#### **Approximate Inference for Neural Networks**

Mauricio A. Álvarez PhD, H.F. Garcia C. Guarnizo (TA)



Universidad Tecnológica de Pereira, Pereira, Colombia

#### **Outline**

EM algorithm

2 Variational EM

- 3 Approximate inference for Neural Networks
  - Bayesian Deep Neural Networks
  - Laplace Approximation



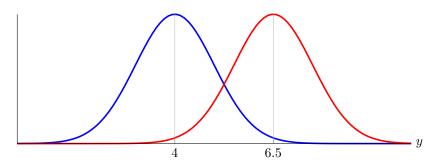
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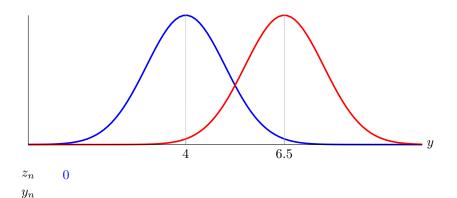




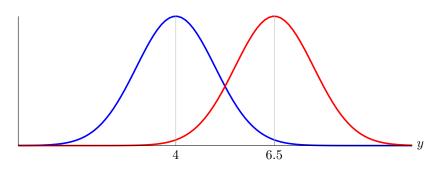
 $z_n$ 

 $y_n$ 



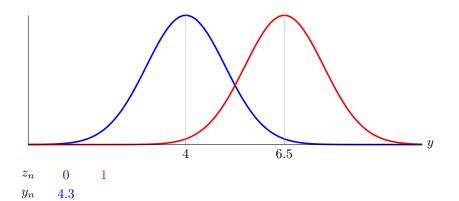




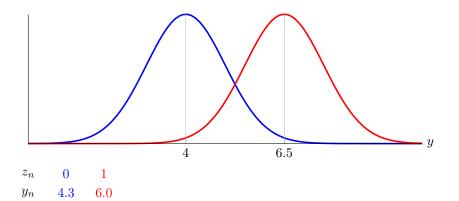


$$z_n = 0$$
  
 $y_n = 4.3$ 

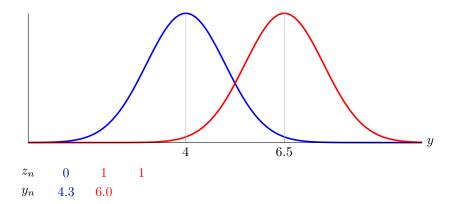




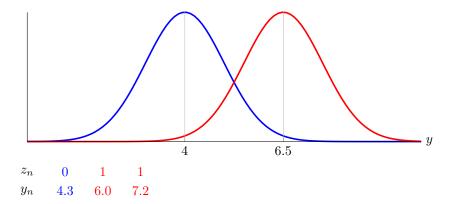




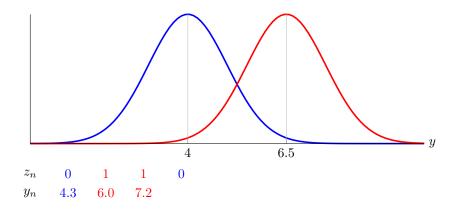




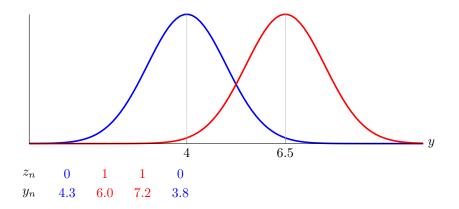














For z:

$$p(z) = \pi^{z} (1 - \pi)^{1 - z}, \quad p(z = 1) = \pi, \quad p(z = 0) = 1 - \pi.$$

For y:

$$p(y|z) = \left( \mathcal{N} \left( y | \mu_1, \sigma_1^2 \right) \right)^z \left( \mathcal{N} \left( y | \mu_2, \sigma_2^2 \right) \right)^{1-z}.$$

The the joint probability:

$$p(y,z) = \left(\pi \mathcal{N}\left(y|\mu_1, \sigma_1^2\right)\right)^z \left((1-\pi)\mathcal{N}\left(y|\mu_2, \sigma_2^2\right)\right)^{1-z}.$$

