

Marginal Treatment Effects Part I: Background

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Treatment Effects: The Big Picture

The Best We Can Do?

- ▶ Ideally, want to learn *individual treatment effects* but we can't: fundamental problem of causal inference!
- ▶ Barring that, want to learn *distribution* of treatment effects, but we can't: fundamental problem of causal inference! ([Can bound them: Notes Chapter 3](#))
- ▶ ATE (or conditional ATE) usually considered best we can do. Identified by “gold standard” placebo controlled, randomized trial with perfect compliance.

We can't force people!

- ▶ Even when treatment is randomly assigned, can't force people to take it: *randomized encouragement design*
- ▶ Intent-to-treat (ITT) effect: causal effect of *offering* treatment. “Diluted” by people offered who don't take (typically assume exclusion restriction).

Better LATE than Nothing?

- ▶ IV allows us to go beyond ITT effects, but if treatment effects are heterogenous, we recover the LATE: average effect for *compliers*
- ▶ Is the LATE an interesting quantity? Maybe, maybe not.
- ▶ Recently: lots of interest in extrapolATE-ing “beyond LATE” to more interesting causal parameters. **That is the topic of this lecture and the next one**
- ▶ Many issues here, but most important: **what causal parameters should we be interested in and why?**

Two Key Questions

1. What is it *possible* to learn from data? (Identification)
2. What do we plan to *do* with our causal effect? (Less commonly asked)

Causal Effects are for *Decisionmaking*

Example Causal Question

- ▶ What is the causal effect of cognitive behavioral therapy (CBT) on anxiety?

Individual's Decision Problem

- ▶ You have anxiety, and need to decide whether to get CBT ($D = 1$) or not ($D = 0$). Weigh the costs against benefits. [Chamberlain \(2011\)](#)
- ▶ You are probably interested in the ATE or conditional ATE: *on average, what is the treatment effect for a person like me?*
- ▶ Side point: experiment only tells you useful information under a *consistency condition*, i.e. *choosing* treatment has the same effect as *being allocated* treatment.
- ▶ Crucial, if obvious, feature: *you can force yourself to take treatment*

Causal Effects are for *Decisionmaking*

Example Causal Question

- ▶ What is the causal effect of cognitive behavioral therapy (CBT) on anxiety?

Policymaker's Decision Problem

- ▶ Should we expand access to CBT on the UK National Health Service (NHS)?
Weigh the costs against the benefits.
- ▶ We can't force people with anxiety to get CBT by making it more widely available so the ATE isn't the relevant quantity.
- ▶ If we expand access, some more people will be treated. Policy question is: **what is the average benefit, per additional person enrolled, of expanding access?**
- ▶ When treatment is *voluntary*, it becomes crucial for policy analysis to understand how treatment effects may correlate with willingness to *take up* treatment.

Causal Effects for Policymaking? TOT and TUT Effects

Treatment on the Treated (TOT aka ATT)

- ▶ Existing program; only some of those eligible choose to enroll. If we **eliminated** the program, how much **worse off** would current participants be?
- ▶ Average effect of a program or policy for those who currently choose to enroll.
- ▶ Equals LATE under one-sided non-compliance: no always-takers

Treatment on the Untreated (TUT aka ATU)

- ▶ Existing program; only some of those eligible choose to enroll. If we **forced** all non-participants to enroll, how much **better off** would they be?
- ▶ Average effect of a program or policy for those who currently choose **not** to enroll.
- ▶ Equals LATE under one-sided non-compliance: no never-takers
- ▶ E.g. increase in UK minimum school-leaving age from 15 to 16 (September 1972).

