

Fully Connected >> & parameter multiplies once

(In x In x I) x (On xon x od)

Convolutional > Parameters are shared across slides

(# of parameters) X (# of slides of convolution)

= (Cfn xfw xId) x N) x COn x Ow)

Layer	# Units	# Weights	# Connections
Convolution Layer 1	55×55×96	1\x11x3×96	(1)x11x3×96)×(55×55)
	= 290900	=34848	= 105415200
Convolution Layer 2	27×27×256	5×5×48×256	(5×5×48×256)×(27×27)
,	= 186624	= 307200	= 223948800
Convolution Layer 3	13x13x384	3×3×256×384	
•	=64896	=884736	= 149520384
Convolution Layer 4	13 ×13×384	3×3×192×384	(3×3×192×384)×(13×13)
30	= 64896	=663552	=112140288
Convolution Layer 5	13×13×256	3×3×192×256	(3×3×142×256) × (13×13)
	=43264	= 442368	=74760192
Fully Connected Layer 2	4096	9216×4096	9216×4096
2.13		= 37748736	= 377 48736
Fully Connected Layer 2	4096	4096×4096	4096 x 4096
		=16777216	= 16777216
Output Layer	1000	4096 X1000	4096×1000
- a put - wider		=4096,000	=4096000

b) Suggest a change to the architecture which will help achieve the dired objective. > Modify one or more layers.

i) Want to reduce memory usage at test time so network can run on a collaphone. This requires reducing number of parameters for the network.

Trained parameters need to be stored in memory.

Fully Connected layers account for majority of the # of parameters.

- : , reduce the size of fully connected layer to reduce # of parameters

=> reduce memory usage.

ii) Network needs to make rapid predictions at test time.

Want to reduce the number of connections, since there is approximately one add-multiply operation per connection.

of connections = # of forward pass computation = time complexity
for predictions at test time.

Convolutional Layers account for majority of the # of connections.

in, reduce number filters for the convolutional layer.

2. Gaussian Naive Bayes

Derive the max. likelihood estimates for Gaussian Naive Bayes, where the features are continuous, and the conditional distribution of each feature given the class is univariate Gaussian rather than Bernoulli.

Start with a generative model for a discrete class label ye [1,2,..., k]

and a real valued vector of d features = (x,

$$P(y=k)=X_k$$

$$P(\overrightarrow{z}|y=k,\overrightarrow{z},\overrightarrow{\sigma}) = \left(\frac{1}{1-2}STG^2\right)^{-\frac{1}{2}} = \frac{\sum_{i=1}^{n} (x_i - u_{ki})^2}{e^{iz}}$$
each feature is anditionally independent given

de = prior on class k the discrete class label.

0; = variances for each feature, shared between all classes

$$U_{ki} = \text{mean of feature is conditioned on class } k$$

$$\vec{\alpha} = \begin{pmatrix} \alpha_1 \\ \alpha_2 \\ \vdots \\ \alpha_k \end{pmatrix}, \quad \vec{\sigma} = \begin{pmatrix} \sigma_1 \\ \sigma_2 \\ \vdots \\ \sigma_k \end{pmatrix}, \quad \vec{\alpha} = \begin{pmatrix} u_{11} \\ u_{21} \\ \vdots \\ u_{k_1} \\ \vdots \\ u_{k_n} \end{pmatrix}$$

D= d in this question

a) Use Bayes rule to derive an expression for p(y=k/x, x, 2). Hint: Use law of total probability to derive an expression for p(\$10, \$) P(=, y== |]= P(== = |],],] P(= |],] = P(=1y=k, x, =) P(y=k)

$$P(y=k|\vec{x},\vec{u},\vec{\sigma}) = \frac{P(\vec{x}|y=k,\vec{u},\vec{\sigma})P(y=k)}{P(\vec{x}|\vec{u},\vec{\sigma})}$$

