

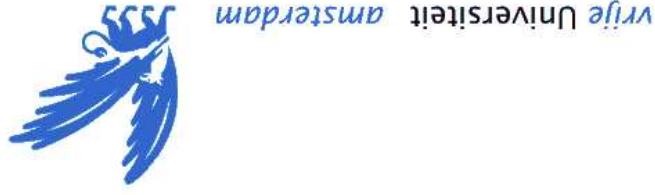
# Convergence rates of posterior distributions for Brownian semimartingale models

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## Outline of talk

Convergence rates of posterior distributions for Brownian semimartingale models

- Brownian semimartingale Model (BSM)  
examples

- Bayesian estimation  
prior, likelihood, posterior - consistency - rate of convergence

- **Main result:** gives lower bounds on the Bayesian rate of convergence for the BSM

- Example: ergodic diffusion

# Brownian semimartingale model

- $\Theta$ : parameter set.  $\theta_0$  true value  
 $B$ : standard Brownian motion
- $\beta_{\theta,n}$  and  $\sigma_n$  (suitably measurable) processes  
 $X_n$ : weak solution of
- $$dX_n^t = \beta_{\theta_0,n}(t, X_n)dt + \sigma_n dB_t, \quad t \in [0, T_n].$$
- Suppose we observe  $X_n$  continuously on  $[0, T_n]$ . Estimate  $\theta_0$ .
- Behavior of **Bayesian estimators** in the large-sample limit  
 $n \rightarrow \infty$ .

## Examples

- Gaussian white-noise model

$$dX_n^t = \theta(t)dt + \sigma_n(t)dB_t, \quad t \in [0, T],$$

where  $(\sigma_n)$  is a sequence of positive numbers tending to zero, as  $n \rightarrow \infty$ .

- Perturbed dynamical system

$$dX_n^t = \theta(X_n^t)dt + \sigma_n(t)dB_t, \quad t \in [0, T],$$

where  $(\sigma_n)$  is a sequence of positive numbers tending to zero, as  $n \rightarrow \infty$ .

- Ergodic diffusion model

$$dX^t = \beta_\theta(X^t)dt + \sigma(X^t)dB_t, \quad t \in [0, T_n],$$

where  $T_n \rightarrow \infty$ .

## Motivation

In practice one never observes such a process continuously, so why study estimators for it?

- Limit model of “high-frequency” data
- Results provide an upper bound on what can be achieved.
- Analysis of the same problem for discrete time observations is harder.

- Bayesian computations for the BSM based on discrete time observations are promising, and feasible by MCMC-algorithms. However, no theoretical justification.
- Certain forms of the Gaussian white-noise model are “statistically equivalent” to regression problems.

# Bayesian estimation (1)

Ingredients:

- Prior distribution  $\Pi$  on  $\Theta$ .
- Likelihood function  $L$ : given observation  $X$ ,  $L(X; \theta)$  is

$$L(\theta; X_n) = \exp \left( \int_{T_n}^0 \beta_{\theta,n} \frac{(\sigma_n)_2}{2} dX_n - \frac{1}{2} \int_{T_n}^0 \left( \frac{\sigma_n}{\beta_{\theta,n}} \right)^2 dt \right)$$

- Posterior distribution  $\Pi(\cdot | X_n)$  obtained via Bayes' Theorem: conditional distribution of  $\Theta$ , given  $X_n$

## Bayesian estimation (2)

- **Consistency:** “The posterior distribution concentrates arbitrarily close to the true one, as we obtain more data”.

$$\forall \varepsilon > 0 \quad \Pi_n(d(\theta, \theta_0) \geq \varepsilon | X^n) \rightarrow 1, \quad P_{\theta_0} - a.s.$$

- **Bayesian convergence rate:** “Maximal rate at which we

can shrink balls around  $\theta_0$ , while still capturing almost

all posterior mass”

Let  $\varepsilon_n \uparrow 0$  ( $n \rightarrow \infty$ ). The sequence of posteriors converges to  $\theta_0$  (at least) at rate  $\varepsilon_n$  if for all sequences  $M_n \rightarrow \infty$ ,

$$\Pi_n(\theta \in \Theta : d(\theta, \theta_0) \geq M_n \varepsilon_n | X^n) \rightarrow 0$$

in  $P_{\theta_0}^{(n)}$ -probability.

