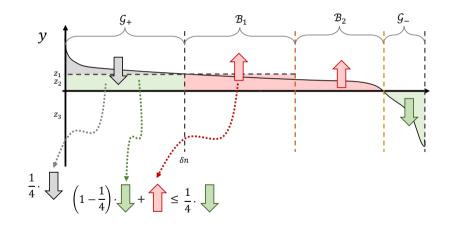
Case A.1: Not too many overloaded bins

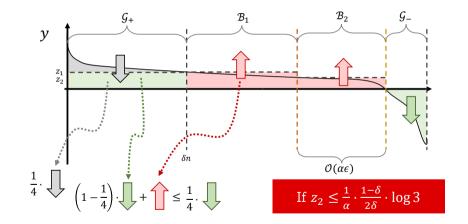


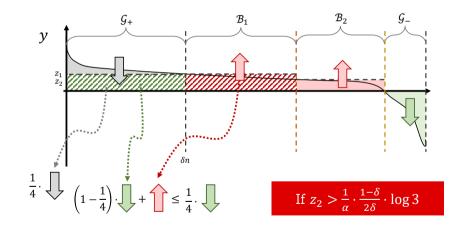
Case A.1: Not too many overloaded bins

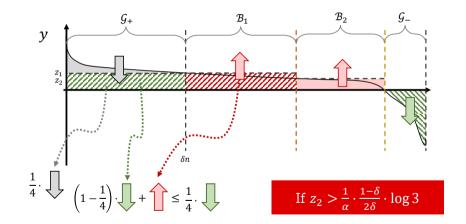


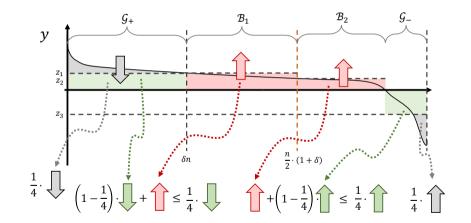
As with the exponential potential, we counteract the bad bins with a fraction of the decrease of the overloaded good bins. *All* underloaded bins are good.











We used the following techniques:

We used the following techniques:

1. Consider only step probability vectors (**Key observation 1**).

We used the following techniques:

- 1. Consider only step probability vectors (**Key observation 1**).
- 2. Consider only the coefficients of α (**Key observation 2**).

We used the following techniques:

- 1. Consider only step probability vectors (**Key observation 1**).
- 2. Consider only the coefficients of α (**Key observation 2**).
- 3. Consider only Φ_i for overloaded bins (and Ψ_i otherwise) (**Key observation 3**).

We used the following techniques:

- 1. Consider only step probability vectors (**Key observation 1**).
- 2. Consider only the coefficients of α (**Key observation 2**).
- 3. Consider only Φ_i for overloaded bins (and Ψ_i otherwise) (**Key observation 3**).
- 4. Consider only Case A by symmetry (**Key observation 4**).

We used the following techniques:

- 1. Consider only step probability vectors (**Key observation 1**).
- 2. Consider only the coefficients of α (**Key observation 2**).
- 3. Consider only Φ_i for overloaded bins (and Ψ_i otherwise) (**Key observation 3**).
- 4. Consider only Case A by symmetry (**Key observation 4**).
- 5. Use the decrease of the underload potential to counteract the increase of bad bins.

The drift theorem

Theorem ([LS22, Corollary 3.2])

Consider any allocation process and a probability vector p being ϵ -biased for some $\epsilon \in (0,1)$ and some constant δ . Further assume that it satisfies for some K>0 and some R>0, for any $t\geq 0$,

$$\mathbf{E}\left[\left.\Phi^{t+1}\mid\mathfrak{F}^{t}\right.\right] \leq \sum_{i=1}^{n}\Phi_{i}^{t}\cdot\left(1+\left(p_{i}-\frac{1}{n}\right)\cdot R\cdot\alpha+K\cdot R\cdot\frac{\alpha^{2}}{n}\right),$$

and

$$\mathbf{E}\left[\left|\Psi^{t+1}\right|\left|\mathfrak{F}^{t}\right]\right] \leq \sum_{i=1}^{n} \Psi_{i}^{t} \cdot \left(1 + \left(\frac{1}{n} - p_{i}\right) \cdot R \cdot \alpha + K \cdot R \cdot \frac{\alpha^{2}}{n}\right).$$

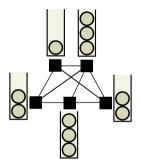
Then, there exists a constant $c := c(\delta) > 0$, such that for $\alpha \in (0, \min\{1, \frac{\epsilon \delta}{8K}\})$

$$\mathbf{E}\left[\left.\Gamma^{t+1} \mid \mathfrak{F}^t\right.\right] \leq \Gamma^t \cdot R \cdot \left(1 - \frac{\alpha \epsilon \delta}{8n}\right) + R \cdot c\alpha \epsilon,$$

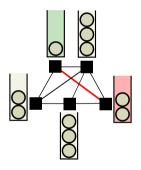
and

$$\mathbf{E}\left[\Gamma^{t}\right] \leq \frac{8c}{\delta} \cdot n.$$

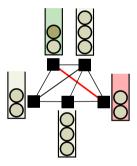
 \blacksquare Given a graph G = (V, E), where the vertices are bins. For each ball [KP06]:



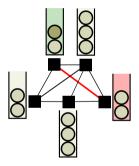
- \blacksquare Given a graph G = (V, E), where the vertices are bins. For each ball [KP06]:
 - ► Sample an edge u.a.r.



