

FEM 11087 - Applied Microeconometrics

Assignment 2: Panel Data Analysis

Empirical Application

Group 33

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Question 1 [0.7 points]

*A central question in labor economics is: **How much more do individuals earn with higher levels of education?** Economists often estimate the returns to education—that is, the increase in earnings associated with completing high school, college, or additional years of schooling.*

*Using the panel data provided, begin by constructing a **bar chart** showing **mean income by education group**. Group individuals based on their **highest level of educational attainment** (e.g., less than high school, high school graduate, some college, college degree or more), and plot the **average income** for each category.*

```
1 gen edyears_cat = .
2 replace edyears_cat = 1 if edyears <= 11 & !missing(edyears)
3 replace edyears_cat = 2 if edyears == 12 & !missing(edyears)
4 replace edyears_cat = 3 if edyears >= 13 & ///
5     edyears <= 15 & !missing(edyears)
6 replace edyears_cat = 4 if edyears >= 16 & !missing(edyears)
```

```
1 graph bar (mean) income, over(edyears_cat)
2 graph bar (mean) income, over(edyears_cat) by(male)
```

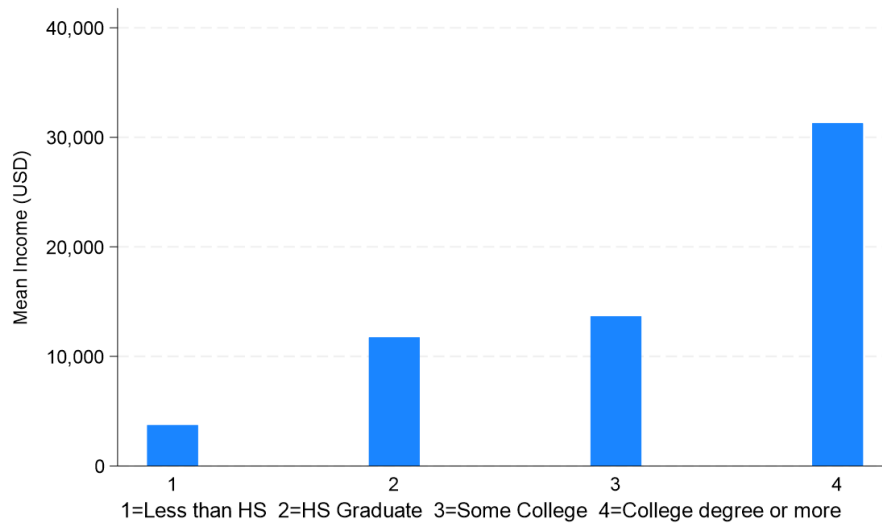


Figure 1.1: Mean income by education level

Recent debates around student debt and the value of higher education often assume that education “pays off” equally for everyone. **Does your analysis support that assumption?** To explore this, create **separate plots by gender** to highlight any differences in the relationship between education and earnings. Discuss your findings.

Note: For this question, create and use a categorical education variable based on each individual’s highest level of education completed across the panel. Construct four categories:

- Less than high school (11 or fewer years)
- High school graduate (exactly 12 years)
- Some college (13 to 15 years)
- College degree or more (16 or more years)

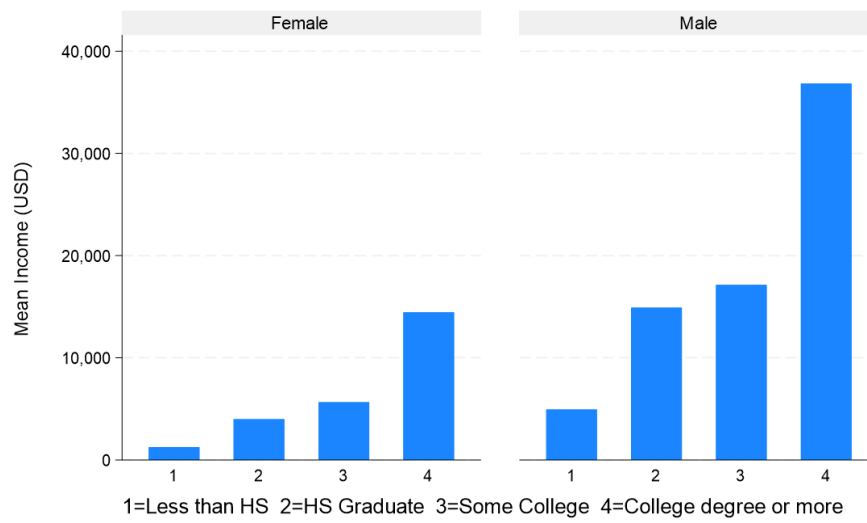


Figure 1.2: Mean income by education level, by gender

Question 2 [1 point]

