

1 **Estimation of spatiotemporal transmission dynamics and**
2 **analysis of management scenarios for sea lice of farmed and**
3 **wild salmon**

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22 Abstract

23 Parasite transmission between farmed and wild salmon affects the sustainability of salmon
24 aquaculture in Pacific Canada. Understanding and managing parasites in aquaculture is
25 challenged by spatial and temporal variation in transmission dynamics. We developed a
26 mechanistic model that connects sea-louse (*Lepeoptheirus salmonis*) outbreak and control
27 on farmed salmon (*Salmo salar*) to spatiotemporal dynamics of sea lice on migrating wild
28 juvenile salmon (*Oncorhynchus keta* and *O. gorbuscha*). We fitted the model to timeseries
29 of sea lice on farmed salmon and spatial surveys of juvenile wild salmon in the Broughton
30 Archipelago. We used the parameterized model to evaluate alternative management sce-
31 narios based on the resulting sea-louse infestations and predicted mortality of wild salmon.
32 Early and coordinated management of sea lice on salmon farms was most effective for
33 controlling outbreaks in wild salmon, while uncoordinated treatments led to a resurgence
34 of sea lice on salmon farms during the juvenile salmon migration. This study highlights
35 the importance of incorporating spatiotemporal variability when considering infectious dis-
36 ease dynamics shared by farmed and wild hosts, particularly when migratory wildlife are
37 involved.

38 **Keywords:** aquaculture, disease, modelling, sea lice, salmon

39 Introduction

40 The effective management of disease depends on a solid understanding of the spatial and
41 temporal processes affecting transmission dynamics (Keeling and Eames 2005; White et al.
42 2018). For wildlife, disease dynamics can be complicated by the movement of hosts over
43 large distances and the associated variability in infection pressure that they experience
44 (Altizer et al. 2011). Further, at any given location, temporal changes in infection pres-
45 sure due to seasonality of both parasite and host lifecycles (Altizer et al. 2006) or disease
46 dynamics in reservoir hosts (Krkošek et al. 2006a) may result in fluctuating sources of infec-
47 tion (Hudson et al. 2002). Examples include social dynamics and dispersal of badgers that
48 underlie spatiotemporal variation of infection risk of bovine tuberculosis to domestic cattle
49 (Delahay et al. 2000), oscillations in measles dynamics in urban centres causing periodic
50 traveling waves of infection through rural communities (Grenfell et al. 2001), and parasite
51 dispersal from and control on salmon farms causing spatiotemporal variation in infection
52 risk along migration routes of wild salmon (Krkošek et al. 2010). Failure to consider the
53 inherent spatiotemporal variability in infection dynamics can lead to erroneous conclu-
54 sions about the risk of infectious disease to host populations and ineffective management
55 recommendations.

56 Sea-louse transmission between farmed salmon and migrating wild salmon in coastal envi-
57 ronments is one example of a system that exhibits strong spatial and temporal variability
58 (Rees et al. 2015; Groner et al. 2016). Sea lice (*Lepeoptheirus salmonis*) are copepod
59 parasites that infect salmonids, feeding on epidermal tissues, muscle, and blood (Costello
60 2006). Sea lice hatch as free-living and non-feeding nauplii that can disperse tens of kilo-
61 metres in ocean currents before finding a suitable host, attaching, and developing through
62 copepodite, chalimus, and motile adult stages (Stucchi et al. 2011; Amundrud and Murray

2009). High infestation intensities on adult hosts may cause host morbidity and mortality (Pike and Wadsworth 2000) and have significant biological and economic impacts on the salmon-farming industry (Costello 2009; Abolofia et al. 2017). Although Pacific salmon (*Oncorhynchus* spp.) have been found to be more resistant to infestation (Jones et al. 2008; Johnson and Albright 1992), even low abundances on small, juvenile salmon may cause mortality (Morton et al. 2005) or sub-lethal effects on physiology (Nendick et al. 2011; Brauner et al. 2012) and behaviour (Krkošek et al. 2011a; Godwin et al. 2015).

Open-net salmon farms provide a reservoir host for sea lice that results in spatial variability in infestation pressure along the migration routes of wild juvenile salmon (Krkošek et al. 2006a). Transmission of sea lice from farmed salmon has been implicated in epizootics of wild salmon in Norway (Bjørn et al. 2001), Ireland (Gargan et al. 2003), Scotland (Butler and Watt 2003), and Canada (Krkošek et al. 2006a). In Pacific Canada, in particular, these epizootics pose a conservation risk to ecologically, culturally, and economically important wild salmon (Krkošek et al. 2011b). Out-migrating juvenile pink (*O. gorbuscha*) and chum (*O. keta*) salmon are most vulnerable due to their small size and underdeveloped scales when they enter the marine environment (Brauner et al. 2012).

