

## Centering in Multilevel Regression

Centering is the rescaling of predictors by subtracting the mean. In OLS regression, rescaling using a linear transformation of a predictor (e.g., subtracting one value from every individual score) has no effect on the significance tests and does not alter the standardized slope values (interpretation of the metric of the unstandardized slope may differ with these linear transformations, but in an understandable way). The effects of centering on multilevel regression are quite complex and deserve more consideration than is possible here, but I would like to make a few general points.

### Intercept Interpretation

With multilevel regression, however, intercepts and intercept variances are of interest and linear transformations impact these values as well as their significance tests. One can see in the formula for the intercept at level one (or similarly in OLS regression) that the intercept depends on the value of  $X$  at its mean.

$$\beta_{0j} = \bar{Y}_j - \beta_{1j} \bar{X}_j$$

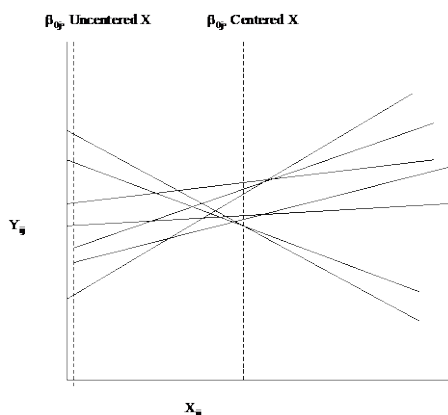
If we recompute the predictor by subtracting the mean from every score ( $x_{ij} = X_{ij} - \bar{X}_j$ ), the value of the intercept will change. The intercept is no longer the expected value of  $Y$  when  $X$  equals the original zero value but, instead, when the value when  $X$  is equal to its original mean. Because the  $\gamma_{00}$  is the average of all of the  $\beta_{0j}$  values, its value and interpretation also is affected by centering.

There are two different versions of centering in multilevel regression, grand mean centering and group mean centering (sometimes called "centering within context"). Grand mean centering subtracts the grand mean of the predictor using the mean from the full sample ( $\bar{X}$ ). Group mean centering subtracts the individual's group mean ( $\bar{X}_j$ ) from the individual's score.

Generally, centering makes this value more interpretable, because the expected value of  $Y$  when  $x$  (centered  $X$ ) is zero represents the expected value of  $Y$  when  $X$  is at its mean. In many cases, such as when age is a predictor, the interpretation of the intercept will be unreasonable or undesirable (e.g., value of the outcome when age equals zero) without some type of centering. It thus appears that raw, uncentered predictors should not be the researcher's default scaling.

### Effects of Centering on Intercept Variance

Not only is the average intercept impacted by centering, but the variance of the intercepts is affected by centering as well. The direction of the effect of centering and the degree of effect depends on the pattern of within-group slopes. The picture below illustrates that moving the location of the  $Y$ -axis by rescaling  $X$ , would lead to a different estimate of the variance of the group intercepts in this case. One can imagine then how the covariance among slopes and intercepts will be affected too.



### Reintroduced means

One special case of a level-2 predictor is a variable that has been computed by averaging the responses for all cases in each group, thereby creating a group-level variable with one value for each group. An example might be calculating the average of the individual student SES values for each classroom and then using these averages as a level-2 predictor in a multilevel regression. When a group-mean centered level-1 predictor and this special type of level-2 variable is used in the model together, it is sometimes referred to as "reintroducing the mean" of the predictor, because the group mean was removed when the level-1 predictor was group-mean centered.

Often the motivation for this type of model is to investigate separate within-group and between-group effects of the predictor. For example, what is the impact of individual-level SES on math achievement as compared with the effect of school-level SES? In this special type of model, when the level-1 and level-2 counterparts are included as predictors, centering approaches can impact the interpretation of the coefficient (Raudenbush & Bryk, 2002, Table 5.11, p. 140, and Snijders & Bosker, 2012, Section 4.6, provide good explanations). If group-mean centering of the level-1 predictor is used, the level-1 predictor coefficient,  $\gamma_{10}$ , will represent the within-group effect and the level-2 predictor,  $\gamma_{01}$ , will represent the between-group effect. In the case of grand-mean centering, however,  $\gamma_{10}$  and  $\gamma_{01}$  are estimates of the within-group effect and the *compositional effect* (difference between the within and between slopes), respectively. The compositional effect coefficient can be derived from the between- and within-group coefficients in the group-mean centered model. The compositional effect may be of interest in some cases, but I suspect that researchers are more often interested in estimating the within-group and between-group effects in order to obtain the independent micro and macro level contributions of a certain predictor.

