

1 Computing the gradient

The loss function is given by

$$\begin{aligned}
L(\tilde{\theta}) &= E \left[\sum_{k=1}^n \min(x_k, w_k) + \sum_{k=1}^n (1 + \gamma) (x_k - w_k) u_k(x) \right] \\
&\quad x \text{ drawn from joint cdf } F(x; \tilde{\theta}) \\
u_k(x) &= \begin{cases} 0 & , \text{ if } k \notin D \\ e_k (I_D - (1 + \gamma) A_D)^{-1} \mathbf{1}_D & , \text{ if } k \in D \end{cases} \\
D &= \{i : p_i < \bar{p}_i\} \\
p_i &= \min \left\{ \bar{p}_i, \max \left\{ (1 + \gamma) \left(\sum_j p_j a_{ji} + c_i - x_i \right) - \gamma \bar{p}_i, 0 \right\} \right\} \\
A_D &= \underbrace{S'}_{|D| \times |D|} \underbrace{A}_{|D| \times n} \underbrace{S}_{n \times |D|} \\
S_{ij} &= \begin{cases} 1 & , \text{ if } j \in D \text{ and } i \leq j \\ 0 & , \text{ otherwise} \end{cases} \\
e_k &= \begin{bmatrix} 0 & \dots & \underbrace{1}_{k^{th} \text{ entry}} & \dots & 0 \end{bmatrix} \\
I_D &= |D| \times |D| \text{ identity matrix} \\
\tilde{\theta} &= (A, b, c, w, \bar{p}, \delta, \gamma)
\end{aligned}$$

Let's write it in vector/matrix form

$$L(\tilde{\theta}) = E \left\{ \underbrace{\min \left(\underbrace{x}_{1 \times n}, \underbrace{w}_{1 \times n} \right) \underbrace{\mathbf{1}_n}_{n \times 1}}_{1 \times 1} + (1 + \gamma) \underbrace{\left(\underbrace{x - w}_{1 \times n} \right) \underbrace{S}_{n \times |D|} \left(\underbrace{I_D}_{|D| \times |D|} - (1 + \gamma) \underbrace{S^T A S}_{|D| \times n \times n \times |D|} \right)^{-1} \underbrace{\mathbf{1}_D}_{|D| \times 1}}_{\underbrace{|D| \times |D|}_{1 \times 1}} \right\}$$

$$\begin{aligned}
& x \text{ drawn from joint cdf } F(x; \tilde{\theta}) \text{ and joint pdf } f(x; \tilde{\theta}) \\
D &= \{i : p_i < \bar{p}_i\} \\
|D| &= \# \text{ elements in } D \\
p &= \min \left\{ \bar{p}, \max \left\{ (1 + \gamma) \left(\underbrace{\underbrace{p}_{1 \times n} \underbrace{A}_{n \times n}}_{1 \times n} + \underbrace{c - x}_{1 \times n} \right) - \gamma \bar{p}, 0 \right\} \right\} \\
S_{ij} &= \begin{cases} 1 & , \text{ if } j \in D \text{ and } i \leq j \\ 0 & , \text{ otherwise} \end{cases} \\
I_D &= |D| \times |D| \text{ identity matrix} \\
1_D &= \begin{bmatrix} 1 & \dots & 1 & \dots & 1 \end{bmatrix}^T_{|D| \times 1} \\
\tilde{\theta} &= (A, b, c, w, \bar{p}, \delta, \gamma)
\end{aligned}$$

$$\begin{aligned}
p_i &< \bar{p}_i \\
p_i &< \bar{p}_i
\end{aligned}$$

$$\begin{aligned}
case1 &: \bar{p} \leq \max \{ (1 + \gamma) (pA + c - x) - \gamma \bar{p}, 0 \} \\
case1a &: (1 + \gamma) (pA + c - x) - \gamma \bar{p} > 0 \\
case1b &: (1 + \gamma) (pA + c - x) - \gamma \bar{p} \leq 0
\end{aligned}$$

$$\begin{aligned}
case2 &: \bar{p} > \max \{ (1 + \gamma) (pA + c - x) - \gamma \bar{p}, 0 \} \\
case2a &: (1 + \gamma) (pA + c - x) - \gamma \bar{p} > 0 \\
case2b &: (1 + \gamma) (pA + c - x) - \gamma \bar{p} \leq 0
\end{aligned}$$

$$\begin{aligned}
A &\text{ is } n \times n \\
b, c, w, \bar{p}, \delta, x &\text{ are } 1 \times n \\
\gamma &\text{ is } 1 \times 1 \\
F(x; \tilde{\theta}) &\text{ is } 1 \times 1 \\
vec(\tilde{\theta}) &\text{ is } (n^2 + 5n + 1) \times 1
\end{aligned}$$

We solve

$$\begin{aligned}
& \max_{\theta} L(x; \tilde{\theta}) \\
& s.t. \\
& \dots
\end{aligned}$$

where θ is a subset of $\tilde{\theta}$.

