
DIGITAL COMMUNICATION

Through Simulations

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Introduction

This book introduces digital communication through probability.

Chapter 1

Maximum Likelihood Detection: BPSK

1.1. Maximum Likelihood

1.1.1 Generate equiprobable $X \in \{1, -1\}$.

1.1.2 Generate

$$Y = AX + N, \tag{1.1}$$

where $A = 5$ dB, and $N \sim \mathcal{N}(0, 1)$.

1.1.3 Plot Y using a scatter plot.

1.1.4 Guess how to estimate X from Y .

1.1.5 Find

$$P_{e|0} = \Pr(\hat{X} = -1 | X = 1) \tag{1.2}$$

and

$$P_{e|1} = \Pr(\hat{X} = 1 | X = -1) \tag{1.3}$$

1.1.6 Find P_e assuming that X has equiprobable symbols.

1.1.7 Verify by plotting the theoretical P_e with respect to A from 0 to 10 dB.

1.1.8 Now, consider a threshold δ while estimating X from Y . Find the value of δ that maximizes the theoretical P_e .

1.1.9 Repeat the above exercise when

$$p_X(0) = p \quad (1.4)$$

1.1.10 Repeat the above exercise using the MAP criterion.

1.2. Examples

1.2.1 Consider a channel over which either symbol x_A or symbol x_B is transmitted. Let the output of the channel Y be the input to a maximum likelihood (ML) detector at the receiver. The conditional probability density for Y given x_A and x_B are:

$$\begin{aligned} f_{Y|x_A}(y) &= e^{-(y+1)} u(y+1) \\ f_{Y|x_B}(y) &= e^{(y-1)} (1 - u(y-1)) \end{aligned}$$

where $u(\cdot)$ is the standard unit step function. the probability of symbol error for this system is

