

## Lecture 13

*Lecturer: Andre Wibisono**Scribe: Tatsuhiko Shimizu*

## 1 Instruction

In the last lecture, we have seen the rejection sampling which can take exponential time to sample from the target distribution  $\nu(x)$  in high dimensions. Today's lecture tries to improve the sampling speed by introducing Metropolis Random Walk and Metropolis-Hasting Algorithm.

## 2 Metropolis Random Walk (MRW)

Suppose we have target distribution  $\nu(x) \propto e^{-f(x)}$ .

**Algorithm 1** Metropolis Random Walk (MRW)

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Given: We start from  $x_0 \sim \rho_0$  on  $\mathcal{X}$ , given step size  $\eta > 0$ .

**for**  $k = 0, 1, \dots, K$  **do**:

$$y_k \leftarrow x_k + \sqrt{\eta} z_k, \quad z_k \sim \mathcal{N}(0, I)$$

Accept  $y_k$  with probability

$$\min \left\{ 1, \frac{\nu(y_k)}{\nu(x_k)} \right\}.$$

**if** Accept **then**

$$x_{k+1} = y_k$$

**else**

$$x_{k+1} = x_k$$

**end if**

**end for**

Return  $x_K \sim \rho_K$ .

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Note that  $x_k \in \mathbb{R}^d$  is the random variable with distribution  $\rho_k \in P(\mathbb{R}^d)$ .

**Example 1** (Ball Walk). Suppose target  $\nu$  is the uniform distribution on convex body  $\mathcal{X} \subset \mathbb{R}^d$ .

Our goal is to sample uniformly from  $\mathcal{X}$ . It is mainly applicable for estimating the volume of  $\mathcal{X}$  practically.

We sample by using MRW. For  $k = 0, 1, \dots, K$ , Let

$$y_k = x_k + \sqrt{\eta} z_k \quad z_k \sim \mathcal{N}(0, I)$$

If  $y_k \in \mathcal{X}$ , then set  $x_{k+1} = y_k$ . Otherwise, set  $x_{k+1} = x_k$ .

For this example, we can say that the sampling speed is more efficient than rejection sampling but to show that, we introduce the notion of isotropic.

**Definition 1.** A convex body  $\mathcal{X} \subseteq \mathbb{R}^d$  is in isotropic position if its center of mass is at 0, i.e.,

$$\mathbb{E}_\nu[X] = 0$$

and its covariance matrix is the identity

$$\text{Cov}_\nu[X] = I$$

where  $\nu$  here is the uniform distribution over  $\mathcal{X}$ . Equivalently, for any unit vector  $\|v\| = 1$ ,

$$\frac{1}{\text{Vol}(\mathcal{X})} \int_{\mathcal{X}} (v^\top x)^2 dx = 1.$$

In general, we can say that a distribution  $\nu$  (not necessarily the uniform distribution) is in isotropic position if it satisfies the above conditions, where the expectations are taken with respect to  $\nu$  (e.g., see Bertsimas and Vempala 2004; Rudelson 1999). Then, using the notion of isotropic, we have the following statement about the sampling speed.

**Theorem 1.** If  $\mathcal{X}$  is isotropic, then the MRW is a polytime algorithm.

For ball walk to produce  $\rho_k$  with  $TV(\rho_k, \nu) \leq \epsilon$ , we need  $K = \tilde{O}(d^2)$  where  $\tilde{O}$  hides the logarithm terms and warm start, that is  $\sup_{x \in \mathcal{X}} \frac{\rho_0(x)}{\nu(x)} \leq M$  where  $M$  is a constant.

### 3 Metropolis-Hastings Algorithm (MH)

We now introduce a generalization of MRW, called the **Metropolis-Hastings** algorithm. Suppose we have a target distribution  $\nu$  on  $\mathcal{X}$ . In this algorithm, we need to choose  $p(x | y)$ .

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**Algorithm 2** Metropolis-Hasting Algorithm (MH)

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Given: We start from any  $x_0 \sim \rho_0$  on  $\mathcal{X}$ .

**for**  $k = 0, 1, \dots, K$  **do**:

    Draw

$$y_k \mid x_k \sim p(y_k \mid x_k)$$

    Accept  $y_k$  with probability

$$\min \left\{ 1, \frac{\nu(y_k)p(x_k \mid y_k)}{\nu(x_k)p(y_k \mid x_k)} \right\}$$

**end for**

Return  $x_K$

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Note that MRW is a special case of the MH algorithm where

$$p(y \mid x) = \mathcal{N}(x, \eta I)(y)$$

$$p(y \mid x) = p(x \mid y)$$

Question: why does the MH algorithm work, that is  $x_K \sim \rho_K \rightarrow \nu$  as  $k \rightarrow \infty$ ? The short answer is that this Markov Chain is reversible on  $\mathcal{X}$ . To answer the question mathematically, we first check the definition of the Markov Chain.

