

## Lecture 10. total prob

Monday, October 9, 2017 10:04 AM

### Properties of $\mathbb{E}\{\cdot | \mathcal{U}\}$

Linearity

$$\mathbb{E}(aX_1 + bX_2 | \mathcal{U}) = a\mathbb{E}(X_1 | \mathcal{U}) + b\mathbb{E}(X_2 | \mathcal{U})$$

Ordering preservation

$$X_1 \leq X_2 \text{ a.s.} \Rightarrow \mathbb{E}(X_1 | \mathcal{U}) \leq \mathbb{E}(X_2 | \mathcal{U}) \text{ a.s.}$$

Conditional probability  $P(A|\mathcal{U}) = \mathbb{E}\{I_A | \mathcal{U}\}$

Probability inequalities

$$X \geq 0, x > 0 \quad P(X \geq x | \mathcal{U}) \leq \frac{\mathbb{E}(X | \mathcal{U})}{x}$$

$$f \uparrow, f > 0 \quad P(X \geq x | \mathcal{U}) \leq \frac{\mathbb{E}(f(X) | \mathcal{U})}{f(x)}$$

$$P(|X - \mathbb{E}(X | \mathcal{U})| \geq x | \mathcal{U}) \leq \frac{\text{Var}(X | \mathcal{U})}{x^2}$$

Monotone convergence Th for  $\mathbb{E}(\cdot | \mathcal{U})$

(TH)

$$0 \leq X_n \uparrow X \text{ a.s.} \Rightarrow \mathbb{E}\{X_n | \mathcal{U}\} \uparrow \mathbb{E}\{X | \mathcal{U}\}$$

Proof.

$$X_{n+1} \geq X_n \text{ a.s.} \Rightarrow \mathbb{E}(X_{n+1} | \mathcal{U}) \geq \mathbb{E}(X_n | \mathcal{U})$$

↓

ordering preservation of  $\mathbb{E}(\cdot | \mathcal{U})$

$$\text{Also } X_n \leq X \text{ a.s.} \Rightarrow \mathbb{E}(X_n | \mathcal{U}) \leq \mathbb{E}(X | \mathcal{U})$$

$\Rightarrow \exists Z : Z \text{ is } \mathcal{U}\text{-measurable} :$

$$Y_n = \mathbb{E}(X_n | \mathcal{U}) \rightarrow Z \text{ a.s.}$$

by the bounded convergence Th from

Note: limits of measurable functions are measurable if they exist

Calculus

Using the following property of Lebesgue integrals from for example Chung

Prop. 1.1.1

Note: limits of measurable functions are measurable if they exist

by the bounded convergence theorem

Using the following property of Lebesgue integrals from for example Shiryaev;  
Proved using usual approximations by simple functions,

Calculus  
for each  $\omega$   
except perhaps  
a set of  
measure zero

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II. Mathematical Foundations of Probability Theory

**Theorem 1** (On Monotone Convergence). Let  $\eta, \xi, \xi_1, \xi_2, \dots$  be random variables.

(a) If  $\xi_n \geq \eta$  for all  $n \geq 1$ ,  $E\eta > -\infty$ , and  $\xi_n \uparrow \xi$ , then

$$E\xi_n \uparrow E\xi.$$

(b) If  $\xi_n \leq \eta$  for all  $n \geq 1$ ,  $E\eta < \infty$ , and  $\xi_n \downarrow \xi$ , then

$$E\xi_n \downarrow E\xi.$$

We have:

$$\begin{aligned} E(Y_n; A) &\rightarrow E(Z; A) \\ E(X_n; A) &\rightarrow E(X; A) \end{aligned} \quad \text{by properties of integrals}$$

$$\begin{aligned} \uparrow \\ \text{Now by definition of } E(\cdot|U) \Rightarrow \\ \Rightarrow E(Y_n; A) = E(X_n; A) \end{aligned}$$

$$\begin{aligned} \Downarrow \\ E(Z; A) = E(X; A) \quad \text{by unique limit} \\ \text{by definition of } E(\cdot|U) \Rightarrow Z = E\{X|U\} \end{aligned}$$

