

Lecture 22 M, Z est

Wednesday, November 29, 2017

10:09 AM

Lyapunov CLT

Lyapunov condition \Rightarrow Lindeberg condition

M and Z estimators

M-estimation

$$M_n(\theta)$$

\downarrow

random functions

$$M_n(\theta) \xrightarrow{P} M(\theta)$$

Estimator is defined as

a maximizer or

near-maximizer (up to $o_p(1)$)

$$\max_{\theta} M_n(\theta) \Rightarrow \hat{\theta}_n$$

$$\text{Consistency: } \hat{\theta}_n \xrightarrow{P} \theta^*$$

LLN +

$$\text{Normality: } \sqrt{n}(\hat{\theta}_n - \theta^*) \rightsquigarrow N(0, \Sigma)$$

CLT

Z-estimation

$\Psi_n(\cdot)$ random function
based on the sample
and approximating Ψ

$$\Psi(\theta)$$

θ^* is a solution to

$$\Psi(\theta) = 0$$

$$\Psi_n(\cdot) \xrightarrow{P} \Psi(\cdot)$$

$$\Psi_n(\theta) = 0 \Rightarrow \hat{\theta}_n$$

$$M_n(\theta) = \frac{1}{n} \sum_i m_\theta(x_i)$$

deterministic functions

$$\Psi_n(\theta) = \frac{1}{n} \sum_i \Psi_\theta(x_i)$$

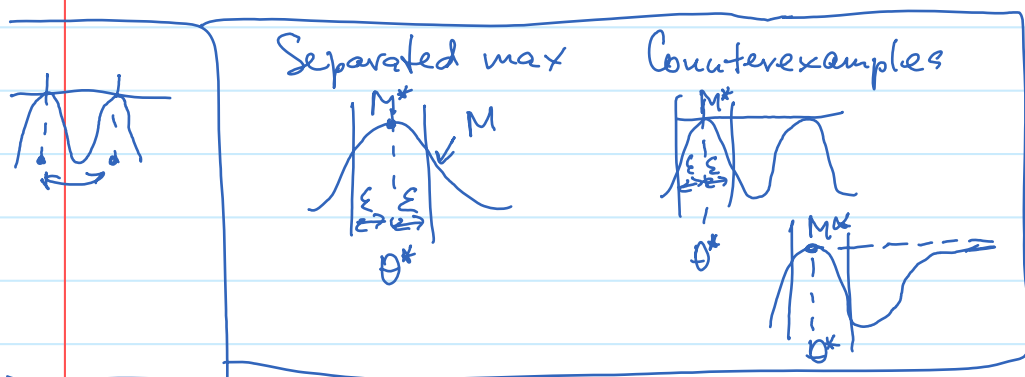
Notation $\mathbb{E}f(x) := Pf$

$$\frac{1}{n} \sum_i f(x_i) := \mathbb{P}_n f$$

↓ empirical distribution

TH

Uniform convergence of $M_n \rightarrow M$
Separated true maximum $\} \Rightarrow$ Consistency



] We have a sequence of estimators $\hat{\theta}_n$ such that
 $M_n(\hat{\theta}_n) \geq M_n(\theta^*) - o_p(1)$ "approx maximizer"

$$\sup_{\theta \in \Theta} |M_n(\theta) - M(\theta)| \xrightarrow{P} 0$$

"Uniform convergence"

$$\forall \varepsilon > 0, \sup_{\theta: \|\theta - \theta^*\| \geq \varepsilon} M(\theta) < M(\theta^*)$$

"separated true" max

Then

$$\hat{\theta}_n \xrightarrow{P} \theta^*$$

$$\hat{\theta}_n \xrightarrow{P} \theta^*$$

Proof: By the unif. conv. cond. $M_n(\theta^*) \xrightarrow{P} M(\theta^*)$

approx. maximizer cond $\Rightarrow M(\theta^*) \leq M_n(\hat{\theta}_n) + o_p(1)$

$$\begin{aligned} M(\theta^*) - M(\hat{\theta}_n) &\leq M_n(\hat{\theta}_n) - M(\hat{\theta}_n) + o_p(1) \leq \\ &\leq \sup_{\theta} |M_n(\theta) - M(\theta)| + o_p(1) \xrightarrow{P} 0 \text{ by unif. conv.} \end{aligned}$$

$$M(\hat{\theta}_n) \xrightarrow{P} M(\theta^*)$$

b/c θ^* is a separated max

$\forall \varepsilon > 0 \exists \eta > 0 :$

$$M(\theta) < M(\theta^*) - \eta, \forall \theta : \|\theta - \theta^*\| \geq \varepsilon$$

$$\|\hat{\theta}_n - \theta^*\| \geq \varepsilon \Rightarrow M(\hat{\theta}_n) < M(\theta^*) - \eta$$

$$P(\|\hat{\theta}_n - \theta^*\| \geq \varepsilon) \leq P(M(\theta^*) - M(\hat{\theta}_n) > \eta) \xrightarrow{n \rightarrow \infty} 0$$

$$\Rightarrow \hat{\theta}_n \xrightarrow{P} \theta^*$$

□

$$M_n(\hat{\theta}_n)$$

(TH)

Z estimator consistency

$$\text{Unif. conv. } \sup_{\theta} \|\Psi_n(\theta) - \Psi(\theta)\| \xrightarrow{P} 0$$

$$\text{Separ. of the root } \inf_{\theta : \|\theta - \theta^*\| \geq \varepsilon} \|\Psi(\theta)\| > 0 = \|\Psi(\theta^*)\|$$

$$\Rightarrow \forall \hat{\theta}_n : \Psi_n(\hat{\theta}_n) = o_p(1)$$

we will have

1

approx root cond

$$\theta_n \xrightarrow{p} \theta$$

Proof:

$$\text{Take } M_n = -\|\Psi_n\|$$

$$M = -\|\Psi\|$$

$$\sup |M_n - M| = \sup |\|\Psi\| - \|\Psi_n\|| \leq$$

$$\leq \sup \|\Psi - \Psi_n\| \xrightarrow{p} 0$$

unif conv. cond

$$\sup_{\theta: \|\theta - \theta^*\| > \varepsilon} M(\theta) = \sup (-\|\Psi\|) = -\inf (\|\Psi\|) < 0$$

$$\parallel \\ M(\theta^*)$$

$$\text{By the previous Th } \Rightarrow \hat{\theta}_n \xrightarrow{p} \theta^*$$

□

Note: if M has a unique sup and M is continuous } Separation cond is satisfied

Functions M_n, Ψ_n that satisfy ULLN are studied in Empirical processes

TH Consistency w/o uniform convergence

$$\theta \in \mathbb{R}$$

Ψ_n random functions Ψ fixed function:

We have pointwise convergence

$$\Psi_n(\theta) \xrightarrow{p} \Psi(\theta), \forall \theta$$

(a) Let Ψ_n be a continuous function with unique zero $\hat{\theta}_n$
or

(b) Ψ_n is a non-decreasing function with $\hat{\theta}_n$ being its asymptotic zero $\Psi(\hat{\theta}_n) = o_p(1)$

Let θ^* be such that $0 \in [\psi(\theta^* \pm \varepsilon)]$, $\forall \varepsilon > 0$
interval $[\psi(\dots - \varepsilon), \psi(\dots + \varepsilon)]$

Proof: