

Homework 0

Issued: Saturday 12th September, 2020

Due: Thursday 17th September, 2020

Tips: It is not a formal homework and will not be graded. The primary goal is to help you remember those basic mathematics you have learnt before.

Calculus & Linear Algebra

0.1. (Inner product) If $\mathbf{x} \in \mathbb{R}^n$ is orthogonal to $\mathbf{y} \in \mathbb{R}^n$, please show that

$$\|\mathbf{x} + \mathbf{y}\|^2 = \|\mathbf{x}\|^2 + \|\mathbf{y}\|^2$$

Solution:

$$\begin{aligned}\|\mathbf{x} + \mathbf{y}\|^2 &= (\mathbf{x} + \mathbf{y})^T (\mathbf{x} + \mathbf{y}) \\ &= \mathbf{x}^T \mathbf{x} + \mathbf{y}^T \mathbf{y} + \mathbf{x}^T \mathbf{y} + \mathbf{y}^T \mathbf{x} \\ &= \|\mathbf{x}\|^2 + \|\mathbf{y}\|^2\end{aligned}$$

0.2. (Orthogonal) Please show that $\|\mathbf{Q}\mathbf{x}\|_2 = \|\mathbf{x}\|_2$.

Solution:

$$\begin{aligned}\|\mathbf{Q}\mathbf{x}\|_2 &= (\mathbf{Q}\mathbf{x})^T \mathbf{Q}\mathbf{x} \\ &= \mathbf{x}^T \mathbf{Q}^T \mathbf{Q}\mathbf{x} \\ &= \|\mathbf{x}\|_2\end{aligned}$$

0.3. (Trace) For any matrices $\mathbf{A}, \mathbf{B}, \mathbf{C} \in \mathbb{R}^{n \times n}$, please show that

- (a) $\text{trace}(\mathbf{AB}) = \text{trace}(\mathbf{BA})$
- (b) $\text{trace}(\mathbf{ABC}) = \text{trace}(\mathbf{CAB}) = \text{trace}(\mathbf{BCA})$
- (c) $\nabla_{\mathbf{A}} \text{tr}(\mathbf{AB}) = \mathbf{B}^T$.

Solution:

- (a) From the definition of trace, we know

$$\begin{aligned}\text{trace}(\mathbf{AB}) &= \sum_{i=1}^n \sum_{j=1}^n a_{ij} b_{ji} \\ &= \sum_{j=1}^n \sum_{i=1}^n b_{ji} a_{ij} \\ &= \text{trace}(\mathbf{BA})\end{aligned}$$

(b) Treat \mathbf{AB} as a new matrix, then we have: $\text{trace}((\mathbf{AB})\mathbf{C}) = \text{trace}(\mathbf{C}(\mathbf{AB}))$

(c) We know $\text{trace}(\mathbf{AB}) = \sum_{i=1}^n \sum_{j=1}^n a_{ij}b_{ji}$, for an arbitray element a_{kl} , we have

$$\frac{\partial \text{tr}(\mathbf{AB})}{\partial a_{kl}} = b_{lk}$$

Hence we have

$$\nabla_{\mathbf{A}} \text{tr}(\mathbf{AB}) = \begin{bmatrix} b_{11} & b_{21} & \cdots & b_{n1} \\ b_{12} & b_{22} & \cdots & b_{n2} \\ \vdots & \vdots & \ddots & \vdots \\ b_{1n} & b_{2n} & \cdots & b_{nn} \end{bmatrix} = \mathbf{B}^T$$

0.4. (Eigenthings) Let \mathbf{x} be an eigenvector of a matrix \mathbf{A} with corresponding eigenvalue λ , then

- (a) Show that for any $\gamma \in \mathbb{R}$, the \mathbf{x} is an eigenvector of $\mathbf{A} + \gamma\mathbf{I}$ with eigenvalue $\lambda + \gamma$.
- (b) If \mathbf{A} is invertible, then \mathbf{x} is an eigenvector of \mathbf{A}^{-1} with eigenvalue λ^{-1} .
- (c) $\mathbf{A}^k \mathbf{x} = \lambda^k \mathbf{x}$ for any $k \in \mathbb{Z}$ ($\mathbf{A}^0 = \mathbf{I}$ by definition)

Solution:

(a) We have

$$(\mathbf{A} + \gamma\mathbf{I})\mathbf{x} = \mathbf{A}\mathbf{x} + \gamma\mathbf{x} = (\lambda + \gamma)\mathbf{x}$$

(b) Suppose \mathbf{A} is invertible, then

$$\mathbf{x} = \mathbf{A}^{-1}\mathbf{A}\mathbf{x} = \mathbf{A}^{-1}(\lambda\mathbf{x}) = \lambda\mathbf{A}^{-1}\mathbf{x}$$

such that we have $\mathbf{A}^{-1}\mathbf{x} = \frac{1}{\lambda}\mathbf{x}$.

