Variational inference for Gaussian processes

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Topics on Deep Probabilistic Models





Acknowledgements





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Posterior inference

■ When using Bayesian inference, we need to compute the posterior distribution of f given the data.

We then use that posterior distribution to compute the predictive distribution.

- Reasons as why computing the posterior distribution is an issue for GPs.
 - Computational complexity.
 - Non-Gaussian likelihood.
 - Both of the above.

Computational complexity

□ To compute the predictive mean and the predictive covariance we need to compute $\left[\mathbf{K}(\mathbf{X},\mathbf{X}) + \sigma_n^2 \mathbf{I}\right]^{-1}$

The usual way to do this is using the Cholesky decomposition which costs $\mathcal{O}(n^3)$.

If n = 1000, then we need to perform 10^9 operations.

Non-Gaussian likelihoods

In Bayesian inference we want to compute

$$p(\mathbf{f}|\mathbf{y}) = \frac{p(\mathbf{y}|\mathbf{f})p(\mathbf{f})}{p(\mathbf{y})},$$

where $p(\mathbf{y}) = \int p(\mathbf{y}|\mathbf{f})p(\mathbf{f})d\mathbf{f}$.

□ When p(y|f) is a Gaussian likelihood, then we can compute p(y) and p(f|y) analytically.

When p(y|f) is non-Gaussian (e.g. Bernoulli with a sigmoid link function) both p(y) and p(f|y) are intractable.

How to address these issues?

 One successful approach is by using the idea of inducing variables or pseudo-variables.

The idea in itself was quite well know in the GP literature. See for example Chapter 8 of the GPML book and in the paper ?.

 However, if we couple this idea with a variational inference approach, we have a powerful tool to build complex GP models.

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- We introduce a new set of M variables $\mathbf{u} = \{u(\mathbf{z}_m)\}_{m=1}^M$ that we refer to as inducing variables or pseudo-variables.
- □ The set of points $\mathbf{Z} = \{\mathbf{z}_m\}_{m=1}^M$ is usually known as inducing inputs.
- □ We augment the original prior $p(\mathbf{f})$ to $p(\mathbf{f}, \mathbf{u})$ such that

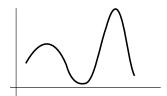
$$p(\mathbf{f}) = \int p(\mathbf{f}, \mathbf{u}) d\mathbf{u} = \int p(\mathbf{f}|\mathbf{u}) p(\mathbf{u}) d\mathbf{u},$$

where $p(\mathbf{u})$ and $p(\mathbf{f}, \mathbf{u})$ are both Gaussians.

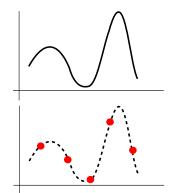
- The auxiliary variables \mathbf{u} can be part of the GP $f(\mathbf{x})$ or they can be linearly related to $f(\mathbf{x})$ (sometimes known as interdomain inducing variables).
- □ In the former case, $p(\mathbf{u}) = \mathcal{N}(\mathbf{u}|\mathbf{0}, \mathbf{K}(\mathbf{Z}, \mathbf{Z}))$.



A sample from p(f)

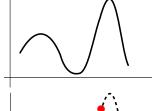


A sample from p(f)

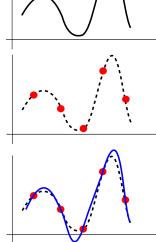


Inducing variables **u**

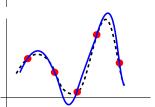
A sample from p(f)



Inducing variables u



A sample from $p(f|\mathbf{u})$



Variational lower-bound for the marginal likelihood (I)

We can write the log marginal probability for y using

$$\log p(\mathbf{y}) = \mathcal{L}(q(\mathbf{f})) + \mathsf{KL}(q(\mathbf{f}) || p(\mathbf{f} | \mathbf{y})),$$

where

$$\mathcal{L}(q(\mathbf{f})) = \int q(\mathbf{f}) \log \left\{ rac{p(\mathbf{y}, \mathbf{f})}{q(\mathbf{f})}
ight\} d\mathbf{f},$$
 $\mathsf{KL}(q(\mathbf{f}) || p(\mathbf{f} | \mathbf{y})) = - \int q(\mathbf{f}) \log \left\{ rac{p(\mathbf{f} | \mathbf{y})}{q(\mathbf{f})}
ight\} d\mathbf{f},$

with $q(\mathbf{f})$ the approximated posterior, KL(q||p) is the Kullback-Leibler divergence between q and p and $p(\mathbf{f}|\mathbf{y})$ is the true posterior.

