# **ML** Cheatsheet

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ABSTRACT: Everything I know about machine learning.

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## 1 Conventions

## **Math Notation**

- $x \in \{0, 1\}^r$ : x is a vector of form e.g. (0, 1, 1, 0, ..., 1, 0) of length r.
- ~: random variable drawn from a distribution
  - $\alpha$ : "proportional to"
- $\mathbf{1}(\cdot)$ : indicator function 1 when arg is true, 0 when arg is false.
- $I_k$ :  $k \times k$  identity matrix
- Boldface capital letters (roman or greek) are matrices, e.g.  $\mathbf{A} = \mathbf{U} \mathbf{\Lambda} \mathbf{V}^T$
- log is base e by default  $\rightarrow$  entropy in nats.
- $\mathbb{E}_{\theta}$ : Expectation value (with respect to some parameter  $\theta$ . Parameter may be omitted if clear from context.)

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• *m* : number of training examples in the dataset

- *n* : dimension of training examples
- $x^{(i)} \in \mathbb{R}^n$ :  $i^{th}$  training example (*column* vector),  $1 \le i \le m$
- $y^{(i)}$ :  $i^{th}$  output (column vector),  $1 \le i \le m$
- $x_j$ :  $j^{th}$  component of a training example,  $1 \le j \le n$ .
- $\mathbf{X} \in \mathbb{R}^{m \times n}$ : (input/design) matrix (training examples are *row* vectors)

### **BISHOP**

- N: number of training examples in the dataset
- *D* : dimension of training examples
- *M* : number of basis functions
- Boldface lower-case letters (roman or greek) are *column* vectors, *e.g.*  $d = \mathbf{x}^T \mathbf{A} \mathbf{x}$ , or  $\mathbf{y} = \mathbf{A} \mathbf{x}$ .
- $\phi$ : M-dim basis function
- $\Phi \in \mathbb{R}^{N \times M}$ : Design matrix (basis functions are *row* vectors)

## 2 Linear Regression

#### 2.1 Basics

- Hypothesis  $h_{\theta}(x) = \theta_0 + \theta_1 x_1 + \theta_2 x_2 + \dots = \sum_{j=0}^n \theta_j x_j \equiv \theta^T x$ , where  $x_0 = 1$ .
- Cost function

$$J(\theta) = \frac{1}{2} \sum_{i=1}^{m} \left( h_{\theta}(x^{(i)}) - y^{(i)} \right)^{2}$$
 (2.1)

• Analytically solve via normal equations, where **X** is the augmented  $m \times (n+1)$  design matrix

$$\vec{\theta} = \left(\mathbf{X}^T \mathbf{X}\right)^{-1} \mathbf{X}^T \vec{y} \tag{2.2}$$

where  $\vec{\theta}$  and  $\vec{y}$  are column vectors of the n+1 weights and m outputs, respectively.

## 3 Logistic Regression

### 3.1 2 classes

With 2 labels  $y \in \{0, 1\}$  and m training examples  $x^{(i)}$ ,  $1 \le i \le m$ , we have

$$h_{\theta}(x) = \frac{1}{1 + e^{-\theta^T x}} \tag{3.1}$$

$$p(y|x;\theta) = (h_{\theta}(x))^{y} (1 - h_{\theta}(x))^{1-y}$$
(3.2)

$$L(\theta) = \prod_{i=1}^{m} (h_{\theta}(x^{(i)}))^{y^{(i)}} (1 - h_{\theta}(x^{(i)}))^{1 - y^{(i)}}$$
(3.3)

$$l(\theta) = \log L(\theta) = \sum_{i=1}^{m} y^{(i)} \log h_{\theta}(x^{(i)}) + (1 - y^{(i)}) \log(1 - h_{\theta}(x^{(i)}))$$
(3.4)

with deriv

$$\frac{\partial}{\partial \theta_j} l(\theta) = x_j^T (y - h_{\theta}(x)) \tag{3.5}$$

We can't minimize the log-likelihood analytically, so we must use numerical optimization.

The log-likelihood is just the cross entropy

$$H(p,q) = -\sum_{i} p_i \log q_i \tag{3.6}$$

with p the actual outputs, and q the hypothesis  $h_{\theta}(x)$ 

### 3.2 2+ classes

One-versus-all: For k classes, do k normal logistic regressions. Each has two classes: the target class, and all the rest. New examples are classified by whichever of these k logistic regressions has highest score.

**Softmax regression:** Model classification cases as *multinomial* distribution. For k classes, hypothesis is k-dim vector

$$h_{\theta}(x) = \frac{1}{\sum_{l=1}^{k} \exp(\theta^{(l)T} x)} \times \left[ \exp(\theta^{(1)T} x), \dots, \exp(\theta^{(k)T} x) \right]$$
(3.7)

so that the final output layer has k units. The log-likelihood is

$$l(\theta) = -\sum_{i=1}^{m} \sum_{l=1}^{k} \mathbf{1}(y^{(i)} = l) \log \frac{\exp(\theta^{(l)T} x)}{\sum_{m=1}^{k} \exp(\theta^{(m)T} x)}$$
(3.8)

TODO: show how 2-class logistic regression cost func can be written this way. write down deriv of cost func. clean up notation.

## 4 Gaussian Discriminant Analysis (GDA) and Naive Bayes

**Generative algorithm**: Instead of modeling p(y|x), model p(y) and p(x|y), then get posterior p(y|x) via Bayes' theorem.

#### 4.1 GDA

Can use GDA for classification problem when input features are continuous-val random vars. Models p(x|y) as multivariate Gaussian. Model is

$$y \sim \text{Bernoulli}(\phi)$$
  
 $x|y = 0 \sim \mathcal{N}(\mu_0, \Sigma)$   
 $x|y = 1 \sim \mathcal{N}(\mu_1, \Sigma)$  (4.1)

TODO: finish —

## 4.2 Naive Bayes (for text classification)

For training examples  $x \in \{0, 1\}^n$  (a *vocabulary*<sup>1</sup> of length n), assume the components of an example,  $x_i$ , are conditionally independent given y (*Naive Bayes assumption*).

$$p(x_1, \dots, x_n | y) = \prod_{i=1}^n p(x_i | y)$$
 (4.2)

<sup>&</sup>lt;sup>1</sup>the  $j^{th}$  entry is 1 if the example contains the  $j^{th}$  word in the vocabulary. **Example**: If the vocab is {cats, rats, bats}, then n = 3, and a training example that contains "cats", "rats", but not "bats", would be x = (1, 1, 0).

Model is parameterized by

$$\phi_y = p(y=1),$$
  $\phi_{j|y=1} = p(x_j = 1|y=1),$   $\phi_{j|y=0} = p(x_j = 1|y=0)$  (4.3)

where  $j \in (1, n)$ , for a total of 2n + 1 parameters. Likelihood is

$$L(\phi_{y}, \{\phi_{j|y=1}, \phi_{j|y=0}\}_{j=1}^{n}) = \prod_{i=1}^{m} p(x^{(i)}, y^{(i)})$$
(4.4)

which has MLE values

 $\phi_{j|y=1}$  = fraction of spam (y=1) in which word j appears  $\phi_{j|y=0}$  = fraction of non-spam (y=0) in which word j appears  $\phi_{v}$  = fraction of training examples that are spam (4.5)

To make predictions, we don't need the evidence p(x); we just need to compare

$$p(y = 1|x) \propto p(x|y = 1)p(y = 1)$$
  
 
$$p(y = 0|x) \propto p(x|y = 0)p(y = 0)$$
 (4.6)

and pick the class that has the higher value (un-normalized posterior).

TODO: finish — show how last two eqs are actually calculated

## 5 Bias/Variance Tradeoff

- **High bias**: underfitting. high training error and test error.
- **High variance**: overfitting. *low* training error and *high* test error

Can't use the test set to decide on values of hypers (*e.g.* number of parameters in model) because "we could just tweak the values of the hypers until the estimator performs optimally [on the test set]".

Cross-validation allows you to do model comparison to find best value of hypers (on frequentist stats), versus using things like computing the evidence in Bayesian stats. ← TODO: more detail here

Can average together multiple models with high capacity (low bias, but high variance). The result of combining is to *reduce* the variance of the combined model.

