

Lecture 24. MLE

Wednesday, December 6, 2017

8:52 AM

Bug in HW8 Solutions for (d) corrected

$$f'(x) = \frac{f(x+2a_n) - f(x)}{2a_n} + o(a_n) \quad \text{Numerical analysis}$$


$$f'(x) = \frac{f(x+a_n) - f(x-a_n)}{2a_n} + o(a_n^2)$$

The asymptotic IN of Z-estimators framework of Lecture 23 does not work for the median b/c ψ' does not always exist (sign is not differentiable everywhere)

(DF) a function f is called Lipschitz if

$$|f(x_1) - f(x_2)| \leq C \cdot \|x_1 - x_2\|$$

(TH) Relaxing differentiability requirements

-]
- $m_\theta(x)$ is measurable as a function of x and differentiable at θ^* , P -a.s. wrt x
 - $m_\theta(x)$ is Lipschitz wrt θ , i.e. $\exists C := \dot{m}(x)$
 - $\dot{m}(x)$ is measurable and $P \dot{m}^2 < \infty$
 - P_{m_θ} has a 2nd order Taylor expansion at θ^* with the 2nd derivative matrix V_{θ^*}
 and 1st derivative $P_{m_\theta}'_{\theta^*}$

$$1 : \theta^* \text{ and } \dots \dots \dots \theta^*$$

$$\bullet \quad P_n m_{\hat{\theta}_n} \geq \sup_{\theta} P_n m_{\theta} - o_p\left(\frac{1}{n}\right), \text{ i.e.}$$

$$\bullet \quad \hat{\theta}_n \text{ is an approx. maximizer of } M_n(\theta) = P_n m_{\theta}$$

$$\bullet \quad \hat{\theta}_n \text{ is consistent for } \theta^*$$

$$\hat{\theta}_n \xrightarrow{P} \theta^*$$

Then $\sqrt{n}(\hat{\theta}_n - \theta) = -V_{\theta^*}^{-1} \underbrace{\sqrt{n} P_n m'_{\theta^*}}_{\uparrow} + o_p(1)$

In particular, using CLT on \uparrow

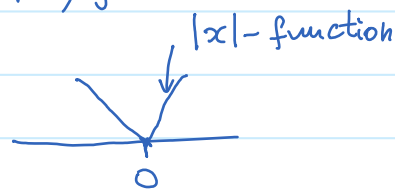
$$\sqrt{n}(\hat{\theta}_n - \theta^*) \rightsquigarrow N(0, V_{\theta^*}^{-1} P(m'_{\theta^*} m'^*_{\theta^*}) V_{\theta^*}^{-1})$$

Proof uses empirical processes

$$(P_n - P)f$$

Example: median

$$M_n(\theta) = -\frac{1}{n} \sum_{i=1}^n |x_i - \theta|$$



$$\Psi_n(\theta) = M'_n(\theta) = -\frac{1}{n} \sum_{i=1}^n \text{sign}(x_i - \theta)$$

$$\text{] pdf, } f, \text{ exists at } \theta^* \Rightarrow \hat{\theta}_n = \arg \max M_n(\theta)$$

$$\sqrt{n}(\hat{\theta}_n - \theta^*) \text{ is asymptotically } N\left(0, \frac{1}{(2f(\theta^*))^2}\right)$$

Proof:

$$|m_{\theta_1}(x) - m_{\theta_2}(x)| = ||x - \theta_1| - |x - \theta_2|| \leq |\theta_1 - \theta_2| \Rightarrow$$

$$\Rightarrow \dot{m}_{\theta}(x) \equiv 1$$

$$m'_{\theta^*}(x) = -\text{sign}(x - \theta^*) \text{ exists a.e. except at } x = \theta^*$$

$$M(\theta) = P m_n = \int |x - \theta| dF(x) =$$

$$\begin{aligned}
 M(\theta) = P_{m_\theta} &= \int |x - \theta| dF(x) = \\
 &= \int_{-\infty}^{\theta} (\theta - x) dF(x) + \int_{\theta+}^{\infty} (x - \theta) dF(x) \quad \begin{array}{l} \uparrow \text{CDF of } X \\ \swarrow \text{integration by parts} \end{array} \\
 &= \theta F(\theta) - \theta [1 - F(\theta)] - \lim_{a \rightarrow \infty} \left[x F(x) \Big|_{-a}^{\theta} - \int_{-a}^{\theta} F(x) dx \right] + \\
 &\quad + \left[x F(x) \Big|_{\theta+}^a - \int_{\theta+}^a F(x) dx \right] = \\
 &= -\theta + \lim_{a \rightarrow \infty} \left[\int_{-a}^{\theta} F(x) dx - \int_{\theta+}^a F(x) dx + 2a F(a) \right]
 \end{aligned}$$

