

Gaussian Mixture Model (GMM)

Assume: $X_1 \sim N(x | \mu_1, \Sigma_1) \quad P(X_1) = \pi_1$
 $X_2 \sim N(x | \mu_2, \Sigma_2) \quad P(X_2) = \pi_2$

$$x = \begin{pmatrix} a \\ b \end{pmatrix} \quad P(x) = P(x \in X_1) + P(x \in X_2)$$

$$= \underbrace{\pi_1 \Phi(x)}_{\text{Probability of belonging to } X_1} + \pi_2 \Phi(x)$$

Probability of observing x .
 \times Probability of observing the value under X_1

Let all observations be X .

$$P(X | \pi, \mu, \Sigma) = \prod_{n=1}^N \left[\sum_{k=1}^K \pi_k N(x_n | \mu_k, \Sigma_k) \right]$$

max Assume i.i.d

$$1_{x \in k}(x_n) = \begin{cases} 1 & \text{if true,} \\ 0 & \text{else} \end{cases}$$

$$\gamma_{nk}(1_{x \in k}(x_n)) = P(1_{x \in k}(x_n) = 1 | x_n)$$

$$= \frac{P(x_n | 1_{x \in k}(x_n) = 1) P(1_{x \in k}(x_n) = 1)}{P(x_n)}$$

$$= \frac{N(x_n | \mu_k, \Sigma_k) \pi_k}{\sum_{j=1}^K N(x_n | \mu_j, \Sigma_j) \pi_j}$$

$$N_k = \sum_{n=1}^N \gamma_{nk} : \text{Expected Number of observations assigned to class } k.$$

Maximize $P(x|\pi, \mu, \Sigma)$ w.r.t.:

$$\mu_k = \frac{1}{N_k} \sum_{n=1}^N \pi_{nk} (1) x_n$$

$$\Sigma_k = \frac{1}{N_k} \sum_{n=1}^N \pi_{nk} (2) (x_n - \mu_k) (x_n - \mu_k)^T$$

$$\pi_k = N_k / N$$

Intuitively:

μ_k = Avg the weighted avg of x_n belonging to class k .

Σ_k = Avg the weighted avg MSE w.r.t class k .

π_k = Proportion of samples in k to total.

Note that everything depends of π_{nk} } $\left. \begin{array}{l} \text{Covarian} \\ \text{dependency} \end{array} \right\}$
but π_{nk} depends on π, μ, Σ .

To solve this: Expectation-Maximization

Mixture Autoregressive Model

$$F(y_t | \mathcal{F}_{t-1}) = \sum_{k=1}^K \underbrace{\alpha_k}_{\text{weight}} \underbrace{\Phi\left(\frac{y_t - \underbrace{\phi_{k0} - \phi_{k1}y_{t-1} - \dots - \phi_{kp}y_{t-p}}_{\text{Intercept}(C_t)}}{\sigma_k}\right)}_{\substack{y_t - \text{AR}(p_k) \\ \sigma_k = \frac{G_{t,k}}{\sigma_k} \rightarrow \frac{\mu + \sigma}{\sigma}}}$$

Conditional CDF of Y_t at y_t

where $\alpha_k > 0$ and $\sum_{k=1}^K \alpha_k = 1$

Intuitively

Conditional CDF given information up to $t-1$ = Weighted sum of the probability of observing the z-scaled observation assuming a Gaussian distribution.