### Recommendations

The primary decisions about centering have to do with the scaling of level-1 variables. Because there is only one score per group, however, there is only one choice for centering of level-2 variables—grand mean centering. Thus, the decision is simple for level-2 variables. In most cases, researchers would likely choose to grand mean center level-2 variables to improve the interpretation of the intercept values. Of course, one should not blindly follow this recommendation, but there will be far more instances where centering the predictor makes more sense than not centering the predictor. Interpreting the intercept as the value when the level-2 predictor is equal to zero may be desirable in some cases, but I venture to guess not in most cases.

Choices made regarding centering level-1 variables are much more difficult. Generally, in most if not nearly all circumstances, intercept interpretation will be more reasonable using some type of centered predictors as compared with using uncentered predictors.

The consequences of choosing grand- or group-mean centering are almost overwhelming and I cannot make any global recommendations. The best I can do is urge a careful reading of the most thorough considerations of this topic (Algina & Swaminathan, 2011; Enders & Tofighi, 2007), but I provide a summary of some of the general conclusions from these sources in the table below. At this point in the course we are considering group-nested designs only, but we will revisit the issue in the context of longitudinal (growth curve) models. Wang and Maxwell (2015) is a source that discusses centering specifically in the longitudinal case.

Raw uncentered variables	Grand mean centering	Group-mean centering
<ul style="list-style-type: none"> <li>Rarely makes sense unless there is a desire to estimate intercept and intercept variance when the predictor is equal to zero (e.g., mean or variance of group means for females only).</li> <li>Does not make sense when <math>X = 0</math> is an unreasonable or undesirable value (e.g., age = 0)</li> <li>Entails assumption that group means are uncorrelated with the predictor (Algina &amp; Swaminathan, 2011).</li> </ul>	<ul style="list-style-type: none"> <li>Effect of level-2 controlling for level-1 variables.</li> <li>If adjusted group mean is desired interpretation of average intercept, and variance of adjusted group means is desired interpretation of intercept variance.</li> <li>Interactions between level-2 variables. (Interaction estimate and test are unaffected, but lower order terms are).</li> <li>Entails assumption that group means are uncorrelated with predictor.</li> <li>Inclusion of level-2 variables in a model without level-1 reintroduced mean variables will not fully control for the covariate.</li> </ul>	<ul style="list-style-type: none"> <li>Use if pure level-1 effect is desired without considering level-2 variables. Group-mean centered variables will be uncorrelated with level-2 variables and therefore estimates of the effect of the level-1 variable will not partial out level-2 variables.</li> <li>When accurate slope variance estimates are desired</li> <li>When relative between and within effects of the same construct at level 1 and level 2 is desired (e.g., SES and average class SES).</li> <li>If effects of level-2 variables only of interest without regard to partialling level-1 variables out.</li> <li>When cross-level interactions are of interest and interpretation of "main effect" is of interest.</li> <li>Inclusion of level-2 variables in a model without level-1 reintroduced mean variables will not fully control for the covariate.</li> </ul>

**Cross-level interactions.** Cross-level interaction coefficients are relatively unaffected by centering decisions, although Algina and Swaminathan point out that this assumes no other covariates have confounding interactions with the independent variables. Lower order terms ("main effects" of the predictors involved in the interaction) are affected by centering, however. Bauer and Curran (2005) recommend group-mean centering level-1 predictors (and grand-mean centering level-2) to improve computation and interpretation of the "main effects" when cross-level interactions are tested.