Now, we turn to formally estimating the effect of years of education on income. Use **pooled OLS** to examine the impact of years of education (*edyears*) on **log(income)**, controlling for age, gender (*male*), marital status categories, ethnicity categories, and childbirth.

```

1 sum income, detail
2 count if income <= 0

1 gen log_income = log(income)
2 reg log_income edyears age i.male ib0.mstatus ib4.ethnicity ///
3   i.child_birth

```

Table 2.1: Pooled OLS model

	Coefficient	Robust std. err.	P> t	Conf. int.
Treatment variable				
Education years	0.1458	(0.0027)	0.000	[0.140,0.151]
Control variables				
Age	0.1263	(0.0014)	0.000	[0.123,0.129]
Male	0.6000	(0.0126)	0.000	[0.575,0.625]
Childbirth	-0.0795	(0.0184)	0.000	[-0.116,-0.043]
Married	0.6189	(0.0185)	0.000	[0.583,0.655]
Separated or divorced	0.0767	(0.0375)	0.041	[0.003,0.150]
Widowed	0.0647	(0.2531)	0.798	[-0.431,0.561]
Black	-0.4286	(0.0139)	0.000	[-0.456,-0.401]
Hispanic	-0.0825	(0.0146)	0.000	[-0.111,-0.054]
Mixed Race (Non-Hispanic)	-0.3744	(0.0583)	0.000	[-0.489,-0.260]
Constant	3.0156	(0.0319)	0.000	[2.953,3.078]
Number of obs	55874			
F-statistic	3864.88			
Prob > F	0.0000			
R-squared	0.4089			
Adj. R-squared	0.4088			
Dependent variable: log(income)				

a) What is the estimated return to an additional year of education? Interpret the coefficient on years of education in terms of its **sign**, **magnitude**, and **statistical significance**.

b) Differences in returns to schooling by gender are sometimes interpreted as potential evidence of **labor market discrimination**. Test whether the effect of years of education using the categorical variable created in Question 1 on $\log(\text{income})$ is the **same for men and women**. Based on your results, do you find any evidence consistent with discrimination?

```
1 reg log_income i.edyears_cat##i.male age ib0.mstatus ib4.ethnicity ///
2 i.child_birth
```

Table 2.2: Pooled OLS model with interaction effects

	Coefficient	Robust std. err.	P> t	Conf. int.
Treatment variable				
HS graduate	0.3802	(0.0256)	0.000	[0.330,0.430]
Some college	0.6488	(0.0277)	0.000	[0.594,0.703]
College degree or more	1.0239	(0.0388)	0.000	[0.948,1.100]
HS graduate × Male	0.3836	(0.0302)	0.000	[0.324,0.443]
Some college × Male	0.1664	(0.0326)	0.000	[0.102,0.230]
College degree or more × Male	0.2206	(0.0432)	0.000	[0.136,0.305]
Control variables				
Age	0.1225	(0.0014)	0.000	[0.120,0.125]
Male	0.4472	(0.0188)	0.000	[0.410,0.484]
Childbirth	-0.0759	(0.0182)	0.000	[-0.112,-0.040]
Married	0.6013	(0.0183)	0.000	[0.565,0.637]
Separated or divorced	0.0631	(0.0372)	0.090	[-0.010,0.136]
Widowed	0.0932	(0.2509)	0.710	[-0.399,0.585]
Black	-0.4143	(0.0138)	0.000	[-0.441,-0.387]
Hispanic	-0.0732	(0.0145)	0.000	[-0.102,-0.045]
Mixed Race (Non-Hispanic)	-0.3752	(0.0578)	0.000	[-0.488,-0.262]
Constant	4.4713	(0.0289)	0.000	[4.415,4.528]
Number of obs	55874			
F-statistic	2690.54			
Prob > F	0.0000			
R-squared	0.4195			
Adj. R-squared	0.4193			
Dependent variable: $\log(\text{income})$				

```
1 test 2.edyears_cat#1.male 3.edyears_cat#1.male 4.edyears_cat#1.male
```

