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# A RELATIONSHIP BETWEEN ARBITRARY POSITIVE MATRICES AND DOUBLY STOCHASTIC MATRICES

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**1.** Introduction. Suppose one observes n transitions of a Markov chain with N states and stochastic matrix  $P = (p_{ij})$ . The usual estimate of  $p_{ij}$  is  $t_{ij} = a_{ij}/\lambda_i$  where  $a_{ij}$  is the number of transitions from i to j which are observed, and  $\lambda_i = \sum_j a_{ij}$ . (Cf. [1].) This amounts to a normalization of the rows of  $A = (a_{ij})$ , and can be expressed as a matrix equation  $T = D_1 A$  where  $T = (t_{ij})$  and  $D_1 = \text{diag}[\lambda_1^{-1}, \dots, \lambda_N^{-1}]$ .

If it is known that the stochastic matrix P is in fact doubly stochastic, (i.e.,  $\sum_{i} p_{ij} = 1$ ), what then is a good estimate of T? The maximum likelihood equations are difficult to solve. One estimate which has been used (for example, by Welch [4]) is to alternately normalize the rows and columns of A, in the belief that this iterative process converges to a doubly stochastic matrix, T, which might be, in some sense, a good estimate.

It is not the intent of this paper to obtain properties of this estimate but only to examine the mechanics of the iteration itself. In the next section we shall study this in detail and show that it is always convergent if the matrix A is strictly positive (i.e.,  $a_{ij} > 0$  for all i, j), and in fact that there exist diagonal matrices  $D_1$  and  $D_2$  (unique up to a scalar factor) with positive diagonals such that  $T = D_1 A D_2$ . T is the only doubly stochastic matrix expressable in this form for a given strictly positive A.

For completeness we shall include a corollary to this result due to Marcus and Newman [3] which states that if A is symmetric and has positive entries, then there exists a diagonal matrix D with positive main diagonal entries such that DAD is doubly stochastic.

Finally in the last section we shall show by example that convergence need not occur at all if some  $a_{ij} = 0$ , or even if it does there need exist no associated diagonal matrices  $D_1$  and  $D_2$  as in the strictly positive case. Even the apparently natural artifice of replacing the zero entries by "small" functions  $a_{ij}(\epsilon)$  of a parameter  $\epsilon$ , getting  $T(\epsilon)$  and letting  $\epsilon \to 0$  leads to difficulties.

## 2. The alternating iteration for positive matrices.

THEOREM 1. To a given strictly positive  $N \times N$  matrix A there corresponds exactly one doubly stochastic matrix  $T_A$  which can be expressed in the form  $T_A = D_1AD_2$  where  $D_1$  and  $D_2$  are diagonal matrices with positive diagonals. The matrices  $D_1$  and  $D_2$  are themselves unique up to a scalar factor.

Proof. We shall establish only the uniqueness part here. The existence will be demonstrated constructively in the proof of Theorem 2.

If there exist two different pairs of diagonal matrices  $D_1$ ,  $D_2$  and  $C_1$ ,  $C_2$  such

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that both  $D_1AD_2$  and  $C_1AC_2$  are doubly stochastic, then this means that there exists a positive doubly stochastic matrix P and matrices  $B_1 = \text{diag}[b_{11}, b_{12}, \cdots, b_{1N}]$ ,  $B_2 = \text{diag}[b_{21}, b_{22}, \cdots, b_{2N}]$  which are not multiples of the identity matrix, for which  $B_1PB_2$  is also doubly stochastic. But this is impossible since by convexity, one obtains

$$\min_{i} b_{2i} \leq 1/b_{1i} \leq \max_{i} b_{2i}$$
;  $\min_{i} b_{1i} \leq 1/b_{2i} \leq \max_{i} b_{1i}$ 

which leads to a contradiction if  $b_{1i}b_{2j} \neq 1$  for some i and j. It follows that  $C_1 = pD_1$ ,  $C_2 = p^{-1}D_2$  for some positive number p.

Theorem 2. The iterative process of alternately normalizing the rows and columns of strictly positive  $N \times N$  matrix is convergent to a strictly positive doubly stochastic matrix.

