

Multi-view Warped Mixtures for Shape Correspondence Analysis

Hernán F. García and Mauricio A. Álvarez

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Abstract

Probabilistic approach for the correspondence problem based on the Single View WMM[6].

1 The model

Based on the Iwatta's paper (see [2, 3]), we set nonlinear approaches for the shape correspondence analysis using probabilistic latent variable models.

Suppose that we are given objects in D domains $\mathcal{X} = \{\mathbf{X}_d\}_{d=1}^D$ mapped to a Hilbert space \mathcal{H} , where $\mathbf{X}_d = \{\mathbf{x}_{dn}\}_{n=1}^{N_d}$ is a set of objects in the d th domain, and $\mathbf{x}_{dn} \in \mathbb{R}^{M_d}$ is the observed vector of the n th object in the d th domain. We can cluster groups of correspondences by using a non-linear function that represents the shape descriptors in the Hilbert space.

As in infinite Gaussian mixture models, our approach assumes that there are an infinite number of clusters related to each correspondence, and each cluster j has a latent vector $\mathbf{z}_j \in \mathbb{R}^K$ in a latent space of dimension K . Descriptors that have the same cluster assignments s_{dn} are related by the same latent vector and considered to match (establish a groupwise correspondence).

Each object in $\mathbf{x}_{dn} \in \mathcal{R}^{M_d}$ in the d th domain is generated depending on the domain-specific projection matrix $\mathbf{W}_d \in \mathbb{R}^{M_d \times K}$ and latent vector $\mathbf{z}_{s_{dn}}$ that is selected from a set of latent vectors $\mathbf{Z} = \{\mathbf{z}_j\}_{j=1}^\infty$. Here, $s_{dn} = \{1, \dots, \infty\}$ is the latent cluster assignment of object \mathbf{x}_{dn} .

The proposed model is based on an infinite mixture model, where the probability of descriptor mapped in a Hilbert space \mathbf{x}_{dn} is given by

$$p(\mathbf{x}_{dn} | \mathbf{Z}, \mathbf{W}, \boldsymbol{\theta}) = \sum_{j=1}^{\infty} \theta_j \mathcal{N}(\mathbf{x}_{dn} | \mathbf{W}_d \mathbf{z}_j, \alpha^{-1} \mathbf{I}), \quad (1)$$

where $\mathbf{W} = \{\mathbf{W}_d\}_{d=1}^D$ is a set of projections matrices, $\boldsymbol{\theta} = (\theta_j)_{j=1}^\infty$ are the mixture weights, θ_j represents the probability that the j th cluster is chosen, and $\mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ denotes a normal distribution with mean $\boldsymbol{\mu}$ and covariance matrix $\boldsymbol{\Sigma}$. One important contribution derived in [2], is that we can analyze multiples structures with different properties and dimensionalities, by employing projection matrices for each brain structure (domain-specific). Figure 1 shows the scheme of the proposed model, in which the relationship between shape descriptors, and latent vectors is described.

2 Model Approaches

The main ideas for the probabilistic correspondence problem are summarized

1. We can relax the assumption that the observations are linear with respect to their latent vectors by using nonlinear matrix factorization techniques ([5]). From these work we point out that by marginalizing out the mapping matrix \mathbf{W} (which goes from the latent space to the observed data space), derived in a Bayesian multi-output regression model.

Our model can be formalized as

$$p(\mathbf{x}_{dn} | \mathbf{Z}, \mathbf{W}, \boldsymbol{\theta}) = \sum_{j=1}^{\infty} \theta_j \prod_{m=1}^{M_d} \mathcal{N}(\mathbf{x}_{dnm} | f_{dm}(\mathbf{z}_j), \alpha^{-1}) \quad (2)$$

where f_{dm} can be handled from two perspectives

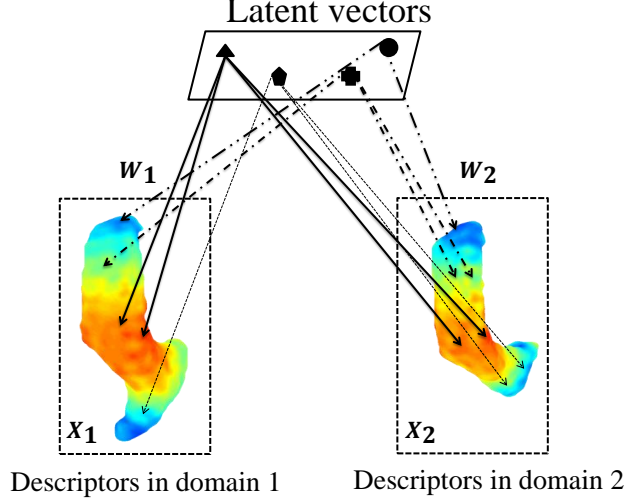


Figure 1: Scheme for the groupwise correspondence method. The figure shows an example of establishing clusters of correspondences in two domains (left ventrals).