Table 2.3: Joint F-test

F-statistic	54.85
Prob > F	0.0000
<hr/>	
(1) HS graduate \times Male = 0	
(2) Some college \times Male = 0	
(3) College degree or more \times Male = 0	

c) Under what conditions is the pooled OLS estimate of the effect of years of education **unbiased and efficient**? Do you believe these conditions are likely to hold in this context?

Question 3 [0.5 points]

So far, the panel structure of the data has been largely unexploited. Random effects (RE) estimation can improve the efficiency of the estimates compared to pooled OLS.

*a) Estimate the effect of years of education (edyears) on **log(income)** using the **random effects** (RE) model, controlling for age, gender (male), marital status categories, ethnicity categories, and childbirth. Interpret the estimated coefficient for years of education in terms of its **sign**, **magnitude**, and **statistical significance**. Then, compare the RE estimate and standard error of the education coefficient with those obtained from the **pooled OLS model**.*

```
1  xtset pid wave
1  xtreg log_income edyears age i.male ib0.mstatus ib4.ethnicity ///
2      i.child_birth, re
3  estimates store random
```

Table 3.1: Random effects (RE) model

	Coefficient	Std. err.	P> t	Conf. int.
Treatment variable				
Education years	0.1509	(0.0031)	0.000	[0.145,0.157]
Control variables				
Age	0.1273	(0.0015)	0.000	[0.124,0.130]
Male	0.5698	(0.0167)	0.000	[0.537,0.603]
Childbirth	-0.0599	(0.0174)	0.001	[-0.094,-0.026]
Married	0.5107	(0.0197)	0.000	[0.472,0.549]
Separated or divorced	0.0454	(0.0392)	0.246	[-0.031,0.122]
Widowed	0.0855	(0.2619)	0.744	[-0.428,0.599]
Black	-0.3833	(0.0192)	0.000	[-0.421,-0.346]
Hispanic	-0.0655	(0.0203)	0.001	[-0.105,-0.026]
Mixed Race (Non-Hispanic)	-0.3244	(0.0813)	0.000	[-0.484,-0.165]
Constant	2.9423	(0.0341)	0.000	[2.875,3.009]
Number of obs	55874			
Number of groups	7126			
R-squared within	0.3380			
R-squared between	0.5546			
R-squared overall	0.4084			
Wald χ^2	34633.21			
Prob > χ^2	0.0000			
σ_u	0.4412			
σ_e	1.1766			
ρ	0.1233			
Dependent variable: log(income)				

Table 3.2: POLS, RE comparison

	POLS	RE
Treatment variable		
Education years	0.1458*** (0.0027)	0.1509*** (0.0031)
Control variables		
Age	0.1263*** (0.0014)	0.1273*** (0.0015)
Male	0.6000*** (0.0126)	0.5698*** (0.0167)
Childbirth	-0.0795*** (0.0184)	-0.0599*** (0.0174)
Black	-0.4286*** (0.0139)	-0.3833*** (0.0192)
Hispanic	-0.0825*** (0.0146)	-0.0655** (0.0203)
Mixed Race (Non-Hispanic)	-0.3744*** (0.0583)	-0.3244*** (0.0813)
Married	0.6189*** (0.0185)	0.5107*** (0.0197)
Separated or divorced	0.0767* (0.0375)	0.0454 (0.0392)
Widowed	0.0647 (0.2531)	0.0855 (0.2619)
Constant	3.0156*** (0.0319)	2.9423*** (0.0341)
Number of obs	55874	55874
Number of groups		7126
F-statistic	3864.88	
Wald χ^2		34633.21
P-value	0.0000	0.0000
R-squared	0.4089	
Adj. R-squared	0.4088	
R-squared within		0.3380
R-squared between		0.5546
R-squared overall		0.4084
σ_u		0.4412
σ_e		1.1766
ρ		0.1233

Standard errors in parentheses.