P(=11, =) = \(\frac{\xi}{\pi} P(=) y= 1 | \alpha \rightarrow \) = \(\frac{\xi}{\pi} P(\frac{\xi}{\pi} | y= 1 | \alpha \rightarrow \frac{\xi}{\pi} \) = \(\frac{\xi}{\pi} P(\frac{\xi}{\pi} | y= 1 | \alpha \rightarrow \frac{\xi}{\pi} \) = \(\frac{\xi}{\pi} P(\frac{\xi}{\pi} | y= 1 | \alpha \rightarrow \frac{\xi}{\pi} \]

$$P(y=k|\vec{x},\vec{x},\vec{\sigma}) = \frac{P(\vec{x}|y=k)\vec{x},\vec{\sigma})P(y=k)}{\sum_{i=1}^{n} P(\vec{x}|y=i)}$$

$$\frac{\xi}{\xi^{2}}\left(\frac{1}{1}2\pi\sigma_{i}^{2}\right)^{-\frac{1}{2}}e^{-\frac{\xi^{2}}{2}\frac{1}{2}(x_{i}-u_{ji})^{2}}\left(\chi_{j}^{2}\right)$$

$$\left(\frac{1}{11}2516^{2}\right)^{-\frac{1}{2}}\sum_{j=1}^{k}\left(e^{-\frac{2}{2}\frac{1}{25}}(x_{i}-u_{ji})^{2}(x_{j})\right)$$
independent of j
$$-\frac{2}{25}\sum_{j=1}^{k}(x_{i}-u_{kj})^{2}(x_{j})$$

a) Use Bayes rule to derive an expression for p(y=k/x, x, 2).

Hint: Use law of total probability to derive an expression for p(\$10, \$)

P(=, y== |], =) = P(y= = |],],] P(= |], =) = P(=1y=1, x, =) P(y=1)

> P(y=k1 =, v, s) = P(=1 y=k, v, s) P(y=k) P(=12, 2)

P(=11,=)= \(\frac{\xi}{\pi} P(=) y= 1 | \alpha , = \(\frac{\xi}{\pi} P(\frac{\xi}{\pi} | y= 1) \alpha , \frac{\xi}{\pi} P(\frac{\xi}{\pi} | y= 1) \\ \frac{\xi

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> = (1 2 st s;2) = - \(\frac{1}{2} \frac{1} \(\left(\frac{1}{11} 2 \si \si^2 \right)^{\frac{1}{2}} = \frac{2}{12} \left(\chi_{i=1}^{2} \left(\chi_{i} - u_{ji} \right)^{2} \left(\chi_{j} \right) \right)

= (T 2 x 62) - 1 - 2 - 2 - 2 - 2 (x; - uki) 2 (Xk)

 $\left(\frac{1}{1} \cdot 2 \cdot 5 \cdot 5^{2}\right)^{-\frac{1}{2}} \left(e^{-\frac{S}{2} \cdot \frac{1}{2} \cdot \frac{1}{2} \cdot \frac{1}{2}} (x_{i} - u_{ji})^{2} (x_{j})\right)$ $= e^{-\frac{S}{2} \cdot \frac{1}{2} \cdot \frac{1}{2} \cdot \frac{1}{2}} (x_{i} - u_{ki})^{2} (x_{k})$

\\ \(\) \(

= (Xk)e-\(\frac{1}{2\varepsilon_{12}}(x_1-u_{ki})^2\) P(y= | = , , ,) $\frac{\mathcal{E}}{\mathcal{E}}\left(\left(\chi_{j}^{\prime}\right)e^{-\frac{\mathcal{E}}{2\varepsilon_{i}^{2}}\left(\chi_{i}-u_{ji}^{\prime}\right)^{2}}\right)$

a specific label = k = K = # of labels

b) Write down an expression for the negative likelihood function 1(0;0)=-log P(y(0), 2(0), y(2), 2(2), ..., y(N), 2(N) (0) Dataset: D = {((1), 2(1)), ((2), 2(2)), ..., ((N), 2(N))}, i.i.d. Parameters=== E & = = = 3

- log(P(yc)=(1), y(1)=(1), y(1)=(1))

 $= -\ln(\frac{\pi}{100} O(6) + 260(3))$

b) Write down an expression for the negative likelihood function $l(\theta; D) = -\log P(y(0), \vec{x}(0), y(2), \vec{x}(2), \dots, y(N), \vec{x}(N) | \vec{\theta})$ Dataset: $D = \{(y(0), \vec{x}(0)), (y(2), \vec{x}(2)), \dots, (y(N), \vec{x}(N))\}, i.i.d.$ Parameters = $\vec{\theta} = \{\vec{x}, \vec{x}, \vec{\sigma}\}$

