Mis-Classified, Binary, Endogenous Regressors: Identification and Inference

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October 11th, 2018

What is the effect of T^* ?

$$y = c(\mathbf{x}) + \beta(\mathbf{x})T^* + \varepsilon$$

- ▶ y − Outcome of interest
- ▶ T* Unobserved, endogenous binary regressor
- ➤ T Observed, mis-measured binary surrogate for T*
- x Exogenous covariates
- ► z Discrete (typically binary) instrumental variable

Using a discrete IV to learn about $\beta(\mathbf{x})$

$$y = c(\mathbf{x}) + \beta(\mathbf{x})T^* + \varepsilon$$

Constributions of This Paper

- Show that only existing point identification result for mis-classified, endogenous T* is incorrect.
- 2. Sharp identified set for β under standard assumptions.
- 3. Point identification of β under slightly stronger assumptions.
- 4. Point out problem of weak identification in mis-classification models, develop identification-robust inference for β .

Example: Smoking and Birthweight (SNAP Trial)

Coleman et al. (N Engl J Med, 2012)

RCT with pregnant smokers in England: half given nicotine patches, the rest given placebo patches. Some given nicotine fail to quit; some given placebo quit.

- ▶ y Birthweight
- ▶ T* True smoking behavior
- ▶ T Self-reported smoking behavior
- x Mother characteristics
- z Indicator of nicotine patch

Baseline Assumptions I – Model & Instrument

Additively Separable Model

$$y = c(\mathbf{x}) + \beta(\mathbf{x})T^* + \varepsilon, \quad \mathbb{E}[\varepsilon] = 0$$

Valid & Relevant Instrument: $z \in \{0, 1\}$

- $ightharpoonup \mathbb{P}(T^* = 1 | \mathbf{x}, z = 1) \neq \mathbb{P}(T^* = 1 | \mathbf{x}, z = 0)$
- $\mathbb{E}[\varepsilon|\mathbf{x},z]=0$
- ▶ $0 < \mathbb{P}(z = 1 | \mathbf{x}) < 1$

Baseline Assumptions II – Measurement Error

Notation

- $\qquad \qquad \alpha_0(\mathbf{x}, z) \equiv \mathbb{P}\left(T = 1 | T^* = 0, \mathbf{x}, z\right)$
- $\qquad \qquad \alpha_1(\mathbf{x}, z) \equiv \mathbb{P}\left(T = 0 | T^* = 1, \mathbf{x}, z\right)$

Mis-classification unaffected by z

$$\alpha_0(\mathbf{x}, z) = \alpha_0(\mathbf{x}), \quad \alpha_1(\mathbf{x}, z) = \alpha_1(\mathbf{x})$$

Extent of Mis-classification

$$\alpha_0(\mathbf{x}) + \alpha_1(\mathbf{x}) < 1$$
 (T is positively correlated with T^*)

Non-differential Mis-classification

$$\mathbb{E}[\varepsilon|\mathbf{x},z,T^*,T] = \mathbb{E}[\varepsilon|\mathbf{x},z,T^*]$$

Existing Results

Correct: Exogenous *T**

- Mahajan (2006), Frazis & Loewenstein (2003)
- ▶ $\mathbb{E}[\varepsilon|\mathbf{x}, z, T^*] = 0 + \text{"Baseline"} \Rightarrow \beta(\mathbf{x}) \text{ identified.}$

Incorrect: Endogenous T*

- ► Mahajan (2006) A.2
- ▶ $\mathbb{E}[\varepsilon|\mathbf{x}, z, T^*, T] = \mathbb{E}[\varepsilon|\mathbf{x}, T^*] + \text{"Baseline"} \Rightarrow \beta(\mathbf{x}) \text{ identified.}$

We show: Mahajan's assumptions imply that the instrument z is uncorrelated with T^* unless T^* is in fact exogenous.

