# Genotype and SNP calling and estimation of allele frequencies

Matteo Fumagalli

We are bioinformaticians thats what we do Sample preparation Sequencing Rawdata Gene identification Novel genes Discoveries...etc http://biocomicals.blogspot.com

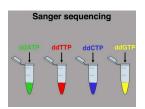
# Presentation outline

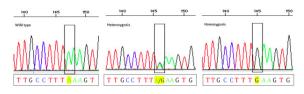
- Motivation
- 2 Genotype likelihoods
- Genotype calling
- 4 SNP calling
- 5 Imputation



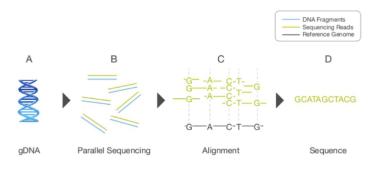
# Sanger sequencing

aka first/former generation sequencing





# **Next Generation Sequencing**



A. Extracted gDNA

B. gDNA is fragmented into a library of small segments that are each sequenced in parallel.

C. Individual sequence reads are reassembled by aligning to a reference genome

D. The whole-genome sequence is derived from the consensus of aligned reads.

# From genomes to variants

# Genome (FASTA)

TAATCCGCACGCTTTAGACTCCCCGGCTGTGATTTTTTGACAATGGCTCGGGGTTCTGCAAAGCGGGCCCTG
TCTGGGGAGTTTGGACCCCGGCACATGGTCAGCTCCATCGTGGGCACCTGAAATTCCAGGCTCCCTCAG

# EASTO)

# Reads (FASTQ) CCAATGATTTTTTCCGTGTTTCAGAATACGGTTAA +SRR038845.41 HWI-EAS038:6:1:0:1474

+SRR038845.41 HWI-EA5038:6:1:0:1474 length=36 BCCBA@BB@BBBBBBBBBBBBBBBBBAB&A:@693:@B= @SRR038845.53 HWI-EA5038:6:1:1:360 length=36 GTTCAAAAAGACTAAATTCTGTCAATAGAAAACTC +SRR038845.53 HWI-EA5038:6:1:1:360 length=36

# Mapped Reads (mpileup, BAM)

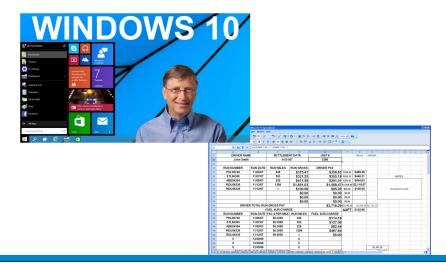
seq1				,.\$,,.,.,
seq1	273	T	23	,,
seql	274	T	23	,.\$,
seq1	275	Α	23	,\$,1
seq1	276	G	22	T,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,
segl	277	T	22	,,,,,,C.,,,,,G. +7<;<<<<<&<<<<<<<<<<<<<<<<<<
seq1	278	G	23	,^k. \\$38*<<;<7<<7<=<<;<<<<
seg1				AT,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,,

### Variants (VCF)

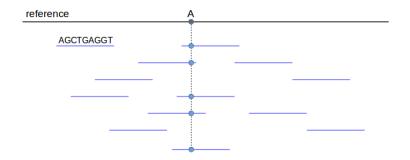
		٠,		,						
	ormat=VC									
	ate=2014									
##sourc	e=23andn	e2vcf.pl	https	://githul	.com/arr	ogantrobo	t/23and	lne2vcf		
##refer	ence=fil	e://23and	ne_v3	hg19_re	f.txt.gz					
##FORMA	T= <id=gt< td=""><td>,Number=1</td><td>,Type</td><td>String,</td><td>Descripti</td><td>on="Genot</td><td>ype"&gt;</td><td></td><td></td><td></td></id=gt<>	,Number=1	,Type	String,	Descripti	on="Genot	ype">			
#CHROM	POS	ID	REF	ALT	QUAL	FILTER	INFO	FORMAT	GENOT	YPE
chr1	82154	rs447721	2	a					GT	•
/0										
chr1	752566	rs309431	5	g	A				GT	- 1
/1										
chr1	752721	rs313197	2	A	G				GT	- 1
/1										
chr1	798959	rs112407	77	9					GT	•
/0										
chr1	800007	rs668104	9	T	C				GT	1
/1										



