
Homework 3

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- **Acknowledgments:** This coursework refers to wikipedia:
<https://en.wikipedia.org>.
 - **Collaborators:** I finish this coursework by myself.
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I use `enumerate` to generate answers for each question:

- 3.1. (a) $p_H(H_0) = p, C_{ij} = \delta_{ij}$ 由似然比检验, 对于给定的数据 y , 当 $2y \geq \frac{p}{1-p}$ 拒绝 H_0 ; 当 $2y \leq \frac{p}{1-p}$ 时接受 H_0 。由于 $y \in [0, 1]$, 当 $p > \frac{2}{3}$ 时, $2y \leq \frac{p}{1-p}$ 恒成立, 所以此时总是接受 H_0 。
- (b) 当 $p < \frac{2}{3}$ 时,

$$P_D = \int_{\frac{p}{2(1-p)}}^1 2y dy = 1 - \left(\frac{p}{2(1-p)} \right)^2$$
$$P_F = \int_{\frac{p}{2(1-p)}}^1 dy = 1 - \frac{p}{2(1-p)}$$
$$\Rightarrow P_D = 2P_F - P_F^2, p \leq \frac{2}{3} \quad (1)$$

当 $p > \frac{2}{3}$ 时, $P_D = P_F = 0$ 。

- (c) i. 令 $P_D > P_F$, 得 $P_F \leq 1 - \epsilon$, 将 $P_F = 1 - \epsilon$ 代入(1)中得 $P_D^{\max}(\epsilon) = 1 - \epsilon^2$
- ii. 令 $0 < P_F < 1$, 得 $0 < \epsilon < 1$ 。
- iii. 令 $1 - \frac{p}{2} = P_F \leq 1 - \epsilon \Rightarrow p \geq \frac{2\epsilon}{1+2\epsilon}$ 。又 $p \leq 1 \Rightarrow p \in [\min\{\frac{2\epsilon}{1+2\epsilon}, \frac{2}{3}\}, 1]$

- 3.2. (a) i. 设 $\tilde{\varphi}(H, H) = \mathbb{E}[C(H, H)|\underline{y} = \underline{y}]$ 当 $\hat{H}(\underline{y}) = H_1$ 时,

$$\begin{aligned} \tilde{\varphi}(H_1, H) &= C_{11}\mathbb{P}(H = H_1|\underline{y} = \underline{y}) + C_{12}\mathbb{P}(H = H_2|\underline{y} = \underline{y}) + C_{13}\mathbb{P}(H = H_3|\underline{y} = \underline{y}) \\ &= \mathbb{P}(H = H_2|\underline{y} = \underline{y}) + 2\mathbb{P}(H = H_3|\underline{y} = \underline{y}) \end{aligned}$$

同理可求出:

$$\begin{aligned} \tilde{\varphi}(H_2, H) &= \mathbb{P}(H = H_1|\underline{y} = \underline{y}) + 2\mathbb{P}(H = H_3|\underline{y} = \underline{y}) \\ \tilde{\varphi}(H_3, H) &= 2\mathbb{P}(H = H_1|\underline{y} = \underline{y}) + 2\mathbb{P}(H = H_2|\underline{y} = \underline{y}) \end{aligned}$$

设

$$j = \arg \min_{i=1,2,3} \tilde{\varphi}(H_i, H)$$

最优的决策准则为 $\hat{H}(\underline{y}) = H_j$

ii. 利用概率的归一化条件有：

$$\tilde{\varphi}(H_1, H) = \pi_2 + 2(1 - \pi_1 - \pi_2)$$

$$\tilde{\varphi}(H_2, H) = \pi_1 + 2(1 - \pi_1 - \pi_2)$$

$$\tilde{\varphi}(H_3, H) = 2\pi_1 + 2\pi_2$$

根据 π_1, π_2 以及 $\pi_3 = 1 - \pi_1 - \pi_2$ 的非负性和 i. 中的结果，决策区域如图 1 所示。

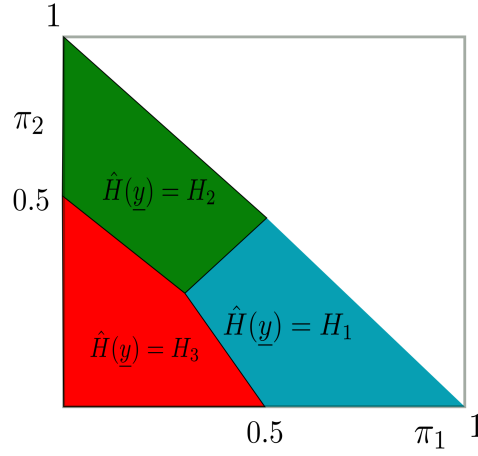


Figure 1: 决策区域

(b) 若 $L_i(\underline{y}), i = 2, 3$ 已知，则

$$\frac{p_{\underline{y}|\mathbf{H}}(\underline{y}|H_2)}{p_{\underline{y}|\mathbf{H}}(\underline{y}|H_3)} = \frac{L_2(\underline{y})}{L_3(\underline{y})}$$

