

1(a) Data is collected as a form of the following table:

subject	item			
	1	2	...	k
1				
2				
...				
n				

$Y_{ij} \in \{1, 0\}$ \sim iid Bernoulli(p_{ij}), $p_{ij} = \frac{e^{\alpha_j \theta_i - \beta_j}}{1 + e^{\alpha_j \theta_i - \beta_j}}$, $\eta_j = (\alpha_j, \beta_j)$, $\eta = (\eta_1, \dots, \eta_n)$
 $0 \leq \theta_i \sim \Phi(\theta)$ unobserved ability
 correct answer (1) / incorrect answer (0)
 subject i / item j

The probability space consists of the following three elements:

$\mathcal{X} = \{0, 1\}^{n \times k}$, a collection of all $n \times k$ matrices of 0 or 1.

$\mathcal{B}_\mathcal{X}$ = a collection of all subsets of matrices of dimensions $n \times k$ or less.

$$\mathcal{P} = \left\{ P_\eta: \prod_{i=1}^n \prod_{j=1}^k p_{ij}^{Y_{ij}} (1-p_{ij})^{1-Y_{ij}} \phi(\theta_i) d\theta_i, \eta = (\alpha_1, \dots, \alpha_k, \beta_1, \dots, \beta_k) \in \mathbb{R}_+^{2k} \right\}$$

where $p_{ij} = P(Y_{ij}=1 | \eta_j, \theta_i)$, $j=1, \dots, k$, $i=1, \dots, n$.

(b) The achm space is (A, \mathcal{B}_A)

$A = \mathbb{R}_+^{2k}$ or \mathbb{R}^{2k} , including impossible values

$\mathcal{B}_A = \{ \text{all possible subsets of } A \}$

(c) The observed likelihood is

$$L(\eta, x) = \prod_{i=1}^n \prod_{j=1}^k p_{ij}^{Y_{ij}}(\eta_j, \theta_i) \{1 - p_{ij}(\eta_j, \theta_i)\}^{1-Y_{ij}} \phi(\theta_i) d\theta_i, \quad x \in \mathcal{X}$$

and the log observed likelihood is

$$\ell(\eta, x) = \sum_{i=1}^n \left(\sum_{j=1}^k Y_{ij} \log p_{ij}(\eta_j, \theta_i) + (1-Y_{ij}) \log \{1 - p_{ij}(\eta_j, \theta_i)\} \right) \phi(\theta_i) d\theta_i, \quad x \in \mathcal{X}$$

The decision function $\delta: (\mathcal{X}, \mathcal{B}_\mathcal{X}) \rightarrow (A, \mathcal{B}_A)$ is given by the MLE:

$$\delta(x) = \underset{\eta}{\operatorname{argmax}} \ell(\eta, x), \quad x \in \mathcal{X}, \quad \text{if } A = \mathbb{R}^{2k}$$

$$\text{or } = \max \{0, \underset{\eta}{\operatorname{argmax}} \ell(\eta, x)\}, \quad \text{if } A = \mathbb{R}_+^{2k}$$

(d) The squared loss function is

$$L(\eta, a) = \sum_{j=1}^k \{ (a_{j1} - \alpha_j)^2 + (a_{j2} - \beta_j)^2 \} = \sum_{j=1}^k (a_j - \eta_j)^T (a_j - \eta_j)$$

Risk function is $E_{\theta} L(\eta, \delta(x)) = \int_{\mathcal{X}} \sum_{j=1}^k (E_{\theta}(\delta_j(x) - \eta_j))^2 dP_{\eta}(dx) = \sum_{j=1}^k \text{bias}_{\eta_j}^2 + \text{var}_{\eta_j}$

(10) **Problem 3** Consider a probability space $(\mathcal{X}, \mathcal{B}_x, \mathcal{P} = \{P_\theta, \theta \in \Theta\})$. Show the following statements.

- 5 (a) If $\delta_0(x)$ is the unique minimax decision rule with respect to the loss function $L(\theta, a)$, then δ_0 is admissible.
- 5 (b) If $\delta_0(x)$ is an admissible decision rule and has a constant risk function on Θ , then δ_0 is the minimax decision rule.

(a) If δ_0 is not admissible, then there exists a $\tilde{\delta} \in \mathcal{D}$ such that

$$R(\theta, \tilde{\delta}) \leq R(\theta, \delta_0), \text{ for all } \theta \in \mathcal{H},$$

and there is at least one $\theta \in \mathcal{H}$ at which $R(\theta, \tilde{\delta}) < R(\theta, \delta_0)$.