## 6 Model Assessment

## 6.1 Classification

#### 6.1.1 Precision/Recall

For 2 classes, with  $\{TN, FN, FP, TP\} = \{\text{true negatives, false negatives, false positives, true positives}\},$  we have the *loss* or *confusion matrix*:

$$\begin{pmatrix}
TN & FP \\
FN & TP
\end{pmatrix},$$
(6.1)

where the rows/columns are actual/predicted numbers of examples not-in-class (negative), or in-class (positive). We define the *precision* and *recall*:

precision = 
$$\frac{TP}{TP + FP}$$
, recall =  $\frac{TP}{TP + FN}$ . (6.2)

**Precision** is ability of classifier to not classify actual negatives as positive (*i.e.* low false positives). So *e.g.* if positive is "person doesn't have cancer", we want high precision.

**Recall** is ability of classifier to find all positive samples.

## 6.1.2 Receiver operating characteristic curve (ROC)

True Positive Rate (TPR): same as recall

*True Negative Rate (TNR):* fraction of actual negatives that are classified as negative TN/(TN + FP)

False Positive Rate (FPR): 1 - TNR

ROC plots TPR versus FPR.

TODO: More on why the ROC is useful. Maybe a picture.

#### 6.1.3 F1 Score

The F1 score is the *harmonic mean* of precision and recall. It's an ok one-number metric for evaluating effectiveness of different classifiers on problem:

$$F1 = \frac{2}{\left(\frac{1}{\text{prec}} + \frac{1}{\text{rec}}\right)} \tag{6.3}$$

## 6.2 Regression

**TODO:** Write this section.

## 7 K-means

- Setup: Select any K points from data to serve as initial centroids of the clusters.
- Repeat until (stopping criteria: *e.g.* no points change clusters, sum of distances is minimized, some total number of iterations..):
  - 1. Assign each datapoint to the cluster with closest centroid.
  - 2. Compute K new centroids, each the mean of the datapoints in its cluster.

## 8 PCA

PCA is unsupervised technique for finding directions of most variance in dataset. The *principal components* are a set of *n* orthonormal *n*-dim basis vectors along the directions of maximum variance, ordered by decreasing variance in their directions.

#### To calc:

- via eigenvectors: Principal components of de-meaned  $m \times n$  design matrix **X** are the eigenvectors of the covariance matrix of **X**,  $\mathbf{C}_{\mathbf{X}} = \frac{1}{m} \mathbf{X}^T \mathbf{X}$ .
- via SVD: Define  $\mathbf{Y} = \frac{1}{\sqrt{m}}\mathbf{X}$ , and take its SVD A.3.2, which is  $\mathbf{Y} = \mathbf{U}\mathbf{W}\mathbf{V}^T$ . Then  $\mathbf{U}^T$  is  $n \times m$  matrix that projects  $\mathbf{X}$  onto its principal components,  $\mathbf{W}$  is the matrix of principal comp values, and  $\mathbf{V}$  is orthonormal matrix of the principal components. TODO: Verify/fix statement on  $\mathbf{U}$  projection.

TODO: re-check everything in PCA section, as I have redefined the design matrix X.

## 9 Time Series

## 9.1 Stationarity

**Definition 1** (Second Order Stationarity). *is when correlation between sequential observations is only a function of the lag.* 

• **Dickey-Fuller test:** tests for stationarity of AR model. Null hypothesis is "series is non-stationary".

### 9.2 ARCH

• **Box-Jenkins:** systematic methodology for identifying and estimating models that can incorporate both AR and MA.

### 10 Gaussian Processes

## 10.1 Prior over functions, from basis functions

Following the treatment in section 6.4 of **BISHOP** (with Bishop's notation).

Let *d* be the number of *basis functions*  $\vec{\phi}$ , which are non-linear functions of the *n*-dim training examples  $\vec{x}$ . We write the hypothesis as

$$y(\vec{x}) = \vec{w}^T \vec{\phi}(\vec{x}) \tag{10.1}$$

where  $\vec{w}$  is a *d*-dim vector of weights.

As in MAP and Bayesian linear regression, we introduce a isotropic Gaussian prior over the weights, with  $\alpha$  its inverse variance:

$$p(\vec{w}) = \mathcal{N}(\vec{w}|\vec{0}, \alpha^{-1}\mathbf{I}_d)$$
(10.2)

The prior over  $\vec{w}$  thus induces a prior over functions  $y(\vec{x})$ .

The design matrix of basis functions is now  $\Phi \in \mathbb{R}^{m \times d}$ , so the hypothesis evaluated on our training examples is an m-dim vector

$$\vec{y} = \Phi \vec{w} \tag{10.3}$$

Each element of  $\vec{y}$  is a linear combination of Gaussians (given by the  $\vec{w}$  prior), where the weights in the linear combination come from elements of the design matrix; thus each element of  $\vec{y}$  is a Gaussian. Then the vector  $\vec{y}$  itself is a multivariate Gaussian, completely specified by its mean vector and covariance matrix.

$$\mathbb{E}_{\vec{w}}[\vec{y}] = \mathbf{\Phi} \, \mathbb{E}_{\vec{w}}[\vec{w}] = \vec{0} \tag{10.4}$$

$$\operatorname{cov}[\vec{y}] = \mathbb{E}_{\vec{w}}[\vec{y}\vec{y}^T] = \mathbf{\Phi} \,\mathbb{E}_{\vec{w}}[\vec{w}\vec{w}^T] \,\mathbf{\Phi}^T = \alpha^{-1}\mathbf{\Phi} \,\mathbf{\Phi}^T \equiv \mathbf{K} \tag{10.5}$$

where **K** is the  $m \times m$  **Gram matrix** with elements

$$\mathbf{K}_{ij} \equiv k(\vec{x}_i, \vec{x}_j) = \alpha^{-1} \vec{\phi}(\vec{x}_i)^T \vec{\phi}(\vec{x}_j). \tag{10.6}$$

Our prior over functions is then

$$p(\vec{y}) = \mathcal{N}(\vec{y}|\vec{0}, \mathbf{K}) \tag{10.7}$$

## 10.2 Basic Gaussian process regression

For regression, we assume our target variables  $t_i$  are given by the hypothesis  $\vec{y}$  with added Gaussian noise

$$p(t_i|y_i) = \mathcal{N}(t_i|y_i, \beta^{-1}) \qquad 1 \le i \le m$$
 (10.8)

where  $\beta$  is the inverse variance of the Gaussian, and the noise is i.i.d for each training example. The *m*-dim vector of target variables,  $\vec{t}$ , is then given by a multivariate Gaussian

$$p(\vec{t}|\vec{y}) = \mathcal{N}(\vec{t}|\vec{y}, \beta^{-1}\mathbf{I}_m)$$
(10.9)

To get the joint distribution  $p(\vec{t})$  over all the target variables, we integrate out  $\vec{y}$  via (A.11), analogous to how we integrate out the weights  $\vec{w}$  in Bayesian regression,

$$p(\vec{t}) = \int p(\vec{t}|\vec{y})p(\vec{y})d\vec{y} = \mathcal{N}(\vec{t}|\vec{0}, \mathbf{C})$$
(10.10)

where **C** is  $m \times m$  with elements  $\mathbf{C}_{ij} = \mathbf{K}_{ij} + \beta^{-1} \delta_{ij}$ .

To make a prediction  $t_{m+1}$ , based on the *m* training examples, we form the joint distribution over m+1 examples,

$$\mathcal{N}(\vec{t}|\vec{0}, \mathbf{C}_{m+1}) \tag{10.11}$$

where  $\vec{t}$  is now (m+1)-dim, and the  $(m+1)\times (m+1)$  matrix  $\mathbb{C}_{m+1}$  is partitioned as

$$\mathbf{C}_{m+1} = \begin{pmatrix} \mathbf{C}_m \ \vec{k} \\ \vec{k}^T \ c \end{pmatrix},\tag{10.12}$$

where

- $\mathbf{C}_m$  is  $m \times m$  with elements  $\mathbf{C}_{ij} = \mathbf{K}_{ij} + \beta^{-1} \delta_{ij}$
- $\vec{k}$  has elements  $k(\vec{x}_i, \vec{x}_{m+1})$  for  $1 \le i \le m$