* p < 0.05, ** p < 0.01, *** p < 0.001

b) Under which conditions and why can the random effects estimator be **more efficient** than pooled OLS?

Question 4 [1.55 points]

Alternatively, the panel structure of the data can be used to perform **fixed effects (FE)** estimation.

a) Based on theoretical considerations, would you **prefer** fixed effects or random effects estimation? Justify your answer.

b) Use a **fixed effects estimator** to examine the impact of years of education (edyears) on **log(income)**, controlling for age, gender (male), marital status categories, ethnicity categories, and childbirth. Interpret the coefficient on years of education in terms of its **sign, magnitude, and statistical significance**. Compare your results with those from the **pooled OLS** and **random effects** models.

```
1 xtreg log_income edyears age i.male ib0.mstatus ib4. ethnicity ///
2     i.child_birth, fe
3 estimates store fixed
```

Table 4.1: Fixed effects (FE) model

	Coefficient	Std. err.	P> t	Conf. int.
Treatment variable				
Education years	0.1610	(0.0038)	0.000	[0.153,0.168]
Control variables				
Age	0.1259	(0.0017)	0.000	[0.123,0.129]
Male	0.0000	(.)	.	[0.000,0.000]
Childbirth	-0.0545	(0.0180)	0.002	[-0.090,-0.019]
Married	0.4178	(0.0223)	0.000	[0.374,0.461]
Separated or divorced	-0.0112	(0.0432)	0.796	[-0.096,0.073]
Widowed	0.1575	(0.2861)	0.582	[-0.403,0.718]
Black	0.0000	(.)	.	[0.000,0.000]
Hispanic	0.0000	(.)	.	[0.000,0.000]
Mixed Race (Non-Hispanic)	0.0000	(.)	.	[0.000,0.000]
Constant	3.1536	(0.0378)	0.000	[3.080,3.228]
Number of obs	55874			
Number of groups	7126			
R-squared within	0.3383			
R-squared between	0.4945			
R-squared overall	0.3713			
Wald χ^2				
Prob > χ^2	0.0000			
σ_u	0.8174			
σ_e	1.1766			
ρ	0.3255			
Dependent variable: log(income)				

Table 4.2: POLS, RE, FE comparison

	POLS	RE	FE
Treatment variable			
Education years	0.1458*** (0.0027)	0.1509*** (0.0031)	0.1610*** (0.0038)
Control variables			
Age	0.1263*** (0.0014)	0.1273*** (0.0015)	0.1259*** (0.0017)
Male	0.6000*** (0.0126)	0.5698*** (0.0167)	0.0000 (.)
Childbirth	-0.0795*** (0.0184)	-0.0599*** (0.0174)	-0.0545** (0.0180)
Black	-0.4286*** (0.0139)	-0.3833*** (0.0192)	0.0000 (.)
Hispanic	-0.0825*** (0.0146)	-0.0655** (0.0203)	0.0000 (.)
Mixed Race (Non-Hispanic)	-0.3744*** (0.0583)	-0.3244*** (0.0813)	0.0000 (.)
Married	0.6189*** (0.0185)	0.5107*** (0.0197)	0.4178*** (0.0223)
Separated or divorced	0.0767* (0.0375)	0.0454 (0.0392)	-0.0112 (0.0432)
Widowed	0.0647 (0.2531)	0.0855 (0.2619)	0.1575 (0.2861)
Constant	3.0156*** (0.0319)	2.9423*** (0.0341)	3.1536*** (0.0378)
Number of obs	55874	55874	55874
Number of groups		7126	7126
F-statistic	3864.88		4154.09
Wald χ^2		34633.21	
P-value	0.0000	0.0000	0.0000
R-squared	0.4089		0.3383
Adj. R-squared	0.4088		0.2415
R-squared within		0.3380	0.3383
R-squared between		0.5546	0.4945
R-squared overall		0.4084	0.3713
σ_u		0.4412	0.8174
σ_e		1.1766	1.1766
ρ		0.1233	0.3255

Standard errors in parentheses.

