Estimating the Effect of a Mis-measured, Endogenous, Binary Treatment

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What is the causal effect of T^* ?

$$y_i = h(T_i^*, \mathbf{x}_i) + \varepsilon_i$$

- ▶ y − Outcome of interest
- ▶ h Unknown function that does not depend on i
- ▶ T* Unobserved, endogenous binary treatment
- ➤ T Observed, mis-measured binary surrogate for T*
- ▶ x − Exogenous covariates
- \triangleright ε Mean-zero error term
- ▶ z − Discrete instrumental variable

Example: Smoking and Birthweight (SNAP Trial)

Coleman et al. (N Engl J Med, 2012)

RCT with 1050 pregnant smokers in England: 521 given nicotine patches, the rest given placebo patches.

- ▶ y Birthweight
- ▶ T* True smoking behavior
- ► T Self-reported smoking behavior
- x Mother characteristics
- z Indicator of nicotine patch

Example: Schooling and Test Scores

Burde & Linden (2013, AEJ Applied)

RCT in Afghanistan: 32 villages divided into 11 clusters. Randomly choose 6 and build a school in each village of these clusters.

- ▶ y Child's score on math and language test
- ► T* Child's true school attendance
- ➤ T Parent's report of child's school attendance
- x Child and household characteristics
- ▶ z School built in village

Related Literature

Continuous Treatment

Lewbel (1997, 2012), Schennach (2004, 2007), Chen et al. (2005), Hu & Schennach (2008), Song (2015), Hu et al. (2015)...

Binary, Exogenous Treatment

Aigner (1973), Bollinger (1996), Kane et al. (1999), Black et al. (2000), Frazis & Loewenstein (2003), Mahajan (2006), Lewbel (2007)

Binary, Endogenous Treatment

Mahajan (2006), Shiu (2015), Ura (2015)

Model: $y = c + \beta T^* + \varepsilon$

First-stage

$$ho_k^* \equiv \mathbb{P}(T^* = 1|z = z_k)
eq \mathbb{P}(T^* = 1|z = z_\ell) \equiv
ho_\ell^*, \ k
eq \ell$$

Measurement Error

- ▶ Non-differential: $\mathbb{E}[\varepsilon|T^*,T,z] = \mathbb{E}[\varepsilon|T^*,z]$
- Does not depend on z:

$$\alpha_0 = \mathbb{P}(T=1|T^*=0,z)$$

$$\alpha_1 = \mathbb{P}(T=0|T^*=1,z)$$

Notation

Define error term that absorbs constant: $u = c + \varepsilon$

Observable Moments: $y = \beta T^* + u$

$$ar{y}_{tk} = \mathbb{E}[y|T=t,z=z_k], \quad p_{tk} = q_k p_k$$
 $q_k = \mathbb{P}(z=z_k), \quad p_k = \mathbb{P}(T=1|z=z_k)$

Unobservable Moments: $y = \beta T^* + u$

$$z = 1 z = 2 ... z = K$$

$$T^* = 0 m_{01}^* m_{02}^* ... m_{0K}^* p_{0K}^*$$

$$T^* = 1 m_{11}^* m_{12}^* ... m_{1K}^* p_{1K}^*$$

$$m_{tk}^* = \mathbb{E}[u|T^* = t, z = z_k], \quad p_{tk}^* = q_k p_k^*$$

 $q_k = \mathbb{P}(z = z_k), \quad p_k^* = \mathbb{P}(T^* = 1|z = z_k)$

Unrestricted System of Equations

$$(1 - p_k)\bar{y}_{0k} \equiv \widetilde{y}_{0k} = (\beta + m_{1k}^*)\alpha_1 p_k^* + (1 - \alpha_0)(1 - p_k^*) m_{0k}^*$$

$$p_k \bar{y}_{1k} \equiv \widetilde{y}_{1k} = (\beta + m_{1k}^*)(1 - \alpha_1) p_k^* + \alpha_0 (1 - p_k^*) m_{0k}^*$$

$$p_k^* = \frac{p_k - \alpha_0}{1 - \alpha_0 - \alpha_1}$$

Possible Restrictions On m_{tk}^*

Joint Exogeneity:
$$\mathbb{E}[\varepsilon|T^*,z]=0$$
 $\implies m_{tk}^*=c \quad \text{for all } t,k$

Exogenous Treatment: $\mathbb{E}[\varepsilon|T^*]=0$
 $\implies \frac{1}{\mathbb{P}(T^*=t)}\sum_k p_{tk}^* m_{tk}^*=c \quad \text{for all } t$

Exogenous Instrument: $\mathbb{E}[\varepsilon|z]=0$
 $\implies (1-p_k^*)m_{0k}^*+p_k^*m_{1k}^*=c \quad \text{for all } k$

Later I'll consider relaxing the assumption that z is exogenous...