"Weak" Bounds

First-Stage

$$\rho_k(\mathbf{x}) \equiv \mathbb{P}(T=1|\mathbf{x},z=k)$$

IV Estimand

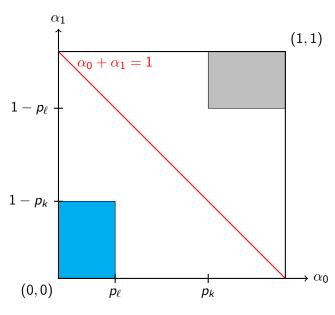
$$\frac{\mathbb{E}[y|\mathbf{x}, z=1] - \mathbb{E}[y|\mathbf{x}, z=0]}{p_1(\mathbf{x}) - p_0(\mathbf{x})} = \frac{\beta(\mathbf{x})}{1 - \alpha_0(\mathbf{x}) - \alpha_1(\mathbf{x})}$$

Bounds for (α_0, α_1)

$$\alpha_0(\mathbf{x}) \leq \min_k \left\{ p_k(\mathbf{x}) \right\}, \quad \alpha_1(\mathbf{x}) \leq \min_k \left\{ 1 - p_k(\mathbf{x}) \right\}$$
 prove

Bounds for β

 $\beta(\mathbf{x})$ is between IV and Reduced form; same sign as IV. \bullet



(Suppress x for simplicity)

Notation

- \triangleright z_k is shorthand for z = k

Iterated Expectations over T^*

$$\mathbb{E}(y|T=0,z_k) = (1-r_{0k})\mathbb{E}(y|T^*=0,T=0,z_k) + r_{0k}\mathbb{E}(y|T^*=1,T=0,z_k)$$

$$\mathbb{E}(y|T=1,z_k) = (1-r_{1k})\mathbb{E}(y|T^*=0,T=1,z_k) + r_{1k}\mathbb{E}(y|T^*=1,T=1,z_k)$$

(Suppress x for simplicity)

Notation

- \triangleright z_k is shorthand for z = k

Adding Non-differential Assumption

$$\mathbb{E}(y|T = 0, z_k) = (1 - r_{0k})\mathbb{E}(y|T^* = 0, z_k) + r_{0k}\mathbb{E}(y|T^* = 1, z_k)$$

$$\mathbb{E}(y|T = 1, z_k) = (1 - r_{1k})\mathbb{E}(y|T^* = 0, z_k) + r_{1k}\mathbb{E}(y|T^* = 1, z_k)$$

2 equations in 2 unknowns \Rightarrow solve for $\mathbb{E}(y|T^*=t^*,z=k)$ given (r_{0k},r_{1k}) .

Mixture Representation

$$F_{tk} = (1 - r_{tk})F_{tk}^0 + r_{tk}F_{tk}^1$$
$$F_{tk} \equiv y|(T = t, z = k)$$

$$F_{tk}^{t^*} \equiv y | (T^* = t^*, T = t, z = k)$$

Restrictions

- $\mathbb{E}(y|T^*,T,z) = \mathbb{E}(y|T^*,z)$ observable given (α_0,α_1)
- r_{tk} observable given (α_0, α_1)

Question

Given (α_0, α_1) can we always find (F_{tk}^0, F_{tk}^1) to satisfy the mixture model?

Equivalent Problem

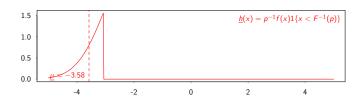
Given a specified CDF F, for what values of p and μ do there exist valid CDFs (G, H) with F = (1 - p)G + pH and $\mu = \text{mean}(H)$?

Necessary and Sufficient Condition if F is Continuous

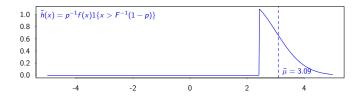
$$\underline{\mu}(F,p) \leq \mu \leq \overline{\mu}(F,p)$$

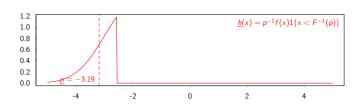
$$\underline{\mu}(F,p) \equiv \int_{-\infty}^{\infty} x \left[p^{-1} f(x) \mathbf{1} \{ x < F^{-1}(p) \} \right] dx = \int_{-\infty}^{\infty} x \underline{h}(x) dx$$

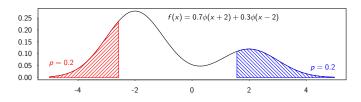
$$\overline{\mu}(F,p) \equiv \int_{-\infty}^{\infty} x \left[p^{-1} f(x) \mathbf{1} \{ x > F^{-1}(1-p) \} \right] dx = \int_{-\infty}^{\infty} x \overline{h}(x) dx$$