# Forget about

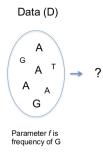


# Imperial College London Why do we need statistics?

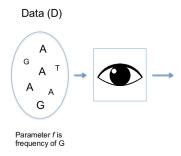


• is a nucleotide/base/allele with a certain quality score

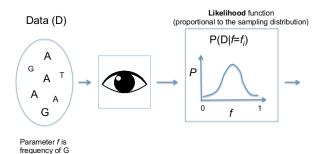
# Imperial College London Statistical inference



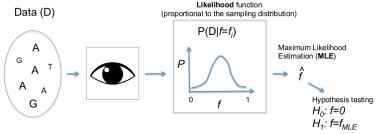
# Imperial College London Statistical inference



# Imperial College London Statistical inference



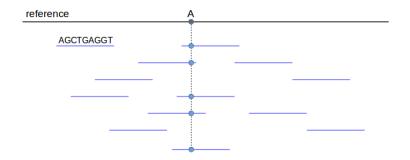
# Statistical inference



### Likelihood approach:

- All the information on the parameter is in the likelihood function (we use all the data!).
- · More data leads to less bias and less variance.
- · Suitable for hypothesis testing.

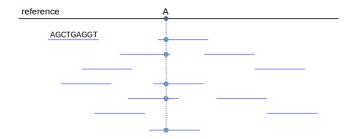
# That is why we need statistics!



• is a nucleotide/base/allele with a certain quality score

# Genotype likelihoods

Given a possible genotype, what is the probability of observing this NGS data?



• is a nucleotide/base/allele with a certain quality score

# Genotype likelihoods - equation

# Likelihood

```
P(D|G = \{A_1, A_2, ..., A_n\}) with
```

 $A_i \in \{A, C, G, T\}$  and n being the ploidy level

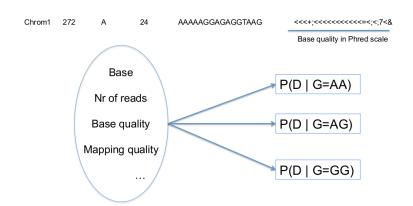
# Genotype likelihoods - equation

### Likelihood

$$P(D|G = \{A_1, A_2, ..., A_n\})$$
 with  $A_i \in \{A, C, G, T\}$  and  $n$  being the ploidy level

How many genotypes likelihoods do we need to calculate for each each diploid individual at each site?

# Genotype likelihoods - rationale



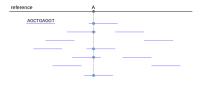
# Genotype likelihoods - calculation

### Likelihood function

$$P(D|G = \{A_1, A_2, ..., A_N\}) = \prod_{i=1}^{R} \sum_{j=1}^{N} \frac{L_{A_j, i}}{N}$$

- $\bullet \ L_{A_j,i} = P(D|A_G = A_j)$
- $A_i \in \{A, C, G, T\}$
- R is the depth (nr. of reads)
- N is the ploidy level (nr. of chromosomal copies)

# Genotype likelihoods - example

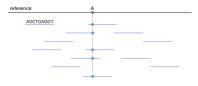


• is a nucleotide/base/allele with a certain quality score

A

with all with quality scores equal to 20 (in phred score)

# Genotype likelihoods - example



• is a nucleotide/base/allele with a certain quality score

Α

Α

Α

J

with all with quality scores equal to 20 (in phred score)

What is 
$$P(D|G = AC) = ?$$

# Genotype likelihoods - example

### Likelihood function

$$P(D|G = \{A_1, A_2, ..., A_N\}) = \prod_{i=1}^{R} \sum_{j=1}^{N} \frac{L_{A_j, i}}{N}$$

A

А

۸

G

& Q=20

$$P(D|G = \{A, C\}) = ...$$

Α

# Genotype likelihoods - example

# Likelihood function

$$P(D|G = \{A_1, A_2, ..., A_N\}) = \prod_{i=1}^{R} \sum_{j=1}^{N} \frac{L_{A_j, i}}{N}$$

```
A

A

G

& Q=20

N = 2; i = 1; A_1 = A; A_2 = C
```

$$P(D|G = \{A, C\}) = (\frac{L_{A,1}}{2} + \frac{L_{C,1}}{2}) \times ...$$

What are  $L_{A,1}$  and  $L_{C,1}$ ?