另一方面，因为 H_i 的先验均相等，基于后验概率的 $\pi(\underline{y})$ 正比于

$$\begin{bmatrix} p_{\underline{y}|\mathbf{H}}(\underline{y}|H_1) \\ p_{\underline{y}|\mathbf{H}}(\underline{y}|H_2) \\ p_{\underline{y}|\mathbf{H}}(\underline{y}|H_3) \end{bmatrix}$$

所以最优决策准则可以用 $L_i(\underline{y}), i = 2, 3$ 给出。

$$p_{\underline{y}|\mathbf{H}}(\underline{y}|H_1) = \frac{\exp(-\frac{(y_1-1)^2+y_2^2+y_3^2}{2\sigma^2})}{\sqrt{2\pi}^3 \sigma^3}$$

$$p_{\underline{y}|\mathbf{H}}(\underline{y}|H_2) = \frac{\exp(-\frac{y_1^2+(y_2-1)^2+y_3^2}{2\sigma^2})}{\sqrt{2\pi}^3 \sigma^3}$$

$$p_{\underline{y}|\mathbf{H}}(\underline{y}|H_3) = \frac{\exp(-\frac{y_1^2+y_2^2+(y_3-1)^2}{2\sigma^2})}{\sqrt{2\pi}^3 \sigma^3}$$

因此

$$L_2(\underline{y}) = \exp\left(-\frac{\ell_2(\underline{y})}{\sigma^2}\right)$$

$$L_3(\underline{y}) = \exp\left(-\frac{\ell_3(\underline{y})}{\sigma^2}\right)$$

3.3. (a) 根据 Bayes 风险最小的准则, $\hat{x} = -1$ 当后验概率满足

$$p_{x|\hat{x}}(x=1|\hat{x}) \leq p_{x|\hat{x}}(x=-1|\hat{x})$$

否则取 $\hat{x} = 1$ 由 $P_x(1) = P_x(-1) = 1/2$, 得

$$p_{\hat{x}|x}(\hat{x}|x=1) \leq p_{\hat{x}|x}(\hat{x}|x=-1)$$

记

$$\phi(k) = \begin{cases} \frac{3}{4} & k > 0 \\ \frac{1}{4} & k \leq 0 \end{cases}$$

则由独立性假设

$$p_{\hat{x}|x}(\hat{x}|x=1) = \prod_{i=1}^n p_{\hat{x}_i|x}(\hat{x}_i|x=1)$$

$$= \prod_{i=1}^n \phi(\hat{x}_i)$$

同理

$$p_{\hat{x}|x}(\hat{x}|x=-1) = \prod_{i=1}^n \phi(-\hat{x}_i)$$

因此最小错误概率的判决准则是:

$$\left(\frac{3}{4}\right)^{\sum_i 1_{\hat{x}_i=1}} \left(\frac{1}{4}\right)^{\sum_i 1_{\hat{x}_i=-1}} \underset{\hat{x}=-1}{\overset{\hat{x}=1}{\geq}} \left(\frac{3}{4}\right)^{\sum_i 1_{\hat{x}_i=-1}} \left(\frac{1}{4}\right)^{\sum_i 1_{\hat{x}_i=1}}$$

(b) 若 $\hat{x}_1(y_1) = 1(H_1)$, 期望损失为:

$$\begin{aligned} \tilde{\varphi}(H_1, y_1) = & C(1, 1, 1)p_{H, \hat{x}_2|y}(\hat{x}_2 = 1, x = 1|y_1) + C(1, 1, -1)p_{H, \hat{x}_2|y}(\hat{x}_2 = 1, x = -1|y_1) \\ & + C(1, -1, 1)p_{H, \hat{x}_2|y}(\hat{x}_2 = -1, x = 1|y_1) + C(1, -1, -1)p_{H, \hat{x}_2|y}(\hat{x}_2 = -1, x = -1|y_1) \\ \Rightarrow \tilde{\varphi}(H_1, y_1)p_{y_1}(y_1) = & C(1, 1, 1)p_x(x=1)p_{\hat{x}_2|x}(\hat{x}_2 = 1|x=1)p_{y_1|x}(y_1|x=1) \\ & + C(1, 1, -1)p_x(x=-1)p_{\hat{x}_2|x}(\hat{x}_2 = 1|x=-1)p_{y_1|x}(y_1|x=-1) \\ & + C(1, -1, 1)p_x(x=1)p_{\hat{x}_2|x}(\hat{x}_2 = -1|x=1)p_{y_1|x}(y_1|x=1) \\ & + C(1, -1, -1)p_x(x=-1)p_{\hat{x}_2|x}(\hat{x}_2 = -1|x=-1)p_{y_1|x}(y_1|x=-1) \end{aligned}$$

