Gaussian Processes Part 2

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Agenda

- 1 Introduction
- Q GP approach
 - Introduction
 - Non-periodic part
 - Periodic part
 - The Model
- Modelling
 - Priors
 - Parametrization
 - Seasonality
- Stochastic Volatility

Time Series, Classical Approach

If data has seasonality, you usually use STL decomposition. However,

- Parameters are not interpretable, only decomposition is available
- No uncertainty estimates
- Quite strict on input values
- Significantly less flexible in modelling

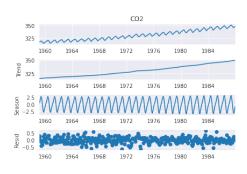


Figure: STL decomposition for CO2 data. Statsmodels

GP decomposition

A Gaussian process can handle a complicated set of assumptions in addition to what STL provides

- Granular seasonality (year + quarter + month + week)
- Changepoint models
- Flexible likelihood Function
- Panel regression models
- Missing values

Typical Model

Typical model is additive

$$x_t \sim \underbrace{g(t)}_{non-periodic} + \underbrace{s(t)}_{periodic} + \underbrace{h(t)}_{holidays}$$

Reference

See more in Prophet preprint [1]. Every time series model is unique

Reminder

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, $y \in \mathbb{R}$

$$Y \sim \mathcal{GP}(m(x), k(x, x'))$$

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1 \mathcal{GP} Gaussian Process - simply, a normal distribution with special mean m(x) and covariance k(x, x')

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$$Y \sim \mathcal{GP}(\mathbf{m}(\mathbf{x}), k(\mathbf{x}, \mathbf{x}'))$$

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- $\mathbf{2} \ m(x)$ mean function, e.g.
 - Linear regression $m(x) = x^{\top} \beta$
 - Simply Constant or Zero m(x) = c
 - Other custom functions $m(x) = \sin(x)$

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 - Linear regression $m(x) = x^{\top} \beta$
 - Simply Constant or Zero m(x) = c
 - Other custom functions $m(x) = \sin(x)$
- 3 k(x,x') kernel function, simply measure of similarity for x and x'
 - $[K]_{ij} = k(x_i, x_j)$ is an SPD matrix

- Growth models
- Linear trend models
- Changepoint models

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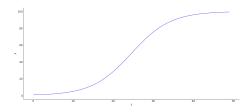


Figure: Growth Model

$$x = \frac{c}{1 + \exp(-k(t - m))}$$

- Growth models
- Linear trend models
- Changepoint models

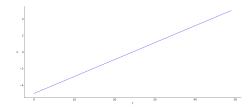


Figure: Linear Trend Model

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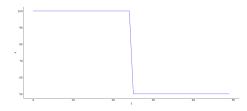


Figure: Changepoint Model

$$x = \begin{cases} c_1, & t < m \\ c_2, & t \ge m \end{cases}$$

- Growth models
- Linear trend models
- Changepoint models

Extentions

Extensions are possible, e.g. time dependent saturation in the growth model. See in [1]

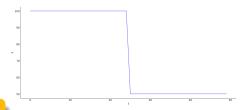


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Holidays

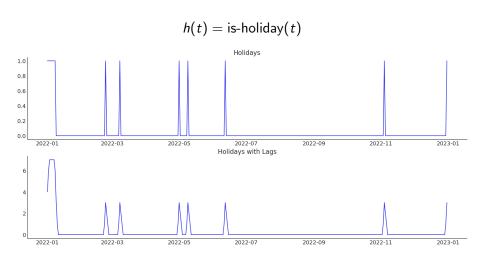
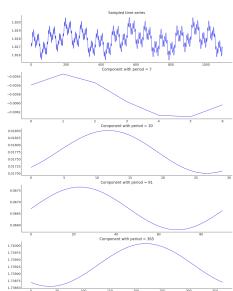


Figure: Holiday features

Periodic part (cov function)

Granularities are important here. Multiple Periodic kernels can be used.