Theorem: β is undentified regardless of K.

(For general case, see paper.)

Proof of special case: $\alpha_0 = 0$

1. System of equations:

$$\widetilde{y}_{0k} = c + p_k \beta \left(\frac{\alpha_1}{1 - \alpha_1}\right) - p_k m_{1k}^*$$

$$\widetilde{y}_{1k} = p_k \beta + p_k m_{1k}^*$$

2. $\beta/(1-\alpha_1) \equiv \mathcal{W}$ is identified and imposing this, algebra gives $\beta \alpha_1/(1-\alpha_1) = \mathcal{W} - \beta$.

Theorem: β is undentified regardless of K.

(For general case, see paper.)

Proof of special case: $\alpha_0 = 0$ continued...

3. Substituting:

$$(c + p_k W - \widetilde{y}_{0k})/p_k = \beta + m_{1k}^*$$
$$\widetilde{y}_{1k}/p_k = \beta + m_{1k}^*$$

- 4. Linear system in (β, m_{1k}^*) no solution or ∞ of solutions.
- 5. Sum original pair of equations $\implies c + p_k W \widetilde{y}_{0k} = \widetilde{y}_{1k}$ thus ∞ of solutions. The model is unidentified.

Conditional Second Moment Independence.

New Assumption

Homoskedastic errors w.r.t. the *instrument*: $E[\varepsilon^2|z] = E[\varepsilon^2]$

Reasonable?

Makes sense in an RCT or a true natural experiment.

New Moment Conditions

Defining
$$\mu_{k\ell}^* = (p_k - \alpha_0)m_{1k}^* - (p_\ell - \alpha_0)m_{k\ell}^*$$
,

$$\mathbb{E}(y^2|z_k) - \mathbb{E}(y^2|z_\ell) \equiv \Delta \overline{y^2} = \beta \mathcal{W}(p_k - p_\ell) + 2\mathcal{W}\mu_{k\ell}^*$$

$$\mathbb{E}(yT|z_k) - \mathbb{E}(yT|z_\ell) \equiv \Delta \overline{yT} = (1 - \alpha_1)\mathcal{W}(p_k - p_\ell) + \mu_{k\ell}^*$$

Theorem: $(\alpha_1 - \alpha_0)$ is Identified if $E[\varepsilon^2|z] = E[\varepsilon^2]$ Requires only binary z

Solve for $\mu_{k\ell}^*$, substitute $\beta = \mathcal{W}(1 - \alpha_0 - \alpha_1)$, rearrange to find

$$lpha_1 - lpha_0 = 1 + \mathcal{R}/\mathcal{W}, \quad ext{where} \quad \mathcal{R} \equiv rac{\Delta \overline{y^2} - 2 \mathcal{W} \Delta \overline{y} \overline{T}}{\mathcal{W}(p_k - p_\ell)}.$$

What good is $(\alpha_1 - \alpha_0)$?

- ▶ Test necessary condition for *no mis-classification*: $\alpha_0 = \alpha_1$
- ightharpoonup Simple, tighter partial identification bounds for eta
- ▶ If α_0 known, e.g. zero $\implies \beta$ point identified

Conditional Third Moment Independence

New Assumption

Third moment independence w.r.t instrument: $E[\varepsilon^3|z] = E[\varepsilon^3]$

New Moment Conditions

Define
$$\lambda_{k\ell}^* = (p_k - \alpha_0)v_{1k}^* - (p_\ell - \alpha_0)v_{1\ell}^*$$

where $v_{tk}^* = \mathbb{E}(u^2|T^* = t, z_k)$. Then

$$\begin{split} \mathbb{E}(y^3|z_k) &- \mathbb{E}(y^3|z_\ell) \equiv \\ \Delta \overline{y^3} &= \beta^2 \mathcal{W}(p_k - p_\ell) + 3\beta \mathcal{W} \mu_{k\ell}^* + 3\mathcal{W} \lambda_{k\ell}^* \\ \mathbb{E}(y^2 T|z_k) &- \mathbb{E}(y^2 T|z_\ell) \equiv \\ \Delta \overline{y^2 T} &= \beta(1 - \alpha_1) \mathcal{W}(p_k - p_\ell) + 2(1 - \alpha_1) \mathcal{W} \mu_{k\ell}^* + \lambda_{k\ell}^* \end{split}$$

Theorem: β , α_0 and α_1 identified

Adding $E[\varepsilon^3|z] = E[\varepsilon^3]$, z need only be binary.