Beyond LATE in a “Textbook” Model

$$Y_0 = \mu_0 + U_0$$

$$Y_1 = \mu_1 + U_1$$

$$D = 1\{\gamma_0 + \gamma_1 Z > V\}$$

$$Y = (1 - D)Y_0 + DY_1$$

$$Z \sim \text{Bernoulli}(q) \perp\!\!\!\perp (V, U_0, U_1)$$

$$\begin{bmatrix} V \\ U_0 \\ U_1 \end{bmatrix} \sim \text{Normal} \left(\begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix}, \begin{bmatrix} 1 & \sigma_0 \rho_0 & \sigma_1 \rho_1 \\ & \sigma_0^2 & \sigma_{01} \\ & & \sigma_1^2 \end{bmatrix} \right)$$

- ▶ Heckman, Tobias & Vytlacil (2001), Angrist (2004)
- ▶ Treatment effects ($Y_1 - Y_0$) are heterogeneous, $\text{ATE} = \mu_1 - \mu_0$.
- ▶ Selection into treatment up D depends on:
 1. Binary instrument / encouragement Z
 2. Heterogeneous cost / **resistance to treatment** V (free normalization)
- ▶ Closed-form expressions: compare ATE, LATE, TOT and TUT.

Simulation: $\mu_1 = \mu_0 = 0$, $\sigma_0 = \sigma_1 = 1$, $\sigma_{01} = 1/2$

```
library(mvtnorm)
library(tidyverse)
rho0 <- 0.5
rho1 <- 0.2

S <- matrix(c(1, rho0, rho1,
              rho0, 1, 0.5,
              rho1, 0.5, 1), 3, 3, byrow = TRUE)
set.seed(1983)
sims <- rmvnorm(5e3, sigma = S)
colnames(sims) <- c('V', 'Y0', 'Y1')
sims <- as_tibble(sims)
sims <- sims %>%
  mutate(Delta = Y1 - Y0)
```



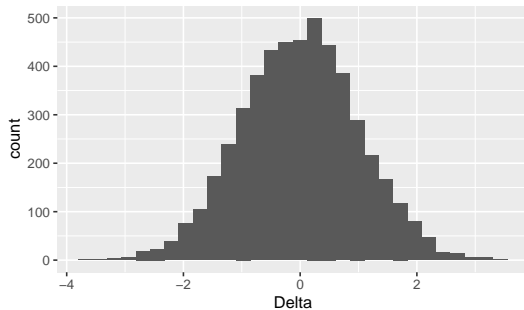
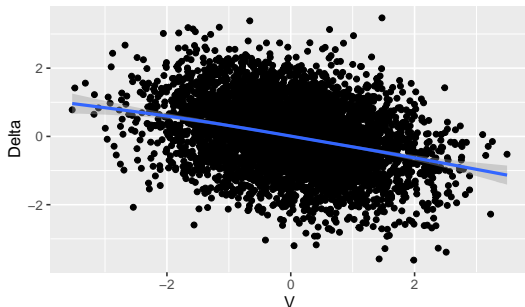
```
sims
```

```
## # A tibble: 5,000 x 4
##           V           Y0           Y1      Delta
##       <dbl>     <dbl>     <dbl>   <dbl>
##  1 -0.122    -0.399     1.08    1.48
##  2 -0.506    -1.10      1.49    2.59
##  3  0.00457  -0.121    -0.456  -0.335
##  4 -0.549    -0.248    -0.899  -0.651
##  5  1.95     -0.0948   -0.675  -0.580
##  6  0.561     0.112    -0.615  -0.726
##  7 -0.238    -0.439    -1.53   -1.10
##  8 -1.46     -1.23     -0.0548  1.17
##  9 -0.336    -0.891     1.53    2.42
## 10 -0.845    -0.274     0.637   0.911
## # ... with 4,990 more rows
```

```
DV_scatter <- sims %>%  
  ggplot(aes(x = V, y = Delta)) +  
  geom_point() +  
  geom_smooth()
```

```
Dhist <- sims %>%  
  ggplot(aes(x = Delta)) +  
  geom_histogram()
```

```
library(gridExtra)
grid.arrange(DV_scatter, Dhist, ncol = 2)
```



Any Parameter values

- ▶ Δ is normally distributed; Δ and V are linearly dependent (jointly normal).