To compute numerically, we want to compute

$$\nabla_{\theta} L(x; \tilde{\theta})$$

Let's make a change of variables so that we can draw the random variables from the uniform distribution U and then transform them

$$x = F^{-1}(u; \tilde{\theta})$$

$$\begin{aligned} L(\tilde{\theta}) &= \int_{[0,1]^n} \left\{ \min(x, w) 1_n + (1 + \gamma) \left[(x - w) S (I_D - (1 + \gamma) S^T A S)^{-1} 1_D \right] \right\} f(x; \tilde{\theta}) dx \\ &= \int_{[0,1]^n} \left\{ \min(F^{-1}(u; \tilde{\theta}), w) 1_n + (1 + \gamma) \left[(F^{-1}(u; \tilde{\theta}) - w) S (I_D - (1 + \gamma) S^T A S)^{-1} 1_D \right] \right\} c(u; \tilde{\theta}) du \end{aligned}$$

where $c(u; \tilde{\theta}) = 1$ for iid uniform. Now let's consider the beta distribution with parameters α and β such that $P(c_i X_i > w_i) = \delta_i$. Then we have

$$\begin{aligned} F(x; \tilde{\theta}) &= \frac{1}{B(\alpha, \beta)} x^{\alpha-1} (1-x)^{\beta-1} \\ &= \frac{1}{B(1, \beta)} (1-x)^{\beta-1} \\ &= 1 - (1-x)^{\beta} \end{aligned}$$

Therefore

$$\begin{aligned} \delta_i &= P(c_i X_i > w_i) \\ &= 1 - P(c_i X_i \leq w_i) \\ &= 1 - F_i\left(\frac{w_i}{c_i}; \tilde{\theta}\right) \\ &= 1 - \left(1 - \left(1 - \frac{w_i}{c_i}\right)^{\beta_i}\right) \\ &= \left(1 - \frac{w_i}{c_i}\right)^{\beta_i} \end{aligned}$$

We then deduce that

$$\beta_i = \frac{\log \delta_i}{\log \left(1 - \frac{w_i}{c_i}\right)}$$

The marginal cdf are

$$F(x; \tilde{\theta}) = 1 - (1-x)^{\frac{\log \delta_i}{\log \left(1 - \frac{w_i}{c_i}\right)}}$$

The inverse marginal cdf is

$$F^{-1}\left(u_i; \tilde{\theta}\right) = 1 - (1 - u_i)^{\frac{\log\left(1 - \frac{w_i}{c_i}\right)}{\log \delta_i}}$$

Taking derivatives, we have

$$\begin{aligned} \frac{\partial F^{-1}\left(u_i; \tilde{\theta}\right)}{\partial c_i} &= -(1 - u_i)^{\frac{\log\left(1 - \frac{w_i}{c_i}\right)}{\log \delta_i}} \frac{w_i \log(1 - u_i)}{c_i^2 \log \delta_i} \left(1 - \frac{w_i}{c_i}\right)^{-1} \\ \frac{\partial F^{-1}\left(u_i; \tilde{\theta}\right)}{\partial c_j} &= 0 \end{aligned}$$

Now for the whole loss

$$\begin{aligned} & \frac{\partial}{\partial c_i} E \left[\sum_{k=1}^n \min(c_k F_k^{-1}(u_k; c_k), w_k) + \sum_{k=1}^n (1 + \gamma) (c_k F_k^{-1}(u_k; c_k) - w_k) u_k(x) \right] \\ & E \left[\sum_{k=1}^n \frac{\partial}{\partial c_i} \{ \min(c_k F_k^{-1}(u_k; c_k), w_k) \} + \sum_{k=1}^n (1 + \gamma) \frac{\partial}{\partial c_i} \{ (c_k F_k^{-1}(u_k; c_k) - w_k) u_k(x) \} \right] \\ & E \left[\sum_{k=1}^n \frac{\partial}{\partial c_i} \{ \min(c_k F_k^{-1}(u_k; c_k), w_k) \} + \sum_{k=1}^n (1 + \gamma) \left[\frac{\partial}{\partial c_i} \{ (c_k F_k^{-1}(u_k; c_k) - w_k) \} u_k(x) + (c_k F_k^{-1}(u_k; c_k) - w_k) \frac{\partial u_k(x)}{\partial c_i} \right] \right] \\ & E \left[\sum_{k=1}^n \frac{\partial [c_k F_k^{-1}(u_k; c_k)]}{\partial c_i} 1 \{ c_k F_k^{-1}(u_k; c_k) < w_k \} + (1 + \gamma) \sum_{k=1}^n \frac{\partial [c_k F_k^{-1}(u_k; c_k)]}{\partial c_i} u_k(x) \right] \\ & E \left[\sum_{k=1}^n \{ 1 \{ c_k x_k < w_k \} + (1 + \gamma) u_k(x) \} \left(c_k \frac{\partial F_k^{-1}(u_k; c_k)}{\partial c_i} + F_k^{-1}(u_k; c_k) \frac{\partial c_k}{\partial c_i} \right) \right] \\ & E \left[\sum_{k=1}^n \{ 1 \{ c_k x_k < w_k \} + (1 + \gamma) u_k(x) \} \left(c_k \frac{\partial F_k^{-1}(u_k; c_k)}{\partial c_k} + x_k \right) \right] \\ \\ & \frac{\partial}{\partial c_i} \{ (c_k F_k^{-1}(u_k; c_k) - w_k) \} = \frac{\partial}{\partial c_i} c_k F_k^{-1}(u_k; c_k) \\ & = c_k \frac{\partial F_k^{-1}(u_k; c_k)}{\partial c_i} + x_k \frac{\partial c_k}{\partial c_i} \\ & = - (1 - u_i)^{\frac{\log\left(1 - \frac{w_i}{c_i}\right)}{\log \delta_i}} \frac{w_i \log(1 - u_i)}{c_i \log \delta_i} \left(1 - \frac{w_i}{c_i}\right)^{-1} + F_i^{-1}(u_i; c_i) \quad , \quad i = k \\ & \quad \quad \quad 0 \quad , \quad \text{otherwise} \\ \\ & \frac{\partial}{\partial c_i} \{ \min(c_k F_k^{-1}(u_k; c_k), w_k) \} = \frac{F_k^{-1}(u_k; c_k) + c_k \frac{\partial}{\partial c_i} F_k^{-1}(u_k; c_k)}{0} \quad , \quad c_k F_k^{-1}(u_k; c_k) < w_k \\ & \quad \quad \quad , \quad \text{otherwise} \\ & = \frac{x_k + c_k \frac{\partial}{\partial c_i} F_k^{-1}(u_k; c_k)}{0} \quad , \quad c_k x_k < w_k \\ & \quad \quad \quad , \quad \text{otherwise} \\ \\ & \frac{\partial u_k(c \circ x)}{\partial c_j} = 0 \end{aligned}$$

$$\begin{aligned}
\frac{\partial [(c_k x_k - w_k) u_k(c \circ x)]}{\partial c_j} &= u_k(c \circ x) \frac{\partial (c_k x_k - w_k)}{\partial c_j} + (c_k x_k - w_k) \frac{\partial u_k(c \circ x)}{\partial c_j} \\
\frac{\partial (c_k x_k - w_k)}{\partial c_j} &= x_k \frac{\partial c_k}{\partial c_j} = \begin{cases} x_k & , \quad j = k \\ 0 & , \quad otherwise \end{cases} \\
\frac{\partial u_k(c \circ x)}{\partial c_j} &= 0
\end{aligned}$$

$$\begin{aligned}
(\nabla_c L)_i &= \left\{ x_i + c_i \frac{\partial F_i^{-1}(u_i; c_i)}{\partial c_i} \right\} 1\{c_i x_i < w_i\} + \left\{ x_i + c_i \frac{\partial F_k^{-1}(u_i; c_i)}{\partial c_i} \right\} \\
&= (1\{c_i x_i < w_i\} + 1) \left\{ x_i + c_i \frac{\partial F_k^{-1}(u_i; c_i)}{\partial c_i} \right\}
\end{aligned}$$

$$\begin{aligned}
\frac{\partial [e_k (I_D - (1 + \gamma) A_D)^{-1} 1_D]}{\partial A_D} &= (1 + \gamma) (I_D - (1 + \gamma) A_D)^{-T} e_k^T 1_D^T (I_D - (1 + \gamma) A_D)^{-T} \\
&= (1 + \gamma) [e_k (I_D - (1 + \gamma) A_D)^{-1}]^T [(I_D - (1 + \gamma) A_D)^{-1} 1_D]^T \\
&= (1 + \gamma) (I_D - (1 + \gamma) A_D)^{-T} e_k^T u_D^T
\end{aligned}$$

2 Old

We would like to compute the gradient of expected losses

$$\nabla_A \left\{ E \left[\sum_{k=1}^n \min(x_k, w_k) + \sum_{k=1}^n (x_k - w_k) u_k(x) \right] \right\}$$

where ∇_A means the gradient with respect to the entries of the matrix A . Since x_k and w_k do not depend on A , and expectations are linear,

$$\nabla_A \left\{ E \left[\sum_{k=1}^n \min(x_k, w_k) + \sum_{k=1}^n (x_k - w_k) u_k(x) \right] \right\} = E \sum_{k=1}^n (x_k - w_k) \nabla_A u_k(x)$$

To compute $\nabla_A u_k(x)$, we use that

$$u_k(x) = \begin{cases} 0 & , \quad \text{if } k \notin D(A) \\ e_k (I_D - (1 + \gamma) A_D)^{-1} 1_D & , \quad \text{if } k \in D(A) \end{cases}$$

where D is the set of defaulting nodes, I_D is the $|D| \times |D|$ identity matrix, e_k is a row vector with a 1 in column k and zeros otherwise, and A_D is the matrix A restricted to D . We can write

$$A_D = S' A S$$

for a matrix S with dimension $n \times |D|$

$$S_{ij} = \begin{cases} 1 & , \text{ if } j \in D \text{ and } i \leq j \\ 0 & , \text{ otherwise} \end{cases}$$

so that A_D has dimension $|D| \times |D|$. For example, if $n = 3$ and $D = \{1, 3\}$

$$S = \begin{bmatrix} 1 & 0 \\ 0 & 0 \\ 0 & 1 \end{bmatrix}$$

and

$$\begin{aligned} A_D &= \begin{bmatrix} 1 & 0 & 0 \\ 0 & 0 & 1 \end{bmatrix} \begin{bmatrix} a_{11} & a_{12} & a_{13} \\ a_{21} & a_{22} & a_{23} \\ a_{31} & a_{32} & a_{33} \end{bmatrix} \begin{bmatrix} 1 & 0 \\ 0 & 0 \\ 0 & 1 \end{bmatrix} \\ &= \begin{bmatrix} a_{11} & a_{13} \\ a_{31} & a_{33} \end{bmatrix} \end{aligned}$$

We also note that D depends on A

$$\begin{aligned} D(A) &= \{i : p_i(A) < \bar{p}_i\} \\ &= \left\{ i : \min \left\{ \bar{p}_i, \max \left\{ \sum_j p_j(A) a_{ji} + c_i - x_i, 0 \right\} \right\} < \bar{p}_i \right\} \end{aligned}$$

However, $p_i(A)$ is continuous in A , so small changes in A can only affect D if

$$p_i(A) = \bar{p}_i$$

We now compute

$$\nabla_A u_k(x) = \begin{cases} 0 & , \text{ if } k \notin D(A) \\ \nabla_A e_k (I_D - S^T A S)^{-1} 1_D & , \text{ if } k \in D(A) \end{cases}$$

Use the formulas

$$\begin{aligned} \frac{\partial a^T X^{-1} b}{\partial X} &= -X^{-T} a b^T X^{-T} \\ \partial(XY) &= (\partial X) Y + X \partial Y \\ \partial(S^T A S) &= \partial(S^T A) S + (S^T A) \partial S \\ &= [\partial(S^T) A + S^T \partial A] S + (S^T A) \partial S \\ &= S^T (\partial A) S \end{aligned}$$

to get

$$\begin{aligned} \frac{\partial \left[e_k (I_D - (1 + \gamma) A_D)^{-1} 1_D \right]}{\partial A_D} &= (1 + \gamma) (I_D - (1 + \gamma) A_D)^{-T} e_k^T 1_D^T (I_D - (1 + \gamma) A_D)^{-T} \\ &= (1 + \gamma) \left[e_k (I_D - (1 + \gamma) A_D)^{-1} \right]^T \left[(I_D - (1 + \gamma) A_D)^{-1} 1_D \right]^T \\ &= (1 + \gamma) (I_D - (1 + \gamma) A_D)^{-T} e_k^T u_D^T \end{aligned}$$

and

$$\frac{\partial \left[e_k (I_D - A_D)^{-1} 1_D \right]}{\partial A_{D^c}} = 0$$

where D^c is the complement of D .