Solve for $\lambda_{k\ell}^*$, substitute and rearrange. After further substitutions:

$$2\mathcal{W}^2(1-\alpha_1)^2 + 2\mathcal{R}\mathcal{W}(1-\alpha_1) + (\mathcal{S} - \mathcal{R}^2) = 0$$

where

$$\mathcal{S} \equiv rac{\Delta \overline{y^3} - 3 \mathcal{W} \left[\Delta \overline{y^2 \, T} + \mathcal{R} \Delta \overline{y \, T}
ight]}{\mathcal{W}(
ho_k -
ho_\ell)}$$

- Quadratic in $(1 \alpha_1)$ and observables only
- ▶ Always two real roots: one is $(1 \alpha_1)$ and the other is α_0 .
- ▶ To tell which is which, need $\alpha_0 + \alpha_1 < 1$.

Recap of Results

- 1. Using first-moment information alone, β is unidentified regardless of how many values the instrument takes on.
- 2. Using second moment information $\alpha_1 \alpha_0$ is identified
 - ightharpoonup Partial identification bound for β
 - ▶ Identifies β if α_0 is known (e.g. smoking/birthweight example)
- 3. Using third moment information β , α_0 and α_1 are identified so long as $\alpha_0 + \alpha_1 < 1$.

Empirical Illustration: Schooling and Test Scores

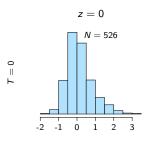
Burde & Linden (2013, AEJ Applied)

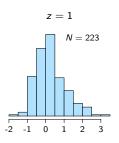
RCT in Afghanistan: 32 villages divided into 11 clusters. Randomly choose 6 and build a school in each village of these clusters (N = 1468).

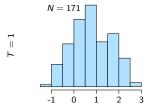
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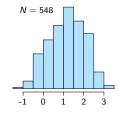
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Empirical Illustration: Schooling and Test Scores

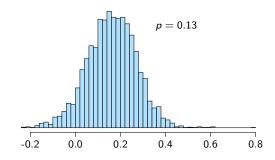
Burde & Linden (2013, AEJ Applied)

$$\widehat{\beta}_{OLS} = 0.88$$

$$\widehat{\beta}_{IV} = 1.27$$

$$\widehat{\alpha}_1 - \widehat{\alpha}_0 = 0.18$$

Cluster Bootstrap Distribution of $\hat{\alpha}_1 - \hat{\alpha}_0$



But what if z is endogenous?

Recall: Unrestricted System

$$\widetilde{y}_{0k} = (\beta + m_{1k}^*)\alpha_1 p_k^* + (1 - \alpha_0)(1 - p_k^*) m_{0k}^*$$

$$\widetilde{y}_{1k} = (\beta + m_{1k}^*)(1 - \alpha_1) p_k^* + \alpha_0(1 - p_k^*) m_{0k}^*$$

Intelligible Quantities

$$\delta_{T^*} \equiv \mathbb{E}[u|T^*=1] - \mathbb{E}[u|T^*=0]$$

$$\delta_z \equiv \mathbb{E}[u|z=1] - \mathbb{E}[u|z=0]$$

... both are linear functions of m_{tk}^* .