(GATE EC 2022) **Solution:**

Decision in favour of x_A when

$$f_{Y|x_A}(y) > f_{Y|x_B}(y) \quad (1.5)$$

Decision in favour of x_B when

$$f_{Y|x_A}(y) < f_{Y|x_B}(y) \quad (1.6)$$

From the figure,

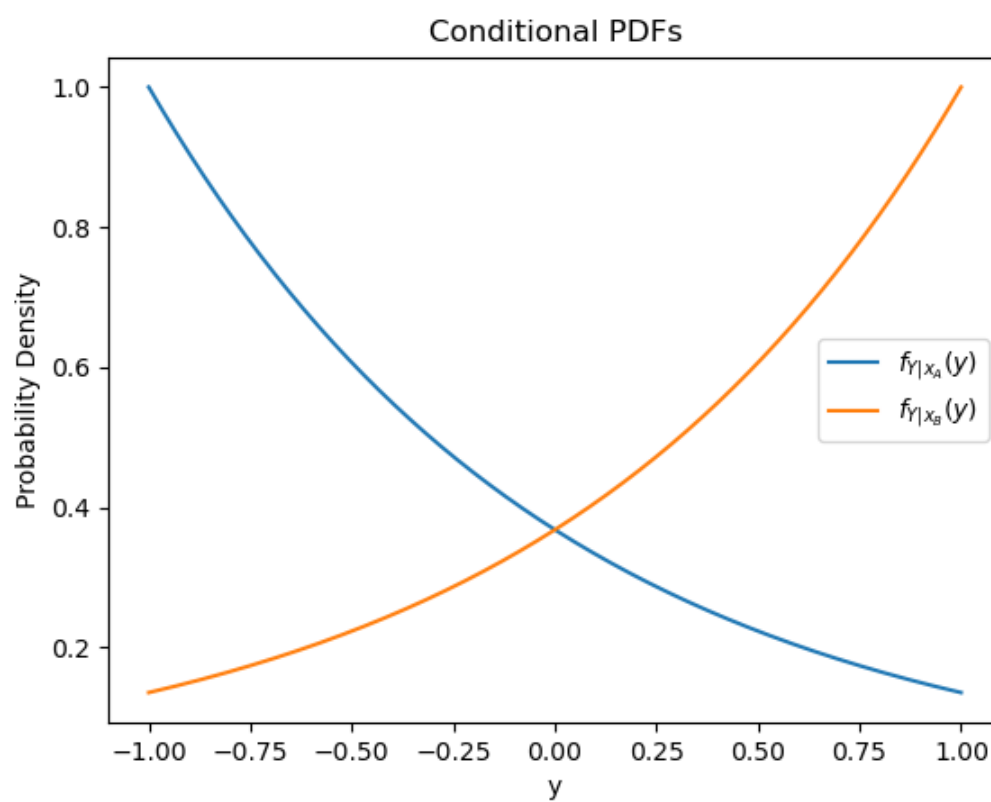


Figure 1.1: Conditional pdf

$$\left\{ \begin{array}{ll} f_{Y|x_A}(y) < f_{Y|x_B}(y) & , y < -1 \\ f_{Y|x_A}(y) > f_{Y|x_B}(y) & , -1 < y < 0 \\ f_{Y|x_A}(y) < f_{Y|x_B}(y) & , 0 < y < 1 \\ f_{Y|x_A}(y) > f_{Y|x_B}(y) & , y > 1 \end{array} \right. \quad (1.7)$$

(a) When $0 < y < 1$.

In this interval, when x_A is transmitted, error occurs because the likelihood of observing Y given x_A is lower than the likelihood of observing Y given x_B ,
Therefore,

$$P_{e_{x_A}} = \int_0^1 f_{Y|x_A}(y) dy \quad (1.8)$$

$$= \int_0^1 e^{-(y+1)} u(y+1) dy \quad (1.9)$$

$$= \int_0^1 e^{-(y+1)} dy \quad (1.10)$$

$$= e^{-1} - e^{-2} \quad (1.11)$$

$$P_{e_{x_A}} = 0.23 \quad (1.12)$$

(b) When $-1 < y < 0$.

In this interval, when x_B is transmitted, error occurs because the likelihood of observing Y given x_A is higher than the likelihood of observing Y given x_B ,

Therefore,

$$P_{e_{x_B}} = \int_{-1}^0 f_{Y|x_B}(y) dy \quad (1.13)$$

$$= \int_{-1}^0 e^{(y-1)} (1 - u(y-1)) \quad (1.14)$$

$$= \int_{-1}^0 e^{(y-1)} \quad (1.15)$$

$$= e^{-1} - e^{-2} \quad (1.16)$$

$$P_{e_{x_B}} = 0.23 \quad (1.17)$$

(c) When $y < -1$ and $y > 1$.

There are no errors in these intervals as the ML detectors can more reliably to make a decision. Therefore,

$$P_e = 0 \quad (1.18)$$

Hence, the total probability of error for this system can be given as,

$$P_e = \Pr(x_A) P_{e_{x_A}} + \Pr(x_B) P_{e_{x_B}} \quad (1.19)$$

$$= 0.23 \times (\Pr(x_A) + \Pr(x_B)) \quad (1.20)$$

$$= 0.23 \quad (1.21)$$

Chapter 2

Transformation of Random Variables

2.1. Gaussian to Other

2.1.1 Let $X_1 \sim \mathcal{N}(0, 1)$ and $X_2 \sim \mathcal{N}(0, 1)$. Plot the CDF and PDF of

$$V = X_1^2 + X_2^2 \tag{2.1}$$

2.1.2 If

$$F_V(x) = \begin{cases} 1 - e^{-\alpha x} & x \geq 0 \\ 0 & x < 0, \end{cases} \tag{2.2}$$

find α .

2.1.3 Plot the CDF and PDF of

$$A = \sqrt{V} \tag{2.3}$$

2.2. Conditional Probability

2.2.1 Plot

$$P_e = \Pr(\hat{X} = -1 | X = 1) \quad (2.4)$$

for

$$Y = AX + N, \quad (2.5)$$

where A is Raleigh with $E[A^2] = \gamma$, $N \sim \mathcal{N}(0, 1)$, $X \in (-1, 1)$ for $0 \leq \gamma \leq 10$ dB.