$$\nabla_A u_k(x) = S [\nabla_{A_D} u_k(x)] S^T$$

Finally,

$$E \sum_{k=1}^n (x_k - w_k) S [\nabla_{A_D} u_k(x)] S^T$$

Now we compute the gradient with respect to c , assuming that x_k is a random variable between 0 and 1 and thus total losses are

$$E \left[\sum_{k=1}^n \min(c_k x_k, w_k) + \sum_{k=1}^n (c_k x_k - w_k) u_k(c \circ x) \right]$$

where $c \circ x$ denotes element-wise multiplication.

We compute

$$\nabla_c \left\{ E \left[\sum_{k=1}^n \min(c_k x_k, w_k) + \sum_{k=1}^n (c_k x_k - w_k) u_k(c \circ x) \right] \right\}$$

Using

$$\begin{aligned} \frac{\partial \min(c_k x_k, w_k)}{\partial c_j} &= \begin{cases} x_k & , \quad c_k < \frac{w_k}{x_k} \text{ and } j = k \\ 0 & , \quad \text{otherwise} \end{cases} \\ \nabla_c \{ \min(c_k x_k, w_k) \} &= \begin{pmatrix} 0 & \cdots & 1_{\{c_k < \frac{w_k}{x_k}\}} x_k & \cdots & 0 \end{pmatrix} \\ \sum_{k=1}^n \nabla_c \{ \min(c_k x_k, w_k) \} &= \begin{pmatrix} 1_{\{c_1 < \frac{w_1}{x_1}\}} x_1 & \cdots & 1_{\{c_k < \frac{w_k}{x_k}\}} x_k & \cdots & 1_{\{c_n < \frac{w_n}{x_n}\}} x_n \end{pmatrix} \end{aligned}$$

and

$$\begin{aligned} \frac{\partial [(c_k x_k - w_k) u_k(c \circ x)]}{\partial c_j} &= u_k(c \circ x) \frac{\partial (c_k x_k - w_k)}{\partial c_j} + (c_k x_k - w_k) \frac{\partial u_k(c \circ x)}{\partial c_j} \\ \frac{\partial (c_k x_k - w_k)}{\partial c_j} &= x_k \frac{\partial c_k}{\partial c_j} = \begin{cases} x_k & , \quad j = k \\ 0 & , \quad \text{otherwise} \end{cases} \\ \frac{\partial u_k(c \circ x)}{\partial c_j} &= 0 \end{aligned}$$

we get

$$\begin{aligned}
& \nabla_c \left\{ E \left[\sum_{k=1}^n \min(c_k x_k, w_k) + \sum_{k=1}^n (c_k x_k - w_k) u_k(c \circ x) \right] \right\} \\
&= E \left[\sum_{k=1}^n \nabla_c \min(c_k x_k, w_k) + \sum_{k=1}^n \nabla_c [(c_k x_k - w_k) u_k(c \circ x)] \right] \\
&= E \left[\begin{pmatrix} 1_{\{c_1 < \frac{w_1}{x_1}\}} x_1 & \cdots & 1_{\{c_k < \frac{w_k}{x_k}\}} x_k & \cdots & 1_{\{c_n < \frac{w_n}{x_n}\}} x_n \\ + (x_1 u_1(c \circ x) & \cdots & x_k u_k(c \circ x) & \cdots & x_n u_n(c \circ x)) \end{pmatrix} \right] \\
&= E \left[\begin{pmatrix} \left[1_{\{c_1 < \frac{w_1}{x_1}\}} + u_1(c \circ x) \right] x_1 & \cdots & \left[1_{\{c_k < \frac{w_k}{x_k}\}} + u_k(c \circ x) \right] x_k & \cdots & \left[1_{\{c_n < \frac{w_n}{x_n}\}} + u_n(c \circ x) \right] x_n \end{pmatrix} \right]
\end{aligned}$$