These Parameter Values

- ▶ ATE is zero; higher cost/resistance $V \implies$ lower treatment effect Δ

Properties of the Textbook Model

$$Y_0 = \mu_0 + U_0$$

$$Y_1 = \mu_1 + U_1$$

$$D = 1\{\gamma_0 + \gamma_1 Z > V\}$$

$$Y = (1 - D)Y_0 + DY_1$$

$$Z \sim \text{Bernoulli}(q) \perp\!\!\!\perp (V, U_0, U_1)$$

$$\begin{bmatrix} V \\ U_0 \\ U_1 \end{bmatrix} \sim \text{Normal} \left(\begin{bmatrix} 0 \\ 0 \\ 0 \end{bmatrix}, \begin{bmatrix} 1 & \sigma_0 \rho_0 & \sigma_1 \rho_1 \\ & \sigma_0^2 & \sigma_{01} \\ & & \sigma_1^2 \end{bmatrix} \right)$$

Implications

- ▶ $\Delta \equiv Y_1 - Y_0 \sim \text{Normal}(\mu_1 - \mu_0, \sigma_0^2 + \sigma_1^2 - 2\sigma_{01})$
- ▶ $\text{Cov}(\Delta_i, V_i) = \text{Cov}(Y_{1i}, V_i) - \text{Cov}(Y_{0i}, V_i) = \sigma_1 \rho_1 - \sigma_0 \rho_0$

LATE for the Textbook Model

- ▶ LATE = average effect for *compliers*: people induced to take treatment by Z .
- ▶ Since $D = 1(\gamma_0 + \gamma_1 Z > V)$, compliers are defined by $\gamma_0 \leq V < \gamma_0 + \gamma_1$
- ▶ **Depends on the particular instrument** through γ_0, γ_1

```
gamma0 <- -1  
gamma1 <- 1.5  
sims <- sims %>%  
  mutate(complier = (V >= gamma0) & (V < gamma0 + gamma1))
```

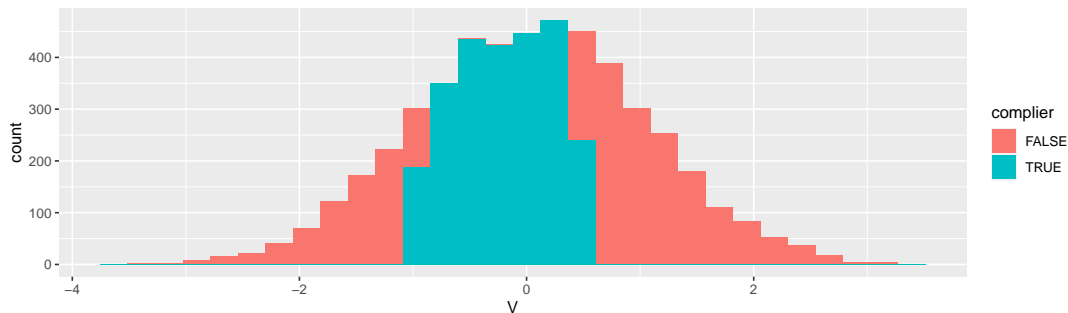
Who's a complier when $\gamma_0 = -1$ and $\gamma_1 = 1.5$?

```
sims
```

```
## # A tibble: 5,000 x 5
##           V           Y0           Y1 Delta complier
##       <dbl>    <dbl>    <dbl>    <dbl> <lgl>
##  1 -0.122    -0.399     1.08     1.48  TRUE
##  2 -0.506    -1.10      1.49     2.59  TRUE
##  3  0.00457  -0.121    -0.456   -0.335 TRUE
##  4 -0.549    -0.248    -0.899   -0.651 TRUE
##  5  1.95     -0.0948   -0.675   -0.580 FALSE
##  6  0.561     0.112    -0.615   -0.726 FALSE
##  7 -0.238    -0.439    -1.53    -1.10  TRUE
##  8 -1.46     -1.23     -0.0548  1.17   FALSE
##  9 -0.336    -0.891     1.53     2.42  TRUE
## 10 -0.845    -0.274     0.637     0.911 TRUE
## # ... with 4,990 more rows
```

Whos's a complier when $\gamma_0 = -1$, $\gamma_1 = 1.5$?

```
ggplot(sims, aes(x = V, fill = complier)) +  
  geom_histogram()
```