* $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$

c) Perform the **Hausman** test. What do the results indicate? Based on the test outcome, **which estimator** (RE or FE) is more appropriate in this context?

1 `hausman` fixed random

Table 4.3: Hausman test

χ^2	141.94
Prob > χ^2	0.0000
H0: Difference in β not systematic	

Question 5 [0.9 points]

Next, estimate a **Correlated Random Effects (CRE)** model to examine the effect of years of education (*edyears*) on ***log(income)***.

```
1 by pid: egen age_mean = mean(age)
2 by pid: egen mstatus_mean = mean(mstatus)
```

```
1 xtreg log_income edyears age i.male ib0.mstatus ib4.ethnicity ///
2 i.child_birth age_mean mstatus_mean, re
```

Table 5.1: Correlated Random Effects (CRE) model

	Coefficient	Std. err.	P> t	Conf. int.
Treatment variable				
Education years	0.1505	(0.0031)	0.000	[0.144,0.156]
Control variables				
Age	0.1282	(0.0016)	0.000	[0.125,0.131]
Male	0.5600	(0.0178)	0.000	[0.525,0.595]
Childbirth	-0.0617	(0.0174)	0.000	[-0.096,-0.028]
Married	0.4463	(0.0221)	0.000	[0.403,0.490]
Separated or divorced	-0.0669	(0.0430)	0.119	[-0.151,0.017]
Widowed	-0.0927	(0.2632)	0.725	[-0.609,0.423]
Black	-0.3703	(0.0194)	0.000	[-0.408,-0.332]
Hispanic	-0.0684	(0.0203)	0.001	[-0.108,-0.029]
Mixed Race (Non-Hispanic)	-0.3117	(0.0813)	0.000	[-0.471,-0.152]
CRE variables				
age_mean	0.0013	(0.0032)	0.676	[-0.005,0.008]
mstatus_mean	0.2142	(0.0340)	0.000	[0.148,0.281]
Constant	2.8795	(0.0581)	0.000	[2.766,2.993]
Number of obs	55874			
Number of groups	7126			
R-squared within	0.3382			
R-squared between	0.5565			
R-squared overall	0.4093			
Wald χ^2	34714.61			
Prob > χ^2	0.0000			
σ_u	0.4412			
σ_e	1.1766			
ρ	0.1233			
Dependent variable: log(income)				

a) What is one advantage of the **CRE estimator** compared to the **random effects (RE) estimator**?

b) What is one advantage of the **CRE estimator** compared to the **fixed effects (FE) estimator**?

c) Compare the estimated coefficient for years of education from the **CRE model** with those from the **RE** and **FE** models. Are the coefficients similar or different? Explain why this is the case.

Table 5.2: POLS, RE, FE, CRE comparison

	POLS	RE	FE	CRE
Treatment variable				
Education years	0.1458*** (0.0027)	0.1509*** (0.0031)	0.1610*** (0.0038)	0.1505*** (0.0031)
Control variables				
Age	0.1263*** (0.0014)	0.1273*** (0.0015)	0.1259*** (0.0017)	0.1282*** (0.0016)
Male	0.6000*** (0.0126)	0.5698*** (0.0167)	0.0000 (.)	0.5600*** (0.0178)
Childbirth	-0.0795*** (0.0184)	-0.0599*** (0.0174)	-0.0545** (0.0180)	-0.0617*** (0.0174)
Black	-0.4286*** (0.0139)	-0.3833*** (0.0192)	0.0000 (.)	-0.3703*** (0.0194)
Hispanic	-0.0825*** (0.0146)	-0.0655** (0.0203)	0.0000 (.)	-0.0684*** (0.0203)
Mixed Race (Non-Hispanic)	-0.3744*** (0.0583)	-0.3244*** (0.0813)	0.0000 (.)	-0.3117*** (0.0813)
Married	0.6189*** (0.0185)	0.5107*** (0.0197)	0.4178*** (0.0223)	0.4463*** (0.0221)
Separated or divorced	0.0767* (0.0375)	0.0454 (0.0392)	-0.0112 (0.0432)	-0.0669 (0.0430)
Widowed	0.0647 (0.2531)	0.0855 (0.2619)	0.1575 (0.2861)	-0.0927 (0.2632)
CRE variables				
age_mean				0.0013 (0.0032)
mstatus_mean				0.2142*** (0.0340)
Constant	3.0156*** (0.0319)	2.9423*** (0.0341)	3.1536*** (0.0378)	2.8795*** (0.0581)
Number of obs	55874	55874	55874	55874
Number of groups		7126	7126	7126
F-statistic	3864.88		4154.09	
Wald χ^2		34633.21		34714.61
P-value	0.0000	0.0000	0.0000	0.0000
R-squared	0.4089		0.3383	
Adj. R-squared	0.4088		0.2415	
R-squared within		0.3380	0.3383	0.3382
R-squared between		0.5546	0.4945	0.5565
R-squared overall		0.4084	0.3713	0.4093
σ_u		0.4412	0.8174	0.4412
σ_e		1.1766	1.1766	1.1766
ρ		0.1233	0.3255	0.1233