- Yearly
- Quarterly
- Monthly
- Weekly



Lengthscales for Periodic Part

Hyperparameters

Common sense driven lenthscale choice

- Week couple of days make a change ($ls \approx 3$)
- Month week makes sense for a dramatic change ($ls \approx 7$)
- Quarter month makes sense for a dramatic change ($ls \approx 30$)
- Year quarter makes sense for a dramatic change ($ls \approx 90$)

Lengthscales for Periodic Part

Hyperparameters

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- Week couple of days make a change ($ls \approx 3$)
- Month week makes sense for a dramatic change ($ls \approx 7$)
- Quarter month makes sense for a dramatic change ($ls \approx 30$)
- Year quarter makes sense for a dramatic change ($ls \approx 90$)

In practice

Hard to infer lengthscales, better just fix them if informed

Putting All Together

$$m(t) = \underbrace{g(t)}_{non-periodic} + \underbrace{h(t)}_{holidays}$$

$$k(*,*) = \operatorname{Periodic}(p = 365, l = 90)$$

$$+ \operatorname{Periodic}(p = 90, l = 30)$$

$$+ \operatorname{Periodic}(p = 30, l = 7)$$

$$+ \operatorname{Periodic}(p = 7, l = 3)$$

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$$m(t) = \underbrace{g(t)}_{non-periodic} + \underbrace{h(t)}_{holidays}$$

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Missing Parts

1 Weights for periodic components

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$$+ \alpha_{7} \operatorname{Periodic}(\mathsf{p} = 7, \mathsf{I} = 3)$$

$$+ \beta \operatorname{ExpQuad}(\mathsf{I} = 90) + \operatorname{WhiteNoise}(\gamma)$$

Missing Parts

- Weights for periodic components
- **2** Trent violations from g(t) that can't be explained by periodic component and observation noise

Priors

Example

CO2 data (Manua Loa, Hawaii)

- Monthly observations
- Clean data
- Unclear trend function

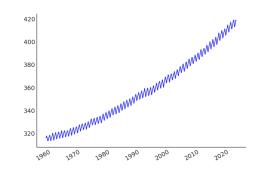


Figure: CO2 level

Example

CO2 data (Manua Loa, Hawaii)

- Monthly observations
- Clean data
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Where to start?

- Decide on a trend model
- 2 Add periodic component

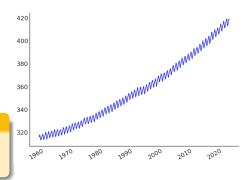


Figure: CO2 level

Priors

The exponential Growth Model

$$y_t = y_0 \cdot g^t$$

- **1** The baseline level for CO2 (data tells it is ≈ 310)
- 2 Growth rate for CO2 (some plausible value)
 - unlikely decay
 - no astronomic values

Prior Predictive Checks (Naive Attempt)

```
g \sim \mathsf{Normal}(1.1, 0.1)

\epsilon \sim \mathsf{LogNormal}(0, 1)

y_0 \sim \mathsf{Normal}(310, 10)
```

 $y_t \sim \mathsf{Normal}(y_0 \cdot g^t, \epsilon)$

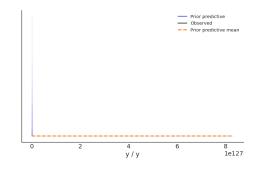


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PPC Observation

Model is unlikely to be correctly specified. Prior predictive ranges to 10^{127} parts per million of CO2. Air has $2.5 \cdot 10^{25}$ molecules in m³

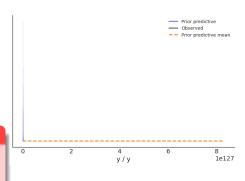


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Thinking loudly

I do not believe in a century we can exceed 400% growth (we would have died already) and I suppose it should be at least 20% (it would not be a problem otherwise)

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- 2 20% in a century is 0.015% per month

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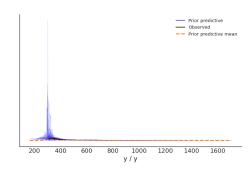


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PPC Observation

Model acts as expected ranging CO2 levels in plausible regions. It seems to be good enough to go to the next step.

