

8: Eksperimenter II

Videregående kvantitative metoder i studiet af politisk adfærd

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- 1 Formalia
- 2 Opsamling fra sidst
- 3 Clustering
- 4 Kovariater
 - Præ-treatment outcome
 - Andre kovariater
 - Blocking
- 5 Noncompliance
- 6 Case: Gerber & Green (2000)
- 7 Kig fremad

Fagets opbygning

Blok 1

Gang	Tema	Litteratur	Case
1	Introduktion til R	Leeper (2016)	
2	R workshop + tidy data	Wickham (2014), Zhang (2017)	
3	Regression I: OLS brush-up	AP kap 3	Newman et al. (2015), Solt et al. (2017)
4	Regression II: Paneldata	AGS kap 4	Larsen et al. (2017)

Fagets opbygning

Blok 2

5	Introduktion til kausal inferens	Hariri (2012), Samii (2016)	Eckles & Bakshy (2017)
6	Matching	Justesen & Klemmensen (2014)	Nall (2015)
<i>Efterårsferie</i>			
7	Eksperimenter I	AP kap 1+2, GG kap 1+2	Gerber, Green & Larimer (2008)
8	Eksperimenter II	GG kap 3+4+5	Gerber & Green (2000)
9	Instrumentvariable	AP kap 4	Lundborg et al. (2017)
10	Difference-in-differences	AP kap 5	Enos (2016)
11	Regressionsdiskontinuitetsdesigns	AP kap 6	Eggers & Hainmueller (2009)

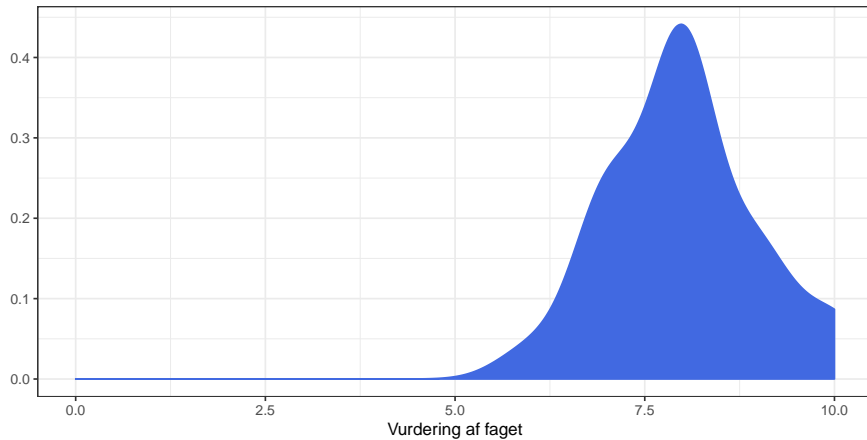
Fagets opbygning

Blok 3

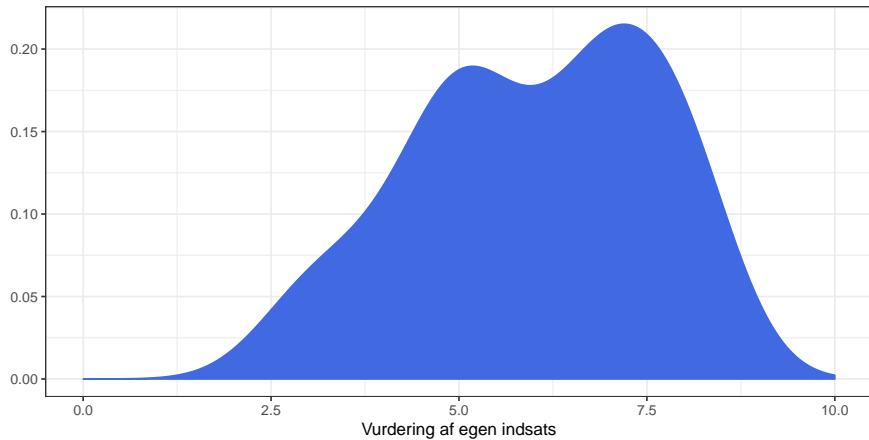
12	Tekst som data	Grimmer & Stewart (2013), Benoit & Nulty (2016)	Baturo & Mikhaylov (2013)
13	Scraping af data fra online-kilder	MRMN kap 9+14	Hjorth (2016)
14	'Big data' og maskinlæring	Varian (2014), Montgomery & Olivella (2017)	Theocharis et al. (2016)

- mere om potential outcomes framework
- randomisering i praksis: `sample()`, `set.seed()`
- forudsætninger eksperimentelle designs: excludability, noninterference
- → med randomisering, excludability og noninterference er den estimerede ATE i stikprøven = den sande ATE (in expectation!)
- case: GGL (2008)

Midtvejsevaluering I



Midtvejsevaluering II



Midtvejsevaluering II

Gode ting:

- generel ros
- underviserengagement
- velvalgt pensum
- intro til R i begyndelsen

