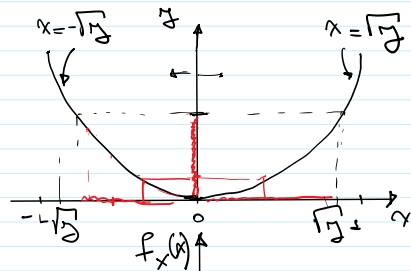


Sea $X \sim U(-1,1)$, y sea $Y=X^2$. Hallar la función de densidad de Y , $g(x)=Y=X^2$



$$f_X(x) = \frac{1}{2} \mathbb{I}\{-1 < x < 1\}, F_X(x) =$$

$$F_Y(y) = ? E_j: P(Y < \frac{1}{2}) = P(X^2 < \frac{1}{2})$$

$$F_Y(y) = P(Y \leq y) = P(X^2 \leq y) = P(|x| \leq \sqrt{y}) = P(-\sqrt{y} \leq x \leq \sqrt{y}) = \Phi$$

$$F_X(x) = \begin{cases} 0 & \text{si } x < -1 \\ \frac{x+1}{2} & \text{si } -1 \leq x < 1 \\ 1 & \text{si } x \geq 1 \end{cases}$$

$$\Phi = F_X(\sqrt{y}) - F_X(-\sqrt{y}) =$$

$$= \frac{\sqrt{y}+1}{2} - \frac{-\sqrt{y}+1}{2} = \sqrt{y}$$

$$F_Y(y) = \begin{cases} 0 & \text{si } y < 0 \\ \sqrt{y} & \text{si } 0 \leq y < 1 \\ 1 & \text{si } y \geq 1 \end{cases}$$

$$F_Y(y) = \sqrt{y} \mathbb{I}\{0 \leq y < 1\} + \mathbb{I}\{y \geq 1\}$$

$$f_Y(y) = \begin{cases} \frac{1}{2\sqrt{y}} & \text{si } 0 \leq y < 1 \\ 0 & \text{en otro caso.} \end{cases}, f_Y(y) = \frac{1}{2\sqrt{y}} \mathbb{I}\{0 \leq y < 1\}$$

Sean X e Y dos v.a. con distribución de Poisson de parámetros μ y λ respectivamente. Hallar la función de probabilidad de $W = X + Y$.

$$X \sim \text{Poi}(\mu) \rightarrow p_X(x) = \frac{\mu^x}{x!} e^{-\mu}, x \in \mathbb{N}_0$$

$$X \sim \text{Po}(\mu), Y \sim \text{Po}(\lambda), p_W(w) = ?$$

$$p_W(w) = P(W=w) = P(X+Y=w) = \sum_{x=0}^w P(X=x, Y=w-x) = \sum_{x=0}^w P(X=x) \cdot P(Y=w-x) = \sum_{x=0}^w \frac{\mu^x \cdot e^{-\mu}}{x!} \cdot \frac{\lambda^{w-x} \cdot e^{-\lambda}}{(w-x)!} =$$

$$= e^{-(\mu+\lambda)} \cdot \sum_{x=0}^w \frac{w!}{x!(w-x)!} \cdot \frac{\mu^x \cdot \lambda^{w-x}}{w!} \cdot \frac{1}{w!} =$$

$$= \frac{e^{-(\mu+\lambda)}}{w!} \sum_{x=0}^w \underbrace{\binom{w}{x}}_{=1} \cdot \underbrace{\left(\frac{\mu}{\mu+\lambda}\right)^x}_{X \sim \text{Bi}(w, \frac{\mu}{\mu+\lambda})} \cdot \underbrace{\left(\frac{\lambda}{\mu+\lambda}\right)^{w-x}}_{=1} \cdot (\mu+\lambda)^w = \frac{e^{-(\mu+\lambda)}}{w!} \cdot (\mu+\lambda)^w$$

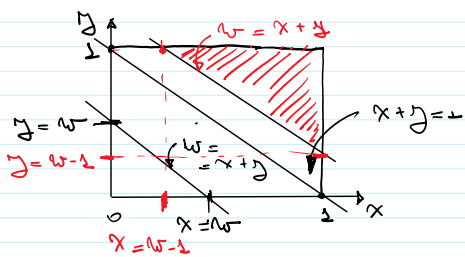
$$T \sim \text{Bi}(n, p), p_T(t) = \binom{n}{t} \cdot p^t \cdot (1-p)^{n-t}$$

$$\Phi = \frac{e^{-(\mu+\lambda)}}{w!} \cdot (\mu+\lambda)^w, W = X + Y \sim \text{Po}(\mu+\lambda)$$