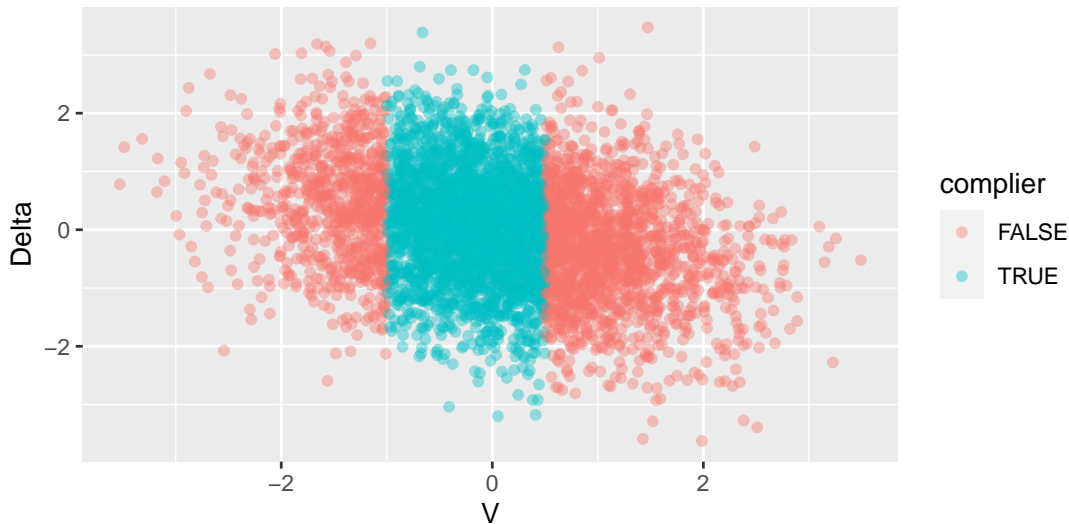


```
# Share of compliers  
pnorm(gamma0 + gamma1) - pnorm(gamma0)
```

```
## [1] 0.5328072
```

Who's a complier when $\gamma_0 = -1$ and $\gamma_1 = 1.5$?

```
ggplot(sims, aes(x = V, y = Delta, col = complier)) +  
  geom_point(alpha = 0.4)
```



Average Treatment Effects by Complier Status: $\gamma_0 = -1$, $\gamma_1 = 1.5$

```
sims %>%  
  group_by(complier) %>%  
  summarize(mean(Y1 - Y0)) %>%  
  knitr::kable(digits = 3)
```

complier	mean(Y1 - Y0)
FALSE	-0.083
TRUE	0.068

Different Instrument, Different LATE: $\gamma_0 = -1$, Varying γ_1

```
get_LATE <- function(gamma1) {  
  sims %>%  
    mutate(complier = (V >= -1) & (V < -1 + gamma1)) %>%  
    filter(complier) %>%  
    summarize(LATE = mean(Y1 - Y0)) %>%  
    pull()  
}  
  
gamma1_seq <- c(0.75, 1, 1.25, 1.5, 1.75, 2)  
LATE <- map_dbl(c(0.75, 1, 1.25, 1.5, 1.75, 2), get_LATE)  
rbind(gamma1_seq, LATE) %>% knitr::kable(digits = 2)
```

gamma1_seq	0.75	1.00	1.25	1.50	1.75	2
LATE	0.21	0.15	0.11	0.07	0.03	0

TOT and TUT in the Textbook Model

$$\begin{aligned}\text{TOT} &\equiv \mathbb{E}(\Delta|D = 1) \\ &= \mathbb{E}(\Delta|D = 1, Z = 0)\mathbb{P}(Z = 0|D = 1) + \mathbb{E}(\Delta|D = 1, Z = 1)\mathbb{P}(Z = 1|D = 1) \\ &= \underbrace{\mathbb{E}(\Delta|V < \gamma_0)}_{\text{Always-takers}} \times (1 - q_1) + \underbrace{\mathbb{E}(\Delta|V < \gamma_0 + \gamma_1)}_{\text{Always-takers \& Compliers}} \times q_1\end{aligned}$$