**Definition 2.** *Markov Chain (MC) on  $\mathcal{X}$  is specified by a family of probability distributions*

$$Q = \{Q_x = p(\cdot \mid x) \mid x \in \mathcal{X}\}$$

Then, given  $x_0 \in \rho_0$  on  $\mathcal{X}$ , we can get Markov Chain  $X_0, X_1, X_2, \dots$  where

$$x_{k+1} \mid x_k \sim Q_{x_k}(x_{k+1}) = p(x_{k+1} \mid x_k).$$

Thus, MC  $Q = \{Q_x : x \in \mathcal{X}\}$  defines a map

$$Q : p(\mathcal{X}) \rightarrow p(\mathcal{X})$$

by

$$\rho_{k+1}(y) = \int_{\mathcal{X}} \rho_k(x)p(y \mid x)dx$$

Now we can introduce the stationary and reversible properties.

**Definition 3.** *A probability distribution  $\nu$  is stationary for the MC defined by*

$$Q = \{Q_x : x \in \mathcal{X}\}$$

*if*

$$Q(\nu) = \nu,$$

equivalently, that is

$$\nu(y) = \int_{\mathcal{X}} \nu(x)p(y|x)dx$$

**Definition 4.** A MC  $Q$  is reversible with respect to some  $\nu \in P(\mathbb{R}^d)$  if

$$\nu(x)p(y|x) = \nu(y)p(x|y) \quad \forall x, y \in \mathcal{X},$$

which is equivalent to that the two joint distributions are symmetric

$$\rho(x, y) = \rho(y, x) \quad \rho(x, y) := \nu(x)p(y|x) \quad \forall x, y \in \mathcal{X}$$

Then we have the lemma about the relationship between two properties.

**Lemma 1.** If  $Q$  is reversible with respect to  $\nu$ , then  $\nu$  is stationary for  $Q$ .

*Proof.* By reversibility, for all  $x, y \in \mathcal{X}$ , we have

$$\nu(x)p(y|x) = \nu(y)p(x|y)$$

Integrating over  $\mathcal{X}$ , we have

$$\begin{aligned} \int_{\mathcal{X}} \nu(x)p(y|x)dx &= \int_{\mathcal{X}} \nu(y)p(x|y)dx \\ \int_{\mathcal{X}} \nu(x)p(y|x)dx &= \nu(y) \int_{\mathcal{X}} p(x|y)dx \\ Q(\nu)(y) &= \nu(y) \end{aligned}$$

□

Then we have a lemma about the properties of MH algorithm.

**Lemma 2.** Let  $Q = \{Q_x = p(\cdot|x); \quad x \in \mathcal{X}\}$  be an arbitrary MC. If we apply MH to get another MC  $\hat{Q}$ , then  $\hat{Q}$  is reversible with respect to  $\nu$ .

We define  $\alpha_x(y)$  as the acceptance probability in MH algorithm.

$$\alpha_x(y) = \min \left\{ 1, \frac{\nu(y)p(x_k|y_k)}{\nu(x_k)p(y_k|x_k)} \right\}$$

Now our question is what  $\hat{Q}$  is.

$$\hat{Q}_x(y) = \underbrace{Q_x(y)}_{\text{original MC}} \times \underbrace{\alpha_x(y)}_{\text{acceptance probability}} + \underbrace{\delta_x(y)}_{\text{indicator}} \times \underbrace{A(x)}_{\text{probability distribution on } \mathcal{X}}$$

where

$$\begin{aligned} A(x) &= Q_x(x) + \int_{\mathcal{X} \setminus \{x\}} (1 - \alpha_x(y))Q_x(y)dy \\ &= 1 - \int_{\mathcal{X} \setminus \{x\}} \alpha_x(y)Q_x(y)dy \end{aligned}$$

**Lemma 3.** *Given MC  $Q$ , let  $\hat{Q} = MH_\nu(Q)$ . Then  $\nu$  is reversible with respect to  $\hat{Q}$ . Thus,  $\nu$  is stationary for  $\hat{Q}$ .*

*Proof.* We want to show that

$$\nu(x)\hat{Q}_x(y) = \nu(y)\hat{Q}_y(x)$$

If  $x = y$ , then the lemma holds obviously, so suppose  $x \neq y$ . Then,

$$\begin{aligned} \nu(x)\hat{Q}_x(y) &= \nu(x)Q_x(y)\alpha_x(y) \\ &= \nu(x)Q_x(y) \min \left\{ 1, \frac{\nu(y)Q_y(x)}{\nu(x)Q_x(y)} \right\} \\ &= \min \{ \nu(x)Q_x(y), \nu(y)Q_y(x) \} \\ &= \nu(y)\hat{Q}_y(x) \end{aligned}$$

□

## References

- Bertsimas, Dimitris and Santosh Vempala (July 2004). “Solving Convex Programs by Random Walks”. In: *J. ACM* 51.4, pp. 540–556. ISSN: 0004-5411. DOI: 10.1145/1008731.1008733. URL: <https://doi.org/10.1145/1008731.1008733>.
- Rudelson, M. (1999). “Random Vectors in the Isotropic Position”. In: *Journal of Functional Analysis* 164.1, pp. 60–72. ISSN: 0022-1236. DOI: <https://doi.org/10.1006/jfan.1998.3384>. URL: <https://www.sciencedirect.com/science/article/pii/S0022123698933845>.