Additionally:

$$p(z = 1|y) = \frac{p(y|z = 1)p(z = 1)}{p(y)}$$

$$= \frac{\mathcal{N}(y|\mu_1, \sigma_1^2)\pi}{\pi \mathcal{N}(y|\mu_1, \sigma_1^2) + (1 - \pi)\mathcal{N}(y|\mu_2, \sigma_2^2)}$$



Let's assume that we have  $(\mathbf{y}, \mathbf{z}) = \{y_n, z_n\}_{n=1}^N$ . The parameters  $\theta = \{\pi, \mu_1, \mu_2, \sigma_1^2, \sigma_2^2\}$ , can be estimated as

$$\theta_{\mathsf{ML}} = \underset{\theta}{\operatorname{arg \, max}} \, \ln \left( p(\mathbf{y}, \mathbf{z}|\theta) \right),$$

$$= \underset{\theta}{\operatorname{arg \, max}} \, \ln \left( \prod_{n=1}^{N} p(y_n, z_n|\theta) \right),$$

$$= \underset{\theta}{\operatorname{arg \, max}} \, \sum_{n=1}^{N} \ln \left( p(y_n, z_n|\theta) \right).$$

Where:

$$\ln p(y_n, z_n | \theta) = z_n \left( \ln(\pi) + \ln \mathcal{N} \left( y_n | \mu_1, \sigma_1^2 \right) \right)$$

$$+ (1 - z_n) \left( \ln(1 - \pi) + \ln \mathcal{N} \left( y_n | \mu_2, \sigma_2^2 \right) \right).$$



Let's assume that  $\mathbf{y} = \{y_n\}_{n=1}^N$ ,  $\mathbf{y}$   $\mathbf{z}$  is unknown (latent). The parameters  $\theta = \{\pi, \mu_1, \mu_2, \sigma_1^2, \sigma_2^2\}$ , can be estimated as

$$\theta_{\mathsf{ML}} = \underset{\theta}{\operatorname{arg\,max}} \ \sum_{n=1}^{N} \ln \left( p(y_n | \theta) \right),$$
$$= \underset{\theta}{\operatorname{arg\,max}} \ \sum_{n=1}^{N} \ln \left( \sum_{z} p(y_n, z_n | \theta) \right).$$

Where:

$$\ln\left(\sum_{z} p(y_n, z_n | \theta)\right) = \ln\left(\pi \mathcal{N}\left(y_n | \mu_1, \sigma_1^2\right) + (1 - \pi) \mathcal{N}\left(y_n | \mu_2, \sigma_2^2\right)\right).$$



#### **Problem definition - General**

The model consists of observations  $\mathbf y$  and a latent random variable  $\mathbf z.$  Then

$$p(\mathbf{y}|\theta) = \prod_{n=1}^{N} p(y_n|\theta)$$
$$= \prod_{n=1}^{N} \sum_{\mathbf{z}} p(y_n, \mathbf{z}|\theta) = \prod_{n=1}^{N} \sum_{\mathbf{z}} p(y_n|\mathbf{z}, \theta) p(\mathbf{z}|\theta)$$

The estimation of  $\theta_{ML}$  is given by

$$\theta_{\mathsf{ML}} = \underset{\theta}{\arg\max} \ln \left( p(\mathbf{y}|\theta) \right)$$
$$= \underset{\theta}{\arg\max} \sum_{n=1}^{N} \ln \left( \sum_{\mathbf{z}} p(y_n|\mathbf{z}, \theta) p(\mathbf{z}|\theta) \right)$$



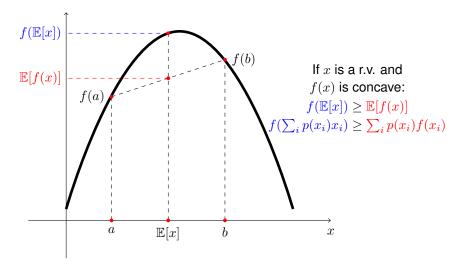
#### **Problem definition**

$$\theta_{\mathsf{ML}} = \underset{\theta}{\arg\max} \sum_{n=1}^{N} \ln \left( \sum_{\mathbf{z}} p(y_n, |\mathbf{z}, \theta) p(\mathbf{z} | \theta) \right)$$

- The sum over z couples the parameters  $\theta$ .
- Gradients have no closed form.
- $\mathbf{z}$  and  $\theta$  are also coupled.