Standard errors in parentheses.

* p < 0.05, ** p < 0.01, *** p < 0.001

d) Based on your CRE estimates, does the assumption of **exogeneity** appear to hold? Which estimator would you consider most appropriate in this context?

Question 6 [0.9 points]

Recent research provides compelling evidence that after the birth of a first child, women's earnings decline sharply and remain persistently lower, while men's earnings remain largely unaffected.

*a) Estimate the effect of childbirth on **log(income)** using the **most appropriate model**. Control for age, gender (male), marital status categories, ethnicity categories, and years of education (edyears). Interpret the estimated coefficient for childbirth in terms of its **sign, magnitude, and statistical significance**.*

```
1 xtreg log_income i.child_birth age i.male ib0.mstatus ib4.ethnicity ///  
2 edyears age_mean mstatus_mean, re
```

Table 6.1: Correlated Random Effects (CRE)

	Coefficient	Std. err.	P> t	Conf. int.
Treatment variable				
Childbirth	-0.0617	(0.0174)	0.000	[-0.096,-0.028]
Control variables				
Age	0.1282	(0.0016)	0.000	[0.125,0.131]
Male	0.5600	(0.0178)	0.000	[0.525,0.595]
Education years	0.1505	(0.0031)	0.000	[0.144,0.156]
Married	0.4463	(0.0221)	0.000	[0.403,0.490]
Separated or divorced	-0.0669	(0.0430)	0.119	[-0.151,0.017]
Widowed	-0.0927	(0.2632)	0.725	[-0.609,0.423]
Black	-0.3703	(0.0194)	0.000	[-0.408,-0.332]
Hispanic	-0.0684	(0.0203)	0.001	[-0.108,-0.029]
Mixed Race (Non-Hispanic)	-0.3117	(0.0813)	0.000	[-0.471,-0.152]
CRE variables				
age_mean	0.0013	(0.0032)	0.676	[-0.005,0.008]
mstatus_mean	0.2142	(0.0340)	0.000	[0.148,0.281]
Constant	2.8795	(0.0581)	0.000	[2.766,2.993]
Number of obs	55874			
Number of groups	7126			
R-squared within	0.3382			
R-squared between	0.5565			
R-squared overall	0.4093			
Wald χ^2	34714.61			
Prob > χ^2	0.0000			
σ_u	0.4412			
σ_e	1.1766			
ρ	0.1233			
Dependent variable: log(income)				

b) Test whether the effect of childbirth on log(income) *differs* between males and females. What conclusions can you draw from your results?

```
1 xtreg log_income i.child_birth##i.male age ib0.mstatus ib4.ethnicity ///
2 edyears age_mean mstatus_mean, re
```

