## Self-paced Example: Correlated random variables

Module 10

MCIT Online - CIT592 - Professor Val Tannen

This is a segment that contains material meant to be learned at your own pace. We are trying to assist you in this endeavor by organizing the material in a manner similar to the way it is outlined in the recorded segments, however with one additional suggestion. When you see the following marker:



we suggest that you stop and make sure you thoroughly understood the material presented so far before you proceed further.

### Product of two r.v.'s

Let  $X,Y:\Omega\to\mathbb{R}$  be two random variables defined on the same probability space  $(\Omega,\Pr)$ . Their **product**, notation XY, is the random variable on the same space defined by:

$$(XY)(w) = X(w) \cdot Y(w) \quad \forall \ w \in \Omega$$

**Example.** A fair coin is flipped twice. Recall that we denoted by  $X_H$  and  $X_T$  the random variables that return, respectively, the number of heads and tails observed.

The probability space of two independent flips of a fair coin is uniform and has the outcomes  $\{HH, HT, TH, TT\}$  each with probability 1/4.

Then  $X_H X_T(HH) = X_H X_T(TT) = 0$  and  $X_H X_T(HT) = X_H X_T(TH) = 1$ .

So  $X_H X_T$  is a Bernoulli r.v. with parameter 1/2.



### Correlated random variables

**Proposition** Var[X + Y] = Var[X] + Var[Y] iff  $E[XY] = E[X] \cdot E[Y]$ .

**Proof** Applying LOE multiple times we have

$$\begin{aligned} \mathsf{Var}[X+Y] &= \mathsf{E}[(X+Y)^2] - (\mathsf{E}[X+Y])^2 \\ &= \mathsf{E}[X^2 + 2XY + Y^2] - (\mathsf{E}[X] + \mathsf{E}[Y])^2 \\ &= \mathsf{E}[X^2] + 2\mathsf{E}[XY] + \mathsf{E}[Y^2] - (\mathsf{E}[X])^2 - 2(\mathsf{E}[X] \cdot \mathsf{E}[Y]) - (\mathsf{E}[Y])^2 \\ &= \mathsf{Var}[X] + \mathsf{Var}[Y] - 2(\mathsf{E}[XY] - \mathsf{E}[X] \cdot \mathsf{E}[Y]) \end{aligned}$$

This suggests that we *define*: the random variables X and Y are **correlated** when  $E[XY] \neq E[X] \cdot E[Y]$ . Thus variance distributes over the sum of uncorrelated r.v.'s.

To test yourself, think of a simple example of two specific correlated random variables.



**Example.** The r.v.'s  $X_H$  and  $Y_H$  defined earlier are correlated. Indeed, we have calculated in a previous lecture that  $\mathsf{E}[X_H] = \mathsf{E}[X_T] = 1$ . However,  $X_H X_T$  is Bernoulli with parameter 1/2 so  $\mathsf{E}[X_H X_T] = 1/2 \neq 1 = \mathsf{E}[X_H] \cdot \mathsf{E}[X_T]$ .

Notice how a higher  $X_H$  means a lower  $X_T$ . This case is referred to as a **negative correlation**. To give this notion mathematical precision, define **covariance** as:

$$Cov(X,Y) = E[XY] - E[X]E[Y]$$

Therefore, two random variables are uncorrelated if and only if their covariance is 0. By contrast, in our example,  $Cov(X_H, X_T) = \frac{1}{2} - 1 = -\frac{1}{2}$ , a negative number. If the covariance had been positive, then we would have had a **positive correlation** between the variables.



# Correlated versus independent random variables

Recall that two random variables are independent when for any x in Val(X) and any y in Val(Y),

$$\Pr[(X = x) \cap (Y = y)] = \Pr[X = x] \cdot \Pr[Y = y]$$

**Proposition** If two random variables are independent then they are uncorrelated.

**Proof** Suppose the random variables X and Y are independent. By the definition of product of random variables,

$$\mathsf{E}[XY] = \sum_{x \in \mathrm{Val}(X)} \sum_{y \in \mathrm{Val}(Y)} \Pr[(X = x) \cap (Y = y)] \cdot x \cdot y$$

By the definition of independent random variables then,

$$\mathsf{E}[XY] = \sum_{x \in \mathrm{Val}(X)} \sum_{y \in \mathrm{Val}(Y)} \Pr[X = x] \cdot \Pr[Y = y] \cdot x \cdot y$$

By rearranging the terms that depend on x and on y:

$$\mathsf{E}[XY] = \sum_{x \in \mathrm{Val}(X)} \Pr[X = x] \cdot x \cdot \sum_{y \in \mathrm{Val}(Y)} \Pr[Y = y] \cdot y$$

Therefore, by definition of expectation,  $E[XY] = E[X] \cdot E[Y]$ .



Now, can you think of two random variables that are uncorrelated but not independent? This is possible! **Example** Let X be a random variable that takes the values -1, 0, and 1 each with probability  $\frac{1}{3}$ . Hence  $\mathsf{E}[X] = 0$ . Let  $Y = X^2$  be our second random variable. Note that  $X^n$  whenever n is odd has the exact same distribution as X. Computing the covariance:

$$Cov(X, X^2) = E[X^3] - E[X] \cdot E[X^2] = 0 - 0 \cdot E[X^2] = 0$$

We have shown X and Y are uncorrelated, but we have to show they are not independent. Consider two events: (X=0) and (Y=1). Note that  $\Pr[X=0 \text{ and } Y=1]=0$ . However,

$$\Pr[X = 0] \cdot \Pr[Y = 1] = \frac{1}{3} \cdot \frac{2}{3} \neq 0$$

Hence the random variables are dependent. This result might seem counterintuitive, but it is explained by the fact that correlation is a measure of *linear* dependence. In our example, the overall covariance between X and Y may be 0, but this is only because the negative contribution from X=-1 exactly cancels the positive contribution from X=1.



#### More than two random variables

We have shown above that if two random variables are uncorrelated then variance distributed over their sum. That is, Var[X + Y] = Var[X] + Var[Y] when (in fact, iff)  $E[XY] = E[X] \cdot E[Y]$ . Now we will extend this to more than two random variables.

**Proposition** If  $X_1, \ldots, X_n$  are pairwise uncorrelated, then:

$$\mathsf{Var}[X_1+\dots+X_n] \ = \ \mathsf{Var}[X_1]+\dots+\mathsf{Var}[X_n]$$

**Proof** To keep the notation lighter we just sketch this for three random variables. Using linearity of expectation, you can see that (verify!):

$$\begin{aligned} \mathsf{Var}[\,X+Y+Z\,] &=& \mathsf{Var}[X] + \mathsf{Var}[Y] + \mathsf{Var}[Z] + \\ &=& 2(\mathsf{E}[XY] - \mathsf{E}[X]\mathsf{E}[Y]) + 2(\mathsf{E}[YZ] - \mathsf{E}[Y]\mathsf{E}[Z]) + 2(\mathsf{E}[ZX] - \mathsf{E}[Z]\mathsf{E}[X]) \end{aligned}$$

Since X, Y, Z are pairwise uncorrelated the last three terms in the sum are 0 and the identity follows.

