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# Combinations

$$\binom{n}{k} = \frac{n(n-1)(n-2)\dots(n-k-1)}{k!} = \frac{n!}{(n-k)!k!}$$

# **Conditional Probability**

 $P(E \mid F) =$  the probability of E given F has already happened

$$P(E \mid F) = \frac{P(E \cap F)}{P(F)}$$

## Chain Rule

$$P(E \cap F) = P(E \mid F)P(F)$$

### Marginalisation

$$P(E) = P(E \cap F_1) + P(E \cap F_2) + \dots + P(E \cap F_n)$$

$$P(E) = P(E \mid F_1)P(F_1) + P(E \mid F_2)P(F_2) + \dots + P(E \mid F_n)P(F_n)$$

given

- $F_1, F_2, \ldots, F_n$  are mutually exclusive
- $F_1 \cup F_2 \cup \cdots \cup F_n = S$

### Bayes Rule

$$P(E \mid F) = \frac{P(F \mid E)P(E)}{P(F)}$$

# **Independent Events**

$$P(E \cap F) = P(E)P(F)$$

$$P(E \mid F) = P(E)$$

### Conditionally Independent

$$P(E \cap F \mid G) = P(E \mid G)P(F \mid G)$$

## Binomial Random Variable

Sum of i successes out of n trials.

$$P(X = i) = \binom{n}{i} p^{i} (1 - p)^{n-i}, i = 0, 1, \dots, n$$

# Expected Value of Random Variable

$$E[X] = \sum_{i=1}^{n} x_i P(X = x_i)$$

### Variance

$$Var(X) = \sum_{i=1}^{n} (x_i - \mu)^2 P(x_i)$$

with  $\mu = E[X]$ 

## Covariance

Say  $E[X] = \mu_x$  and  $E[Y] = \mu_y$  then

$$Cov(X, Y) = E[(X - \mu_x)(Y - \mu_y)] = E[XY] - E[X]E[Y]$$

# Correlation

$$Corr(X,Y) = \frac{Cov(X,Y)}{\sqrt{Var(X)Var(Y)}}$$

# Inequalities

Markov

$$P(X \ge a) \le \frac{E(X)}{a}$$
 for all  $a > 0$ 

Chebyshev

$$P(\mid X - \mu \mid \geq k) \leq \frac{\sigma^2}{k^2}$$
 for all  $k > 0$ 

Chernoff

$$P(X \ge a) \le \min_{t>0} e^{ta} e^{\log E(e^{tX})}$$

Binomial RV

$$P(X \ge (1 - \delta)np) \le e^{-np((1+\delta)\log(1+\delta) - \delta)}$$
$$P(X \ge (1 - \delta)\mu) \le e^{-\mu((1+\delta)\log(1+\delta) - \delta)}$$

# Distribution of Sample Mean

$$\bar{X} = \frac{1}{N} \sum_{k=1}^{N} X_k$$

**Expected Value** 

$$E[\bar{X}] = \frac{1}{N} \sum_{k=1}^{N} E[X_k]$$

Variance

$$Var(\bar{X}) = \frac{\sigma^2}{N}$$

# Weak Law of Large Numbers

$$P(|\bar{X} - \mu| \ge \epsilon) \to 0 \text{ as } N \to \infty$$

By Chebyshev's inequality:

$$P(\mid \bar{X} - \mu \mid \geq \epsilon) \leq \frac{\sigma^2}{N\epsilon^2}$$

# Continuous Random Variables

## Cumulative and Probability Distribution Function

CDF is  $F_Y(y)$ , PDF is  $f_Y(y)$ 

$$P(a < Y \le b) = F_Y(b) - F_Y(a)$$
$$F_Y(y) = \int_{-\infty}^y f_Y(t)dt$$
$$P(a < Y \le b) = \int_a^b f_Y(t)dt$$

#### Independent CDF

$$P(X \leq x \land Y \leq y) = P(X \leq x)P(Y \leq y)$$

i.e.