(c) The case $k > 0$ follows immediately by induction on k , as

$$\begin{aligned} \mathbf{A}\mathbf{x} &= \lambda\mathbf{x} \\ \mathbf{A}^2\mathbf{x} &= \mathbf{A} \cdot \mathbf{A}\mathbf{x} = \lambda^2\mathbf{x} \\ \mathbf{A}^3\mathbf{x} &= \mathbf{A} \cdot \mathbf{A}^2\mathbf{x} = \lambda^3\mathbf{x} \\ &\dots \end{aligned}$$

0.5. (Chain rule) $x \in \mathbb{R}$ is a scalar, we have

$$\begin{aligned} y &= ax + b \\ z &= \frac{1}{1 + e^{-y}} \end{aligned}$$

Please give the $\frac{\partial z}{\partial x}$.

Solution: According to the chain rule, we have

$$\frac{\partial z}{\partial x} = \frac{\partial z}{\partial y} \times \frac{\partial y}{\partial x} = z(1 - z) \times a$$

0.6. (Matrix derivative) $\mathbf{x}, \mathbf{w} \in \mathbb{R}^n$, and $\mathbf{A} \in \mathbb{R}^{n \times n}$. We have $f : \mathbb{R}^n \rightarrow \mathbb{R}$ as

$$f(\mathbf{x}) = \mathbf{x}^T \mathbf{A} \mathbf{x} + \mathbf{w}^T \mathbf{x}$$

Please give the $\nabla_{\mathbf{x}} f(\mathbf{x})$.

Solution: The standard solution is, first, we give the differential of $f(\mathbf{x})$:

$$\begin{aligned} df(\mathbf{x}) &= \sum_{i=1}^n \frac{\partial f(\mathbf{x})}{\partial x_i} dx_i \\ &= \frac{\partial f(\mathbf{x})}{\partial \mathbf{x}}^T d\mathbf{x} \\ &= \text{tr} \left(\frac{\partial f(\mathbf{x})}{\partial \mathbf{x}}^T d\mathbf{x} \right) \end{aligned}$$

Here we use the trace trick, that is, for a scalar a we have $\text{tr}(a) = a$. Then, for the function above we derive its differential

$$\begin{aligned} df(\mathbf{x}) &= d\mathbf{x}^T \mathbf{A} \mathbf{x} + \mathbf{x}^T \mathbf{A} d\mathbf{x} + \mathbf{w}^T d\mathbf{x} \\ &= \text{tr}(d\mathbf{x}^T \mathbf{A} \mathbf{x}) + \text{tr}(\mathbf{x}^T \mathbf{A} d\mathbf{x} + \mathbf{w}^T d\mathbf{x}) \\ &= \text{tr}(\mathbf{x}^T \mathbf{A}^T d\mathbf{x}) + \text{tr}(\mathbf{x}^T \mathbf{A} d\mathbf{x} + \mathbf{w}^T d\mathbf{x}) \\ &= \text{tr}((\mathbf{x}^T \mathbf{A}^T + \mathbf{x}^T \mathbf{A} + \mathbf{w}^T) d\mathbf{x}) \end{aligned}$$

Refer to the above two equations, we have

$$\frac{\partial f(\mathbf{x})}{\partial \mathbf{x}}^T d\mathbf{x} = (\mathbf{x}^T \mathbf{A}^T + \mathbf{x}^T \mathbf{A} + \mathbf{w}^T) d\mathbf{x}$$

which means

$$\frac{\partial f(\mathbf{x})}{\partial \mathbf{x}} = \mathbf{A} \mathbf{x} + \mathbf{A}^T \mathbf{x} + \mathbf{w}$$

