

# Estimating the Effect of a Mis-measured, Endogenous, Binary Treatment

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## What is the causal effect of $T^*$ ?

$$y_i = h(T_i^*, \mathbf{x}_i) + \varepsilon_i$$

- ▶  $y$  – Outcome of interest
- ▶  $h$  – Unknown function that *does not depend on  $i$*
- ▶  $T^*$  – Unobserved, endogenous binary treatment
- ▶  $T$  – Observed, mis-measured binary surrogate for  $T^*$
- ▶  $\mathbf{x}$  – Exogenous covariates
- ▶  $\varepsilon$  – Mean-zero error term
- ▶  $z$  – Discrete (typically binary) instrumental variable

### Target of Inference:

ATE function:  $\tau(\mathbf{x}) = h(1, \mathbf{x}) - h(0, \mathbf{x})$

## Example: Smoking and Birthweight (SNAP Trial)

Coleman et al. (N Engl J Med, 2012)

RCT with 1050 pregnant smokers in England: 521 given nicotine patches, the rest given placebo patches.

- ▶  $y$  – Birthweight
- ▶  $T^*$  – True smoking behavior
- ▶  $T$  – Self-reported smoking behavior
- ▶  $\mathbf{x}$  – Mother characteristics
- ▶  $z$  – Indicator of nicotine patch

## Example: Schooling and Test Scores

Burde & Linden (2013, AEJ Applied)

RCT in Afghanistan: 32 villages divided into 11 clusters. Randomly choose 6 and set up school in each village of these clusters.

- ▶  $y$  – Child's score on math and language test
- ▶  $T^*$  – Child's true school attendance
- ▶  $T$  – Parent's report of child's school attendance
- ▶  $\mathbf{x}$  – Child and household characteristics
- ▶  $z$  – School built in village

## Example: Returns to Schooling

Oreopoulos (2006, AER)

Fuzzy RD: minimum school-leaving age in UK increased from 14 to 15 in 1947 but some already stayed until 15 before the law and others failed to comply after it.

- ▶  $y$  – Log wage
- ▶  $T^*$  – School attendance at age 15
- ▶  $T$  – Self-report of school attendance at age 15
- ▶  $x$  – Individual characteristics
- ▶  $z$  – Indicator: born in or after 1933

# Related Literature

## Continuous Treatment

Lewbel (1997, 2012), Schennach (2004, 2007), Chen et al. (2005), Hu & Schennach (2008), Song (2015), Hu et al. (2015)...

## Binary, Exogenous Treatment

Aigner (1973), Bollinger (1996), Kane et al. (1999), Black et al. (2000), Frazis & Loewenstein (2003), Mahajan (2006), Lewbel (2007)

## Binary, Endogenous Treatment

**Mahajan (2006)**, Shiu (2015), Ura (2015)

► Mahajan Details

Model:  $y = c + \beta T^* + \varepsilon$

Valid Instrument

$$\mathbb{E}[\varepsilon|z] = 0.$$

First-stage

$$p_k^* \equiv \mathbb{P}(T^* = 1|z = z_k) \neq \mathbb{P}(T^* = 1|z = z_\ell) \equiv p_\ell^*, \quad k \neq \ell$$

Non-differential Measurement Error

- ▶  $\mathbb{E}[\varepsilon|T^*, T, z] = \mathbb{E}[\varepsilon|T^*, z]$
- ▶  $\alpha_0 = \mathbb{P}(T = 1|T^* = 0, z)$
- ▶  $\alpha_1 = \mathbb{P}(T = 0|T^* = 1, z)$
- ▶  $\alpha_0 + \alpha_1 < 1$

Observable Moments:  $y = c + \beta T^* + \varepsilon$

	$z = 1$	$z = 2$	$\dots$	$z = K$
$T = 0$	$\bar{y}_{01}$ $p_{01}$	$\bar{y}_{02}$ $p_{02}$	$\dots$	$\bar{y}_{0K}$ $p_{0K}$
$T = 1$	$\bar{y}_{11}$ $p_{11}$	$\bar{y}_{12}$ $p_{12}$	$\dots$	$\bar{y}_{1K}$ $p_{1K}$

$$\bar{y}_{tk} = \mathbb{E}[y | T = t, z = z_k], \quad p_{tk} = q_k p_k$$

$$q_k = \mathbb{P}(z = z_k), \quad p_k = \mathbb{P}(T = 1 | z = z_k)$$



## Unobservable Moments: $y = \beta T^* + u$

Define error term that absorbs constant:  $u = c + \varepsilon$

	$z = 1$	$z = 2$	$\dots$	$z = K$
$T^* = 0$	$m_{01}^*$ $p_{01}^*$	$m_{02}^*$ $p_{02}^*$	$\dots$	$m_{0K}^*$ $p_{0K}^*$
$T^* = 1$	$m_{11}^*$ $p_{11}^*$	$m_{12}^*$ $p_{12}^*$	$\dots$	$m_{1K}^*$ $p_{1K}^*$

