

Lecture 25. MLE norm, Emp Proc, Final practice

Monday, December 11, 2017 9:20 AM

Final Date and Time: Thursday 1:30-3:30, M4318

Exam is cumulative as far as knowledge is concerned. However, problems will come from topics discussed in lectures 12-25. These include

Convergence

Laws of large numbers

Limit theorems

M/Z Estimation

MLE

Exam is closed book

Two 2-sided help-sheets are allowed

Sample practice exam:

① State and prove the Borel-Cantelli TH

② State and prove the Cramer-Wold device

③ Let $\{X_i\}$ be i.i.d. with zero mean and unit variance

Find the asymptotic distribution of

$$Y_n = \frac{X_1 + \dots + X_n}{\sqrt{X_1^2 + \dots + X_n^2}}$$

④ Let X follow a Power-Series distribution with pmf $f(x; \theta) = a(x) \theta^x / f(\theta)$.

where

$$f(\theta) = \sum_{x=1}^{\infty} a(x) \theta^x \quad \text{converges for } 0 < \theta \leq R$$

for some $R, a > 0$.

(a) Prove that MLE of θ , $\hat{\theta}_n$ satisfies the equation

$$\mathbb{E} X = \frac{1}{n} \sum_{i=1}^n X_i$$

(b) Provide a weak convergence statement

(c) Use the result to show asymptotic normality and consistency of \hat{p}_n of a negatively Binomial distribution $\bar{B}_i(r, p)$

Note: \bar{B}_i has expectation $\frac{rp}{1-p}$

$$\text{variance } \frac{rp}{(1-p)^2}$$

$$P(X=k) = C_{r+k-1}^k p^k (1-p)^r$$

⑤ $\{X_i\}$ is a sequence of r.v.s

\exists There exist r.v. X and a sequence of integers $n_i > 0$:

$$\max_{n_{k-1} < m \leq n_k} |X_m - X_{n_{k-1}}| \xrightarrow[k \rightarrow \infty]{} 0 \quad \text{a.s.}$$

$$\text{and } X_{n_k} \rightarrow X \quad \text{a.s.}$$

$$\text{Prove: } X_n \xrightarrow[\text{a.s.}]{} X$$

MLEs

$$l_n = \sum_{i=1}^n \log \underbrace{f_{\theta}(X_i)}_{\text{pdf}}$$

$$\max_{\theta} l_n \Rightarrow \hat{\theta}_n \quad \text{MLE}$$

$$M_n(\theta) = \mathbb{P}_n \log \frac{f_{\theta}}{f_{\theta^*}}$$

Model identifiability $\Rightarrow \theta^*$
true parameter value
unique solution to $\max_{\theta} M_{\theta}$

$$M_n \xrightarrow{r} M_{\theta} \text{ uniformly ; } M_{\theta} = \mathbb{P} m_{\theta}$$

Asymptotic variance of MLEs

From M-estimation theory

Asymptotic variance is

$$\frac{\mathbb{P}(m'_{\theta^*})^2}{(\mathbb{P}(m''_{\theta^*}))^2}$$

$$\psi_{\theta} = m'_{\theta} \quad (\text{MLE as a Z-estimation})$$

Normalization condition

$$\mathbb{P}(1) = 1 = \int f_{\theta} dx$$

$$\int f'_\theta dx = \int f''_\theta dx = 0$$

$$m_\theta = \log f_\theta$$

$$m'_\theta = \frac{f'_\theta}{f_\theta}$$

$$m''_\theta = \frac{f''_\theta}{f_\theta} - \left(\frac{f'_\theta}{f_\theta} \right)^2$$

$$\mathbb{P} m'_{\theta^*} = \int \frac{f'_{\theta^*}}{f_{\theta^*}} \cdot f_{\theta^*} dx = 0$$

Expected value of the score is zero at the true param. value

$$\mathbb{P}(m''_{\theta^*}) = \underbrace{\int f''_{\theta^*} dx}_{=0 \text{ by deriv. of normalization condition}} - \mathbb{P}(m'_{\theta^*})^2$$

$$\underbrace{\mathbb{P}(m'_{\theta^*})^2}_{\text{variance of the true score}} = \underbrace{-\mathbb{P}(m''_{\theta^*})}_{\text{negative Hessian}} = I_{\theta^*}$$

\Rightarrow

Asymptotic variance $\frac{I_{\theta^*}}{[I_{\theta^*}]^2} = I_{\theta^*}^{-1}$

$$\sqrt{n} (\hat{\theta}_n - \theta^*) \rightsquigarrow N(0, I_{\theta^*}^{-1})$$

Under reg. conditions
Observed information matrix is a consistent estimate of I_{θ^*}

End of material that goes into the final exam

Overview of empirical process theory

Empirical Distribution function

$$F_n(t) = \frac{1}{n} \sum_{i=1}^n \mathbb{1}(X_i \leq t)$$

$$n F_n \sim \text{Bin}(n F(t)) \Rightarrow \text{unbiased}$$

↳ CDF

$$\text{LLN} \Rightarrow \text{consistency} \quad F_n(t) \xrightarrow{p} F(t), \forall t$$

$$\text{CLT} \Rightarrow \sqrt{n}(F_n - F) \rightsquigarrow N(0, F(t)(1-F(t)))$$

Random functions of t are called Stochastic Processes

Glivenko - Cantelli Theorem

Uniform Law of Large Numbers for F_n

$\|\cdot\|_\infty$ Norm

$$\|F_n - F\|_\infty := \sup_t |F_n(t) - F(t)|$$

Kolmogorov - Smirnov
statistic

(TH) Glivenko - Cantelli:

$\exists X_i$ be i.i.d. r.v.s

\Rightarrow

$$\|F_n - F\|_\infty \xrightarrow{\text{a.s.}} 0$$

↳ uniform convergence statement
is a functional norm

Proof:

Take $\varepsilon > 0$

$\exists t_i$ are points where F jumps by
more than ε

Then, for $t_{i-1} \leq t < t_i$, we have

$$F_n(t_{i-1}) - F(t_{i-1}) - \varepsilon \leq F_n(t) - F(t) \leq F_n(t_i-) - F(t_i-) + \varepsilon$$

F can have at most a countable set of jumps < ε

b/c F is monotonic and bounded

b/c jumps @ t_i are more than ε , and F is bounded \Rightarrow at most finite # of t_i s

\Rightarrow any pointwise result over a finite number of points is a uniform result over those points

\Downarrow

$$\text{b/c } F_n(t_{i-1}) - F(t_{i-1}) \xrightarrow{\text{a.s.}} 0$$

$$F_n(t_i-) - F(t_i-) \xrightarrow{\text{a.s.}} 0$$

$$\Rightarrow \sup_t |F_n - F| < \varepsilon$$

$$\Rightarrow P(\|F_n - F\|_\infty \leq \varepsilon) \rightarrow 1 \quad \square$$

$$\textcircled{\text{DF}} \quad \sqrt{n} (P_n - P)f = \sqrt{n} \left(\frac{1}{n} \sum_{i=1}^n f(x_i) - \int f(x) dF(x) \right)$$

is called an empirical process evaluated at f

$\textcircled{\text{TH}}$ Donsker

$$G_n := \sqrt{n} (P_n - F) \rightsquigarrow G_F$$

G_F is a Brownian Bridge with covariance

$$\text{Cov}(G_F(x), G_F(y)) = F(\min(x, y)) - F(x)F(y)$$

$\textcircled{\text{DF}}$ Brownian Motion and Bridge

... \uparrow

$\sqrt{n}(F_n - F)f$

$W(t)$ is \uparrow or Wiener Process
when

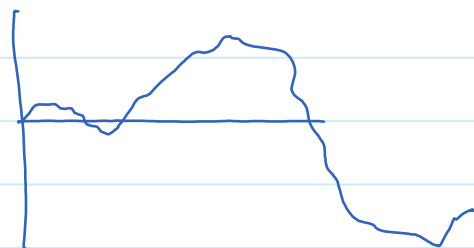
$$W(0) = 0$$

independent increments $dW(t)$

$\Delta_s^t W$ depends only on $t-s$

\downarrow increment on $[s, t]$

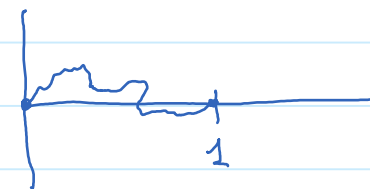
$$W(t) \sim N(0, t)$$



Bridge is $w(t)$ given that $w(1) = 0$

\downarrow

$$w(t) - w(1)$$



(DF)

Class of functions f is called
 \uparrow
measur.

Glivenko-
Cantelli (GC)

\downarrow
 \mathcal{G}

$$\text{if } \sup_{f \in \mathcal{G}} |(P_n - P)f| \xrightarrow{\text{a.s.}} 0$$

(DF)

A class \mathcal{D}_P is called Donsker if
(P -Donsker)

$$\mathcal{G}_n f \rightsquigarrow \mathcal{G}_P \text{ where}$$

\mathcal{G}_P is zero-mean Gaussian
w. cov.

$$\mathbb{E}(G_p f G_p g) = \mathbb{P}(fg) - \mathbb{P}f \cdot \mathbb{P}g$$

Donsker and Glivenko - Cantelli will be fulfilled when entropy of \mathcal{C} is not growing too fast as $\varepsilon \rightarrow 0$ (see below)

(DF) Bracketing number $N_{[]} \varepsilon$ is the min number of balls in L_p norm, centered around functions $f \in \mathcal{C}$, with a radius of $\varepsilon > 0$ are needed to cover all of \mathcal{C} .

(DF) Entropy = $\log N_{[]} \varepsilon$

(DF) Bracketing integral $J_{[]} = \int_0^\delta \sqrt{\log N_{[]} \varepsilon} d\varepsilon$

(TH) $\int N_{[]} \varepsilon < \infty, \forall \varepsilon > 0 \Rightarrow \mathcal{C}$ is GC

(TH) $\int J_{[]} < \infty \Rightarrow \mathcal{C}$ is \mathbb{D}_p

Note: $N_{[]} \varepsilon \rightarrow \infty$ slower than $\frac{1}{\varepsilon^2} \Rightarrow \mathbb{D}_p$