#### Jensen's Inequality





## EM algorithm - Jensen's Inequality

$$\log p(\mathbf{y}|\theta) = \sum_{n=1}^{N} \ln \left( \sum_{\mathbf{z}} p(y_n, \mathbf{z}|\theta) \right)$$



## EM algorithm - Jensen's Inequality

$$\log p(\mathbf{y}|\theta) = \sum_{n=1}^{N} \ln \left( \sum_{\mathbf{z}} p(y_n, \mathbf{z}|\theta) \right)$$
$$= \sum_{n=1}^{N} \ln \left( \sum_{\mathbf{z}} q(\mathbf{z}) \frac{p(y_n, \mathbf{z}|\theta)}{q(\mathbf{z})} \right)$$



## EM algorithm - Jensen's Inequality

$$\log p(\mathbf{y}|\theta) = \sum_{n=1}^{N} \ln \left( \sum_{\mathbf{z}} p(y_n, \mathbf{z}|\theta) \right)$$
$$= \sum_{n=1}^{N} \ln \left( \sum_{\mathbf{z}} q(\mathbf{z}) \frac{p(y_n, \mathbf{z}|\theta)}{q(\mathbf{z})} \right)$$
$$\geq \sum_{n=1}^{N} \sum_{\mathbf{z}} q(\mathbf{z}) \ln \left( \frac{p(y_n, \mathbf{z}|\theta)}{q(\mathbf{z})} \right)$$



Assuming a value of  $\theta^{(t)}$ , what form  $q(\mathbf{z})$  should have to maximize

$$\ln p(\mathbf{y}|\theta^{(t)}) \ge \sum_{n=1}^{N} \sum_{\mathbf{z}} q(\mathbf{z}) \ln \left( \frac{p(\mathbf{z}|y_n, \theta^{(t)}) p(y_n|\theta^{(t)})}{q(\mathbf{z})} \right)$$



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$$\ge \sum_{n=1}^{N} \sum_{\mathbf{z}} q(\mathbf{z}) \left[ \ln \left( \frac{p(\mathbf{z}|y_n, \theta^{(t)})}{q(\mathbf{z})} \right) + \ln p(y_n|\theta^{(t)}) \right]$$



Assuming a value of  $\theta^{(t)}$ , what form  $q(\mathbf{z})$  should have to maximize

$$\ln p(\mathbf{y}|\boldsymbol{\theta}^{(t)}) \ge \sum_{n=1}^{N} \sum_{\mathbf{z}} q(\mathbf{z}) \ln \left( \frac{p(\mathbf{z}|y_n, \boldsymbol{\theta}^{(t)}) p(y_n|\boldsymbol{\theta}^{(t)})}{q(\mathbf{z})} \right)$$
$$\ge \sum_{n=1}^{N} \sum_{\mathbf{z}} q(\mathbf{z}) \left[ \ln \left( \frac{p(\mathbf{z}|y_n, \boldsymbol{\theta}^{(t)})}{q(\mathbf{z})} \right) + \ln p(y_n|\boldsymbol{\theta}^{(t)}) \right]$$

organizing,

$$\sum_{n=1}^{N} \ln p(y_n|\boldsymbol{\theta}^{(t)}) \geq \sum_{n=1}^{N} \sum_{\mathbf{z}} q(\mathbf{z}) \ln \left( \frac{p(\mathbf{z}|y_n, \boldsymbol{\theta}^{(t)})}{q(\mathbf{z})} \right) + \sum_{n=1}^{N} \ln p(y_n|\boldsymbol{\theta}^{(t)}),$$