Mindre gode ting:

- mere R-øvelse ønskes
- flere øvelsestimer
- matematikken er for teknisk
- materialetrængsel (→ tidspres)

Andet:

- "Nææ - håber, du får billetter! fuck Irland"

Formlen for \widehat{ATE} 's standardfejl:

$$SE(\widehat{ATE}) = \sqrt{\frac{1}{N-1} \left\{ \frac{m \text{Var}(Y_{i0})}{N-m} + \frac{(N-m) \text{Var}(Y_{i1})}{m} + 2 \text{Cov}(Y_{i0} Y_{i1}) \right\}} \quad (1)$$

Kilder til mindre standardfejl:

- ① $N \uparrow$
- ② $\text{Var}(Y_{i0})$ eller $\text{Var}(Y_{i1}) \downarrow$
- ③ $\text{Cov}(Y_{i0} Y_{i1}) \downarrow$
- ④ m balanceret / allokeret til potential outcome med størst varians

I nogle situationer måles *outcome* individuelt, men *assignment* sker på cluster-niveau

- fx. mediemarkeder, kommuner, klasseværelser
- grundlæggende implikation: ingen bias, men svækket præcision

\widehat{ATE} 's standardfejl med k clusters:

$$SE(\widehat{ATE}) = \sqrt{\frac{1}{k-1} \left\{ \frac{m \text{Var}(\bar{Y}_{j0})}{N-m} + \frac{(N-m) \text{Var}(\bar{Y}_{j1})}{m} + 2 \text{Cov}(\bar{Y}_{j0} \bar{Y}_{j1}) \right\}} \quad (2)$$

Kilder til mindre standardfejl:

- $k \uparrow$
- $\text{Var}(\bar{Y}_{j0})$ eller $\text{Var}(\bar{Y}_{j1}) \downarrow$

p-værdier kan (med eller uden cluster assignment) beregnes med *randomiseringsinferens*

- simulér alle tænkelige assignments af treatment
- for hver hypotetisk assignment, estimér ATE
- beregn p-værdi pba. faktisk ATE ift. fordelingen af estimater

Særlig type kovariat: præ-treatment observationer på outcome

- outcome kan måles som *ændring* præ-post
- i st. for difference-in-means, *difference-in-differences* estimator (jf. u. 10)
- lad X_i være præ-treatment observation af Y_i ; diff-in-diffs estimator mere efficient når

$$\frac{\text{Cov}(Y_{i0}, X_i)}{\text{Var}(X_i)} + \frac{\text{Cov}(Y_{i1}, X_i)}{\text{Var}(X_i)} > 1 \quad (3)$$

- m.a.o., stærkt øget præcision når præ-treatment kovariater korrelerer m. potential outcomes

Regression af Y_i på d_i og kovariat X_i :

$$Y_i = Y_{i0}(1 - d_i) + Y_{i1}d_i = a + bd_i + cX_i + (u_i - cX_i) \quad (4)$$

→ prædiktive kovariater reducerer residualer → $\sigma_{\hat{b}} \downarrow$

Men: øger også ‘researcher degrees of freedom’

» This type of analysis introduces an element of discretion in terms of what results are reported. Perhaps unconsciously, the researcher may settle on a regression model that makes the estimated ATE look impressive or interesting, a decision rule that jeopardizes the unbiasedness of the estimator.« (105)

Table 1. Explaining support for socially protective policies with physiological reactions to threatening images. Results of ordinary least squares (OLS) regression with support for socially protective policies (possible range from 0 to 18), with higher numbers indicating attitudes more supportive of policies thought to protect the social unit regressed on five explanatory variables: gender (0 = male; 1 = female), age (in years), education (six categories ranging from “did not finish high school” to “college degree plus”), income (six categories ranging from an annual salary of less than \$20,000 to an annual salary of more than \$100,000), and changes in skin conductance level (SCL) occasioned by the viewing of threatening images. Descriptive statistics on the variables and further discussion of the regression techniques are available in the SOM. * $P < 0.05$, two-tailed t test.

Variable	Unstandardized coefficient (SE)	Standardized coefficient
SCL	92.2* (29.03)	0.377
Income	-0.395 (0.471)	-0.10
Education	-1.63* (0.465)	-0.42
Age	0.19 (0.10)	0.235
Gender	-2.34 (1.3)	-0.20
Constant	-353* (193)	
N	46	
Adj. R-square	0.37	

Table 2. Explaining support for socially protective policies with physiological reactions to nonthreatening images. Results of regression (OLS) with support for socially protective policies regressed on five explanatory variables. Variables are the same as those described for Table 1 except that skin conductance (SCL) is the change in skin conductance occasioned by the viewing of nonthreatening images. Descriptive statistics and further discussion of the regression techniques are available in the SOM. * $P < 0.05$, two-tailed t test.

