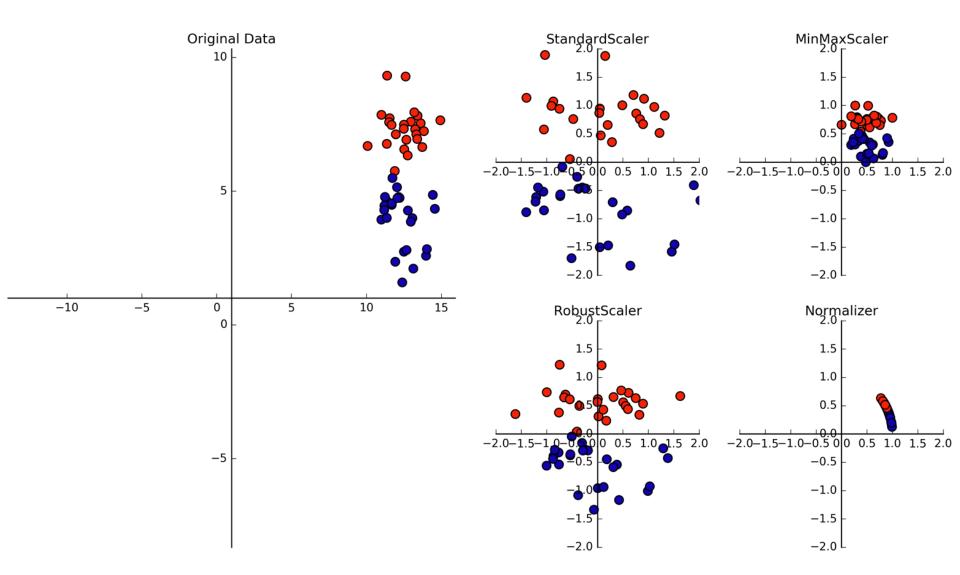
DIMENSIONALITY REDUCTION

Machine Learning Problems

_	Supervised Learning	Unsupervised Learning
Discrete	classification or categorization	clustering
Continuous	regression	dimensionality reduction



Why Reduce Dimensionality?

- Reduces time complexity: Less computation
- Reduces space complexity: Fewer parameters
- Saves the cost of observing the feature
- Simpler models are more robust on small datasets
- More interpretable; simpler explanation
- Data visualization (structure, groups, outliers, etc) if plotted in 2 or 3 dimensions (just have to)

Feature Selection vs Extraction

- Feature selection: Choosing k<d important features, ignoring the remaining d k
 Subset selection algorithms
- □ Feature extraction: Project the original x_i , i = 1,...,d dimensions to new k < d dimensions, z_i , j = 1,...,k



PRINCIPAL

COMPONENT ANALYSIS

PCA projects the features onto the principal components. The motivation is to reduce the features dimensionality while only losing a small amount of information.

Second principal component First principal component

ChrisAlbon

Principal Components Analysis

- □ Find a low-dimensional space such that when **x** is projected there, information loss is minimized.
- \square The projection of x on the direction of w is: $z = w^T x$
- \square Find w such that Var(z) is maximized

$$Var(z) = Var(\mathbf{w}^{\mathsf{T}}\mathbf{x}) = E[(\mathbf{w}^{\mathsf{T}}\mathbf{x} - \mathbf{w}^{\mathsf{T}}\boldsymbol{\mu})^{2}]$$

$$= E[(\mathbf{w}^{\mathsf{T}}\mathbf{x} - \mathbf{w}^{\mathsf{T}}\boldsymbol{\mu})(\mathbf{w}^{\mathsf{T}}\mathbf{x} - \mathbf{w}^{\mathsf{T}}\boldsymbol{\mu})]$$

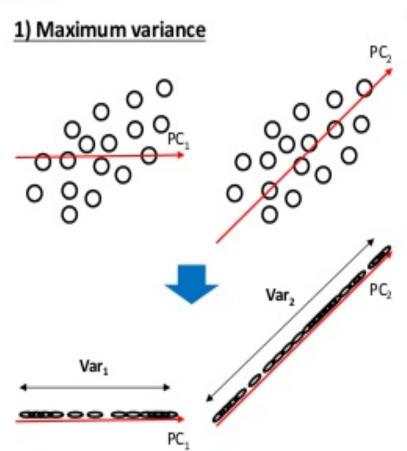
$$= E[\mathbf{w}^{\mathsf{T}}(\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^{\mathsf{T}}\mathbf{w}]$$

$$= \mathbf{w}^{\mathsf{T}} E[(\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^{\mathsf{T}}]\mathbf{w} = \mathbf{w}^{\mathsf{T}} \sum \mathbf{w}$$
where $Var(\mathbf{x}) = E[(\mathbf{x} - \boldsymbol{\mu})(\mathbf{x} - \boldsymbol{\mu})^{\mathsf{T}}] = \sum$

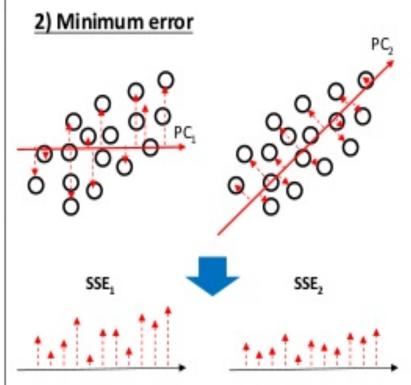


Principles of PCA



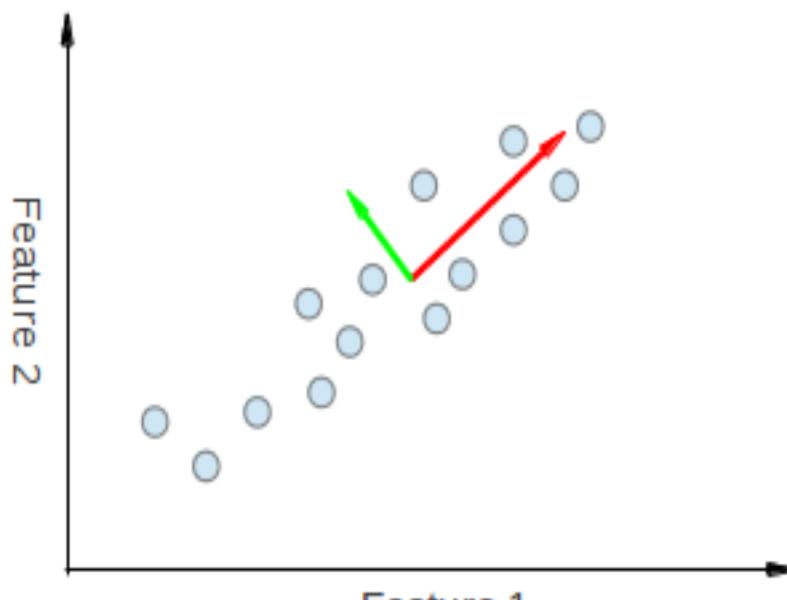


 To maximize the variance of the projected data on the certain dimension.

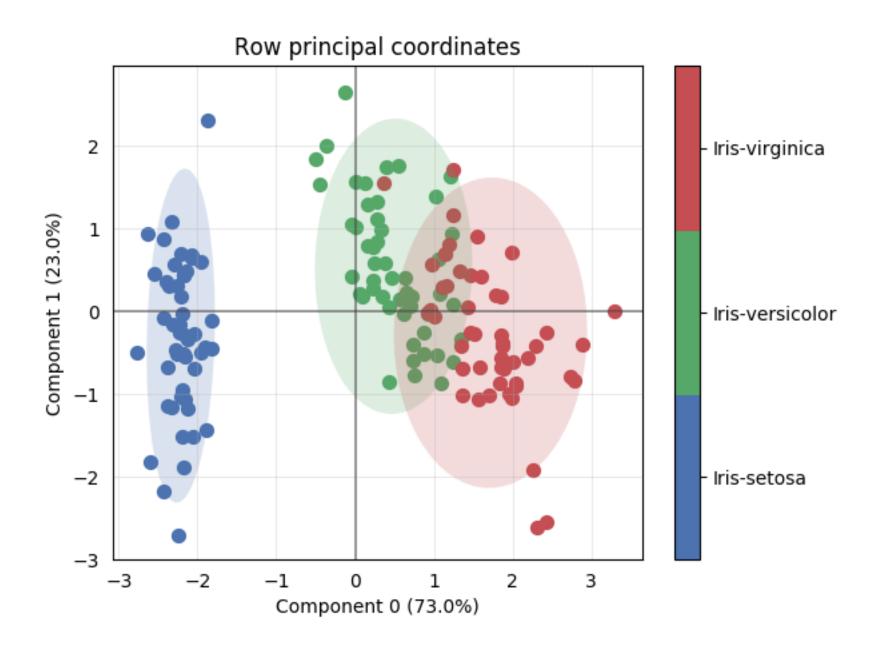


 To minimize the mean squared distance between the data and their projections.