• 
$$c = k(\vec{x}_{m+1}, \vec{x}_{m+1}) + \beta^{-1}$$

The conditional distribution  $p\left(t_{m+1}|\vec{t}\right)$  is then given by a Gaussian with

$$m(\vec{x}_{m+1}) = \vec{k}^T \mathbf{C}_m^{-1} \vec{t}$$
 (10.13)

$$\sigma^2(\vec{x}_{m+1}) = c - \vec{k}^T \mathbf{C}_m^{-1} \, \vec{k}. \tag{10.14}$$

which we obtain by using (A.7).  $\leftarrow$  **TODO:** check to make sure this is the right eq.

## 11 Neural Networks

## 11.1 Hyperparameters

Hypers: learning rate, num iterations, num hidden layers, num units in each layer, choice of activation function, amount of momentum, minibatch size, regularization type and amount.

## 11.2 Bias/Variance

Bias/variance is less of a *tradeoff* in neural network training, since we have a number of methods to reduce both independently.

- Reduce bias (train set perf): Bigger network, train longer, (NN arch search)
- **Reduce variance (dev set perf):** More data, regularization.. (when changed go back and retune bias)

## 12 Variational Methods

### **Preliminary section!**

If our prior is not a conjugate prior to our likelihood, we can't compute posterior in closed form. Try:

- approximate intractable posterior  $p(\vec{\theta}|\vec{y})$  with tractable  $q(\vec{\theta}|\gamma)$ .
- Adjust  $\gamma$  to minimize  $KL\left[p(\vec{\theta}|\vec{y})||q(\vec{\theta}|\gamma)\right] \to \text{complicated}$ , since we don't know  $p(\vec{\theta}|\vec{y})$ . Eventually will use *expectation propagation* to solve.
- Adjust  $\gamma$  to minimize  $KL[q(\vec{\theta}|\gamma)||p(\vec{\theta}|\vec{y})] \rightarrow \text{simpler: } variational \ Bayes.$
- Variational Bayes: minimizing  $KL\left[q(\vec{\theta}|\gamma)||p(\vec{\theta}|\vec{y})\right] \iff$  maximizing ELBO  $\mathscr L$
- note the *latent/hidden variables* helping us are the  $\vec{\theta}$

## Appendix A Math Stuff

### A.1 Matrix Stuff

## A.1.1 Matrix Identities

(NB More useful stuff in **BARBER** chapter 29.)

### **Trace Identities**

- $Tr(\mathbf{A}^T) = Tr(\mathbf{A})$
- $Tr(\mathbf{A} + \mathbf{B}) = Tr(\mathbf{A}) + Tr(\mathbf{B})$
- Tr(aA) = aTr(A) (with a a scalar)
- Tr(AB) = Tr(BA)
- Tr(ABC) = Tr(BCA) = Tr(CAB) (cyclicity of trace)

## **Transpose Identities**

$$\bullet \ (\mathbf{A}\mathbf{B})^T = \mathbf{B}^T \mathbf{A}^T$$

#### **Determinant Identities**

- $\bullet ||AB|| = |A||B|$
- $\bullet |\mathbf{A}^{-1}| = \frac{1}{|\mathbf{A}|}$
- $\bullet |\mathbf{A}^T| = |\mathbf{A}|$

## **Inverse Indentities**

- $(AB)^{-1} = B^{-1}A^{-1}$
- $\bullet \ (\mathbf{A}^T)^{-1} = (\mathbf{A}^{-1})^T$

## A.1.2 Matrix derivatives

$$\frac{\partial}{\partial \vec{x}} \left( \vec{x}^T \vec{a} \right) = \frac{\partial}{\partial \vec{x}} \left( \vec{a}^T \vec{x} \right) = \vec{a}^2 \tag{A.1}$$

For a matrix  $\mathbf{A}(x)$  as a function of a scalar x (equivalently a component  $x_i$  of a vector  $\vec{x}$ )

$$\frac{\partial}{\partial x} \log |\mathbf{A}| = \text{Tr}\left(\mathbf{A}^{-1} \frac{\partial \mathbf{A}}{\partial x}\right) \tag{A.2}$$

<sup>&</sup>lt;sup>2</sup>TODO: review understanding of this (I recall this expression was not enough to answer my questions on a formula derivation) and give examples.