2.2.2 Assuming that N is a constant, find an expression for P_e . Call this $P_e(N)$

2.2.3 For a function g ,

$$E[g(X)] = \int_{-\infty}^{\infty} g(x)p_X(x) dx \quad (2.6)$$

Find $P_e = E[P_e(N)]$.

2.2.4 Plot P_e in problems 2.2.1 and 2.2.3 on the same graph w.r.t γ . Comment.

Chapter 3

Bivariate Random Variables: FSK

3.1. Two Dimensions

Let

$$\mathbf{y} = A\mathbf{x} + \mathbf{n}, \quad (3.1)$$

where

$$x \in (\mathbf{s}_0, \mathbf{s}_1), \mathbf{s}_0 = \begin{pmatrix} 1 \\ 0 \end{pmatrix}, \mathbf{s}_1 = \begin{pmatrix} 0 \\ 1 \end{pmatrix} \quad (3.2)$$

$$\mathbf{n} = \begin{pmatrix} n_1 \\ n_2 \end{pmatrix}, n_1, n_2 \sim \mathcal{N}(0, 1). \quad (3.3)$$

3.1.1 Plot

$$\mathbf{y}|\mathbf{s}_0 \text{ and } \mathbf{y}|\mathbf{s}_1 \quad (3.4)$$

on the same graph using a scatter plot.

3.1.2 For the above problem, find a decision rule for detecting the symbols \mathbf{s}_0 and \mathbf{s}_1 .

3.1.3 Plot

$$P_e = \Pr(\hat{\mathbf{x}} = \mathbf{s}_1 | \mathbf{x} = \mathbf{s}_0) \quad (3.5)$$

with respect to the SNR from 0 to 10 dB.

- 3.1.4 Obtain an expression for P_e . Verify this by comparing the theory and simulation plots on the same graph.

Chapter 4

Exercises

4.1. BPSK

1. The signal constellation diagram for BPSK is given by Fig. 4.1. The symbols s_0 and s_1 are equiprobable. $\sqrt{E_b}$ is the energy transmitted per bit. Assuming a zero mean additive white gaussian noise (AWGN) with variance $\frac{N_0}{2}$, obtain the symbols that are received.

Solution: The possible received symbols are

$$y|s_0 = \sqrt{E_b} + n \quad (4.1)$$

$$y|s_1 = -\sqrt{E_b} + n \quad (4.2)$$

where the AWGN $n \sim \mathcal{N}\left(0, \frac{N_0}{2}\right)$.