Assuming a value of  $\theta^{(t)}$ , what form  $q(\mathbf{z})$  should have to maximize

$$\ln p(\mathbf{y}|\theta^{(t)}) \ge \sum_{n=1}^{N} \sum_{\mathbf{z}} q(\mathbf{z}) \ln \left( \frac{p(\mathbf{z}|y_n, \theta^{(t)}) p(y_n|\theta^{(t)})}{q(\mathbf{z})} \right)$$
$$\ge \sum_{n=1}^{N} \sum_{\mathbf{z}} q(\mathbf{z}) \left[ \ln \left( \frac{p(\mathbf{z}|y_n, \theta^{(t)})}{q(\mathbf{z})} \right) + \ln p(y_n|\theta^{(t)}) \right]$$

organizing,

$$\sum_{n=1}^N \ln p(y_n|\boldsymbol{\theta}^{(t)}) \geq \sum_{n=1}^N \sum_{\mathbf{z}} q(\mathbf{z}) \ln \left( \frac{p(\mathbf{z}|y_n,\boldsymbol{\theta}^{(t)})}{q(\mathbf{z})} \right) + \sum_{n=1}^N \ln p(y_n|\boldsymbol{\theta}^{(t)}),$$

then,  $q(\mathbf{z}) = p(\mathbf{z}|y_n, \theta^{(t)}).$ 

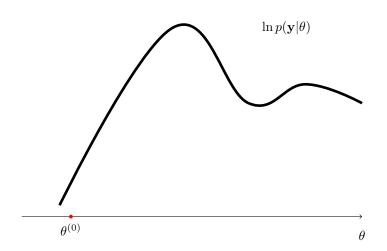


#### **Parameter estimation**

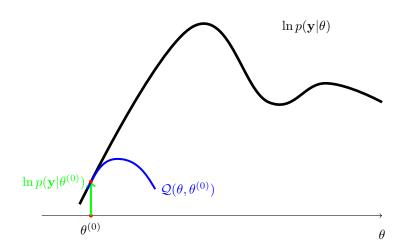
If 
$$q(\mathbf{z}) = p(\mathbf{z}|y_n, \theta^{(t)})$$
,

$$\theta^{(t+1)} = \arg\max_{\theta} \sum_{n=1}^{N} \sum_{\mathbf{z}} p(\mathbf{z}|y_n, \theta^{(t)}) \ln\left(\frac{p(y_n, \mathbf{z}|\theta)}{p(\mathbf{z}|y_n, \theta^{(t)})}\right)$$
$$= \arg\max_{\theta} \sum_{n=1}^{N} \mathbb{E}_{q(\mathbf{z})} \left[\ln\left(p(y_n, \mathbf{z}|\theta)\right)\right]$$
$$= \arg\max_{\theta} \mathcal{Q}\left(\theta, \theta^{(t)}\right)$$

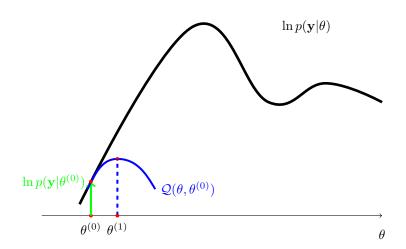




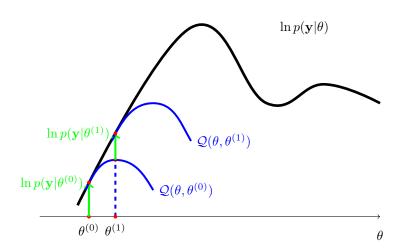




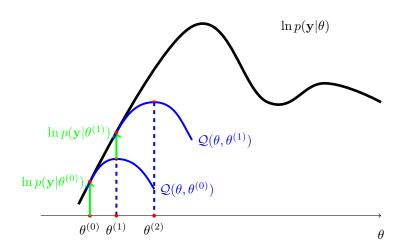




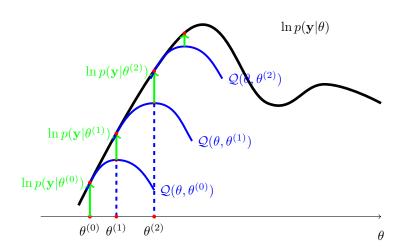














#### **EM Algorithm - Pseudo-code**

Given the joint distribution  $p(\mathbf{y}, \mathbf{z}|\theta)$ , the objective is to maximize  $p(\mathbf{y}|\theta)$  w.r.t.  $\theta$ .