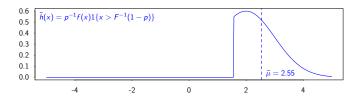


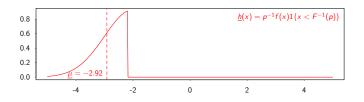


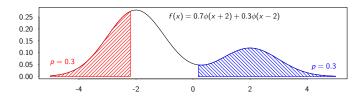


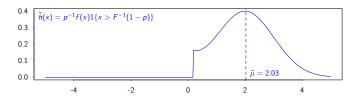


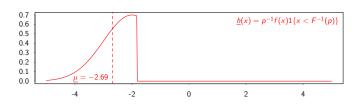


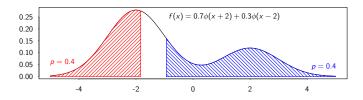


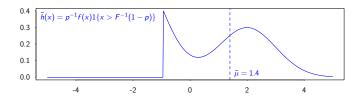


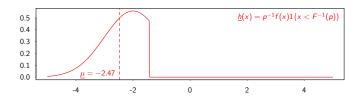


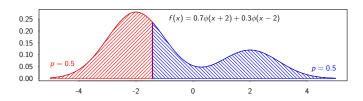


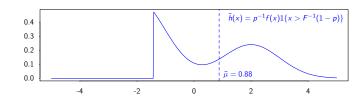


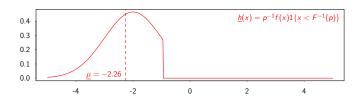






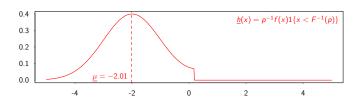


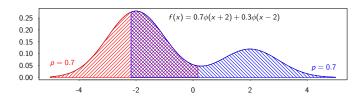


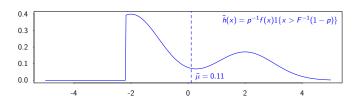


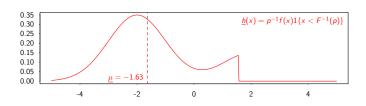


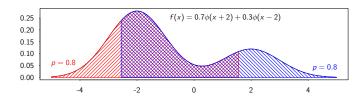


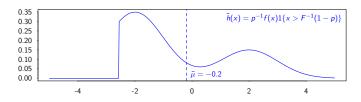


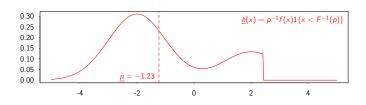


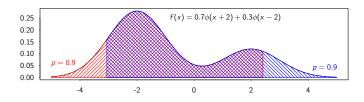


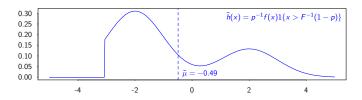


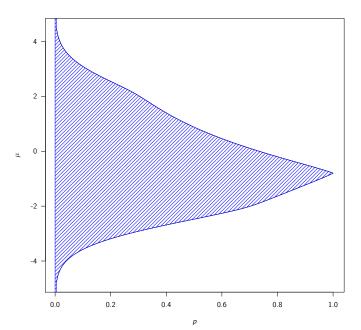












Sharp Identified Set under Baseline Assumptions

Theorem

- (i) As long as $\mathbb{E}[y|\mathbf{x}, T=0, z=k] \neq \mathbb{E}[y|\mathbf{x}, T=1, z=k]$ for some k, non-differential measurement error strictly improves the weak bounds for α_0, α_1 , and β .
- (ii) Under the baseline assumptions, β is not point identified, regardless of how many (discrete) values z takes on.

Corollary

Our bounds for α_0, α_1 , and β remain valid in a LATE model, although they may not be sharp, since they do not incorporate the testable implications of the LATE assumptions.