The strength of infestation pressure from salmon farms changes depending on environmental conditions such as temperature and salinity that affect sea-louse development and survival, but more importantly on the management of sea lice by the industry (Rogers et al. 2013). The cost of sea lice to the salmon-farming industry is on the order of hundreds of millions of US dollars annually (Abolofia et al. 2017), and there have been numerous management strategies to reduce louse abundances on farmed salmon. In Pacific Canada, in-feed treatments with emamectin benzoate (EMB; trade-name SLICE) are the most common treatment for sea lice, and farms are required to treat (or harvest) if the number of sea lice per farmed salmon exceeds three motile lice (Fisheries and Oceans Canada 2018).

88 For the most part, EMB has been effective at reducing sea-louse infestations of farmed
89 salmon (Saksida et al. 2010), although sporadic and localized reports of resistance to the
90 drug among Pacific sea lice indicate this may change (Messmer et al. 2018). Nonetheless,
91 chemical treatments have and continue to be a strong driver of sea-louse population dy-
92 namics on salmon farms (e.g., Krkošek et al. 2010), which in turn influence infestations of
93 sympatric wild salmon.

94 Although the environmental and management factors affecting sea-louse dynamics on
95 farmed and wild salmon have received considerable attention (e.g., Revie et al. 2005; Jansen
96 et al. 2012; Rogers et al. 2013; Bateman et al. 2016), no studies have integrated the spatial
97 dynamics of wild salmon migration and temporal dynamics of sea lice on farmed salmon.
98 Inter-annual changes in average parasite abundance show a positive correlation between
99 sea lice on farmed salmon and infestations of wild juvenile salmon (Marty et al. 2010;
100 Peacock et al. 2013). However, within-year transmission dynamics that govern if and how
101 an outbreak will emerge are mediated by ocean currents, the lifecycle of sea lice, and mi-
102 gration patterns of wild salmon in relation to farms. In order to understand how chemical
103 treatments on farms influence transmission dynamics of sea lice from farm to wild salmon,
104 we must explicitly consider such complexity.

105 Some of these complex physical and biological processes have been modelled with respect to
106 sea-louse dispersal from salmon farms. At a fine resolution, Stucchi et al. (2011) conducted
107 particle-tracking simulations that captured the three-dimensional dispersal of sea-louse
108 nauplii from a salmon farm, including the effects of wind, tides, and freshwater input
109 as well as the vertical migration of sea louse nauplii and the effects of temperature and
110 salinity on sea-louse survival and development. These simulations were more recently
111 expanded on to examine the roles of individual farms within the Broughton Archipelago,
112 Canada as “emitters” or “receivers” of sea-louse infestations (Cantrell et al. 2018). At a

113 coarser scale, Aldrin et al. (2013) modelled dispersal using simple seaway distance metrics,
 114 but included transmission of lice among numerous salmon farms and from non-specified
 115 reservoirs for sea lice such as wild salmonids. Complex simulation models like the former
 116 can suggest major drivers of spatiotemporal variability in infestations and yield specific,
 117 detailed predictions, but one advantage of simpler models is that they can be fit to data to
 118 infer unknown parameters. This simpler approach has been taken to quantify the relative
 119 importance of salmon farms in driving sea-louse infestations of wild salmon. Krkošek
 120 et al. (2005a, 2006a) modelled the broad-scale ocean currents as an advection-diffusion
 121 process, yielding a steady-state spatial distribution of infectious sea lice around salmon
 122 farms, and the subsequent attachment and development of sea lice on migrating wild
 123 salmon. However, unlike the previously mentioned studies, Krkošek et al. (2005a, 2006a)
 124 ignored the sea-louse population dynamics on salmon farms and considered farms to be a
 125 constant source of infectious-stage sea lice. Due to the large temporal fluctuations in sea-
 126 louse numbers on farmed salmon (e.g., Krkošek et al. 2010; Jansen et al. 2012), ignoring
 127 the source dynamics could lead to erroneous conclusions about how farm management
 128 influences parasite population dynamics as well as survival of wild salmon.

129 In this study, we develop a mechanistic model that connects temporal dynamics of sea-louse
 130 populations on farmed salmon to the spatiotemporal infestations of wild juvenile salmon
 131 and fit this model to data from sea-louse monitoring of both farmed and wild salmon
 132 in the Broughton Archipelago, Canada. We use the fitted model to evaluate different
 133 farm management scenarios by the resulting juvenile salmon infestation dynamics, building
 134 on previous studies that have evaluated treatment timing based on effective control on
 135 salmon farms alone (e.g., Revie et al. 2005). The results may inform the management of
 136 salmon farms for the benefit of wild salmon, but are also an example of how spatiotemporal
 137 variability in infection pressure can be incorporated into models used to inform management

138 of diseases.

139 **Methods**

140 **Model**

141 We connected sea-louse infestations of farmed salmon to observed louse abundances on
 142 juvenile wild salmon using a mechanistic model that includes sea-louse population dynamics
 143 on farms in response to parasite control (Krkošek et al. 2010), dispersal of sea lice from
 144 farms, and infestation and development of lice on wild juvenile salmon (Krkošek et al.
 145 2005a) (Figure 1).