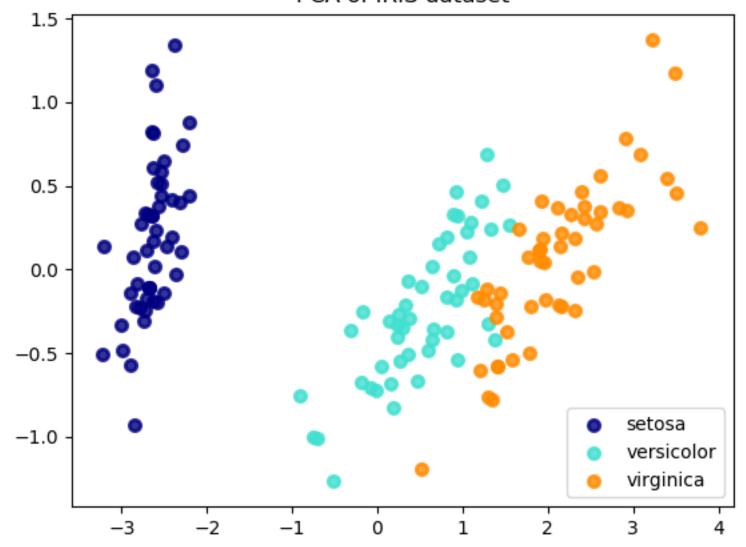
SSE: Sum or squared errors



Feature 1



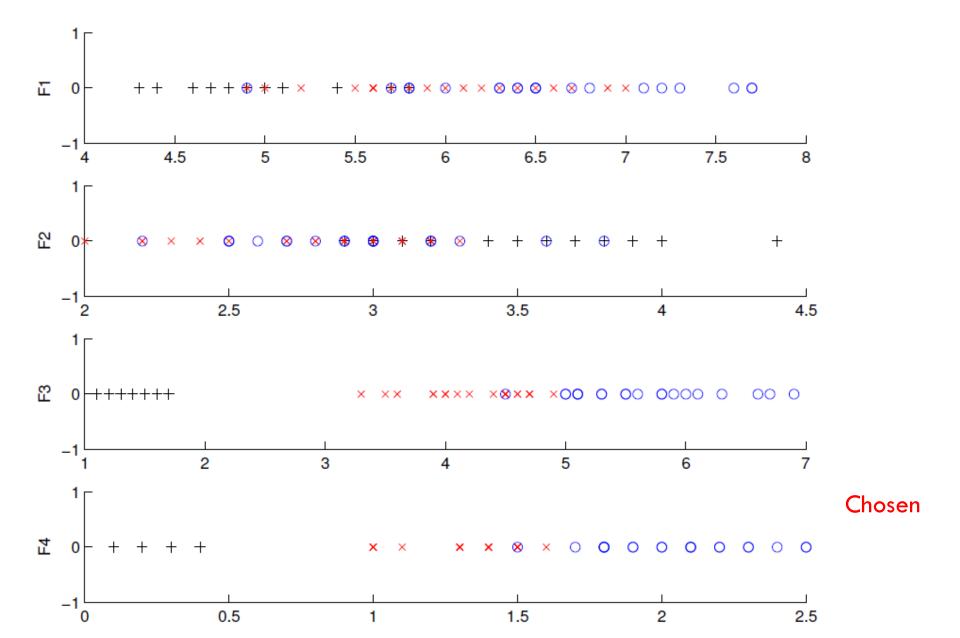
PCA of IRIS dataset



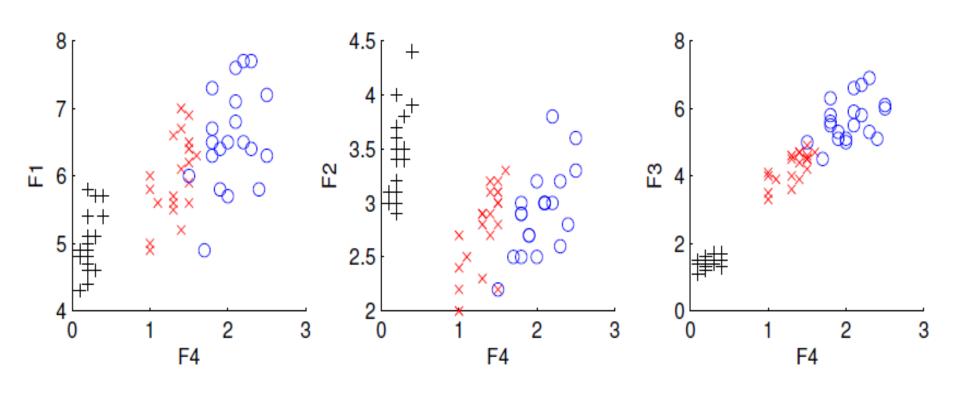
Subset Selection

- \Box There are 2^d subsets of d features
- Forward search: Add the best feature at each step
 - \blacksquare Set of features F initially \emptyset .
 - At each iteration, find the best new feature $j = \operatorname{argmin}_i E(F \cup x_i)$
 - Add x_i to F if $E(F \cup x_i) < E(F)$
- □ Hill-climbing $O(d^2)$ algorithm
- Backward search: Start with all features and remove one at a time, if possible.
- □ Floating search (Add k, remove I)

Iris data: Single feature



Iris data: Add one more feature to F4



Chosen

14

□ Maximize Var(z) subject to ||w|| = 1

$$\max_{\mathbf{w}_1} \mathbf{w}_1^T \mathbf{\Sigma} \mathbf{w}_1 - \alpha (\mathbf{w}_1^T \mathbf{w}_1 - 1)$$

 $\sum w_1 = \alpha w_1$ that is, w_1 is an eigenvector of \sum Choose the one with the largest eigenvalue for Var(z) to be max

□ Second principal component: Max $Var(z_2)$, s.t., $||w_2||=1$ and orthogonal to w_1

$$\max_{\mathbf{w}_2} \mathbf{w}_2^\mathsf{T} \mathbf{\Sigma} \mathbf{w}_2 - \alpha (\mathbf{w}_2^\mathsf{T} \mathbf{w}_2 - 1) - \beta (\mathbf{w}_2^\mathsf{T} \mathbf{w}_1 - 0)$$

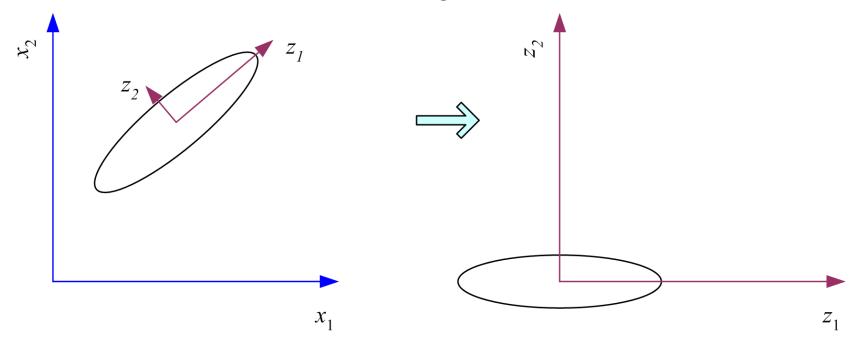
 $\sum w_2 = \alpha w_2$ that is, w_2 is another eigenvector of \sum and so on.

What PCA does

$$z = \mathbf{W}^{\mathsf{T}}(\mathbf{x} - \mathbf{m})$$

where the columns of \mathbf{W} are the eigenvectors of \sum and \mathbf{m} is sample mean

Centers the data at the origin and rotates the axes



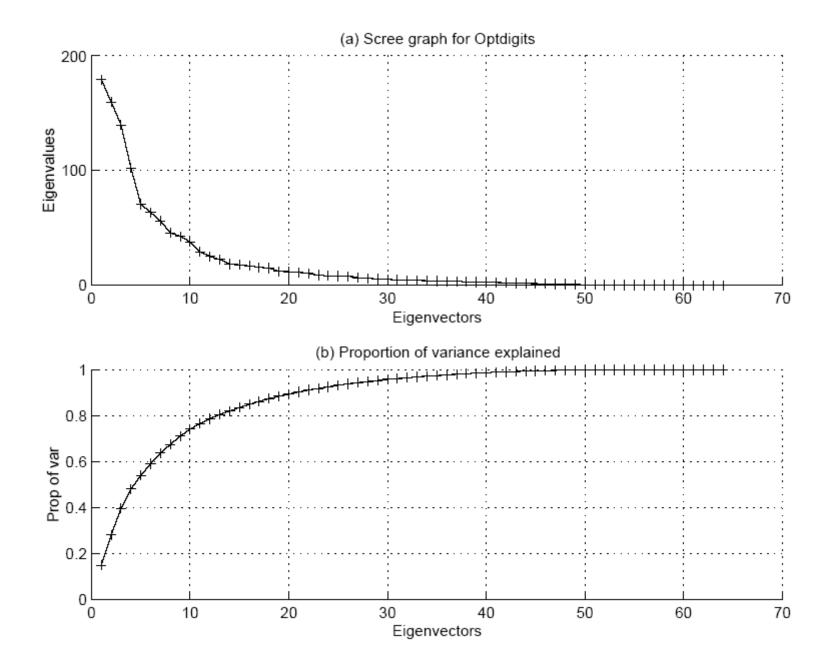
How to choose k?