- **1** Select the initial parameters  $\theta^{(t)} \leftarrow \theta^{(0)}$ .
- **2** Evaluate E-step,  $q(\mathbf{z}) = p(\mathbf{z}|\mathbf{y}, \theta^{(t)})$ .
- 3 Evaluate M-step,

$$\boldsymbol{\theta}^{(t+1)} = \operatorname*{arg\,max}_{\boldsymbol{\theta}} \mathcal{Q}\left(\boldsymbol{\theta}, \boldsymbol{\theta}^{(t)}\right)$$

Verify convergence. If it doesn't satisfy, then

$$\theta^{(t)} \leftarrow \theta^{(t+1)},$$

return to step 2.



#### EM algorithm - Remark

$$\log p(\mathbf{y}|\theta) \ge \sum_{n=1}^{N} \sum_{\mathbf{z}} q(\mathbf{z}) \ln \left( \frac{p(y_n, \mathbf{z}|\theta)}{q(\mathbf{z})} \right)$$

$$\sum_{n=1}^{N} \ln p(y_n|\theta) \ge \sum_{n=1}^{N} \sum_{\mathbf{z}} q(\mathbf{z}) \ln \left( \frac{p(\mathbf{z}|y_n, \theta)}{q(\mathbf{z})} \right) + \sum_{n=1}^{N} \ln p(y_n|\theta),$$

$$\sum_{n=1}^{N} \ln p(y_n|\theta) \ge \sum_{n=1}^{N} \sum_{\mathbf{z}} q(\mathbf{z}) \ln \left( p(y_n|\mathbf{z},\theta) \right) + \sum_{\mathbf{z}} q(\mathbf{z}) \ln \left( \frac{p(\mathbf{z}|\theta)}{q(\mathbf{z})} \right).$$

$$\sum_{n=1}^{N} \ln p(y_n|\theta) \ge \sum_{n=1}^{N} \sum_{\mathbf{z}} q(\mathbf{z}) \ln \left( p(y_n, \mathbf{z}|\theta) \right) - \sum_{\mathbf{z}} q(\mathbf{z}) \ln \left( q(\mathbf{z}) \right).$$



#### **EM algorithm - Summary**

- We just need to calculate the expected value of the joint distribution w.r.t. the posterior of the latent variables.
- **2** We are able to estimate the parameters  $\theta$  by maximizing an easier objective function.



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#### **Problem formulation**

In this case,

$$p(\mathbf{z}|\mathbf{y}) = \frac{p(\mathbf{y}, \mathbf{z})}{p(\mathbf{y})}$$

is intractable. Most of the time because

$$\begin{aligned} p(\mathbf{y}) &= \int_{\mathbf{z}} p(\mathbf{y}, \mathbf{z}) \, \mathrm{d} \, \mathbf{z} \\ &= \int_{\mathbf{z}} p(\mathbf{y} | \mathbf{z}) p(\mathbf{z}) \, \mathrm{d} \, \mathbf{z} \\ &= \mathbb{E}_{p(\mathbf{z})} \left[ p(\mathbf{y} | \mathbf{z}) \right] \end{aligned}$$

is difficult to calculate.



## **Variational EM - Objective**

We are interested in estimating  $p(\mathbf{z}|\mathbf{y})$ . This can be achieved by solving the following optimization problem,

$$q^*(\mathbf{z}) = \text{arg min } \mathrm{KL}(q(\mathbf{z})||p(\mathbf{z}|\mathbf{y}))$$

where

$$\mathrm{KL}(q(\mathbf{z})||p(\mathbf{z}|\mathbf{y})) = \int q(\mathbf{z}) \ln \frac{q(\mathbf{z})}{p(\mathbf{z}|\mathbf{y})} \mathsf{d}\mathbf{z}$$