146 Sea-louse populations on farmed salmon tend to grow exponentially in the absence of
 147 control and decline exponentially after treatment with parasiticide (Krkošek et al. 2010;
 148 Rogers et al. 2013). These temporal fluctuations in louse abundance on farmed salmon
 149 impact the infestation pressure on juvenile wild salmon migrating past salmon farms. We
 150 modelled the average number of motile sea lice per farmed salmon as:

$$(1) \quad f(t) = f_0 \begin{cases} e^{r_1(t-t_0)} & t < t_0 \\ e^{r_2(t-t_0)} & t \geq t_0, \end{cases}$$

151 where $f(t)$ is the average number of motile sea lice per farmed salmon at time t , r_1 is the
 152 population growth of lice before treatment and r_2 is the rate of decay after a treatment,
 153 f_0 is the average number of motile sea lice per farmed salmon at the time of treatment,
 154 and t_0 is the treatment date. We assumed that host population size on salmon farms is

approximately constant when farms are stocked, and thus the effect of host population size on sea louse population growth is included in the growth parameter for each farm, r_1 (Krkošek et al. 2010). Equation (1) does not include negative or positive density dependence of sea lice. Negative density dependence is unlikely on farmed hosts due to management interventions at low to moderate louse densities. Although several modelling studies have included positive density dependence (i.e., mate limitation at low densities; Groner et al. 2014; McEwan et al. 2019), we found that the exponential growth model fit the data well, perhaps because mate limitation is less likely on farmed salmon that are larger and in higher densities than juvenile wild salmon (Cox 2017). We also found that sea lice were overdispersed on farmed hosts (see Results), which reduces the probability of mate limitation (Stormoen et al. 2013).

Several studies have found that in-feed treatments with EMB are effective for a period of approximately three months (e.g., Saksida et al. 2010; Rogers et al. 2013). Therefore, at time $t_0 + 90$ days, we assumed the treatment efficacy to have declined to the extent that sea-louse population growth was again possible, and the growth rate returned to r_1 .

Sea-louse nauplii hatch from gravid motile sea lice and can disperse tens of kilometres from the open-net pens containing farmed salmon (Foreman et al. 2009). The temporal dynamics of sea-louse populations at discrete farm locations can therefore lead to spatiotemporal patterns of infestation pressure on juvenile wild salmon migrating past farms. To capture this spatial dimension, we considered the dispersal of naupliar sea lice from salmon farms along the migration corridor of juvenile salmon through Knight Inlet-Tribune Channel (Figure 2). This migration corridor is much longer (>100 km) than it is wide (≈ 1 km), and in this model we follow Krkošek et al. (2005a, 2006a) and consider the migration corridor to be a one-dimensional domain along which sea lice disperse and juvenile salmon migrate. The dispersal of nauplii described by,