 $= -\log(P(y^{(i)}, z^{(i)}, y^{(i)}, z^{(i)}, y^{(i)}, z^{(i)}))$ $= -\log(\frac{N}{1} P(y^{(i)}, z^{(i)} | \vec{\theta})) \quad \text{, singe data are inited.}$

 $= -\sum_{i=1}^{N} \log \left(P(y^{(i)}, \hat{x}^{(i)} | \vec{\theta}) \right)$ $= -\sum_{i=1}^{N} \log \left(P(\hat{x}^{(i)} | \vec{\theta}) \right) P(y^{(i)} | \vec{\theta})$ $= -\sum_{i=1}^{N} \log \left(P(\hat{x}^{(i)} | \vec{\theta}) P(y^{(i)} | \vec{\theta}) \right)$

 $= -\frac{2}{2} \left(\log \left(P(\vec{x}^{(i)} | \vec{0}) + \log \left(P(\vec{y}^{(i)}) \right) \right) - P(\vec{y}^{(i)} | \vec{0}) = P(\vec{y}^{(i)}) = Q_{(i)}$ $= -\frac{2}{2} \left(\log \left(\frac{1}{12} 2 \pi \sigma_{i}^{2} \right)^{-\frac{1}{2}} - \frac{2}{2} \frac{1}{2\sigma_{i}^{2}} (x_{i}^{(i)} - y_{i}^{(i)})^{2} \right) + \log \left(\alpha_{y^{(i)}} \right)$ $= -\frac{2}{2} \left(-\frac{1}{2} \sum_{j=1}^{2} \log (2 \pi \sigma_{j}^{2}) + \left(-\frac{2}{2} \frac{1}{2\sigma_{i}^{2}} (x_{i}^{(i)} - y_{i}^{(i)})^{2} \right) + \log (\alpha_{y^{(i)}}) \right)$ $= -\frac{2}{2} \left(-\frac{1}{2} \sum_{j=1}^{2} \log (2 \pi \sigma_{j}^{2}) + \left(-\frac{2}{2} \frac{1}{2\sigma_{i}^{2}} (x_{i}^{(i)} - y_{i}^{(i)})^{2} \right) + \log (\alpha_{y^{(i)}}) \right)$

Negative Log Likelihood $\begin{pmatrix}
(\theta, 0) = -\sum_{i=1}^{N} \left(-\frac{1}{2}\sum_{j=1}^{N} \log(2\pi\sigma_{j}^{2}) - \sum_{j=1}^{N} \frac{1}{2\sigma_{j}^{2}} (x_{j}^{(i)})^{2} + \log(\alpha_{y_{cij}})
\end{pmatrix}$

c) Take partial derivatives of the likelihood wirt each of the parameters Uki

and w.r.t. the shared variances 5? Based on this, find the max. likelihood estimates for it and of.

Assume each class appears at least once in the dataset.

Assume each class appears at least once in the datase Partial derivatives wint. Uk:

Partial derivatives w.r.t. Uk;