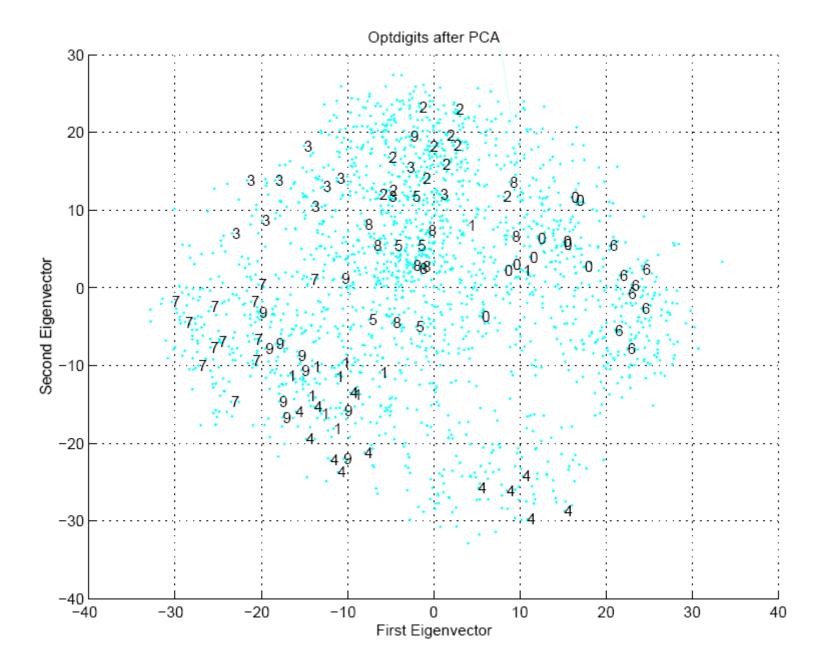
Proportion of Variance (PoV) explained

$$\frac{\lambda_1 + \lambda_2 + \dots + \lambda_k}{\lambda_1 + \lambda_2 + \dots + \lambda_k + \dots + \lambda_d}$$

when λ_i are sorted in descending order

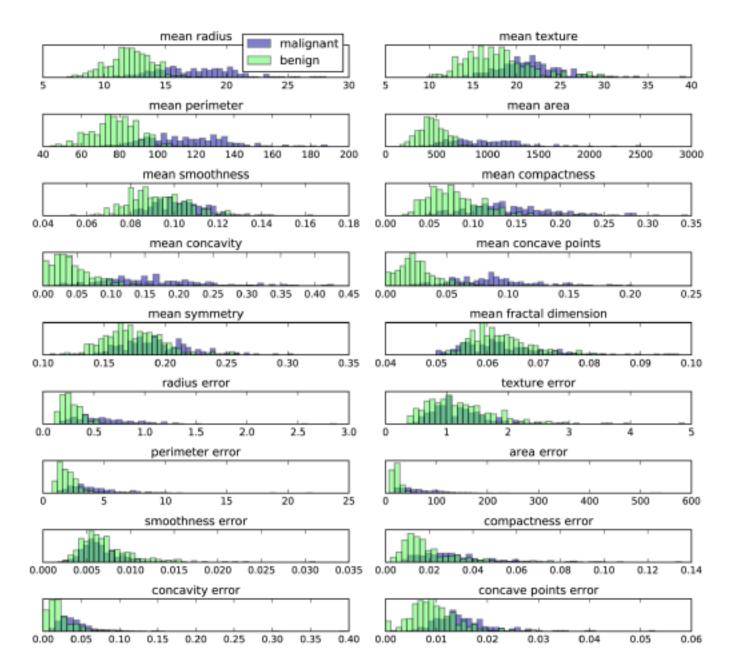
- □ Typically, stop at PoV>0.9
- □ Scree graph plots of PoV vs k, stop at "elbow"

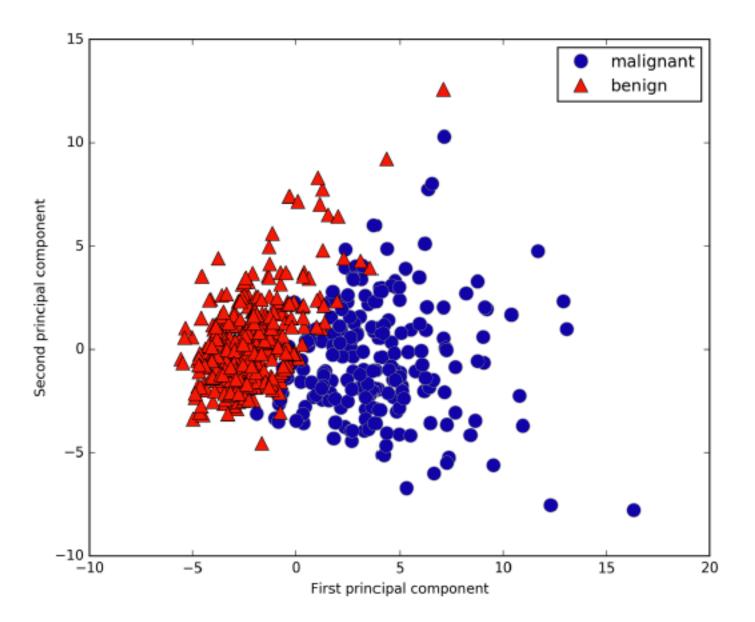


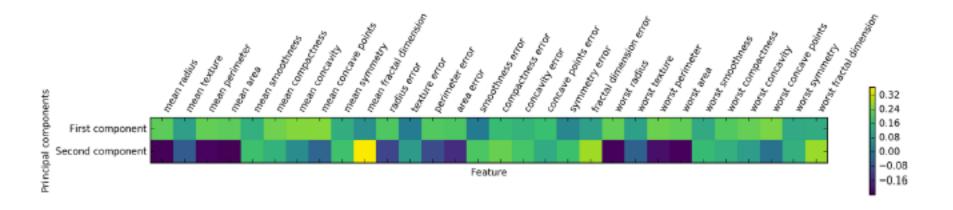


Feature Embedding

- \square When **X** is the Nxd data matrix,
- $\mathbf{X}^T\mathbf{X}$ is the dxd matrix (covariance of features, if mean-centered)
- XX^T is the NxN matrix (pairwise similarities of instances)
- \square PCA uses the eigenvectors of $\mathbf{X}^T\mathbf{X}$ which are d-dim and can be used for projection
- □ Feature embedding uses the eigenvectors of XX^T which are N-dim and which give directly the coordinates after projection
- Sometimes, we can define pairwise similarities (or distances) between instances, then we can use feature embedding without needing to represent instances as vectors.







Factor Analysis

□ Find a small number of factors z, which when combined generate x :

$$x_i - \mu_i = v_{i1}z_1 + v_{i2}z_2 + ... + v_{ik}z_k + \varepsilon_i$$

where z_i , i = 1,...,k are the latent factors with $E[z_i]=0$, $Var(z_i)=1$, $Cov(z_i, z_i)=0$, $i \neq j$,

 ε_i are the noise sources

E[ε_i]= ψ_i , Cov(ε_i , ε_j) =0, $i \neq j$, Cov(ε_i , z_j) =0 , and v_{ij} are the factor loadings