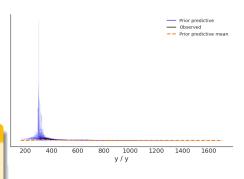


Figure: Prior Predictive Check

Inference

- In sampling any value could appear
- Here we got a very large value for some parameter
- 3 We can try to reparametrize

```
: with model:
    trace = pm.sample()
Auto-assigning NUTS sampler...
Initializing NUTS using jitter+adapt_diag...
/home/ferres/.miniconda3/envs/bayes-econ/lib/python3.10/site-pric.py:57: RuntimeWarning: overflow encountered in multiply
    return bound(*args, **kwds)
Multiprocess sampling (4 chains in 4 jobs)
MUTS: [initial, g, eps]
```

Figure: Overflow in sampling

Log Model

$$\log y_t = \log y_0 + g \cdot t$$
 $g \sim \mathsf{Normal}(1.00065, 0.0005)$ $\epsilon \sim \mathsf{LogNormal}(-2, 0.5)$ $\log y_0 \sim \mathsf{Normal}(\log 310, 10/310)$ $y_t \sim \mathsf{LogNormal}(\log y_0 + g \cdot t, \epsilon)$

Important

 ϵ has a different meaning in the model. It is now a relative error.

Prior Predictive (Reparametrized)

```
\begin{split} g &\sim \mathsf{Normal}(1.00065, 0.0005) \\ \epsilon &\sim \mathsf{LogNormal}(-2, 0.5) \\ \mathsf{log}\, y_0 &\sim \mathsf{Normal}(\mathsf{log}\, 310, 10/310) \\ y_t &\sim \mathsf{LogNormal}(\mathsf{log}\, y_0 + g \cdot t, \epsilon) \end{split}
```

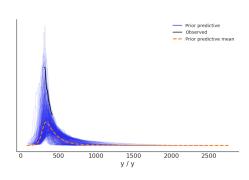


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PPC Observation

Model acts as expected ranging CO2 levels in plausible regions. It seems to be good enough to go to proceed with sampling.

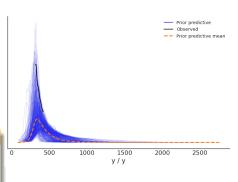


Figure: Prior Predictive Check

Inference

- Everything is good
- We can visualize



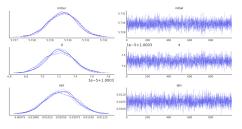
Sampling 4 chains, 0 divergences1

Sampling 4 chains for 1_000 tune and 1_000 draw iterations (4 s total) took 6 seconds.

Figure: Good sampling

Inference

- Everything is good
- We can visualize
- Trace plot is a good visual inspection if sampling went well



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with model:
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Auto-assigning NUTS sampler...
Initializing NUTS using jitter+adapt_diag...
Multiprocess sampling (4 chains in 4 jobs)
NUTS: [initial, g, eps]

100.00% [8000/8000

Sampling 4 chains, 0 divergences]

Sampling 4 chains for 1_000 tune and 1_000 draw iterations (4 s total) took 6 seconds.
```

Figure: Good sampling

Posterior Predictive

Another way to validate your model is posterior predictive.

- How model performs vs observed data?
- Is there any missing effect to include?

Let's see

Posterior Predictive

Another way to validate your model is posterior predictive.

- How model performs vs observed data?
- Is there any missing effect to include?

Let's see

- 1 The fit is good
- Model does not reflect seasonalities
- We can fix it with a better model

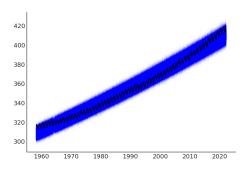


Figure: Posterior Predictive Plot

Conclusions so far

- The reparametrized model is decent
- It does not account for seasonality
- We have the mean function for the next modelling step

```
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- Monthly observations
- Yearly seasonality makes sense
- Trend function is clear from a simpler model

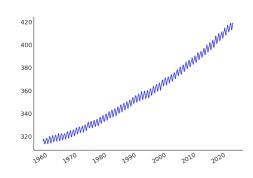


Figure: CO2 level

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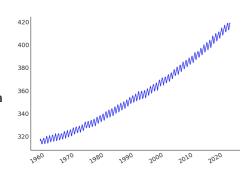


Figure: CO2 level

$$k(*,*) = Periodic(p = , l =)$$

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- Period is 12 months.