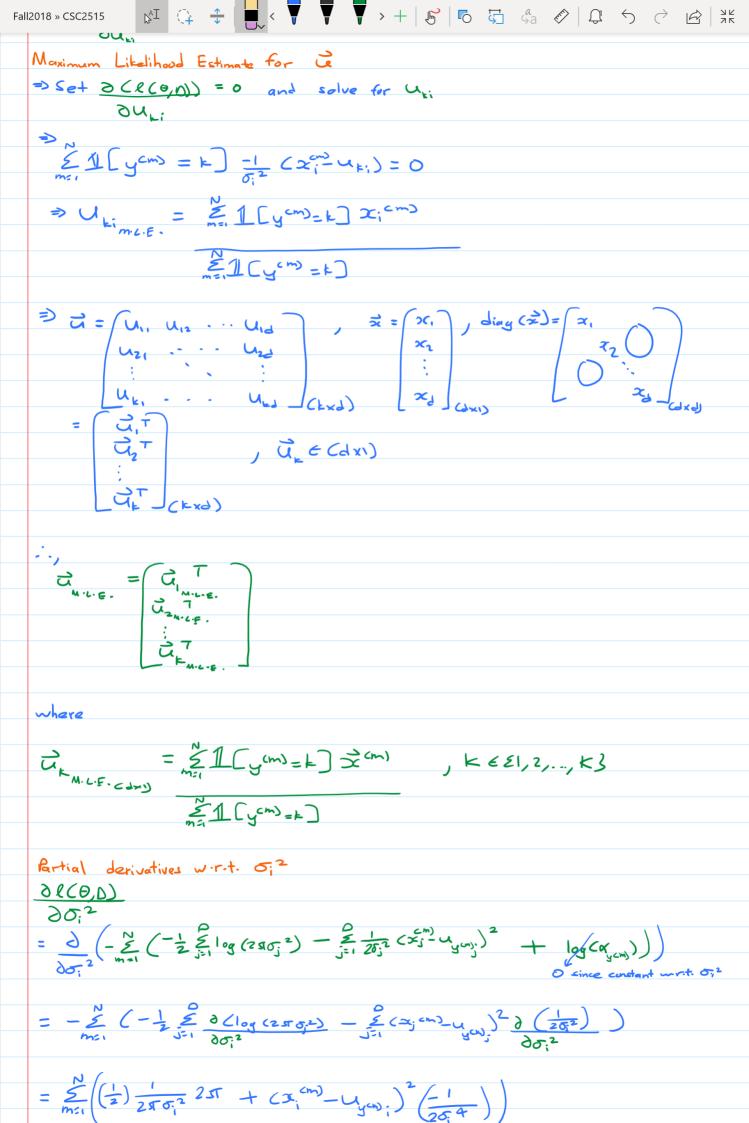
gar:

 $= \frac{\partial}{\partial u_{ki}} \left(-\frac{1}{2} \sum_{j=1}^{k} \log (2\pi \sigma_{j}^{2}) - \sum_{j=1}^{k} \frac{1}{2\delta_{j}^{2}} (x_{j}^{cm}) - u_{ym_{j}}^{cm} \right)^{2} + \log (\alpha_{ym_{j}}^{cm})$ $= \frac{\partial}{\partial u_{ki}} \left(\sum_{m=1}^{k} \sum_{j=1}^{k} \frac{1}{2\sigma_{j}^{2}} (x_{j}^{cm}) - u_{ym_{j}}^{cm} \right)^{2}$ $= \frac{\partial}{\partial u_{ki}} \left(\sum_{m=1}^{k} \sum_{j=1}^{k} \frac{1}{2\sigma_{j}^{2}} (x_{j}^{cm}) - u_{ym_{j}}^{cm} \right)^{2}$

 $= \sum_{m=1}^{N} 1 \left[y^{cm} = k \right] \frac{2}{25i^{2}} \left(x_{i}^{cm} - u_{ki} \right) \left(-1 \right)$

 $\frac{\partial (\ell(\theta, 0))}{\partial u_{ki}} = \sum_{m=1}^{N} I \left[y^{cm} \right] = \sum_{j=1}^{N} \frac{1}{\sigma_{i}^{2}} \left(x_{i}^{cm} - u_{ki} \right)$ Maximum Likelihood Estimate for u_{ki}^{cm}

=> Set 3 (R(G,N)) = 0 and solve for Uk;



Rartial derivatives w.r.t. 5;2

$$= \frac{\partial}{\partial \sigma_{i}^{2}} \left(-\frac{2}{2} \sum_{j=1}^{2} \log (2 \pi \sigma_{j}^{2}) - \sum_{j=1}^{2} \frac{1}{2 \sigma_{j}^{2}} (x_{j}^{cm})^{2} + \log (x_{j}^{cm}) \right)$$

$$= \frac{\partial}{\partial \sigma_{i}^{2}} \left(-\frac{2}{2} \sum_{j=1}^{2} \log (2 \pi \sigma_{j}^{2}) - \sum_{j=1}^{2} \frac{1}{2 \sigma_{j}^{2}} (x_{j}^{cm})^{2} + \log (x_{j}^{cm}) \right)$$

$$= \frac{\partial}{\partial \sigma_{i}^{2}} \left(-\frac{2}{2} \sum_{j=1}^{2} \log (2 \pi \sigma_{j}^{2}) - \sum_{j=1}^{2} \frac{1}{2 \sigma_{j}^{2}} (x_{j}^{cm}) \right)$$

