lme4 extras

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May 26, 2012

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| | Zero-inflation via the EM algorithm This vignette is intended to document some extra tricks that can be us | |
| with lme4 models. Some of them are included here because they are statistically | | |
| non-rigorous and we didn't want to build them into automatic functions that | | |
| could be applied unthinkingly, but recipes are supplied here for use at your | | |
| own risk and assuming that you know what you're doing | | |

1 To do

In principle, we should be able to get confidence intervals on parameters and confidence intervals on predictions via (1) quadratic/Wald approximation (ignoring uncertainty of θ , and possibly of u, for CIs of prediction); (2) cheesy MCMC (in this case, for LMMs, we need a way to retrieve new values of sigma and the fixed effects conditional on θ); (3) parametric bootstrap.

The basic machinery for this is (1) functions for converting among parameterizations of the random effects, i.e. from a (standard deviation, correlation) vectors to θ (concatenated Cholesky-factor vector); (2) a way to extract a deviance function from a fit (i.e. mkdevfun) and (3) a way to simulate values from a fit (i.e. simulate), along with basic components (matrix inversion, etc.).

We'll see how far I get.

2 Fit basic models

In this section we simply fit a few basic models to use as examples later on.

3 Quadratic confidence intervals on random effects parameters

Extract the deviance function and the ML (or REML) parameters:

```
fm1Fun <- update(fm1,devFunOnly=TRUE)
fm1_thpar <- getME(fm1,"theta")</pre>
```

Extract internal functions for converting (standard deviation, correlation) vectors to θ (concatenated Cholesky-factor) parameterization, and vice versa (this is temporary, until we finalize the definitions/names of these functions...)

```
Sv_to_Cv <- lme4:::Sv_to_Cv ## standard dev vector to cholesky
vector
Cv_to_Sv <- lme4:::Cv_to_Sv ## vice versa</pre>
```

Test the round-trip of these functions: does converting from standard deviation scale to Cholesky scale, and back, work?

```
fm1_spar <- Cv_to_Sv(fm1_thpar,s=sigma(fm1))
all(abs(Sv_to_Cv(fm1_spar,s=fm1_spar[4])-fm1_thpar)<1e-6)
## [1] TRUE</pre>
```

A wrapper around the deviance function that we extracted.

```
fm1FunS <- function(spar) {
   thpar <- Sv_to_Cv(c(spar,NA),s=fm1_spar[4])
   fm1Fun(thpar)
}</pre>
```

Use the numDeriv package to compute the Hessian (second derivative) matrix at the MLE:

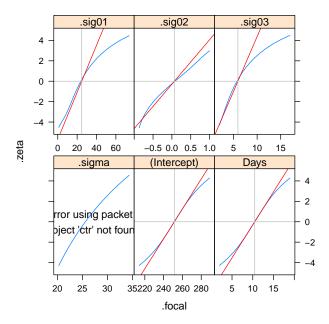
```
library(numDeriv)
h <- hessian(fm1FunS,fm1_spar[-4])</pre>
```

Variance-covariance matrix of the random-effects (standard deviation) parameters:

```
vcov_ran <- solve(h)
```

Compare profiles to their quadratic approximations, and profile confidence intervals to these approximate (Wald) confidence intervals:

pp <- profile(fm1)</pre>



FIXME: These slopes are not quite right. Why? (Are the profiles correct?) Should we be recomputing σ for each set of θ values? Think about this sometime when I have a brain ...

Profile confidence intervals:

```
(ci_prof <- confint(pp))</pre>
##
                   2.5 % 97.5 %
## .sig01
                 14.3815
                          37.716
## .sig02
                 -0.4815
                           0.685
## .sig03
                  3.8012
                           8.753
## .sigma
                 22.8983 28.858
## (Intercept) 237.6807 265.130
## Days
                  7.3587
                          13.576
```

Wald confidence intervals:

```
c(fm1_spar,fixef(fm1))+
    1.96*outer(c(sqrt(diag(vcov_ran)),
                 sqrt(diag(as.matrix(vcov(fm1))))),
               c(-1,1)
                   [,1]
                            [,2]
##
##
                16.2243 33.257
                -0.3849
##
                         0.516
##
                 4.0969
                          7.747
##
                     NA
                             NA
## (Intercept) 238.0289 264.781
## Days
                 7.4375 13.497
```

4 Approximate confidence intervals on predictions

5 Poor man's MCMC

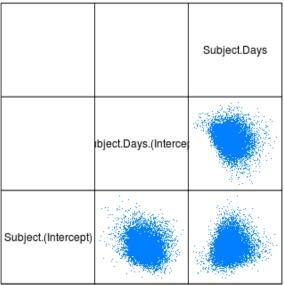
```
library(MCMCpack)
```

MCMCpack expects a function that gives a value proportional to the log posterior density for any specified set of parameters. We can get lme4 to give us a function for the deviance (by using devFunOnly=TRUE. If we assume all-improper priors (i.e. flat on the scale on which we have defined the parameters), then the log posterior density is -D/2:

```
fm1_metropfun <- function(x) {
    ## getME(.,"lower")?
    if (any(x<fm1@lower)) -Inf else -fm1Fun(x)/2
}</pre>
```

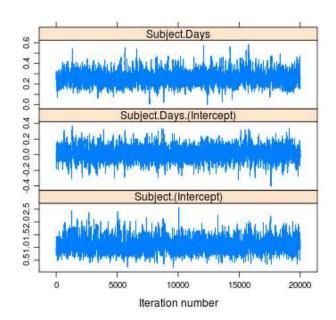
library(coda)

splom(fm1_mcmc_out)