同理给出 $\hat{x}_1(y_1) = -1(H_0)$, 期望损失 $\tilde{\varphi}(H_0, y_1)$ 满足:

$$\begin{aligned} \tilde{\varphi}(H_0, y_1)p_{y_1}(y_1) = & C(-1, 1, 1)p_x(x=1)p_{\hat{x}_2|x}(\hat{x}_2 = 1|x=1)p_{y_1|x}(y_1|x=1) \\ & + C(-1, 1, -1)p_x(x=-1)p_{\hat{x}_2|x}(\hat{x}_2 = 1|x=-1)p_{y_1|x}(y_1|x=-1) \\ & + C(-1, -1, 1)p_x(x=1)p_{\hat{x}_2|x}(\hat{x}_2 = -1|x=1)p_{y_1|x}(y_1|x=1) \\ & + C(-1, -1, -1)p_x(x=-1)p_{\hat{x}_2|x}(\hat{x}_2 = -1|x=-1)p_{y_1|x}(y_1|x=-1) \end{aligned}$$

当 $\tilde{\varphi}(H_0, y_1) \leq \tilde{\varphi}(H_1, y_1)$ 时接受 H_0 , 得
 $\gamma_2 p_{y_1|x}(y_1|x=1) \leq \gamma_3 p_{y_1|x}(y_1|x=-1)$ 其中

$$\gamma_2 = C(-1, 1, 1) p_x(x=1) p_{\hat{x}_2|x}(\hat{x}_2=1|x=1) + C(-1, -1, 1) p_x(x=1) p_{\hat{x}_2|x}(\hat{x}_2=-1|x=1) \\ - C(1, 1, 1) p_x(x=1) p_{\hat{x}_2|x}(\hat{x}_2=1|x=1) - C(1, -1, 1) p_x(x=1) p_{\hat{x}_2|x}(\hat{x}_2=-1|x=1)$$

$$\gamma_3 = C(-1, 1, -1) p_x(x=-1) p_{\hat{x}_2|x}(\hat{x}_2=1|x=-1) \\ + C(-1, -1, -1) p_x(x=-1) p_{\hat{x}_2|x}(\hat{x}_2=-1|x=-1) \\ - C(1, 1, -1) p_x(x=-1) p_{\hat{x}_2|x}(\hat{x}_2=1|x=-1) - C(1, -1, -1) p_x(x=-1) p_{\hat{x}_2|x}(\hat{x}_2=-1|x=-1)$$

根据 $C(\hat{x}_1, \hat{x}_2, x)$ 的性质,
 $C(-1, 1, 1) > C(1, 1, 1), C(-1, -1, 1) > C(1, -1, 1)$, 所以 $\gamma_2 > 0$, 因此

$$\frac{p_{y_1|x}(y_1|x=1)}{p_{y_1|x}(y_1|x=-1)} \leq \frac{\gamma_3}{\gamma_2} \triangleq \gamma_1$$

另解:

$$\begin{aligned} \bar{R} &= \sum_{x_1, x_2, \tilde{x}=-1, 1} C(x_1, x_2, \tilde{x}) P(\hat{x}_1 = x_1, \hat{x}_2 = x_2, x = \tilde{x}) \\ &= \sum_{x_1, x_2, \tilde{x}=-1, 1} C(x_1, x_2, \tilde{x}) P_x(\tilde{x}) P_{\hat{x}_1|x}(x_1|\tilde{x}) P_{\hat{x}_2|x}(x_2|\tilde{x}) \\ &= \sum_{x_2, \tilde{x}=-1, 1} C(1, x_2, \tilde{x}) P_x(\tilde{x}) P_{\hat{x}_2|x}(x_2|\tilde{x}) \int_{y_1 \in \mathcal{Y}_1} p_{y_1|x}(y_1|\tilde{x}) dy_1 \\ &\quad + \sum_{x_2, \tilde{x}=-1, 1} C(-1, x_2, \tilde{x}) P_x(\tilde{x}) P_{\hat{x}_2|x}(x_2|\tilde{x}) \int_{y_1 \in \mathcal{Y}_0} p_{y_1|x}(y_1|\tilde{x}) dy_1 \\ &= \sum_{x_2, \tilde{x}=-1, 1} C(-1, x_2, \tilde{x}) P_x(\tilde{x}) P_{\hat{x}_2|x}(x_2|\tilde{x}), \mathcal{Y}_1 \cup \mathcal{Y}_0 = \mathbb{R} \\ &\quad + \int_{y_1 \in \mathcal{Y}_1} \sum_{x_2, \tilde{x}=-1, 1} (C(1, x_2, \tilde{x}) - C(-1, x_2, \tilde{x})) P_{\hat{x}_2|x}(x_2|\tilde{x}) P_x(\tilde{x}) p_{y_1|x}(y_1|\tilde{x}) dy_1 \end{aligned}$$