#### A.1.3 Partitioned Matrix Inversion

Given block decomposition of matrix into submatrices A, B, C, D, the inverse is:

$$\begin{pmatrix} \mathbf{A} & \mathbf{B} \\ \mathbf{C} & \mathbf{D} \end{pmatrix}^{-1} = \begin{pmatrix} \mathbf{M} & -\mathbf{M}\mathbf{B}\mathbf{D}^{-1} \\ -\mathbf{D}^{-1}\mathbf{C}\mathbf{M} & \mathbf{D}^{-1} + \mathbf{D}^{-1}\mathbf{C}\mathbf{M}\mathbf{B}\mathbf{D}^{-1} \end{pmatrix}, \tag{A.3}$$

where

$$\mathbf{M} = \left(\mathbf{A} - \mathbf{B}\mathbf{D}^{-1}\mathbf{C}\right)^{-1} \tag{A.4}$$

(NB compare with Matt Headrick's compendium inversion formula)

## A.1.4 Symmetric Matrix Properties

**TODO:** Do this section.

#### A.2 Gaussians

### A.2.1 Partitioned Gaussians: Conditional and Marginal Distributions

From **BISHOP**, sections 2.3.1 - 2.3.2.

Consider a joint Gaussian distribution,  $\mathcal{N}(\vec{x}|\vec{\mu}, \Sigma)$  with  $\Lambda \equiv \Sigma^{-1}$  and

$$\vec{x} = \begin{pmatrix} \vec{x}_a \\ \vec{x}_b \end{pmatrix}, \quad \vec{\mu} = \begin{pmatrix} \vec{\mu}_a \\ \vec{\mu}_b \end{pmatrix}$$
 (A.5)

$$\Sigma = \begin{pmatrix} \Sigma_{aa} & \Sigma_{ab} \\ \Sigma_{ba} & \Sigma_{bb} \end{pmatrix}, \quad \Lambda = \begin{pmatrix} \Lambda_{aa} & \Lambda_{ab} \\ \Lambda_{ba} & \Lambda_{bb} \end{pmatrix}. \tag{A.6}$$

(That is, the vector  $\vec{x}$  is partitioned into  $\vec{x}_a$  and  $\vec{x}_b$ , which induces a block partitioning of the covariance matrix  $\Sigma$ .)

The conditional distribution is:

$$p(\vec{x}_a|\vec{x}_b) = \mathcal{N}(\vec{x}|\vec{\mu}_{a|b}, \mathbf{\Lambda}_{aa}^{-1}), \quad \text{where} \quad \vec{\mu}_{a|b} \equiv \vec{\mu}_a - \mathbf{\Lambda}_{aa}^{-1} \mathbf{\Lambda}_{ab} (\vec{x}_b - \vec{\mu}_b). \tag{A.7}$$

The marginal distribution is:

$$p(\vec{\mathbf{x}}_a) = \mathcal{N}(\vec{\mathbf{x}}_a | \vec{\boldsymbol{\mu}}_a, \boldsymbol{\Sigma}_{aa}) \tag{A.8}$$

### A.2.2 Bayes' theorem for Gaussian variables

From **BISHOP**, section 2.3.3. NB This is known as a *linear Gaussian model*.