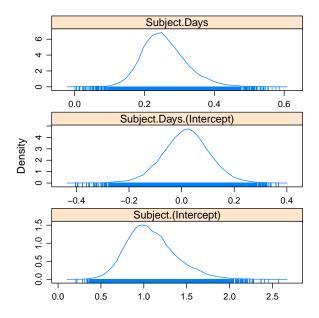


Scatter Plot Matrix

xyplot(fm1_mcmc_out)



densityplot(fm1_mcmc_out,layout=c(1,3))



```
HPDinterval(fm1_mcmc_out)

## lower upper

## Subject.(Intercept) 0.5646 1.6712

## Subject.Days.(Intercept) -0.1693 0.1903

## Subject.Days 0.1416 0.3966

## attr(,"Probability")

## [1] 0.95
```

FIXME: if we want to get this on the sd/corr scale we have to figure out how to recalculate sigma for each set of θ values ... for now, just use a fixed sigma

Highest posterior density intervals:

```
HPDinterval(sdmat)

## lower upper
## 11 14.4493 42.7685

## 12 -0.5457 0.6567

## 13 4.0532 10.5283

## 25.5918 25.5918

## attr(,"Probability")
## [1] 0.95
```

Or quantile-based estimates:

```
t(apply(sdmat,2,quantile,c(0.025,0.975)))

## 2.5% 97.5%

## 11 15.4823 44.2549

## 12 -0.5194 0.6922

## 13 4.2457 10.8785

## 25.5918 25.5918
```

The latter *should* be translation-invariant, and hence (???) the same as:

(Only true for variable 1, although not terribly different: think about this (i.e. the effect of $\theta\theta^T$) some more ...)

If we have the Cholesky form

$$\left(\begin{array}{cc} c_1 & 0 \\ c_2 & c_3 \end{array}\right)$$

and take the cross-product, we get

$$\left(\begin{array}{cc}
c_1^2 & c_1c_2 \\
c_1c_2 & c_2^2 + c_3^2
\end{array}\right)$$

so it's natural that only element 1 scales as we would expect: all the other terms are not just scale translations of a single element, but combinations of multiple elements.

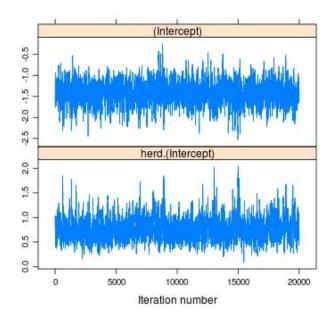
Should work for GLMMs as well:

```
gm1Fun <- update(gm1,devFunOnly=TRUE)
gm1_par <- c(getME(gm1,"theta"),fixef(gm1))
nt <- length(getME(gm1,"theta"))
gm1_metropfun <- function(x) {
   if (any(x[seq(nt)] < gm1@lower)) return(-Inf)
   r <- try(gm1Fun(x),silent=TRUE)
   if (inherits(r,"try-error")) return(-Inf)
   -r/2
}</pre>
```

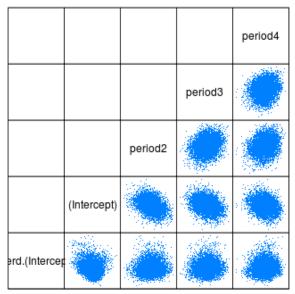
FIXME: PIRLS failure, restore when this is fixed

```
set.seed(101)
gm1_mcmc_out <- MCMCmetrop1R(gm1_metropfun,gm1_par)</pre>
```

```
colnames(gm1_mcmc_out) <- names(gm1_par)
xyplot(gm1_mcmc_out[,1:2])</pre>
```



splom(gm1_mcmc_out)



Scatter Plot Matrix

```
HPDinterval(gm1_mcmc_out)

## lower upper
## herd.(Intercept) 0.3505 1.2576
## (Intercept) -2.0225 -0.9204
## period2 -1.6026 -0.3795
## period3 -1.7782 -0.5182
## period4 -2.5010 -0.7952
## attr(,"Probability")
## [1] 0.95
```

6 Confidence intervals on predictions etc. via parametric bootstrap

FIXME: do we want a as.data.frame.boot function to retrieve stuff from bootMer output?

7 Zero-inflation via the EM algorithm

The zipme function, adapted from code by Mihoko Minami and Cleridy Lennert, is available . . . its form is

Where:

cformula the conditional formula, i.e. the glmer formula for the GLMM part of the model

zformula the zero-inflation formula, i.e. the glm formula for the zero-inflation probability. The response variable *must* be specified as ..z

cfamily the family variable for the conditional part of the model

data a data frame

maxitr the maximum number of EM iterations

tol convergence tolerance

verbose print out information on EM iterations?

An example of EM use:

```
set.seed(101)
zprob <- 0.2
nblock <- 20
ntot <- 500
nperblock <- 25
d <-
data.frame(x=runif(ntot),f=factor(rep(1:nblock,each=nperblock)))
u <- rnorm(nblock,sd=0.75)
d$eta <- with(d,1+2*x+u[as.numeric(f)])
d$resp <- ifelse(runif(ntot) < zprob,0,rpois(ntot,exp(d$eta)))</pre>
```

```
zfit1 <- zipme(resp~x+(1|f),..z~1,data=d,verbose=FALSE)</pre>
```

FIXME: suppress non-integer #successes in a binomial glm! warnings ... preferably without suppressing other possible warnings. Will it be possible to adapt this for zero-inflated binomial GLMMs, or will the dual use of weights get in the way?