- (a) Set the same covariance matrix for all domains

$$f_{dm} \sim \mathcal{GP}(\mu_{dm}(\mathbf{z}_j), k(\mathbf{z}_j, \mathbf{z}'_j)) \quad (3)$$

- (b) One covariance matrix for each domain

$$f_{dm} \sim \mathcal{GP}(\mu_{dm}(\mathbf{z}_j), k_d(\mathbf{z}_j, \mathbf{z}'_j)) \quad (4)$$

2. As in the in the linear model of coregionalization (LMC), the outputs are expressed as linear combinations of independent random functions [1]. Consider a set of D outputs $\{\mathbf{f}_d(\mathbf{z}_j)\}_{d=1}^D$ with $\mathbf{f}_d(\mathbf{z}_j) \in \mathbb{R}^{M_d}$. By adopting this framework, our model can be formulated as

$$p(\mathbf{x}_{dn} | \mathbf{Z}, \mathbf{W}, \boldsymbol{\theta}) = \sum_{j=1}^{\infty} \theta_j \mathcal{N}(\mathbf{x}_{dn} | \mathbf{f}_d(\mathbf{z}_j), \alpha^{-1} \mathbf{I}), \quad (5)$$

where each component can be expressed as:

- (a) By using the linear model of coregionalization

$$\mathbf{f}_d(\mathbf{z}_j) \sim \mathcal{GP}\left(\mathbf{0}, \sum_{q=1}^Q \mathbf{B}_q k_d(\mathbf{z}_j, \mathbf{z}'_j)\right), \quad (6)$$

where $\mathbf{B}_q = \mathbf{L}_q^\top \mathbf{L}_q \in \mathbb{R}^{M_d \times M_d}$ is the coregionalization matrix (computed from the Cholesky decomposition)

- (b) By using simplified version of the LMC, known as the intrinsic coregionalization model (ICM) (see [1]), assumes that the elements of the coregionalization matrix \mathbf{B}_q can be written as a scaled version of the elements b_q , which do not depend on the particular output functions $f_d(\mathbf{z}_j)$.

$$\mathbf{f}_d(\mathbf{z}_j) \sim \mathcal{GP}(\mathbf{0}, \mathbf{B} k(\mathbf{z}_j, \mathbf{z}'_j)), \quad (7)$$

3 Multiview Warped Mixture Models

We can use the single view problem of the warped mixture model from [6] in which they warp a latent mixture of Gaussians into nonparametric cluster shapes. The low-dimensional latent mixture model summarizes the properties of the high-dimensional density manifolds describing the data.

Our idea is to introduce a model which warps a multiview latent mixture of Gaussians (possibly MRD) to produce nonparametric cluster shapes.¹

WARPED MIXTURE MODEL

- ▶ An extension of GP-LVM, where $p(x)$ is a mixture of Gaussians.
- ▶ Or: An extension of iGMM, where mixture is warped.
- ▶ Given mixture assignments, likelihood has only two parts: GP-LVM and GMM

$$\begin{aligned}
 p(\mathbf{Y}|\mathbf{X}, \mathbf{Z}, \boldsymbol{\theta}) = & \underbrace{(2\pi)^{-\frac{DN}{2}} |\mathbf{K}|^{-\frac{D}{2}} \exp\left(-\frac{1}{2} \text{tr}(\mathbf{K}^{-1} \mathbf{Y} \mathbf{Y}^\top)\right)}_{\text{GP-LVM Likelihood}} \\
 & \times \underbrace{\prod_i \sum_{c=1}^{\infty} \lambda_c \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_c, \mathbf{R}_c^{-1}) I(\mathbf{x}_i \in \mathbf{Z}_c)}_{\text{Mixture of Gaussians Likelihood}}
 \end{aligned}$$

Figure 2: WMM from [6]

3.1 The model

Let us define a multi-view data set as $\mathcal{Y} = \{\mathbf{Y}^v\}_{v=1}^V$, where each view is defined as $\mathbf{Y}^v \in \mathbb{R}^{N_v \times D_v}$. This leads to the likelihood. Given mixture assignments, likelihood has only two parts: GP-LVM and GMM

$$p(\mathbf{Y}^\mathcal{V} | \mathbf{X}, \mathbf{Z}, \boldsymbol{\theta}) = \prod_{v=1}^V p(\mathbf{Y}^v | \mathbf{X}, \boldsymbol{\theta}) \times \prod_i \sum_{c=1}^{\infty} \lambda_c \mathcal{N}(\mathbf{x}_i | \boldsymbol{\mu}_c, \mathbf{R}_c^{-1}), \quad \mathbf{x}_i \in \mathbf{Z}_c \quad (8)$$