Point Identification: 1st Ingredient

Reparameterization

$$\begin{aligned} \theta_1(\mathbf{x}) &= \beta(\mathbf{x}) / \left[1 - \alpha_0(\mathbf{x}) - \alpha_1(\mathbf{x}) \right] \\ \theta_2(\mathbf{x}) &= \left[\theta_1(\mathbf{x}) \right]^2 \left[1 + \alpha_0(\mathbf{x}) - \alpha_1(\mathbf{x}) \right] \\ \theta_3(\mathbf{x}) &= \left[\theta_1(\mathbf{x}) \right]^3 \left[\left\{ 1 - \alpha_0(\mathbf{x}) - \alpha_1(\mathbf{x}) \right\}^2 + 6\alpha_0(\mathbf{x}) \left\{ 1 - \alpha_1(\mathbf{x}) \right\} \right] \\ & \boxed{\beta(\mathbf{x}) = 0 \iff \theta_1(\mathbf{x}) = \theta_2(\mathbf{x}) = \theta_3(\mathbf{x}) = 0} \end{aligned}$$

Lemma

Baseline Assumptions $\implies Cov(y, z|\mathbf{x}) = \theta_1(\mathbf{x})Cov(z, T|\mathbf{x}).$

Point Identification: 2nd Ingredient

Assumption (II)

$$\mathbb{E}[\varepsilon^2|\mathbf{x},z] = \mathbb{E}[\varepsilon^2|\mathbf{x}]$$

Lemma

(Baseline) + (II)
$$\Longrightarrow$$
 $Cov(y^2, z|\mathbf{x}) = 2Cov(yT, z|\mathbf{x})\theta_1(\mathbf{x}) - Cov(T, z|\mathbf{x})\theta_2(\mathbf{x})$

Corollary

(Baseline) + (II) + $[\beta(\mathbf{x}) \neq 0] \implies [\alpha_1(\mathbf{x}) - \alpha_0(\mathbf{x})]$ is identified.

Point Identification: 3rd Ingredient

Assumption (III)

- (i) $\mathbb{E}[\varepsilon^2|\mathbf{x}, z, T^*, T] = \mathbb{E}[\varepsilon^2|\mathbf{x}, z, T^*]$
- (ii) $\mathbb{E}[\varepsilon^3|\mathbf{x},z] = \mathbb{E}[\varepsilon^3|\mathbf{x}]$

Lemma

$$(Baseline) + (II) + (III) \implies$$

$$Cov(y^3, z|\mathbf{x}) = 3Cov(y^2T, z|\mathbf{x})\theta_1(\mathbf{x}) - 3Cov(yT, z|\mathbf{x})\theta_2(\mathbf{x}) + Cov(T, z|\mathbf{x})\theta_3(\mathbf{x})$$

Point Identification Result

Theorem

(Baseline) + (II) + (III) $\implies \beta(\mathbf{x})$ is point identified. If $\beta(\mathbf{x}) \neq 0$, then $\alpha_0(\mathbf{x})$ and $\alpha_1(\mathbf{x})$ are likewise point identified.

Explicit Solution

$$\beta(\mathbf{x}) = \operatorname{sign} \left[\theta_1(\mathbf{x})\right] \sqrt{3 \left[\theta_2(\mathbf{x})/\theta_1(\mathbf{x})\right]^2 - 2 \left[\theta_3(\mathbf{x})/\theta_1(\mathbf{x})\right]}$$

Sufficient for (II) and (III)

- (a) T is conditionally independent of (ε, z) given (T^*, \mathbf{x})
- (b) z is conditionally independent of ε given ${\bf x}$

Inference for a Mis-classified Regressor

Weak Identification

- ▶ β small \Rightarrow moment equalities uninformative about (α_0, α_1) \bigcirc more
- (α_0, α_1) could be on the boundary of the parameter space
- ▶ Also true of existing estimators that assume *T** exogenous

Our Approach

- Sharp identified set yields *inequality* moment restrictions that remain informative even if $\beta \approx 0$.
- ▶ Identification-robust inference with equality and inequality MCs.

Inference with Moment Equalities and Inequalities

Moment Conditions

$$\mathbb{E}\left[m_j(\mathbf{w}_i, \vartheta_0)\right] \ge 0, \quad j = 1, \cdots, J$$

$$\mathbb{E}\left[m_j(\mathbf{w}_i, \vartheta_0)\right] = 0, \quad j = J + 1, \cdots, J + K$$

Test Statistic

$$T_{n}(\vartheta) = \sum_{j=1}^{J} \left[\frac{\sqrt{n} \ \bar{m}_{n,j}(\vartheta)}{\widehat{\sigma}_{n,j}(\vartheta)} \right]_{-}^{2} + \sum_{j=J+1}^{J+K} \left[\frac{\sqrt{n} \ \bar{m}_{n,j}(\vartheta)}{\widehat{\sigma}_{n,j}(\vartheta)} \right]^{2}$$