# Genotype likelihoods - example

# Likelihood function

$$P(D|G = \{A_1, A_2, ..., A_N\}) = \prod_{i=1}^{R} \sum_{j=1}^{N} \frac{L_{A_j, i}}{N}$$

# Genotype likelihoods - example

# Likelihood function

$$P(D|G = \{A_1, A_2, ..., A_N\}) = \prod_{i=1}^{R} \sum_{j=1}^{N} \frac{L_{A_j, i}}{N}$$

$$L_{C,1} = \frac{\epsilon}{3}$$

$$L_{A,1} =$$

# Genotype likelihoods - example

### Likelihood function

$$P(D|G = \{A_1, A_2, ..., A_N\}) = \prod_{i=1}^{R} \sum_{j=1}^{N} \frac{L_{A_j, i}}{N}$$

$$L_{C,1} = \frac{\epsilon}{3}$$

$$L_{A,1} = 1 - \epsilon$$

$$P(D|G = \{A, C\}) = (\frac{1-\epsilon}{2} + \frac{\epsilon}{6}) \times \dots$$

# Genotype likelihoods - example

### Likelihood function

$$P(D|G = \{A_1, A_2, ..., A_N\}) = \prod_{i=1}^R \sum_{j=1}^N \frac{L_{A_j,i}}{N}$$

# Genotype likelihoods - example

# Likelihood function

$$P(D|G = \{A_1, A_2, ..., A_N\}) = \prod_{i=1}^R \sum_{j=1}^N \frac{L_{A_j,i}}{N}$$

$$L_{C,4} = \frac{\epsilon}{3}$$

$$L_{A,4} =$$

# Genotype likelihoods - example

# Likelihood function

$$P(D|G = \{A_1, A_2, ..., A_N\}) = \prod_{i=1}^R \sum_{i=1}^N \frac{L_{A_i, i}}{N}$$

$$L_{C,4} = \frac{\epsilon}{3}$$

$$L_{A,4}=\frac{\epsilon}{3}$$

$$P(D|G = \{A, C\}) = \left(\frac{1-\epsilon}{2} + \frac{\epsilon}{6}\right)^3 \times \frac{\epsilon}{3}$$

# Genotype likelihoods - example

Genotype	Likelihood (log10)	
AA	-2.49	
AC	-3.38	
AG	-1.22	Α
AT	-3.38	Α
CC	-9.91	Α
CG	-7.74	G
CT	-9.91	$\epsilon = 0.01$
GG	-7.44	
GT	-7.74	
TT	-9.91	

# Genotype calling

# Genotype calling

Genotype	Likelihood (log10)
AA	-2.49
AC	-3.38
AG	-1.22
AT	-3.38
CC	-9.91
CG	-7.74
CT	-9.91
GG	-7.44
GT	-7.74
TT	-9.91

AAAG &  $\epsilon = 0.01$ 

What is the genotype here?

# Genotype calling

Genotype	Likelihood (log10)
AA	-2.49
AC	-3.38
AG	-1.22
AT	-3.38
CC	-9.91
CG	-7.74
CT	-9.91
GG	-7.44
GT	-7.74
TT	-9.91

AAAG &  $\epsilon = 0.01$  What is the genotype? AG.

# Maximum Likelihood

The simplest genotype caller: choose the genotype with the highest likelihood.

# Major and minor alleles

# Likelihood function

$$\log P(D|G = A) = \sum_{i=1}^{R} \log L_{A_i,i}$$

AAAG & 
$$\epsilon = 0.01$$

Allele	Likelihood
Α	-2.49
C	-3.38
G	-1.22
Т	-3.38

We can reduce the genotype space to 3 entries (from 10, for diploids).