Then,
$$\sup_{\theta \in \mathcal{H}} R(\theta, \tilde{\delta}) \leq \sup_{\theta \in \mathcal{H}} R(\theta, \delta_0) \leq \sup_{\theta \in \mathcal{H}} R(\theta, \delta), \delta \in \mathcal{D}$$

↓ minimax rule

It implies that $\tilde{\delta}$ is a minimax decision rule because δ_0 is a minimax rule. Due to the uniqueness, $\tilde{\delta} \equiv \delta_0$, which is contradictory to the fact that their risk functions are not the exactly same.

Thus, δ_0 must be admissible.

(b) If δ_0 is not the minimax decision rule, there exists a $\tilde{\delta} \in \mathcal{D}$, such that

$$\sup_{\theta} R(\theta, \tilde{\delta}) < \sup_{\theta} R(\theta, \delta_0) = c.$$

Then,
$$R(\theta, \tilde{\delta}) < R(\theta, \delta_0) = c, \text{ all } \theta \in \mathcal{H}.$$

That is, δ_0 is not admissible. Contradiction!

Therefore, δ_0 is a minimax decision rule.

2(a) The loss function is

$$L(\theta, \delta_c(x, y)) = \theta 1(\theta > 0) 1(\bar{x} \leq \bar{y} + c) - \theta 1(\theta \leq 0) 1(\bar{x} > \bar{y} + c)$$

2(b) $\bar{X} \sim N(\mu_1, \frac{1}{m})$ and $\bar{Y} \sim N(\mu_2, \frac{1}{n})$, and they are independent. Then,
 $\bar{X} - \bar{Y} \sim N(\mu_1 - \mu_2, \frac{1}{m} + \frac{1}{n}) = N(\theta, \frac{1}{m} + \frac{1}{n})$.

Thus the risk function is

$$\begin{aligned} R(\theta, \delta_c) &= \theta 1(\theta > 0) P(\bar{X} - \bar{Y} \leq c) - \theta 1(\theta < 0) P(\bar{X} - \bar{Y} > c) \\ &= \theta 1(\theta > 0) \Phi\left(\frac{c - \theta}{\sqrt{\frac{1}{m} + \frac{1}{n}}}\right) - \theta 1(\theta < 0) \left\{1 - \Phi\left(\frac{c - \theta}{\sqrt{\frac{1}{m} + \frac{1}{n}}}\right)\right\} \\ &= \theta \{1(\theta > 0) + 1(\theta < 0)\} \Phi\left(\frac{c - \theta}{\sqrt{\frac{1}{m} + \frac{1}{n}}}\right) - \theta 1(\theta < 0) \\ &= \theta \Phi\left(\frac{c - \theta}{\sqrt{\frac{1}{m} + \frac{1}{n}}}\right) - \theta 1(\theta < 0) \end{aligned}$$

2(c) For any $c \in (-\infty, \infty)$, because $\Phi\left(\frac{c - \theta}{\sqrt{\frac{1}{m} + \frac{1}{n}}}\right) \uparrow c \uparrow$ there always exist $c^* < c$
 $c^{**} > c$
 (e.g. $c^* = c - 1$ and $c^{**} = c + 1$).

$$\begin{aligned} R(\theta, \delta_c) &= \theta \Phi\left(\frac{c - \theta}{\sqrt{\frac{1}{m} + \frac{1}{n}}}\right) - \theta 1(\theta < 0) \\ &> \underbrace{\theta \Phi\left(\frac{c^* - \theta}{\sqrt{\frac{1}{m} + \frac{1}{n}}}\right) 1(\theta > 0) + \theta \Phi\left(\frac{c^{**} - \theta}{\sqrt{\frac{1}{m} + \frac{1}{n}}}\right) 1(\theta < 0) - \theta 1(\theta < 0)} \\ &= R(\theta, \delta^*) \end{aligned}$$

where $\delta^* = \begin{cases} a_0, & \text{if } \bar{X} < \bar{Y} + c^*, \\ a_1, & \text{if } \bar{X} > \bar{Y} + c^{**}, \\ \emptyset, & \text{o/w.} \end{cases}$

So, δ_c is not admissible.

Note that no decision made incurs no loss or risk.