Naturally:

$$E[y_t | \mathcal{F}_{t-1}] = \sum_{k=1}^K \alpha_k \mu_{k,t}$$

$$E[V(y_t | k, \mathcal{F}_{t-1})] = \sum_{k=1}^K \alpha_k \sigma_k^2$$

$$V[E(y_t | k, \mathcal{F}_{t-1})] = \sum_{k=1}^K \alpha_k \mu_{k,t}^2 - \left(\sum_{k=1}^K \alpha_k \mu_{k,t} \right)^2$$

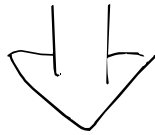
$$\therefore V[y_t | \mathcal{F}_{t-1}] = \sum_{k=1}^K \alpha_k \sigma_k^2 + \underbrace{\sum_{k=1}^K \alpha_k \mu_{k,t}^2 - \left(\sum_{k=1}^K \alpha_k \mu_{k,t} \right)^2}_{\geq 0 \text{ (variance cannot be negative)}}$$

* Law of Total Variance:

$$V(y) = \underbrace{E[V(y|x)]}_{\text{Average variance / random}} + \underbrace{V(E(y|x))}_{\text{variance of } \mu \text{ / within each group}}$$

$$F(y_{t+2} | \mathcal{F}_t) = \underbrace{\int}_{\text{CDF of } y_{t+2}} \underbrace{F(y_{t+2} | \mathcal{F}_t, y_{t+1})}_{\text{probability mass}} \underbrace{dF(y_{t+1} | \mathcal{F}_t)}_{\substack{\text{point probabilities} \\ \text{(curtain)}}}$$

weighted average of the CDF
for y_{t+2} given information up to time t ,
integrating w.r.t all possible y_{t+1} values.



* Note that the above is intractable

\therefore Approximate using Monte Carlo

$$F(y_{t+2} | \mathcal{F}_t) \approx \frac{1}{N} \sum_{j=1}^N F(y_{t+2} | \mathcal{F}_t, y_{t+1}^j) \quad (??)$$

sampled from $F(y_{t+1} | \mathcal{F}_t)$

Perhaps they use

Expectation-Maximization to
get the CDF.

$$\underline{\underline{F(y_{t+1} | \mathcal{F}_t; \theta)}}$$

$$f(y_e | \mathcal{F}_{e-1}) = \frac{1}{\sqrt{2\pi} \sigma_k} e^{-\frac{(y_e - \mu_{k,e})^2}{2\sigma_k^2}}$$

$$Z_{k,e} = \begin{cases} 1 & \text{if } y_e \in k \\ 0 & \text{else} \end{cases}$$

if $y_e \notin k$,
we would get
 $f(x)^0 = 1$

$$f(y_e | Z_e, \mathcal{F}_{e-1}) = \prod_{k=1}^K [f_k(y_e | \mathcal{F}_{e-1})]^{Z_{k,e}}$$

$$f(y_e, z_e | \mathcal{F}_{e-1}) = \prod_{k=1}^K [\alpha_k f_k(y_e | \mathcal{F}_{e-1})]^{Z_{k,e}}$$

Remember that conditional PDFs:

$$p(z | \theta, x) = \frac{h(x, z | \theta)}{g(x | \theta)}$$

$$= \frac{\text{complete Likelihood}}{\text{observed Likelihood}}$$

$$= \frac{L^c(\theta | x, z)}{L(\theta | x)}$$

Does not depend
on random variable z

$$\log L(\theta | x) = \log L^c(\theta | x, z) - \log p(z | \theta, x)$$

$$\log L(\theta | x) = \underbrace{E_{\theta_0} [\log L^c(\theta | x, z)]}_{\text{Maximize this}} - E_{\theta_0} [\log p(z | \theta, x)]$$

$$E_{\theta_0} [\log L^c(\theta | x, z)] = E_{\theta_0} [\log f(y_e, z_e | \mathcal{F}_{e-1}; \theta)]$$

Note the link



$$p(z_e | y_e, \mathcal{F}_{e-1}; \theta)$$

$$= \frac{f(y_e, z_e | \mathcal{F}_{e-1}; \theta)}{f(y_e | \mathcal{F}_{e-1}; \theta)}$$

??