Variable	Unstandardized coefficient (SE)	Standardized coefficient
SCL	-1.8 (35.08)	-0.007
Income	-0.438 (0.533)	-0.115
Education	-1.57* (0.53)	-0.408
Age	0.165 (0.11)	0.204
Gender	-2.23 (1.52)	-0.196
Constant	-304* (217)	
N	46	
Adj. R-square	0.21	

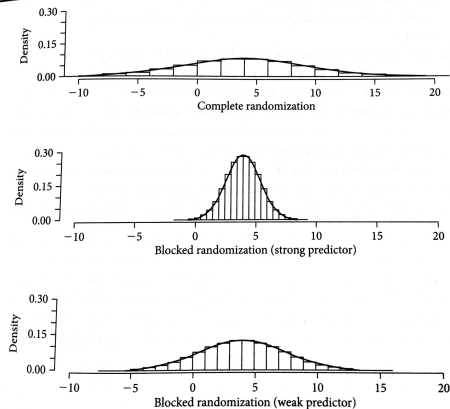
Blocking på kovariat X hjælper når:

- N er relativt lille
- X er ubalanceret
- X er stærk prädiktor for Y

God pakke til block random assignment:
`randomizr`

FIGURE 4.2

Comparison of sampling distributions based on completely randomized and block randomized designs



Konceptuelt: to typer subjekter under ensidet noncompliance

- compliers: $d_i(z = 1) = 1$
- never-takers: $d_i(z = 1) = 0$

→ tre grupper:

- ① treated compliers
- ② non-treated compliers
- ③ never-takers

For hvert subjekt i defineres:

$$ITT_{i,D} \equiv d_i(1) - d_i(0) \quad (5)$$

$$ITT_{i,Y} \equiv Y_i(1) - Y_i(0) \quad (6)$$

CACE er forholdet mellem $\overline{ITT_{i,Y}}$ og $\overline{ITT_{i,D}}$:

$$CACE = \frac{ITT}{ITT_D} \quad (7)$$

- $ATE = (2+6+4+2+4+8+3+3+4)/9 = 4$
- $ITT = (2+0+4+2+4+0+3+3+0)/9 = 2$
- $CACE = (2+4+2+4+3+3)/6 = 3$
- øvelse: hvad er ITT_D ?

TABLE 5.1

Hypothetical schedule of potential outcomes assuming one-sided noncompliance

Observation	$Y_i(d=0)$	$Y_i(d=1)$	$d_i(z=0)$	$d_i(z=1)$	Type
1	4	6	0	1	Complier
2	2	8	0	0	Never-Taker
3	1	5	0	1	Complier
4	5	7	0	1	Complier
5	6	10	0	1	Complier
6	2	10	0	0	Never-Taker
7	6	9	0	1	Complier
8	2	5	0	1	Complier
9	5	9	0	0	Never-Taker

Direkte sammenligning af treatede og nontreatede v. noncompliance estimerer flg.:

$$CACE + \{E[Y_i(d=0)|D_i(1)=1] - E[Y_i(d=0)|D_i(1)=0]\}(1 - ITT_D) \quad (8)$$

→ bias hvis compliers og never-takers har uens untreated potential outcomes

FIGURE 2. Picture Side



FIGURE 1. Text Side

**DON'T LET THEIR
SACRIFICE GO TO
WASTE.**

*The whole point of
democracy is that citizens are
active participants in government,
that we have a voice in government.
Your voice starts with your vote.*

**ON NOVEMBER 3RD
REMEMBER YOUR
RIGHTS AND
RESPONSIBILITIES
AS A CITIZEN.**

Remember to Vote.

Vote Form 16000-119
11/01 Revised 11/01
New Mexico, 478 01023802P



Don't

**Remember to Vote.
VOTE on November 3rd.**

11/01/01 11/01/01 11/01/01 11/01/01 11/01/01 11/01/01 11/01/01 11/01/01 11/01/01 11/01/01

»to find the treatment effect, subtract the turnout rate of the control group from the turnout rate of the experimental group and divide this difference by the observed "contact rate," which is 28%. Using this formula, we find that personal contact raises the probability of turnout by 8.7 percentage points« (658)

»Despite limitations, this experiment provides important new clues in the ongoing mystery of why turnout has declined even as the average age and education of the population has risen. A certain segment of the electorate tends not to vote unless encouraged to do so through face-to-face contact. As voter mobilization grows more impersonal, fewer people receive this kind of encouragement. This point is of great practical significance for those who seek to reverse the declining trend in turnout. Many of the recent policy innovations designed to encourage voter participation (e.g., absentee balloting) focus on reducing the costs of voting. Our findings suggest the importance of focusing as well on the personal connection between voters and the electoral process.« (662)

Næste gang:

- instrumentvariable
- pensum: AP kap. 4
- fokus på afsnit 4.1
- case: Lundborg et al.
 - lang, svær AER-artikel
 - men: helt ny dansk topforskning om hot button-issue
 - fokus på hvordan de retfærdiggør instrumentet

Tak for i dag!