- \blacksquare Given a graph G = (V, E), where the vertices are bins. For each ball [KP06]:
 - ► Sample an edge u.a.r.
 - ightharpoonup Allocate the ball to the least loaded of its two adjacent bins.

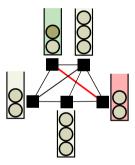


- \blacksquare Given a graph G = (V, E), where the vertices are bins. For each ball [KP06]:
 - ► Sample an edge u.a.r.
 - ightharpoonup Allocate the ball to the least loaded of its two adjacent bins.



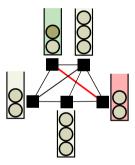
■ For any d-regular graph with conductance ϵ ,

- Given a graph G = (V, E), where the vertices are bins. For each ball [KP06]:
 - ► Sample an edge u.a.r.
 - ▶ Allocate the ball to the least loaded of its *two* adjacent bins.



For any d-regular graph with conductance ϵ , p^t is majorized by an ϵ -biased probability vector.

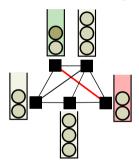
- Given a graph G = (V, E), where the vertices are bins. For each ball [KP06]:
 - ► Sample an edge u.a.r.
 - ▶ Allocate the ball to the least loaded of its *two* adjacent bins.



For any d-regular graph with conductance ϵ , p^t is majorized by an ϵ -biased probability vector. \rightsquigarrow gap is w.h.p. $\mathcal{O}(\frac{\log n}{\epsilon})$ [PTW15].

Example 1: The Graphical Setting

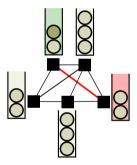
- Given a graph G = (V, E), where the vertices are bins. For each ball [KP06]:
 - ► Sample an edge u.a.r.
 - ▶ Allocate the ball to the least loaded of its *two* adjacent bins.



- For any d-regular graph with conductance ϵ , p^t is majorized by an ϵ -biased probability vector. \rightsquigarrow gap is w.h.p. $\mathcal{O}(\frac{\log n}{\epsilon})$ [PTW15].
- Majorization does not apply for weights.

Example 1: The Graphical Setting

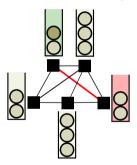
- Given a graph G = (V, E), where the vertices are bins. For each ball [KP06]:
 - ► Sample an edge u.a.r.
 - ▶ Allocate the ball to the least loaded of its *two* adjacent bins.



- For any d-regular graph with conductance ϵ , p^t is majorized by an ϵ -biased probability vector. \rightsquigarrow gap is w.h.p. $\mathcal{O}(\frac{\log n}{\epsilon})$ [PTW15].
- Majorization does not apply for weights. But the refined drift theorem applies for the majorized vector.

Example 1: The Graphical Setting

- Given a graph G = (V, E), where the vertices are bins. For each ball [KP06]:
 - ► Sample an edge u.a.r.
 - ▶ Allocate the ball to the least loaded of its *two* adjacent bins.



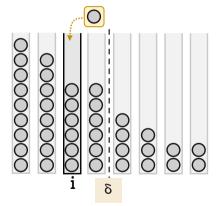
- For any d-regular graph with conductance ϵ , p^t is majorized by an ϵ -biased probability vector. \rightsquigarrow gap is w.h.p. $\mathcal{O}(\frac{\log n}{\epsilon})$ [PTW15].
- Majorization does not apply for weights. But the refined drift theorem applies for the majorized vector. \rightsquigarrow Resolves [PTW15, Open problem 1]

■ The TWINNING process at quantile δ , samples one bin i u.a.r. and:

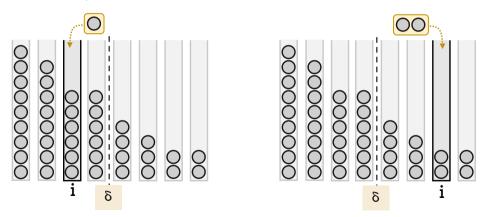
- The TWINNING process at quantile δ , samples one bin i u.a.r. and:
 - ▶ If i is among the heaviest $n \cdot \delta$ bins, it allocates one ball there.

- The Twinning process at quantile δ , samples one bin i u.a.r. and:
 - ▶ If i is among the heaviest $n \cdot \delta$ bins, it allocates one ball there.
 - ightharpoonup Otherwise, it allocates two balls.

- The TWINNING process at quantile δ , samples *one* bin i u.a.r. and:
 - ▶ If i is among the heaviest $n \cdot \delta$ bins, it allocates one ball there.
 - ▶ Otherwise, it allocates *two* balls.