$$f(y_c, z_c | \mathcal{F}_{c-1}) = \prod_{k=1}^K [\alpha_k f_k(y_c | \mathcal{F}_{c-1})]^{z_{k,c}}$$

$$\log f(y_c, z_c | \mathcal{F}_{c-1}) = \sum_{k=1}^K z_{k,c} [\log \alpha_k + \log f_k(y_c | \mathcal{F}_{c-1})]$$

$$f(y_c | \mathcal{F}_{c-1}) = \frac{1}{\sqrt{2\pi} \sigma_k} e^{-\frac{(y_c - \mu_{k,c})^2}{2\sigma_k^2}}$$

$$\begin{aligned} \log f(y_c | \mathcal{F}_{c-1}) &= \log (\sqrt{2\pi} \sigma_k)^{-1} + \log e^{-\frac{(y_c - \mu_{k,c})^2}{2\sigma_k^2}} \\ &= -\left[\log(2\pi)^{\frac{1}{2}} + \log \sigma_k\right] - \frac{(y_c - \mu_{k,c})^2}{2\sigma_k^2} \\ &= -\frac{1}{2} \log(2\pi) - \log \sigma_k - \frac{(y_c - \mu_{k,c})^2}{2\sigma_k^2} \end{aligned}$$

$$\log f(y_c, z_c | \mathcal{F}_{c-1}) = \sum_{k=1}^K z_{k,c} \left[\underbrace{\log \alpha_k - \frac{1}{2} \log(2\pi) - \log \sigma_k}_{\text{removed. constant will not contribute any information for maximization.}} - \frac{(y_c - \mu_{k,c})^2}{2\sigma_k^2} \right]$$

$$l_c = \sum_{k=1}^K z_{k,c} \left[\log \alpha_k - \log \sigma_k - \frac{\varepsilon_{k,c}^2}{2\sigma_k^2} \right]$$

$$L = \sum_{c=p+1}^n l_c = \sum_{c=p+1}^n \left\{ \sum_{k=1}^K z_{k,c} \left[\log \alpha_k - \log \sigma_k - \frac{\varepsilon_{k,c}^2}{2\sigma_k^2} \right] \right\}$$

where $\varepsilon_{k,c} = y_c - \phi_{k,0} - \phi_{k,1} y_{c-1} - \dots - \phi_{k,p_k} y_{c-p_k}$

θ refers to our unknown parameters: α, ϕ, σ

$$J = \sum_{t=p+1}^n \left\{ \sum_{k=1}^K Z_{k,t} \left[\log \alpha_k - \log \sigma_k - \frac{\varepsilon_{k,t}^2}{2\sigma_k^2} \right] \right\}$$

$$\frac{\partial J}{\partial \alpha_k} = \frac{\partial}{\partial \alpha_k} \sum_{t=p+1}^n \sum_{k=1}^K Z_{k,t} \log \alpha_k, \quad \text{s.t.} \quad \sum_{k=1}^K \alpha_k = 1$$

$$= \frac{\partial}{\partial \alpha_k} \sum_{t=p+1}^n \left(\sum_{k=1}^{K-1} Z_{k,t} \log \alpha_k + Z_{K,t} \log \alpha_K \right)$$

$$= \frac{\partial}{\partial \alpha_k} \sum_{t=p+1}^n \left[\underbrace{\sum_{k=1}^{K-1} Z_{k,t} \log \alpha_k + Z_{K,t} \log \left(1 - \sum_{k=1}^{K-1} \alpha_k \right)}_{\text{simplex reparameterization}} \right]$$

note that only one term belongs to k .

If $k=1$, we only care about α_1 .

$$= \sum_{t=p+1}^n \frac{\partial}{\partial \alpha_k} \left[Z_{k,t} \log \alpha_k + Z_{K,t} \log \left(1 - \sum_{k=1}^{K-1} \alpha_k \right) \right]$$

$$= \sum_{t=p+1}^n \left(\frac{Z_{k,t}}{\alpha_k} + \frac{Z_{K,t}}{1 - \sum_{k=1}^{K-1} \alpha_k} (-1) \right)$$

$$= \sum_{t=p+1}^n \left(\frac{Z_{k,t}}{\alpha_k} - \frac{Z_{K,t}}{\alpha_K} \right) \quad \text{where } k=1, \dots, K-1$$