2. From Fig. 4.1 obtain a decision rule for BPSK

Solution: The decision rule is

$$y \underset{s_1}{\overset{s_0}{\gtrless}} 0 \quad (4.3)$$

3. Repeat the previous exercise using the MAP criterion.

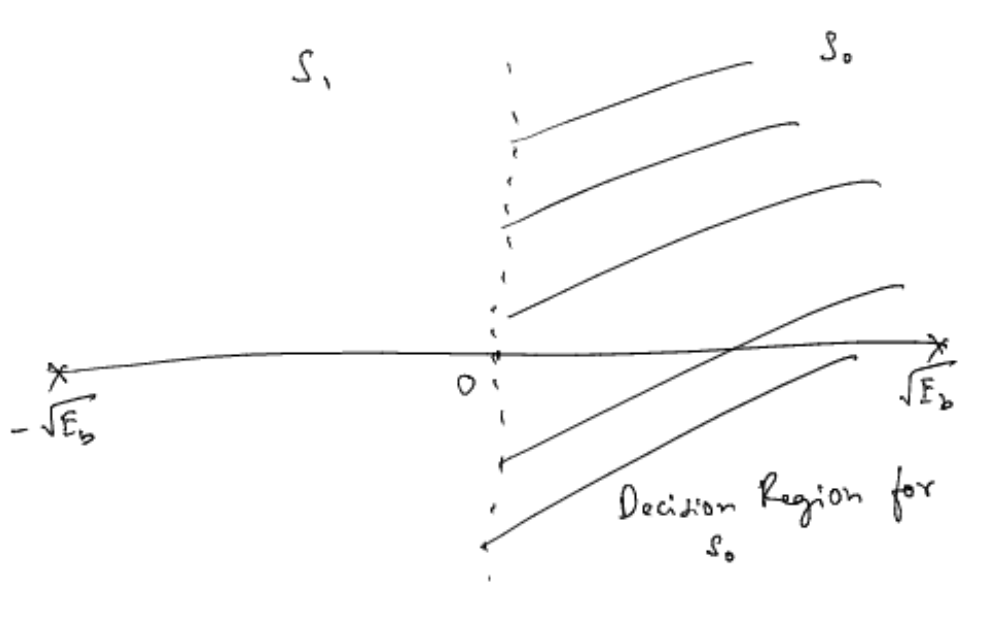


Figure 4.1:

4. Using the decision rule in Problem 2, obtain an expression for the probability of error for BPSK.

Solution: Since the symbols are equiprobable, it is sufficient if the error is calculated assuming that a 0 was sent. This results in

$$P_e = \Pr(y < 0 | s_0) = \Pr(\sqrt{E_b} + n < 0) \quad (4.4)$$

$$= \Pr(-n > \sqrt{E_b}) = \Pr(n > \sqrt{E_b}) \quad (4.5)$$

since n has a symmetric pdf. Let $w \sim \mathcal{N}(0, 1)$. Then $n = \sqrt{\frac{N_0}{2}}w$. Substituting this in

(4.5),

$$P_e = \Pr \left(\sqrt{\frac{N_0}{2}} w > \sqrt{E_b} \right) = \Pr \left(w > \sqrt{\frac{2E_b}{N_0}} \right) \quad (4.6)$$

$$= Q \left(\sqrt{\frac{2E_b}{N_0}} \right) \quad (4.7)$$

where $Q(x) \triangleq \Pr(w > x), x \geq 0$.

5. The PDF of $w \sim \mathcal{N}(0, 1)$ is given by

$$p_w(x) = \frac{1}{\sqrt{2\pi}} \exp \left(-\frac{x^2}{2} \right), -\infty < x < \infty \quad (4.8)$$

and the complementary error function is defined as

$$\text{erfc}(x) = \frac{2}{\sqrt{\pi}} \int_x^\infty e^{-t^2} dt. \quad (4.9)$$

Show that

$$Q(x) = \frac{1}{2} \text{erfc} \left(\frac{x}{\sqrt{2}} \right) \quad (4.10)$$

6. Verify the bit error rate (BER) plots for BPSK through simulation and analysis for 0 to 10 dB.

Solution: The following code

```
codes/bpsk_ber.py
```

yields Fig. 4.2

7. Show that

$$Q(x) = \frac{1}{\pi} \int_0^{\frac{\pi}{2}} e^{-\frac{x^2}{2 \sin^2 \theta}} d\theta \quad (4.11)$$