PCA vs FA

PCA

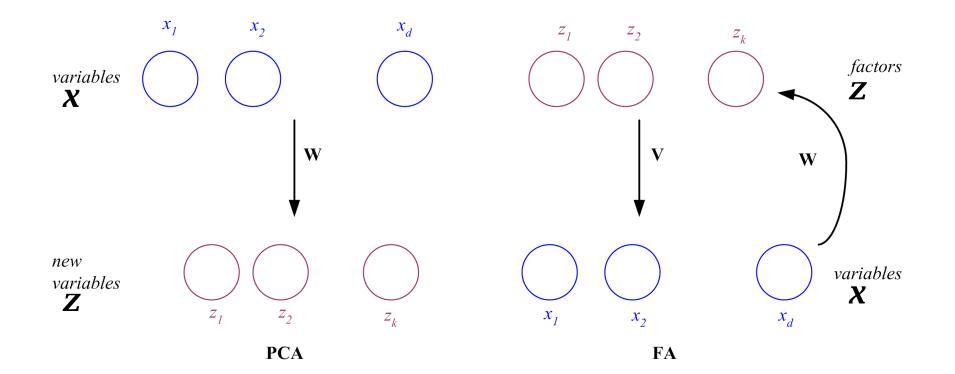
From x to z

 $z = W^T(x - \mu)$

□ FA

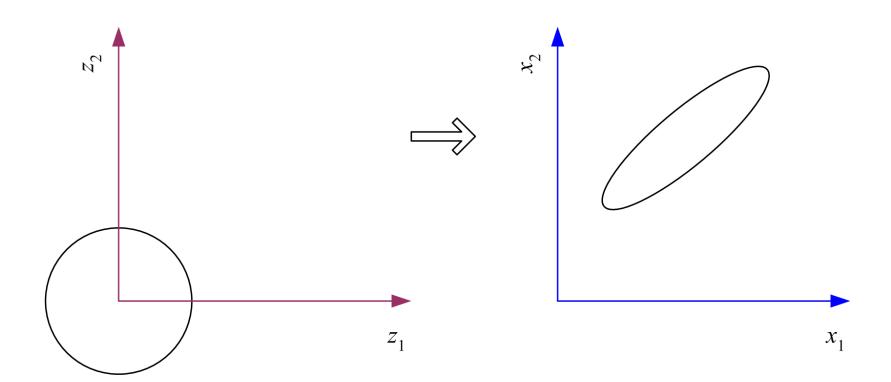
From z to x

 $x - \mu = Vz + \varepsilon$



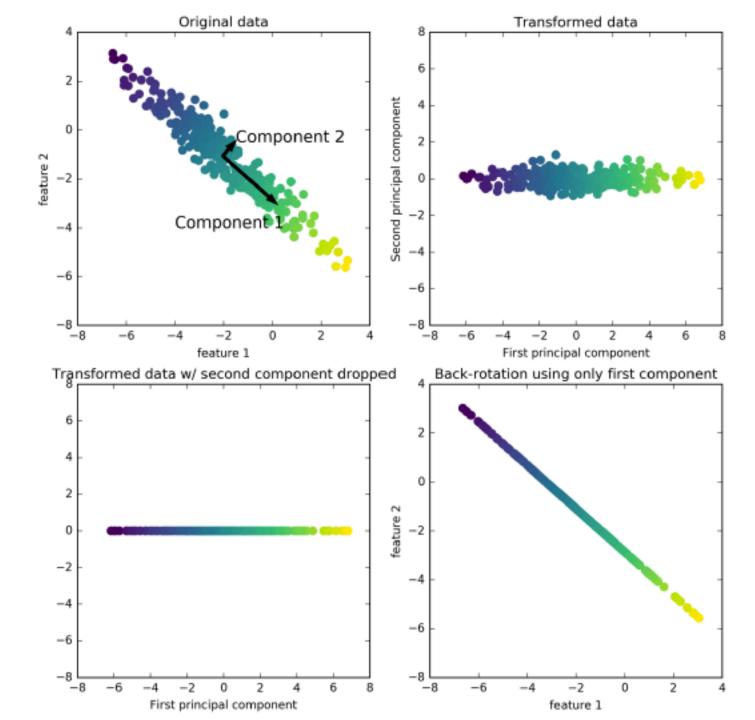
Factor Analysis

□ In FA, factors z_i are stretched, rotated and translated to generate x



Singular Value Decomposition and Matrix Factorization

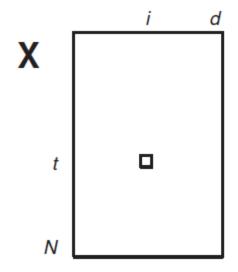
- \square Singular value decomposition: $X = VAW^T$
 - V is NxN and contains the eigenvectors of XX^T W is dxd and contains the eigenvectors of X^TX and A is Nxd and contains singular values on its first k diagonal
- $\square X = \mathbf{u}_1 \mathbf{a}_1 \mathbf{v}_1^T + ... + \mathbf{u}_k \mathbf{a}_k \mathbf{v}_k^T$ where k is the rank of X

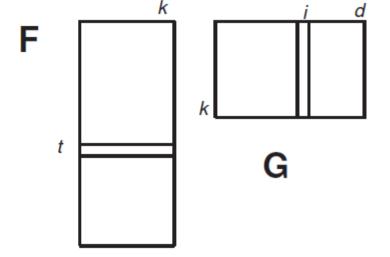


Matrix Factorization

■ Matrix factorization: X=FG

F is Nxk and G is kxd

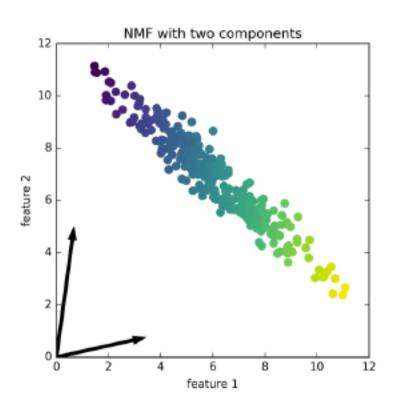


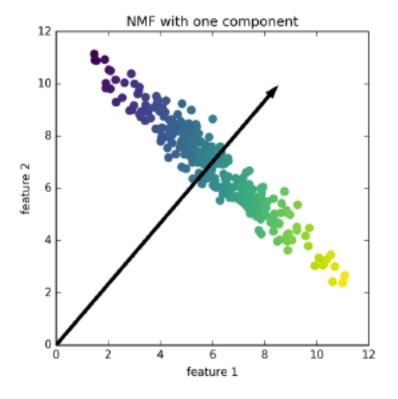


$$\mathbf{X}_{ti} = \mathbf{F}_t^T \mathbf{G}_i = \sum_{j=1}^k \mathbf{F}_{tj} \mathbf{G}_{ji}$$

Latent semantic indexing

Non-negative Matrix Factorization





Multidimensional Scaling (MDS)

□ Given pairwise distances between N points,

$$d_{ij}$$
, $i,j = 1,...,N$

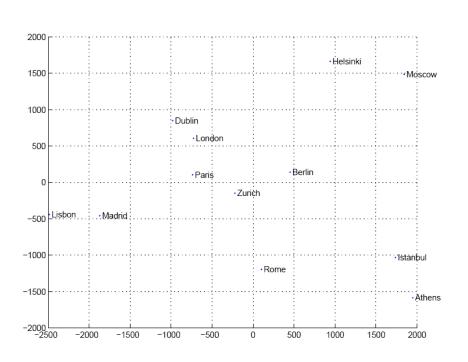
place on a low-dim map s.t. distances are preserved (by feature embedding)

 $\mathbf{z} = \mathbf{g}(\mathbf{x} \mid \vartheta)$ Find ϑ that min Sammon stress

$$E(\theta \mid \mathcal{X}) = \sum_{r,s} \frac{\left\| \mathbf{z}^r - \mathbf{z}^s \right\| - \left\| \mathbf{x}^r - \mathbf{x}^s \right\|^2}{\left\| \mathbf{x}^r - \mathbf{x}^s \right\|^2}$$

$$= \sum_{r,s} \frac{\left\| \mathbf{g}(\mathbf{x}^r \mid \theta) - \mathbf{g}(\mathbf{x}^s \mid \theta) \right\| - \left\| \mathbf{x}^r - \mathbf{x}^s \right\|^2}{\left\| \mathbf{x}^r - \mathbf{x}^s \right\|^2}$$

Map of Europe by MDS



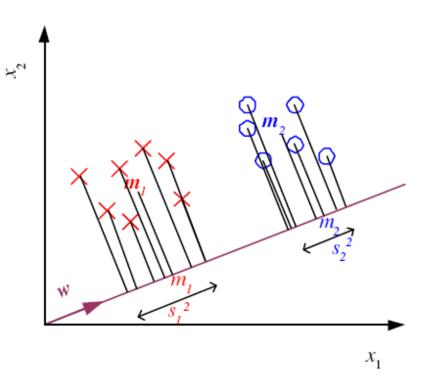


Map from CIA – The World Factbook: http://www.cia.gov/

Linear Discriminant Analysis (LDA)

- Find a low-dimensional space such that when x is projected, classes are well-separated.
- Find w that maximizes

$$J(\mathbf{w}) = \frac{(m_1 - m_2)^2}{s_1^2 + s_2^2}$$



$$m_1 = \frac{\sum_t \mathbf{w}^T \mathbf{x}^t r^t}{\sum_t r^t} \quad s_1^2 = \sum_t (\mathbf{w}^T \mathbf{x}^t - m_1)^2 r^t$$

Between-class scatter:

$$(\mathbf{m}_1 - \mathbf{m}_2)^2 = (\mathbf{w}^T \mathbf{m}_1 - \mathbf{w}^T \mathbf{m}_2)^2$$

$$= \mathbf{w}^T (\mathbf{m}_1 - \mathbf{m}_2) (\mathbf{m}_1 - \mathbf{m}_2)^T \mathbf{w}$$

$$= \mathbf{w}^T \mathbf{S}_B \mathbf{w} \text{ where } \mathbf{S}_B = (\mathbf{m}_1 - \mathbf{m}_2) (\mathbf{m}_1 - \mathbf{m}_2)^T$$

Within-class scatter:

$$s_1^2 = \sum_t (\mathbf{w}^T \mathbf{x}^t - m_1)^2 r^t$$

$$= \sum_t \mathbf{w}^T (\mathbf{x}^t - \mathbf{m}_1) (\mathbf{x}^t - \mathbf{m}_1)^T \mathbf{w} r^t = \mathbf{w}^T \mathbf{S}_1 \mathbf{w}$$
where $\mathbf{S}_1 = \sum_t (\mathbf{x}^t - \mathbf{m}_1) (\mathbf{x}^t - \mathbf{m}_1)^T r^t$

$$s_1^2 + s_1^2 = \mathbf{w}^T \mathbf{S}_W \mathbf{w} \text{ where } \mathbf{S}_W = \mathbf{S}_1 + \mathbf{S}_2$$

Fisher's Linear Discriminant (LDA)

Find w that max

$$J(\mathbf{w}) = \frac{\mathbf{w}^T \mathbf{S}_B \mathbf{w}}{\mathbf{w}^T \mathbf{S}_W \mathbf{w}} = \frac{\left| \mathbf{w}^T (\mathbf{m}_1 - \mathbf{m}_2) \right|^2}{\mathbf{w}^T \mathbf{S}_W \mathbf{w}}$$

□ LDA soln:

$$\mathbf{w} = \mathbf{c} \cdot \mathbf{S}_{W}^{-1} (\mathbf{m}_{1} - \mathbf{m}_{2})$$

Parametric soln:

$$\mathbf{w} = \Sigma^{-1}(\mu_1 - \mu_2)$$
when $p(\mathbf{x} \mid C_i) \sim \mathcal{N}(\mu_i, \Sigma)$