**Centering dichotomous predictors?** It may seem odd to center a dichotomous predictor like gender, but if original coding of 0,1 is used, then the intercept and variance of the intercept represents the mean of the 0 group and the variance of the zero group. There is nothing incorrect about this, but it may not be desirable to simply estimate the variance of the intercepts for the 0 group in many cases. It makes sense then to consider centering a binary variable, so that the mean represents the average of the two groups. Note that coding a binary predictor as 1,2 would rarely, if ever, make sense because the intercepts are interpreted as the 0 values of the predictor, which would be a group that does not exist. Deciding whether to group-mean or grand-mean center a binary level-1 predictor is complicated, however. Group mean centering will produce intercepts weighted by the proportion of 1 to 0 values for each group, whereas grand-mean centering will produce intercepts weighted by the proportion of 1 to 0 in the entire sample. The grand-mean centering is analogous to using a sample weight adjustment to make the sample mean (here, each group's mean) be proportionate to the population mean (here, the full sample).

**General comments.** Most of the above conclusions are based on fairly simple models and the structure of the model, such as whether both level-1 and level-2 predictors are included and whether there are cross-level interactions, can make a difference on the consequences of centering choices. There are a number of other complexities that have not been thoroughly considered in the literature, such as the consequences of mixing different centering approaches and the impact of large variability of group sample sizes.

#### References

- Algina, J., & Swaminathan, H. (2011). Centering in two-level nested designs. In J. Hox, & K. Roberts (Eds.), *The Handbook of Advanced Multilevel Data Analysis* (pp 285-312). New York: Routledge.
- Bauer, D.J., & Curran, P.J. (2005). Probing interactions in fixed and multilevel regression: Inferential and graphical techniques. *Multivariate Behavioral Research*, 40, 373-400.
- Enders, C.K., & Tofighi, D. (2007). Centering predictor variables in cross-sectional multilevel models: A new look at an old issue. *Psychological Methods*, 12, 121-138.
- Wang, L. P., & Maxwell, S. E. (2015). On disaggregating between-person and within-person effects with longitudinal data using multilevel models. *Psychological methods*, 20(1), 63.

## Centering Example Using HSB Data Without Reintroduced Means

To provide some illustration of the impact of centering, I have tested two different models using the HSB data set. One set of models (p. 4) includes a Level-1 predictor, TSES, a transformed version of the SES variable, and a Level-2 variable, SECTOR, the variable for type of school (0=public, 1=catholic). The transformed variable TSES was used because the original SES variable was standardized with a mean of 0, which interferes with the ability to compare the effects of centering choices. Although results presented here are from the HLM package, the consequences of centering will not be different using SPSS, R, or other packages. The second set of models (p. 5) examines the effects of reintroducing the school mean of the socioeconomic variable, MEANTSES, into the model at Level-2. Centering effects are complex and the pattern with other models may differ.

### Separate Equations (centering options for TSES are varied)

$$MATHACH = \beta_{0j} + \beta_{1j}(TSES) + R_{ij}$$

$$\beta_{0j} = \gamma_{00} + \gamma_{01}(SECTOR) + U_{0j}$$

$$\beta_{1j} = \gamma_{10} + \gamma_{11}(SECTOR) + U_{1j}$$

### Single Equation

$$MATHACH = \gamma_{00} + \gamma_{01}(SECTOR) + \gamma_{10}(TSES) + \gamma_{11}(TSES * SECTOR) + U_{0j} + U_{1j}(TSES) + R_{ij}$$

	Description	Uncentered	Grand Mean Centered	Group Mean Centered
$\gamma_{00}$	Adjusted grand mean of MATHACH	- 3.044 (.701) , $t = -4.340, p < .001$	11.751 (.292), $t = 50.596, p < .001$	11.394 (.293), $t = 38.915, p < .001$
$\gamma_{01}$	Effect of SECTOR	8.694 (1.089), $t = 7.982, p < .001$	2.128 (.347), $t = 6.140, p < .001$	2.807 (.439), $t = 6.392, p < .001$
$\gamma_{10}$	Average effect of TSES	.296 (.015), $t = 20.341, p < .001$	.296 (.015), $t = 20.341, p < .001$	.280 (.016), $t = 17.904, p < .001$
$\gamma_{11}$	Interaction of SECTOR with TSES	-.131 (.022), $t = -5.994, p < .001$	-.131 (.022), $t = -5.994, p < .001$	-.134 (.024), $t = -5.680, p < .001$
$\tau_0^2$	Variance of adjusted intercepts across schools	1.668, $\chi^2 = 164.874, p = .338$	3.833, $\chi^2 = 756.043, p < .001$	6.740, $\chi^2 = 1383.785, p < .001$
$\tau_1^2$	Variance of TSES slopes across schools	.001, $\chi^2 = 178.091, p = .131$	.001, $\chi^2 = 178.091, p = .131$	.003, $\chi^2 = 175.312, p = .164$
$\sigma^2$	Variance within schools	36.763	36.763	36.690

