10708 Homework 3

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1 Variational Autoencoders

1.1 Derivations

1.1.1

$$\log P(x) = \int_{z} q_{\phi}(z|x) \log p_{\theta}(x) dz$$

$$= \int_{z} q_{\phi}(z|x) \log \frac{p_{\theta}(x,z)}{p_{\theta}(z|x)}$$

$$= \int_{z} q_{\phi}(z|x) \log \frac{q_{\phi}(z|x)}{p_{\theta}(z|x)} \frac{p_{\theta}(x,z)}{q_{\phi}(z|x)}$$

$$= \int_{z} q_{\phi(z|x)} \log \frac{q_{\phi}(z|x)}{p_{\theta}(z|x)} + \int_{z} q_{\phi(z|x)} \log \frac{p_{\theta}(x,z)}{q_{\phi}(z|x)}$$

$$= KL(q_{\phi}(z|x)||p_{\theta}(z|x)) + \int_{z} q_{\phi(z|x)} \log \frac{p_{\theta}(x,z)}{q_{\phi}(z|x)}$$

$$\geq \int_{z} q_{\phi}(z|x) \log \frac{p_{\theta}(x,z)}{q_{\phi}(z|x)}$$

$$:= L(\theta, \phi; x)$$

1.1.2

In the WAKE phase,

Generate examples from $q_{\phi}(z|x^i)$, from training data x^i .

Use the sample z and input x^i as target to update the generator network parameter θ , i.e. performs one step of gradient ascent update with respect to maximum likelihood.

The **optimization objective** for this phase is:

$$\begin{aligned} \max_{\theta} L(\theta, \phi; x) &= \max_{\theta} \int_{z} q_{\phi}(z|x) \log \frac{p_{\theta}(x, z)}{q_{\phi}(z|x)} \\ &= \max \int_{z} q_{\phi}(z|x) \log p_{\theta}(z|x) - \int_{z} q_{\theta}(z|x) \log q_{\phi}(z|x) \\ &= \max_{\theta} E_{q_{\phi}(z|x)} [\log p_{\theta}(x|z)] \\ &= \max_{\theta} \frac{1}{N} \sum_{i=1}^{N} \log p_{\theta}(x^{i}|z) \end{aligned}$$

In the SLEEP phase,

Start with random hidden variables z^i drawn from prior p(z) in the top layer, then use top-down to generate each following layer. At the end, generate an unbiased sample x^i from the generative model. Then train recognition weights using the data generated $(x^i, z^i)_{i=1}^N$.

The optimization objective for this phase is to maximize

$$F'(\theta, \phi; x) = -\log p(x) + KL(p(z|x)||q_{\phi}(z|x))$$

w.r.t. $q_{\phi}(z|x)$:

$$\max_{\phi} E_{p_{\theta}(z,x)}[\log q_{\phi}(z|x)] = \max_{\phi} \frac{1}{N} \sum_{i=1}^{N} \log q_{\phi}(z^{i}|x^{i})$$

Advantages:

- 1. This algorithm is unsupervised. It does not need any labels, all weights are updated iteratively by the samples generated in the model.
- 2. This algorithm does not require communicating methods that sending error information to all of the connections. Instead, each layer compare the input and the top-down reconstruction, and try to minimize the "description length".

Disadvantages:

- 1. At first few iterations, the data generated might be very different true data, but we still use the generated data to train the recognition weights. This is wasteful.
- 2. The recognition weights update is the gradient of the variational bound on the log probability. This can lead to mode-averaging.
- 3. We are assuming the prior is independent when generating the generative weights in the top layer, but it might not be the case because of explaining away effects.
- 4. This algorithm may not converge.

1.1.3

The stochastic estimate of the ELBO used as the objective: From question 1.1.1,

$$\begin{split} \log p(x) &\geq \int_z q_\phi(z|x) \log \frac{p_\theta(x,z)}{q_\phi(z|x)} \\ &= E_{q_\phi(z|x)} [\log p_\theta(x|z) p(z)] - \int_z q_\phi(z|x) \log q_\phi(z|x) \\ &= E_{q_\phi(z|x)} [\log p_\theta(x|z)] - KL(q_\phi(z|x)||p(z)) \\ &:= L(\theta,\phi;x). \end{split}$$

Optimize L w.r.t. $p_{\theta}(x|z)$ is the same with the wake phase. $\max_{\theta} E_{q_{\phi}(z|x)}[\log p_{\theta}(x|z)]$

Optimize L w.r.t. $q_{\phi}(z|x)$ uses the reparameterization trick.

Make $q_{\phi}(z^i|x^i) = N(z^i; \mu^i, \sigma^{2i}I)$, then $z^{(i,l)} = \mu^i + \sigma^i \epsilon^l$, where $\epsilon \sim N(0, I)$.