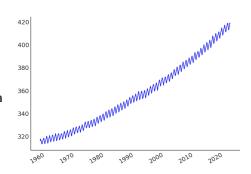


Figure: CO2 level

$$k(*,*) = Periodic(p = 12, l =)$$

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- Period is 12 months
- Appropriate lengthscale is ≈ 3 months

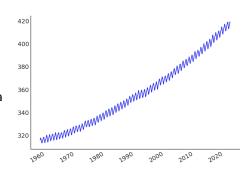


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- Do not forget to scale the covariance!



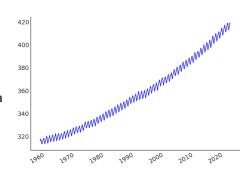


Figure: CO2 level

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- Seasonality pattern is assumed constant (what if ...)

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- We need a decay in correlation between distant year patterns

$$K(*,*) = \alpha_k \operatorname{Periodic}(p = 12, l = 3) \cdot \operatorname{Exponential}(l = 1200)$$

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- We need a decay in correlation between distant year patterns
- Multiplication with the Exponential kernel should work just fine
- Lengthscale in the Exponential kernel is how strong patterns correlate across months
- I choose 1200, century long patterns

Decay Visualized

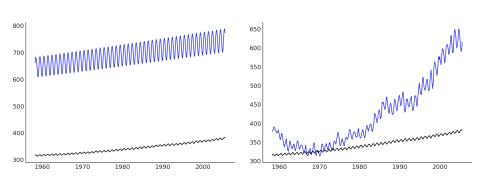


Figure: Mean without decay

Figure: Mean with decay

Decay model

No decay model is the limit case of decay model with lengthscale $ightarrow \infty$

```
g \sim \text{Normal}(1.00065, 0.0005)
         \epsilon \sim \text{LogNormal}(-2, 0.5)
  \log y_0 \sim \text{Normal}(\log 310, 10/310)
       \alpha_k \sim \text{LogNormal}(-2, 0.5)
K(*,*) = \alpha_k Periodic(p = 12, l = 3) · Exponential(l = 1200)
  \log \bar{\mathbf{y}}_t = \mathcal{GP}(\log \mathbf{y}_0 + \mathbf{g} \cdot \mathbf{t}, K)
        v_t \sim \text{LogNormal}(\log \bar{v}_t, \epsilon)
```

Improve the model

- **1** Figure out a better ϵ prior
- 2 Infer better lengthscales (hyperpriors)
- Infer structural medium term shift (additive kernel)

Prior Predictive Check

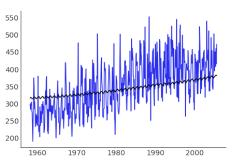


Figure: $\epsilon \sim \text{LogNormal}(-2, 0.5)$

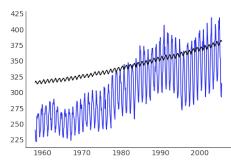


Figure: $\epsilon \sim \text{LogNormal}(-4, 0.5)$

Inference

- pm.sample() is slow
- GPU sampling is also slow
- pm.find_MAP() makes sense and fast

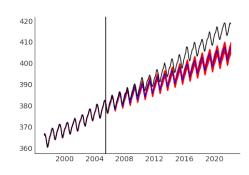


Figure: Predictions with MAP

Key takeaways

- Look at prior predictive to validate initial model
- Interpretable building blocks create interpretable model
- Sasonality model is done with Periodic kernel
- Seasonality might be flexible

Stochastic Volatility Model

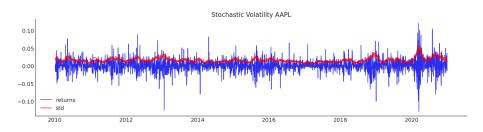


Figure: Stochastic Volatility Estimation for AAPL

Read More

https://github.com/quantopian/bayesalpha

Key Ideas

- Returns are constant
- Volatility is not constant
- Model volatility as a Gaussian process
- Use approximations to speed up the model
- Use GPU to do fast inference