Proof. The iteration produces a sequence of positive matrices which alternately have row and column sums one. We shall show that the two subsequences which are composed respectively of the matrices with row sums one and the matrices with column sums one each converge to a positive doubly stochastic limit of the form  $D_1AD_2$  where each  $D_i$  has a positive diagonal. The uniqueness part of Theorem 1 will complete the proof. Since the terms of either of the subsequences are generated in the same way as are the transposes of the terms of the other, only one convergence proof is required.

Let  $\{A_n\} = \{(a_{nij})\}\$  be the sequence with column sums one and let  $a_n$  be the minimal element of  $A_n$ . We shall show that  $\{a_n\}$  is bounded away from zero.

Let  $A_n$  have row sums  $\lambda_{n1}$ ,  $\cdots$ ,  $\lambda_{nN}$  and set

$$\delta_{nj} = \sum_i a_{nij}/\lambda_{ni}$$
.

Since  $\delta_{nj}$  is a convex combination of the  $1/\lambda_{ni}$  and  $\lambda_{n+1,i}$  is a convex combination of the  $1/\delta_{nj}$ , it follows that

$$(1) \quad \lambda_n(m) \leq 1 \leq \lambda_n(M) \Rightarrow \lambda_n(m) \leq \lambda_{n+1}(m) \leq 1 \leq \lambda_{n+1}(M) \leq \lambda_n(M)$$

where the m and M respectively label minimal and maximal quantities relative to a given  $A_n$ . Similarly

$$(2) \quad \delta_n(m) \leq 1 \leq \delta_n(M) \Rightarrow \delta_n(m) \leq \delta_{n+1}(m) \leq 1 \leq \delta_{n+1}(M) \leq \delta_n(M).$$

Therefore the maximum and minimum row and column sums are monotone sequences and hence have limits. To complete the proof, it is necessary to show that these limits all equal one.

Let  $x_{ni} = [\lambda_{1i}\lambda_{2i}\cdots\lambda_{ni}]^{-1}$  and  $y_{nj} = [\delta_{1j}\delta_{2j}\cdots\delta_{nj}]^{-1}$ . Then if  $A_1 = (a_{ij})$  has minimal element  $a_i$ 

$$y_{nj} = 1/\sum_{i} a_{ij}x_{ni} \le 1/a_{ij}x_{ni} \le 1/ax_{ni}$$

for all i and j. Thus in particular  $y_{nj} \leq 1/ax_n(M)$ . Since  $\sum_j x_{ni}a_{ij}y_{nj} = \lambda_{n+1,i} \geq \lambda_{n+1}(m) \geq \lambda_1(m) = \lambda$ , it follows that

$$x_{ni} \ge \lambda / \sum_{j} a_{ij} y_{nj} \ge a \lambda x_n(M) / N.$$

Also  $y_{nj} \ge 1/\sum_i a_{ij}x_{ni} \ge 1/Nx_n(M)$  and we see that  $a_{n+1, ij} = x_{ni}a_{ij}y_{nj} \ge a\lambda/N^2 = \mu$ ; whence  $a_n \ge \mu > 0$  for all n.

It is clear from (1) that  $\lambda_n(M) \to 1 + c$  where  $c \ge 0$ . For convenience set  $\lambda_n(M) = 1 + c_n$ . Then if  $\mu_i[\lambda_{ni} \le 1] = \sum a_{nij}$  where the sum is taken over all i for which  $\lambda_{ni} \le 1$ , and if  $\mu_i[\lambda_{ni} > 1]$  has a corresponding meaning,

$$\delta_{nj} \ge \mu_j[\lambda_{ni} \le 1] + \frac{1}{1+c_n} \mu_j[\lambda_{ni} > 1] = \frac{1+c_n \mu_j[\lambda_{ni} \le 1]}{1+c_n} \ge \frac{1+c_n a_n}{1+c_n}.$$

Then if  $\lambda_{n+1}(M) = \lambda_{n+1,i_0}$ ,

$$1 + c \le \lambda_{n+1,i_0} = \sum_{j} a_{ni_0j}/\lambda_{ni_0}\delta_{nj} \le (1 + c_n)/(1 + c_na_n);$$

if c > 0,  $a_n \to 0$ , a contradiction. Thus c = 0 and  $\lambda_n(M) \to 1$ . It readily follows that  $\lambda_n(m) \to 1$ .

COROLLARY. (Marcus and Newman [3]) If A is symmetric and has positive entries there exists a diagonal matrix D with positive main diagonal entries such that DAD is doubly stochastic.