- The TWINNING process at quantile δ , samples *one* bin i u.a.r. and:
 - ▶ If i is among the heaviest $n \cdot \delta$ bins, it allocates one ball there.
 - ▶ Otherwise, it allocates *two* balls.



For processes allocating more than one balls we have that:

$$\mathbf{E}\left[\left.\Delta\Phi_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]\leq\Phi_{i}^{t}\cdot\left(\alpha\cdot\mathbf{E}\left[\left.\Delta y_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]+\alpha^{2}\cdot\mathbf{E}\left[\left.(\Delta y_{i}^{t+1})^{2}\mid\mathfrak{F}^{t}\right.\right]\right),$$

For processes allocating more than one balls we have that:

$$\mathbf{E}\left[\left.\Delta\Phi_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]\leq\Phi_{i}^{t}\cdot\left(\alpha\cdot\mathbf{E}\left[\left.\Delta y_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]+\alpha^{2}\cdot\mathbf{E}\left[\left.(\Delta y_{i}^{t+1})^{2}\mid\mathfrak{F}^{t}\right.\right]\right),$$

$$\mathbf{E}\left[\left.\Delta\Psi_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]\leq\Psi_{i}^{t}\cdot\left(-\alpha\cdot\mathbf{E}\left[\left.\Delta y_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]+\alpha^{2}\cdot\mathbf{E}\left[\left.(\Delta y_{i}^{t+1})^{2}\mid\mathfrak{F}^{t}\right.\right]\right).$$

For processes allocating more than one balls we have that:

$$\mathbf{E}\left[\Delta\Phi_{i}^{t+1}\mid \mathfrak{F}^{t}\right] \leq \Phi_{i}^{t} \cdot \left(\alpha \cdot \mathbf{E}\left[\Delta y_{i}^{t+1}\mid \mathfrak{F}^{t}\right] + \alpha^{2} \cdot \mathbf{E}\left[(\Delta y_{i}^{t+1})^{2}\mid \mathfrak{F}^{t}\right]\right),$$

$$\mathbf{E}\left[\left.\Delta\Psi_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]\leq\Psi_{i}^{t}\cdot\left(-\alpha\cdot\mathbf{E}\left[\left.\Delta y_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]+\alpha^{2}\cdot\mathbf{E}\left[\left.(\Delta y_{i}^{t+1})^{2}\mid\mathfrak{F}^{t}\right.\right]\right).$$

For TWINNING, for any heavy bin $i \leq n \cdot \delta$:

$$\mathbf{E}\left[\right.\Delta y_i^{t+1}\left|\right.\mathfrak{F}^t\left.\right] = \underbrace{\left(1-\frac{1}{n}\right)\cdot\frac{1}{n}}_{\text{Allocate one ball to bin }i} + \underbrace{\left(-\frac{1}{n}\right)\cdot\left(\delta-\frac{1}{n}\right)}_{\text{Allocate one ball}} + \underbrace{\left(-\frac{2}{n}\right)\cdot(1-\delta)}_{\text{Allocate two balls to any other heavy bin}}$$

For processes allocating more than one balls we have that:

$$\mathbf{E} \left[\Delta \Phi_i^{t+1} \mid \mathfrak{F}^t \right] \le \Phi_i^t \cdot \left(\alpha \cdot \mathbf{E} \left[\Delta y_i^{t+1} \mid \mathfrak{F}^t \right] + \alpha^2 \cdot \mathbf{E} \left[(\Delta y_i^{t+1})^2 \mid \mathfrak{F}^t \right] \right),$$

$$\mathbf{E} \left[\Delta \Psi_i^{t+1} \mid \mathfrak{F}^t \right] \le \Psi_i^t \cdot \left(-\alpha \cdot \mathbf{E} \left[\Delta y_i^{t+1} \mid \mathfrak{F}^t \right] + \alpha^2 \cdot \mathbf{E} \left[(\Delta y_i^{t+1})^2 \mid \mathfrak{F}^t \right] \right).$$

For TWINNING, for any heavy bin $i \leq n \cdot \delta$:

$$\mathbf{E}\left[\right.\Delta y_i^{t+1}\left|\right.\mathfrak{F}^t\left.\right] = \underbrace{\left(1-\frac{1}{n}\right)\cdot\frac{1}{n}}_{\text{Allocate one ball to bin }i} + \underbrace{\left(-\frac{1}{n}\right)\cdot\left(\delta-\frac{1}{n}\right)}_{\text{Allocate one ball}} + \underbrace{\left(-\frac{2}{n}\right)\cdot(1-\delta)}_{\text{Allocate two balls to any other heavy bin}} = \frac{\delta}{n} - \frac{1}{n}.$$

For processes allocating more than one balls we have that:

$$\mathbf{E} \left[\Delta \Phi_i^{t+1} \mid \mathfrak{F}^t \right] \le \Phi_i^t \cdot \left(\alpha \cdot \mathbf{E} \left[\Delta y_i^{t+1} \mid \mathfrak{F}^t \right] + \alpha^2 \cdot \mathbf{E} \left[(\Delta y_i^{t+1})^2 \mid \mathfrak{F}^t \right] \right),$$

$$\mathbf{E} \left[\Delta \Psi_i^{t+1} \mid \mathfrak{F}^t \right] \le \Psi_i^t \cdot \left(-\alpha \cdot \mathbf{E} \left[\Delta y_i^{t+1} \mid \mathfrak{F}^t \right] + \alpha^2 \cdot \mathbf{E} \left[(\Delta y_i^{t+1})^2 \mid \mathfrak{F}^t \right] \right).$$

For TWINNING, for any heavy bin $i \leq n \cdot \delta$:

$$\mathbf{E}\left[\right.\Delta y_i^{t+1}\left|\right.\mathfrak{F}^t\left.\right] = \underbrace{\left(1-\frac{1}{n}\right)\cdot\frac{1}{n}}_{\text{Allocate one ball to bin }i} + \underbrace{\left(-\frac{1}{n}\right)\cdot\left(\delta-\frac{1}{n}\right)}_{\text{Allocate one ball}} + \underbrace{\left(-\frac{2}{n}\right)\cdot\left(1-\delta\right)}_{\text{Allocate two balls}} = \frac{\delta}{n} - \frac{1}{n}.$$

Similarly for a light bin $i > n \cdot \delta$:

$$\mathbf{E}\left[\left.\Delta y_i^{t+1}\,\right|\,\mathfrak{F}^t\,
ight]=rac{\delta}{n}.$$

For processes allocating more than one balls we have that:

$$\mathbf{E}\left[\left.\Delta\Phi_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]\leq\Phi_{i}^{t}\cdot\left(\alpha\cdot\mathbf{E}\left[\left.\Delta y_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]+\alpha^{2}\cdot\mathbf{E}\left[\left.(\Delta y_{i}^{t+1})^{2}\mid\mathfrak{F}^{t}\right.\right]\right),$$

$$\mathbf{E}\left[\left.\Delta\Psi_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]\leq\Psi_{i}^{t}\cdot\left(-\alpha\cdot\mathbf{E}\left[\left.\Delta y_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]+\alpha^{2}\cdot\mathbf{E}\left[\left.(\Delta y_{i}^{t+1})^{2}\mid\mathfrak{F}^{t}\right.\right]\right).$$

For TWINNING, for any heavy bin $i \leq n \cdot \delta$:

$$\mathbf{E}\left[\right.\Delta y_i^{t+1}\left|\right.\mathfrak{F}^t\left.\right] = \underbrace{\left(1-\frac{1}{n}\right)\cdot\frac{1}{n}}_{\text{Allocate one ball to bin }i} + \underbrace{\left(-\frac{1}{n}\right)\cdot\left(\delta-\frac{1}{n}\right)}_{\text{Allocate one ball}} + \underbrace{\left(-\frac{2}{n}\right)\cdot(1-\delta)}_{\text{Allocate two balls}} = \frac{\delta}{n} - \frac{1}{n}.$$

to any other heavy bin

to any light bin

Similarly for a light bin $i > n \cdot \delta$:

$$\mathbf{E}\left[\left.\Delta y_i^{t+1} \mid \mathfrak{F}^t\right.\right] = \frac{\delta}{n}.$$

So, we can apply the drift theorem with probability vector

$$p = \begin{cases} \frac{\delta}{n} & \text{for } i \le n \cdot \delta, \\ \frac{1+\delta}{n} & \text{otherwise,} \end{cases}$$

For processes allocating more than one balls we have that:

$$\mathbf{E}\left[\left.\Delta\Phi_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]\leq\Phi_{i}^{t}\cdot\left(\alpha\cdot\mathbf{E}\left[\left.\Delta y_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]+\alpha^{2}\cdot\mathbf{E}\left[\left.(\Delta y_{i}^{t+1})^{2}\mid\mathfrak{F}^{t}\right.\right]\right),$$

$$\mathbf{E}\left[\right. \Delta \Psi_{i}^{t+1} \mid \mathfrak{F}^{t}\left.\right] \leq \Psi_{i}^{t} \cdot \left(-\alpha \cdot \mathbf{E}\left[\right. \Delta y_{i}^{t+1} \mid \mathfrak{F}^{t}\left.\right] + \alpha^{2} \cdot \mathbf{E}\left[\left. (\Delta y_{i}^{t+1})^{2} \mid \mathfrak{F}^{t}\left.\right]\right).$$

For TWINNING, for any heavy bin $i \leq n \cdot \delta$:

$$\mathbf{E}\left[\right.\Delta y_i^{t+1}\left|\right.\mathfrak{F}^t\left.\right] = \underbrace{\left(1-\frac{1}{n}\right)\cdot\frac{1}{n}}_{\text{Allocate one ball to bin }i} + \underbrace{\left(-\frac{1}{n}\right)\cdot\left(\delta-\frac{1}{n}\right)}_{\text{Allocate one ball}} + \underbrace{\left(-\frac{2}{n}\right)\cdot(1-\delta)}_{\text{Allocate two balls}} = \frac{\delta}{n} - \frac{1}{n}.$$

