Estimating the Effect of a Mis-measured, Endogenous, Binary Treatment

Francis J. DiTraglia Camilo Garcia-Jimeno

University of Pennsylvania

January 19th, 2016

What is the causal effect of T^* ?

$$y_i = h(T_i^*, \mathbf{x}_i) + \varepsilon_i$$

- ▶ y − Outcome of interest
- ▶ h Unknown function that does not depend on i
- ▶ T* Unobserved, endogenous binary treatment
- ➤ T Observed, mis-measured binary surrogate for T*
- ▶ x − Exogenous covariates
- \triangleright ε Mean-zero error term
- ➤ z Discrete instrumental variable

Example 1: Smoking and Birthweight (SNAP Trial)

Coleman et al. (N Engl J Med, 2012)

RCT with 1050 pregnant smokers in England: 521 given nicotine patches, the rest given placebo patches.

- ▶ y Birthweight
- ▶ T* True smoking behavior
- ► T Self-reported smoking behavior
- x Mother characteristics
- z Indicator of nicotine patch

Example 2: Schooling and Test Scores

Burde & Linden (2013, AEJ Applied)

RCT in Afghanistan: 32 villages divided into 11 clusters. Randomly choose 6 and build a school in each village of these clusters.

- ▶ y Child's score on math and language test
- ► T* Child's true school attendance
- ➤ T Parent's report of child's school attendance
- x Child and household characteristics
- ▶ z School built in village

Example 3: Job Training Partnership Act (JPTA)

Heckman et al. (2000, QJE)

Randomized offer of job training, but about 30% of those *not* offered also obtain training and about 40% of those offered training don't attend. Estimate causal effect of *training* rather than *offer* of training.

- y − Log wage
- ▶ T* True training attendence
- ➤ T Self-reported training attendance
- x Individual characteristics
- \triangleright z Offer of job training

Related Literature

Continuous Treatment

Lewbel (1997, 2012), Schennach (2004, 2007), Chen et al. (2005), Hu & Schennach (2008)...

Binary, Exogenous Treatment

Aigner (1973), Bollinger (1996), Kane et al. (1999), Black et al. (2000), Frazis & Loewenstein (2003), Mahajan (2006), Lewbel (2007)

Binary, Endogenous Treatment

Only existing result in Mahajan (2006)

Model: $y = c + \beta T^* + \varepsilon$

First-stage

$$ho_k^* \equiv \mathbb{P}(T^* = 1|z = z_k)
eq \mathbb{P}(T^* = 1|z = z_\ell) \equiv
ho_\ell^*, \ k
eq \ell$$

Measurement Error

- ▶ Non-differential: $\mathbb{E}[\varepsilon|T^*, T, z] = \mathbb{E}[\varepsilon|T^*, z]$
- Does not depend on z:

$$\alpha_0 = \mathbb{P}(T = 1 | T^* = 0, z)$$

$$\alpha_1 = \mathbb{P}(T=0|T^*=1,z)$$

Notation

Define error term that absorbs constant: $u = c + \varepsilon$

Observable Moments: $y = \beta T^* + u$

$$ar{y}_{tk} = \mathbb{E}[y|T=t,z=z_k], \quad p_{tk} = q_k p_k$$
 $q_k = \mathbb{P}(z=z_k), \quad p_k = \mathbb{P}(T=1|z=z_k)$

Unobservable Moments: $y = \beta T^* + u$

$$m_{tk}^* = \mathbb{E}[u|T^* = t, z = z_k], \quad p_{tk}^* = q_k p_k^*$$

$$p_k^* = \mathbb{P}(T^* = 1|z = z_k) = \frac{p_k - \alpha_0}{1 - \alpha_0 - \alpha_1}$$

Possible Assumptions On m_{tk}^*

Joint Exogeneity:
$$\mathbb{E}[\varepsilon|T^*,z]=0$$
 $\Rightarrow m_{tk}^*=c \quad \text{for all } t,k$

Exogenous Treatment: $\mathbb{E}[\varepsilon|T^*]=0$
 $\Rightarrow \frac{1}{\mathbb{P}(T^*=t)}\sum_k p_{tk}^* m_{tk}^*=c \quad \text{for all } t$