□

(TH)  $U$ -measurable r.v.s behave like Const inside  $E(\cdot|U)$

□

$X \sim \mathcal{F}$ -measurable  $U \subset \mathcal{F}$

$Y \sim U$ -measurable

$$E(|X|) < \infty$$

$$E(|X \cdot Y|) < \infty$$

Then

$$E(X \cdot Y | U) = Y \cdot E(X | U)$$

Proof.

$$\exists Y = I_B, \quad B \in U$$

Take  $\forall A \in U$

$$\begin{aligned} \int_A X \cdot Y \, dP &= \int_A X \cdot I_B \, dP \\ &= \int_{A \cap B} X \, dP \\ &= \int_A X \cdot E(X|U) \cdot I_B \, dP \\ &= \int_A E(X|U) \cdot I_B \, dP \\ &= \int_A E(X|U) \cdot Y \, dP \end{aligned}$$

$$\begin{aligned}
 \mathbb{E}(X \cdot I_B; A) &= \mathbb{E}(X \cdot I_B \cdot I_A) = \\
 &= \mathbb{E}(X; AB) := \mathbb{E}\{ \mathbb{E}(X|u); AB \} = \\
 &= \mathbb{E}\{ I_B \cdot \mathbb{E}(X|u); A \} \Rightarrow \\
 &\quad \underbrace{I_B \mathbb{E}(X|u)}_{u\text{-meas. r.v.}} = \mathbb{E}(X \cdot I_B | u) \quad \text{by def of } \mathbb{E}(\cdot|u)
 \end{aligned}$$

Suppose all variables are  $\geq 0$

Construct  $Y_n$  - simple and :  $0 \leq Y_n \uparrow Y, X \geq 0$

The fact we are trying to prove is valid for simple functions & c

$$\mathbb{E}(X \cdot Y_n | u) = Y_n \mathbb{E}(X | u)$$

Going to  $\lim_{n \rightarrow \infty}$  in  $\uparrow$  and monotone conv. Th for  $\mathbb{E}(\cdot|u)$

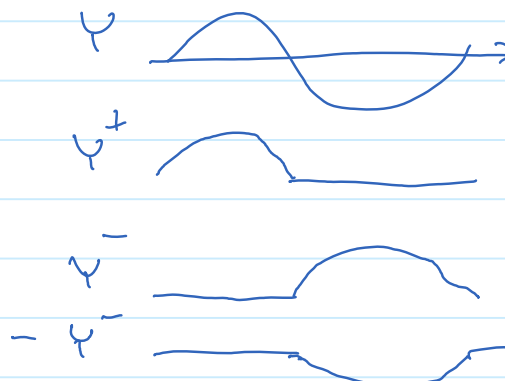
we will have the result for  $\geq 0$  r.v.s.

Any  $Y = \underbrace{Y^+}_{\geq 0} - \underbrace{Y^-}_{\geq 0} \Rightarrow$  get the validity of the Th for arbitrary  $Y$

$$X = \underbrace{X^+}_{\geq 0} - \underbrace{X^-}_{\geq 0}$$

□

$$Y^\pm = \max(0, \pm Y)$$



Formula of total probability for  $\mathbb{E}$

(TH)  $\mathbb{E}(X) = \mathbb{E}\{ \underbrace{\mathbb{E}\{X|u\}}_Y \}$

$$\pi(V, A) = \pi(V \cdot A) \quad \dots \quad \pi(A)$$

$$\mathbb{E}(Y; A) \stackrel{Y}{=} \mathbb{E}(X; A), \quad \forall A \in \mathcal{U}$$

Take  $A = \Omega \Rightarrow \mathbb{E}(Y) = \mathbb{E}(X)$   $\square$

(TH)  $\exists \mathbb{E} X^2 < \infty$

$$\arg \min_{Y: \mathcal{U}\text{-measurable}} \underbrace{\mathbb{E} \{ (X - Y)^2 \}}_{\text{distance}^2 \text{ between } X \text{ \& } Y} = \mathbb{E}(X | \mathcal{U})$$