Similarly for a light bin $i > n \cdot \delta$:

$$\mathbf{E}\left[\left.\Delta y_i^{t+1} \mid \mathfrak{F}^t\right.\right] = \frac{\delta}{n}.$$

to any other heavy bin

to any light bin

So, we can apply the drift theorem with probability vector

$$p = \begin{cases} \frac{\delta}{n} & \text{for } i \le n \cdot \delta, \\ \frac{1+\delta}{n} & \text{otherwise,} \end{cases}$$

For processes allocating more than one balls we have that:

$$\mathbf{E}\left[\left.\Delta\Phi_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]\leq\Phi_{i}^{t}\cdot\left(\alpha\cdot\mathbf{E}\left[\left.\Delta y_{i}^{t+1}\mid\mathfrak{F}^{t}\right.\right]+\alpha^{2}\cdot\mathbf{E}\left[\left.(\Delta y_{i}^{t+1})^{2}\mid\mathfrak{F}^{t}\right.\right]\right),$$

$$\mathbf{E}\left[\left.\Delta \Psi_{i}^{t+1}\mid \mathfrak{F}^{t}\right.\right] \leq \Psi_{i}^{t}\cdot\left(-\alpha\cdot\mathbf{E}\left[\left.\Delta y_{i}^{t+1}\mid \mathfrak{F}^{t}\right.\right] + \alpha^{2}\cdot\mathbf{E}\left[\left.(\Delta y_{i}^{t+1})^{2}\mid \mathfrak{F}^{t}\right.\right]\right).$$

For TWINNING, for any heavy bin $i \leq n \cdot \delta$:

$$\mathbf{E}\left[\right.\Delta y_i^{t+1}\left|\right.\mathfrak{F}^t\left.\right] = \underbrace{\left(1-\frac{1}{n}\right)\cdot\frac{1}{n}}_{\text{Allocate one ball to bin }i} + \underbrace{\left(-\frac{1}{n}\right)\cdot\left(\delta-\frac{1}{n}\right)}_{\text{Allocate one ball}} + \underbrace{\left(-\frac{2}{n}\right)\cdot(1-\delta)}_{\text{Allocate two balls}} = \frac{\delta}{n} - \frac{1}{n}.$$

to any other heavy bin

to any light bin

Similarly for a light bin $i > n \cdot \delta$:

$$\mathbf{E}\left[\left.\Delta y_i^{t+1} \mid \mathfrak{F}^t\right.\right] = \frac{\delta}{n}.$$

So, we can apply the drift theorem with probability vector

$$p = \begin{cases} \frac{\delta}{n} & \text{for } i \le n \cdot \delta, \\ \frac{1+\delta}{n} & \text{otherwise,} \end{cases}$$

Potential functions

33

■ Mitzenmacher, Prabhakar and Shah [MPS02] introduced the Memory process which is allowed to store a bin in a *cache*.

■ Mitzenmacher, Prabhakar and Shah [MPS02] introduced the MEMORY process which is allowed to store a bin in a *cache*. In each step, it can either allocate to the cache or to a random bin.

■ Mitzenmacher, Prabhakar and Shah [MPS02] introduced the MEMORY process which is allowed to store a bin in a *cache*. In each step, it can either allocate to the cache or to a random bin.

 \blacksquare In Memory with resets the cache is *emptied* every r steps.

- Mitzenmacher, Prabhakar and Shah [MPS02] introduced the MEMORY process which is allowed to store a bin in a *cache*. In each step, it can either allocate to the cache or to a random bin.
- \blacksquare In Memory with resets the cache is *emptied* every r steps.
- For r = 2, a similar analysis to that of the TWINNING process, gives rise to the probability vector p of the TWO-CHOICE process, i.e., $p_i = \frac{2i-1}{n^2}$.

- Mitzenmacher, Prabhakar and Shah [MPS02] introduced the MEMORY process which is allowed to store a bin in a *cache*. In each step, it can either allocate to the cache or to a random bin.
- \blacksquare In Memory with resets the cache is *emptied* every r steps.
- For r = 2, a similar analysis to that of the TWINNING process, gives rise to the probability vector p of the TWO-CHOICE process, i.e., $p_i = \frac{2i-1}{n^2}$.
- Again, applying the drift theorem gives w.h.p. an $\mathcal{O}(\log n)$ upper bound on the gap.

Recall that in the **b-Batched setting**, a batch of b balls is allocated in parallel, using the allocation vector at the start of the batch.

- Recall that in the b-BATCHED setting, a batch of b balls is allocated in parallel, using the allocation vector at the start of the batch.
- Consider a process with an ϵ -biased allocation vector p which further satisfies $\max_{i \in [n]} p_i \leq \frac{C}{n}$.