$$\begin{aligned}\text{TUT} &\equiv \mathbb{E}(\Delta|D = 0) \\ &= \mathbb{E}(\Delta|D = 0, Z = 0)\mathbb{P}(Z = 0|D = 0) + \mathbb{E}(\Delta|D = 0, Z = 1)\mathbb{P}(Z = 1|D = 0) \\ &= \underbrace{\mathbb{E}(\Delta|V > \gamma_0)}_{\text{Never-takers \& Compliers}} (1 - q_1) + \underbrace{\mathbb{E}(\Delta|V > \gamma_0 + \gamma_1)}_{\text{Never-takers}} q_1\end{aligned}$$

TOT and TUT in the Textbook Model

- ▶ TOT is a weighted average of $\mathbb{E}(\Delta|V < \gamma_0)$ and $\mathbb{E}(\Delta|V < \gamma_0 + \gamma_1)$.
- ▶ TUT is a weighted average of $\mathbb{E}(\Delta|V > \gamma_0)$ and $\mathbb{E}(\Delta|V > \gamma_0 + \gamma_1)$.
- ▶ Need to be able to calculate $\mathbb{E}(\Delta|V > c)$ and $\mathbb{E}(\Delta|V < c)$.
- ▶ TOT and TUT depend on Z through γ_0 and γ_1 : defines “the treated”

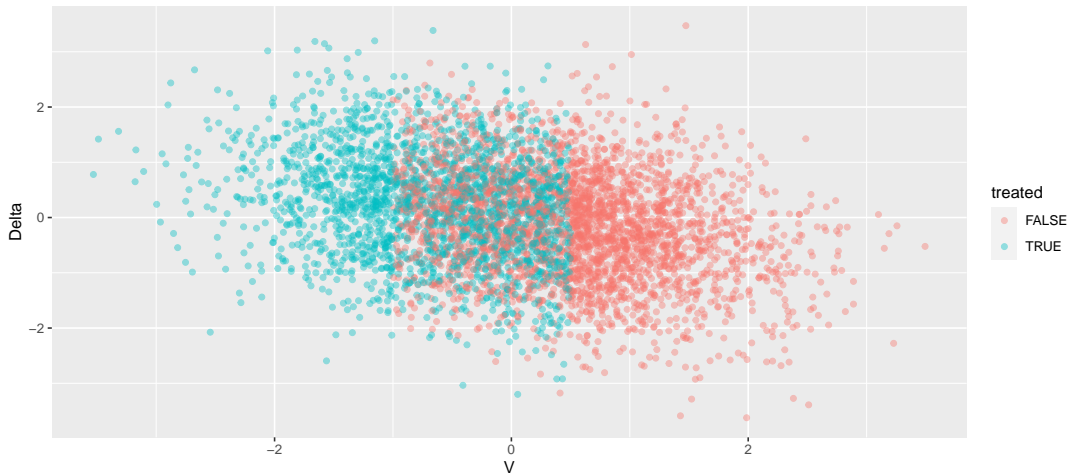
```
# Need Z to define "the treated"
sims <- sims %>%
  select(-complier) %>%
  mutate(Z = rbinom(nrow(sims), 1, 0.5),
         treated = gamma0 + gamma1 * Z > V)
sims
```

```
## # A tibble: 5,000 x 6
```

```
##           V           Y0           Y1   Delta     Z treated
##      <dbl>    <dbl>    <dbl>  <dbl>  <int>  <lgl>
##  1 -0.122    -0.399     1.08    1.48      0 FALSE
##  2 -0.506    -1.10      1.49    2.59      0 FALSE
##  3  0.00457  -0.121    -0.456  -0.335     0 FALSE
##  4 -0.549    -0.248    -0.899  -0.651     0 FALSE
##  5  1.95     -0.0948   -0.675  -0.580     1 FALSE
##  6  0.561     0.112    -0.615  -0.726     0 FALSE
##  7 -0.238    -0.439    -1.53   -1.10      1 TRUE
##  8 -1.46     -1.23    -0.0548  1.17      1 TRUE
##  9 -0.336    -0.891     1.53    2.42      1 TRUE
```