# Genotype calling

AAAG &  $\epsilon = 0.01$  & A,G alleles

Genotype	Likelihood
AA	-5.73
AG	-2.80
GG	-17.12

At what extent is the data affecting the called genotype and its **confidence**?

Open jupyter-notebook.

## Genotype likelihood ratio

$$\log_{10} \frac{L_{G(1)}}{L_{G(2)}} > t$$

i.e. t=1 meaning that the most likely genotype is 10 times more likely than the second most likely one Pros and cons?

Yes:

## Genotype likelihood ratio

$$\log_{10}\frac{L_{G(1)}}{L_{G(2)}} > t$$

i.e. t=1 meaning that the most likely genotype is 10 times more likely than the second most likely one Pros and cons?

- Yes: genotype are called with higher confidence
- No:

# Genotype likelihood ratio

$$\log_{10} \frac{L_{G(1)}}{L_{G(2)}} > t$$

i.e. t=1 meaning that the most likely genotype is 10 times more likely than the second most likely one Pros and cons?

- Yes: genotype are called with higher confidence
- No: more missing data

Practical: genotype likelihoods and (basic) genotype calling https://github.com/mfumagalli/Copenhagen

### The monster dilemma



Figure 1: Nessie, the Loch Ness Monster. True or fake news?

### The monster dilemma - likelihood

Let's denote D (data) as the set of observations specifying whether I tell you that I saw Nessie (D=1) or not (D=0).

D is our sample space, the set of all possible outcomes of the experiment, and  $D=\{0,1\}$ .

We want to make some inferences on the probability that Nessie exists, or that it is true that I saw it (her?). Let's denote this probability as N.

- $D = \{0, 1\}$ , whether I tell you I saw Nessie or not.
- $N = \{0, 1\}$ , whether Nessie exists or not.

### The monster dilemma - likelihood

### Questions

- What are p(D = 1|N = 1) and p(D = 1|N = 0)?
- What is a Maximum Likelihood Estimate of N?
- What is a statistical test for N = 1?

### The monster dilemma - likelihood

```
Let's assume that p(D=1|N=0)=0.01 and p(D=1|N=1)=0.90 are valid for each observer I, with I=3. Then the log-likelihood of N=0 is given by \sum_{l=1}^{3} log(p(D=1|N=0))=-6.91 while the log-likelihood of N=1 is given by \sum_{l=1}^{3} log(p(D=1|N=1))=-0.32. With 3 observations of D=1 we obtained a likelihood ratio (LR, of N=1 vs N=0) of 6.59.
```

Does the Loch ness monster exist?

### Imperial College London "Eyes" thinking

What's "wrong"?

Our inference on N, our parameter, is driven solely by our observations, given by our likelihood function.

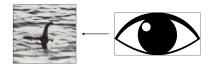


Figure 2: The eye: a "likelihood" organ.

"Blind Brain" thinking
In real life we take many decisions based not only on what we observe but also on some believes of ours\*.

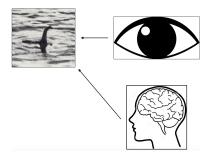


Figure 3: The brain: a "non-likelihood" organ.

\* unfortunately in many cases

# Eyes+Brain thinking

- with "eyes only" our intuition is that  $p(N|D) \approx p(D|N)$
- with "the brain" our intuition is that  $p(N|D) \approx p(D|N)p(N)$

Our "belief" expresses the probability p(N) unconditional of the data.

#### Question

How can we define p(N)?

# "Eyes + Blind Brain"thinking

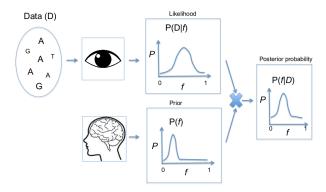
The "belief" function p(N) is called **prior probability** and the joint product of the likelihood p(D|N) and the prior is proportional to the **posterior probability** p(N|D).

The use of posterior probabilities for inferences is called Bayesian statistics.