$$\ell = \sum_{t=p+1}^n \left\{ \sum_{k=1}^K z_{k,t} \left[\log \alpha_k - \log \sigma_k - \frac{\varepsilon_{k,t}^2}{2\sigma_k^2} \right] \right\}$$

$$\begin{aligned} \frac{\partial \ell}{\partial \phi_{k,i}} &= \frac{\partial}{\partial \phi_{k,i}} \sum_{t=p+1}^n \sum_{k=1}^K z_{k,t} \left(-\frac{\varepsilon_{k,t}^2}{2\sigma_k^2} \right) \\ &= \sum_{t=p+1}^n \frac{\partial \ell}{\partial \varepsilon_{k,t}} \frac{\partial \varepsilon_{k,t}}{\partial \phi_{k,i}} \end{aligned}$$

$$\frac{\partial \ell}{\partial \varepsilon_{k,t}} = z_{k,t} \left(-\frac{\varepsilon_{k,t}}{\sigma_k^2} \right) \quad \text{we only care about a particular } k.$$

$$\begin{aligned} \frac{\partial \varepsilon_{k,t}}{\partial \phi_{k,i}} &= \frac{\partial}{\partial \phi_{k,i}} \left(y_t - \sum_{r=0}^{p_k} \phi_{k,i} y_{t-r} \right) \quad \text{we only care about a particular } i. \\ &= -y_{t-i} \end{aligned}$$

$$\text{Let } x_{t,i} = u(y_{t,i})$$

$$u(y_{t,i}) = \begin{cases} 1 & \text{if } i=0 \\ y_{t-i} & \text{if } i>0 \end{cases}$$

$$\begin{aligned} \frac{\partial \ell}{\partial \phi_{k,i}} &= \sum_{t=p+1}^n z_{k,t} \left(-\frac{\varepsilon_{k,t}}{\sigma_k^2} \right) \left(-u(y_{t,i}) \right) \\ &= \sum_{t=p+1}^n \frac{z_{k,t} u(y_{t,i}) \varepsilon_{k,t}}{\sigma_k^2} \quad \text{where } k=1, \dots, K \\ &\quad \text{where } i=0, \dots, p_k \end{aligned}$$

$$l = \sum_{t=p+1}^n \left\{ \sum_{k=1}^K z_{k,t} \left[\log \sigma_k - \log \sigma_k - \frac{\varepsilon_{k,t}^2}{2\sigma_k^2} \right] \right\}$$

$$\begin{aligned} \frac{\partial l}{\partial \sigma_k} &= \sum_{t=p+1}^n -\frac{z_{k,t}}{\sigma_k} - \left[-\frac{z_{k,t} \varepsilon_{k,t}^2}{2\sigma_k^3} (-2) \right] \\ &= \sum_{t=p+1}^n -\frac{z_{k,t}}{\sigma_k} + \left(\frac{z_{k,t} \varepsilon_{k,t}^2}{\sigma_k^3} \right) \\ &= \sum_{t=p+1}^n \frac{z_{k,t}}{\sigma_k} \left(\frac{\varepsilon_{k,t}^2}{\sigma_k^2} - 1 \right) \quad k=1, \dots, K \end{aligned}$$

Suppose $\theta(\alpha, \phi, \sigma)$ is known.

$$z \leftarrow E[z | Y; \theta]$$

$$\begin{aligned} z_{k,t} &\leftarrow E[z_k | Y; \theta] = P(y_t = k | Y, \theta) \\ &= P(z_{k,t} = 1 | y_t, \theta) \\ &= z_{k,t} \end{aligned}$$

$$\begin{aligned} Q(\theta | \theta_0, x) &= E_{\theta_0} [\log L^t(\theta | x, z)] \\ &= E_{\theta_0} [\log f(y_t, z_t | \mathcal{F}_{t-1}; \theta)] \end{aligned}$$

$$Q = E_{Z_k|Y,\theta} \left[\sum_{k=1}^K Z_{k,e} \left[\log \alpha_k + \log f_k(y_e | \mathcal{F}_{e-1}) \right] \right]$$

$$Q = E_{Z_k|Y,\theta} \left[\sum_{k=1}^K Z_{k,e} \left[\log \alpha_k - \frac{1}{2} \log(2\pi) - \log \sigma_k - \frac{(y_e - \mu_{k,e})^2}{2\sigma_k^2} \right] \right]$$