Or simply, you can remember the result for convenience

$$\begin{aligned} \frac{\partial \mathbf{x}^T \mathbf{A} \mathbf{x}}{\partial \mathbf{x}} &= (\mathbf{A}^T + \mathbf{A}) \mathbf{x} \\ \frac{\partial \mathbf{w}^T \mathbf{x}}{\partial \mathbf{x}} &= \mathbf{w} \end{aligned}$$

Probability Theory Part

0.7. (Conditional Probability) Explain that $\mathbb{E}[\mathbb{E}[X|Y]] = \mathbb{E}[X]$

Solution:

$$\begin{aligned}
 \mathbb{E}[\mathbb{E}[X|Y]] &= \mathbb{E}[g(Y)] \\
 &= \sum_{y \in \mathcal{Y}} p(Y = y) \cdot \left[\sum_{x \in \mathcal{X}} x \cdot p(X = x|Y = y) \right] \\
 &= \sum_{x \in \mathcal{X}} \sum_{y \in \mathcal{Y}} x \cdot p(X = x, Y = y) \\
 &= \mathbb{E}[X]
 \end{aligned}$$

0.8. (Bayes) A city has a 50% chance to rain everyday and the weather report has a 90% chance to correctly forecast.

You will take an umbrella when the report says it will rain and you have a 50% chance to take an umbrella when the report says it will not rain.

Compute

- (a) the probability of raining when you don't take an umbrella;
- (b) the probability of not raining when you take an umbrella.

Solution: Let's evaluate the question. \bar{A} denotes the opposite events of A . Let A be the event **Rain**.

$$p(A) = p(\bar{A}) = 0.5$$

Let B be the event **Forecasting Rain**.

$$p(B|A) = p(\bar{B}|\bar{A}) = 0.9$$

Let C be the event **Taking Umbrella**.

$$p(C|B) = 1$$

$$p(C|\bar{B}) = 0.5$$

OK, now let's come to the questions.

- (a) the probability of raining when you don't take an umbrella $= p(A|\bar{C})$

$$p(A|\bar{C}) = \frac{p(A)p(\bar{C}|A)}{p(A)p(\bar{C}|A) + p(\bar{A})p(\bar{C}|\bar{A})}$$

$$p(\bar{C}|A) = p(\bar{C}|AB)p(B|A) + p(\bar{C}|A\bar{B})p(\bar{B}|A) = 0 * 0.9 + 0.5 * 0.1 = 0.05$$

$$p(\bar{C}|\bar{A}) = p(\bar{C}|\bar{A}B)p(B|\bar{A}) + p(\bar{C}|\bar{A}\bar{B})p(\bar{B}|\bar{A}) = 0 * 0.1 + 0.5 * 0.9 = 0.45$$

Here, we use that

$$p(\bar{C}|AB) = p(\bar{C}|B)$$

$$p(A|\bar{C}) = \frac{0.5 * 0.05}{0.5 * 0.05 + 0.5 * 0.45} = 0.1$$

- (b) the probability of not raining when you take an umbrella $= p(\bar{A}|C)$
The deduction is the same, so let me omit some steps.

$$p(\bar{A}|C) = \frac{0.5 * 0.55}{0.5 * 0.55 + 0.5 * 0.95} = \frac{11}{30}$$

0.9. (Joint Distribution) Random Variables X and Y have a joint distribution with joint probability density function

$$f(x, y) = \begin{cases} Ce^{-(2x+y)} & x > 0, y > 0 \\ 0 & \text{ow.} \end{cases}$$

Please find C by

$$\int_0^\infty \int_0^\infty f(x, y) dx dy = 1$$

Solution:

$$C \int_0^\infty \int_0^\infty e^{-(2x+y)} dx dy = C \cdot \frac{1}{2} \cdot 1 = 1$$

$$C = 2$$

0.10. (Covariance) Now we have a joint pdf

$$f(x, y) = \begin{cases} 4xy & 0 < x < 1, 0 < y < 1 \\ 0 & \text{ow.} \end{cases}$$

Please show that the covariance of X and Y is 0.

Solution:

$$\mathbb{E}[X] = \int_0^1 \int_0^1 x \cdot 4xy dx dy = \frac{2}{3}$$

$$\mathbb{E}[Y] = \int_0^1 \int_0^1 y \cdot 4xy dx dy = \frac{2}{3}$$

$$\mathbb{E}[XY] = \int_0^1 \int_0^1 xy \cdot 4xy dx dy = \frac{4}{9}$$

Thus,

$$\text{Cov}[X, Y] = 0$$

Of course, if you are clever enough, you will see that they are independent.