### Bayesian vs. Likelihoodist

- we derive "proper" probability distributions of our parameters rather than deriving a point estimate;
- a probability is assigned to a hypothesis rather than a hypothesis is tested;
- we can "accept" the null hypothesis rather than "fail to reject" it;
- parsimony imposed in model choice rather than correcting for multiple tests.

# Bayesian inference



# Bayes' Theorem

$$p(G|D) = \frac{f(D|G)\pi(G)}{\int f(D|G)\pi(G)dG}$$

- G is not a fixed parameter but a random quantity with prior distribution  $\pi(G)$
- p(G|D) is the posterior probability distribution of G
- $\int p(G|D)dG = 1$

# Genotype posterior probability

Genotype	Likelihood (log)	Prior	Posterior
AA	-5.73		

# Genotype posterior probability

Genotype	Likelihood (log)	Prior	Posterior
AA	-5.73	1/3	

# Genotype posterior probability

Genotype	Likelihood (log)	Prior	Posterior
AA	-5.73	1/3	0.05
AG	-2.80		•

# Genotype posterior probability

Genotype	Likelihood (log)	Prior	Posterior
AA	-5.73	1/3	0.05
AG	-2.80	1/3	0.95
GG	-17.12		

# Genotype posterior probability

AAAG &  $\epsilon = 0.01$  & A,G alleles

Genotype	Likelihood (log)	Prior	Posterior
AA	-5.73	1/3	0.05
AG	-2.80	1/3	0.95
GG	-17.12	1/3	0

What is the called genotype? What's its confidence?

# Genotype posterior probability

AAAG &  $\epsilon = 0.01$  & A,G alleles

Genotype	Likelihood (log)	Prior	Posterior
AA	-5.73	1/3	0.05
AG	-2.80	1/3	0.95
GG	-17.12	1/3	0

What is the called genotype? What's its confidence? Only call genotypes if the largest probability is above a certain threshold (e.g. 0.95).

## Genotype posterior probability

AAAG & 
$$\epsilon=0.01$$
 & A,G alleles & **A** is the reference allele  $P(AA)>P(AG)>P(GG)$ 

Genotype	Likelihood (log)	Prior	Posterior
AA	-5.73		

## Genotype posterior probability

AAAG &  $\epsilon=0.01$  & A,G alleles & **A** is the reference allele P(AA)>P(AG)>P(GG)

Genotype	Likelihood (log)	Prior	Posterior
AA	-5.73	0.80	0.22
AG	-2.80		•

## Genotype posterior probability

AAAG &  $\epsilon=0.01$  & A,G alleles & **A** is the reference allele P(AA)>P(AG)>P(GG)

Genotype	Likelihood (log)	Prior	Posterior
AA	-5.73	0.80	0.22
AG	-2.80	0.15	0.78
GG	-17.12		

## Genotype posterior probability

AAAG &  $\epsilon = 0.01$  & A,G alleles & **A** is the reference allele P(AA) > P(AG) > P(GG)

Genotype	Likelihood (log)	Prior	Posterior
AA	-5.73	0.80	0.22
AG	-2.80	0.15	0.78
GG	-17.12	0.05	0

Warning: the reference allele is just one of the possible alleles, often chosen arbitrarily: why so much weight???

Genotype posterior probability AAAG &  $\epsilon = 0.01$  & A,G alleles & f(A) = 0.7 from a reference panel

$$P(AA) = ?$$

$$P(AG) = ?$$

$$P(GG) = ?$$

Genotype	Likelihood (log)	Prior	Posterior
AA	-5.73		

Genotype posterior probability AAAG &  $\epsilon = 0.01$  & A,G alleles & f(A) = 0.7 from a reference panel

$$P(AA) = ?$$

$$P(AG) = ?$$

$$P(GG) = ?$$

Genotype	Likelihood (log)	Prior	Posterior
AA	-5.73	0.49	0.06
AG	-2.80		•

Genotype posterior probability AAAG &  $\epsilon = 0.01$  & A,G alleles & f(A) = 0.7 from a reference panel

$$P(AA) = ?$$

$$P(AG) = ?$$

$$P(GG) = ?$$

Genotype	Likelihood (log)	Prior	Posterior
AA	-5.73	0.49	0.06
AG	-2.80	0.42	0.94
GG	-17.12		