$$m_{tk}^* = \mathbb{E}[u | T^* = t, z = z_k], \quad p_{tk}^* = q_k p_k^*$$

$$q_k = \mathbb{P}(z = z_k), \quad p_k^* = \mathbb{P}(T^* = 1 | z = z_k)$$

## Possible Restrictions On $m_{tk}^*$

Joint Exogeneity:  $\mathbb{E}[\varepsilon | T^*, z] = 0$

$$\implies m_{tk}^* = c \quad \text{for all } t, k$$

Exogenous Treatment:  $\mathbb{E}[\varepsilon | T^*] = 0$

$$\implies \frac{1}{\mathbb{P}(T^* = t)} \sum_k p_{tk}^* m_{tk}^* = c \quad \text{for all } t$$

Exogenous Instrument:  $\mathbb{E}[\varepsilon | z] = 0$

$$\implies (1 - p_k^*) m_{0k}^* + p_k^* m_{1k}^* = c \quad \text{for all } k$$

## System of Equations given $E[\varepsilon|z] = 0$

$\mathbb{E}[\varepsilon|z] = 0 \implies$  *pair of equations for each  $k = 1, \dots, K$*

$$\hat{y}_{0k} = \alpha_1(p_k - \alpha_0) \left( \frac{\beta}{1 - \alpha_0 - \alpha_1} \right) + (1 - \alpha_0)c - (p_k - \alpha_0)m_{1k}^*$$

$$\hat{y}_{1k} = (1 - \alpha_1)(p_k - \alpha_0) \left( \frac{\beta}{1 - \alpha_0 - \alpha_1} \right) + \alpha_0 c + (p_k - \alpha_0)m_{1k}^*$$

where  $\hat{y}_{0k} = (1 - p_k)\bar{y}_{0k}$  and  $\hat{y}_{1k} = p_k\bar{y}_{1k}$

**2K Equations in  $K + 4$  Unknowns**

$\beta$  is unidentified regardless of  $K$ .

Proof of special case:  $\alpha_0 = 0$

1. System of equations:

$$\tilde{y}_{0k} = c + p_k \left( \frac{\beta \alpha_1}{1 - \alpha_1} \right) - p_k m_{1k}^*$$

$$\tilde{y}_{1k} = p_k \beta + p_k m_{1k}^*$$

2.  $\beta/(1 - \alpha_1) \equiv \beta_{IV}$  identified,  $\beta \alpha_1/(1 - \alpha_1) = \beta_{IV} - \beta \implies$

$$(c + p_k \beta_{IV} - \tilde{y}_{0k})/p_k = \beta + m_{1k}^*$$

$$\tilde{y}_{1k}/p_k = \beta + m_{1k}^*$$

3. Sum equations from 1.  $\implies (c + p_k \beta_{IV} - \tilde{y}_{0k}) = \tilde{y}_{1k}$

# Bounds for Mis-classification Probabilities

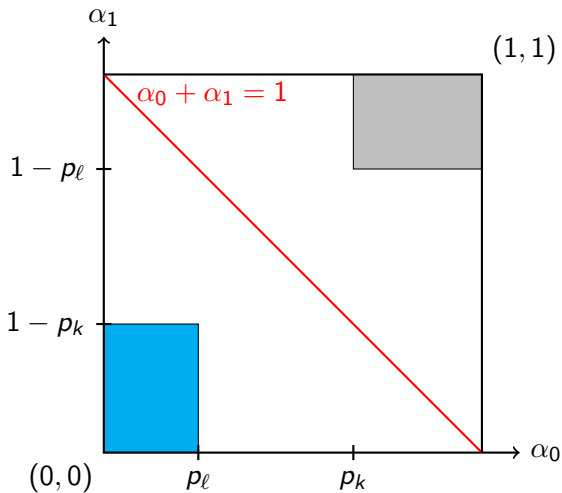
$$\alpha_0(z) = \alpha_0, \alpha_1(z) = \alpha_1$$

$$\implies p_k^* = \frac{p_k - \alpha_0}{1 - \alpha_0 - \alpha_1}, \quad 1 - p_k^* = \frac{1 - p_k - \alpha_1}{1 - \alpha_0 - \alpha_1}$$

$$\alpha_0 + \alpha_1 < 1 \iff \text{Cor}(T, T^*) > 0 \iff (1 - \alpha_0 - \alpha_1) > 0$$

$$\alpha_0 < \min_k \{p_k\}, \alpha_1 < \min_k \{1 - p_k\}$$

$$\alpha_0 \leq \min_k \{p_k\}, \quad \alpha_1 \leq \min_k \{1 - p_k\}$$



## Bounds for $\beta$

$$\mathbb{E}[\varepsilon|z] = 0$$

$$\implies \beta_{RF} = \mathbb{E}[y|z_k] - \mathbb{E}[y|z_\ell] = \beta(p_k^* - p_\ell^*)$$