Suppose $\theta(\alpha, \phi, \sigma)$ is known. **E-STEP**

$$\begin{aligned} Q &= \sum_{k=1}^K E[Z_k | Y; \theta] \left[\log \alpha_k - \frac{1}{2} \log(2\pi) - \log \sigma_k - \frac{(y_e - \mu_{k,e})^2}{2\sigma_k^2} \right] \\ &= \underbrace{\sum_{k=1}^K \tau_{k,e} \log \alpha_k + \sum_{k=1}^K \tau_{k,e} \log f_k(y_e | \mathcal{F}_{e-1})}_{\text{Note that } Q \text{ depends on } \tau_{k,e}.} \end{aligned}$$

Note that Q depends on $\tau_{k,e}$.

\therefore our E-step should be $\tau_{k,e}$.

$\tau_{k,e} = P(Z_{k,e}=1 | y_e; \theta)$, obtaining exact value \rightarrow PDF

$$= \frac{f(y_e, Z_{k,e}=1 | \theta)}{f(Z_{k,e} | \theta)}, \quad \begin{array}{l} \text{Joint distribution} \\ \text{Marginal distribution} \end{array}$$

$$= \frac{f(y_e | Z_{k,e}=1, \theta) P(Z_{k,e}=1 | \theta)}{\sum_{k=1}^K f(y_e | Z_{k,e}=1, \theta) P(Z_{k,e}=1 | \theta)}, \quad P(Z_{k,e}=1 | \theta) = \alpha_k \text{ by design of mixture models.}$$

$$f(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} \quad \phi(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} \quad f(x) = \frac{1}{\sigma} \phi\left(\frac{x-\mu}{\sigma}\right)$$

$$= \frac{\alpha_k (1/\sigma_k) \phi(e_{k,e}/\sigma_k)}{\sum_{k=1}^K \alpha_k (1/\sigma_k) \phi(e_{k,e}/\sigma_k)} \quad \text{where } e = 1, \dots, K$$

M-Step Assume Z is known

$$\frac{\partial \ell}{\partial \alpha_k} = \sum_{t=p+1}^n \left(\frac{z_{t,k}}{\alpha_k} - \frac{z_{t,k}}{\alpha_k} \right) = 0 \text{ where } k=1, \dots, K-1$$

$$\frac{\sum_{t=p+1}^n z_{t,k}}{\alpha_k} - \frac{\sum_{t=p+1}^n z_{t,k}}{\alpha_k} = 0,$$