Genotype posterior probability AAAG &  $\epsilon = 0.01$  & A,G alleles & f(A) = 0.7 from a reference panel

$$P(AA) = ?$$

$$P(AG) = ?$$

$$P(GG) = ?$$

Genotype	Likelihood (log)	Prior	Posterior
AA	-5.73	0.49	0.06
AG	-2.80	0.42	0.94
GG	-17.12	0.09	0

If the assumption of Hardy Weinberg Equilibrium can be reasonably met.

What happens if that's not the case?

Genotype posterior probability AAAG &  $\epsilon = 0.01$  & A,G alleles & f(A) = 0.7 from a reference panel

$$P(AA) = ?$$

$$P(AG) = ?$$

$$P(GG) = ?$$

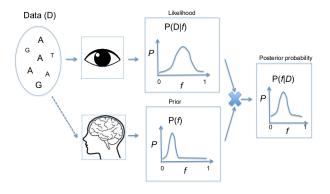
Genotype	Likelihood (log)	Prior	Posterior
AA	-5.73	0.49	0.06
AG	-2.80	0.42	0.94
GG	-17.12	0.09	0

If the assumption of Hardy Weinberg Equilibrium can be reasonably met.

What happens if that's not the case?

Inbreeding can be incorporated:  $f_{AA} = (1 - f)^2 + (1 - f)fF$  ...

# "Eyes + non-Blind Brain" inference



**Empirical Bayesian** 

Genotype posterior probability AAAG &  $\epsilon = 0.01$  & A,G alleles & f(A) = 0.7 from the data itself

Genotype	Likelihood (log)	Prior	Posterior
AA	-5.73		

Genotype posterior probability AAAG &  $\epsilon = 0.01$  & A,G alleles & f(A) = 0.7 from the data itself

Genotype	Likelihood (log)	Prior	Posterior
AA	-5.73	0.49	0.04
AG	-2.80	0.42	0.96
GG	-17.12	0.09	0

- if the assumption of HWE(+-F) can be met (no population structure)
- if enough samples to have a robust estimate of the allele frequencies

Practical: (advanced) genotype calling https://github.com/mfumagalli/Copenhagen

## Genotype posterior probability

AAAG &  $\epsilon = 0.01$  & A,G alleles & f(A) = 0.7 from the data itself

Genotype	Likelihood (log)	Prior	Posterior
AA	-5.73	0.49	0.04
AG	-2.80	0.42	0.96
GG	-17.12	0.09	0

How can we estimate allele frequencies from NGS data?

### SNP calling

# Estimating allele frequencies

Assuming 2 alleles (A,G) with true allele frequency of 0.50

Sample	True genotype	Reads allele A	Read allele G
1	AA	7	0
2	AA	25	1
3	AG	5	3
4	AG	4	4
5	GG	0	2
6	GG	0	4

What is the simplest estimator of allele frequencies?

Estimating allele frequencies
Assuming 2 alleles (A,G) with true allele frequency of 0.50

Sample	True genotype	Reads allele A	Read allele G
1	AA	7	0
2	AA	25	1
3	AG	5	3
4	AG	4	4
5	GG	0	2
6	GG	0	4
Total		41	14

$$\hat{f} = \frac{\sum_{i=1}^{N} n_{A,i}}{\sum_{i=1}^{N} (n_{A,i} + n_{G,i})}$$

 $\hat{f} = 0.75$ 

What is wrong with this estimator?