Returns

Volatility

- Returns
 - \bullet Expect year return at orders $\pm 100\%$

Volatility

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 - Daily return $\pm (2^{\frac{1}{250}} 1)$
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- Volatility
 - \bullet Expect signal to noise $\frac{\text{mean}}{\text{std}} \approx 1$
 - $\log \text{std} = \log(2^{\frac{1}{250}} 1)$
- Time Component
 - After the model is framed

$$\begin{split} & \mathsf{return} \sim \mathsf{Normal}(0, 2^{\frac{1}{250}} - 1) \\ & \mathsf{log}\,\mathsf{std} \sim \mathsf{Normal}(\mathsf{log}(2^{\frac{1}{250}} - 1), 0.05) \\ & \mathsf{ls} \sim \mathsf{Gamma}(30, 5) \\ & \alpha_{\mathsf{vol}} \sim \mathsf{Exponential}(100) \\ & \mathcal{K}(*, *) = \alpha_{\mathsf{vol}}\,\mathsf{Matern32}(\mathsf{ls}) \\ & \Delta_t^{\mathsf{log}\,\mathsf{std}} \sim \mathcal{GP}(0, \mathcal{K}) \\ & \mathsf{obs}_t \sim \mathsf{Normal}(\mathsf{return}, \mathsf{exp}(\mathsf{log}\,\mathsf{std} + \Delta_t^{\mathsf{log}\,\mathsf{std}})) \end{split}$$

$$\begin{array}{l} {\sf return} \sim {\sf Normal}(0,2^{\frac{1}{250}}-1) \\ {\sf log\,std} \sim {\sf Normal}({\sf log}(2^{\frac{1}{250}}-1),0.05) \\ {\sf ls} \sim {\sf Gamma}(30,5) \\ {\alpha_{\sf vol}} \sim {\sf Exponential}(100) \\ {\cal K}(*,*) = {\alpha_{\sf vol}} \, {\sf Matern32}({\sf ls}) \\ {\Delta_t^{\sf log\,std}} \sim {\cal GP}(0,{\cal K}) \\ {\sf obs}_t \sim {\sf Normal}({\sf return}, {\sf exp}({\sf log\,std}+\Delta_t^{\sf log\,std})) \end{array}$$

$$\begin{split} & \mathsf{return} \sim \mathsf{Normal}(0, 2^{\frac{1}{250}} - 1) \\ & \mathsf{log}\,\mathsf{std} \sim \mathsf{Normal}(\mathsf{log}(2^{\frac{1}{250}} - 1), 0.05) \\ & \mathsf{ls} \sim \mathsf{Gamma}(30, 5) \\ & \alpha_{\mathsf{vol}} \sim \mathsf{Exponential}(100) \\ & \mathcal{K}(*, *) = \alpha_{\mathsf{vol}}\,\mathsf{Matern32}(\mathsf{ls}) \\ & \Delta_t^{\mathsf{log}\,\mathsf{std}} \sim \mathcal{GP}(0, \mathcal{K}) \\ & \mathsf{obs}_t \sim \mathsf{Normal}(\mathsf{return}, \mathsf{exp}(\mathsf{log}\,\mathsf{std} + \Delta_t^{\mathsf{log}\,\mathsf{std}})) \end{split}$$

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Splines

Make the model faster!

Spline interpolation

pmx.utils.spline.bspline_interpolation

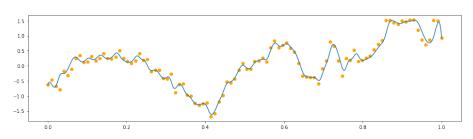


Figure: B-Spline interpolation

$$y' = B(x, x', d)y,$$

where $\{x, y\}$ original function, d - degree, x' - new domain, y' interpolated function

Spline interpolation

pmx.utils.spline.bspline_interpolation

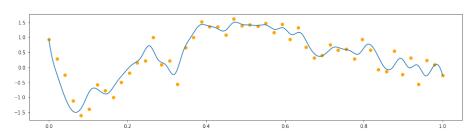


Figure: B-Spline interpolation (fewer points)

$$y' = B(x, x', d)y,$$

where $\{x, y\}$ original function, d - degree, x' - new domain, y' interpolated function