$$F_{XY}(x,y) = F_X(x)F_Y(y)$$
 
$$F_{XY}(x,y) = \int_{-\infty}^x \int_{-\infty}^y f_{XY}(u,v)du dv$$

#### **Conditional PDF**

$$f_{X|Y}(x \mid y) = \frac{f_{XY}(x,y)}{f_Y(y)}$$

### Chain Rule for PDF

$$f_{XY}(x,y) = f_{X|Y}(x \mid y)f_Y(y) = f_{Y|X}(y \mid x)f_X(x)$$

### Marginalisation of PDFs

$$f_X(x) = \int_{-\infty}^{\infty} f_{XY}(x, y) dy$$

#### Bayes Rule for PDFs

$$f_{Y\mid X}(y\mid x) = \frac{f_{X\mid Y}(x\mid y)f_{Y}(y)}{f_{X}(x)}$$

## Normal Distribution

 $Y \sim N(\mu, \sigma^2)$  when it has PDF

$$f_Y(y) = \frac{1}{\sigma\sqrt{(2\pi)}}e^{-\frac{(y-\mu)^2}{2\sigma^2}}$$

### Central Limit Theorem

$$\bar{X} \sim N(\mu, \frac{\sigma^2}{N})$$
 as  $N \to \infty$ 

## Linear Regression Model

$$Y = \sum_{i=1}^{m} \Theta^{(i)} x^{(i)} + M$$

where  $\vec{\Theta}$  is a vector of unknown (random) parameters and M is random noise

- M is Gaussian with mean 0 and variance 1,  $M \sim N(0, 1)$
- $\Theta^{(i)}$  is Gaussian with mean 0 and variance  $\lambda$  (where  $\lambda$  is known),  $\Theta^{(i)} \sim N(0,\lambda)$

This is equivalent to

$$f_{Y|X,\vec{\Theta}}(y \mid x, \vec{\theta}) = \frac{1}{\sqrt{2\pi}} \exp(-(y - \sum_{i=1}^{m} \theta^{(i)} x^{(i)})^2 / 2)$$

given  $\vec{\Theta} = \vec{\theta}$ . Model also assumes  $\Theta^{(i)} \sim N(0, \lambda)$  i.e.

$$f_{\Theta^{(i)}}(\theta) \propto \exp(-\theta^2/2\lambda)$$

#### **Parameter Estimation**

$$f_{\Theta|D}(\vec{\theta} \mid d) = \frac{f_{D|\Theta}(d \mid \vec{\theta}) f_{\Theta}(\vec{\theta})}{f_{D}(d)}$$

### Maximum Likelihood Estimation

$$\log f_{D|\Theta}(d \mid \vec{\theta}) \propto \log L(\theta) = -\frac{1}{2} \sum_{j=1}^{n} (y_j - \sum_{i=1}^{m} \theta^{(i)} x_j^{(i)})^2$$
$$\theta = \frac{\sum_{j=1}^{n} y_j x_j}{\sum_{j=1}^{n} x_j^2}$$

### Maximum a Posteriori (MAP) Estimation

$$\theta = \frac{\sum_{j=1}^{n} y_j x_j}{\frac{1}{\lambda} + \sum_{j=1}^{n} x_j^2}$$

# Logistic Regression Model

$$P(Y = 1 \mid \Theta = \vec{\theta}, \vec{X} = \vec{x}) = \frac{1}{1 + \exp(-z)}, z = \sum_{i=1}^{m} \theta^{(i)} x^{(i)}$$
 
$$P(Y = 0 \mid \Theta = \vec{\theta}, \vec{X} = \vec{x}) = \frac{\exp(-z)}{1 + \exp(-z)}$$

### Parameter Estimation

$$P(D=d\mid\Theta=\vec{\theta}) = \prod_{k=1}^{n} (\frac{1}{1+\exp(-z_k)})^{y_k} (\frac{\exp(-z_k)}{1+\exp(-z_k)})^{1-y_k}$$
 where  $z_k = \sum_{i=1}^{m} \theta^{(i)} x_k^{(i)}$ 

### **Maximum Likelihood Estimation**

$$P(Y = 1 \mid \Theta = \infty, \vec{X} = \vec{x}) = \frac{1}{1 + \exp(-z)}, z = \theta^{(1)} x^{(1)} \begin{cases} 1 & x^{(1)} = -1 \\ 0 & x^{(1)} = 0 \end{cases}$$
$$P(Y = 0 \mid \Theta = \infty, \vec{X} = \vec{x}) = 1 - P(Y = 1 \mid \Theta = \infty, \vec{X} = \vec{x})$$