Exogenous Instrument: $\mathbb{E}[\varepsilon|z]=0$
 $\Rightarrow (1-p_k^*)m_{0k}^*+p_k^*m_{1k}^*=c \quad \text{for all } k$

Mahajan (2006, Econometrica)

$$y = \mathbb{E}[y|T^*] + \nu$$

$$\mathbb{E}[
u|T^*] = 0$$
 by construction

Causal Model

$$y = c + \beta T^* + \varepsilon$$

$$\mathbb{E}[\varepsilon|T^*]\neq 0$$

Main Result (Correct) – Exogenous Treatment

Relevant binary instrument z $(p_1^* \neq p_2^*)$ identifies α_0, α_1 and

$$\mathbb{E}[y|T^*]$$
 provided that $\mathbb{E}[\nu|T^*, T, z] = 0$.

Mahajan (2006, Econometrica)

$$y = \mathbb{E}[y|T^*] + \nu$$

$$\mathbb{E}[
u|T^*]=0$$
 by construction

Causal Model

$$y = c + \beta T^* + \varepsilon$$

$$\mathbb{E}[arepsilon|T^*]
eq 0$$

Additional Result (Incorrect) – Endogenous Treatment

$$\mathbb{E}[\varepsilon|z] = 0, \ \rho_1^* \neq \rho_2^*, \ \mathbb{E}[\varepsilon|T,T^*,z] = \mathbb{E}[\varepsilon|T^*] \implies \beta \ \text{identified}.$$

Mahajan's Argument

Regression Model

$$y = \mathbb{E}[y|T^*] + \nu$$

$$\mathbb{E}[\nu|T^*] = 0$$
 by construction

Causal Model

$$y = c + \beta T^* + \varepsilon$$

$$\mathbb{E}[\varepsilon|T^*]\neq 0$$

Ingredients

- 1. If $p_1^* \neq p_2^*$, $\mathbb{E}[\varepsilon|z] = 0$ then, since $\beta_{IV} = \beta/(1 \alpha_0 \alpha_1)$, knowledge of α_0, α_1 is sufficient to recover β . (Correct)
- 2. If $p_1^* \neq p_2^*$, $\mathbb{E}[\nu|T^*,T,z]=0$, α_0,α_1 are identified. (Correct) How to satisfy both 1 and 2 while allowing $\mathbb{E}[\varepsilon|T^*]\neq 0$?
- 3. Assume that $\mathbb{E}[arepsilon|T^*,T,z]=\mathbb{E}[arepsilon|T^*]$ (i.e. $m_{01}^*=m_{02}^*$ and $m_{11}^*=m_{12}^*$)

Mahajan's Argument

Proposition

If $\mathbb{E}[\varepsilon|T^*] \neq 0$ then $\mathbb{E}[\varepsilon|T^*, T, z] = \mathbb{E}[\varepsilon|T^*]$ combined with $\mathbb{E}[\varepsilon|z] = 0$ implies $p_1^* = p_2^*$, i.e. z is irrelevant for T^* .

Proof

Recall that
$$\mathbb{E}[arepsilon|z]=0$$
 implies
$$(1-p_1^*)m_{01}^*+p_1^*m_{11}^*=c$$

$$(1-p_2^*)m_{02}^*+p_2^*m_{12}^*=c$$

while Mahajan's assumption implies $m_{01}^*=m_{02}^*$ and $m_{11}^*=m_{12}^*$. Therefore either $m_{01}^*=m_{02}^*=m_{11}^*=m_{12}^*=c$, which is ruled out by $E[\varepsilon|T^*]=0$, or $p_1^*=p_2^*$.