因此

$$\begin{aligned} y_1 \in \mathcal{Y}_1 &\iff \sum_{x_2, \tilde{x}=-1, 1} (C(1, x_2, \tilde{x}) - C(-1, x_2, \tilde{x})) P_{\hat{x}_2|x}(x_2|\tilde{x}) P_x(\tilde{x}) p_{y_1|x}(y_1|\tilde{x}) < 0 \\ &\iff \sum_{x_2=-1, 1} (C(1, x_2, 1) - C(-1, x_2, 1)) P_{\hat{x}_2|x}(x_2|1) P_x(1) p_{y_1|x}(y_1|1) \\ &\quad < \sum_{x_2=-1, 1} (C(-1, x_2, -1) - C(1, x_2, -1)) P_{\hat{x}_2|x}(x_2|-1) P_x(-1) p_{y_1|x}(y_1|-1) \\ &\iff \frac{p_{y_1|x}(y_1|1)}{p_{y_1|x}(y_1|-1)} > \gamma_1 \\ \gamma_1 &= \frac{P_x(-1) \sum_{x_2=-1, 1} (C(-1, x_2, -1) - C(1, x_2, -1)) P_{\hat{x}_2|x}(x_2|-1)}{P_x(1) \sum_{x_2=-1, 1} (C(1, x_2, 1) - C(-1, x_2, 1)) P_{\hat{x}_2|x}(x_2|1)} \end{aligned}$$

(c) 由对称性, 根据 (b) 中的结果, 可得最优的决策准则 $\hat{x}_2^*(\cdot)$ 是如下形式的似然比检验:

$$\frac{p_{y_2|x}(y_2|x=1)}{p_{y_2|x}(y_2|x=-1)} \underset{\hat{x}_2(y_1)=1}{\overset{\hat{x}_2(y_1)=-1}{\geq}} \frac{\gamma'_2}{\gamma'_3} \triangleq \gamma'_1$$

其中

$$\begin{aligned} \gamma'_2 = & C(1, -1, 1)p_x(x=1)p_{\hat{x}_1|x}(\hat{x}_1=1|x=1) + C(-1, -1, 1)p_x(x=1)p_{\hat{x}_1|x}(\hat{x}_1=-1|x=1) \\ & - C(1, 1, 1)p_x(x=1)p_{\hat{x}_1|x}(\hat{x}_1=1|x=1) - C(-1, 1, 1)p_x(x=1)p_{\hat{x}_1|x}(\hat{x}_1=-1|x=1) \end{aligned}$$

$$\begin{aligned} \gamma'_3 = & C(1, -1, -1)p_x(x=-1)p_{\hat{x}_1|x}(\hat{x}_1=1|x=-1) \\ & + C(-1, -1, -1)p_x(x=-1)p_{\hat{x}_1|x}(\hat{x}_1=-1|x=-1) \\ & - C(1, 1, -1)p_x(x=-1)p_{\hat{x}_1|x}(\hat{x}_1=1|x=-1) - C(-1, 1, -1)p_x(x=-1)p_{\hat{x}_1|x}(\hat{x}_1=-1|x=-1) \end{aligned}$$

(d) (b) 中 γ_2, γ_3 可化简为:

$$\begin{aligned} \gamma_2 = & p_x(x=1)p_{\hat{x}_2|x}(\hat{x}_2=1|x=1) + Lp_x(x=1)p_{\hat{x}_2|x}(\hat{x}_2=-1|x=1) \\ & - p_x(x=1)p_{\hat{x}_2|x}(\hat{x}_2=-1|x=1) \\ = & p_x(x=1)(1 + (L-2)p_{\hat{x}_2|x}(\hat{x}_2=-1|x=1)) \end{aligned}$$

$$\begin{aligned} \gamma_3 = & p_x(x=-1)p_{\hat{x}_2|x}(\hat{x}_2=1|x=-1) \\ & - Lp_x(x=-1)p_{\hat{x}_2|x}(\hat{x}_2=1|x=-1) - p_x(x=-1)p_{\hat{x}_2|x}(\hat{x}_2=-1|x=-1) \\ = & -p_x(x=-1)(1 + (L-2)p_{\hat{x}_2|x}(\hat{x}_2=1|x=-1)) \end{aligned}$$

因此, 当 $L=2$ 时, $\gamma = -\frac{p_x(x=1)}{p_x(x=-1)}$, 与 \hat{x}_2 无关; 同理求出 $\gamma' = \gamma$.

3.4. 上机作业

3.5. Thanks to 陆石, who gives me this template.