# Variational EM - Objective

Analysing the KL divergence between  $q(\mathbf{z})$  and  $p(\mathbf{z}|\mathbf{y})$ ,

$$\begin{split} \mathrm{KL}(q(\mathbf{z})||p(\mathbf{z}|\mathbf{y})) &= \int q(\mathbf{z}) \ln \frac{q(\mathbf{z})}{p(\mathbf{z}|\mathbf{y})} \mathsf{d}\mathbf{z} \\ &= \int q(\mathbf{z}) \ln q(\mathbf{z}) \mathsf{d}\mathbf{z} - \int q(\mathbf{z}) \ln p(\mathbf{z}|\mathbf{y}) \mathsf{d}\mathbf{z} \\ &= \int q(\mathbf{z}) \ln q(\mathbf{z}) \mathsf{d}\mathbf{z} - \int q(\mathbf{z}) \ln \frac{p(\mathbf{z},\mathbf{y})}{p(\mathbf{y})} \mathsf{d}\mathbf{z} \\ &= \int q(\mathbf{z}) \ln q(\mathbf{z}) \mathsf{d}\mathbf{z} - \int q(\mathbf{z}) \ln p(\mathbf{z},\mathbf{y}) \mathsf{d}\mathbf{z} + \ln p(\mathbf{y}) \end{split}$$



#### Variational EM - The evidence lower bound

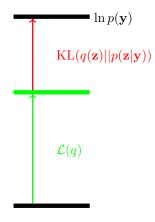
From previous equation we have,

$$\ln p(\mathbf{y}) = \text{KL}(q(\mathbf{z})||p(\mathbf{z}|\mathbf{y})) + \int q(\mathbf{z}) \ln p(\mathbf{z}, \mathbf{y}) d\mathbf{z} - \int q(\mathbf{z}) \ln q(\mathbf{z}) d\mathbf{z}$$
$$= \text{KL}(q(\mathbf{z})||p(\mathbf{z}|\mathbf{y})) + \mathcal{L}(q).$$

Given that  $\mathrm{KL}(q(\mathbf{z})||p(\mathbf{z}|\mathbf{y})) \geq 0$ , then the lower bound of  $\ln p(\mathbf{y})$  is  $\mathcal{L}(q)$ .



## Variational EM - The evidence lower bound





# **Variational EM - Summary**

- If the posterior  $p(\mathbf{z}|\mathbf{y})$  is intractable or complex (main issue is to calculate  $p(\mathbf{y})$ ).
- **2** We can approximate the posterior by minimizing  $KL(q(\mathbf{z})||p(\mathbf{z}|\mathbf{y}))$ , which results in maximizing  $\mathcal{L}(q)$ .
- 3 The E-step is an iterative process.



#### **Variational EM - Remark**

$$\mathrm{KL}(q(\mathbf{z})||p(\mathbf{z}|\mathbf{y})) = \int q(\mathbf{z}) \ln \frac{q(\mathbf{z})}{p(\mathbf{z}|\mathbf{y})} \mathsf{d}\mathbf{z}$$

$$\mathrm{KL}(q(\mathbf{z})||p(\mathbf{z}|\mathbf{y})) = \int q(\mathbf{z}) \ln q(\mathbf{z}) d\mathbf{z} - \int q(\mathbf{z}) \ln p(\mathbf{z}, \mathbf{y}) d\mathbf{z} + \ln \frac{p(\mathbf{y})}{p(\mathbf{z})}$$

$$\mathrm{KL}(q(\mathbf{z})||p(\mathbf{z}|\mathbf{y})) = -\int q(\mathbf{z}) \ln p(\mathbf{y}|\mathbf{z}) \mathsf{d}\mathbf{z} + \int q(\mathbf{z}) \ln \frac{q(\mathbf{z})}{p(\mathbf{z})} \mathsf{d}\mathbf{z} + \ln \frac{p(\mathbf{y})}{p(\mathbf{z})} \mathsf{d}\mathbf{z}$$



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# **Approx. NNs - Problem definition**

Let's assume we have a dataset  $\mathcal{D}=\{\mathbf{X},\mathbf{Y}\}$ , with  $\mathbf{X}=\{\mathbf{x}_1,\ldots,\mathbf{x}_N\}$  and  $\mathbf{Y}=\{\mathbf{y}_1,\ldots,\mathbf{y}_N\}$ 

$$KL(q_{\theta}(\boldsymbol{\omega}) \| p(\boldsymbol{\omega} | \mathcal{D})) \propto -\int q_{\theta}(\boldsymbol{\omega}) \log p(\mathbf{Y} | \mathbf{X}, \boldsymbol{\omega}) d\boldsymbol{\omega} + KL(q_{\theta}(\boldsymbol{\omega}) \| p(\boldsymbol{\omega})),$$

$$= -\sum_{i=1}^{N} \int q_{\theta}(\boldsymbol{\omega}) \log p(\mathbf{y}_{i} | \mathbf{f}^{\boldsymbol{\omega}}(\mathbf{x}_{i})) d\boldsymbol{\omega} + KL(q_{\theta}(\boldsymbol{\omega}) \| p(\boldsymbol{\omega})),$$

where  $\omega$  is a vector composed by the stacking of all weights.



