



Lancaster University College  
at Beijing Jiaotong University

## 2019/20 Examinations

Course code: [WB73L004Q](#)

Course name: [Probability Theory and Mathematical Statistics \(B\)](#)

Resit examination (March)

### INSTRUCTIONS TO STUDENTS

- 1) Duration of the exam: [120 minutes](#)
- 2) This paper contains [3](#) pages. There are [8](#) questions.
- 3) You must answer all questions.
- 4) This is a closed book exam. No books or notes may be brought into the exam room.
- 5) A scientific calculator is allowed in the examination. Other electronic devices are not allowed in the exam room.
- 6) Some values that might be useful:

$$\Phi(0.5634) = 0.7134, \quad \Phi(1) = 0.8413,$$

$$\Phi(1.3363) = 0.9093, \quad t_{0.05}(4) = 2.1318.$$

**1. (20pt)**

- (1) The joint PDF of  $X$  and  $Y$  is  $f(x, y) = \begin{cases} ae^{-2x-4y}, & x > 0, y > 0, \\ 0, & \text{otherwise.} \end{cases}$  then  $a = 8$ .
- (2) A family is chosen at random from all three-child families. If it is known that the family has a boy among the three children, then the conditional probability that the chosen family has one boy and two girls is  $3/7$ .
- (3) If random variables  $X \sim B(4, 0.3)$ ,  $Y \sim B(11, 0.3)$ , and  $X, Y$  are independent, then  $X + Y \sim B(15, 0.3)$ .
- (4) If random variables  $X, Y$  have a joint continuous distribution, then  $P(X = Y) = 0$ .
- (5) If random variable  $X \sim U(1, 3)$ , then  $\text{Var}(X) = 1/3$ .
- (6) Suppose that  $X_1, X_2$  have a bivariate normal distribution. Then  $X_1, X_2$  are independent if and only if  $\text{Cov}(X_1, X_2) = 0$ . (or  $\rho_{X_1, X_2} = 0$ )
- (7) Let  $X_1, X_2, \dots, X_n$  be a random sample from  $X$  with  $E(X) = \mu$  and  $\text{Var}(X) = \sigma^2$ . Then  $E[(X_2 - X_1)^2] = 2\sigma^2$ , and if  $c \sum_{i=1}^{n-1} (X_{i+1} - X_i)^2$  is an unbiased estimator of  $\sigma^2$  then  $c = \frac{1}{2(n-1)}$ .
- (8) Let  $X_1, X_2, X_3, X_4$  be a random sample from distribution  $N(\mu, \sigma^2)$ .  $\bar{X}$  is the sample mean. Suppose that  $c \sum_{i=1}^4 (X_i - \bar{X})^2 \sim \chi^2(m)$ . Then  $c = \frac{1}{\sigma^2}$  and  $m = 3$ .

- 2. (10pt)** In a world, 40% of products are highly successful, 35% are moderately successful, and 25% are poor products. In addition, 95% of highly successful products receive good reviews, 60% of moderately successful products receive good reviews, and 10% of poor products receive good reviews.

- (1) What is the probability that a product attains a good review?
- (2) If a new product attains a good review, what is the probability that it will be a highly successful product?
- (3) If a product does not attain a good review, what is the probability that it will be a highly successful product?

**Solution:** Let H denote the event of being highly successful, M the event of being moderately successful, and P the event of being poor. Let G denote the event of having a good review.

$$(1) P(G) = P(G|H)P(H) + P(G|M)P(M) + P(G|P)P(P) \\ = 0.95 \times 0.4 + 0.6 \times 0.35 + 0.1 \times 0.25 = 0.615.$$

$$(2) P(H|G) = \frac{P(HG)}{P(G)} = \frac{P(G|H)P(H)}{P(G)} = \frac{0.95 \times 0.4}{0.615} = 0.6179.$$

$$(3) P(H|\bar{G}) = \frac{P(H\bar{G})}{P(\bar{G})} = \frac{P(H)-P(HG)}{1-P(G)} = \frac{0.4-0.38}{1-0.615} = 0.0325.$$

3. **(10pt)** Two random variables  $X$  and  $Y$  are independent,  $X \sim N(2,4)$ ,  $Y \sim N(3,5)$ . Determine  $P(3X + 2Y \leq 22)$ .