# Estimating allele frequencies Assuming 2 alleles (A,G) with true allele frequency of 0.50

Sample	True genotype	Reads allele A	Read allele G
1	AA	7	0
2	AA	25	1
3	AG	5	3
4	AG	4	4
5	GG	0	2
6	GG	0	4
Total		41	14

$$\hat{n_A} = \sum_{i=1}^{N} (1 - \epsilon) n_{A,i} + \epsilon n_{G,i} - \epsilon n_{A,i} - (1 - \epsilon) n_{G,i}$$

$$\hat{f} = 0.77$$

# Estimating allele frequencies

### Maximum Likelihood estimator

$$P(D|f) = \prod_{i=1}^{N} \sum_{g \in \{0,1,2\}} P(D|G = g)P(G = g|f)$$

# Estimating allele frequencies

#### Maximum Likelihood estimator

$$P(D|f) = \prod_{i=1}^{N} \sum_{g \in \{0,1,2\}} P(D|G = g)P(G = g|f)$$

P(D|G = g) is the genotype likelihood and P(G = g|f) is given by HWE (for instance).

In our previous example,  $\hat{f}=0.46$  which is much closer to the true value than previous estimators.

### Imperial College London SNP calling

### Challenges

- If high levels of missing data, then genotypes can be lost.
- Rare variants are hard to detect.
- Trade off between false positive and false negative rates.

### How to call SNPs?

- If at least one heterozygous genotype has been called.
- If the estimated allele frequency is above a certain threshold.

### Imperial College London SNP calling

Call a SNP if

$$\hat{f} \geq t$$

where t can be the minimum sample allele frequency detectable (e.g. t = 1/2N with N diploids).

### Likelihood Ratio Test

A Likelihood Ratio Test (LRT) compares the goodness of fit between the null and the alternative model:

- Null model: f = 0
- Alternative model:  $f \neq 0$

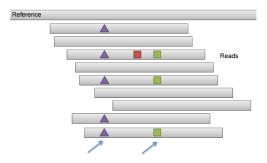
$$T = -2\log\frac{L(f=0)}{L(f=\hat{f}_{MLE})}$$

where T is  $\chi^2$  distributed with 1 degree of freedom.

Practical: allele frequencies and SNP calling https://github.com/mfumagalli/Copenhagen

# SNP calling procedures

Alignment-based caller



We completely rely on how reads have been mapped

Figure from Erik Garrison

# SNP calling procedures

- Assembly-based caller (as in GATK)
- Local re-alignment around putative variants; better resolution for INDELs detection.
- Haplotype-based caller (as in freebayes)

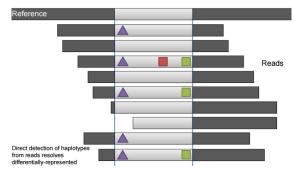
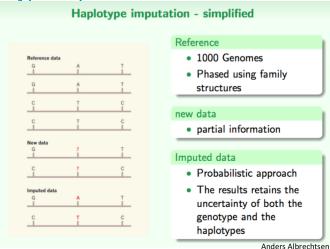


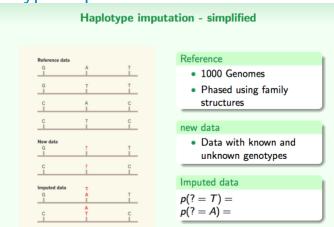
Figure from Erik Garrison



## Haplotype imputation



# Haplotype imputation



## Haplotype imputation





#### Reference

· haplotype frequencies

#### new data

 Data with known and unknown genotypes

### first haplotype

$$p(? = T) = \frac{0.56}{0.56 + 0.03} = 0.95$$
$$p(? = A) = \frac{0.03}{0.56 + 0.03} = 0.05$$

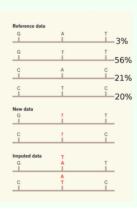
#### second haplotype

$$p(? = T) = \frac{0.21}{0.21 + 0.2} = 0.51$$
  
 $p(? = A) = \frac{0.2}{0.21 + 0.2} = 0.49$ 

Anders Albrechtsen

## Haplotype imputation

### Haplotype imputation - simplified



#### Bayes formula

$$p(H = h|f,G) = \frac{P(G|H=h)P(H=h|f)}{\sum_{h'} P(G|H=h')P(H=h'|f)}$$

#### P(G|H=h)

1 if consistent

0 otherwise

#### first haplotype

$$p(? = T) = \frac{0.56}{0.56 + 0.03} = 0.95$$
  
 $p(? = A) = \frac{0.03}{0.56 + 0.03} = 0.05$