### Mis-classification

$$\implies p_k^* - p_\ell^* = (p_k - p_\ell)/(1 - \alpha_0 - \alpha_1)$$

$$\text{Combining: } \beta_{IV} = \beta/(1 - \alpha_0 - \alpha_1)$$

$$\alpha_0 + \alpha_1 < 1 \implies$$

- ▶  $\beta$  is between  $\beta_{RF}$  and  $\beta_{IV}$
- ▶  $\beta_{IV}$  *inflated* but has correct sign
- ▶  $\beta_{RF}$  bound equivalent to substituting  $\alpha_0, \alpha_1$  bounds

# Strengthening the Measurement Error Assumptions

- ▶  $\alpha_0 = \mathbb{P}(T = 1 | T^* = 0, z)$
- ▶  $\alpha_1 = \mathbb{P}(T = 0 | T^* = 1, z)$
- ▶  $\alpha_0 + \alpha_1 < 1$
- ▶  $\mathbb{E}[\varepsilon | T^*, T, z] = \mathbb{E}[\varepsilon | T^*, z]$

## Additional Assumption

$$\mathbb{E}[\varepsilon^2 | T^*, T, z] = \mathbb{E}[\varepsilon^2 | T^*, z]$$

Improve bounds for  $\alpha_0, \alpha_1$  to tighten lower bound for  $\beta \dots$



# Tighter Bounds for $\alpha_0, \alpha_1$ from Conditional Variances

Assume

$$\mathbb{E}[\varepsilon^2 | T^*, T, z] = \mathbb{E}[\varepsilon^2 | T^*, z]$$

Observables

$$\sigma_{tk}^2 = \text{Var}(y | T = t, z = k)$$

Constrain Unobservables

$$s_{tk}^{*2} = \text{Var}(u | T^* = t, z_k) > 0$$

$$\begin{aligned} (p_k - \alpha_0) \left[ (1 - \alpha_0)p_k\sigma_{1k}^2 - \alpha_0(1 - p_k)\sigma_{0k}^2 \right] &> \alpha_0(1 - \alpha_0)p_k(1 - p_k)(\bar{y}_{1k} - \bar{y}_{0k})^2 \\ (1 - p_k - \alpha_1) \left[ (1 - \alpha_1)(1 - p_k)\sigma_{0k}^2 - \alpha_1p_k\sigma_{1k}^2 \right] &> \alpha_1(1 - \alpha_1)p_k(1 - p_k)(\bar{y}_{1k} - \bar{y}_{0k})^2 \end{aligned}$$

# Schooling and Test Scores – Afghan RCT

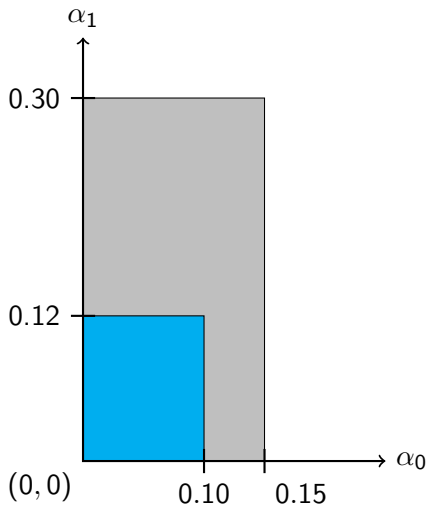
Burde & Linden (2013, AEJ Applied)

“Weak” Bounds

$$\beta \in [0.65 \times \beta_{IV}, \beta_{IV}]$$

Add 2nd Moments

$$\beta \in [0.78 \times \beta_{IV}, \beta_{IV}]$$



Independence Assumption:  $\varepsilon \perp T|(T^*, z)$

Define  $F_{tk}(\tau) = \mathbb{P}(Y \leq \tau | T = t, z_k)$  and  $F_k(\tau) = \mathbb{P}(Y \leq \tau | z_k)$