**Solution:** It is clear that  $3X + 2Y \sim N(3 \times 2 + 2 \times 3, 9 \times 4 + 4 \times 5) = N(12, 56)$

$$P(3X + 2Y \leq 22) = P\left(\frac{3X + 2Y - 12}{\sqrt{56}} \leq \frac{22 - 12}{\sqrt{56}}\right) = \Phi(1.3363) \\ = 0.9093.$$

4. **(12pt)** The joint PDF of  $X$  and  $Y$  is given by

$$f(x, y) = \begin{cases} 12y^2, & 0 \leq x \leq 1, 0 \leq y \leq x, \\ 0, & \text{otherwise.} \end{cases}$$

- (1) Find  $\text{Cov}(X, Y)$ .  
(2) Find  $P(X + Y \leq 1)$ .

**Solution:** (1)

$$E(X) = \iint xf(x, y) dx dy = \int_0^1 \int_0^x x \cdot 12y^2 dy dx = \int_0^1 4x^4 dx = \frac{4}{5} \\ E(Y) = \iint yf(x, y) dx dy = \int_0^1 \int_0^x y \cdot 12y^2 dy dx = \int_0^1 3x^4 dx = \frac{3}{5} \\ E(XY) = \iint xyf(x, y) dx dy = \int_0^1 \int_0^x xy \cdot 12y^2 dy dx = \int_0^1 3x^5 dx = \frac{1}{2} \\ \text{Cov}(X, Y) = E(XY) - E(X)E(Y) = \frac{1}{2} - \frac{4}{5} \cdot \frac{3}{5} = \frac{1}{50}.$$

(2)

$$P(X + Y \leq 1) = \int_0^{\frac{1}{2}} \int_y^{1-y} 12y^2 dx dy = \int_0^{\frac{1}{2}} 12y^2(1-2y) dy \\ = [4y^3 - 6y^4]_0^{\frac{1}{2}} = \frac{1}{8}.$$

5. **(13pt)** Let  $X$  denote the total number of successes in 15 Bernoulli trials, with probability of success  $p = 0.3$  on each trial.

- (1) Determine  $P(X = 4)$ .  
(2) Determine approximately the value of  $P(X = 4)$  by using the central limit theorem with the correction for continuity.

**Solution:** (1) It is clear that  $X \sim B(15, 0.3)$ . So

$$\begin{aligned}
 P(X = 4) &= C_{15}^4 0.3^4 0.7^{11} = \frac{15 \cdot 14 \cdot 13 \cdot 12}{24} 0.3^4 0.7^{11} \\
 &= \frac{15 \cdot 7 \cdot 13 \cdot 3^4 \cdot 7^{11}}{10^{15}} = \frac{15 \cdot 13 \cdot 3^4 \cdot 7^{12}}{10^{15}} = 0.2186
 \end{aligned}$$

- (2)  $E(X) = 15 \cdot 0.3 = 4.5$ ,  $\text{Var}(X) = 15 \cdot 0.3 \cdot 0.7 = 3.15$ . By the central limit theorem,  $X$  approximately has  $N(4.5, 3.15)$ . Using correction for continuity,

$$\begin{aligned}
 P(X = 4) &= P(3.5 < X < 4.5) = P\left(\frac{3.5 - 4.5}{\sqrt{3.15}} < \frac{X - 4.5}{\sqrt{3.15}} < 0\right) \\
 &\approx \Phi(0) - \Phi\left(\frac{-1}{\sqrt{3.15}}\right) = \Phi\left(\frac{1}{\sqrt{3.15}}\right) - 0.5 = \Phi(0.5634) - 0.5 \\
 &= 0.7134 - 0.5 = 0.2134.
 \end{aligned}$$

6. **(12pt)** Suppose that  $X_1, X_2, X_3, X_4, X_5$  are a random sample from distribution  $N(\mu, \sigma^2)$ , where  $\mu$  and  $\sigma^2$  are both unknown. We have the following observed data of the sample:  $(x_1, x_2, x_3, x_4, x_5) = (29, 20, 11, 16, 14)$ . Find a 90% confidence interval for  $\mu$ .