- □ The KL divergence is zero when q = p. In that case, $\log p(\mathbf{y}) = \mathcal{L}(q(\mathbf{f}))$.
- □ If this is not the case $KL(q(\mathbf{f})||p(\mathbf{f}|\mathbf{y}))$ and $\log p(\mathbf{y}) > \mathcal{L}(q(\mathbf{f}))$.



Variational lower-bound for the marginal likelihood (II)

- \Box We have two ways to find the optimal $q(\mathbf{f})$
 - 1. We find $q(\mathbf{f})$ by minimising $KL(q(\mathbf{f})||p(\mathbf{f}|\mathbf{y}))$.
 - 2. We find $q(\mathbf{f})$ by maximising $\mathcal{L}(q(\mathbf{f}))$.
- Option 1 is not possible since $p(\mathbf{f}|\mathbf{y})$ is unknown.
- So, in general, we appeal to option 2

$$\log p(\mathbf{y}) \geq \mathcal{L}(q(\mathbf{f})).$$

Lower-bound with inducing variables

For our augmented model we want to find an approximated posterior $q(\mathbf{f}, \mathbf{u})$ by maximising

$$\mathcal{L}(q(\mathbf{f}, \mathbf{u})) = \int \int q(\mathbf{f}, \mathbf{u}) \log \left\{ \frac{p(\mathbf{y}, \mathbf{f}, \mathbf{u})}{q(\mathbf{f}, \mathbf{u})} \right\} d\mathbf{u} d\mathbf{f},$$

where $p(\mathbf{y}, \mathbf{f}, \mathbf{u}) = p(\mathbf{y}|\mathbf{f})p(\mathbf{f}|\mathbf{u})p(\mathbf{u})$ and we will approximate the posterior $q(\mathbf{f}, \mathbf{u})$ as $q(\mathbf{f}, \mathbf{u}) \approx p(\mathbf{f}|\mathbf{u})q(\mathbf{u})$.

Since we know that $q(\mathbf{f}) = \int_{\mathbf{u}} p(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) d\mathbf{u}$, the bound above really only depends on $q(\mathbf{u})$

$$\begin{split} \mathcal{L}(q(\mathbf{u})) &= \int \int \rho(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) \log \left\{ \frac{\rho(\mathbf{y}|\mathbf{f}) \rho(\mathbf{f}|\mathbf{u}) \rho(\mathbf{u})}{\rho(\mathbf{f}|\mathbf{u}) q(\mathbf{u})} \right\} d\mathbf{u} d\mathbf{f}, \\ &= \int \int \rho(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) \log \left\{ \frac{\rho(\mathbf{y}|\mathbf{f}) \rho(\mathbf{u})}{q(\mathbf{u})} \right\} d\mathbf{u} d\mathbf{f}. \end{split}$$



Two approaches for optimising $\mathcal{L}(q(\mathbf{u}))$

- □ There are two approaches for optimising $q(\mathbf{u})$ in $\mathcal{L}(q(\mathbf{u}))$.
- □ First approach (?):
 - We assume a multi-variate Gaussian form for $q(\mathbf{u}) = \mathcal{N}(\mathbf{u}|\boldsymbol{\mu}, \mathbf{S})$ with $\boldsymbol{\mu} \in \mathbb{R}^{M \times 1}$ and $\mathbf{S} \in \mathbb{R}^{M \times M}$.
 - We then find μ and **S** by numerically optimising $\mathcal{L}(q(\mathbf{u}))$.
- Second approach (?):
 - We marginalise q(u) from the bound and then compute it by using Jensen's inequality.
 - We then find μ and **S** by using the rules of probability.

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Stochastic variational inference

Stochastic variational inference (SVI) allows (?) the use of stochastic gradients over variational lower bounds.

? proposed the use of SVI for sparse GPs.

The idea is to use stochastic gradients for optimising $\mathcal{L}(q(\mathbf{u}))$ with respect to $q(\mathbf{u})$, this is, μ and \mathbf{S} .

