S&DS 551 Midterm Cheat Sheet

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Random Walks

- $-S_n: +S_0 + \sum_{i=1}^n X_i$
- T stopping time boundary a < b
- $\mathbb{P}\{T = \infty \mid S_0 = i\} = 0$
- $-\mathbb{P}\{S_T = b \mid S_0 = i\} = \frac{i-a}{b-a}$
- $-\mathbb{E}[T \mid S_0 = i] = (i a)(b i)$
- $-T_a := \min\{n : S_n = a\}$
- For 0 = a < b, $\{S_T = 0\} \subset \{T_0 < \infty\}$
- $-\mathbb{P}\{T_0 < \infty \mid S_0 = i\} \ge \mathbb{P}\{S_T = 0 \mid S_0 = i\} = 1 \frac{i}{b} \to 1$
- $-\mathbb{E}[T_0 \mid S_0 = i] = \lim_{b \to \infty} \mathbb{E}[T \mid S_0 = i] \to \infty \cdot \delta_{i,0}$

Branching Processes

- $-X_0 = 1, X_{n+1} = \sum_{i=1}^{X_n} X_{n,i} \text{ where } X_{n,i} : \Omega \to \mathbb{N} \text{ is iid}$
- $-\mathbb{E}[X_{n+1}] = \mu, \mathbb{E}[X_n] = \mu^{n+1}, \mu = \mathbb{E}[X_{n,i}]$
- $-\eta_n := \mathbb{P}\{X_n = 0\}$
- $\eta_{n+1} = \sum_{k \ge 0} p_k \eta_n^k =: g(\eta_n)$ $\eta = g(\eta)$ fixed point
- $-q(1) = 1, q'(1) = \mu, q''(x) > 0$

Poisson Processes

- $Z \sim Po(\lambda)$
- $\mathbb{P}\{Z = k\} = e^{-\lambda} \lambda^k / k!$
- $\mathbb{E}[Z] = \lambda$, $\operatorname{Var}[Z] = \lambda$
- Counting Process:
- 1) $N_0 = 0, 2$ N_t increasing in t, 3 Increase at most 1
- Poisson Process
- 1) $N_t \sim \text{Po}(\lambda t)$, 2) independent increments, 3) $N_t N_s \stackrel{d}{=}$ N_{t-s}
- T_k : k-th arrival time

- N_t : count at time t
- $-\{T_k < t\} \equiv \{N_t > k\}$
- $T_k \sim \Gamma(k,\lambda)$
- $\mathbb{E}[T_k] = k/\lambda$, $\operatorname{Var}[T_k] = k/\lambda^2$
- $N_t^{(i)}$ iid Poisson process with params λ_i , $\sum_{i} N_{t}^{(i)}$ Poisson process with param $\sum_{i} \lambda_{i}$
- R_t Poisson process rate λ ,

subsample k-th arrival with prob $p \in \Delta^k$,

 $R_t^{(1)}, \ldots, R_t^{(k)}$ ind. Poisson process with params $p_i \lambda$

- $X \sim \text{Exp}(\mu), Y \sim \text{Exp}(\gamma) \text{ ind.}, \mathbb{P}\{X \wedge Y = X\} = \mu/(\mu + \gamma)$
- N_t Poisson process param $\lambda(t)$ if
- 1) independent increments, 2) $N_t N_s \sim \text{Po}\left(\int_s^t \lambda(u)du\right)$ for $t \ge s$

Markov Chains

- Time-homogeneous Markov chain,
- $\mathbb{P}\{X_{n+1} = j \mid \forall 0 < k < n, X_k = i_k\}$
- $= \mathbb{P}\{X_{n+1} = j \mid X_n = i_n\}$
- thm: $P_{ij}(s) = \delta_{ij} + F_{ij}(s)P_{ij}(s)$
- recurrent state i if $\mathbb{P}_i\{\exists n \geq 1, X_n = i\} = 1$
- <u>lem</u>: If $\sum_{n} P_{jj}(n) = \infty$,
- a) j is recurrent, b) $\sum_{n} P_{ij}(n) = \infty$ for all $i f_{ij} > 0$
- trasient state i if $\mathbb{P}_i\{\exists n \geq 1, X_n = i\} < 1$ lem: If $\sum_{n} P_{ij}(n) < \infty$
- a) j is transient, b) $\sum_{n} P_{ij}(n) < \infty$ for all i
- If j transient, $\lim_{n} P_{ij}(n) = 0$ for all i
- recurrent j null-recurrent if $\mathbb{E}_i \min\{n \geq 1 : X_n = j\} = \infty$, else positive-recurrent
- j null-recurrent $\iff P_{ii}(n) \to 0$
- period state $j: d(j) = \gcd\{n: P_{ij}(n) > 0\}$
- i aperiodic if d(i) = 1
- i communicates with $j, i \to j$ if $\exists m, P_{ij}(m) > 0$
- thm: If $i \leftrightarrow j$, 1) d(i) = d(j), 2) i transient $\iff j$ 3) i Generating func of $X : \Omega \to \mathbb{N}$, null-recurrent $\iff i$
- $C \subseteq S$ is *closed* if cannot transition out,

- irreducible if $i \leftrightarrow j \ \forall i, j \in C$ lem: If $|S| < \infty$, a) ≥ 1 recurrent state, b) all recurrent states are positive
- thm: Irreducible MC has stationary distribution ←⇒ all states positive recurrent, then $\pi_i = 1/\mu_i$ uniquely (mean recurrent time)
- thm: Irreducible, aperiodic MC satisfies $P_{ij}(n) \to 1/\mu_i$
- thm: Irrreducible, finite, aperiodic MC has $\lambda_1 = 1$ and $|\lambda_r| < 1$ for all r > 1
- $\forall T \subseteq S, \sum_{i \in T, j \notin T} \pi_i P_{ij} = \sum_{i \in T, j \notin T} \pi_j P_{ji}$
- Irreducible, aperiodic, positive-recurrent MC is ergodic
- thm: RW on G aperiodic \iff G not bipartite
- thm: RW on conneted non-bipartite graph $\rightarrow \pi_v =$ d(v)/(2|E|)
- X_n Irreducible, positive-recurrent MC, stationary π , $X_0 \sim$ π , reversed chain $Y_n := X_{N-m}$
- Y_n is MC with transition $Q_{ij} = \pi_i P_{ji}/\pi_i$
- X_n reversible if $\pi_i P_{ii} = \pi_i P_{ij}$
- X_n reversible wrt π then π stationary distribution

MCMC

test

Useful Facts

- $-\sum_{i=1}^{n} i = n(n+1)/2$
- $-\sum_{i=1}^{n} i^2 = (n+1)(2n+1)/6$
- For $N:\Omega\to\mathbb{N}$,
- $\mathbb{E}[N] = \sum_{n>1} \mathbb{P}\{N \ge n\}$
- $-\mathbb{E}[N^2] = \sum_{n>1}^{-} \mathbb{P}\{N \ge n\} \left[n^2 (n-1)^2\right]$
- Catalan number $C_n = \frac{1}{n+1} {2n \choose n} = \frac{(2n)!}{(n+1)!n!} = \prod_{k=2}^n \frac{n+k}{k}$
- Convolution of seq a_n, b_n is $c_n := \sum_{i=0}^n a_i b_{n-i}$ generating func $G_c(s) = G_a(s) \cdot G_b(s)$
- $G_X(s) := \mathbb{E}[s^X] = \sum_{i>0} \mathbb{P}\{X=i\}s^i$
- Sum of MC not necessarily a MC