**Solution:** We have  $n = 5$ ,  $\bar{x} = \frac{29+20+11+16+14}{5} = 18$ .  $\alpha = 0.1$

$$\begin{aligned}
 s^2 &= \frac{1}{4} [(29 - 18)^2 + (20 - 18)^2 + (11 - 18)^2 + (16 - 18)^2 + (14 - 18)^2] \\
 &= 48.5.
 \end{aligned}$$

$$s = \sqrt{s^2} = \sqrt{48.5} = 6.9642$$

$$t_{\frac{\alpha}{2}}(n - 1) = t_{0.05}(4) = 2.1318$$

$$\frac{s}{\sqrt{n}} t_{\frac{\alpha}{2}}(n - 1) = \frac{6.9642}{\sqrt{5}} \cdot 2.1318 = 6.6395.$$

Therefore a 90% confidence interval for  $\mu$  is

$$\begin{aligned}
 \left( \bar{x} - \frac{s}{\sqrt{n}} t_{\frac{\alpha}{2}}(n - 1), \bar{x} + \frac{s}{\sqrt{n}} t_{\frac{\alpha}{2}}(n - 1) \right) &= (18 - 6.6395, 18 + 6.6395) = \\
 &= (11.3605, 24.6395).
 \end{aligned}$$

7. (10pt) Let  $X_1, X_2$  be a random sample from  $X$  with  $E(X) = \mu$  and  $\text{Var}(X) = \sigma^2$ . We want  $Y = a_1X_1 + a_2X_2$  to be an estimator of  $\mu$ , where  $a_1, a_2$  are real constants.

(1) For  $Y$  to be unbiased, what requirement must  $a_1, a_2$  fulfil?

(2) Find  $\text{Var}(Y)$ .

(3) Determine  $a_1, a_2$  if we want  $Y$  to be unbiased and the most efficient.

**Solution:** (1)  $E(Y) = a_1E(X_1) + a_2E(X_2) = (a_1 + a_2)\mu$ . For  $E(Y) = \mu$ , it must be satisfied that  $a_1 + a_2 = 1$ .

(2)  $\text{Var}(Y) = a_1^2\text{Var}(X_1) + a_2^2\text{Var}(X_2) = (a_1^2 + a_2^2)\sigma^2$ .

(3) Now suppose that  $Y$  is unbiased and the most efficient. Then  $a_1 + a_2 = 1$ , and

$$\text{Var}(Y) = (a_1^2 + a_2^2)\sigma^2 = (a_1^2 + (1 - a_1)^2)\sigma^2 = (2a_1^2 - 2a_1 + 1)\sigma^2.$$

It is easy to verify that when  $a_1 = \frac{1}{2}$ ,  $\text{Var}(Y)$  attains its minimum, i.e. the

most efficient. Therefore in that situation,  $a_1 = a_2 = \frac{1}{2}$ .

8. (13pt) Let  $X_1, X_2, \dots, X_n$  be a random sample from  $X$  whose PDF is

$$f(x) = \begin{cases} \theta x^{-(\theta+1)}, & x > 1, \\ 0, & \text{otherwise,} \end{cases}$$

where  $\theta > 1$  is an unknown parameter.

(1) Find the moment estimator of  $\theta$ .

(2) Find the maximum likelihood estimator of  $\theta$ .

**Solution:** (1)  $\mu_1 = E(X) = \int_1^\infty x\theta x^{-(\theta+1)}dx = \int_1^\infty \theta x^{-\theta}dx = \frac{\theta}{-\theta+1}x^{-\theta+1}|_1^\infty = \frac{\theta}{\theta-1}$ .

Therefore,  $\theta = \frac{\mu_1}{\mu_1 - 1}$ . And the moment estimator of  $\theta$  is

$$\hat{\theta}_{MOM} = \frac{A_1}{A_1 - 1}, \text{ where } A_1 = \frac{1}{n} \sum_{i=1}^n X_i.$$

(2) Given observed data  $x_1, \dots, x_n$ , the likelihood function

$$L(\theta) = \prod_{i=1}^n \theta x_i^{-(\theta+1)} = \theta^n \left( \prod_{i=1}^n x_i \right)^{-(\theta+1)}$$

$$\ln L(\theta) = n \ln \theta - (\theta + 1) \sum_{i=1}^n \ln(x_i)$$

$$\frac{d \ln L(\theta)}{d\theta} = \frac{n}{\theta} - \sum_{i=1}^n \ln(x_i)$$

Obviously,  $\frac{d \ln L(\theta)}{d\theta} = 0$  if and only if  $\theta = \frac{n}{\sum_{i=1}^n \ln(x_i)}$ .

Therefore, the maximum likelihood estimator of  $\theta$  is

$$\hat{\theta}_{ML} = \frac{n}{\sum_{i=1}^n \ln(X_i)}.$$