Lower bound $\mathcal{L}(q(\mathbf{u}))$ (I)

From a previous slide,

$$\mathcal{L}(q(\mathbf{u})) = \int \int \rho(\mathbf{f}|\mathbf{u})q(\mathbf{u}) \log \left\{ \frac{\rho(\mathbf{y}|\mathbf{f})\rho(\mathbf{u})}{q(\mathbf{u})} \right\} d\mathbf{u} d\mathbf{f}.$$

We can re-arrange the expression above using

$$\begin{split} \mathcal{L}(q(\mathbf{u})) &= \int \int \rho(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) \log \left\{ \frac{\rho(\mathbf{y}|\mathbf{f}) \rho(\mathbf{u})}{q(\mathbf{u})} \right\} d\mathbf{u} d\mathbf{f}, \\ &= \int \int \rho(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) \left[\log p(\mathbf{y}|\mathbf{f}) + \log \frac{\rho(\mathbf{u})}{q(\mathbf{u})} \right] d\mathbf{u} d\mathbf{f}, \\ &= \int \int \rho(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) \log p(\mathbf{y}|\mathbf{f}) d\mathbf{u} d\mathbf{f} + \int q(\mathbf{u}) \log \frac{\rho(\mathbf{u})}{q(\mathbf{u})} d\mathbf{u}, \\ &= \int \log p(\mathbf{y}|\mathbf{f}) \int p(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) d\mathbf{u} d\mathbf{f} + \int q(\mathbf{u}) \log \frac{\rho(\mathbf{u})}{q(\mathbf{u})} d\mathbf{u}, \\ &= \int \log p(\mathbf{y}|\mathbf{f}) q(\mathbf{f}) d\mathbf{f} - \mathrm{KL}(q(\mathbf{u}) \| p(\mathbf{u})) d\mathbf{u}, \\ &= \mathbb{E}_{q(\mathbf{f})}[\log p(\mathbf{y}|\mathbf{f})] - \mathrm{KL}(q(\mathbf{u}) \| p(\mathbf{u})). \end{split}$$



Lower bound $\mathcal{L}(q(\mathbf{u}))$ (II)

The lower bound is

$$\mathcal{L}(q(\mathbf{u})) = \mathbb{E}_{q(\mathbf{f})}[\log p(\mathbf{y}|\mathbf{f})] - \mathsf{KL}(q(\mathbf{u})||p(\mathbf{u})).$$

We can find an estimate for μ and **S** by maximising $\mathcal{L}(q(\mathbf{u}))$ using numerical optimisation.

We need to compute the gradients

$$rac{\partial \mathcal{L}}{\partial oldsymbol{\mu}}, \; rac{\partial \mathcal{L}}{\partial oldsymbol{\mathsf{S}}}$$

Exercises

□ Recall that $q(\mathbf{u}) = \mathcal{N}(\mathbf{u}|\boldsymbol{\mu}, \mathbf{S})$ and $p(\mathbf{u}) = \mathcal{N}(\mathbf{u}|\mathbf{0}, \mathbf{K}(\mathbf{Z}, \mathbf{Z}))$. Now, for the regression case,

$$p(\mathbf{y}|\mathbf{f}) = \mathcal{N}(\mathbf{y}|\mathbf{f}, \sigma_n^2 \mathbf{I}).$$

- 1. Write the expression for the bound in terms of μ , **S** and σ_n^2 .
- 2. Compute $\frac{\partial \mathcal{L}}{\partial \boldsymbol{\mu}}$ and $\frac{\partial \mathcal{L}}{\partial \mathbf{S}}$.
- For the case above, what is the computational complexity of this method?

Stochastic gradient descend

It can be shown that the lower bound can be written as

$$\mathcal{L}(q(\mathbf{u})) = \sum_{i=1}^{n} \ell(y_i, \mathbf{x}_i, \theta) - \mathsf{KL}(q(\mathbf{u}) || p(\mathbf{u})),$$

where $\ell(y_i, \mathbf{x}_i, \theta)$ is a function that depends on the data, the variational parameters μ , **S**, and any other (hyper) parameters in the model (e.g. the hyperparameters of the kernel).

- □ For *n* large, we could only use a subset of the data to compute the gradients to be used in numerical optimisation.
- ☐ This is usually known as *stochastic gradient descend*.
- □ The computational complexity is this model is $\mathcal{O}(nM^2)$, where M is the number of inducing points.