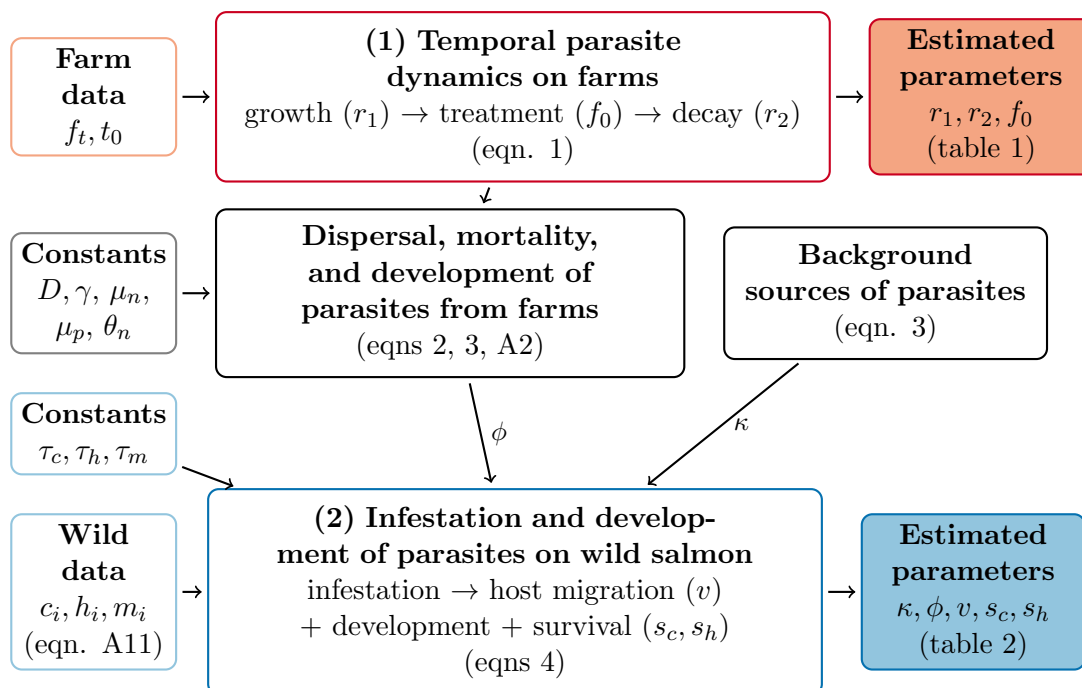


Figure 1: Schematic of the modelling framework for the sea-louse transmission model incorporating temporal dynamics of lice on salmon farms. The model was fit to data in two stages; (1) the farm dynamics were fit to counts of lice on farmed salmon (red), and (2) the final predictions of lice per wild juvenile salmon were confronted with data of sea-louse abundances on wild salmon throughout their migration route (blue). [Colour online.]

$$(2) \quad \frac{\partial n_i}{\partial t} = \underbrace{f_i(t) \omega S_i(t) \delta(x - y_i)}_{(1) \text{ production}} + \underbrace{D \frac{\partial^2 n_i}{\partial x^2} - \gamma \frac{\partial n_i}{\partial x}}_{(2) \text{ diffusion-advection}} - \underbrace{(\mu_n + \theta_n) n_i}_{(3) \text{ mortality-development}},$$

where $n_i(x, t)$ is the density of nauplii at location x and day t , originating from farm i . Nauplii are produced by motile sea lice at farm i , where $f_i(t)$ is the average number of motile lice per salmon on farm i from equation 1, ω is an unknown parameter for the fecundity of motile lice times the dilution of nauplii in three dimensions, and $S_i(t)$ is the number of farmed salmon on farm i at time t (Marty et al. 2010). In our case, the number of salmon in each farm was similar and relatively constant throughout the period considered, and so we assume that $\omega S_i(t)$ is constant (see Figure S1 for comparison of average vs. total motile *L. salmonis* per farmed salmon). The production of sea lice from farmed salmon occurs at exactly location y_i and is described by a Delta function, $\delta(x - y_i)$, which assumes that the length of the farm is small (i.e., on the scale of metres) relative to the spatial domain of the juvenile salmon migration route (tens of kilometres). The inclusion of a time-varying source of sea lice from salmon farms is a novel development from previous work (Krkošek et al. 2005a), and one that is necessary when lice on farmed salmon are changing dramatically in response to treatment.

The second part of equation 2 captures random diffusion due to winds and tides, where D is the diffusion coefficient, and a general seaward advective flow due to high freshwater influx at the heads of inlets, where γ is the advection coefficient. These parameters have been estimated in previous studies as $D = 22.67 \text{ km}^2 \cdot \text{day}^{-1}$ (Krkošek et al. 2006a) and $\gamma = 1.56 \text{ km} \cdot \text{day}^{-1}$ (Brooks 2005), and we fix them at these values to avoid identifiability problems with other parameters in our model (e.g., the migration speed of salmon - see

below), which are directly confounded.

The third part of equation 2 describes natural mortality of nauplii at rate μ_n and development into copepodites at rate θ_n . Experimental data of nauplii mortality and development rates indicate $(\mu_n + \theta_n) = 4/5 \text{ day}^{-1}$ (Krkošek et al. 2006b). We calculated the solution to equation 2, yielding the density of nauplii at any point along the migration x and time t , by numerically convolving the solution for advection-diffusion-decay (parts 2 and 3 of equation 2), known as a Green's function, with the production term (part 1) (Polyanin and Nazaikinskii 2016). Fixing the diffusion, advection, mortality, and development parameters allowed us to calculate the distribution of nauplii outside of the estimation of parameters for infestation and survival of wild salmon (see below) and thus increased computational efficiency and feasibility of model fitting. In the Discussion, we further justify these assumptions and consider errors they may introduce. Details of this solution are given in the Appendix.

Nauplii develop into copepodites which can attach to susceptible juvenile salmon in the vicinity. These copepodites diffuse and advect via the same process described by equation 2, except with mortality $\mu_p = 1/5 \text{ day}^{-1}$ (Krkošek et al. 2006b) and production $\theta_n n_i(x, t)$. We calculated the distribution of copepodites numerically by convolving the distribution of nauplii with the Green's function described above (see Appendix for details). The total distribution of farm-source copepodites is the sum of the copepodid densities from all farms along the migration route, which we call $L_1(x, t)$. The infestation pressure for migrating wild salmon is a combination of background and farm sources of sea lice,