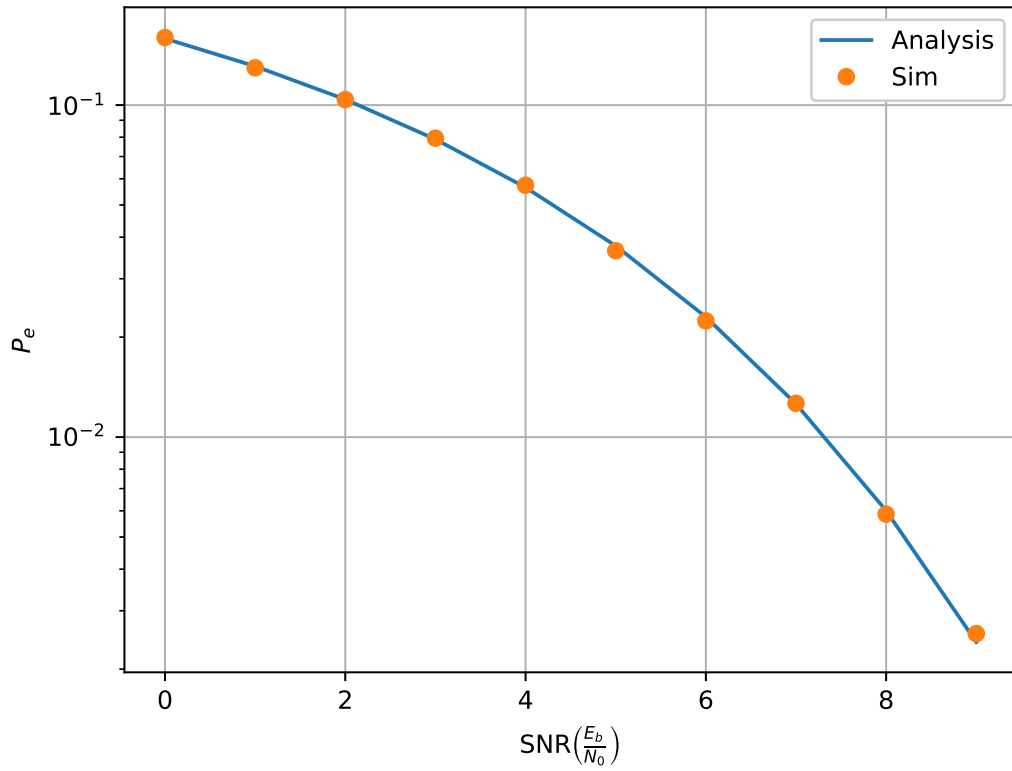


Figure 4.2:

4.2. Coherent BFSK

1. The signal constellation for binary frequency shift keying (BFSK) is given in Fig. 4.3.

Obtain the equations for the received symbols.

Solution: The received symbols are given by

$$\mathbf{y}|s_0 = \begin{pmatrix} \sqrt{E_b} \\ 0 \end{pmatrix} + \begin{pmatrix} n_1 \\ n_2 \end{pmatrix}, \quad (4.12)$$

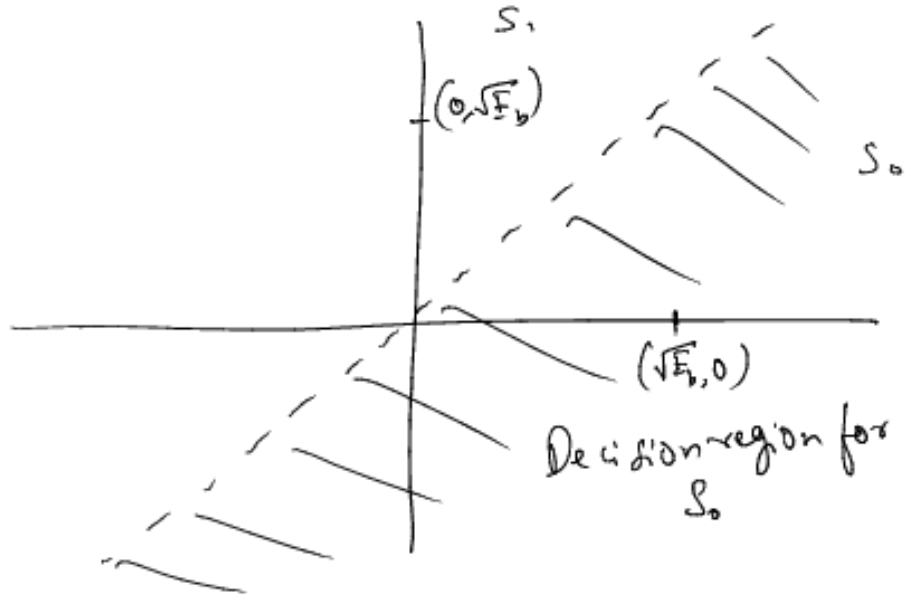


Figure 4.3:

and

$$\mathbf{y}|s_1 = \begin{pmatrix} 0 \\ \sqrt{E_b} \end{pmatrix} + \begin{pmatrix} n_1 \\ n_2 \end{pmatrix}, \quad (4.13)$$

where $n_1, n_2 \sim \mathcal{N}(0, \frac{N_0}{2})$. and $\mathbf{y} = \begin{pmatrix} y_1 \\ y_2 \end{pmatrix}$.