Based on the iWMM (see [6]) our generative model generates multiple observations $\mathbf{Y}^\mathcal{V}$ according to the following generative process:

1. Draw mixture weights $\boldsymbol{\lambda} \sim \text{GEM}(\eta)$
2. For each cluster $c = 1, \dots, \infty$
 - (a) Draw precision $\mathbf{R}_c \sim \mathcal{W}(\mathbf{S}^{-1}, v)$
 - (b) Draw mean $\boldsymbol{\mu}_c \sim \mathcal{N}(\mathbf{u}, (r\mathbf{R}_c)^{-1})$

¹The possibly low-dimensional latent mixture model allows us to summarize the properties of the high-dimensional clusters (or density manifolds) describing the data. The number of manifolds, as well as the shape and dimension of each manifold is automatically inferred.

3. For each view $v = 1, \dots, \mathcal{V}$
 - (a) For each observation $n = 1, \dots, N_v$
 - i. Draw latent assignment $z_{nv} \sim \text{Mult}(\boldsymbol{\lambda})$
 - ii. Draw latent coordinates $\mathbf{x}_{nv} \sim \mathcal{N}(\boldsymbol{\mu}_{z_{nv}}, \mathbf{R}_{z_{nv}}^{-1})$
4. For each view $v = 1, \dots, \mathcal{V}$
 - (a) For each observed dimension $d = 1, \dots, D_v$
 - i. Draw function $\mathbf{f}_d^v \sim \mathcal{GP}(\mathbf{0}, \mathbf{K}^v)$
5. For each view $v = 1, \dots, \mathcal{V}$
 - (a) For each observed dimension $d = 1, \dots, D_v$
 - i. Draw projection variable $w_d^v \sim \mathcal{N}(\rho_d^v, \omega_d^v)$
 - ii. For each observation $n = 1, \dots, N_v$
 - A. Draw feature $y_{nd}^v \sim \mathcal{N}(w_d^v \mathbf{f}_d^v(\mathbf{x}_{nv}), \beta^{-1})$

We define the dimensionalities of our variables as:

- K : real number of clusters
- Q : dimensionality of the Latent Space
- D_v : dimensionality of the input data in the v -th view
- $\boldsymbol{\lambda} \in \mathbb{R}^{K \times 1}$
- $\mathbf{R}_c \in \mathbb{R}^{Q \times Q}$
- $\boldsymbol{\mu}_c \in \mathbb{R}^{Q \times 1}$

3.2 Latent Multi-view Warped Mixture Model

Our model is set as a multi-view Gaussian Process Latent Variable model as [4]. First we assume that observations are generated by mapping the latent coordinates through a set of smooth functions, over which Gaussian process priors are placed. Under the GPLVM, the probability of observations given the latent coordinates, integrating out the mapping functions, is defined as

$$p(\mathbf{Y}|\mathbf{X}, \boldsymbol{\theta}) = \prod_{v=1}^V p(\mathbf{Y}^v|\mathbf{X}^v, \boldsymbol{\theta}^v) \quad (9)$$

$$= \prod_{v=1}^V \prod_{d=1}^{D_v} p(y_d^v|\mathbf{X}^v, \boldsymbol{\theta}^v), \quad (10)$$

where \mathbf{y}_d^v represents the d th column of \mathbf{Y}^v and

$$p(\mathbf{y}_d^v|\mathbf{X}^v, \boldsymbol{\theta}^v) = \mathcal{N}(\mathbf{y}_d^v|\mathbf{0}, \beta^{-1}\mathbf{I} + w_d^2\mathbf{K}^v). \quad (11)$$

Our multi-view iWMM assumes that the latent coordinates (per View) are generated from a Dirichlet process mixture model. In particular, we use the following infinite Gaussian mixture model,

$$p(\mathbf{x}^v|\lambda_c, \boldsymbol{\mu}_c, \mathbf{R}_c) = \sum_{c=1}^{\infty} \lambda_c \mathcal{N}(\mathbf{x}^v|\boldsymbol{\mu}_c, \mathbf{R}_c^{-1}), \quad (12)$$

where λ_c , $\boldsymbol{\mu}_c$ and \mathbf{R}_c are the mixture weight, mean, and precision matrix of the c th mixture component.