Example SVI for Sparse GPs

Lab SVI for Sparse GPs (GPFlow)

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Marginalising $q(\mathbf{u})$ from the bound (I)

- Instead of finding particular parameters μ and **S** as before, we can marginalise $q(\mathbf{u})$ from the lower bound and then use probability rules to compute $q(\mathbf{u})$.
- Let us go back to the general expression for the bound

$$\mathcal{L}(q(\mathbf{u})) = \int \int p(\mathbf{f}|\mathbf{u})q(\mathbf{u}) \log \left\{ \frac{p(\mathbf{y}|\mathbf{f})p(\mathbf{u})}{q(\mathbf{u})} \right\} d\mathbf{u}d\mathbf{f}.$$

- Let us assume that $q(\mathbf{u}) = \mathcal{N}(\mathbf{u}|\boldsymbol{\mu}, \mathbf{S})$ and find expressions for \mathbf{u} and \mathbf{S} .
- We first integrate over f.

Marginalising $q(\mathbf{u})$ from the bound (II)

The bound can be expressed as

$$\begin{split} \mathcal{L}(q(\mathbf{u})) &= \int \int \rho(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) \log \left\{ \frac{\rho(\mathbf{y}|\mathbf{f}) \rho(\mathbf{u})}{q(\mathbf{u})} \right\} d\mathbf{u} d\mathbf{f} \\ &= \int q(\mathbf{u}) \int \rho(\mathbf{f}|\mathbf{u}) \left\{ \log \rho(\mathbf{y}|\mathbf{f}) + \log \left[\frac{\rho(\mathbf{u})}{q(\mathbf{u})} \right] \right\} d\mathbf{f} d\mathbf{u}. \end{split}$$

Let us focus on the integral over f

$$\log T(\mathbf{y}, \mathbf{u}) = \int \log p(\mathbf{y}|\mathbf{f})p(\mathbf{f}|\mathbf{u})d\mathbf{f}.$$

Exercise

 $\rho(\mathbf{f}|\mathbf{u})$ is a conditional Gaussian distribution given as,

$$\mathcal{N}(\mathbf{f}|\mathbf{K}(\mathbf{X},\mathbf{Z})\mathbf{K}^{-1}(\mathbf{Z},\mathbf{Z})\mathbf{u},\mathbf{K}(\mathbf{X},\mathbf{X})-\mathbf{K}(\mathbf{X},\mathbf{Z})\mathbf{K}^{-1}(\mathbf{Z},\mathbf{Z})\mathbf{K}(\mathbf{X},\mathbf{Z})^{\top}),$$
 and $p(\mathbf{y}|\mathbf{f})$ is again a Gaussian given as $\mathcal{N}(\mathbf{y}|\mathbf{f},\sigma_n^2\mathbf{I})$.

Compute the expression $\log T(\mathbf{y}, \mathbf{u})$

Marginalising $q(\mathbf{u})$ from the bound (III)

The bound can now be expressed as

$$\mathcal{L}(q(\mathbf{u})) = \int q(\mathbf{u}) \left\{ \log \mathcal{N}(\mathbf{y} | \alpha, \sigma_n^2 \mathbf{I}) - \frac{1}{2} \operatorname{trace}(\sigma_n^{-2} \widetilde{\mathbf{K}}) + \log \left[\frac{p(\mathbf{u})}{q(\mathbf{u})} \right] \right\} d\mathbf{u},$$

where

$$\begin{split} &\alpha = \mathbf{K}(\mathbf{X},\mathbf{Z})\mathbf{K}^{-1}(\mathbf{Z},\mathbf{Z})\mathbf{u} \\ &\widetilde{\mathbf{K}} = \mathbf{K}(\mathbf{X},\mathbf{X}) - \mathbf{K}(\mathbf{X},\mathbf{Z})\mathbf{K}^{-1}(\mathbf{Z},\mathbf{Z})\mathbf{K}(\mathbf{X},\mathbf{Z})^{\top}. \end{split}$$

It follows that

$$\mathcal{L}(q(\mathbf{u})) = \int q(\mathbf{u}) \left\{ \log \left[\frac{\mathcal{N}(\mathbf{y}|lpha, \sigma_n^2 \mathbf{I}) p(\mathbf{u})}{q(\mathbf{u})}
ight]
ight\} d\mathbf{u} - \frac{1}{2} \operatorname{trace}(\sigma_n^{-2} \widetilde{\mathbf{K}})$$



Jensen's inequality

 \Box A function φ is *convex* if

$$\varphi(\lambda a + (1 - \lambda)b) \le \lambda \varphi(a) + (1 - \lambda)\varphi(b).$$

 \Box A function φ is *concave* if

$$\varphi(\lambda a + (1 - \lambda)b) \ge \lambda \varphi(a) + (1 - \lambda)\varphi(b).$$

 \Box Let φ be a convex function. It can be shown that

$$arphi(\mathbb{E}(\mathbf{x})) \leq \mathbb{E}(arphi(\mathbf{x}))$$
 $arphi\left(\int \mathbf{x} p(\mathbf{x}) d\mathbf{x}\right) \leq \int arphi(\mathbf{x}) p(\mathbf{x}) d\mathbf{x}.$

This inequality is known as the Jensen's inequality.