# **Approx. NNs - Problem definition**

- The summed-over terms log-likelihood are not tractable for BNNs with more than a single hidden layer.
- **2** This objective requires us to perform computations over the entire dataset, which can be too costly for large N.



## Approx. NNs - Data size solution

We can reduce the data size problem by data sub-sampling (also referred to as mini-batch optimization).

$$\widehat{\mathcal{L}}_{VI}(\theta) := -\frac{N}{M} \sum_{i \in S} \int q_{\theta}(\boldsymbol{\omega}) \log p\left(\mathbf{y}_{i} | \mathbf{f}^{\boldsymbol{\omega}}\left(\mathbf{x}_{i}\right)\right) d\boldsymbol{\omega} + KL\left(q_{\theta}(\boldsymbol{\omega}) \| p(\boldsymbol{\omega})\right)$$

with a random index set S of size M.



## Approx. NNs - Expected value solution

We can reduce the data size problem by data sub-sampling (also referred to as mini-batch optimization).

$$\mathbb{E}_{q_{\theta}(\boldsymbol{\omega})}[\log p\left(\mathbf{y}_{i}|\mathbf{f}^{\boldsymbol{\omega}}\left(\mathbf{x}_{i}\right)\right)] = \int q_{\theta}(\boldsymbol{\omega})\log p\left(\mathbf{y}_{i}|\mathbf{f}^{\boldsymbol{\omega}}\left(\mathbf{x}_{i}\right)\right)d\boldsymbol{\omega}$$

$$q_{\theta}(\boldsymbol{\omega}) = \prod_{i=1}^{L} q_{\theta}\left(\mathbf{W}_{i}\right) = \prod_{i=1}^{L} \prod_{j=1}^{K_{i}} \prod_{k=1}^{K_{i+1}} q\left(w_{ijk}\right) = \prod_{i,j,k} \mathcal{N}\left(w_{ijk}; m_{ijk}, \sigma_{ijk}^{2}\right)$$

$$\mathbb{E}_{q_{\theta}(\boldsymbol{\omega})}[\log p\left(\mathbf{y}_{i}|\mathbf{f}^{\boldsymbol{\omega}}\left(\mathbf{x}_{i}\right)\right)] \approx \frac{1}{M} \sum_{m=1}^{M} q_{\theta}(\boldsymbol{\omega}_{j})\log p\left(\mathbf{y}_{i}|\mathbf{f}^{\boldsymbol{\omega}}\left(\mathbf{x}_{i}\right)\right)$$



## Approx. NNs - Expected value solution

A better way is to use Pathwise Gradient Estimator

$$\mathbb{E}_{q_{\theta}(\boldsymbol{\omega})} [f_{\theta}(\boldsymbol{\omega})] \longrightarrow \mathbb{E}_{q_{0}(\boldsymbol{\epsilon})} [f_{\theta}(g(\boldsymbol{\theta}; \boldsymbol{\epsilon}))]$$
$$q(w_{ijk}) = g(\theta_{ijk}, \epsilon_{ijk}) = m_{ijk} + \sigma_{ijk} \epsilon_{ijk}$$

then

$$\widehat{\mathcal{L}}_{VI}(\theta) := -\frac{N}{M} \sum_{i \in S} \int q_{\theta}(\boldsymbol{\omega}) \log p\left(\mathbf{y}_{i} | \mathbf{f}^{\boldsymbol{\omega}}\left(\mathbf{x}_{i}\right)\right) d\boldsymbol{\omega} + KL\left(q_{\theta}(\boldsymbol{\omega}) \| p(\boldsymbol{\omega})\right)$$

becomes

$$\widehat{\mathcal{L}}_{\mathrm{MC}}(\theta) = -\frac{N}{M} \sum_{i \in S} \log p \left( \mathbf{y}_i | \mathbf{f}^{g(\theta, \epsilon)} \left( \mathbf{x}_i \right) \right) + \mathrm{KL} \left( q_{\theta}(\boldsymbol{\omega}) || p(\boldsymbol{\omega}) \right)$$

where 
$$\mathbb{E}_{S,\epsilon}\left[\hat{\mathcal{L}}_{\mathrm{MC}}(\theta)
ight] = \mathcal{L}_{\mathrm{VI}}(\theta).$$