$$\text{Let } \sum_{t=p+1}^n z_{t,k} = N_k$$

$$\sum_{t=p+1}^n \bar{z}_{t,k} = N_k$$

$$\alpha_k = \frac{N_k}{N_k} \triangle \alpha_k$$

$$\sum_{k=1}^{K-1} \alpha_k + \alpha_K = 1,$$

$$\alpha_1 + \alpha_2 + \alpha_3 + \dots + \alpha_K = 1$$

$$\sum_{k=1}^{K-1} \frac{N_k}{N_k} \alpha_k + \alpha_K = 1$$

$$\alpha_K \left(\sum_{k=1}^{K-1} \frac{N_k}{N_k} + 1 \right) = 1,$$

$$1 = \frac{N_K}{N_K}$$

$$\alpha_K \left(\frac{1}{N_K} \sum_{k=1}^K N_k \right) = 1$$

$$\triangle \alpha_K = \frac{N_K}{\sum_{k=1}^K N_k}$$

$$\alpha_k = \frac{N_k}{\sum_{p=1}^K N_p},$$

$$\sum_{p=1}^K \sum_{t=p+1}^n z_{t,k} = \sum_{t=p+1}^n 1$$

finite sums are commutative

$$\sum_{p=1}^K P(z_{t,k}=1 | y_t, \theta)$$

$$\alpha_k = \frac{N_k}{\sum_{t=p+1}^n 1}$$

$$\therefore \hat{\alpha}_k = \frac{\sum_{t=p+1}^n z_{t,k}}{n-p} \quad k=1, \dots, K$$

M-Step Assume Z is known

$$\frac{\partial \mathcal{L}}{\partial \sigma_k} = \sum_{t=p+1}^n \frac{Z_{t,k}}{\sigma_k} \left(\frac{\varepsilon_{t,k}^2}{\sigma_k^2} - 1 \right) \quad k=1, \dots, K$$

where $\varepsilon_{t,k} = y_t - \phi_{k,0} - \phi_{k,1}y_{t-1} - \dots - \phi_{k,p_k}y_{t-p_k}$

$$\sum_{t=p+1}^n \frac{Z_{t,k}}{\sigma_k} \left(\frac{\varepsilon_{t,k}^2}{\sigma_k^2} - 1 \right) = 0$$

$$\frac{1}{\sigma_k^2} \sum_{t=p+1}^n Z_{t,k} \varepsilon_{t,k}^2 = \sum_{t=p+1}^n Z_{t,k}$$

$$\hat{\sigma}_k = \left(\frac{\sum_{t=p+1}^n Z_{t,k} \varepsilon_{t,k}^2}{\sum_{t=p+1}^n Z_{t,k}} \right)^{\frac{1}{2}}$$

$$\frac{\partial \mathcal{L}}{\partial \phi_{k,i}} = \sum_{t=p+1}^n \frac{Z_{t,k} u(y_t, i) \varepsilon_{t,k}}{\sigma_k^2} \quad \text{where } k=1, \dots, K \quad i=0, \dots, p_k$$

$$\sum_{t=p+1}^n Z_{t,k} u(y_t, i) \varepsilon_{t,k} = 0 \quad \boxed{u(y_t, i) = \begin{cases} 1 & \text{if } i=0 \\ y_{t-i} & \text{if } i>0 \end{cases}}$$

$$\sum_{t=p+1}^n Z_{t,k} u(y_t, i) y_t = \sum_{t=p+1}^n Z_{t,k} u(y_t, i) \sum_{j=0}^{p_k} \phi_{k,j} u(y_t, j)$$

$$\sum_{t=p+1}^n Z_{t,k} u(y_t, i) y_t = \sum_{j=0}^{p_k} \phi_{k,j} \sum_{t=p+1}^n Z_{t,k} u(y_t, j) u(y_t, i)$$

where $i=0, \dots, p_k$

$$\sum_{t=p+1}^n \tau_{t,k} u(y_t, i) y_t = \sum_{j=0}^{p_k} \hat{\phi}_{k,j} \sum_{t=p+1}^n \tau_{t,k} u(y_t, j) u(y_t, i)$$

where $i = 0, \dots, p_k$

$$\begin{bmatrix} \sum \tau y_t u_0 \\ \sum \tau y_t u_1 \\ \vdots \\ \sum \tau y_t u_{p_k} \end{bmatrix}_{b_k} = \begin{bmatrix} \sum \tau u_0 u_0 & \sum \tau u_0 u_1 & \dots & \sum \tau u_0 u_{p_k} \\ \sum \tau u_1 u_0 & & & \\ \vdots & & & \\ \sum \tau u_{p_k} u_0 & \dots & \dots & \sum \tau u_{p_k} u_{p_k} \end{bmatrix}_{A_k} \begin{bmatrix} \hat{\phi}_{k_0} \\ \vdots \\ \hat{\phi}_{k_{p_k}} \end{bmatrix}_{\hat{\phi}_k}$$