$$(3) \quad L(x, t) = \kappa + \phi L_1(x, t),$$

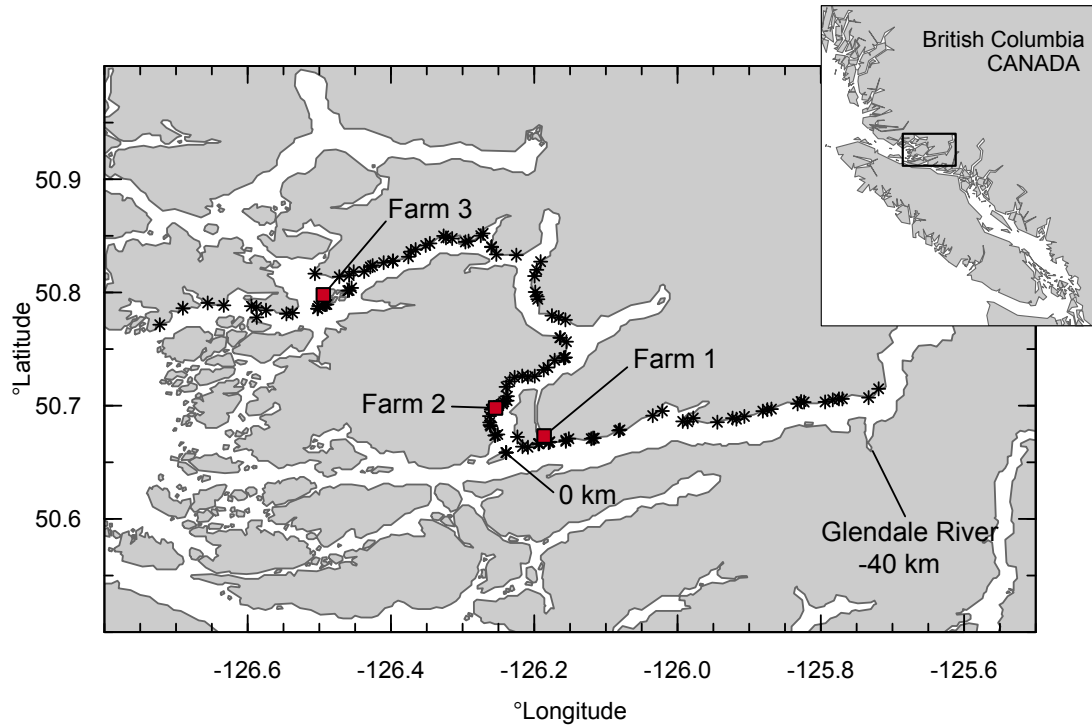


Figure 2: Models were fit to data from the Broughton Archipelago, on the south central coast of British Columbia, Canada. Sampling of farmed salmon took place on the three active salmon farms in the study region in 2006 (red squares). Approximately 100 juvenile wild pink and chum salmon were sampled every ~ 1 km (black stars) along their migration past these farms along the Knight Inlet - Tribune Channel migration corridor. The model was simulated from -60 km to 80 km along the migration corridor, with 0 km being a reference point at the confluence of Knight Inlet and Tribune Channel. The map was produced using the R package PBSmapping (Schnute et al. 2018) with shoreline data from Wessel and Smith (2016). [Colour online.]

where κ is the background density of copepodites from distributed sources (e.g., returning wild adult salmon) and $\phi = \omega S_i(t)$ is the unknown fecundity/dilution parameter for sea lice from farmed salmon from equation 2.

Copepodites attach to juvenile salmon and subsequently develop through chalimus and motile stages. The expected number of attached copepodid, chalimus and motile lice on a wild juvenile salmon at any point (x, t) is proportional to the density of infectious copepodites that fish previously encountered along its migration (Krkošek et al. 2005a). We assumed that wild salmon migrate at a constant speed v , and calculate the expected number of sea lice on a wild juvenile salmon as the integral of the distribution of infectious copepodites, $L(x, t)$, along the migration path of juvenile salmon through space and time (i.e., the line integral),

$$(4a) \quad C(x, t) = \beta \int_0^{\tau_c} L(x - vu, t - u) du$$

$$(4b) \quad H(x, t) = \beta s_c \int_{\tau_c}^{\tau_c + \tau_h} L(x - vu, t - u) du$$

$$(4c) \quad M(x, t) = \beta s_c s_h \int_{\tau_c + \tau_h}^{\tau_c + \tau_h + \tau_m} L(x - vu, t - u) du,$$

where $C(x, t)$, $H(x, t)$ and $M(x, t)$ are the expected number of attached copepodid, chalimus and motile lice on a juvenile salmon at km x and day t , τ_c , τ_h and τ_m are the developmental times of copepodites, chalimus and motiles, and s_c and s_h are the survival of copepodites and chalimus to the next stage. The developmental times of copepodites, chalimus and motiles (τ_c , τ_h and τ_m) have been previously estimated, and therefore we assumed developmental to be constant at their 10° C averages (Stien et al. 2005, Table A1). This assumption seemed reasonable given that the average temperature over the period of

239 wild salmon sampling was $9.5^{\circ} \text{ C} \pm 0.12$ (mean \pm SE).