As in the iWMM [6], we place Gaussian-Wishart priors on the Gaussian parameters $\{\boldsymbol{\mu}_c, \mathbf{R}_c\}$.

$$p(\boldsymbol{\mu}_c, \mathbf{R}_c) = \mathcal{N}(\boldsymbol{\mu}_c | \mathbf{u}, (r\mathbf{R}_c)^{-1}) \mathcal{W}(\mathbf{R}_c | \mathbf{S}^{-1}, \nu) \quad (13)$$

where \mathbf{u} is the mean of $\boldsymbol{\mu}_c$, r is the relative precision of $\boldsymbol{\mu}_c$, \mathbf{S}^{-1} is the scale matrix for \mathbf{R}_c , and ν is the number of degrees of freedom for \mathbf{R}_c .

By using conjugate Gaussian-Wishart priors for the parameters of the Gaussian mixture components, we can analytically integrate out those parameters, given the assignments of points to components. Let z_n^v be the latent assignment of the n th object in the v th view. The probability of latent coordinates \mathbf{X} given latent assignments $\mathbf{Z}^v = (z_1, \dots, z_{N_v})$ is obtained by integrating out the Gaussian parameters $\{\boldsymbol{\mu}_c, \mathbf{R}_c\}$ as follows:

$$p(\mathbf{X} | \mathbf{Z}, \mathbf{S}, \nu, r) = \prod_{c=1}^{\infty} \pi^{-\frac{\sum_v N_{vc} Q}{2}} \frac{r^{Q/2} |\mathbf{S}|^{\nu/2}}{r_c^{Q/2} |\mathbf{S}_c|^{\nu_c/2}} \prod_{q=1}^Q \frac{\Gamma(\frac{\nu_c+1-q}{2})}{\Gamma(\frac{\nu+1-q}{2})}, \quad (14)$$

where N_{vc} is the number of objects in the v th view assigned to the c th cluster, $\Gamma(\cdot)$ is the Gamma function and

$$\begin{aligned} r_c &= r + \sum_v N_{vc}, \quad \nu_c = \nu + \sum_v N_{vc}, \\ \mathbf{u}_c &= \frac{r\mathbf{u} + \sum_{v=1}^V \sum_{n: z_{nv}=c} \mathbf{x}_{nv}}{r + \sum_v N_{vc}}, \\ \mathbf{S}_c &= \mathbf{S} + \sum_{v=1}^V \sum_{n: z_{nv}=c} \mathbf{x}_{nv} \mathbf{x}_{nv}^\top + r\mathbf{u}\mathbf{u}^\top - r_c \mathbf{u}_c \mathbf{u}_c^\top, \end{aligned}$$

are the posterior Gaussian-Wishart parameters of the c th component. As in [6], we use a Dirichlet process with concentration parameter η for infinite mixture modeling in the latent space.

The probability of \mathbf{Z} is given as follows

$$p(\mathbf{Z} | \eta) = \prod_{v=1}^V \frac{\eta^C \prod_{c=1}^C (N_{vc} - 1)!}{\eta(\eta + 1) \cdots (\eta + N_v - 1)}, \quad (15)$$

where C is the number of components for which $N_{vc} > 0$. The joint distribution is given by

$$p(\mathbf{Y}, \mathbf{X}, \mathbf{Z} | \boldsymbol{\theta}, \nu, \mathbf{u}, r\eta) = \prod_{v=1}^V p(\mathbf{Y}^v | \mathbf{X}^v, \boldsymbol{\theta}^v) p(\mathbf{X}^v | \mathbf{Z}^v, \mathbf{S}_v, \nu, \mathbf{u}, r) p(\mathbf{Z}^v | \eta). \quad (16)$$

3.3 Bayesian Multi-view Warped Mixture Model

Since we have a multi-view data set defined as $\mathcal{Y} = \{\mathbf{Y}^v\}_{v=1}^V$, where each view is set as $\mathbf{Y}^v \in \mathbb{R}^{N_v \times D_v}$, the joint likelihood is given by

$$p(\mathbf{Y}, \mathbf{X}, \mathbf{Z}) = \prod_{v=1}^V p(\mathbf{Y}^v | \mathbf{F}^v, \mathbf{X}^v, \boldsymbol{\theta}) p(\mathbf{F}^v | \mathbf{X}^v) p(\mathbf{X}^v | \mathbf{Z}^v) p(\mathbf{Z}^v) \quad (17)$$

We assume that we have observations (inputs) in different views, in which observations are generated by mapping the latent coordinates through a set of smooth functions, over which Gaussian process priors are placed.

$$p(\mathbf{Y}^v | \mathbf{F}^v, \mathbf{X}^v, \mathbf{Z}^v, \boldsymbol{\theta}) = \mathcal{N}(\mathbf{Y}^v | \mathbf{W}^v \mathbf{F}^v, \beta^{-1} \mathbf{I}) \quad (18)$$

References

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