Then the update rule for $q_{\phi}(z|x)$ is:

$$L(\theta, \phi; x) = E_{q_{\phi}(z|x)}[\log p_{\theta}(x, z)] - KL(q_{\phi}(z|x)||p(z))$$

= $E_{\epsilon \sim N(0, I)}[\log p_{\theta}(x, z_{\phi}(\epsilon))] - KL(q_{\phi}(z|x)||p(z))$

$$\nabla_{\phi} E_{q_{\phi}(z|x)}[\log p_{\theta}(x,z)] = E_{\epsilon \sim N(0,I)}[\nabla_{\phi} \log p_{\theta}(x,z_{\phi}(\epsilon))]$$

KL distance can be computed and differentiated analytically.

Advantages:

- 1. VAE reparameterization of the variational lower bound yields a simple differentiable unbiased estimator of the lower bound, which is easy to optimize using standard stochastic gradient ascent techniques.
- 2. The Approximate posterior inference is easier because we can use simple ancestral sampling, instead of using expensive iterative inference schemes (such as MCMC) per datapoint.

Disadvantages:

- 1. This model is not applicable to discrete latent variables.
- 2. Each element experience reconstruction error. Also this model is sensitive to irrelevant variance, for examples, translations.
- 3. For the simplicity of inference and learning, usually use a fixed standard normal distribution as prior.

1.1.4

By Jensen's inequality,

$$logp(x) = logE_{z^{i} \sim q_{\phi}(z|x)} \left[\frac{1}{k} \sum_{i=1}^{k} \frac{p_{\theta}(x, z_{i})}{q_{\phi}(z_{i}|x)} \right]$$

$$\geq E_{z^{i} \sim q_{\phi}(z|x)} \left[log \frac{1}{k} \sum_{i=1}^{k} \frac{p_{\theta}(x, z_{i})}{q_{\phi}(z_{i}|x)} \right]$$

$$= L_{k}(x)$$

We have shown that $logp(x) \ge L_k(x)$ for any given k > 0, next we have to show that $L_{k+1}(x) \ge L_k(x)$. Let $M \subset \{1,...,k+1\}$ and M has k elements. Then $E_{M=\{m1,...,m_k\}}[\frac{a_{m_1}+...+a_{m_k}}{k}] = \frac{a_1+...+a_{k+1}}{k+1}$. By Jensen's inequality, we got the following:

$$\begin{split} L_{k+1}(x) &= E_{z^1, \dots z^{k+1} \sim q_{\phi}(z|x)}[\log \frac{1}{k+1} \sum_{i=1}^{k+1} \frac{p_{\theta(x, z_i)}}{q_{\phi}(z^i|x)}] \\ &= E_{z^1, \dots z^{k+1} \sim q_{\phi}(z|x)}[\log E_{m^1, \dots m^k} \frac{1}{k} \sum_{j=1}^k \frac{p_{\theta(x, z_{m^i})}}{q_{\phi}(z^{m^i}|x)}] \\ &\geq E_{z^1, \dots z^{k+1} \sim q_{\phi}(z|x)}[E_{m^1, \dots m^k}[\log \frac{1}{k} \sum_{j=1}^k \frac{p_{\theta(x, z_{m^i})}}{q_{\phi}(z^{m^i}|x)}]] \\ &= E_{z^1, \dots z^k \sim q_{\phi}(z|x)}[\log \frac{1}{k} \sum_{i=1}^k \frac{p_{\theta(x, z_i)}}{q_{\phi}(z^i|x)}] \\ &= L_k(x) \end{split}$$

Combining above two parts, we can get $log p(x) \ge L_{k+1}(x) \ge L_k(x)$.

1.1.5

In order for $L_k(x) \to \log p(x)$ as $k \to \infty$, $\frac{p_{\theta}(x,z^i)}{q_{\phi}(z^i|x)}$ has to be bounded. $L_{k\to\infty}(x) < \log p(x)$

2 Markov Chain Monte Carlo

2.1 Metropolis-Hastings

The proposal distribution I choose: $p(x'|x) = N(x'-x; 0, \sigma^2 I)$. The acceptance can be calculated as:

$$\begin{split} A(x'|x) &= min(1, \frac{P(x')Q(x|x')}{P(x)Q(x'|x)}) \\ &= min(1, \frac{P(x')N(x-x'; 0, \sigma^2 I)}{P(x)N(x'-x; 0, \sigma^2 I)}) \\ &= min(1, \frac{P(x')}{P(x)}) \end{split}$$

The algorithm:

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\begin{array}{l} \text{Initialize starting state } x^0, \, \text{set } \mathbf{t} = \mathbf{0}; \\ \textbf{while } samples \, have \, not \, converged \, \textbf{do} \\ & x = x^t, \, \mathbf{t} = \mathbf{t} + \mathbf{1}; \\ & \text{sample } x^* \sim Q(x^*|x) \\ & \text{sample } u \sim Uniform(\mathbf{0},\mathbf{1}); \\ & \textbf{if } u < A(x^*|x) = min(\mathbf{1},\frac{P(x')}{P(x)}) \, \textbf{then} \\ & \mid x^t = x^* \\ & \textbf{else} \\ & \mid x^t = x. \\ & \textbf{end} \\ & \textbf{end} \end{array}
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2.2 Hamiltonian MCMC

2.2.1

$$H(q,p) = \frac{1}{Z} \exp(-U(q)/T) \exp(-K(p)/T),$$

where $U(q) = -log[\pi(q)]$ and $K(p) = \sum_{i=1}^{d} \frac{p_i^2}{2}$. where $\pi(q)$ indicates the mixture Gaussian model - $\sum_{i=1}^{m} \pi_i N(x; \mu_i, \Sigma_i)$. P are assumed to be independent standard Gaussians.

