# Conditional Power

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## 1 Introduction

The conditional power is defined as the probability to reject the null hypothersis at **Final** analysis given the current collected data.

## 2 Mean Test

## 2.1 Normal distribution, variance known

#### 2.1.1 One-sample mean test

Let the sample  $x_1, \dots, x_n \sim N(\mu, \sigma^2)$ , where  $\sigma^2$  is known. To simplify the discussion here, we assume an one-sided test

$$H_0: \mu \leq \mu_0, \quad H_1: \mu > \mu_0$$

The reject rule is rejecting  $H_0$  if  $\bar{x} > c$  for some c. Since  $\bar{x} \sim N(\mu, \sigma^2/n)$ , c can be determined by type-I error  $\alpha$ :

$$P\left(\bar{x} > c | H_0\right) < \alpha.$$

By Neyman-Pearson Lemma, and stochastical monotone property of normal distribution, to fully utilize the type-I error, it is

$$P\left(\bar{x} > c \middle| \mu = \mu_0\right) = P\left(\frac{\bar{x} - \mu_0}{\sqrt{\sigma^2/n}} > \frac{c - \mu_0}{\sqrt{\sigma^2/n}}\right) = P\left(Z > \frac{c - \mu_0}{\sqrt{\sigma^2/n}}\right) = \alpha.$$

Hence  $c = \mu_0 - z_\alpha \sqrt{\sigma^2/n}$ , where  $z_\alpha = \Phi^{-1}(\alpha)$  is the (left-) $\alpha$  quantile of normal distribution. For more details about how this test is constructed, one can refer to the **ttest.pdf** notes in this repo.

The probability to reject the null hypothersis at some given  $\mu_1$  can be computed as

$$P(\bar{x} > c | \mu = \mu_1)$$

$$= P\left(\frac{\bar{x} - \mu_1}{\sqrt{\sigma^2/n}} > \frac{c - \mu_1}{\sqrt{\sigma^2/n}}\right)$$

$$= P\left(Z > -z_\alpha + \frac{\mu_0 - \mu_1}{\sqrt{\sigma^2/n}}\right)$$

$$= 1 - \Phi\left(-z_\alpha + \frac{\mu_0 - \mu_1}{\sqrt{\sigma^2/n}}\right) = 1 - \beta,$$

where  $\beta$  is the type-II error. Then the samplesize is determined by

$$n = \frac{\sigma^2 (z_{\alpha} + z_{\beta})^2}{(\mu_0 - \mu_1)^2}.$$

Now, assume we have collected  $n_1$  samples, which means there are  $n_2 = n - n_1$  more samples to come. And the current statistics is  $\bar{x}_1 = \frac{1}{n_1} \sum_{i=1}^{n_1} x_i$  and the final statistics is  $\bar{x} = \frac{1}{n_1 + n_2} \sum_{i=1}^{n_1 + n_2} x_i$ . The statistics from the second part is  $\bar{x}_2 = \frac{1}{n_2} \sum_{i=n_1+1}^{n_1+n_2} x_i$ . Note that these  $(\bar{x}_1, \bar{x}_2 \text{ and } \bar{x})$  are all sufficient statistics for  $\mu$  (based on part1, part1 and all parts). Here we know

$$\bar{x}_1 \sim N\left(\mu, \sigma^2/n_1\right), \quad \bar{x}_2 \sim N\left(\mu, \sigma^2/n_2\right), \quad \bar{x}_1 \perp \bar{x}_2.$$

And  $\bar{x} = \frac{n_1}{n_1 + n_2} \bar{x}_1 + \frac{n_2}{n_1 + n_2} \bar{x}_2$ . So the conditional power, which is the probability to reject the hypothesis at final analysis given the current collected data is

$$P(\bar{x} > c | \bar{x}_1)$$

$$= P\left(\frac{n_1}{n_1 + n_2} \bar{x}_1 + \frac{n_2}{n_1 + n_2} \bar{x}_2 > c | \bar{x}_1\right)$$

$$= P\left(\bar{x}_2 > \frac{(n_1 + n_2) c - n_1 \bar{x}_1}{n_2}\right)$$

$$= P\left(\frac{\bar{x}_2 - \mu}{\sqrt{\sigma^2/n_2}} > \frac{\frac{(n_1 + n_2) c - n_1 \bar{x}_1}{n_2} - \mu}{\sqrt{\sigma^2/n_2}}\right)$$

$$= P\left(Z > \frac{-z_{\alpha} \sqrt{(n_1 + n_2) \sigma^2} + n_1 (\mu_0 - \bar{x}_1) + n_2 (\mu_0 - \mu)}{\sqrt{n_2 \sigma^2}}\right)$$

$$= 1 - \Phi\left(\frac{-z_{\alpha} \sqrt{(n_1 + n_2) \sigma^2} + n_1 (\mu_0 - \bar{x}_1) + n_2 (\mu_0 - \mu)}{\sqrt{n_2 \sigma^2}}\right)$$

**Note:** in previous discussions there are some abuse of the notation.  $\bar{x}_1$  and  $\bar{x}_2$  can both refer to the random variable, and there realization. But I hope the context is clear and not much confusion.

- ${\bf 2.1.2}\quad {\bf Two\text{-}sample\ mean\ test}$
- 3 Rate Test
- 4 Log Rank Test