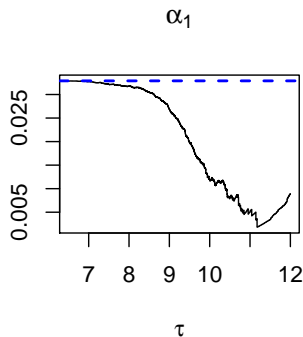
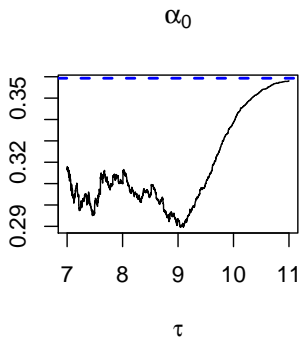
$$\alpha_0 \leq p_k \inf_{\tau} \left\{ \left[ \frac{F_{1k}(\tau)}{F_k(\tau)} \right] \wedge \left[ \frac{1 - F_{1k}(\tau)}{1 - F_k(\tau)} \right] \right\} \leq p_k$$

$$\alpha_1 \leq (1 - p_k) \inf_{\tau} \left\{ \left[ \frac{F_{0k}(\tau)}{F_k(\tau)} \right] \wedge \left[ \frac{1 - F_{0k}(\tau)}{1 - F_k(\tau)} \right] \right\} \leq (1 - p_k)$$

Bounds for  $(\alpha_0, \alpha_1)$  do *not* require  $z$  to be a valid instrument!

# Upper Bounds for Mis-Classification Rates

Returns to Schooling Example: Oreopoulos (2006)



# Sufficient Conditions To Identify $\alpha_0, \alpha_1$ , and $\beta$

## Baseline Assumptions

- ▶  $\mathbb{E}[\varepsilon|z] = 0$
- ▶  $\mathbb{E}[\varepsilon|T^*, T, z] = \mathbb{E}[\varepsilon|T^*, z]$
- ▶  $\alpha_0 = \mathbb{P}(T = 1|T^* = 0, z)$ ,  $\alpha_1 = \mathbb{P}(T = 0|T^* = 1, z)$ ,  $\alpha_0 + \alpha_1 < 1$

## Strengthen IV Assumption

- ▶  $\mathbb{E}[\varepsilon^2|z] = \mathbb{E}[\varepsilon^2]$
- ▶  $\mathbb{E}[\varepsilon^3|z] = \mathbb{E}[\varepsilon^3]$

## Strengthen Measurement Error Assumption

- ▶  $\mathbb{E}[\varepsilon^2|T^*, T, z] = \mathbb{E}[\varepsilon^2|T^*, z]$
- ▶  $\mathbb{E}[\varepsilon^3|T^*, T, z] = \mathbb{E}[\varepsilon^3|T^*, z]$

# First Moment Condition

## Assumptions

- ▶  $\mathbb{E}[\varepsilon|z] = 0$
- ▶  $\mathbb{E}[\varepsilon|T^*, T, z] = \mathbb{E}[\varepsilon|T^*, z]$
- ▶  $\alpha_0 = \mathbb{P}(T = 1|T^* = 0, z)$
- ▶  $\alpha_1 = \mathbb{P}(T = 0|T^* = 1, z)$

## Moment Condition

$$\text{Cov}(y, z) - \left( \frac{\beta}{1 - \alpha_0 - \alpha_1} \right) \text{Cov}(T, z) = 0$$

MC # 1 identifies  $\beta/(1 - \alpha_0 - \alpha_1)$

# Second Moment Condition

## Additional Assumptions

- ▶  $\mathbb{E}[\varepsilon^2|z] = \mathbb{E}[\varepsilon^2]$
- ▶  $\mathbb{E}[\varepsilon^2|T^*, T, z] = \mathbb{E}[\varepsilon|T^*, z]$

## Moment Condition

$$\text{Cov}(y^2, z) - \frac{\beta}{1 - \alpha_0 - \alpha_1} \left\{ 2\text{Cov}(yT, z) - \beta\text{Cov}(T, z) \left( \frac{1 + \alpha_0 - \alpha_1}{1 - \alpha_0 - \alpha_1} \right) \right\} = 0$$

Given MC #1, MC #2 identifies  $(\alpha_1 - \alpha_0)$

# Third Moment Condition

## Additional Assumptions

- ▶  $\mathbb{E}[\varepsilon^3|z] = \mathbb{E}[\varepsilon^2]$
- ▶  $\mathbb{E}[\varepsilon^3|T^*, T, z] = \mathbb{E}[\varepsilon|T^*, z]$

## Moment Condition

$$\text{Cov}(y^3, z) - \left( \frac{\beta}{1 - \alpha_0 - \alpha_1} \right) \left\{ \beta^2 \left[ 1 + \frac{6\alpha_0(1 - \alpha_1)}{(1 - \alpha_0 - \alpha_1)^2} \right] \text{Cov}(T, z) \right. \\ \left. - 3\beta \left[ \frac{1 - (\alpha_1 - \alpha_0)}{1 - \alpha_0 - \alpha_1} \right] \text{Cov}(yT, z) + 3\text{Cov}(y^2 T, z) \right\} = 0$$



# Sketch of Identification Argument

Very tedious algebra. . .