Who's treated if $q = 0.5$, $\gamma_0 = -1$ and $\gamma_1 = 1.5$?

```
ggplot(sims, aes(x = V, y = Delta, col = treated)) +  
  geom_point(alpha = 0.4)
```



TOT and TUT Effects: $q = 0.5$, $\gamma_0 = -1$ and $\gamma_1 = 1.5$

```
sims %>%  
  group_by(treated) %>%  
  summarize(mean(Y1 - Y0)) %>%  
  knitr::kable(digits = 3)
```

treated	mean(Y1 - Y0)
FALSE	-0.170
TRUE	0.223

- ▶ Different values of q , γ_0 , γ_1 , would give different TUT and TOT.
- ▶ In this example we have **selection on gains**: $TUT < ATE < TOT$

Analytical Results for the Textbook Model

$$ATE = \mu_1 - \mu_0$$

$$LATE = ATE - (\sigma_1\rho_1 - \sigma_0\rho_0) \left[\frac{\varphi(\gamma_0 + \gamma_1) - \varphi(\gamma_0)}{\Phi(\gamma_0 + \gamma_1) - \Phi(\gamma_0)} \right]$$

$$TOT = ATE - (\sigma_1\rho_1 - \sigma_0\rho_0) \left[\frac{(1 - q)\varphi(\gamma_0) + q\varphi(\gamma_0 + \gamma_1)}{(1 - q)\Phi(\gamma_0) + q\Phi(\gamma_0 + \gamma_1)} \right]$$

$$TUT = ATE + (\sigma_1\rho_1 - \sigma_0\rho_0) \left[\frac{(1 - q)\varphi(\gamma_0) + q\varphi(\gamma_0 + \gamma_1)}{(1 - q)\{1 - \Phi(\gamma_0)\} + q\{1 - \Phi(\gamma_0 + \gamma_1)\}} \right]$$

Example: $\sigma_0 = \sigma_1 = 1$ and $q = 1/2$

Formulas Simplify ($\delta \equiv \rho_1 - \rho_0$)

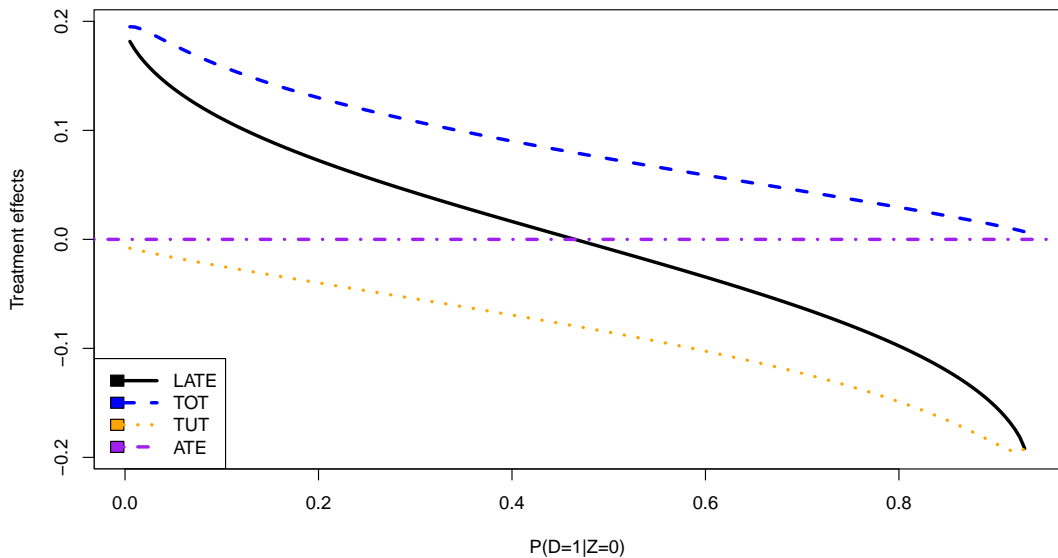
$$\text{LATE} = -\delta \left[\frac{\varphi(\gamma_0 + \gamma_1) - \varphi(\gamma_0)}{\Phi(\gamma_0 + \gamma_1) - \Phi(\gamma_0)} \right]$$