2. Obtain a decision rule for BFSK from Fig. 4.3.

Solution: The decision rule is

$$y_1 \underset{s_1}{\overset{s_0}{\gtrless}} y_2 \quad (4.14)$$

3. Repeat the previous exercise using the MAP criterion.

4. Derive and plot the probability of error. Verify through simulation.

4.3. QPSK

1. Let

$$\mathbf{r} = \mathbf{s} + \mathbf{n} \quad (4.15)$$

where $\mathbf{s} \in \{s_0, s_1, s_2, s_3\}$ and

$$\mathbf{s}_0 = \begin{pmatrix} \sqrt{E_b} \\ 0 \end{pmatrix}, \mathbf{s}_1 = \begin{pmatrix} 0 \\ \sqrt{E_b} \end{pmatrix}, \quad (4.16)$$

$$\mathbf{s}_2 = \begin{pmatrix} -\sqrt{E_b} \\ 0 \end{pmatrix}, \mathbf{s}_3 = \begin{pmatrix} 0 \\ -\sqrt{E_b} \end{pmatrix}, \quad (4.17)$$

$$E[\mathbf{n}] = \mathbf{0}, E[\mathbf{n}\mathbf{n}^T] = \sigma^2 \mathbf{I} \quad (4.18)$$

- (a) Show that the MAP decision for detecting \mathbf{s}_0 results in

$$|r|_2 < r_1 \quad (4.19)$$

- (b) Express $\Pr(\hat{\mathbf{s}} = \mathbf{s}_0 | \mathbf{s} = \mathbf{s}_0)$ in terms of r_1, r_2 . Let $X = n_2 - n_1, Y = -n_2 - n_1$, where $\mathbf{n} = (n_1, n_2)$. Their correlation coefficient is defined as

$$\rho = \frac{E[(X - \mu_x)(Y - \mu_y)]}{\sigma_x \sigma_y} \quad (4.20)$$

X and Y are said to be uncorrelated if $\rho = 0$

- (c) Show that if X and Y are uncorrelated Verify this numerically.

- (d) Show that X and Y are independent, i.e. $p_{XY}(x, y) = p_X(x)p_Y(y)$.
- (e) Show that $X, Y \sim \mathcal{N}(0, 2\sigma^2)$.
- (f) Show that $\Pr(\hat{\mathbf{s}} = \mathbf{s}_0 | \mathbf{s} = \mathbf{s}_0) = \Pr(X < A, Y < A)$.
- (g) Find $\Pr(X < A, Y < A)$.
- (h) Verify the above through simulation.

4.4. M -PSK

1. Consider a system where $\mathbf{s}_i = \begin{pmatrix} \cos(\frac{2\pi i}{M}) \\ \sin(\frac{2\pi i}{M}) \end{pmatrix}, i = 0, 1, \dots, M-1$. Let

$$\mathbf{r}|s_0 = \begin{pmatrix} r_1 \\ r_2 \end{pmatrix} = \begin{pmatrix} \sqrt{E_s} + n_1 \\ n_2 \end{pmatrix} \quad (4.21)$$

where $n_1, n_2 \sim \mathcal{N}(0, \frac{N_0}{2})$.

- (a) Substituting

$$r_1 = R \cos \theta \quad (4.22)$$

$$r_2 = R \sin \theta \quad (4.23)$$

show that the joint pdf of R, θ is

$$p(R, \theta) = \frac{R}{\pi N_0} \exp\left(-\frac{R^2 - 2R\sqrt{E_s} \cos \theta + E_s}{N_0}\right) \quad (4.24)$$

(b) Show that

$$\lim_{\alpha \rightarrow \infty} \int_0^{\infty} (V - \alpha) e^{-(V-\alpha)^2} dV = 0 \quad (4.25)$$

$$\lim_{\alpha \rightarrow \infty} \int_0^{\infty} e^{-(V-\alpha)^2} dV = \sqrt{\pi} \quad (4.26)$$

(c) Using the above, evaluate

$$\int_0^{\infty} V \exp \{ - (V^2 - 2V\sqrt{\gamma} \cos \theta + \gamma) \} dV \quad (4.27)$$

for large values of γ .