Identified Set for $(\alpha_0, \alpha_1, \delta_{T^*}, \delta_z)$

First Moment Information

$$\delta_{\mathbf{z}} = C(\alpha_0, \alpha_1 | \mathbf{p}, \mathbf{q}, \overline{\mathbf{y}}) - \left(\frac{p_1 - p_2}{1 - \alpha_0 - \alpha_1}\right) \delta_{T^*}$$

Second Moment Information

$$Var(u|T=t,z=k)>0$$

$$\Longrightarrow \left[Var(y|T=t,z=k)-Q_{tk}(\delta_{T^*},\delta_z,\alpha_0,\alpha_1|\mathbf{p},\mathbf{q},\mathbf{\bar{y}})\right]>0$$

Approaches to (Partial) Identification

Identification

- $\delta_z = 0, \alpha_0 = \alpha_1 = 0 \Rightarrow \text{Wald Estimator}$
- ▶ Joint Exogeneity ($\Rightarrow \delta_{T^*} = \delta_z = 0$) Kane et al. (1999), Black et al. (2000), Mahajan (2006)...

Partial Identification

- ► Frazis & Loewenstein (2003): $\delta_z = 0$, $(\alpha_0 + \alpha_1) \in [\ell, u]$
- ▶ Conley et al. (2012): $\delta_z \in [\underline{\delta}_z, \overline{\delta}_z]$, $\alpha_0 = \alpha_1 = 0$
- Nevo & Rosen (2012): $\delta_T^* > \delta_z$, $\delta_T^* \delta_z > 0$, $\alpha_0 = \alpha_1 = 0$

Our Proposed Approach

Elicit Beliefs

Ask researcher for bounds on $\alpha_0, \alpha_1, \delta_{T^*}, \delta_z$

Discipline Beliefs

Are these beliefs mutually consistent? Explore joint constraints implied by identified set.

Incorporate Beliefs

Carry out (Bayesian) inference for β using beliefs, constraints, and accounting for sampling uncertainty.

Example: Vouchers for Private Schooling (PACES)

Angrist et al. (2002, AER)

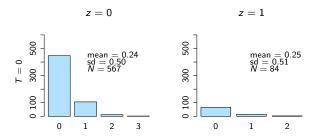
Data from Colombia: vouchers to attend private school awarded by lottery to poor, primary school-aged children (N=1577).

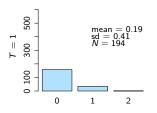
- ▶ y # of grades repeated after lottery
- ▶ T* Scholarship use
- ► T Self-reported Scholarship use
- ▶ x − Demographic controls
- ▶ z Offered scholarship through lottery

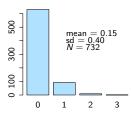
Authors raise concerns about the lottery in one of the two cities. . .

Example: Vouchers for Private Schooling (PACES)

Overall: Mean = 0.19, SD = 0.45

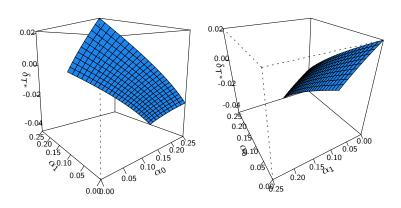


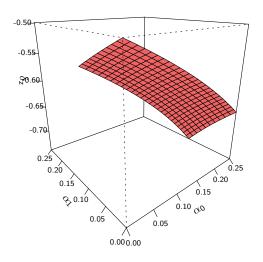


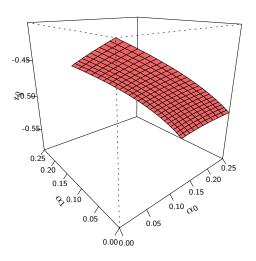


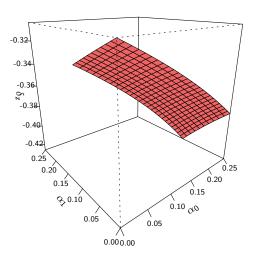
Implications of Valid IV: $\delta_z = 0$

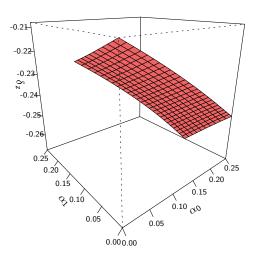
Angrist et al. (2002)