We pick the parameter β of the beta distribution with pdf

$$f(x) = \frac{1}{B(\alpha, \beta)} x^{\alpha-1} x^{\beta-1}$$

and cdf

$$F(x) = \frac{1}{B(\alpha, \beta)} \int_0^x x^{\alpha-1} x^{\beta-1} dx$$

so that

$$\begin{aligned}
\delta &= P(cx > w) \\
&= 1 - F\left(\frac{w}{c}, \alpha, \beta\right) \\
&= 1 - \frac{1}{B(\alpha, \beta)} \int_0^{\frac{w}{c}} x^{\alpha-1} x^{\beta-1} dx
\end{aligned}$$

Solving for β gives a solution

$$\beta = \beta(c, w, \delta, \alpha)$$

The CDF and PDF are then

$$\begin{aligned}
F(x; c) &= \frac{1}{B(\alpha, \beta(c, w, \delta, \alpha))} \int_0^x x^{\alpha-1} x^{\beta(c, w, \delta, \alpha)-1} dx \\
f(x; c) &= \frac{1}{B(\alpha, \beta(c, w, \delta, \alpha))} x^{\alpha-1} x^{\beta(c, w, \delta, \alpha)-1}
\end{aligned}$$

To compute

$$\nabla_c \left\{ E \left[\sum_{k=1}^n \min(c_k x_k, w_k) + \sum_{k=1}^n (c_k x_k - w_k) u_k(c \circ x) \right] \right\}$$

we then note that

$$\begin{aligned}
& \nabla_c \left\{ E \left[\sum_{k=1}^n \min(c_k x_k, w_k) + \sum_{k=1}^n (c_k x_k - w_k) u_k(c \circ x) \right] \right\} \\
&= \nabla_c \int_0^1 \dots \int_0^1 \left[\sum_{k=1}^n \min(c_k x_k, w_k) + \sum_{k=1}^n (c_k x_k - w_k) u_k(c \circ x) \right] f(x; c) dx_1 \dots dx_n \\
&= \int_0^1 \dots \int_0^1 \nabla_c \left\{ \left[\sum_{k=1}^n \min(c_k x_k, w_k) + \sum_{k=1}^n (c_k x_k - w_k) u_k(c \circ x) \right] f(x; c) \right\} dx_1 \dots dx_n \\
&= \int_0^1 \dots \int_0^1 \nabla_c \left[\sum_{k=1}^n \min(c_k x_k, w_k) + \sum_{k=1}^n (c_k x_k - w_k) u_k(c \circ x) \right] f(x; c) dx_1 \dots dx_n \\
&\quad + \int_0^1 \dots \int_0^1 \left[\sum_{k=1}^n \min(c_k x_k, w_k) + \sum_{k=1}^n (c_k x_k - w_k) u_k(c \circ x) \right] \nabla_c f(x; c) dx_1 \dots dx_n \\
&= \int_0^1 \dots \int_0^1 \nabla_c \left[\sum_{k=1}^n \min(c_k x_k, w_k) + \sum_{k=1}^n (c_k x_k - w_k) u_k(c \circ x) \right] f(x; c) dx_1 \dots dx_n \\
&\quad + \int_0^1 \dots \int_0^1 L(x) \frac{\nabla_c f(x; c)}{f(x; c)} f(x; c) dx_1 \dots dx_n \\
&= E \left[\nabla_c \left[\sum_{k=1}^n \min(c_k x_k, w_k) + \sum_{k=1}^n (c_k x_k - w_k) u_k(c \circ x) \right] \right] + E \left[L(x) \frac{\nabla_c f(x; c)}{f(x; c)} \right]
\end{aligned}$$

For independent random variables x

$$f(x; c) = \prod_{k=1}^n f(x_k; c_k)$$

so that

$$\frac{\nabla_c f(x; c)}{f(x; c)} = \left(\frac{1}{f(x_1; c_1)} \frac{df(x_1; c_1)}{dc_1}, \dots, \frac{1}{f(x_k; c_k)} \frac{df(x_k; c_k)}{dc_k}, \dots, \frac{1}{f(x_n; c_n)} \frac{df(x_n; c_n)}{dc_n} \right)$$