$$= \frac{\partial}{\partial \sigma_{i}^{2}} \left(-\frac{2}{2} \sum_{j=1}^{2} \log (2 \pi \sigma_{j}^{2}) - \sum_{j=1}^{2} \frac{1}{2 \sigma_{j}^{2}} (x_{j}^{cm}) \right)$$

$$= \frac{\partial}{\partial \sigma_{i}^{2}} \left(-\frac{2}{2} \sum_{j=1}^{2} \log (2 \pi \sigma_{j}^{2}) - \sum_{j=1}^{2} \frac{1}{2 \sigma_{j}^{2}} (x_{j}^{cm}) \right)$$

$$= \frac{\partial}{\partial \sigma_{i}^{2}} \left(-\frac{2}{2} \sum_{j=1}^{2} \log (2 \pi \sigma_{j}^{2}) - \sum_{j=1}^{2} \frac{1}{2 \sigma_{j}^{2}} (x_{j}^{cm}) \right)$$

$$= -\frac{5}{5} \left(-\frac{1}{2} \frac{2}{5} \frac{\partial (\log(2\pi\sigma_{i}^{2}))}{\partial \sigma_{i}^{2}} - \frac{5}{5} (2\pi^{2}) \frac{1}{2} \frac{\partial (\log(2\pi\sigma_{i}^{2}))}{\partial \sigma_{i}^{2}} \right)$$

$$= \sum_{m=1}^{N} \left(\left(\frac{1}{2} \right) \frac{1}{2 \pi \sigma_{i}^{2}} 2 \pi + \left(2 \pi_{i}^{(m)} - u_{y^{(m)}} \right)^{2} \left(\frac{1}{2 \sigma_{i}^{4}} \right) \right)$$

$$= \frac{8}{20i^{2}} - \frac{1}{20i^{4}} - \frac{1}{20i^{4}}$$

$$\frac{\partial l(\theta, D)}{\partial \sigma_i^2} = \frac{8}{m_{51}} \frac{1}{2\sigma_i^2} - \frac{\left(x_i^{cm} - u_{yem}_i\right)^2}{2\sigma_i^2}$$
More Althority of Eath 1 for σ_i^2 4 try

$$\frac{3 \text{ Set } 3 \text{ R}(\Theta_{i}, 0)}{30^{2}} = 0 \text{ and solve for } \sigma_{i}$$

$$\Rightarrow \underbrace{\frac{1}{25i^2}}_{mc_1} \underbrace{\frac{(x_1^{cm}) - u_{yms_1}^{cm}}{25i^2}}_{25i^4} = 0$$

$$\sum_{m=1}^{N} \sigma_{1}^{2} - (x_{1}^{cm})^{2} - 0$$

d) Show that the M.L.E. for ok is given by

Assume each class appears at least once.

Hint: Use Lagrange Multipliers.

Since α_k is the prior probability, P(y=k), will need to satisfy constraint $\sum_{i=1}^{k} \alpha_i = 1$

to be a valid probability distribution.

Lagrange Multiplier

Use Lagrange Multiplier to turn this constrained optimization probl into an

unconstrained optimization problem.

Minimize f(z)

st. q(2)=0

$$L(z,\lambda) = f(z) + \lambda g(z)$$

$$9(2) = \ell(\theta, 0)$$

$$g(2) = \frac{\xi}{2} \alpha_{3} - 1 = 0$$

9 F (5/7)

$$= \sum_{k=1}^{N} (1[y^{(k)} = k](-\frac{1}{x_{k}}) + 1)$$

$$\frac{\partial L(z,x)}{\partial x} = \sum_{i=1}^{N} \left(\frac{1}{2} \left(\frac{1}{2} \left(\frac{1}{2} \right) + \frac{1}{2} \right) \right)$$

$$\sum_{i=1}^{N} \left(\mathbb{1} \left[y^{(i)} = k \right] \left(-\frac{1}{\alpha_{k}} \right) \right) + 2 \right) = 0$$

$$\frac{1}{M_{E}} = \frac{1}{N}$$

=>
$$\alpha_{k} = \frac{1}{N} \stackrel{\sim}{=} 1 [y^{ci}] = k$$
, .., proven