Critical Value

- $\sqrt{n}\, ar{m}_n(\vartheta_0) o_d$ normal limit with covariance matrix $\Sigma(\vartheta_0)$
- ▶ Use this to bootstrap the limit dist. of $T_n(\vartheta)$ under $H_0: \vartheta = \vartheta_0$

Generalized Moment Selection

Andrews & Soares (2010)

- ▶ Inequalities that don't bind reduce power of test, so eliminate those that are "far from binding" before calculating critical value.
- ▶ Uniformly valid test of H_0 : $\vartheta = \vartheta_0$ even if ϑ_0 is not point identified.
- Not asymptotically conservative.

Problem

Joint test for the whole parameter vector but we're only interested in β . Projection is conservative and computationally intensive.

Our Solution: Bonferroni-Based Inference

Special Structure

- β only enters MCs through $\theta_1 = \beta/(1 \alpha_0 \alpha_1)$
- ▶ Strong instrument \Rightarrow inference for θ_1 is standard.
- ▶ Nuisance pars γ strongly identified under null for (α_0, α_1)

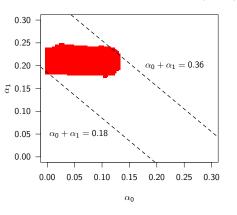
Procedure

- 1. Concentrate out $(\theta_1, \gamma) \Rightarrow$ joint GMS test for (α_0, α_1)
- 2. Invert test \Rightarrow $(1 \delta_1) \times 100\%$ confidence set for (α_0, α_1)
- 3. Project \Rightarrow CI for $(1 \alpha_0 \alpha_1)$
- 4. Construct standard $(1 \delta_2) \times 100\%$ IV CI for θ_1
- 5. Bonferroni \Rightarrow $(1 \delta_1 \delta_2) \times 100\%$ CI for β

Example

(sim data:
$$\beta = 1, \alpha_0 = 0.1, \alpha_1 = 0.2, n = 5000$$
)

97.5% GMS Confidence Region for (α_0, α_1)



Bonferroni Interval

- 1. 97.5% CI for $(1 \alpha_0 \alpha_1) = (0.64, 0.82)$
- 2. 97.5% CI for $\theta_1 = (1.20, 1.47)$
- 3. > 95% CI for β : $(0.64 \times 1.20, 0.82 \times 1.47) = (0.77, 1.21)$

Comparisons

- \triangleright (0.88, 1.04) for IV if T^* were observed
- \blacktriangleright (1.22,1.45) for naive IV interval using T

Conclusion

- Identification and inference for effect of binary, mis-classified, endogenous regressor.
- Only existing point identification result is incorrect.
- ▶ Sharp identified set for $\beta(\mathbf{x})$ under standard assumptions.
- ▶ Point identification of $\beta(\mathbf{x})$ under slightly stronger assumptions.
- Point out weak identification problem in mis-classification models, develop identification-robust inference for $\beta(\mathbf{x})$.

Simple Bounds for Mis-classification from First-stage

Unobserved Observed
$$ho_k^*(\mathbf{x}) \equiv \mathbb{P}(T^*=1|\mathbf{x},z=k)$$
 $p_k(\mathbf{x}) \equiv \mathbb{P}(T=1|\mathbf{x},z=k)$

Relationship

$$p_k^*(\mathbf{x}) = \frac{p_k(\mathbf{x}) - \alpha_0(\mathbf{x})}{1 - \alpha_0(\mathbf{x}) - \alpha_1(\mathbf{x})}, \quad k = 0, 1$$

z does not affect (α_0, α_1) ; denominator $\neq 0$

Bounds for Mis-classification

$$\alpha_0(\mathbf{x}) \leq p_k(\mathbf{x}) \leq 1 - \alpha_1(\mathbf{x}), \quad k = 0, 1$$

$$\alpha_0(\mathbf{x}) + \alpha_1(\mathbf{x}) < 1$$



What does IV estimate under mis-classification?