1. Use 1st MC to eliminate  $\beta/(1 - \alpha_0 - \alpha_1)$  from others
2. Use 2nd MC to solve for  $\alpha_1$  in terms of  $\alpha_0$
3. 3rd MC becomes a quadratic in  $(1 - \alpha_1)$  and observables only.
4. The quadratic always has two real roots:  $(1 - \alpha_1)$  and  $\alpha_0$ .
5. To tell which root is which, use  $\alpha_0 + \alpha_1 < 1$ .
6. Calculate  $\alpha_0 + \alpha_1$  and substitute into 1st MC to obtain  $\beta$ .

Unfortunately, identification of  $\alpha_0, \alpha_1$  fails if  $\beta = 0$ . . .

## Simple Special Case: $\alpha_0 = 0$

$$\text{Cov}(y, z) - \left( \frac{\beta}{1 - \alpha_0} \right) \text{Cov}(T, z) = 0$$

$$\text{Cov}(y^2, z) - \frac{\beta}{1 - \alpha_0} \{2\text{Cov}(yT, z) - \beta\text{Cov}(T, z)\} = 0$$

$$\beta = \frac{2\text{Cov}(yT, z)}{\text{Cov}(T, z)} - \frac{\text{Cov}(y^2, z)}{\text{Cov}(y, z)}$$

## Simulation Example: $y = \beta T^* + \varepsilon$

### Errors

$(\varepsilon, \eta) \sim$  jointly normal, mean 0, variance 1, correlation 0.5.

### First-Stage

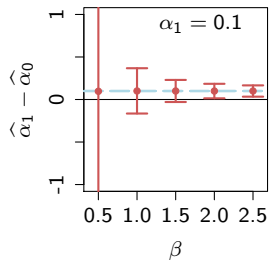
- ▶ Half of subjects have  $z = 1$ , the rest have  $z = 0$ .
- ▶  $T^* = \mathbf{1}\{\gamma_0 + \gamma_1 z + \eta > 0\}$
- ▶  $\delta = \mathbb{P}(T^* = 0|z = 1) = \mathbb{P}(T^* = 1|z = 0)$

### Mis-classification

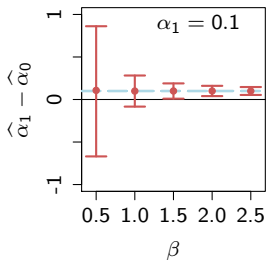
- ▶ Set  $\alpha_0 = 0$  so  $T^* = 0 \implies T = 0$
- ▶  $T|T^* = 1 \sim \text{Bernoulli}(1 - \alpha_1)$

Sampling Distribution of  $\hat{\alpha}_1 - \hat{\alpha}_0$

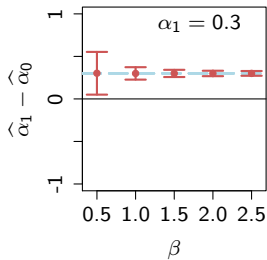
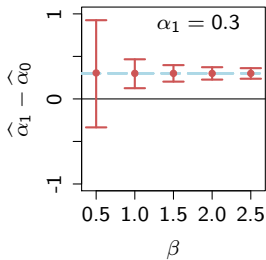
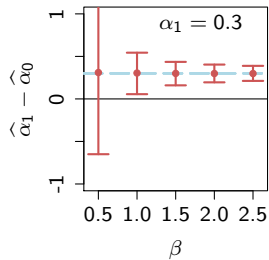
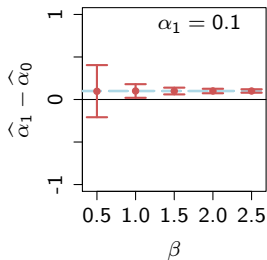
(a)  $N = 500, \delta = 0.1$