K>2 Classes

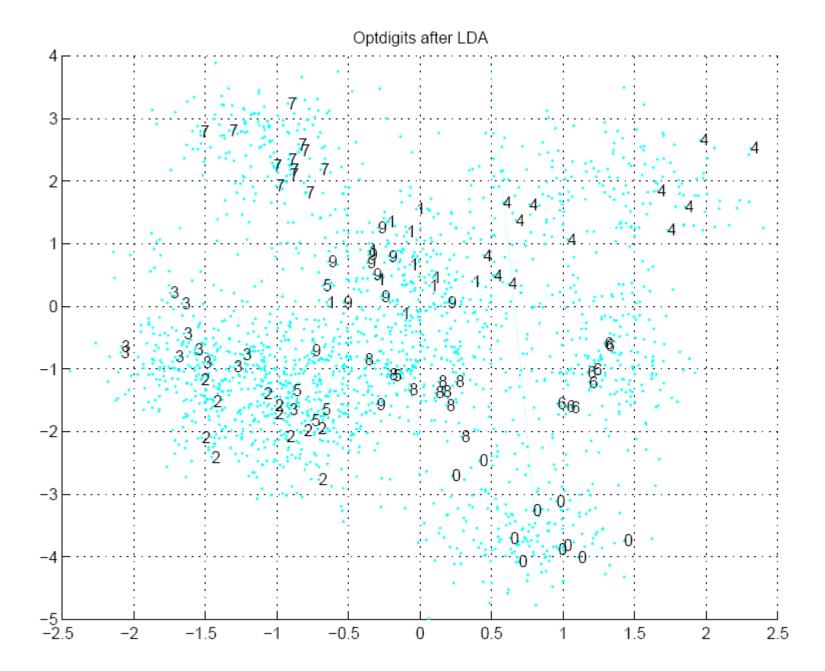
Within-class scatter:

$$\mathbf{S}_{W} = \sum_{i=1}^{K} \mathbf{S}_{i} \qquad \mathbf{S}_{i} = \sum_{t} r_{i}^{t} \left(\mathbf{x}^{t} - \mathbf{m}_{i} \right) \left(\mathbf{x}^{t} - \mathbf{m}_{i} \right)^{T}$$

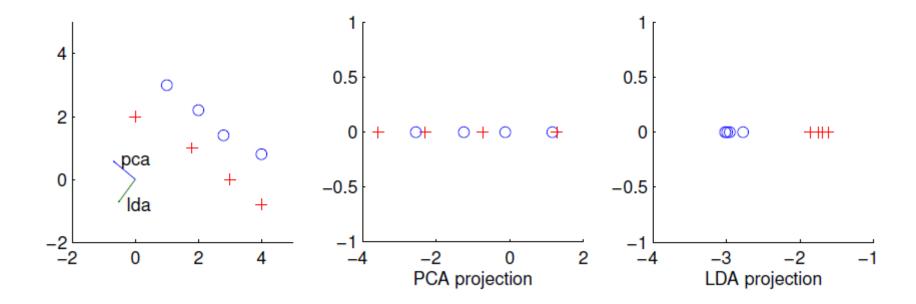
Between-class scatter:

$$\mathbf{S}_{B} = \sum_{i=1}^{K} N_{i} (\mathbf{m}_{i} - \mathbf{m}) (\mathbf{m}_{i} - \mathbf{m})^{T} \qquad \mathbf{m} = \frac{1}{K} \sum_{i=1}^{K} \mathbf{m}_{i}$$

The largest eigenvectors of $S_{W}^{-1}S_{B_{i}}$ maximum rank of K-1



PCA vs LDA



Canonical Correlation Analysis

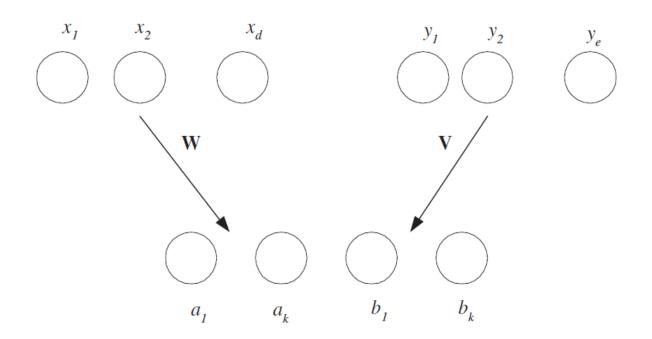
- $\square X = \{x^t, y^t\}_t$; two sets of variables x and y x
- We want to find two projections w and v st when x is projected along w and y is projected along v, the correlation is maximized:

$$\rho = \operatorname{Corr}(\boldsymbol{w}^{T}\boldsymbol{x}, \boldsymbol{v}^{T}\boldsymbol{y}) = \frac{\operatorname{Cov}(\boldsymbol{w}^{T}\boldsymbol{x}, \boldsymbol{v}^{T}\boldsymbol{y})}{\sqrt{\operatorname{Var}(\boldsymbol{w}^{T}\boldsymbol{x})}} \sqrt{\operatorname{Var}(\boldsymbol{v}^{T}\boldsymbol{y})}$$

$$= \frac{\boldsymbol{w}^{T}\operatorname{Cov}(\boldsymbol{x}, \boldsymbol{y})\boldsymbol{v}}{\sqrt{\boldsymbol{w}^{T}\operatorname{Var}(\boldsymbol{x})\boldsymbol{w}}\sqrt{\boldsymbol{v}^{T}\operatorname{Var}(\boldsymbol{y})\boldsymbol{v}}} = \frac{\boldsymbol{w}^{T}\operatorname{S}_{xy}\boldsymbol{v}}{\sqrt{\boldsymbol{w}^{T}\operatorname{S}_{xx}\boldsymbol{w}}\sqrt{\boldsymbol{v}^{T}\operatorname{S}_{yy}\boldsymbol{v}}}$$

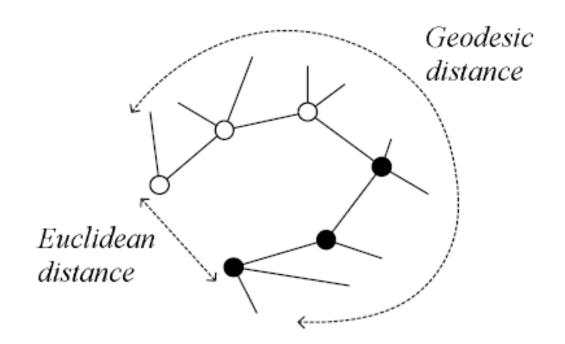
CCA

x and y may be two different views or modalities;
 e.g., image and word tags, and CCA does a joint mapping



Isomap

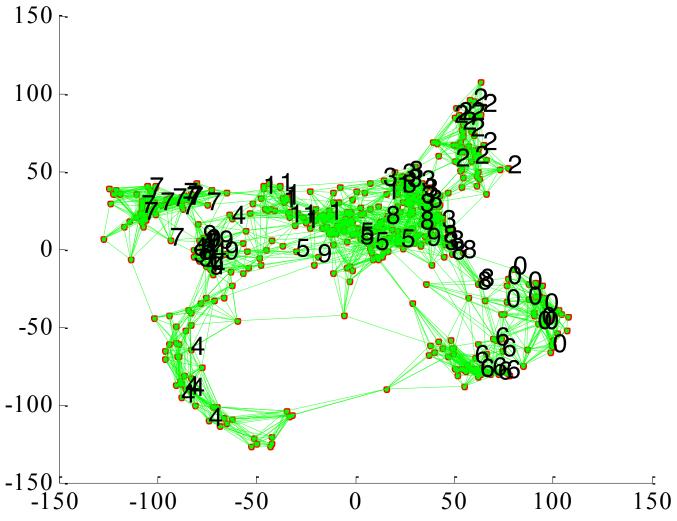
 Geodesic distance is the distance along the manifold that the data lies in, as opposed to the Euclidean distance in the input space



Isomap

- □ Instances r and s are connected in the graph if $||\mathbf{x}^r \mathbf{x}^s|| < \varepsilon$ or if \mathbf{x}^s is one of the k neighbors of \mathbf{x}^r . The edge length is $||\mathbf{x}^r \mathbf{x}^s||$
- For two nodes r and s not connected, the distance is equal to the shortest path between them
- Once the NxN distance matrix is thus formed, use
 MDS to find a lower-dimensional mapping

Optdigits after Isomap (with neighborhood graph).



 ${\tt Matlab\ source\ from\ http://web.mit.edu/cocosci/isomap/isomap.html}$

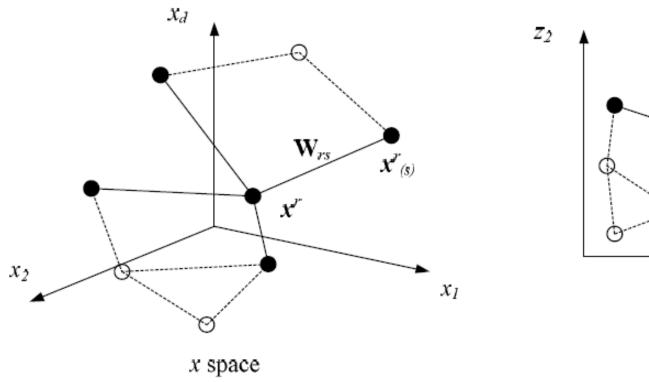
Locally Linear Embedding

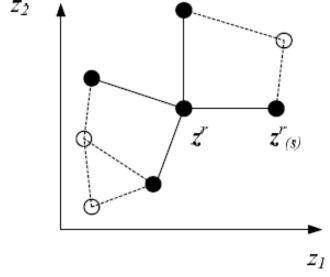
- 1. Given \mathbf{x}^r find its neighbors $\mathbf{x}^s_{(r)}$
- 2. Find W_{rs} that minimize

$$E(\mathbf{W} \mid X) = \sum_{r} \left\| \mathbf{x}^{r} - \sum_{s} \mathbf{W}_{rs} \mathbf{x}_{(r)}^{s} \right\|^{2}$$