240 We modelled the transmission coefficient, β , to be a Gamma random variable with mean β_0
 241 and shape parameter r . Variability in transmission coefficients among individual juvenile
 242 salmon within a school may occur due to heterogeneity in host susceptibility and/or small-
 243 scale patchiness in the distribution of copepodites due to swarming (Murray 2002), and
 244 leads to overdispersion of parasites among hosts. We assumed that infestation occurred as
 245 a Poisson process, which (with the Gamma-distributed transmission coefficient) gave rise
 246 to a Gamma-Poisson process that captures this overdispersion (Greenwood and Yule 1920)
 247 and is equivalent to a negative binomial distribution of sea lice among hosts. Previous
 248 models have assumed that sea lice are evenly dispersed among hosts according to the
 249 Poisson distribution with constant rate parameter (Krkošek et al. 2005a, 2006a). Details
 250 of the infestation model are given in the Appendix.

251 Model fitting

252 The sea-louse transmission model was fit to data collected in the Broughton Archipelago
 253 (BA), on the west coast of British Columbia, Canada (Figure 2), in 2006. In this region,
 254 wild juvenile pink and chum salmon migrate through narrow fjords to the open ocean each
 255 spring, and in the year of the study, these migrating salmon passed by several active salmon
 256 farms. The model was fit to data of sea-louse abundances on both farmed and wild salmon
 257 in two steps.

258 First, we obtained counts of motile-stage *L. salmonis* on salmon from three farms along the
 259 Knight Inlet - Tribune Channel migration corridor of the BA (Figure 2) and were active in
 260 2006. The number of motile lice on individual farmed salmon was available from November
 261 2005 through June 2006 from a previous study of sea-louse dynamics on Farm 1 ($n = 1659$

262 fish sampled) and Farm 2 ($n = 1080$ fish sampled) (Krkošek et al. 2010). For Farm 3, we
263 had the total number of motile lice per sampling event (which included 20 fish), totalling 39
264 sampling events from July 2005 through June 2006 (Cohen Commission 2011). Sea-louse
265 counts were done at irregular intervals of anywhere from 7 to 44 days, with more frequent
266 sampling during the juvenile salmon outmigration (Figure 4).

267 We fit the model of exponential growth and decay of motile lice on farmed salmon (equation
268 1) to counts of sea lice on each farm separately. The different format of the data for Farms
269 1 and 2 versus Farm 3 required slightly different assumptions in the statistical analyses; for
270 Farms 1 and 2 we assumed that the number of lice per fish was distributed according to
271 the negative binomial with mean predicted by equation 1 and overdispersion parameter, k
272 to be estimated, and for Farm 3 we assumed normally distributed error between the mean
273 number of lice per fish from the 20-fish sample and the model prediction, with the residual
274 variance to be estimated. Maximum likelihood estimates for growth rate parameters (r_1
275 and r_2) and the average number of lice per farmed salmon at the time of treatment (f_0)
276 were then used to simulate a distribution of infectious copepodites originating from all
277 three salmon farms throughout the migration corridor for the period of the juvenile salmon
278 migration, using the advection-diffusion-decay processes described by equation 2 (with
279 details in the Appendix).

280 We fit the model of infestation and development of sea lice on wild juvenile salmon (equa-
281 tions 4-5) to spatiotemporal data of sea-louse abundance on wild juvenile salmon. Wild
282 juvenile pink and chum salmon were collected by beach seine along 60 km of the Knight
283 Inlet-Tribune Channel corridor (Figure 2). At each site, a maximum of 100 pink salmon
284 and 100 chum salmon were live-assayed for sea lice (Krkošek et al. 2005a,b). Fish collection
285 and examination protocols were approved by the University of Alberta Animal Care Com-
286 mittee and carried out in accordance with the Guide to the Care and Use of Experimental

287 Animals (www.ccac.ca). A total of 128 sites were sampled from April 10 - May 22, 2006,
 288 totalling 6593 pink, 6016 chum salmon samples, and 6428 associated *L. salmonis*. The
 289 surface water temperature at sample locations ranged from 7 – 16°C (mean 9.5°C), and
 290 salinities ranged from 9 – 33‰ (mean 28.3‰). Some of these data have been analysed
 291 previously as part of a larger project involving sea-louse data from several sources (the
 292 Broughton Archipelago Monitoring Program; Rees et al. 2015; Patanasatienkul et al. 2015;
 293 Cox 2017).