Sean $X, Y \sim U(0,1)$ e independientes. Hallar la función de densidad de $W = X+Y$ si X e Y son independientes.



$$f_{X,Y}(x,y) = f_X(x) \cdot f_Y(y) = 1 \cdot 1 = 1$$



S: $0 < w < 1$

$$F_w(w) = P(W \leq w) = \frac{w^2}{2}$$

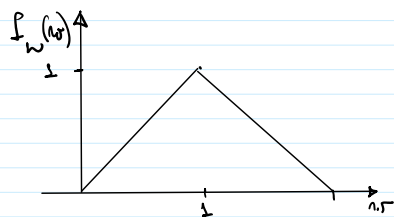
S: $1 \leq w \leq 2$

$$F_w(w) = 1 - P(W > w) = 1 - \frac{[1 - (w-1)] \cdot [1 - (w-1)]}{2} = 1 - \frac{(2-w)^2}{2}$$

$$F_w(w) = \begin{cases} 0 & \text{si } w < 0 \\ \frac{w^2}{2} & \text{si } 0 \leq w < 1 \\ 1 - \frac{(2-w)^2}{2} & \text{si } 1 \leq w < 2 \\ 1 & \text{si } w \geq 2 \end{cases}, \quad f_w(w) = \begin{cases} w & \text{si } 0 \leq w < 1 \\ 2-w & \text{si } 1 \leq w < 2 \\ 0 & \text{en otro caso} \end{cases}$$

$$F_w(w) = \frac{w^2}{2} \mathbb{I}\{0 \leq w < 1\} + \left[1 - \frac{(2-w)^2}{2}\right] \mathbb{I}\{1 \leq w < 2\} + \mathbb{I}\{w \geq 2\}$$

$$f_w(w) = w \mathbb{I}\{0 \leq w < 1\} + (2-w) \mathbb{I}\{1 \leq w < 2\}$$

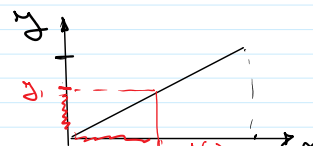


- Sea X una v.a.c. con función de densidad $f_X(x)$,
- Sea $Y=g(X)$.
- $g(x)$ es una función 1 a 1 (existe $g^{-1}(y)$)

$$f_Y(y) = f_X(g^{-1}(y)) \left| \frac{dg^{-1}(y)}{dy} \right|$$

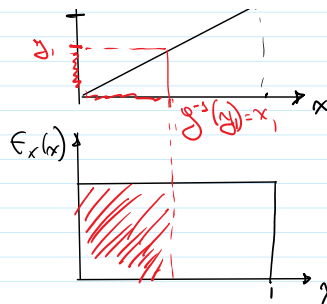
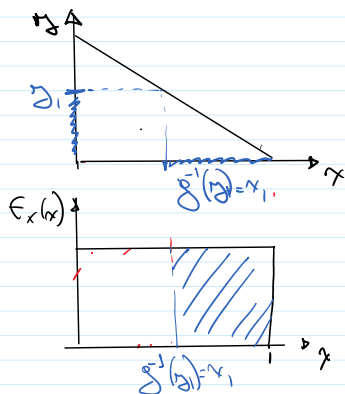
$$F_Y(y) = P(Y \leq y) = P(g(X) \leq y) = P(X \leq g^{-1}(y)) = F_X(g^{-1}(y))$$

$$f_Y(y) = \frac{dF_X(\overset{x}{g^{-1}(y)})}{dx} \cdot \frac{d\overset{x}{g^{-1}(y)}}{dy}$$



$$f_y(y) = \frac{dF_x(\overset{x}{g^{-1}(y)})}{dx} \cdot \frac{d\overset{x}{g^{-1}(y)}}{dy}$$

$$f_y(y) = f_x(g^{-1}(y)) \cdot \frac{dg^{-1}(y)}{dy}$$



$$F_y(y) = P(Y \leq y) = P(X > g^{-1}(y)) = 1 - F_x(g^{-1}(y))$$

$$f_y(y) = \frac{d(1 - F_x(\overset{x}{g^{-1}(y)}))}{dx} \cdot \frac{d\overset{x}{g^{-1}(y)}}{dy}$$

$$f_y(y) = -f_x(g^{-1}(y)) \cdot \frac{dg^{-1}(y)}{dy}$$

↗ va resultar negativo por la relación inversa entre X e Y

Entonces $f_y(y)$ siempre va resultar positivo:

Por lo tanto se define: $f_y(y) = f_x(g^{-1}(y)) \cdot \left| \frac{dg^{-1}(y)}{dy} \right|$ → para cualquier tipo de relación entre X e Y.