0.11. (Uncorrelated and independent RVs) We have a uniform distribution of X and Y on a disk. The pdf is

$$f(x, y) = \frac{1}{\pi} \quad x^2 + y^2 \leq 1$$

Please show that X and Y are uncorrelated but not independent.

Solution:

$$f_X(x) = \int_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}} \frac{1}{\pi} dy = \frac{2\sqrt{1-x^2}}{\pi} \quad -1 < x < 1$$

Similarly,

$$f_Y(y) = \frac{2\sqrt{1-y^2}}{\pi} \quad -1 < y < 1$$

Obviously,

$$f(x, y) \neq f_X(x)f_Y(y) \Rightarrow \text{Not Independent}$$

$$\mathbb{E}[X] = \int_{-1}^1 x \frac{2\sqrt{1-x^2}}{\pi} dx = 0$$

$$\mathbb{E}[Y] = \int_{-1}^1 y \frac{2\sqrt{1-y^2}}{\pi} dy = 0$$

$$\mathbb{E}[XY] = \int_{x^2+y^2 \leq 1} \frac{xy}{\pi} dx dy = 0$$

$$\mathbb{E}[XY] = \mathbb{E}[X] \mathbb{E}[Y] \Rightarrow \text{Uncorrelated}$$

0.12. (Guassian Distribution) There is a famous integral here

$$\int_{-\infty}^{\infty} e^{-x^2} dx = \sqrt{\pi}$$

It is called Guassian Integral. Based on it, please find some results of the Normal Distribution

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right) \quad -\infty < x < \infty$$

- (a) Prove it is a pdf ($\sigma > 0$)
- (b) Compute the expectation and variance

Solution:

(a)

$$\int_{-\infty}^{\infty} f(x) dx = \int_{-\infty}^{\infty} \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\left(\frac{x-\mu}{\sqrt{2}\sigma}\right)^2\right) d\left(\frac{x-\mu}{\sqrt{2}\sigma}\right) = 1$$

(b)

$$\begin{aligned}
\int_{-\infty}^{\infty} x f(x) dx &= \int_{-\infty}^{\infty} (x - \mu + \mu) \frac{1}{\sqrt{\pi}} \exp \left(- \left(\frac{x - \mu}{\sqrt{2}\sigma} \right)^2 \right) d \left(\frac{x - \mu}{\sqrt{2}\sigma} \right) \\
&= \mu + \int_{-\infty}^{\infty} \frac{\sigma}{\sqrt{2\pi}} \exp \left(- \left(\frac{x - \mu}{\sqrt{2}\sigma} \right)^2 \right) d \left(\frac{x - \mu}{\sqrt{2}\sigma} \right)^2 \\
&= \mu
\end{aligned}$$

$$\begin{aligned}
\int_{-\infty}^{\infty} x^2 e^{-x^2} dx &= \int_{-\infty}^{\infty} \frac{1}{2} x e^{-x^2} dx^2 \\
&= - \int_{-\infty}^{\infty} \frac{1}{2} x de^{-x^2} \\
&= \int_{-\infty}^{\infty} \frac{1}{2} e^{-x^2} dx - \frac{1}{2} x e^{-x^2} \Big|_{-\infty}^{+\infty} \\
&= \frac{\sqrt{\pi}}{2}
\end{aligned}$$

$$\begin{aligned}
\int_{-\infty}^{\infty} x^2 f(x) dx &= \int_{-\infty}^{\infty} \frac{(x - \mu)^2 + 2\mu x - \mu^2}{\sqrt{\pi}} \exp \left(- \left(\frac{x - \mu}{\sqrt{2}\sigma} \right)^2 \right) d \left(\frac{x - \mu}{\sqrt{2}\sigma} \right) \\
&= \frac{\sqrt{\pi}}{2} \cdot \frac{2\sigma^2}{\sqrt{\pi}} + 2\mu \cdot \mu - \mu \cdot \mu \\
&= \sigma^2 + \mu^2
\end{aligned}$$

Therefore,

$$\mathbb{E}[X] = \mu$$

$$\text{Var}[X] = \sigma^2 + \mu^2 - \mu^2 = \sigma^2$$