What about increasing the support of z?

$$\mathbb{E}[arepsilon|z] = 0 \implies \mathit{pair} \ \mathsf{of} \ \mathsf{equations} \ \mathsf{for} \ \mathsf{each} \ k = 1, \dots, K$$

$$\begin{split} \hat{y}_{0k} &= \alpha_1(p_k - \alpha_0) \left(\frac{\beta}{1 - \alpha_0 - \alpha_1}\right) + (1 - \alpha_0)c - (p_k - \alpha_0)m_{1k}^* \\ \hat{y}_{1k} &= (1 - \alpha_1)(p_k - \alpha_0) \left(\frac{\beta}{1 - \alpha_0 - \alpha_1}\right) + \alpha_0c + (p_k - \alpha_0)m_{1k}^* \\ \text{where } \hat{y}_{0k} &= (1 - p_k)\bar{y}_{0k} \text{ and } \hat{y}_{0k} = p_k\bar{y}_{1k} \end{split}$$

2K Equations in K + 4 Unknowns

Theorem: β is undentified regardless of K.

(For general case, see paper.)

Proof of special case: $\alpha_0 = 0$

1. System of equations simplifies to

$$\hat{y}_{0k} = c + p_k \beta \left(\frac{\alpha_1}{1 - \alpha_1}\right) - p_k m_{1k}^*$$

$$\hat{y}_{1k} = p_k \beta + p_k m_{1k}^*$$

2. $\beta/(1-\alpha_1) \equiv \mathcal{W}$ is identified and imposing this, algebra gives $\beta \alpha_1/(1-\alpha_1) = \mathcal{W} - \beta$.

Theorem: β is undentified regardless of K.

(For general case, see paper.)

Proof of special case: $\alpha_0 = 0$ continued...

3. Substituting:

$$(c + p_k W - \hat{y}_{0k})/p_k = \beta + m_{1k}^*$$

 $\hat{y}_{1k}/p_k = \beta + m_{1k}^*$

- 4. Linear system in (β, m_{1k}^*) no solution or ∞ of solutions.
- 5. Sum original pair of equations $\implies c + p_k W \hat{y}_{0k} = \hat{y}_{1k}$ thus ∞ of solutions. The model is unidentified.

Conditional Second Moment Independence.

New Assumption

Homoskedastic errors w.r.t. the *instrument*: $E[\varepsilon^2|z] = E[\varepsilon^2]$

Not Crazy!

Holds in an RCT or a true natural experiment.

New Moment Conditions

Defining
$$\mu_{k\ell}^* = (p_k - \alpha_0)m_{1k}^* - (p_\ell - \alpha_0)m_{k\ell}^*$$
,

$$\mathbb{E}(y^2|z_k) - \mathbb{E}(y^2|z_\ell) \equiv \Delta \overline{y^2} = \beta \mathcal{W}(p_k - p_\ell) + 2\mathcal{W}\mu_{k\ell}^*$$

$$\mathbb{E}(yT|z_k) - \mathbb{E}(yT|z_\ell) \equiv \Delta \overline{yT} = (1 - \alpha_1)\mathcal{W}(p_k - p_\ell) + \mu_{k\ell}^*$$

Theorem: $(\alpha_1 - \alpha_0)$ is Identified.

(Requires only binary z)

Proof

$$\Delta \overline{y^2} = \beta \mathcal{W}(p_k - p_\ell) + 2\mathcal{W}\mu_{k\ell}^*$$

$$\Delta \overline{yT} = (1 - \alpha_1)\mathcal{W}(p_k - p_\ell) + \mu_{k\ell}^*$$

Solve for $\mu_{k\ell}^*$, substitute and rearrange:

$$\mathcal{R} \equiv \beta - 2(1 - \alpha_1)\mathcal{W} = \frac{\Delta \overline{y^2} - 2\mathcal{W}\Delta \overline{yT}}{\mathcal{W}(p_k - p_\ell)}.$$

Rearrange and substitute $\beta = \mathcal{W}(1 - \alpha_0 - \alpha_1)$ to find

$$\alpha_1 - \alpha_0 = 1 + \mathcal{R}/\mathcal{W}.$$

What Good is $(\alpha_1 - \alpha_0)$?