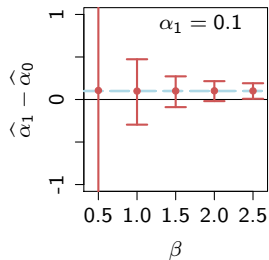
(b)  $N = 1000, \delta = 0.1$



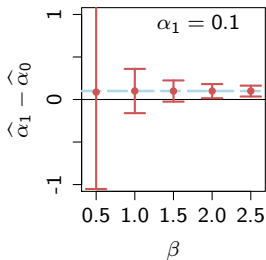
(c)  $N = 5000, \delta = 0.1$



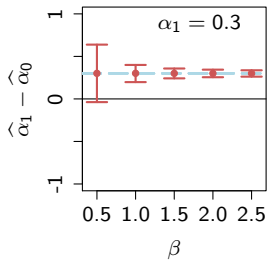
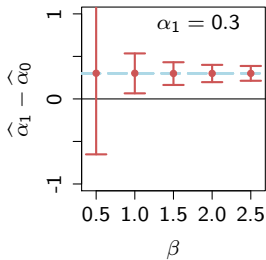
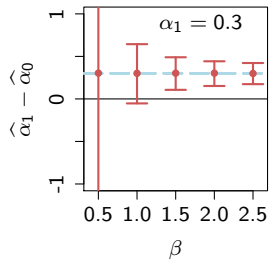
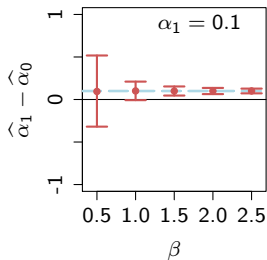
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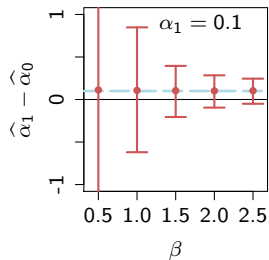
(b)  $N = 1000, \delta = 0.2$



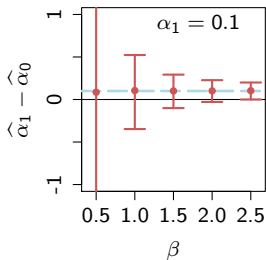
(c)  $N = 5000, \delta = 0.2$



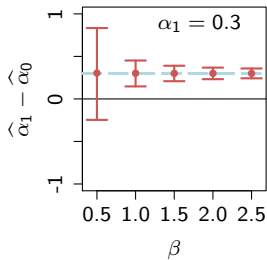
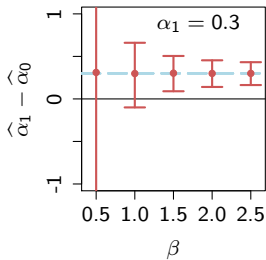
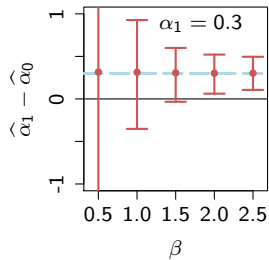
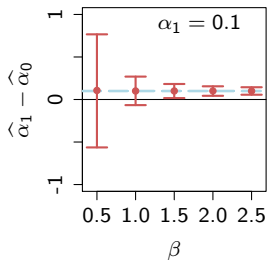
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(b)  $N = 1000, \delta = 0.3$



(c)  $N = 5000, \delta = 0.3$



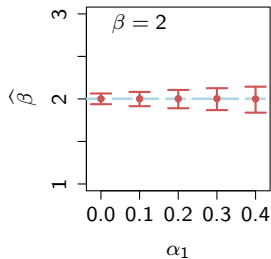
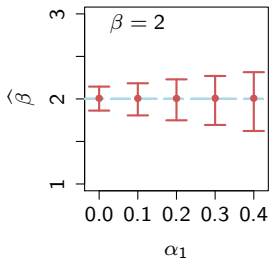
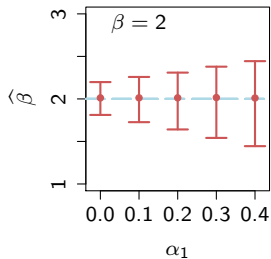
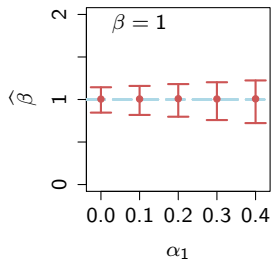
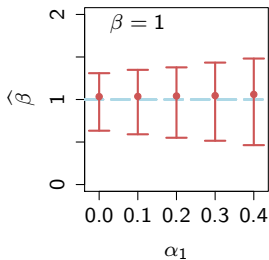
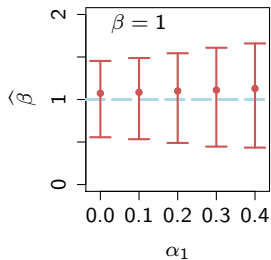
Sampling Distribution of  $\hat{\beta} = (1 - \hat{\alpha}_0 - \hat{\alpha}_1)\hat{\beta}_{IV}$