294 By assuming infestation occurred as a Gamma-Poisson process (see Appendix) we were able
 295 to assign probabilities to each of our observations of sea lice on wild juvenile salmon and
 296 calculate the likelihood of these data given a parameter set. Maximum likelihood estimates
 297 for both the growth and decay of sea-louse populations on farmed salmon and transmission
 298 of sea lice to wild salmon were obtained using a statistical tool called data cloning (Lele
 299 et al. 2010, 2007). Briefly, data cloning applies the Markov Chain Monte Carlo (MCMC)
 300 algorithm in a Bayesian framework to obtain maximum likelihood parameter estimates
 301 (Lele et al. 2007). Due to the complexity of the model and potential for parameters to be
 302 confounded and thus non-identifiable or non-estimable given our data, we also investigated
 303 the estimability of model parameters using data cloning (Peacock et al. 2017). If param-
 304 eters are estimable, the theory of data cloning posits that the variance in the posterior
 305 distribution should decline at a rate of $1/K$ when the likelihood is raised to the power K
 306 (or, equivalently, the data are “cloned” K times)(Lele et al. 2010). We implemented data
 307 cloning in R (R Development Core Team 2018) using the software JAGS (Plummer 2003)
 308 and package *dclone* (Sólymos 2010). In fitting the transmission model to sea-louse data
 309 from wild salmon, we used the simulated distribution of infectious sea lice from salmon
 310 farms over a 150-day period from January 1, 2006 to May 31, 2006 with a time-step of
 311 0.05 days and over a 140-km-long migration corridor (Figure 1) with a grid space of 0.05

312 km. This grid was sufficient to cover the period and locations of wild-salmon sampling
313 for model fitting. Details of the model fitting methodology and results are provided as
314 supplementary material online and R code is available (see Data Accessibility).

315 The free parameters that we estimated were the background louse density (κ), the fecun-
316 dity/dilution parameter controlling the rate of production of nauplii at farm locations (ϕ),
317 juvenile salmon migration speed (v), survival of copepodites and chalimus to the next stage
318 (s_c and s_h), and the shape parameter (i.e., dispersion parameter of the negative binomial
319 distribution; r) (Figure 1, Table A1). We were unable to estimate the mean transmission
320 coefficient β_0 because we lack data on planktonic sea-louse densities and thus this param-
321 eter was confounded. Therefore we could only estimate the parameter groupings of $\beta_0\kappa$
322 and $\beta_0\phi$ representing the background infestation pressure and farm infestation pressure.
323 The survival of attached lice on pink and chum salmon was not assumed to be the same
324 because, for example, the immune response may differ between host species (Jones et al.
325 2007; Sutherland et al. 2014). To consider the impact of this on the attachment process,
326 we also fit a model allowing $\beta_0\kappa$ to differ for pink and chum hosts, but the estimates were
327 not significantly different between host species. All other parameters were the same for
328 both pink and chum salmon.

329 Simulations

330 We used the parameterized model to explore the effect of the timing of treatments rela-
331 tive to the wild salmon migration and relative to treatments on other farms on sea-louse
332 infestations of juvenile pink salmon in a simulation framework. Previous studies of farm
333 networks have found that the timing of treatments among salmon farms influences the rate
334 of sea-louse population recovery, and thus the frequency of treatments needed within a

production cycle (e.g., Revie et al. 2005; Peacock et al. 2016). However, the influence of treatment timing relative to juvenile salmon migration has not been investigated (although see Bateman et al. (2016)), and we aimed to understand whether coordinated treatments are beneficial for wild salmon and, if so, when treatments should occur. We investigated four different treatment scenarios (Figure 3):

- (A) independent treatments on farms at the observed date (Figure 4; Figure 3a),
- (B) independent but immediate treatment of each farm when the louse abundance reached the treatment threshold of three motile lice per fish (Fisheries and Oceans Canada 2018, Figure 3b),
- (C) coordinated treatments of the three farms when the first farm reaches the treatment threshold on November 18, 2005 (even though the two other farms are below the threshold at that time; Figure 3c), and
- (D) coordinated treatments of all three farms on February 1, 2006, prior to the juvenile salmon outmigration (Rogers et al. 2013), even if it means delaying treatment of some farms after they have reached the threshold (Figure 3d).

For each scenario, louse abundances on the three farms were simulated using the growth rates, r_1 and r_2 , estimated for each farm (Table 1). For each farm, simulations started at the predicted louse abundance on September 1, 2005 (Figure 4) and ran to July 1, 2006, with one treatment per farm at the date specified by the scenario (Figure S5). We assumed that treatment efficacy lasted 90 days (Saksida et al. 2010), after which time the growth rate changed from r_2 back to r_1 (Rogers et al. 2013).