- Derive conditions from which we can infer the convergence rate.

## Type of conditions

To obtain consistency and rates of convergence:

- **Condition on the prior** (sufficient mass near the true value)

If the prior  $\Pi$  excludes  $\theta_0$  consistency cannot even occur!

- **Entropy conditions**

Birgé and Le Cam:

- $d_n, e_n$ : semimetrics on  $\Theta^n$ .
- $\varepsilon_n$ : sequence of positive numbers  $\downarrow 0$ .
- There exist estimators  $\hat{\theta}_n = \hat{\theta}_n(X^n)$  such that  $d_n(\hat{\theta}_n, \theta_0) = O_P(\varepsilon_n)$  if  $\sup_{\varepsilon > \varepsilon_n} \log N(\varepsilon, \{\theta : d_n(\theta, \theta_0) \leq \varepsilon\}, e_n) \leq n\varepsilon_n^2$ .



## Working assumption

- Measure distance on  $\Theta$  by the **Hellinger** (semi)metric
 
$$h_2^n(\theta, \psi) = \int_0^T \left( \frac{\sigma_n}{\beta_{\theta,n} - \beta_{\psi,n}} \right)^2 dt, \quad \theta, \psi \in \Theta.$$

- There exist semimetrics  $d_n$  and  $\bar{d}_n$  on  $\Theta$  such that, as  $n$  gets large, with large probability

$$c_n d(\theta, \psi) \lesssim h_n(\theta, \psi) \lesssim c_n \bar{d}(\theta, \psi),$$

for  $c_n$  a sequence of positive numbers.

- $B(\theta_0, \varepsilon)$ : ball of  $d$ -radius  $\varepsilon$  around  $\theta_0$ .  $\bar{B}(\theta_0, \varepsilon)$ : for ball of  $\bar{d}$ -radius  $\varepsilon$  around  $\theta_0$

## Main result

- $\varepsilon^n$ : sequence of positive numbers such that  $c_n \varepsilon_n$  is bounded away from zero.

- Assume for every  $a > 0$  there exists a constant  $g(a) < \infty$  such that

$$\sup_{\varepsilon > \varepsilon_n} \log N(a\varepsilon, B(\theta_0, \varepsilon), \bar{d}) \leq (c_n \varepsilon_n)^2 g(a).$$

If for every  $\xi > 0$  there is a constant  $J$  such that for  $j \geq J$

$$\frac{\Pi_n(B(\theta_0, j\varepsilon_n))}{\Pi_n(B(\theta_0, \varepsilon_n))} \leq e^{\xi c_n^2 \varepsilon_n j^2}.$$

- Then for every  $M_n \rightarrow \infty$ , we have

$$P_{\theta_0}^{(n)}[\Pi_n(\theta) \in \Theta_n : h_n(\theta, \theta_0) \geq M_n c_n \varepsilon_n | X_n] \rightarrow 0.$$

## Example: Ergodic diffusion

- Under “conditions”

$$dX_t = \beta_\theta(X_t)dt + \sigma(X_t)dB_t, \quad t \in [0, T_n],$$

defines a regular diffusion in an interval  $I$

- **semimetrics:** measure

$$\bar{d}(\theta, \psi) = \left\| \frac{\sigma}{\theta - \psi} \right\|_{L^2(\mu_0)}, \quad d(\theta, \psi) = \left\| \frac{\sigma}{\theta - \psi} 1_J \right\|_{L^2(\mu_0)}.$$

( $\mu_0$  is the invariant probability measure,  $J \subseteq I$  compact)

- **Lipschitz condition on drift.** Assume

$$\bar{b}(x)\|\theta - \psi\| \leq |b_\theta(x) - b_\psi(x)| \leq \bar{b}(x)\|\theta - \psi\|, \quad \forall x \in I.$$

- If  $\Theta \subseteq \mathbb{R}^k$  bounded and the prior density is bounded away from zero near  $\theta_0$ , then as  $T_n \rightarrow \infty$

$$P_{\theta_0, n} \Pi_n(\theta) : \|\theta - \theta_0\| \geq M_n / \sqrt{T_n} |X_n| \rightarrow 0.$$