- ▶ Test a necessary condition for *no mis-classification*: $\alpha_0 = \alpha_1$
- ▶ Simple, tighter partial identification bounds for β
- In some settings, one of the mis-classification probabilities is known to be zero $\implies \beta$ point identified

Conditional Third Moment Independence

New Assumption

Third moment independence w.r.t instrument: $E[\varepsilon^3|z] = E[\varepsilon^3]$

New Moment Conditions

Define
$$\lambda_{k\ell}^* = (p_k - \alpha_0)v_{1k}^* - (p_\ell - \alpha_0)v_{1\ell}^*$$

where $v_{tk}^* = \mathbb{E}(u^2|T^* = t, z_k)$. Then

$$\begin{split} \mathbb{E}(y^3|z_k) &- \mathbb{E}(y^3|z_\ell) \equiv \\ \Delta \overline{y^3} &= \beta^2 \mathcal{W}(p_k - p_\ell) + 3\beta \mathcal{W} \mu_{k\ell}^* + 3\mathcal{W} \lambda_{k\ell}^* \\ \mathbb{E}(y^2 T|z_k) &- \mathbb{E}(y^2 T|z_\ell) \equiv \\ \Delta \overline{y^2 T} &= \beta(1 - \alpha_1) \mathcal{W}(p_k - p_\ell) + 2(1 - \alpha_1) \mathcal{W} \mu_{k\ell}^* + \lambda_{k\ell}^* \end{split}$$

Theorem: β , α_0 and α_1 are identified.

Requires $\alpha_0 + \alpha_1 < 1$, but z need only be binary.

Proof

$$\Delta \overline{y^3} = \beta^2 \mathcal{W}(p_k - p_\ell) + 3\beta \mathcal{W} \mu_{k\ell}^* + 3\mathcal{W} \lambda_{k\ell}^*$$

$$\Delta \overline{y^2 T} = \beta (1 - \alpha_1) \mathcal{W}(p_k - p_\ell) + 2(1 - \alpha_1) \mathcal{W} \mu_{k\ell}^* + \lambda_{k\ell}^*$$

Solve for $\lambda_{k\ell}^*$, substitute and rearrange:

$$\mathcal{S} \equiv eta^2 - 3\mathcal{W}(1-lpha_1)(eta+\mathcal{R}) = rac{\Delta\overline{y^3} - 3\mathcal{W}\left[\Delta\overline{y^2T} + \mathcal{R}\Delta\overline{yT}
ight]}{\mathcal{W}(p_k-p_\ell)}.$$

Theorem: β , α_0 and α_1 are identified.

Requires $\alpha_0 + \alpha_1 < 1$, but z need only be binary.

Proof continued...

$$\mathcal{S} \equiv \beta^2 - 3\mathcal{W}(1 - \alpha_1)(\beta + \mathcal{R}) = \frac{\Delta \overline{y^3} - 3\mathcal{W}\left[\Delta \overline{y^2 T} + \mathcal{R} \Delta \overline{y T}\right]}{\mathcal{W}(p_k - p_\ell)}$$

Use the fact that $\mathcal{R} = \beta - 2(1 - \alpha_1)\mathcal{W}$ to eliminate β from \mathcal{S} :

$$2\mathcal{W}^2(1-\alpha_1)^2 + 2\mathcal{R}\mathcal{W}(1-\alpha_1) + (\mathcal{S} - \mathcal{R}^2) = 0$$

which is a quadratic in $(1-\alpha_1)$ and observables only! Can show that there are always two real roots: one is $(1-\alpha_1)$ and the other is α_0 . To tell which is which, need $\alpha_0+\alpha_1<1$.

Recap of Results

- 1. Using first-moment information alone, β is unidentified regardless of how many values the instrument takes on.
- 2. Using second moment information $\alpha_1 \alpha_0$ is identified
 - ▶ Partial identification bound for β
 - ▶ Identifies β if α_0 is known (e.g. smoking/birthweight example)
- 3. Using third moment information β , α_0 and α_1 are identified so long as $\alpha_0 + \alpha_1 < 1$.

Simulation Study: $y = \beta T^* + \varepsilon$

- (ε, η) \sim jointly normal, mean 0, variance 1, corr. 0.3.
- ▶ First stage: $T^* = \mathbf{1} \{ \gamma_0 + \gamma_1 z + \eta > 0 \}$
 - ▶ Half of subjects have z = 1, the rest have z = 0.
 - ho $\gamma_0 = \Phi^{-1}(\delta)$
 - $\gamma_1 = \Phi^{-1}(1 \delta) \Phi(\delta)$
 - $m \delta$ equals fraction of those offered treatment who fail to take it up *and* fraction of those not offered treatment who do.
- Generate T as follows:
 - $T^*=0 \implies T=0$, i.e. $\alpha_0=0$
 - $T \mid T^* = 1 \sim \text{Bernoulli}(1 \alpha_1)$
 - $ightharpoonup \alpha_0, \alpha_1$ unknown to econometrician.