3. Find the new coordinates z^r that minimize

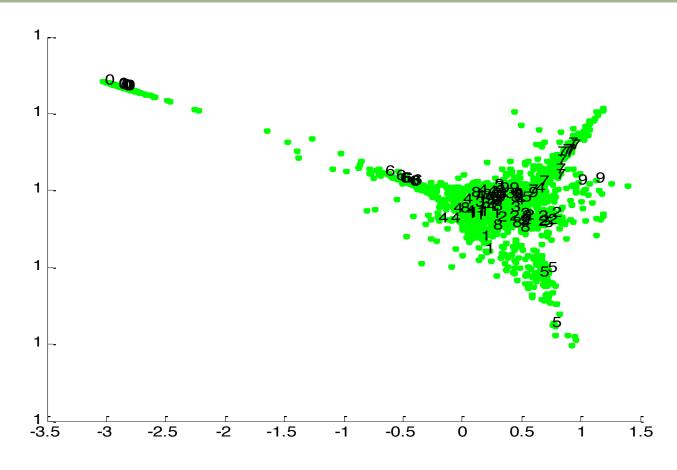
$$E(\mathbf{z} \mid \mathbf{W}) = \sum_{r} \left\| z^{r} - \sum_{s} \mathbf{W}_{rs} z_{(r)}^{s} \right\|^{2}$$





z space

LLE on Optdigits



 ${\tt Matlab\ source\ from\ http://www.cs.toronto.edu/}{\sim} roweis/Ile/code.html$

Laplacian Eigenmaps

□ Let r and s be two instances and B_{rs} is their similarity, we want to find \mathbf{z}^r and \mathbf{z}^s that

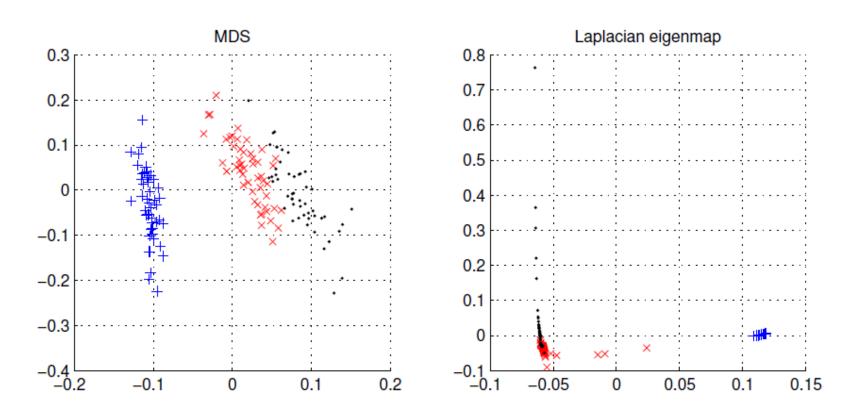
$$\min \sum_{r,s} \|\boldsymbol{z}^r - \boldsymbol{z}^s\|^2 B_{rs}$$

 B_{rs} can be defined in terms of similarity in an original space: 0 if x^r and x^s are too far, otherwise

$$B_{rs} = \exp\left[-\frac{\|\mathbf{x}^r - \mathbf{x}^s\|^2}{2\sigma^2}\right]$$

 Defines a graph Laplacian, and feature embedding returns z^r

Laplacian Eigenmaps on Iris



Spectral clustering (chapter 7)

CHAPTER 10:

LINEAR DISCRIMINATION

Likelihood- vs. Discriminant-based Classification

Likelihood-based: Assume a model for $p(x | C_i)$, use Bayes' rule to calculate $P(C_i | x)$

$$g_i(\mathbf{x}) = \log P(C_i | \mathbf{x})$$

- □ Discriminant-based: Assume a model for $g_i(\mathbf{x} \mid \Phi_i)$; no density estimation
- Estimating the boundaries is enough; no need to accurately estimate the densities inside the boundaries

Linear Discriminant

Linear discriminant:

$$g_i(\mathbf{x} \mid \mathbf{w}_i, \mathbf{w}_{i0}) = \mathbf{w}_i^T \mathbf{x} + \mathbf{w}_{i0} = \sum_{j=1}^a \mathbf{w}_{ij} \mathbf{x}_j + \mathbf{w}_{i0}$$

- Advantages:
 - Simple: O(d) space/computation
 - Knowledge extraction: Weighted sum of attributes; positive/negative weights, magnitudes (credit scoring)
 - Optimal when $p(x | C_i)$ are Gaussian with shared cov matrix; useful when classes are (almost) linearly separable

Generalized Linear Model

Quadratic discriminant:

$$g_i(\mathbf{x} \mid \mathbf{W}_i, \mathbf{w}_i, \mathbf{w}_{i0}) = \mathbf{x}^T \mathbf{W}_i \mathbf{x} + \mathbf{w}_i^T \mathbf{x} + \mathbf{w}_{i0}$$

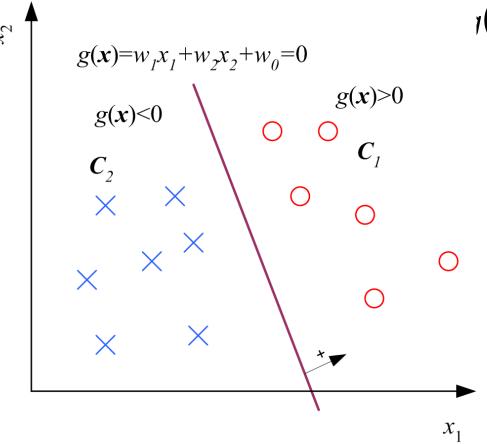
Higher-order (product) terms:

$$Z_1 = X_1$$
, $Z_2 = X_2$, $Z_3 = X_1^2$, $Z_4 = X_2^2$, $Z_5 = X_1X_2$

Map from x to z using nonlinear basis functions and use a linear discriminant in z-space

$$g_i(\mathbf{x}) = \sum_{j=1}^k w_{ij} \phi_j(\mathbf{x})$$

Two Classes



$$f(\mathbf{x}) = g_1(\mathbf{x}) - g_2(\mathbf{x})$$

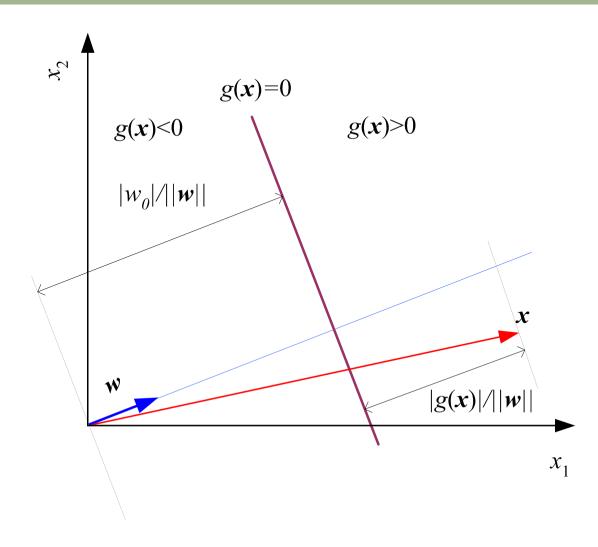
$$= (\mathbf{w}_1^T \mathbf{x} + \mathbf{w}_{10}) - (\mathbf{w}_2^T \mathbf{x} + \mathbf{w}_{20})$$

$$= (\mathbf{w}_1 - \mathbf{w}_2)^T \mathbf{x} + (\mathbf{w}_{10} - \mathbf{w}_{20})$$

$$= \mathbf{w}^T \mathbf{x} + \mathbf{w}_0$$

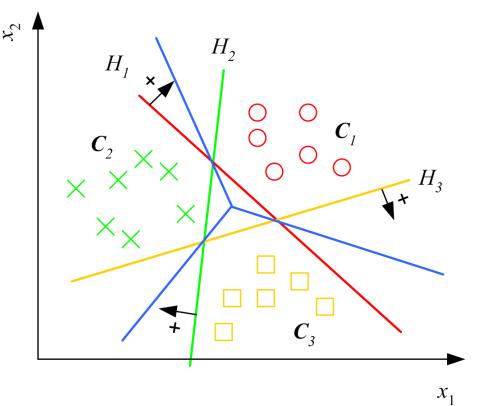
choose
$$\begin{cases} C_1 & \text{if } g(\mathbf{x}) > 0 \\ C_2 & \text{otherwise} \end{cases}$$

Geometry



Multiple Classes

$$g_i(\mathbf{x} \mid \mathbf{w}_i, \mathbf{w}_{i0}) = \mathbf{w}_i^T \mathbf{x} + \mathbf{w}_{i0}$$