(a)  $N = 500, \delta = 0.1$

(b)  $N = 1000, \delta = 0.1$

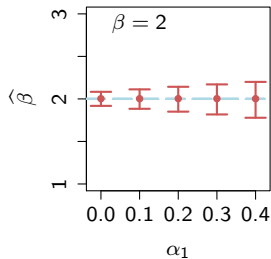
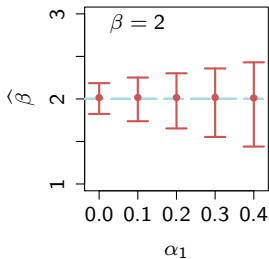
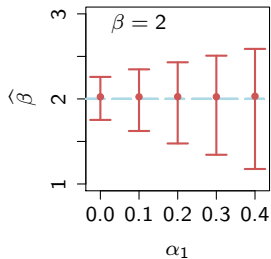
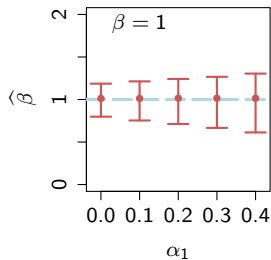
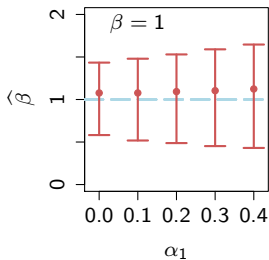
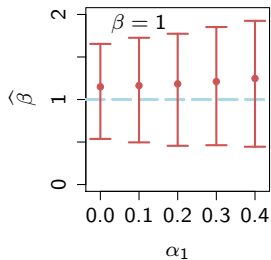
(c)  $N = 5000, \delta = 0.1$



(a)  $N = 500, \delta = 0.2$

(b)  $N = 1000, \delta = 0.2$

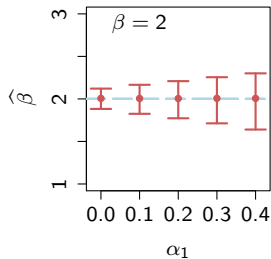
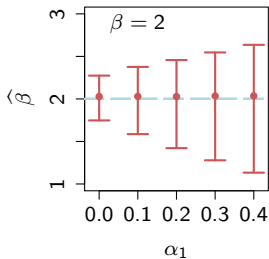
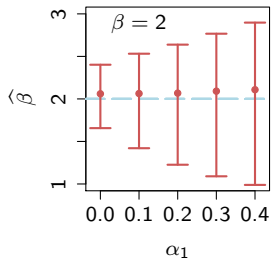
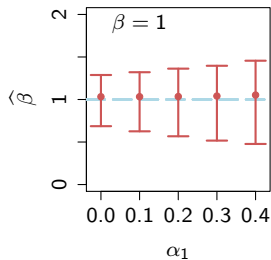
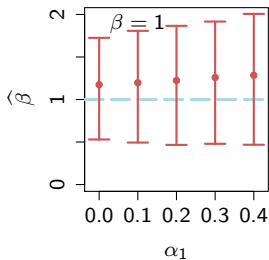
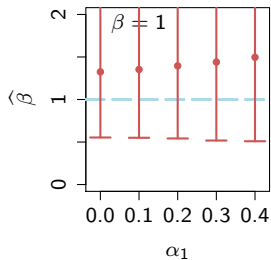
(c)  $N = 5000, \delta = 0.2$



(a)  $N = 500, \delta = 0.3$

(b)  $N = 1000, \delta = 0.3$

(c)  $N = 5000, \delta = 0.3$



$(z \perp \varepsilon)$  and  $(T \perp \varepsilon | T^*, z) \Rightarrow$  Continuum of MCs

## Characteristic Functions

$$e^{i\omega\beta} [(1 - \alpha_1) - \xi(\omega)] = \alpha_0 - \xi(\omega)$$

$$\xi(\omega) \equiv \frac{\varphi_k(\omega) - \varphi_\ell(\omega)}{p_k \varphi_{1k}(\omega) - p_\ell \varphi_{1\ell}(\omega)}$$

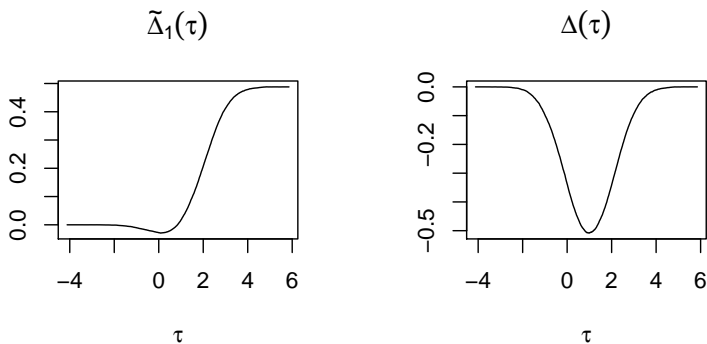
## Distribution Functions

$$\tilde{\Delta}_1(\tau + \beta) - \tilde{\Delta}_1(\tau) = \alpha_0 \Delta(\tau + \beta) - (1 - \alpha_1) \Delta(\tau)$$

$$\Delta(\tau) = F_k(\tau) - F_\ell(\tau)$$

$$\tilde{\Delta}_1(\tau) = p_k F_{1k}(\tau) - p_\ell F_{1\ell}(\tau)$$

## CDF Conditions for Simulation DGP



$$\tilde{\Delta}_1(\tau + \beta) - \tilde{\Delta}_1(\tau) = \alpha_0 \Delta(\tau + \beta) - (1 - \alpha_1) \Delta(\tau)$$