Sampling Distribution of $\hat{lpha}_1 - \hat{lpha}_0$

(a)
$$N = 500, \delta = 0.1$$
 (b) $N = 1000, \delta = 0.1$ (c) $N = 5000, \delta = 0.1$ (d) $\alpha_1 = 0.1$ (e) $\alpha_1 = 0.1$ (f) $\alpha_1 = 0.1$ (f) $\alpha_1 = 0.1$ (f) $\alpha_1 = 0.1$ (f) $\alpha_1 = 0.1$ (g) α

(a)
$$N = 500, \delta = 0.3$$
 (b) $N = 1000, \delta = 0.3$ (c) $N = 5000, \delta = 0.3$ (d) $N = 5000, \delta = 0.3$ (e) $N = 5000, \delta = 0.3$ (f) $N = 5000, \delta = 0.3$ (f) $N = 5000, \delta = 0.3$ (f) $N = 5000, \delta = 0.3$ (g) N

Sampling Distribution of $\widehat{eta}=(1-\widehat{lpha}_0-\widehat{lpha}_1)\widehat{eta}_{IV}$

(a)
$$N = 500, \delta = 0.1$$
 (b) $N = 1000, \delta = 0.1$ (c) $N = 5000, \delta = 0.1$ (d) $N = 5000, \delta = 0.1$ (e) $N = 5000, \delta = 0.1$ (f) $N = 1000, \delta = 0.1$ (f) $N = 1000, \delta = 0.1$ (g) N

(a)
$$N = 500$$
, $\delta = 0.3$ (b) $N = 1000$, $\delta = 0.3$ (c) $N = 5000$, $\delta = 0.3$ (d) $N = 5000$, $\delta = 0.3$ (e) $N = 5000$, $\delta = 0.3$ (f) $N = 5000$, $\delta = 0.3$ (f) $N = 5000$, $\delta = 0.3$ (g) $N = 0.3$ (g)

Empirical Illustration: Schooling and Test Scores

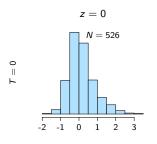
Burde & Linden (2013, AEJ Applied)

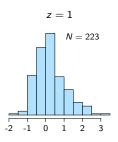
RCT in Afghanistan: 32 villages divided into 11 clusters. Randomly choose 6 and build a school in each village of these clusters (N = 1468).

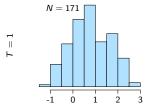
- ▶ y Child's score on math and language test
- ► T* Child's true school attendance
- T Parent's report of child's school attendance
- x Child and household characteristics
- ▶ z School built in village

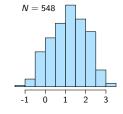
Empirical Illustration: Schooling and Test Scores

Burde & Linden (2013, AEJ Applied)







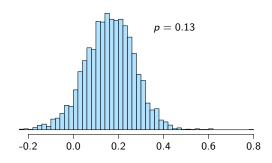


Empirical Illustration: Schooling and Test Scores

Burde & Linden (2013, AEJ Applied)

$$\widehat{\beta}_{OLS} = 0.88$$
 $\widehat{\beta}_{IV} = 1.27$
 $\widehat{\alpha}_1 - \widehat{\alpha}_0 = 0.18$

Cluster Bootstrap Distribution of $\hat{\alpha}_1 - \hat{\alpha}_0$



Conclusion

- Study causal effect of an endogenous, mis-measured, binary treatment.
- Important case in applied work, only existing result is incorrect. Identification requires going beyond first moments.
- New partial and point identification results by exploiting higher moments of outcome variable.
- Test necessary condition for absence of measurement error.
- Explore sampling distribution of our simple closed-form method of moments estimator in a simulation experiment.
- ▶ Detect evidence of measurement error in real-world example.