Why B-Spline

$$y'=B(x,x',d)y,$$

- $\mathbf{1}$ x, x' are usually given
- 2 y is random (Gaussian Process)
- 3 y' is a smooth interpolation
- **4** B(x, x', d) is static
- **5** B(x, x', d) is sparse

To remember

1 Bad extrapolation when degree d > 2

Spline Interpolation

- ullet d=1 linear interpolation
- d = 2 quadratic interpolation
- d = 3 cubic interpolation

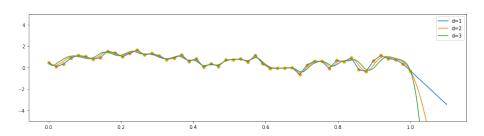


Figure: Spline interpolation

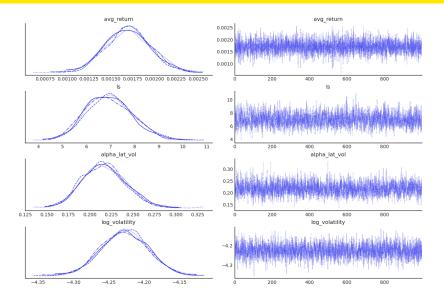
Splines for Stochastic volatility

$$\begin{split} & \mathsf{return} \sim \mathsf{Normal}(0, 2^{\frac{1}{250}} - 1) \\ & \mathsf{log}\,\mathsf{std} \sim \mathsf{Normal}(2^{\frac{1}{250}} - 1, 0.05) \\ & \mathsf{ls} \sim \mathsf{Gamma}(30, 5) \\ & \alpha_{\mathsf{vol}} \sim \mathsf{Exponential}(100) \\ & \mathcal{K}(*, *) = \alpha_{\mathsf{vol}}\,\mathsf{Matern32}(\mathsf{ls}) \\ & \bar{\Delta}_{t'}^{\mathsf{log}\,\mathsf{std}} \sim \mathcal{GP}(0, \mathcal{K}) \\ & \Delta_{t}^{\mathsf{log}\,\mathsf{std}} = \mathsf{B-Spline}(\bar{\Delta}_{t'}^{\mathsf{log}\,\mathsf{std}}, \mathsf{d} = 3) \\ & \mathsf{obs}_{t} \sim \mathsf{Normal}(\mathsf{return}, \mathsf{exp}(\mathsf{log}\,\mathsf{std} + \Delta_{t}^{\mathsf{log}\,\mathsf{std}})) \end{split}$$

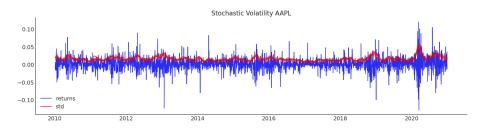
Splines for Stochastic volatility

```
return \sim \text{Normal}(0, 2^{\frac{1}{250}} - 1)
 \log \text{ std} \sim \text{Normal}(2^{\frac{1}{250}} - 1, 0.05)
             Is \sim Gamma(30, 5)
       \alpha_{\rm vol} \sim {\sf Exponential}(100)
K(*,*) = \alpha_{\text{vol}} \text{ Matern32(ls)}
\bar{\Delta}_{\star\prime}^{\log \text{std}} \sim \mathcal{GP}(0,K)
\Delta_{\star}^{\log \text{ std}} = \text{B-Spline}(\bar{\Delta}_{\star}^{\log \text{ std}}, d = 3)
      \mathsf{obs}_t \sim \mathsf{Normal}(\mathsf{return}, \mathsf{exp}(\mathsf{log}\,\mathsf{std} + \Delta_t^{\mathsf{log}\,\mathsf{std}}))
```

Results



Takeaways



- Interpretable parameters
- Splines make model much faster
- Uncertainties for
 - average volatility
 - average returns
 - stochastic variation (α_{vol})

References I



T. SJ and L. B. Forecasting at scale. 2017.