Unobserved

$$\beta(\mathbf{x}) = \frac{\mathbb{E}[y|\mathbf{x}, z=1] - \mathbb{E}[y|\mathbf{x}, z=0]}{p_1^*(\mathbf{x}) - p_0^*(\mathbf{x})}$$

Wald (Observed)

$$\frac{\mathbb{E}[y|\mathbf{x},z=1] - \mathbb{E}[y|\mathbf{x},z=0]}{p_1(\mathbf{x}) - p_0(\mathbf{x})} = \beta(\mathbf{x}) \left[\frac{p_1^*(\mathbf{x}) - p_0^*(\mathbf{x})}{p_1(\mathbf{x}) - p_0(\mathbf{x})} \right] = \frac{\beta(\mathbf{x})}{1 - \alpha_0(\mathbf{x}) - \alpha_1(\mathbf{x})}$$

$$p_1^*(\mathbf{x}) - p_0^*(\mathbf{x}) = \frac{p_1(\mathbf{x}) - \alpha_0(\mathbf{x})}{1 - \alpha_0 - \alpha_1(\mathbf{x})} - \frac{p_0(\mathbf{x}) - \alpha_0(\mathbf{x})}{1 - \alpha_0 - \alpha_1(\mathbf{x})} = \frac{p_1(\mathbf{x}) - p_0(\mathbf{x})}{1 - \alpha_0(\mathbf{x}) - \alpha_1(\mathbf{x})}$$



Partial Identification Bounds for $\beta(\mathbf{x})$

$$\beta(\mathbf{x}) = [1 - \alpha_0(\mathbf{x}) - \alpha_1(\mathbf{x})] \left[\frac{\mathbb{E}[y|\mathbf{x}, z = 1] - \mathbb{E}[y|\mathbf{x}, z = 0]}{\rho_1(\mathbf{x}) - \rho_0(\mathbf{x})} \right]$$

$$0 \le \alpha_0 \le \min_k \{ p_k(\mathbf{x}) \}, \quad 0 \le \alpha_1 \le \min_k \{ 1 - p_k(\mathbf{x}) \}$$

No Mis-classification

$$\alpha_0(\mathbf{x}) = \alpha_1(\mathbf{x}) = 0 \implies \beta(\mathbf{x}) = \mathsf{Wald}$$

Maximum Mis-classification

$$\alpha_0(\mathbf{x}) = p_{\min}(\mathbf{x}), \ \alpha_1(\mathbf{x}) = 1 - p_{\max}(\mathbf{x})$$

$$\Rightarrow 1 - \alpha_0(\mathbf{x}) - \alpha_1(\mathbf{x}) = p_{\text{max}}(\mathbf{x}) - p_{\text{min}}(\mathbf{x}) = |p_1(\mathbf{x}) - p_0(\mathbf{x})|$$
$$\Rightarrow \beta(\mathbf{x}) = \text{sign} \{p_1(\mathbf{x}) - p_0(\mathbf{x})\} \times (\text{Reduced Form})$$

Just-Identified System of Moment Equalities

Suppress dependence on x...

$$\mathbb{E}\left[\left\{\mathbf{\Psi}(\boldsymbol{\theta})\mathbf{w}_{i}-\boldsymbol{\kappa}\right\} \otimes \begin{pmatrix} 1\\z \end{pmatrix}\right] = \mathbf{0}$$

$$\mathbf{\Psi}(\boldsymbol{\theta}) \equiv \begin{bmatrix} -\theta_{1} & 1 & 0 & 0 & 0 & 0\\ \theta_{2} & 0 & -2\theta_{1} & 1 & 0 & 0\\ -\theta_{3} & 0 & 3\theta_{2} & 0 & -3\theta_{1} & 1 \end{bmatrix}$$

$$\begin{aligned} \mathbf{w}_{i} &= (T_{i}, y_{i}, y_{i}T_{i}, y_{i}^{2}, y_{i}^{2}T_{i}, y_{i}^{3})' & \theta_{1} &= \beta/(1 - \alpha_{0} - \alpha_{1}) \\ \kappa &= (\kappa_{1}, \kappa_{2}, \kappa_{3})' & \theta_{2} &= \theta_{1}^{2}(1 + \alpha_{0} - \alpha_{1}) \\ \theta_{3} &= \theta_{1}^{3} \left[(1 - \alpha_{0} - \alpha_{1})^{2} + 6\alpha_{0}(1 - \alpha_{1}) \right] \end{aligned}$$