- Recall that in the b-BATCHED setting, a batch of b balls is allocated in parallel, using the allocation vector at the start of the batch.
- Consider a process with an ϵ -biased allocation vector p which further satisfies $\max_{i \in [n]} p_i \leq \frac{C}{n}$.
- With a few *Taylor estimates*, we get

$$\mathbf{E}\left[\Phi_i^{t+b} \mid \Phi_i^t\right] \leq \Phi_i^t \cdot \left(1 + \left(p_i - \frac{1}{n}\right) \cdot b \cdot \alpha + \frac{5C^2b}{n} \cdot b \cdot \frac{\alpha^2}{n}\right),$$

- Recall that in the **b-Batched setting**, a batch of b balls is allocated in parallel, using the allocation vector at the start of the batch.
- Consider a process with an ϵ -biased allocation vector p which further satisfies $\max_{i \in [n]} p_i \leq \frac{C}{n}$.
- With a few *Taylor estimates*, we get

$$\mathbf{E}\left[\Phi_i^{t+b} \mid \Phi_i^t\right] \le \Phi_i^t \cdot \left(1 + \left(p_i - \frac{1}{n}\right) \cdot b \cdot \alpha + \frac{5C^2b}{n} \cdot b \cdot \frac{\alpha^2}{n}\right),$$

and similarly,

$$\mathbf{E}\left[\Psi_i^{t+b} \mid \Psi_i^t\right] \leq \Psi_i^t \cdot \left(1 + \left(\frac{1}{n} - p_i\right) \cdot b \cdot \alpha + \frac{5C^2b}{n} \cdot b \cdot \frac{\alpha^2}{n}\right).$$

- Recall that in the **b-Batched setting**, a batch of b balls is allocated in parallel, using the allocation vector at the start of the batch.
- Consider a process with an ϵ -biased allocation vector p which further satisfies $\max_{i \in [n]} p_i \leq \frac{C}{n}$.
- With a few *Taylor estimates*, we get

$$\mathbf{E}\left[\Phi_i^{t+b} \mid \Phi_i^t\right] \le \Phi_i^t \cdot \left(1 + \left(p_i - \frac{1}{n}\right) \cdot b \cdot \alpha + \frac{5C^2b}{n} \cdot b \cdot \frac{\alpha^2}{n}\right),$$

and similarly,

$$\mathbf{E}\left[\Psi_i^{t+b} \mid \Psi_i^t\right] \leq \Psi_i^t \cdot \left(1 + \left(\frac{1}{n} - p_i\right) \cdot b \cdot \alpha + \frac{5C^2b}{n} \cdot b \cdot \frac{\alpha^2}{n}\right).$$

Therefore, by the *drift theorem* over b steps, we get w.h.p. an $\mathcal{O}(\frac{b}{n} \cdot \log n)$ gap.

- Recall that in the **b-Batched setting**, a batch of b balls is allocated in parallel, using the allocation vector at the start of the batch.
- Consider a process with an ϵ -biased allocation vector p which further satisfies $\max_{i \in [n]} p_i \leq \frac{C}{n}$.
- With a few *Taylor estimates*, we get

$$\mathbf{E}\left[\Phi_i^{t+b} \mid \Phi_i^t\right] \leq \Phi_i^t \cdot \left(1 + \left(p_i - \frac{1}{n}\right) \cdot b \cdot \alpha + \frac{5C^2b}{n} \cdot b \cdot \frac{\alpha^2}{n}\right),$$

and similarly,

$$\mathbf{E}\left[\Psi_i^{t+b} \mid \Psi_i^t\right] \leq \Psi_i^t \cdot \left(1 + \left(\frac{1}{n} - p_i\right) \cdot b \cdot \alpha + \frac{5C^2b}{n} \cdot b \cdot \frac{\alpha^2}{n}\right).$$

- Therefore, by the *drift theorem* over b steps, we get w.h.p. an $\mathcal{O}(\frac{b}{n} \cdot \log n)$ gap.
- This is tight up to a $\log n$ factor for constant C > 1.

Conclusions

The drift theorem in $[\mathrm{PTW}15]$ can be used to:

The drift theorem in [PTW15] can be used to:

■ Analyze processes with an ϵ -biased allocation vector with weights.

The drift theorem in [PTW15] can be used to:

- Analyze processes with an ϵ -biased allocation vector with weights.
- Analyze graphical allocations via majorization (without weights).

The drift theorem in [PTW15] can be used to:

- Analyze processes with an ϵ -biased allocation vector with weights.
- Analyze graphical allocations via majorization (without weights).

The refined drift theorem can be used to:

The drift theorem in [PTW15] can be used to:

- Analyze processes with an ϵ -biased allocation vector with weights.
- Analyze graphical allocations via majorization (without weights).

The refined drift theorem can be used to:

Analyze processes allocating to more than one bins in each step.

The drift theorem in [PTW15] can be used to:

- Analyze processes with an ϵ -biased allocation vector with weights.
- Analyze graphical allocations via majorization (without weights).

The refined drift theorem can be used to:

- Analyze processes allocating to *more than one* bins in each step.
- Provide tighter bounds on $\mathbf{E}[\Gamma^t]$ (and so tighter characterization of the load vector).

The drift theorem in [PTW15] can be used to:

- Analyze processes with an ϵ -biased allocation vector with weights.
- Analyze graphical allocations via majorization (without weights).

The refined drift theorem can be used to:

- Analyze processes allocating to more than one bins in each step.
- Provide tighter bounds on $\mathbf{E}[\Gamma^t]$ (and so tighter characterization of the load vector).
- Analyze a wider range of settings (including outdated information, noise, etc.).