$$\text{TOT} = -\delta \left[\frac{\varphi(\gamma_0) + \varphi(\gamma_0 + \gamma_1)}{\Phi(\gamma_0) + \Phi(\gamma_0 + \gamma_1)} \right]$$

$$\text{TUT} = \delta \left[\frac{\varphi(\gamma_0) + \varphi(\gamma_0 + \gamma_1)}{\{1 - \Phi(\gamma_0)\} + \{1 - \Phi(\gamma_0 + \gamma_1)\}} \right]$$

- In the practical session you will reproduce some plots from Angrist (2004).

First-stage effect 0.07, $q = 1/2$, $\delta = -0.1$



Why do we care about any of this?

- ▶ In the textbook model we can see how the ATE, LATE, TOT and TUT compare.
- ▶ The key parameters of the textbook model **are point identified**.
- ▶ This allows us to use data to go *beyond LATE* to other causal effects: ATE, TOT and TUT, and more (next time).
- ▶ **Next Time:** Marginal Treatment Effects methods are a modern “update” of this textbook model.

Heckman Two-step Estimator

We will show that:

$$\mathbb{E}[Y|D = 1, Z = z] = \mu_1 + \delta_1 \mathbb{E}(V|D = 1, Z = z)$$

$$\mathbb{E}(V|D = 1, Z = z) = \frac{-\varphi(\gamma_0 + \gamma_1 z)}{\Phi(\gamma_0 + \gamma_1 z)}$$

$$\mathbb{E}[Y|D = 0, Z = z] = \mu_0 + \delta_0 \mathbb{E}(V|D = 0, Z = z)$$

$$\mathbb{E}(V|D = 0, Z = z) = \frac{\varphi(\gamma_0 + \gamma_1 z)}{1 - \Phi(\gamma_0 + \gamma_1 z)}$$

Heckman Two-step Estimator

Define the following shorthand:

$$\lambda(z) \equiv \mathbb{E}(V|D = 0, Z = z) = \frac{\varphi(\gamma_0 + \gamma_1 z)}{1 - \Phi(\gamma_0 + \gamma_1 z)}$$
$$\kappa(z) \equiv \mathbb{E}(V|D = 1, Z = z) = \frac{-\varphi(\gamma_0 + \gamma_1 z)}{\Phi(\gamma_0 + \gamma_1 z)}.$$

Then we have

$$\mathbb{E}[Y|D = 0, Z] = \mu_0 + \delta_0 \lambda(Z)$$

$$\mathbb{E}[Y|D = 1, Z] = \mu_1 + \delta_1 \kappa(Z)$$

- ▶ Use D and Z to estimate γ_0 and γ_1
- ▶ To estimate μ_0 and δ_0 regress Y on $\lambda(Z)$ and a constant for obs with $D = 0$
- ▶ To estimate μ_1 and δ_1 regress Y on $\kappa(Z)$ and a constant for obs with $D = 1$

Step 1: $(U_0, U_1) \perp\!\!\!\perp Z \mid V$

Axioms of Conditional Independence

- See <https://expl.ai/LXPVDDN> or chapter 2 of the [lecture notes](#)

$$\text{(Assumption)} \quad Z \perp\!\!\!\perp (U_0, U_1, V) \implies Z \perp\!\!\!\perp (U_0, U_1, V) \mid V \quad \text{(Weak Union)}$$

$$\implies Z \perp\!\!\!\perp (U_0, U_1) \mid V \quad \text{(Decomposition)}$$

$$\implies (U_0, U_1) \perp\!\!\!\perp Z \mid V \quad \text{(Symmetry)}$$

Step 2: $\mathbb{E}(U_0|V)$ and $\mathbb{E}(U_1|V)$.