Note

τ is the probability of observing y_t in cluster k .
 $u_0 = u(y_t, 0) = 1 \quad \& \quad u_i = u(y_t, i) = y_{t-i}$
 \therefore weighted least squares. $A_k \hat{\phi}_k = b_k$

$$A_k = X_k^T W_k X_k \quad b_k = X_k^T W_k Y$$

$$X_k = \begin{bmatrix} 1 & y_{t-1} & y_{t-2} & \dots & y_{t-p_k} \\ 1 & y_t & y_{t-1} & \dots & y_{t+1-p_k} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & y_{n-1} & y_{n-2} & \dots & y_{n-p_k} \end{bmatrix} \rightarrow \begin{matrix} n-p \text{ rows} \\ p_k+1 \text{ cols} \end{matrix}$$

$$Y = [y_t \ y_{t+1} \ \dots \ y_n] \quad W_k = \text{diag}(\tau_{t,k})$$

where $t = p+1, p+2, \dots, n$

$$\boxed{\hat{\phi}_k = A_k^{-1} b_k}$$

Log-Likelihood

$$f(y_e | \mathbf{x}_{e-1}) = \frac{1}{\sqrt{2\pi} \sigma_k} e^{-\frac{(y_e - \mu_{\epsilon,k})^2}{2\sigma_k^2}}$$

$$L(\theta) = \prod_{t=p+1}^n \left[\sum_{k=1}^K \alpha_k \frac{1}{\sqrt{2\pi} \sigma_k} e^{-\frac{(y_t - \mu_{\epsilon,k})^2}{2\sigma_k^2}} \right]$$

$$\ell(\theta) = \sum_{t=p+1}^n \log \left[\sum_{k=1}^K \alpha_k \frac{1}{\sqrt{2\pi} \sigma_k} e^{-\frac{(y_t - \mu_{\epsilon,k})^2}{2\sigma_k^2}} \right]$$

$$= \sum_{t=p+1}^n \log \left[\sum_{k=1}^K \alpha_k \frac{1}{\sigma_k} e^{-\frac{(y_t - \mu_{\epsilon,k})^2}{2\sigma_k^2}} \right] + \boxed{\log(2\pi)^{-\frac{n-p}{2}}}$$

Constant can be removed.

$$\therefore \ell(\theta) = \sum_{t=p+1}^n \log \left[\sum_{k=1}^K \alpha_k \frac{1}{\sigma_k} e^{-\frac{\epsilon_{t,k}^2}{2\sigma_k^2}} \right]$$

$$\varepsilon_t = \Delta y_t - \phi_0 - \phi_1 \Delta y_{t-1} - \phi_2 \Delta y_{t-2}$$

$$= y_t - y_{t-1} - \phi_0 - \phi_1 (y_{t-1} - y_{t-2}) - \phi_2 (y_{t-2} - y_{t-3})$$

$$= y_t - \phi_0 - y_{t-1} - \phi_1 y_{t-1} + \phi_1 y_{t-2} - \phi_2 y_{t-2} + \phi_2 y_{t-3}$$

$$= y_t + (-1 - \phi_1) y_{t-1} + (\phi_1 - \phi_2) y_{t-2} + \phi_2 y_{t-3} - \phi_0$$

$$\text{Let } \phi_{11} = -0.3208, \phi_{21} = 0.6711$$

$$\phi_{0k} = 0$$

$$\varepsilon_{t,1} = y_t - \phi_{01} - (1 + \phi_{11}) y_{t-1} + \phi_{11} y_{t-2}$$

$$= y_t - 0.6792 y_{t-1} - 0.3208 y_{t-2}$$

$$\varepsilon_{t,2} = y_t - 1.6711 y_{t-1} + 0.6711 y_{t-2}$$

$$\varepsilon_{t,3} = \Delta y_t - \phi_0$$

$$= y_t - y_{t-1}$$