▶ back

Moment Inequalities I – First-stage Probabilities

$$\alpha_0 \leq p_k \leq 1 - \alpha_1$$
 becomes $\mathbb{E}[m(\mathbf{w}_i, \boldsymbol{\vartheta})] \geq \mathbf{0}$ for all k where

$$m(\mathbf{w}_i, \vartheta) \equiv \left[\begin{array}{c} \mathbf{1}(z_i = k)(T - \alpha_0) \\ \mathbf{1}(z_i = k)(1 - T_i - \alpha_1) \end{array} \right]$$

Moment Inequalities II – Non-differential Assumption

For all k, we have $\mathbb{E}[m(\mathbf{w}_i, \boldsymbol{\vartheta}, \mathbf{q}_k)] \geq 0$ where

$$m(\mathbf{w}_{i}, \boldsymbol{\vartheta}, \mathbf{q}_{k}) \equiv \begin{bmatrix} y_{i} \mathbf{1} \left(z_{i} = k\right) \left\{ \left(T_{i} - \alpha_{0}\right) - \mathbf{1} \left(y_{i} \leq \underline{q}_{0k}\right) \left(1 - T_{i}\right) \left(\frac{1 - \alpha_{0} - \alpha_{1}}{\alpha_{1}}\right) \right\} \\ -y_{i} \mathbf{1} \left(z_{i} = k\right) \left\{ \left(T_{i} - \alpha_{0}\right) - \mathbf{1} \left(y_{i} \geq \overline{q}_{0k}\right) \left(1 - T_{i}\right) \left(\frac{1 - \alpha_{0} - \alpha_{1}}{\alpha_{1}}\right) \right\} \\ y_{i} \mathbf{1} \left(z_{i} = k\right) \left\{ \left(T_{i} - \alpha_{0}\right) - \mathbf{1} \left(y_{i} \leq \underline{q}_{1k}\right) T_{i} \left(\frac{1 - \alpha_{0} - \alpha_{1}}{1 - \alpha_{1}}\right) \right\} \\ -y_{i} \mathbf{1} \left(z_{i} = k\right) \left\{ \left(T_{i} - \alpha_{0}\right) - \mathbf{1} \left(y_{i} \geq \overline{q}_{1k}\right) T_{i} \left(\frac{1 - \alpha_{0} - \alpha_{1}}{1 - \alpha_{1}}\right) \right\} \end{bmatrix}$$

and $\mathbf{q}_k \equiv (\underline{q}_{0k},\,\overline{q}_{0k},\,\underline{q}_{1k},\,\overline{q}_{1k})'$ defined by $\mathbb{E}[h(\mathbf{w}_i,\vartheta,\mathbf{q}_k)]=0$ with

$$h(\mathbf{w}_{i}, \vartheta, \mathbf{q}_{k}) = \begin{bmatrix} \mathbf{1}(y_{i} \leq \underline{q}_{0k})\mathbf{1}(z_{i} = k)(1 - T_{i}) - \left(\frac{\alpha_{1}}{1 - \alpha_{0} - \alpha_{1}}\right)\mathbf{1}(z_{i} = k)(T_{i} - \alpha_{0}) \\ \mathbf{1}(y_{i} \leq \overline{q}_{0k})\mathbf{1}(z_{i} = k)(1 - T_{i}) - \left(\frac{1 - \alpha_{1}}{1 - \alpha_{0} - \alpha_{1}}\right)\mathbf{1}(z_{i} = k)(1 - T_{i} - \alpha_{1}) \\ \mathbf{1}(y_{i} \leq \underline{q}_{1k})\mathbf{1}(z_{i} = k)T_{i} - \left(\frac{1 - \alpha_{1}}{1 - \alpha_{0} - \alpha_{1}}\right)\mathbf{1}(z_{i} = k)(T_{i} - \alpha_{0}) \\ \mathbf{1}(y_{i} \leq \overline{q}_{1k})\mathbf{1}(z_{i} = k)T_{i} - \left(\frac{\alpha_{0}}{1 - \alpha_{0} - \alpha_{1}}\right)\mathbf{1}(z_{i} = k)(1 - T_{i} - \alpha_{1}) \end{bmatrix}$$

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