Table 6.2: Correlated Random Effects (CRE) with interaction effects

	Coefficient	Std. err.	P> t	Conf. int.
Treatment variable				
Childbirth	-0.3390	(0.0312)	0.000	[-0.400,-0.278]
Childbirth × Male	0.3959	(0.0370)	0.000	[0.323,0.468]
Control variables				
Age	0.1272	(0.0016)	0.000	[0.124,0.130]
Male	0.5165	(0.0183)	0.000	[0.481,0.552]
Education years	0.1512	(0.0031)	0.000	[0.145,0.157]
Married	0.4458	(0.0221)	0.000	[0.403,0.489]
Separated or divorced	-0.0692	(0.0429)	0.107	[-0.153,0.015]
Widowed	-0.0878	(0.2629)	0.739	[-0.603,0.428]
Black	-0.3674	(0.0194)	0.000	[-0.405,-0.329]
Hispanic	-0.0680	(0.0203)	0.001	[-0.108,-0.028]
Mixed Race (Non-Hispanic)	-0.3043	(0.0813)	0.000	[-0.464,-0.145]
CRE variables				
age_mean	0.0021	(0.0032)	0.515	[-0.004,0.008]
mstatus_mean	0.2211	(0.0340)	0.000	[0.154,0.288]
Constant	2.9032	(0.0582)	0.000	[2.789,3.017]
Number of obs	55874			
Number of groups	7126			
R-squared within	0.3396			
R-squared between	0.5563			
R-squared overall	0.4106			
Wald χ^2	34887.51			
Prob > χ^2	0.0000			
σ_u	0.4418			
σ_e	1.1753			
ρ	0.1238			
Dependent variable: log(income)				

```
1 test 1.child_birth#1.male
```

Table 6.3: Single coeff. test

χ^2	114.41
Prob > χ^2	0.0000
(1) Childbirth × Male = 0	

Question 7 [1.2 points]

Without conducting any empirical analysis:

- a) Compare the key assumptions underlying **pooled OLS**, **fixed effects (FE)**, and **random effects (RE)** estimators. Discuss theoretically in which scenarios you would prefer to use each method.*
- b) Within the practical context of this assignment (effect of education on earnings), provide an example situation for each estimator in the form of a **Directed Acyclic Graph (DAG)**. For each case (Pooled OLS, FE, and RE), explain why the assumptions required for the respective method hold in that example, and why that method would be preferred.*

Question 8 [0.75 points]

Finally, revisit your data and evaluate whether **attrition** is present in your sample. Based on your preferred model, discuss the likelihood of **attrition bias**. What conclusions can you draw regarding its presence, and how might it affect the validity of your results?

```
1 bysort pid (wave): gen n_waves = _N
2 gen all_waves = n_waves == 17
```

```
1 xtreg log_income i.child_birth##i.male age ib0.mstatus ib4.ethnicity ///
2 edyears age_mean mstatus_mean all_waves, re
```

Table 8.1: Attrition bias: *all waves* indicator

	Coefficient	Std. err.	P> t	Conf. int.
Treatment variable				
Childbirth	-0.3376	(0.0312)	0.000	[-0.399,-0.276]
Childbirth × Male	0.3971	(0.0370)	0.000	[0.325,0.470]
Control variables				
Age	0.1272	(0.0016)	0.000	[0.124,0.130]
Male	0.5146	(0.0183)	0.000	[0.479,0.550]
Education years	0.1511	(0.0031)	0.000	[0.145,0.157]
Married	0.4451	(0.0221)	0.000	[0.402,0.488]
Separated or divorced	-0.0687	(0.0429)	0.110	[-0.153,0.015]
Widowed	-0.0913	(0.2629)	0.728	[-0.607,0.424]
Black	-0.3660	(0.0194)	0.000	[-0.404,-0.328]
Hispanic	-0.0673	(0.0203)	0.001	[-0.107,-0.027]
Mixed Race (Non-Hispanic)	-0.3058	(0.0813)	0.000	[-0.465,-0.147]
CRE variables				
age_mean	-0.0001	(0.0033)	0.988	[-0.007,0.006]
mstatus_mean	0.2220	(0.0340)	0.000	[0.155,0.289]
Bias indicator				
all_waves	0.0586	(0.0252)	0.020	[0.009,0.108]
Constant	2.9418	(0.0605)	0.000	[2.823,3.060]
Number of obs	55874			
Number of groups	7126			
R-squared within	0.3396			
R-squared between	0.5565			
R-squared overall	0.4108			
Wald χ^2	34898.11			
Prob > χ^2	0.0000			
σ_u	0.4416			
σ_e	1.1753			
ρ	0.1237			
Dependent variable: log(income)				

```
1 bysort pid (wave): gen next_wave = (wave[_n+1] == wave + 1)
```

```
1 xtreg log_income i.child_birth##i.male age ib0.mstatus ib4.ethnicity ///
2 edyears age_mean mstatus_mean next_wave, re
```


