# Survival Analysis

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#### November 1, 2022

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# 1 Basic knowledge

#### 1.1 Survival and hazard

Let T denote the time to an event that we are interested in. Then we know the c.d.f.

$$F_T(t) = P(T < t)$$
,

and the corresponding p.d.f.

$$f_T(t) = \frac{\mathrm{d}}{\mathrm{d}t} F_T(t).$$

Here to simplify the discussion, we assume T is a continuous random variable. In the context of survival analysis, the *event* often refers to death. Then T represents the lifespan of the subject. So  $F_T(t)$  represents the probability that the death occurs before t. In another word, we know the probability that the subject survives passes t is

$$S_T(t) = 1 - F_T(t) = P(T > t).$$

 $S_T(t)$  is often called the survival function? and clearly

$$f_T(t) = -\frac{\mathrm{d}}{\mathrm{d}t} S_T(t).$$

The **hazard function** h(t) is defined as

$$h\left(t\right) = \lim_{\Delta \to 0} \frac{P\left(T \le t + \Delta \middle| T > t\right)}{\Delta} = \lim_{\Delta \to 0} \frac{F_T\left(t + \Delta\right) - F_T\left(t\right)}{\Delta \cdot S_T\left(t\right)} = \frac{f_T\left(t\right)}{S_T\left(t\right)}.$$

h(t) represents the instant hazard? unified probability? that the subject will be dead instantly after t given the fact that it's alive at t. And the **cummulative hazard function** is

$$H(t) = \int_0^t h(x) dx = \int_0^t \frac{f_T(x)}{S_T(x)} dx = \int_0^t \frac{-dS_T(x)}{S_X(t)} = -\log(S_T(x))|_0^t = -\log(S_T(t)).$$

**Proposition 1.** The random variable H(T) follows unit exponential distribution EXP(1).

Proof.

$$\begin{split} P\left(H\left(T\right) \leq t\right) = & P\left(-\log S\left(T\right) \leq t\right) \\ = & P\left(1 - F\left(T\right) \geq e^{-t}\right) \\ = & P\left(T \leq F^{-1}\left(1 - e^{-t}\right)\right) \\ = & F\left(F^{-1}\left(1 - e^{-t}\right)\right) \\ = & 1 - e^{-t}, \end{split}$$

which is the c.d.f of EXP(1). Here to simplify the deduction we make some assumptions that

- F(t) is continuous.
- $F^{-1}(t)$  is well defined.

Also to simplify the notation and avoid confusion, we use  $S(\cdot)$  and  $F(\cdot)$  instead of  $S_T(\cdot)$  and  $F_T(\cdot)$  like before.

1. Exponential distribution: Denote  $T \sim EXP(\lambda)$ . Then

$$f(t) = \lambda e^{-\lambda t}$$

$$F(t) = 1 - e^{-\lambda t} \qquad S(t) = e^{-\lambda t}$$

$$h(t) = \lambda \qquad \text{constant hazard}$$

$$H(t) = \lambda t$$

$$E(T) = 1/\lambda \qquad Var(T) = 1/\lambda^2$$

2. Weibull distribution: Denote  $T \sim W(p, \lambda)$ . Then

$$f(t) = p\lambda^{p}t^{p-1}\exp\left(-\left(\lambda t\right)^{p}\right)$$

$$F(t) = 1 - \exp\left(-\left(\lambda t\right)^{p}\right) \qquad S(t) = \exp\left(-\left(\lambda t\right)^{p}\right)$$

$$h(t) = p\lambda^{p}t^{p-1}$$

$$H(t) = \left(\lambda t\right)^{p}$$

$$E(T) = \frac{1}{\lambda} \cdot \Gamma\left(1 + \frac{1}{p}\right) \qquad \operatorname{Var}(T) = \frac{1}{\lambda^{2}}\left(\Gamma\left(1 + \frac{2}{p}\right) - \Gamma\left(1 + \frac{1}{p}\right)\right)$$

$$E(T^{m}) = \frac{1}{\lambda^{m}}\Gamma\left(1 + \frac{m}{p}\right)$$

#### 1.2 Censor

#### 1.2.1 Right censor

• Type I: an i.i.d sample  $T_1, \dots, T_n \sim F$  and a fixed constant c. And the observed data is  $(U_i, \delta_i)$  for  $i = 1, \dots, n$  where

$$U_i = \min (T_i, c)$$
$$\delta_i = 1_{T_i \le c}.$$

So the observed data consists of a random number, r, of uncensored observations, all of which are less than c. And n-r censored observations, all are c.

• Type II: an i.i.d sample  $T_1, \dots, T_n \sim F$  and a pre-defined number of failure r. The observation is stopped when r failure occurs and the stopping time is c. The observed data is still the form  $(U_i, \delta_i)$  for  $i = 1, \dots, n$ , the same as that in Type I censor. But in actuality, we observe the first r order statistics

$$T_{(1,n)},\cdots,T_{(r,n)}$$
.

Note that here  $(U_1, \delta_1), \dots, (U_n, \delta_n)$  are dependent whereas they are independent for Type I.

• Type III (Random censor): The underlying data is

$$c_1, \dots, c_n$$
 constant  $T_1, \dots, T_n \sim F$ .

And the observed data is  $(U_i, \delta_i)$  for  $i = 1, \dots, n$ , where

$$U_i = \min (T_i, c_i)$$
  
$$\delta_i = 1_{T_i < c_i}.$$

**Note:** for inference,  $c_i$  is often treated as constant. For study design or studying the asymptotic property, they are often treated as i.i.d random variables  $C_1, \dots, C_n$ .

#### 1.2.2 Left censor

 $T_i$  is censored when  $T_i \leq l_i$ .

#### 1.2.3 Interval censor

 $l_i \leq T_i \leq u_i$ , but only  $l_i$  and  $u_i$  are observed.

## References