2.2.2

$$\begin{split} \frac{\partial U(q)}{\partial q} &= \frac{\partial - \log \left(\sum_{i=1}^{m} \pi_{i} N(q; \mu_{i}, \Sigma_{i}) \right)}{\partial q} \\ &= -\frac{1}{\sum_{i=1}^{m} \pi_{i} N(q; \mu_{i}, \Sigma_{i})} \frac{\partial \sum_{i=1}^{m} \pi_{i} N(q; \mu_{i}, \Sigma_{i})}{\partial q} \\ &= -\frac{1}{\sum_{i=1}^{m} \pi_{i} N(q; \mu_{i}, \Sigma_{i})} \frac{\sum_{i=1}^{m} \partial \pi_{i} N(q; \mu_{i}, \Sigma_{i})}{\partial q} \\ &= -\frac{1}{\sum_{i=1}^{m} \pi_{i} N(q; \mu_{i}, \Sigma_{i})} \frac{\sum_{i=1}^{m} \pi_{i} \partial N(q; \mu_{i}, \Sigma_{i})}{\partial q} \end{split}$$

Take derivative of a normal distribution with respect to q,

$$\begin{split} \frac{\partial N(q;\mu_i,\Sigma_i)}{\partial q} &= \frac{\partial \frac{1}{\sqrt{2\pi}\sigma_i} \exp\left(-\frac{(q-\mu_i)^2}{2\sigma_i^2}\right)}{\partial q} \\ &= \frac{1}{\sqrt{2\pi}\sigma_i} \frac{\partial \exp\left(-\frac{(q-\mu_i)^2}{2\sigma_i^2}\right)}{\partial q} \\ &= \frac{1}{\sqrt{2\pi}\sigma_i} \exp\left(-\frac{(q-\mu_i)^2}{2\sigma_i^2}\right) \frac{\partial -\frac{(q-\mu_i)^2}{2\sigma_i^2}}{\partial q} \\ &= \frac{1}{\sqrt{2\pi}\sigma_i} \exp\left(-\frac{(q-\mu_i)^2}{2\sigma_i^2}\right) \left(-\frac{1}{2\sigma_i^2}\right) \frac{\partial ((q-\mu_i)^2)}{\partial q} \\ &= \frac{1}{\sqrt{2\pi}\sigma_i} \exp\left(-\frac{(x-\mu_i)^2}{2\sigma_i^2}\right) \left(-\frac{1}{2\sigma_i^2}\right) 2(q-\mu_i) \\ &= \left(-\frac{q-\mu_i}{\sqrt{2\pi}\sigma_i^3}\right) \exp\left(-\frac{(q-\mu_i)^2}{2\sigma_i^2}\right) \end{split}$$

Plug the derivative of normal distribution back to the original derivative, we get

$$\frac{\partial U(q)}{\partial q} = -\frac{\sum_{i=1}^{m} \pi_i \left(-\frac{q - \mu_i}{\sqrt{2\pi}\sigma_i^3}\right) \exp\left(-\frac{(q - \mu_i)^2}{2\sigma_i^2}\right)}{\sum_{i=1}^{m} \pi_i N(q; \mu_i, \Sigma_i)}$$

$$= \frac{\sum_{i=1}^{m} \pi_i \frac{q - \mu_i}{\sqrt{2\pi}\sigma_i^3} \exp\left(-\frac{(q - \mu_i)^2}{2\sigma_i^2}\right)}{\sum_{i=1}^{m} \pi_i \frac{1}{\sqrt{2\pi}\sigma_i} \exp\left(-\frac{(q - \mu_i)^2}{2\sigma_i^2}\right)}$$

2.2.3

The algorithm leap frog is given by the following:

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\begin{array}{l} q = currentq \\ p = random \ normal(shape(q), \ mean=0, \ stddev=1) \\ currentp = p \\ p = p \ - \ epsilon \ * \ gradient \ of \ U(q) \ / \ 2 \\ \textbf{for} \ i \ in \ 1:L \ \textbf{do} \\ q = q \ + \ epsilon \ * \ p \\ \textbf{if} \ i \ !=L \ \textbf{then} \\ p = p \ - \ epsilon \ * \ gradient \ of \ U(q) \\ \textbf{end if} \\ \textbf{end for} \\ p = p \ - \ epsilon \ * \ gradient \ of \ U(q) \ / \ 2 \end{array}
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2.3 Effective sample size