Sean $X_1, X_2 \stackrel{i.i.d}{\sim} \mathcal{E}(\lambda)$ y sean $U = X_1 + X_2$ y $V = \frac{X_1}{X_1 + X_2}$. Hallar $f_{U,V}(u,v)$ ¿Qué puede decir al respecto?

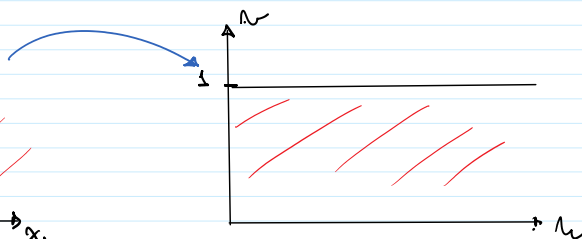
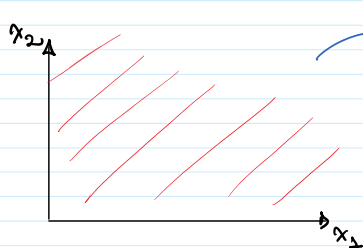
$$X_1 \sim \mathcal{E}(\lambda), X_2 \sim \mathcal{E}(\lambda), \quad \begin{cases} u = x_1 + x_2 > 0 \\ v = \frac{x_1}{x_1 + x_2} > 0 \end{cases}$$

$$\begin{cases} u = x_1 + x_2 \Rightarrow x_2 = u - x_1 \Rightarrow x_2 = u - u \cdot v = u(1-v) \\ v = \frac{x_1}{x_1 + x_2} \Rightarrow v = \frac{x_1}{x_1 + u - x_1} = \frac{x_1}{u} \Rightarrow x_1 = u \cdot v \end{cases}$$

$$|J| = \begin{vmatrix} v & u \\ 1-v & -u \end{vmatrix} = -u \cdot v - (1-v)u = -u$$

Por X_1 y X_2 independientes

$$f_{Y_1, Y_2} = \overbrace{f_{X_1, X_2}(x_1, x_2)}^{h_1^{-1}(u_1, u_2), h_2^{-1}(u_1, u_2)} |J|, \quad f_{X_1, X_2}(x_1, x_2) = \lambda e^{-\lambda x_1} \cdot \lambda e^{-\lambda x_2} \mathbb{I}\{x_1 > 0, x_2 > 0\} = \lambda^2 e^{-\lambda(x_1 + x_2)} \mathbb{I}\{x_1 > 0, x_2 > 0\}$$



$$\begin{aligned} u &= x_1 + x_2 \\ v &= \frac{x_1}{x_1 + x_2} \end{aligned}$$

$$f_{Y_1, Y_2}(y_1, y_2) = \lambda^2 e^{-\lambda(x_1 + x_2)} \mathbb{I}\{x_1 > 0, x_2 > 0\} \bigg|_{\substack{x_1 = u \cdot v \\ x_2 = u - u \cdot v}} \cdot |-u| =$$

$$= u \cdot \lambda^2 \cdot e^{-\lambda(u \cdot v + u - u \cdot v)} \cdot \mathbb{I} \{u \cdot v > 0, u(1-v) > 0\} =$$

$$= \underbrace{u \cdot \lambda^2 \cdot e^{-\lambda u}}_{\substack{\uparrow \\ \text{Gamma}}} \cdot \mathbb{I} \{u > 0, 0 < v < 1\}$$

$$= \underbrace{u \cdot \lambda^2 \cdot e^{-\lambda u}}_{U \sim \Gamma(\lambda, 2)} \cdot \mathbb{I} \{u > 0\} \cdot \underbrace{\mathbb{I} \{0 < v < 1\}}_{V \sim U(0, 1)}$$

$$f_x(x) = \frac{\lambda^k}{\Gamma(k)} x^{k-1} \cdot e^{-\lambda x}$$

\uparrow
Gamma, $X \sim \Gamma(\lambda, k)$