The drift theorem in [PTW15] can be used to:

- Analyze processes with an ϵ -biased allocation vector with weights.
- Analyze graphical allocations via majorization (without weights).

The refined drift theorem can be used to:

- Analyze processes allocating to *more than one* bins in each step.
- Provide tighter bounds on $\mathbf{E}[\Gamma^t]$ (and so tighter characterization of the load vector).
- Analyze a wider range of settings (including outdated information, noise, etc.).
- It is *agnostic* of the balanced allocations setting.

Still many open problems, including:

Still many open problems, including:

Proving sublogarithmic bounds for the weighted graphical setting.

Still many open problems, including:

- Proving sublogarithmic bounds for the weighted graphical setting.
- \blacksquare Gaps between $\log \log n$ and $\log n$ for the MEMORY process with resets.

Still many open problems, including:

- Proving sublogarithmic bounds for the weighted graphical setting.
- \blacksquare Gaps between $\log \log n$ and $\log n$ for the MEMORY process with resets.
- Analyse various load balancing algorithms used in practice (cf. envoy, nginx).

Still many open problems, including:

- Proving sublogarithmic bounds for the weighted graphical setting.
- \blacksquare Gaps between $\log \log n$ and $\log n$ for the MEMORY process with resets.
- Analyse various load balancing algorithms used in practice (cf. envoy, nginx).
- Analyse balls-into-bins with deletions.

Still many open problems, including:

- Proving sublogarithmic bounds for the weighted graphical setting.
- \blacksquare Gaps between $\log \log n$ and $\log n$ for the MEMORY process with resets.
- Analyse various load balancing algorithms used in practice (cf. envoy, nginx).
- Analyse balls-into-bins with deletions.
- Apply the drift theorem to other dynamic processes.

Still many open problems, including:

- Proving sublogarithmic bounds for the weighted graphical setting.
- \blacksquare Gaps between $\log \log n$ and $\log n$ for the MEMORY process with resets.
- Analyse various load balancing algorithms used in practice (cf. envoy, nginx).
- Analyse balls-into-bins with deletions.
- Apply the drift theorem to *other dynamic processes*.

Regarding techniques, there are also a few open questions:

Still many open problems, including:

- Proving sublogarithmic bounds for the weighted graphical setting.
- \blacksquare Gaps between $\log \log n$ and $\log n$ for the Memory process with resets.
- Analyse various load balancing algorithms used in practice (cf. envoy, nginx).
- Analyse balls-into-bins with deletions.
- Apply the drift theorem to other dynamic processes.

Regarding techniques, there are also a few open questions:

■ Is there a simpler way to argue that ϵ -biased processes have $\Theta(n)$ bins above the average load?

Still many open problems, including:

- Proving sublogarithmic bounds for the weighted graphical setting.
- \blacksquare Gaps between $\log \log n$ and $\log n$ for the Memory process with resets.
- Analyse various load balancing algorithms used in practice (cf. envoy, nginx).
- Analyse balls-into-bins with deletions.
- Apply the drift theorem to other dynamic processes.

Regarding techniques, there are also a few open questions:

- Is there a simpler way to argue that ϵ -biased processes have $\Theta(n)$ bins above the average load?
- Is there a way to adapt the drift theorem to work for processes with thresholds?

Still many open problems, including:

- Proving sublogarithmic bounds for the weighted graphical setting.
- \blacksquare Gaps between $\log \log n$ and $\log n$ for the Memory process with resets.
- Analyse various load balancing algorithms used in practice (cf. envoy, nginx).
- Analyse balls-into-bins with deletions.
- Apply the drift theorem to other dynamic processes.

Regarding techniques, there are also a few open questions:

- Is there a simpler way to argue that ϵ -biased processes have $\Theta(n)$ bins above the average load?
- Is there a way to adapt the drift theorem to work for processes with thresholds?
- Can we use a *single potential* to prove sublogarithmic bounds (e.g., the $\log_2 \log n + \Theta(1)$ bound for TWO-CHOICE)?

Questions?

 $More\ visualisations:\ {\tt dimitrioslos.com/phdthesis}$

Bibliography I

- D. Alistarh, J. Aspnes, and R. Gelashvili, *Space-optimal majority in population protocols*, 29th Annual ACM-SIAM Symposium on Discrete Algorithms (SODA'18), SIAM, 2018, pp. 2221–2239.
- ▶ D. Alistarh, t. Brown, J. Kopinsky, J. Z. Li, and G. Nadiradze, *Distributionally linearizable data structures*, 30th Annual ACM Symposium on Parallel Algorithms and Architectures (SPAA'18), ACM, 2018, pp. 133–142.
- Y. Azar, A. Z. Broder, A. R. Karlin, M. Mitzenmacher, and E. Upfal, *The ACM Paris Kanellakis Theory and Practice Award*, 2020, https://www.acm.org/media-center/2021/may/technical-awards-2020.
- Y. Azar, A. Z. Broder, A. R. Karlin, and E. Upfal, *Balanced allocations*, SIAM J. Comput. **29** (1999), no. 1, 180–200. MR 1710347
- D. Alistarh, R. Gelashvili, and J. Rybicki, Fast graphical population protocols, 25th International Conference on Principles of Distributed Systems (OPODIS'21), vol. 217, Schloss Dagstuhl Leibniz-Zentrum für Informatik, 2021, pp. 14:1–14:18.