General Result: $(X, Y) \sim \text{Bivariate Normal}$

$$\mathbb{E}(Y|X = x) = \mathbb{E}(Y) + \frac{\text{Cov}(Y, X)}{\text{Var}(X)} [x - \mathbb{E}(X)]$$

Our Setting: $V \sim N(0, 1)$

$$\mathbb{E}(Y_1 - Y_0|V) = (\mu_1 - \mu_0) + \mathbb{E}(U_1 - U_0)$$

$$\mathbb{E}(U_1|V) = \sigma_1 \rho_1 V \equiv \delta_0 V$$

$$\mathbb{E}(U_0|V) = \sigma_0 \rho_0 V \equiv \delta_1 V$$

$$\mathbb{E}(U_1 - U_0|V) = (\sigma_1 \rho_1 - \sigma_0 \rho_0) V \equiv (\delta_1 - \delta_0) V$$

Step 3: $\mathbb{E}(Y|D, Z, V)$

$$\begin{aligned}\mathbb{E}(Y|D = 0, Z, V) &= \mathbb{E}(Y_0|D = 0, Z, V) \\ &= \mu_0 + \mathbb{E}(U_0|D = 0, Z, V) && \text{(Defn. of } Y_0\text{)} \\ &= \mu_0 + \mathbb{E}(U_0|Z, V) && (D = f(Z, V)) \\ &= \mu_0 + \mathbb{E}(U_0|V) && \text{(Step 1)} \\ &= \mu_0 + \delta_0 V && \text{(Step 2)}\end{aligned}$$

$$\mathbb{E}(Y|D = 1, Z, V) = \mu_1 + \delta_1 V \quad \text{(Same Steps)}$$

Step 4: $\mathbb{E}(Y, D, Z)$

$$\begin{aligned}\mathbb{E}(Y|D=0, Z) &= \mathbb{E}_{V|(D=0, Z)} [\mathbb{E}(Y|D=0, Z, V)] && \text{(Iterated } \mathbb{E}) \\ &= \mathbb{E}(\mu_0 + \delta_0 V|D=0, Z) && \text{(Step 3)} \\ &= \mu_0 + \delta_0 \mathbb{E}(V|D=0, Z) && \text{(Linearity of } \mathbb{E})\end{aligned}$$

$$\mathbb{E}(Y|D=1, Z) = \mu_1 + \delta_1 \mathbb{E}(V|D=1, Z) \quad \text{(Same Steps)}$$

The Mean of a Truncated Normal Distribution

- ▶ We will need these results on the next slide!
- ▶ Derivation of the first result: <https://expl.ai/VFARCYE>.

Suppose that $Z \sim N(0, 1)$. Then for any constants a, b, c

$$E(Z|Z > c) = \frac{\varphi(c)}{1 - \Phi(c)}$$

$$E(Z|Z < c) = \frac{-\varphi(c)}{\Phi(c)}$$

$$E(Z|a < Z < b) = \frac{-[\varphi(b) - \varphi(a)]}{\Phi(b) - \Phi(a)}$$

Step 5: $\mathbb{E}(V|D, Z)$

$$\begin{aligned}\mathbb{E}(V|D=1, Z=1) &= \mathbb{E}(V|\gamma_0 + \gamma_1 > V, Z=1) && (D = f(Z, V)) \\ &= \mathbb{E}(V|\gamma_0 + \gamma_1 > V) && (V \perp\!\!\!\perp Z) \\ &= \frac{-\varphi(\gamma_0 + \gamma_1)}{\Phi(\gamma_0 + \gamma_1)} && (\text{Trunc. Normal})\end{aligned}$$

$$\mathbb{E}(V|D=1, Z=0) = \frac{-\varphi(\gamma_0 + \gamma_1)}{\Phi(\gamma_0 + \gamma_1)} \quad (\text{Similar Steps})$$

$$\mathbb{E}(V|D=0, Z=1) = \frac{\varphi(\gamma_0 + \gamma_1)}{1 - \Phi(\gamma_0 + \gamma_1)} \quad (\text{Similar Steps})$$

$$\mathbb{E}(V|D=0, Z=0) = \frac{\varphi(\gamma_0)}{1 - \Phi(\gamma_0)} \quad (\text{Similar Steps})$$