

HW4

1 a)

$Y_{1k}$ : response from subject  $k$  (treatment)

$Y_{2k}$ : "  $\approx$  (placebo)

$$\pi_1 = P(Y_{1k} = 1) \quad \pi_2 = P(Y_{2k} = 1).$$

$$H_0: \pi_1 = \pi_2 \quad \text{vs} \quad H_1: \pi_1 \neq \pi_2$$

Test statistic

$$\chi^2_M = \frac{(18 - 36)^2}{(18 + 36)} = 6$$

Since  $6 > 3.84 = \chi^2_{1, 0.95}$ , we can reject  $H_0$ .

Treatment has different effects on the kidney disease.

b) Since  $X_{1k} = 1$  and  $X_{2k} = 0$ ,  $X_{1k} - X_{2k} = 0$ . For all  $k$ ,

$$L(\beta) = \left\{ \frac{e^\beta}{1+e^\beta} \right\}^{Y_{1k}(1-Y_{2k})} \cdot \left\{ \frac{1}{1+e^\beta} \right\}^{Y_{2k}(1-Y_{1k})}$$

$$l_k(\beta) = Y_{1k}(1-Y_{2k})\beta - (Y_{1k}(1-Y_{2k}) + Y_{2k}(1-Y_{1k})) \ln(1+e^\beta)$$

$$l(\beta) = \sum_{k=1}^{19} l_k(\beta) = 36\beta - 54 \ln(1+e^\beta)$$

$$U(\beta) = \frac{\partial l(\beta)}{\partial \beta} = 36 - 54 \frac{e^\beta}{1+e^\beta} = 0$$

$$e^\beta = \frac{36}{18} = 2$$

$$\beta = \ln 2.$$

$$c) H_0: \beta = 0 \text{ vs } H_1: \beta \neq 0.$$

$$U(\beta) \Big|_{\beta=0} = 36 - \frac{54}{2} = 9.$$

$$J(\beta) = 54 \frac{e^\beta}{(1+e^\beta)^2}$$

$$J(\beta) \Big|_{\beta=0} = \frac{54}{4}$$

$$X_S^2 = U(0) J'(0) U(0) = 6 \sim \chi^2_1$$

Since  $6 > \chi^2_{1,0.95} = 3.84$ , we can reject  $H_0$ .

⇒ The test statistic is identical to the test statistic of the McNemar test.

$$d) H_0: \beta = 0 \text{ vs } H_1: \beta \neq 0.$$

$$X_W^2 = \frac{\hat{\beta}^2}{SE(\hat{\beta})^2} = \frac{(\log 2)^2}{J''(\log 2)} = 5.765 \sim \chi^2_1$$

$$5.765 > 3.84 = \chi^2_{1,0.95}, \text{ reject } H_0.$$

This test is not identical to the McNemar test.

2.

(a)  $\hat{\beta}_{01} = \log\left(\frac{100}{150}\right) = -0.405$

$$\hat{\beta}_{11} = \log\left(\frac{70}{80}\right) - \log\left(\frac{100}{150}\right) = 0.272$$

$$\hat{\beta}_{21} = \log\left(\frac{30}{20}\right) - \log\left(\frac{100}{150}\right) = 0.811$$

(b)  $\hat{\beta}_{01}$ : Estimated log odds of ~~getting~~ moderate obesity for a non-drinker, given that the patient is not severely obese.

$e^{\hat{\beta}_{21}}$ : Estimated odds ratio of moderate obesity comparing heavy drinkers to non-drinkers, given that the patient is not severely obese.

(c)  $SE(\hat{\beta}_{01}) = \sqrt{\frac{1}{150} + \frac{1}{100}}$

$$95\% \text{ CI for } \beta_{01} : \hat{\beta}_{01} \pm 1.96 \cdot SE(\hat{\beta}_{01}) = (-0.66, -0.15)$$

$$SE(\hat{\beta}_{21}) = \sqrt{\frac{1}{30} + \frac{1}{20} + \frac{1}{100} + \frac{1}{150}}$$

$$95\% \text{ CI for } \beta_{21} : \hat{\beta}_{21} \pm 1.96 \cdot SE(\hat{\beta}_{21}) = (0.19, 1.43)$$

$$95\% \text{ CI for } e^{\hat{\beta}_{21}} : (e^{0.19}, e^{1.43}) = (1.21, 4.18)$$

3.

$$(a) \text{ Model : } \log \left\{ \frac{P(Y_i \leq j)}{P(Y_i > j)} \right\} = \gamma_0 j + \gamma_1 \text{Treat}_i + \gamma_2 \text{Sex}_i$$

$$\hat{\gamma}_{01} = -0.196 \quad \hat{\gamma}_{02} = 1.371 \quad \hat{\gamma}_{03} = 2.422$$

$$\hat{\gamma}_1 = -0.581 \quad \hat{\gamma}_2 = -0.541$$

$Y_i = 0$	progression
$Y_i = 1$	no change
$Y_i = 2$	partial remission
$Y_i = 3$	complete remission
$\text{Treat} = 0$	Alt+, $\text{Treat} = 1$ Seq
$\text{Sex} = 0$	Female, $\text{Sex} = 1$ male

Estimated probabilities :

$$P(Y_i = j \mid \text{Treat}, \text{Sex}) = P(Y_i \leq j \mid \text{Treat}, \text{Sex}) - P(Y_i \leq j-1 \mid \text{Treat}, \text{Sex}),$$

$$\text{where } P(Y_i \leq j \mid \text{Treat}, \text{Sex}) = \text{expit}(\gamma_{0j} + \gamma_1 \text{Treat}_i + \gamma_2 \text{Sex}_i)$$

Trt	Sex	Progression	No change	Partial remission	Complete remission
Alt	F	0.45	0.35	0.12	0.08
Alt	M	0.32	0.37	0.17	0.13
Seq	F	0.32	0.37	0.18	0.14
Seq	M	0.21	0.35	0.22	0.21

$$(b) \text{ Pearson Chi-square statistic} = 5.35 \sim \chi^2_7.$$

$$\chi^2_{7, 0.95} = 14.07 > 5.35 \Rightarrow \text{Fail to reject } H_0.$$

So the model fits well according to Pearson Chi-square statistic.

$$(c) H_0: \gamma_{\text{Treat}} = 0 \quad \text{v.s.} \quad H_1: \gamma_{\text{Treat}} \neq 0$$

$$\chi^2_w = 7.51 \sim \chi^2 \quad p\text{-value} = 0.0061$$

Reject  $H_0$ .

So there is significant difference in responses for the 2 treatments.