#### **Algorithm 1** Minimise divergence between $q_{\theta}(\boldsymbol{\omega})$ and $p(\boldsymbol{\omega}|X,Y)$

- 1: Given dataset X, Y,
- 2: Define learning rate schedule  $\eta$ ,
- 3: Initialise parameters  $\theta$  randomly.
- 4: repeat
- 5: Sample M random variables  $\hat{\epsilon}_i \sim p(\epsilon)$ , S a random subset of  $\{1,...,N\}$  of size M.
- 6: Calculate stochastic derivative estimator w.r.t.  $\theta$ :

$$\widehat{\Delta\theta} \leftarrow -\frac{N}{M} \sum_{i \in S} \frac{\partial}{\partial \theta} \log p(\mathbf{y}_i | \mathbf{f}^{g(\theta, \widehat{\epsilon}_i)}(\mathbf{x}_i)) + \frac{\partial}{\partial \theta} \mathrm{KL}(q_{\theta}(\boldsymbol{\omega}) | | p(\boldsymbol{\omega})).$$

7: Update  $\theta$ :

$$\theta \leftarrow \theta + \eta \widehat{\Delta \theta}$$
.

8: **until**  $\theta$  has converged.



# **Approx. NNs - Dropout**

$$\begin{split} \widehat{\mathbf{y}} &= \widehat{\mathbf{h}} \mathbf{M}_2 \\ &= (\mathbf{h} \odot \widehat{\boldsymbol{\epsilon}}_2) \mathbf{M}_2 \\ &= (\mathbf{h} \cdot \operatorname{diag}(\widehat{\boldsymbol{\epsilon}}_2)) \mathbf{M}_2 \\ &= \mathbf{h}(\operatorname{diag}(\widehat{\boldsymbol{\epsilon}}_2) \mathbf{M}_2) \\ &= \sigma \Big( \widehat{\mathbf{x}} \mathbf{M}_1 + \mathbf{b} \Big) (\operatorname{diag}(\widehat{\boldsymbol{\epsilon}}_2) \mathbf{M}_2) \\ &= \sigma \Big( (\mathbf{x} \odot \widehat{\boldsymbol{\epsilon}}_1) \mathbf{M}_1 + \mathbf{b} \Big) (\operatorname{diag}(\widehat{\boldsymbol{\epsilon}}_2) \mathbf{M}_2) \\ &= \sigma \Big( \mathbf{x} (\operatorname{diag}(\widehat{\boldsymbol{\epsilon}}_1) \mathbf{M}_1) + \mathbf{b} \Big) (\operatorname{diag}(\widehat{\boldsymbol{\epsilon}}_2) \mathbf{M}_2) \end{split}$$



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# Approx. NNs - Laplace approximation

The standard Laplace approximation is obtained by taking the second-order Taylor expansion around a mode of a distribution.

If we approximate the log posterior over the weights of a network given some data  $\mathcal{D}$  around a MAP estimate  $\omega^*$ , we obtain

$$\log p(\boldsymbol{\omega}|\mathcal{D}) \approx \log p(\boldsymbol{\omega}^*|\mathcal{D}) - \frac{1}{2} (\boldsymbol{\omega} - \boldsymbol{\omega}^*)^{\top} \overline{H} (\boldsymbol{\omega} - \boldsymbol{\omega}^*),$$

where  $\overline{H} = \mathbb{E}[H]$  is the average Hessian of the negative log posterior. The posterior over the weights is then approximated as Gaussian:

$$oldsymbol{\omega} \sim \mathcal{N}\left(oldsymbol{\omega}^*, \overline{H}^{-1}
ight)$$



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