Given a marginal Gaussian distribution for  $\vec{x}$  and a conditional Gaussian distribution for  $\vec{y}$  given  $\vec{x}$  in the form:

$$p(\vec{x}) = \mathcal{N}(\vec{x}|\vec{\mu}, \mathbf{\Lambda}^{-1}) \tag{A.9}$$

$$p(\vec{y}|\vec{x}) = \mathcal{N}(\vec{y}|\mathbf{A}\vec{x} + \vec{b}, \mathbf{L}^{-1})$$
(A.10)

we have

$$p(\vec{y}) = \mathcal{N}(\vec{y}|\mathbf{A}\vec{\mu} + \vec{b}, \mathbf{L}^{-1} + \mathbf{A}\mathbf{\Lambda}^{-1}\mathbf{A}^{T})$$
(A.11)

$$p(\vec{x}|\vec{y}) = \mathcal{N}(\vec{x}|\mathbf{\Sigma}\{\mathbf{A}^T\mathbf{L}(\vec{y} - \vec{b}) + \mathbf{\Lambda}\vec{\mu}\}, \mathbf{\Sigma})$$
(A.12)

where

$$\Sigma \equiv (\mathbf{\Lambda} + \mathbf{A}^T \mathbf{L} \mathbf{A})^{-1}. \tag{A.13}$$

### A.3 Matrix Factorization

### A.3.1 Diagonalization (Eigendecomposition)

A square  $N \times N$  symmetric matrix **A** has *real* eigenvalues from its symmetricity, and N linearly independent eigenvectors. We can write  $\mathbf{A} = \mathbf{U}\mathbf{D}\mathbf{U}^{-1} = \mathbf{U}\mathbf{D}\mathbf{U}^{T}$ , with **U** an orthogonal<sup>3</sup> matrix with columns  $\vec{u}_i$  equal to **A**'s eigenvectors, and **D** a diagonal matrix of **A**'s eigenvalues,  $\lambda_i$ .

We can also write down the *spectral decomposition* of A:

$$\mathbf{A} = \mathbf{U}\mathbf{D}\mathbf{U}^T = \sum_{i=1}^{N} \lambda_i \, \vec{u}_i \, \vec{u}_i^T. \tag{A.14}$$

### A.3.2 Singular value decomposition

Using conventions of **PRESS**, section 2.6.

An  $M \times N$  matrix **A** can be written as  $\mathbf{UDV}^T$ , where **U** is  $M \times N$ , **D** is  $N \times N$  and diagonal, and **V** is  $N \times N$  and orthogonal.

If **A** is square, the columns of **V** are its eigenvectors, and **D** are its eigenvalues.

**Uses**:

• If  $M \ge N$ , cols of **U** are an orthonormal basis for space spanned by cols of **A** 

TODO: finish —

## **Appendix B** The Linear Regression Hierarchy

Linear regression starting from maximum likelihood, through variational Bayes.

### **B.1** Maximum likelihood

Maximize likelihood  $p(y|\theta; \beta)$  wrt  $\theta, \beta$ .

<sup>&</sup>lt;sup>3</sup>An *orthogonal* matrix should really be called an *orthonormal* matrix, as its columns are not just linearly independent, but also normalized.

## B.2 Maximum a posteriori

Introduce prior  $p(\theta; \alpha) = \exp(-\alpha \theta^T \theta)$ . Find  $\theta$  by maximizing  $p(y|\theta; \beta)p(\theta; \alpha)$  wrt  $\theta$ .

- Q: what is proper way to estimate  $\beta$  and  $\alpha$  here?
- half-A: not sure if you can? You typically use the evidence approx (see "Full Bayesian", below) to find best values for  $\beta$  and  $\alpha$ .. OTOH we found most likely  $\beta$  in previous section ("Maximum Likelihood"), so why can't we do it here?  $\leftarrow$  **TODO:** finish answering this.

## **B.3** Full Bayesian (stationary prior)

Calculate posterior  $p(\theta|y) = p(y|\theta; \beta)p(\theta; \alpha)/p(y; \alpha, \beta)$ , where  $p(y; \alpha, \beta) = \int_{\theta} p(y|\theta; \beta)p(\theta; \alpha)d\theta$  is the *evidence*. Determine values of hypers  $\alpha, \beta$  by maximizing evidence wrt them. This is difficult to do directly, as derivs wrt  $\alpha, \beta$  are hard to compute, so use EM algorithm.

## **B.4** Variational Bayes (non-stationary prior)

## References

- **BARBER** David Barber. *Bayesian Reasoning and Machine Learning*. Cambridge University Press, 2012. ISBN: 0521518148.
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