Differentiating under the limit, we have

Note: Differentiation is justified if derivatives converge uniformly wrt a . In our situation this is the case because derivatives do not depend on a .

$$\begin{aligned}
 \left. \frac{d^2 P_{m_\theta}}{d\theta^2} \right|_{\theta^*} &= 2 \cdot f(\theta) \Big|_{\theta^*} = V_{\theta^*} ; & \left. \frac{d P_{m_\theta}}{d\theta} \right|_{\theta^*} &= 2 F(\theta) - 1 \Big|_{\theta^*} = 0 \\
 & & & \uparrow \\
 & & & F(\theta^*) = \frac{1}{2}
 \end{aligned}$$

$$P(\hat{m}_{\theta^*}^2) = P(1) = 1 - \sqrt{n} (\hat{\theta}_n - \theta^*) \rightsquigarrow N\left(0, \frac{1}{(2f(\theta^*))^2}\right)$$

MLE - framework

$$\begin{aligned}
 \ell_n &= \sum_{i=1}^n \log f_\theta(x_i) & \max_{\theta} \ell_n(\theta, \text{data}) &\Rightarrow \hat{\theta}_n \text{ MLE} \\
 M_n(\theta) &= \mathbb{P}_n \log \frac{f_\theta}{f_{\theta^*}} = \frac{1}{n} \sum_{i=1}^n \left[\log f_\theta(x_i) - \log f_{\theta^*}(x_i) \right] \\
 & & & \underbrace{\hspace{10em}}_{\text{doesn't depend on } \theta} \\
 LLN &\Rightarrow M_n(\theta) \rightarrow P \log \frac{f_\theta}{f_{\theta^*}} =: M_\theta
 \end{aligned}$$

$$LLN \Rightarrow M_n(\theta) \rightarrow P \log \frac{f_\theta}{f_{\theta^*}} := M_\theta$$

Jensen inequality

$$P \log \frac{f_\theta}{f_{\theta^*}} \leq \log P \frac{f_\theta}{f_{\theta^*}} = \log \int \frac{f_\theta}{f_{\theta^*}} \cdot f_{\theta^*} dx = 0$$

$P(\cdot)$ is taken over the X_i 's $\sim P_{\theta^*}$

$$\left. \begin{array}{l} \text{i.e. } M_\theta \leq 0 \\ M_{\theta^*} = 0 \end{array} \right\} M_\theta \text{ is maximized by } \theta^*$$

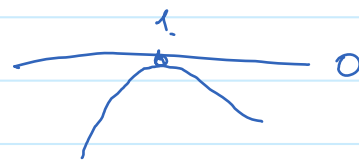
Note: $-M_n(\theta)$ is called Kullback-Leibler divergence

Model identifiability: $f_\theta(\cdot) \equiv f_{\theta^*}(\cdot) \Leftrightarrow \theta = \theta^*$

\Downarrow
 θ^* as $\arg\max M_\theta$ is unique

Proof:

$$\begin{aligned} \text{Consider } \varphi(x) &= \log x - x + 1 \\ \varphi'(x) &= \frac{1}{x} - 1 \end{aligned}$$



$$\Rightarrow \log x \leq x - 1 \quad \text{"=" iff } x = 1$$

$$x \leftarrow \sqrt{\frac{f_\theta}{f_{\theta^*}}}$$

$$\log \frac{f_\theta}{f_{\theta^*}} \leq 2 \left(\sqrt{\frac{f_\theta}{f_{\theta^*}}} - 1 \right)$$

Now:

$$P \log \frac{f_\theta}{f_{\theta^*}} = M_\theta \leq 2 \int \left(\sqrt{\frac{f_\theta}{f_{\theta^*}}} - 1 \right) \cdot f_{\theta^*} dx =$$

$$\begin{aligned}
&= 2 \int \sqrt{f_\theta \cdot f_{\theta^*}} - \underbrace{2}_{\text{"}} = \int (\sqrt{f_\theta})^2 dx + \int (\sqrt{f_{\theta^*}})^2 dx \\
&= - \int (\sqrt{f_\theta} - \sqrt{f_{\theta^*}})^2 dx \leq 0 \Rightarrow \\
&\quad \text{wt} = 0 \text{ iff. } f_\theta \equiv f_{\theta^*} \\
&\quad \text{b/c of identif. } \Leftrightarrow \theta = \theta^* \quad \square
\end{aligned}$$

Note: $\frac{1}{2} \int (\sqrt{f_1} - \sqrt{f_2})^2 dx$ is Hellinger distance between f_1 and f_2