PROOF. Let  $S = D_1 A D_2$  be doubly stochastic where the  $D_i$  are as in Theorem 1. Then  $A = D_1^{-1} S D_2^{-1}$  and  $A^T = A$  implies that  $D_2 D_1^{-1} S D_2^{-1} D_1 = S^T$ , and since  $S^T$  is doubly stochastic,  $D_2 D_1^{-1}$  is a scalar multiple of the identity by the uniqueness part of Theorem 1. Thus we can take  $D_1 = D_2 = D$ .

**3.** Remarks concerning matrices with zero entries. When A contains zero elements Theorems 1 and 2 need no longer hold. If

$$A = \begin{pmatrix} 0 & 0 & 1 \\ 0 & 0 & 1 \\ 1 & 1 & 0 \end{pmatrix}$$

the iteration oscillates and there certainly exists no  $D_1$  and  $D_2$ . If

$$A = \begin{pmatrix} 0 & 1 \\ 1 & 1 \end{pmatrix}$$
 the iteration converges to  $\begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix}$ ,

but again there is no  $D_1$  and  $D_2$ .

One might try to overcome these difficulties by replacing the zero elements in A by small quantities. But this approach may be questionable. For instance if

$$A = \begin{pmatrix} 0 & 0 & 1 \\ 0 & 0 & 1 \\ 1 & 1 & 1 \end{pmatrix} \text{ is approximated by } A(\epsilon) = \begin{pmatrix} \epsilon & \epsilon & 1 \\ \epsilon & \epsilon & 1 \\ 1 & 1 & 1 \end{pmatrix},$$

the limit of  $D_1(\epsilon)A(\epsilon)D_2(\epsilon)$  is

$$\begin{pmatrix} .25 & .25 & .5 \\ .25 & .25 & .5 \\ .5 & .5 & 0 \end{pmatrix} \quad \text{as} \quad \epsilon \to 0.$$

If A is approximated by

$$A'(\epsilon) = \begin{pmatrix} \epsilon & \epsilon^2 & 1 \\ \epsilon & \epsilon^2 & 1 \\ 1 & 1 & 1 \end{pmatrix},$$

the doubly stochastic limit becomes

$$\begin{pmatrix} .5 & 0 & .5 \\ .5 & 0 & .5 \\ 0 & 1 & 0 \end{pmatrix} \quad \text{as} \quad \epsilon \to 0,$$

something quite different. In fact, it is possible to have  $A(\epsilon) \to A$  without having  $D_1(\epsilon)A(\epsilon)D_2(\epsilon)$  converge at all as  $\epsilon \to 0$ . If

$$A = \begin{pmatrix} 0 & 0 \\ 1 & 1 \end{pmatrix}$$
 and  $A(\epsilon) = \begin{pmatrix} \epsilon & \epsilon \sin^2 1/\epsilon \\ 1 & 1 \end{pmatrix}$ 

where  $\epsilon > 0$ ,  $\epsilon \neq 1/n\pi$ ,

$$D_1(\epsilon)A(\epsilon)D_2(\epsilon) = \begin{pmatrix} \alpha_{\epsilon} & 1 - \alpha_{\epsilon} \\ 1 - \alpha_{\epsilon} & \alpha_{\epsilon} \end{pmatrix}$$

where  $\alpha_{\epsilon}^{-1} = 1 + |\sin 1/\epsilon|$ . This has no limit as  $\epsilon \to 0$ .

Whence any attempt to estimate the transition matrix from an observation matrix by a double normalization or by an alternating row-column iteration may well result in failure when the observation contains zero entries. It may also be a poor policy to use a strictly positive approximation for an observation with zeros in hopes of finding an approximate transition matrix, unless there is some very good reason for a particular selection.

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#### REFERENCES

- BILLINGSLEY, PATRICK (1962). Statistical Inference for Markov Processes. Univ. of Chicago Press.
- [2] KEMENY, JOHN G. and SNELL, J. LAURIE (1960). Finite Markov Chains. Van Nostrand, Princeton.
- [3] MARCUS, MARVIN and NEWMAN, MORRIS (1961). The permanent of a symmetric matrix, Abstract 587-85. Amer. Math. Soc. Notices 8 595.
- [4] Welch, Lloyd, Unpublished Report of the Institute of Defense Analysis. Princeton, New Jersey.