Table 8.2: Attrition bias: *next wave* indicator

	Coefficient	Std. err.	P> t	Conf. int.
Treatment variable				
Childbirth	-0.3361	(0.0312)	0.000	[-0.397,-0.275]
Childbirth × Male	0.3941	(0.0370)	0.000	[0.322,0.467]
Control variables				
Age	0.1290	(0.0016)	0.000	[0.126,0.132]
Male	0.5165	(0.0183)	0.000	[0.481,0.552]
Education years	0.1509	(0.0031)	0.000	[0.145,0.157]
Married	0.4435	(0.0221)	0.000	[0.400,0.487]
Separated or divorced	-0.0704	(0.0429)	0.101	[-0.155,0.014]
Widowed	-0.0932	(0.2629)	0.723	[-0.608,0.422]
Black	-0.3668	(0.0194)	0.000	[-0.405,-0.329]
Hispanic	-0.0679	(0.0203)	0.001	[-0.108,-0.028]
Mixed Race (Non-Hispanic)	-0.3061	(0.0812)	0.000	[-0.465,-0.147]
CRE variables				
age_mean	-0.0014	(0.0033)	0.672	[-0.008,0.005]
mstatus_mean	0.2226	(0.0340)	0.000	[0.156,0.289]
Bias indicator				
next_wave	0.0641	(0.0139)	0.000	[0.037,0.091]
Constant	2.8931	(0.0582)	0.000	[2.779,3.007]
Number of obs	55874			
Number of groups	7126			
R-squared within	0.3398			
R-squared between	0.5567			
R-squared overall	0.4109			
Wald χ^2	34925.32			
Prob > χ^2	0.0000			
σ_u	0.4414			
σ_e	1.1752			
ρ	0.1236			
Dependent variable: log(income)				

```

1 xtreg log_income i.child_birth##i.male age ib0.mstatus ib4.ethnicity ///
2 edyears age_mean mstatus_mean n_waves, re

```

Table 8.3: Attrition bias: *number of waves* indicator

	Coefficient	Std. err.	P> t	Conf. int.
Treatment variable				
Childbirth	-0.3358	(0.0312)	0.000	[-0.397,-0.275]
Childbirth × Male	0.3979	(0.0370)	0.000	[0.325,0.470]
Control variables				
Age	0.1272	(0.0016)	0.000	[0.124,0.130]
Male	0.5103	(0.0185)	0.000	[0.474,0.546]
Education years	0.1510	(0.0031)	0.000	[0.145,0.157]
Married	0.4451	(0.0221)	0.000	[0.402,0.488]
Separated or divorced	-0.0699	(0.0429)	0.104	[-0.154,0.014]
Widowed	-0.0865	(0.2629)	0.742	[-0.602,0.429]
Black	-0.3666	(0.0194)	0.000	[-0.405,-0.329]
Hispanic	-0.0676	(0.0203)	0.001	[-0.107,-0.028]
Mixed Race (Non-Hispanic)	-0.3058	(0.0812)	0.000	[-0.465,-0.147]
CRE variables				
age_mean	-0.0058	(0.0045)	0.197	[-0.015,0.003]
mstatus_mean	0.2210	(0.0340)	0.000	[0.154,0.288]
Bias indicator				
n_waves	0.0065	(0.0027)	0.014	[0.001,0.012]
Constant	3.0107	(0.0728)	0.000	[2.868,3.153]
Number of obs	55874			
Number of groups	7126			
R-squared within	0.3396			
R-squared between	0.5567			
R-squared overall	0.4108			
Wald χ^2	34902.86			
Prob > χ^2	0.0000			
σ_u	0.4412			
σ_e	1.1753			
ρ	0.1235			
Dependent variable: log(income)				