# Conclusion

## Summary

- ▶ Endogenous, mis-measured binary treatment.
- ▶ Important in applied work but no solution in the literature.
- ▶ Usual (1st moment) IV assumption fails to identify  $\beta$
- ▶ Bounds for mis-classification probabilities and  $\beta$ .
- ▶ Higher moment / independence restrictions identify  $\beta$

## Extensions / Work in Progress

- ▶ Efficient estimation w/ continuum of MCs
- ▶ Inference / Specification Testing
- ▶ LATE interpretation

# Mahajan (2006, ECTA)

## Regression Model

$$y = \mathbb{E}[y|T^*] + \nu$$

$$\mathbb{E}[\nu|T^*] = 0 \text{ by construction}$$

## Causal Model

$$y = c + \beta T^* + \varepsilon$$

$$\mathbb{E}[\varepsilon|T^*] \neq 0$$

## Main Result (Correct) – Exogenous Treatment

Relevant binary instrument  $z$  ( $p_k^* \neq p_\ell^*$ ) identifies  $\alpha_0, \alpha_1$  and  $\mathbb{E}[y|T^*]$  provided that  $\mathbb{E}[\nu|T^*, T, z] = 0$  and  $\alpha_0 + \alpha_1 < 1$ .

## Extension (Incorrect) – Endogenous Treatment

$$\mathbb{E}[\varepsilon|z] = 0, p_k^* \neq p_\ell^*, \mathbb{E}[\varepsilon|T, T^*, z] = \mathbb{E}[\varepsilon|T^*] \implies \beta \text{ identified.}$$

# Mahajan (2006, ECTA)

## Regression Model

$$y = \mathbb{E}[y | T^*] + \nu$$

$$\mathbb{E}[\nu | T^*] = 0 \text{ by construction}$$

## Causal Model

$$y = c + \beta T^* + \varepsilon$$

$$\mathbb{E}[\varepsilon | T^*] \neq 0$$

## Ingredients

1. If  $p_k^* \neq p_\ell^*$ ,  $\mathbb{E}[\varepsilon | z] = 0$  then, since  $\beta_{IV} = \beta / (1 - \alpha_0 - \alpha_1)$ , knowledge of  $\alpha_0, \alpha_1$  is sufficient to recover  $\beta$ . (Correct)
2. If  $p_k^* \neq p_\ell^*$ ,  $\mathbb{E}[\nu | T^*, T, z] = 0$ ,  $\alpha_0, \alpha_1$  are identified. (Correct)

How to satisfy both 1 and 2 while allowing  $\mathbb{E}[\varepsilon | T^*] \neq 0$ ?

3. Assume that  $\mathbb{E}[\varepsilon | T^*, T, z] = \mathbb{E}[\varepsilon | T^*]$   
(i.e.  $m_{0k}^* = m_{0\ell}^*$  and  $m_{1k}^* = m_{1\ell}^*$ )



# Flaw in the Argument

## Proposition

If  $\mathbb{E}[\varepsilon | T^*] \neq 0$  then  $\mathbb{E}[\varepsilon | T^*, T, z] = \mathbb{E}[\varepsilon | T^*]$  combined with  $\mathbb{E}[\varepsilon | z] = 0$  implies  $p_k^* = p_\ell^*$ , i.e.  $z$  is irrelevant for  $T^*$ .

## Proof

$\mathbb{E}[\varepsilon | z] = 0$  implies

$$(1 - p_1^*)m_{0k}^* + p_1^*m_{1k}^* = c$$

$$(1 - p_2^*)m_{0\ell}^* + p_2^*m_{1\ell}^* = c$$

while Mahajan's assumption implies  $m_{0k}^* = m_{0\ell}^*$  and  $m_{1k}^* = m_{1\ell}^*$ .

Therefore either  $m_{0k}^* = m_{0\ell}^* = m_{1k}^* = m_{1\ell}^* = c$ , which is ruled out by  $E[\varepsilon | T^*] = 0$ , or  $p_k^* = p_\ell^*$ .