 \Box If φ is a concave function then

$$arphi(\mathbb{E}(\mathbf{x})) \geq \mathbb{E}(arphi(\mathbf{x}))$$
 $arphi\left(\int \mathbf{x} p(\mathbf{x}) d\mathbf{x}\right) \geq \int arphi(\mathbf{x}) p(\mathbf{x}) d\mathbf{x}.$

Jensen's inequality applied to $\mathcal{L}(q(\mathbf{u}))$

Reversing Jensen's inequality, we can write

$$\begin{split} &\log\left[\int q(\mathbf{u})\left[\frac{\mathcal{N}(\mathbf{y}|\alpha,\sigma_n^2\mathbf{I})p(\mathbf{u})}{q(\mathbf{u})}\right]d\mathbf{u}\right] \geq \\ &\int q(\mathbf{u})\left\{\log\left[\frac{\mathcal{N}(\mathbf{y}|\alpha,\sigma_n^2\mathbf{I})p(\mathbf{u})}{q(\mathbf{u})}\right]\right\}d\mathbf{u} \end{split}$$

The expression above can be simplified as

$$\log \left[\int \mathcal{N}(\mathbf{y}|\boldsymbol{\alpha}, \sigma_n^2 \mathbf{I}) p(\mathbf{u}) d\mathbf{u} \right] \ge \int q(\mathbf{u}) \left\{ \log \left[\frac{\mathcal{N}(\mathbf{y}|\boldsymbol{\alpha}, \sigma_n^2 \mathbf{I}) p(\mathbf{u})}{q(\mathbf{u})} \right] \right\} d\mathbf{u}$$

A tighter bound $\mathcal{L}(q(\mathbf{u}))$

Reversing Jensen's inequality, we can write

$$egin{aligned} \log \left[\int q(\mathbf{u}) \left[rac{\mathcal{N}(\mathbf{y} | oldsymbol{lpha}, \sigma_n^2 \mathbf{I}) p(\mathbf{u})}{q(\mathbf{u})}
ight] d\mathbf{u}
ight] & \geq \ \int q(\mathbf{u}) \left\{ \log \left[rac{\mathcal{N}(\mathbf{y} | oldsymbol{lpha}, \sigma_n^2 \mathbf{I}) p(\mathbf{u})}{q(\mathbf{u})}
ight]
ight\} d\mathbf{u} \end{aligned}$$

The expression above can be simplified as

$$\log \left[\int \mathcal{N}(\mathbf{y}|\alpha, \sigma_n^2 \mathbf{I}) p(\mathbf{u}) d\mathbf{u} \right] \ge \int q(\mathbf{u}) \left\{ \log \left[\frac{\mathcal{N}(\mathbf{y}|\alpha, \sigma_n^2 \mathbf{I}) p(\mathbf{u})}{q(\mathbf{u})} \right] \right\} d\mathbf{u}$$

If we define

$$\mathcal{L}_2 = \log \left[\int \mathcal{N}(\mathbf{y} | \alpha, \sigma_n^2 \mathbf{I}) p(\mathbf{u}) d\mathbf{u} \right] - \frac{1}{2} \operatorname{trace}(\sigma_n^{-2} \widetilde{\mathbf{K}}),$$

then

$$\mathcal{L}_2 \geq \mathcal{L}(\textit{q}(\textbf{u})).$$

And \mathcal{L}_2 is closer to $\log p(\mathbf{y})$ than $\mathcal{L}(q(\mathbf{u}))$.



Exercise

 \Box Show that \mathcal{L}_2 can be written as

$$\mathcal{L}_2 = \log \mathcal{N}(\mathbf{y}|\mathbf{0}, \mathbf{K}(\mathbf{X}, \mathbf{Z})\mathbf{K}^{-1}(\mathbf{Z}, \mathbf{Z})\mathbf{K}^{\top}(\mathbf{X}, \mathbf{Z}) + \sigma_n^2 \mathbf{I}) - \frac{1}{2}\operatorname{trace}(\sigma_n^{-2}\widetilde{\mathbf{K}}).$$

References I