(d) Find a compact expression for

$$I = 1 - \sqrt{\frac{\gamma}{\pi}} \int_{-\frac{\pi}{M}}^{\frac{\pi}{M}} e^{-\gamma \sin^2 \theta} \cos \theta d\theta \quad (4.28)$$

(e) Find $P_{e|s_0}$.

4.5. Noncoherent BFSK

4.5.1 Show that

$$I_0(x) = \frac{1}{2\pi} \int_0^{2\pi} e^{x \cos \theta} d\theta \quad (4.29)$$

$$I_0(x) = \frac{1}{2\pi} \int_0^{2\pi} e^{x \cos(\theta-\phi)} d\theta \quad (4.30)$$

$$\frac{1}{2\pi} \int_0^{2\pi} e^{m_1 \cos \theta + m_2 \sin \theta} d\theta = I_0 \left(\sqrt{m_1^2 + m_2^2} \right) \quad (4.31)$$

where the modified Bessel function of order n (integer) is defined as

$$I_n(x) = \frac{1}{\pi} \int_0^\pi e^{x \cos \theta} \cos n\theta d\theta \quad (4.32)$$

4.5.2 Let

$$\mathbf{r}|0 = \sqrt{E_b} \begin{pmatrix} \cos \phi_0 \\ \sin \phi_0 \\ 0 \\ 0 \end{pmatrix} + \mathbf{n}_0, \mathbf{r}|1 = \sqrt{E_b} \begin{pmatrix} 0 \\ 0 \\ \cos \phi_1 \\ \sin \phi_1 \end{pmatrix} + \mathbf{n}_1 \quad (4.33)$$

where $\mathbf{n}_0, \mathbf{n}_1 \sim \mathcal{N}(\mathbf{0}, \frac{N_0}{2} \mathbf{I})$.

- (a) Taking $\mathbf{r} = (r_1, r_2, r_3, r_4)^T$, find the pdf $p(\mathbf{r}|0, \phi_0)$ in terms of $r_1, r_2, r_3, r_4, \phi, E_b$ and N_0 . Assume that all noise variables are independent.
- (b) If ϕ_0 is uniformly distributed between 0 and 2π , find $p(\mathbf{r}|0)$. Note that this expression will no longer contain ϕ_0 .
- (c) Show that the ML detection criterion for this scheme is

$$I_0 \left(k \sqrt{r_1^2 + r_2^2} \right) \stackrel{0}{\underset{1}{\gtrless}} I_0 \left(k \sqrt{r_3^2 + r_4^2} \right) \quad (4.34)$$

where k is a constant.

- (d) The above criterion reduces to something simpler. Can you guess what it is? Justify your answer.
- (e) Show that

$$P_{e|0} = \Pr(r_1^2 + r_2^2 < r_3^2 + r_4^2 | 0) \quad (4.35)$$

(f) Show that the pdf of $Y = r_3^2 + r_4^2$ is

$$p_Y(y) = \frac{1}{N_0} e^{-\frac{y}{N_0}}, y > 0 \quad (4.36)$$

(g) Find

$$g(r_1, r_2) = \Pr(r_1^2 + r_2^2 < Y | 0, r_1, r_2). \quad (4.37)$$

(h) Show that $E \left[e^{-\frac{X^2}{2\sigma^2}} \right] = \frac{1}{\sqrt{2}} e^{-\frac{\mu^2}{4\sigma^2}}$ for $X \sim \mathcal{N}(\mu, \sigma^2)$.