$\mathcal{U} \subset \mathcal{F} = \sigma(X)$

Proof.

$$\begin{aligned} \mathbb{E}((X - Y)^2) &= \mathbb{E}(\underbrace{\mathbb{E}[(X - Y)^2 | \mathcal{U}]}_{\substack{\text{formula of} \\ \text{total prob. for } \mathbb{E}}} = \\ &\quad \underbrace{X^2 - 2XY + Y^2}_{\substack{\text{\mathcal{U}\text{-measurable}}}} \\ &= \mathbb{E} \{ \mathbb{E}(X^2 | \mathcal{U}) - 2Y \mathbb{E}(X | \mathcal{U}) + Y^2 \} \text{ achieves} \\ &\quad \text{min of zero when } Y = \underbrace{\mathbb{E}(X | \mathcal{U})}_{\substack{\text{\mathcal{U}\text{-meas. by def.}}} \end{aligned}$$

$\square$

Formula of total probability for Var

$$\text{Var}(X) = \mathbb{E}(\text{Var}(X | \mathcal{U})) + \text{Var}(\mathbb{E}(X | \mathcal{U}))$$

Model  $(a, b) \Rightarrow \hat{a}, \hat{b}$  MLE

$$\text{Var } \hat{a} = \underbrace{\text{Var } \hat{a}}_{\substack{\uparrow \\ \text{Model}(a, b) | b \text{ known}}} + \text{Var due to } b \text{ being estimated}$$

Proof

$$\text{Var}(X) = \underbrace{\mathbb{E}(X^2)}_{\substack{\text{...} \\ \text{...}}} - \underbrace{\mathbb{E}^2(X)}_{\substack{\text{...} \\ \text{...}}} =$$

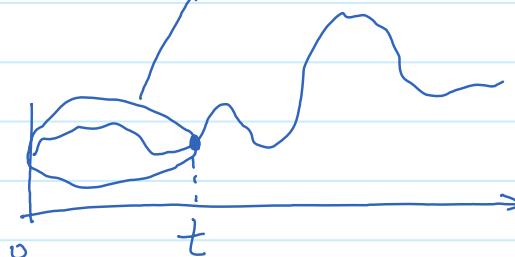
$$\text{Var}(X) = \mathbb{E}(X^2) - \mathbb{E}(X)^2 =$$

$$= \underbrace{\mathbb{E}(\mathbb{E}(X^2|u)) - \mathbb{E}(\mathbb{E}(X|u))^2}_{\text{"}} + \underbrace{\mathbb{E}(\mathbb{E}^2(X|u))}_{Y^2}$$

$$\mathbb{E}(\text{Var}(X|u)) + \text{Var}(\mathbb{E}(X|u)) \quad \square$$

Random process is a collection of r.v.s  $X_t$   
Stochastic

Trajectory of  $X$  on  $[0, t]$  is  $\bar{X}_t$



$$X(t, \omega), \quad \omega \in \mathcal{F}$$

Filtration is a sequence of increasing  $\sigma$ -algebras indexed by  $t$

$$\mathcal{F}_s \subset \mathcal{F}_t \subset \mathcal{F} \\ s < t$$

Increment

$$dX_t = X_{t+dt} - X_t$$

(DF) Take  $\mathcal{F}_t = \sigma(\bar{X}_t) \Rightarrow$  The process  $X_t$  is adapted to filtration

(DF) Martingale is a process :  $\mathbb{E}\{dX_t | \mathcal{F}_{t-}\} = 0$

(TH) Increments of a Martingale are uncorrelated

$$\mathbb{E}(dX_t \cdot dX_\tau) = \mathbb{E}\left(dX_\tau \cdot \overbrace{\mathbb{E}(dX_t | \mathcal{F}_{t-})}^{:=0}\right) = 0$$

$\tau < t$

$X_t = \int_0^t dX_x$  is a "sum" of uncorrelated increments