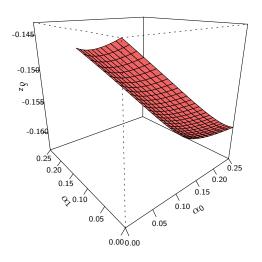


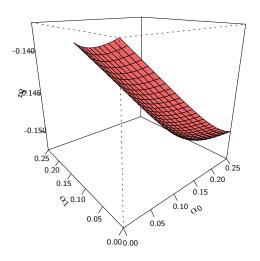


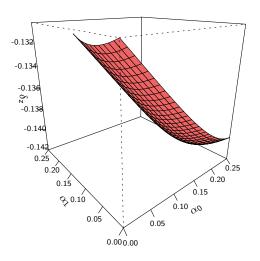


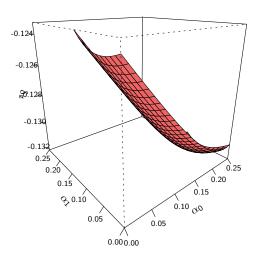


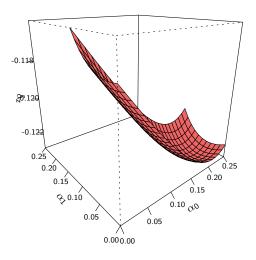


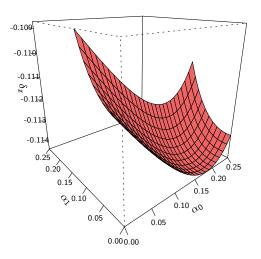


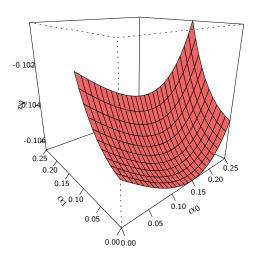


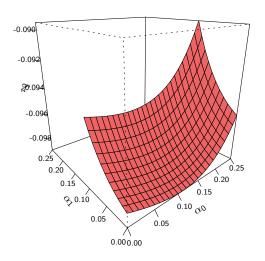


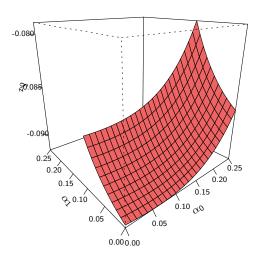


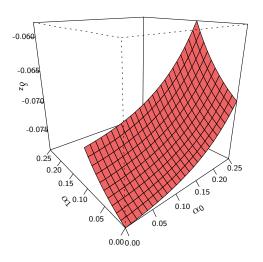


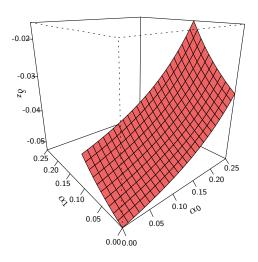


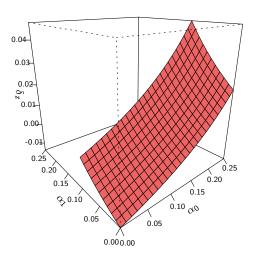












(Bayesian) Inference via Transparent Parameterization

Reduced Form Parameters: $(\mathbf{p}, \mathbf{q}, \overline{\mathbf{y}}, \sigma^2)$

▶ Draw reduced form parameters from Bayesian posterior constructed to match usual large-sample frequentist inference.

Structural Parameters $(\delta_{T^*}, \delta_z, \alpha_0, \alpha_1)$

- ▶ Updated by data only *through* reduced from parameters.
- Impose prior beliefs on structural parameters
- ► E.g. $\alpha_0, \alpha_1 \sim \text{ iid } U(0, 0.25), \ \delta_{T^*} | (\alpha_0, \alpha_1) \sim U(\ell, u)$

"Blue Bar" Picture

Implications for Beta

Bayes histogram next to Frequentist histogram (picture of lower bound)

Probability that frequentist identified sets include zero? Something like that

Conclusion

Simulation Study

Simulation Study: $y = \beta T^* + \varepsilon$

- (ε, η) \sim jointly normal, mean 0, variance 1, corr. 0.3.
- ▶ First stage: $T^* = \mathbf{1} \{ \gamma_0 + \gamma_1 z + \eta > 0 \}$
 - ▶ Half of subjects have z = 1, the rest have z = 0.
 - ho $\gamma_0 = \Phi^{-1}(\delta)$

 - $m \delta$ equals fraction of those offered treatment who fail to take it up *and* fraction of those not offered treatment who do.
- Generate T as follows:
 - $T^*=0 \implies T=0$, i.e. $\alpha_0=0$
 - $ightharpoonup T | T^* = 1 \sim \mathsf{Bernoulli}(1 \alpha_1)$
 - α_0, α_1 unknown to econometrician.

Sampling Distribution of $\hat{\alpha}_1 - \hat{\alpha}_0$