Assume $\alpha = 1$. Then

$$\begin{aligned}
\delta &= 1 - \frac{1}{B(1, \beta)} \int_0^{\frac{w}{c}} x^{\beta-1} dx \\
\delta &= \left(1 - \frac{w}{c}\right)^\beta \\
\beta &= \frac{\log \delta}{\log \left(1 - \frac{w}{c}\right)}
\end{aligned}$$

and

$$\begin{aligned}
f_k(x; 1, \beta) &= \beta (1-x)^{\beta-1} \\
f_k\left(x; 1, \frac{\log \delta}{\log\left(1-\frac{w}{c}\right)}\right) &= \frac{\log \delta}{\log\left(1-\frac{w}{c}\right)} (1-x)^{\frac{\log \delta}{\log\left(1-\frac{w}{c}\right)}-1} \\
\frac{1}{f(x_k; c_k)} \frac{df(x_k; c_k)}{dc_k} &= \frac{1}{\frac{\log \delta}{\log\left(1-\frac{w}{c}\right)} (1-x)^{\frac{\log \delta}{\log\left(1-\frac{w}{c}\right)}-1}} \left(\left\{ \frac{\log \delta}{\log\left(1-\frac{w}{c}\right)} \right\} \frac{d}{dc} \left\{ (1-x)^{\frac{\log \delta}{\log\left(1-\frac{w}{c}\right)}-1} \right\} + \left\{ (1-x)^{\frac{\log \delta}{\log\left(1-\frac{w}{c}\right)}-1} \right\} \right) \\
&= \frac{1}{(1-x)^{\frac{\log \delta}{\log\left(1-\frac{w}{c}\right)}}} \frac{d}{dc} \left\{ (1-x)^{\frac{\log \delta}{\log\left(1-\frac{w}{c}\right)}} \right\} - \frac{1}{\log\left(1-\frac{w}{c}\right)} \frac{1}{(c-w)} \frac{w}{c} \\
&= -\frac{w \left(\log\left(1-\frac{w}{c}\right) + \log(1-x) \log(\delta) \right)}{c(c-w) \log\left(1-\frac{w}{c}\right)^2}
\end{aligned}$$

If we use a gaussian copula with correlation matrix R and marginals

$$f(x_k; c_k) = \frac{1}{B(\alpha_k, \beta(c_k, w_k, \delta_k, \alpha_k))} x^{\alpha_k-1} x^{\beta(c_k, w_k, \delta_k, \alpha_k)-1}$$

the density of the copula is

$$c_R^{Gauss}(u) = \frac{1}{\sqrt{\det R}} \exp \left(-\frac{1}{2} \begin{pmatrix} \Phi^{-1}(u_1) \\ \vdots \\ \Phi^{-1}(u_n) \end{pmatrix}^T (R^{-1} - I) \begin{pmatrix} \Phi^{-1}(u_1) \\ \vdots \\ \Phi^{-1}(u_n) \end{pmatrix} \right)$$

Thus,

$$f(x; c) = c_R^{Gauss}(F(x_1; c_1), \dots, F(x_n; c_n)) f(x_1; c_1) \dots f(x_n; c_n)$$

We can now compute the i^{th} component of $\frac{\nabla_c f(x; c)}{f(x; c)}$:

$$c_{R,k}^{Gauss}(F(x_1; c_1), \dots, F(x_n; c_n)) f(x_k; c_k) \frac{df(x_k; c_k)}{dc_k} + c_R^{Gauss}(F(x_1; c_1), \dots, F(x_n; c_n)) \frac{1}{f(x_k; c_k)} \frac{df(x_k; c_k)}{dc_k}$$

Try

$$\begin{aligned}
&\nabla_c \left\{ E \left[\sum_{k=1}^n \min(c_k F^{-1}(u_k; c_k), w_k) + \sum_{k=1}^n (c_k F^{-1}(u_k; c_k) - w_k) u_k (c \circ F^{-1}(x)) \right] \right\} \\
&= \nabla_c \left\{ E \left[\sum_{k=1}^n \min(c_k x_k, w_k) + \sum_{k=1}^n (c_k x_k - w_k) u_k (c \circ x) \right] \right\} \\
&+ \left\{ E \left[\sum_{k=1}^n \min(c_k \nabla_c F^{-1}(u_k; c_k), w_k) + \sum_{k=1}^n (c_k \nabla_c F^{-1}(u_k; c_k) - w_k) u_k (c \circ F^{-1}(x)) \right] \right\}
\end{aligned}$$