Bibliography II

- D. Alistarh, J. Kopinsky, J. Li, and g. Nadiradze, The power of choice in priority scheduling, 36th Annual ACM-SIGOPT Principles of Distributed Computing (PODC'17), ACM, 2017, pp. 283–292.
- ▶ P. Berenbrink, A. Czumaj, M. Englert, T. Friedetzky, and L. Nagel, *Multiple-choice balanced allocation in (almost) parallel*, 16th International Workshop on Randomization and Computation (RANDOM'12), Springer-Verlag, 2012, pp. 411–422.
- P. Berenbrink, A. Czumaj, A. Steger, and B. Vöcking, Balanced allocations: the heavily loaded case, SIAM J. Comput. 35 (2006), no. 6, 1350–1385. MR 2217150
- ▶ M. Dahlin, *Interpreting stale load information*, IEEE Trans. Parallel Distributed Syst. **11** (2000), no. 10, 1033–1047.
- P. Delgado, D. Didona, F. Dinu, and W. Zwaenepoel, *Job-aware scheduling in eagle: Divide and stick to your probes*, 7th ACM Symposium on Cloud Computing (SoCC'16), ACM, 2016, pp. 497–509.

Bibliography III

- P. Delgado, F. Dinu, A. M. Kermarrec, and W. Zwaenepoel, *Hawk: Hybrid datacenter scheduling*, 2015 USENIX Annual Technical Conference (USENIX'15), USENIX, 2015, pp. 499–510.
- A. Gupta, R. Krishnaswamy, A. Kumar, and S. Singla, *Online carpooling using expander decompositions*, 40th IARCS Annual Conference on Foundations of Software Technology and Theoretical Computer Science (FSTTCS'20), vol. 182, Schloss Dagstuhl Leibniz-Zentrum für Informatik, 2020, pp. 23:1–23:14.
- ▶ G. H. Gonnet, Expected length of the longest probe sequence in hash code searching, J. Assoc. Comput. Mach. 28 (1981), no. 2, 289–304. MR 612082
- ▶ M. Khelghatdoust and V. Gramoli, *Peacock: Probe-based scheduling of jobs by rotating between elastic queues*, 24th International Conference on Parallel and Distributed Computing (Euro-Par'18), vol. 11014, Springer, 2018, pp. 178–191.
- R. M. Karp, M. Luby, and F. Meyer auf der Heide, Efficient PRAM simulation on a distributed memory machine, Algorithmica 16 (1996), no. 4-5, 517–542. MR 1407587

Bibliography IV

- K. Kenthapadi and R. Panigrahy, Balanced allocation on graphs, 17th Annual ACM-SIAM Symposium on Discrete Algorithms (SODA'06), ACM, 2006, pp. 434–443. MR 2368840
- D. Los and T. Sauerwald, Balanced allocations in batches: Simplified and generalized, 34th Annual ACM Symposium on Parallel Algorithms and Architectures (SPAA'22), ACM, 2022, p. 389–399.
- Y. Lu, Q. Xie, G. Kliot, A. Geller, J. R. Larus, and A. G. Greenberg, *Join-idle-queue: A novel load balancing algorithm for dynamically scalable web services*, Perform. Evaluation **68** (2011), no. 11, 1056–1071.
- ▶ M. Mitzenmacher, *The power of two choices in randomized load balancing*, Ph.D. thesis, University of California at Berkeley, 1996.
- How useful is old information?, IEEE Trans. Parallel Distributed Syst. 11 (2000), no. 1, 6–20.

Bibliography V

- ▶ M. Mitzenmacher, B. Prabhakar, and D. Shah, *Load balancing with memory*, The 43rd Annual IEEE Symposium on Foundations of Computer Science, 2002. Proceedings., IEEE, 2002, pp. 799–808.
- ▶ G. Nadiradze, On achieving scalability through relaxation, Ph.D. thesis, IST Austria, 2021.
- ▶ K. Ousterhout, P. Wendell, M. Zaharia, and I. Stoica, *Sparrow: distributed, low latency scheduling*, 24th ACM SIGOPS Symposium on Operating Systems Principles (SOSP'13), ACM, 2013, pp. 69–84.
- Y. Peres, K. Talwar, and U. Wieder, Graphical balanced allocations and the $(1+\beta)$ -choice process, Random Structures & Algorithms 47 (2015), no. 4, 760–775. MR 3418914
- ▶ M. Raab and A. Steger, "Balls into bins"—a simple and tight analysis, 2nd International Workshop on Randomization and Computation (RANDOM'98), vol. 1518, Springer, 1998, pp. 159–170. MR 1729169

Bibliography VI

- J. Spencer, Balancing games, J. Combinatorial Theory Ser. B 23 (1977), no. 1, 68–74. MR 526057
- ▶ W. Whitt, Deciding which queue to join: Some counterexamples, Oper. Res. **34** (1986), no. 1, 55–62.