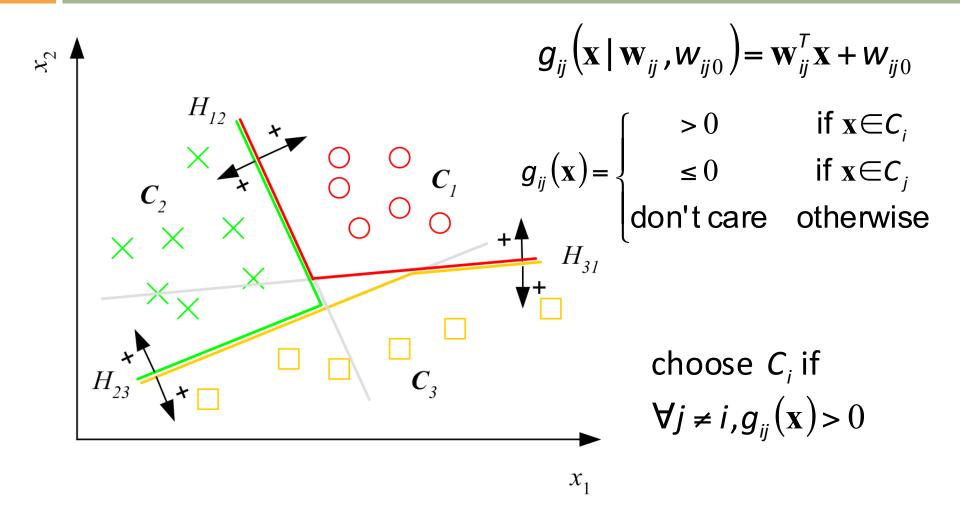


Choose C_i if

$$g_i(\mathbf{x}) = \max_{j=1}^K g_j(\mathbf{x})$$

Classes are linearly separable

Pairwise Separation



From Discriminants to Posteriors

When
$$p(\mathbf{x} \mid C_i) \sim N(\boldsymbol{\mu}_i, \boldsymbol{\Sigma})$$

$$g_i(\mathbf{x} \mid \mathbf{w}_i, \boldsymbol{w}_{i0}) = \mathbf{w}_i^T \mathbf{x} + \boldsymbol{w}_{i0}$$

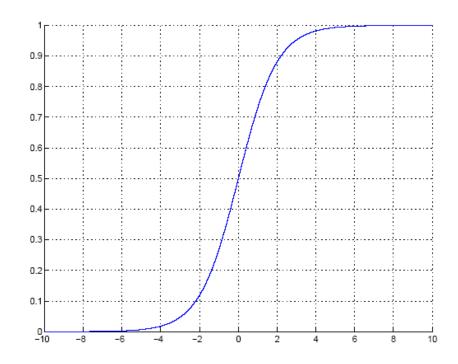
$$\mathbf{w}_i = \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_i \quad \boldsymbol{w}_{i0} = -\frac{1}{2} \boldsymbol{\mu}_i^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}_i + \log P(C_i)$$

$$y = P(C_1 \mid \mathbf{x}) \text{ and } P(C_2 \mid \mathbf{x}) = 1 - y$$

$$choose C_1 \text{ if } \begin{cases} y > 0.5 \\ y/(1-y) > 1 & \text{and } C_2 \text{ otherwise} \\ \log[y/(1-y)] > 0 \end{cases}$$

$$\begin{split} \log & \mathrm{idgit}(P(C_1 \mid \mathbf{x})) = \log \frac{P(C_1 \mid \mathbf{x})}{1 - P(C_1 \mid \mathbf{x})} = \log \frac{P(C_1 \mid \mathbf{x})}{P(C_2 \mid \mathbf{x})} \\ &= \log \frac{p(\mathbf{x} \mid C_1)}{p(\mathbf{x} \mid C_2)} + \log \frac{P(C_1)}{P(C_2)} \\ &= \log \frac{(2\pi)^{-d/2} |\Sigma|^{-1/2} \exp \left[-(1/2)(\mathbf{x} - \mu_1)^T \Sigma^{-1}(\mathbf{x} - \mu_1) \right]}{(2\pi)^{-d/2} |\Sigma|^{-1/2} \exp \left[-(1/2)(\mathbf{x} - \mu_2)^T \Sigma^{-1}(\mathbf{x} - \mu_2) \right]} + \log \frac{P(C_1)}{P(C_2)} \\ &= \mathbf{w}^T \mathbf{x} + \mathbf{w}_0 \\ &= \mathbf{w}^T \mathbf{x} + \mathbf{w}_0 \\ &\text{where } \mathbf{w} = \Sigma^{-1} (\mu_1 - \mu_2) \quad \mathbf{w}_0 = -\frac{1}{2} (\mu_1 + \mu_2)^T \Sigma^{-1} (\mu_1 - \mu_2) \\ &\qquad \qquad \text{The inverse of logit} \\ &\log \frac{P(C_1 \mid \mathbf{x})}{1 - P(C_1 \mid \mathbf{x})} = \mathbf{w}^T \mathbf{x} + \mathbf{w}_0 \\ &\qquad \qquad P(C_1 \mid \mathbf{x}) = \mathrm{sigmoid}(\mathbf{w}^T \mathbf{x} + \mathbf{w}_0) = \frac{1}{1 + \exp[-(\mathbf{w}^T \mathbf{x} + \mathbf{w}_0)]} \end{split}$$

Sigmoid (Logistic) Function



Calculate $g(\mathbf{x}) = \mathbf{w}^T \mathbf{x} + w_0$ and choose C_1 if $g(\mathbf{x}) > 0$, or Calculate $y = \text{sigmoid}(\mathbf{w}^T \mathbf{x} + w_0)$ and choose C_1 if y > 0.5

Gradient-Descent

□ $E(w \mid X)$ is error with parameters w on sample X $w^* = \arg \min_{w} E(w \mid X)$

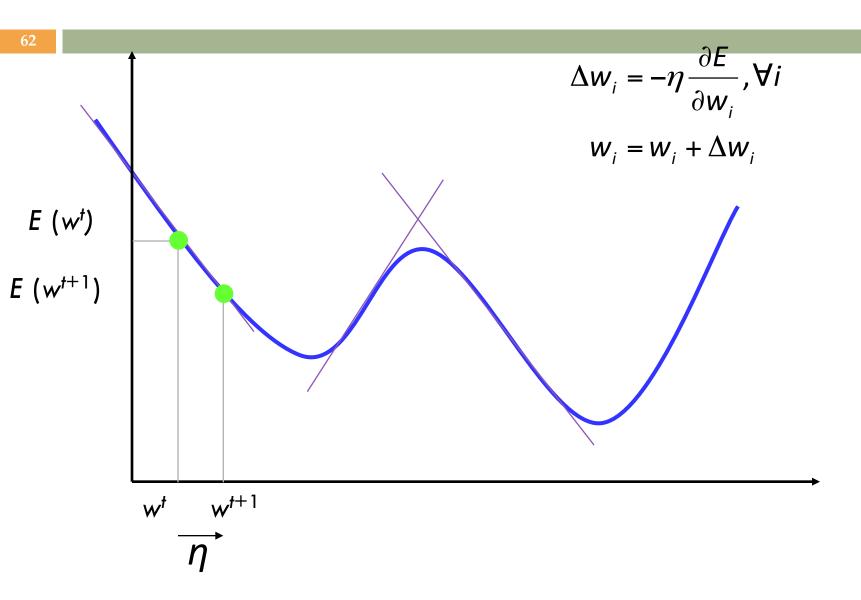
Gradient

$$\nabla_{w}E = \left[\frac{\partial E}{\partial w_{1}}, \frac{\partial E}{\partial w_{2}}, \dots, \frac{\partial E}{\partial w_{d}}\right]^{T}$$

Gradient-descent:

Starts from random w and updates w iteratively in the negative direction of gradient

Gradient-Descent



Logistic Discrimination

Two classes: Assume log likelihood ratio is linear

$$\log \frac{p(\mathbf{x} \mid C_1)}{p(\mathbf{x} \mid C_2)} = \mathbf{w}^T \mathbf{x} + \mathbf{w}_0^o$$

$$\log \operatorname{it}(P(C_1 \mid \mathbf{x})) = \log \frac{P(C_1 \mid \mathbf{x})}{1 - P(C_1 \mid \mathbf{x})} = \log \frac{p(\mathbf{x} \mid C_1)}{p(\mathbf{x} \mid C_2)} + \log \frac{P(C_1)}{P(C_2)}$$

$$= \mathbf{w}^T \mathbf{x} + \mathbf{w}_0$$

$$\text{where } \mathbf{w}_0 = \mathbf{w}_0^o + \log \frac{P(C_1)}{P(C_2)}$$

$$y = \hat{P}(C_1 \mid \mathbf{x}) = \frac{1}{1 + \exp[-(\mathbf{w}^T \mathbf{x} + \mathbf{w}_0)]}$$

Training: Two Classes

$$\mathcal{X} = \left\{ \mathbf{x}^{t}, r^{t} \right\}_{t} \quad r^{t} \mid \mathbf{x}^{t} \sim \text{Bernoulli}(y^{t})$$

$$y = P(C_{1} \mid \mathbf{x}) = \frac{1}{1 + \exp[-(\mathbf{w}^{T}\mathbf{x} + \mathbf{w}_{0})]}$$

$$I(\mathbf{w}, \mathbf{w}_{0} \mid \mathcal{X}) = \prod_{t} (y^{t})^{(r^{t})} (1 - y^{t})^{(1 - r^{t})}$$

$$E = -\log I$$

$$E(\mathbf{w}, \mathbf{w}_{0} \mid \mathcal{X}) = -\sum_{t} r^{t} \log y^{t} + (1 - r^{t}) \log (1 - y^{t})$$