(i) Now show that

$$E[g(r_1, r_2)] = \frac{1}{2} e^{-\frac{E_b}{2N_0}}. \quad (4.38)$$

4.5.3 Let $U, V \sim \mathcal{N}(0, \frac{k}{2})$ be i.i.d. Assuming that

$$U = \sqrt{R} \cos \Theta \quad (4.39)$$

$$V = \sqrt{R} \sin \Theta \quad (4.40)$$

(a) Compute the jacobian for U, V with respect to R and Θ defined by

$$J = \det \begin{pmatrix} \frac{\partial U}{\partial R} & \frac{\partial U}{\partial \Theta} \\ \frac{\partial V}{\partial R} & \frac{\partial V}{\partial \Theta} \end{pmatrix} \quad (4.41)$$

(b) The joint pdf for R, Θ is given by,

$$p_{R,\Theta}(r, \theta) = p_{U,V}(u, v) J|_{u=\sqrt{r} \cos \theta, v=\sqrt{r} \sin \theta} \quad (4.42)$$

Show that

$$p_R(r) = \begin{cases} \frac{1}{k} e^{-\frac{r}{k}} & r > 0, \\ 0 & r < 0, \end{cases} \quad (4.43)$$

assuming that Θ is uniformly distributed between 0 to 2π .

- (c) Show that the pdf of $Y = R_1 - R_2$, where R_1 and R_2 are i.i.d. and have the same distribution as R is

$$p_Y(y) = \frac{1}{2k} e^{-\frac{|y|}{k}} \quad (4.44)$$

- (d) Find the pdf of

$$Z = p + \sqrt{p} [U \cos \phi + V \sin \phi] \quad (4.45)$$

where ϕ is a constant.

- (e) Find $\Pr(Y > Z)$.

- (f) If $U \sim \mathcal{N}(m_1, \frac{k}{2})$, $V \sim \mathcal{N}(m_2, \frac{k}{2})$, where m_1, m_2, k are constants, show that the pdf of

$$R = \sqrt{U^2 + V^2} \quad (4.46)$$

is

$$p_R(r) = \frac{e^{-\frac{r+m}{k}}}{k} I_0\left(\frac{2\sqrt{mr}}{k}\right), \quad m = \sqrt{m_1^2 + m_2^2} \quad (4.47)$$

(g) Show that

$$I_0(x) = \sum_{n=0}^{\infty} \frac{x^{2n}}{4^n (n!)^2} \quad (4.48)$$

(h) If

$$p_Z(z) = \begin{cases} \frac{1}{k} e^{-\frac{z}{k}} & z \geq 0 \\ 0 & z < 0 \end{cases} \quad (4.49)$$

find $\Pr(R < Z)$.

4.6. Craig's Formula and MGF

4.6.1 The Moment Generating Function (MGF) of X is defined as

$$M_X(s) = E[e^{sX}] \quad (4.50)$$

where X is a random variable and $E[\cdot]$ is the expectation.

(a) Let $Y \sim \mathcal{N}(0, 1)$. Define

$$Q(x) = \Pr(Y > x), x > 0 \quad (4.51)$$

Show that

$$Q(x) = \frac{1}{\pi} \int_0^{\frac{\pi}{2}} e^{-\frac{x^2}{2 \sin^2 \theta}} d\theta \quad (4.52)$$

(b) Let $h \sim \mathcal{CN}(0, \frac{\Omega}{2})$, $n \sim \mathcal{CN}(0, \frac{N_0}{2})$. Find the distribution of $|h|^2$.

(c) Let

$$P_e = \Pr(\Re\{h^*y\} < 0), \text{ where } y = \left(\sqrt{E_s}h + n\right), \quad (4.53)$$

Show that

$$P_e = \int_0^\infty Q\left(\sqrt{2x}\right) p_A(x) dx \quad (4.54)$$

where $A = \frac{E_s|h|^2}{N_0}$.

(d) Show that

$$P_e = \frac{1}{\pi} \int_0^{\frac{\pi}{2}} M_A\left(-\frac{1}{\sin^2\theta}\right) d\theta \quad (4.55)$$

(e) compute $M_A(s)$.

(f) Find P_e .

(g) If $\gamma = \frac{\Omega E_s}{N_0}$, show that $P_e < \frac{1}{2\gamma}$.