Note: TSES has a mean of 50 and a standard deviation of 7.79; it is a transformed version of the SES variable found in the original Raudenbush and Bryk HSB (2002) data set using the T-score formula so that the mean would not equal zero for uncentered scores. Level-2 predictors are entered as uncentered variables for all models (not usually recommended). Standard REML estimates (not using robust standard errors) are presented here.

VARIABLE NAME	N	MEAN	SD	MINIMUM	MAXIMUM
MEANTSES	160	49.94	4.14	38.06	58.25
TSES	7185	50.00	7.79	12.42	76.92

### Centering Example Using HSB Data With Reintroduced Means

**Separate Equations** (centering options for TSES are varied)

$$MATHACH = \beta_{0j} + \beta_{1j}(TSES) + R_{ij}$$

$$\beta_{0j} = \gamma_{00} + \gamma_{01}(MEANTSES) + \gamma_{02}(SECTOR) + U_{0j}$$

$$\beta_{1j} = \gamma_{10} + \gamma_{11}(MEANTSES) + \gamma_{12}(SECTOR) + U_{1j}$$

### Single Equation

$$MATHACH = \gamma_{00} + \gamma_{01}(MEANTSES) + \gamma_{02}(SECTOR) + \gamma_{10}(TSES) + \gamma_{11}(TSES * MEANTSES) + \gamma_{12}(TSES * SECTOR) + U_{0j} + U_{1j}(TSES) + R_{ij}$$

	Description	Uncentered	Grand Mean Centered	Group Mean Centered
$\gamma_{00}$	Adjusted grand mean of MATHACH	1.983 (6.562), $t = .302$ , $p = .763$	-4.518 (1.911), $t = -2.364$ , $p < .05$	-14.537 (1.805), $t = -8.053$ , $p < .001$
$\gamma_{01}$	Effect of MEANTSES	-.088 (.136), $t = -.648$ , $p = .518$	.333 (.039), $t = 8.546$ , $p < .001$	.533 (.037), $t = 14.446$ , $p < .001$
$\gamma_{02}$	Effect of SECTOR	9.074 (1.150), $t = 7.888$ , $p < .001$	1.193 (.308), $t = 3.870$ , $p < .001$	1.227 (.306), $t = 4.005$ , $p < .001$
$\gamma_{10}$	Average effect of TSES	-.130 (.134), $t = -.969$ , $p = .334$	-.130 (.134), $t = -.969$ , $p = .334$	-.223 (.147), $t = -1.512$ , $p = .132$
$\gamma_{11}$	Interaction of MEANTSES with TSES	.008 (.002), $t = 3.058$ , $p < .01$	.008(.003), $t = 3.058$ , $p < .01$	.010 (.003), $t = 3.420$ , $p < .01$
$\gamma_{12}$	Interaction of SECTOR with TSES	-.158 (.023), $t = -6.929$ , $p < .001$	-.158 (.023), $t = -6.929$ , $p < .001$	-.164 (.024), $t = -6.756$ , $p < .001$
$\tau_0^2$	Variance of adjusted intercepts across schools	1.922, $\chi^2 = 160.950$ , $p = .398$	2.411, $\chi^2 = 573.179$ , $p < .001$	2.380, $\chi^2 = 605.306$ , $p < .001$
$\tau_1^2$	Variance of TSES slopes across schools	.001, $\chi^2 = 162.623$ , $p = .363$	.001, $\chi^2 = 162.623$ , $p = .362$	.001, $\chi^2 = 162.302$ , $p = .369$
$\sigma^2$	Variance within schools	36.740	36.740	36.703

Note: TSES has a mean of 50 and a standard deviation of 7.79; it is a transformed version of the SES variable found in the original Raudenbush and Bryk HSB (2002) data set using the T-score formula so that the mean would not equal zero for uncentered scores. Level-2 predictors entered as uncentered variables for all models (not usually recommended). Standard REML estimates (not using robust standard errors) are presented here.