Training: Gradient-Descent

$$E(\mathbf{w}, w_0 \mid \mathcal{X}) = -\sum_{t} r^t \log y^t + (1 - r^t) \log (1 - y^t)$$

$$If y = \text{sigmoid}(a) \quad \frac{dy}{da} = y(1 - y)$$

$$\Delta w_j = -\eta \frac{\partial E}{\partial w_j} = \eta \sum_{t} \left(\frac{r^t}{y^t} - \frac{1 - r^t}{1 - y^t} \right) y^t (1 - y^t) x_j^t$$

$$= \eta \sum_{t} (r^t - y^t) x_j^t, j = 1, ..., d$$

$$\Delta w_0 = -\eta \frac{\partial E}{\partial w_0} = \eta \sum_{t} (r^t - y^t)$$

For
$$j=0,\ldots,d$$

$$w_j \leftarrow \operatorname{rand}(-0.01,0.01)$$
 Repeat
$$\operatorname{For}\ j=0,\ldots,d$$

$$\Delta w_j \leftarrow 0$$

$$\operatorname{For}\ t=1,\ldots,N$$

$$o\leftarrow 0$$

$$\operatorname{For}\ j=0,\ldots,d$$

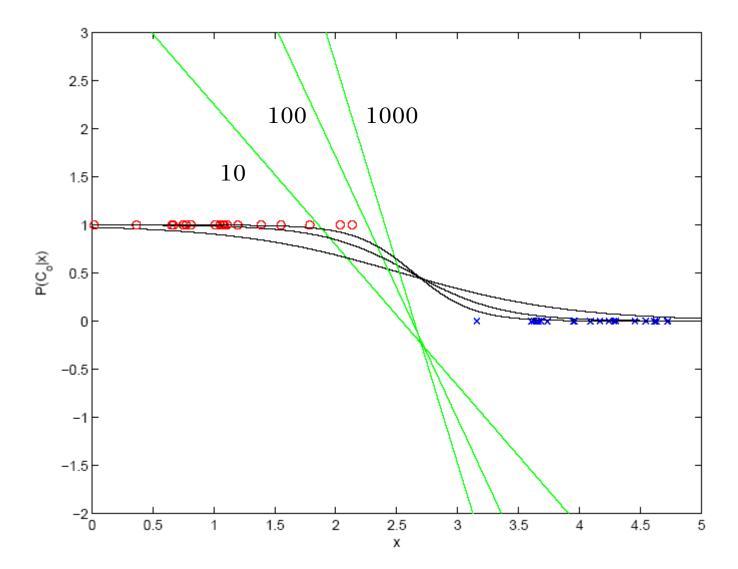
$$o\leftarrow o+w_jx_j^t$$

$$y\leftarrow\operatorname{sigmoid}(o)$$

$$\Delta w_j\leftarrow \Delta w_j+(r^t-y)x_j^t$$

$$\operatorname{For}\ j=0,\ldots,d$$

$$w_j\leftarrow w_j+\eta\Delta w_j$$
 Until convergence



K>2 Classes

$$\mathcal{X} = \left\{\mathbf{x}^{t}, \mathbf{r}^{t}\right\}_{t} \quad r^{t} \mid \mathbf{x}^{t} \sim \mathsf{Mult}_{K}(1, \mathbf{y}^{t})$$

$$\log \frac{\rho(\mathbf{x} \mid C_{i})}{\rho(\mathbf{x} \mid C_{K})} = \mathbf{w}_{i}^{T} \mathbf{x} + \mathbf{w}_{i0}^{o}$$

$$y = \hat{P}(C_{i} \mid \mathbf{x}) = \frac{\exp[\mathbf{w}_{i}^{T} \mathbf{x} + \mathbf{w}_{i0}]}{\sum_{j=1}^{K} \exp[\mathbf{w}_{j}^{T} \mathbf{x} + \mathbf{w}_{j0}]}, i = 1, ..., K \quad \text{softmax}$$

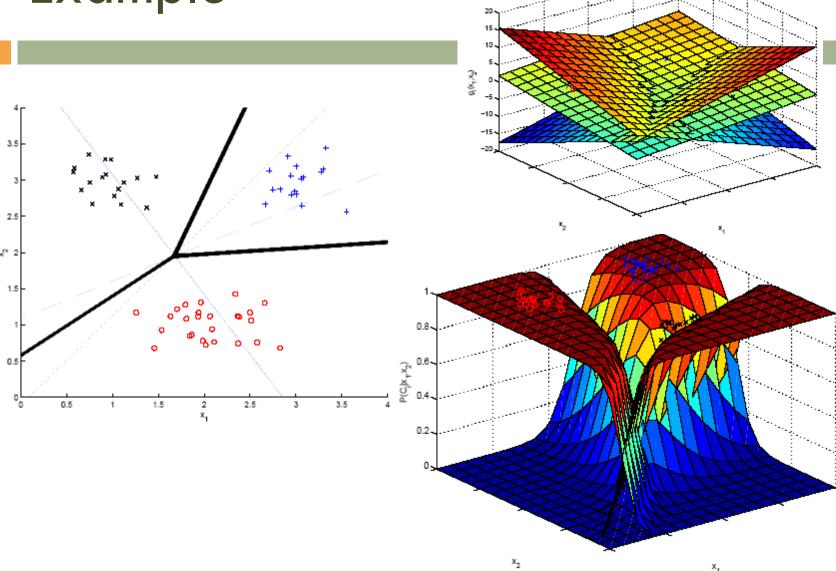
$$I(\left\{\mathbf{w}_{i}, \mathbf{w}_{i0}\right\}_{i} \mid \mathcal{X}) = \prod_{t} \prod_{i} \left(y_{i}^{t}\right)^{r_{i}^{t}}$$

$$E(\left\{\mathbf{w}_{i}, \mathbf{w}_{i0}\right\}_{i} \mid \mathcal{X}) = -\sum_{t} r_{i}^{t} \log y_{i}^{t}$$

$$\Delta \mathbf{w}_{j} = \eta \sum_{t} \left(r_{j}^{t} - y_{j}^{t}\right) \mathbf{x}^{t} \quad \Delta \mathbf{w}_{j0} = \eta \sum_{t} \left(r_{j}^{t} - y_{j}^{t}\right)$$

```
For i = 1, ..., K, For j = 0, ..., d, w_{ij} \leftarrow \text{rand}(-0.01, 0.01)
Repeat
      For i = 1, ..., K, For j = 0, ..., d, \Delta w_{ij} \leftarrow 0
      For t = 1, ..., N
             For i = 1, \ldots, K
                   o_i \leftarrow 0
                   For j = 0, \ldots, d
                          o_i \leftarrow o_i + w_{ij} x_i^t
             For i = 1, \ldots, K
                   y_i \leftarrow \exp(o_i) / \sum_k \exp(o_k)
             For i = 1, \ldots, K
                   For j = 0, \ldots, d
                          \Delta w_{ij} \leftarrow \Delta w_{ij} + (r_i^t - y_i) x_j^t
      For i = 1, \ldots, K
             For j = 0, \ldots, d
                   w_{ij} \leftarrow w_{ij} + \eta \Delta w_{ij}
Until convergence
```

Example



Generalizing the Linear Model

Quadratic:

$$\log \frac{p(\mathbf{x} \mid C_i)}{p(\mathbf{x} \mid C_K)} = \mathbf{x}^T \mathbf{W}_i \mathbf{x} + \mathbf{w}_i^T \mathbf{x} + \mathbf{w}_{i0}$$

Sum of basis functions:

$$\log \frac{p(\mathbf{x} \mid C_i)}{p(\mathbf{x} \mid C_K)} = \mathbf{w}_i^T \phi(\mathbf{x}) + \mathbf{w}_{i0}$$

where $\varphi(\mathbf{x})$ are basis functions. Examples:

- Hidden units in neural networks (Chapters 11 and 12)
- Kernels in SVM (Chapter 13)

Discrimination by Regression

Classes are NOT mutually exclusive and exhaustive

$$r^{t} = y^{t} + \varepsilon \text{ where } \varepsilon \sim \mathcal{N}(0, \sigma^{2})$$

$$y^{t} = \operatorname{sigmoid}(\mathbf{w}^{T}\mathbf{x}^{t} + w_{0}) = \frac{1}{1 + \exp[-(\mathbf{w}^{T}\mathbf{x}^{t} + w_{0})]}$$

$$I(\mathbf{w}, w_{0} \mid \mathcal{X}) = \prod_{t} \frac{1}{\sqrt{2\pi\sigma}} \exp\left[-\frac{(r^{t} - y^{t})^{2}}{2\sigma^{2}}\right]$$

$$E(\mathbf{w}, w_{0} \mid \mathcal{X}) = \frac{1}{2} \sum_{t} (r^{t} - y^{t})^{2}$$

$$\Delta \mathbf{w} = \eta \sum_{t} (r^{t} - y^{t}) y^{t} (1 - y^{t}) \mathbf{x}^{t}$$

Learning to Rank

- Ranking: A different problem than classification or regression
- Let us say x^u and x^v are two instances, e.g., two movies

We prefer u to v implies that $g(x^u) > g(x^v)$

where g(x) is a score function, here linear:

$$g(\mathbf{x}) = \mathbf{w}^{\mathsf{T}} \mathbf{x}$$

 Find a direction w such that we get the desired ranks when instances are projected along w

Ranking Error

■ We prefer u to v implies that $g(x^v) > g(x^v)$, so error is $g(x^v) - g(x^v)$, if $g(x^v) < g(x^v)$

$$E(\boldsymbol{w}|\{r^{u},r^{v}\}) = \sum_{r^{u} < r^{v}} [g(\boldsymbol{x}^{v}|\theta) - g(\boldsymbol{x}^{u}|\theta)]_{+}$$

where a_+ is equal to a if $a \ge 0$ and 0 otherwise.