The migration path taken by a juvenile salmon through the simulated density of infectious larvae will influence the infestation pressure they encounter and the intensity of the result-

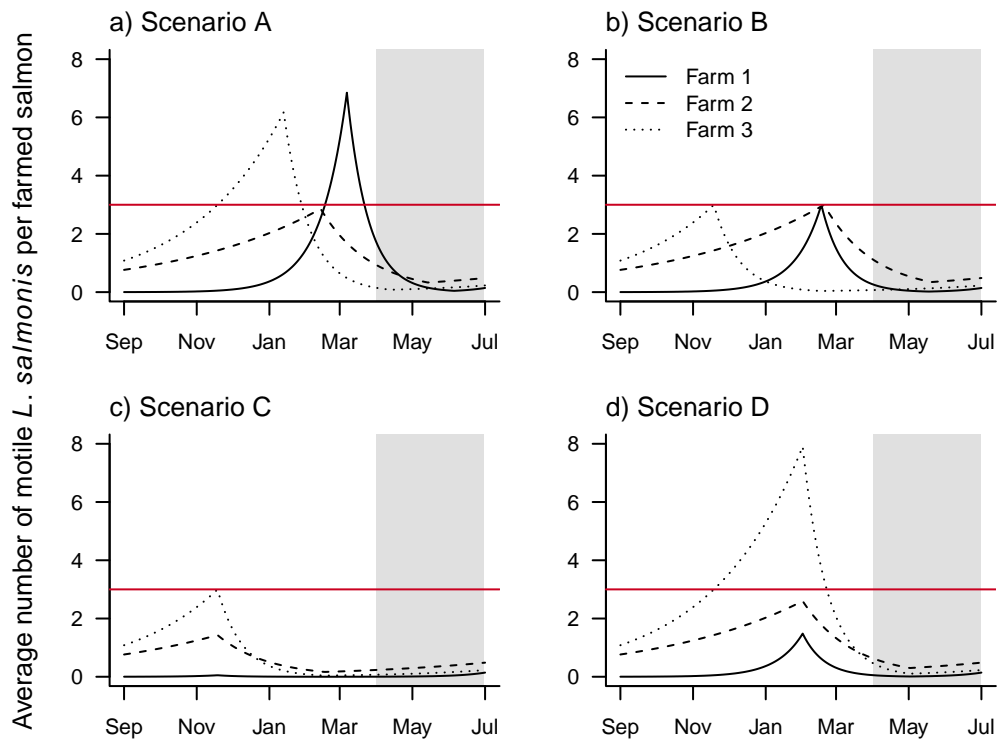


Figure 3: The average number of motile *L. salmonis* per farmed salmon on three salmon farms under four different treatment scenarios: (a) independent treatments on the observed date, (b) independent but immediate treatment when threshold is reached, (c) coordinated treatment of all farms when the first farm reaches the threshold, and (d) coordinated treatments of all three farms prior to the juvenile salmon migration. The horizontal (red) line indicates the treatment threshold of three motile lice per farmed salmon (Fisheries and Oceans Canada 2018). [Colour online.]

ing infestation (Figure 6). We estimated the migration path of salmon based on data of emergence timing from Glendale River, the major salmon-bearing river in the Broughton Archipelago located at approximately -40 km along our migration corridor (Figure 2). We obtained dates of emergence of pink salmon fry from Glendale River from 2007-2012 from the Glendale Creek Juvenile Downstream Program (Pieter vanWill, DFO, pers. comm.). In 2010, pink salmon fry seemed to emerge about two weeks earlier than in other years (Figure S1) and early migration timing was anecdotally reported to have exacerbated sea louse exposure of juvenile salmon in the 2015 outbreak in the BA (Bateman et al. 2016). To account for variability in migration timing, we considered an “early” migration scenario in which the population of juvenile salmon emerged as observed in 2010 and “normal” migration timing based on emergence data for 2007-2009 and 2011-2012 (Figure 6). Within each of those scenarios, we incorporated stochasticity in migration timing by resampling with replacement from the emergence dates of fry (Figure S1) 1000 times, yielding 1000 different migration paths through the spatiotemporal distribution of infectious sea lice. All migration paths starting at -40 km along the migration route (Glendale River; Figure 2). We assumed the migration speed of the juvenile salmon was constant at the speed estimated in the model fitting.

We summarized the effect of treatment timing on wild juvenile salmon using three metrics: (1) total infestation pressure encountered along the migration, (2) the maximum expected number of sea lice per juvenile salmon, and (3) the expected mortality of juvenile pink salmon due to infestation. For each migration path we simulated, we calculated the overall infestation pressure as the line integral of the migration path over the spatiotemporal distribution of infectious copepodids. The maximum number of lice per juvenile salmon was the maximum sum of copepodid, chalimus, and motile lice at any point along the migration path. We calculated the expected host mortality using previous estimates of

louse-induced mortality from Peacock et al. (2013). That study used a time series of the mean number of sea lice (copepodid, chalimus, and motile stage) at three locations in the Broughton that have been monitored since 2001 (Peacock et al. 2016) together with salmon spawner and recruitment data over 60 years to estimate the per-sea-louse mortality rate, c , for pink salmon populations. To estimate population-level mortality under our four treatment scenarios, we simulated the mean louse abundance at those same three monitoring locations on the migration route. We calculated the mortality of wild salmon due to sea lice per generation of salmon as $1 - e^{-cL}$ (Krkošek et al. 2011b) where $c = 0.190$ is the estimated louse-induced mortality from Peacock et al. (2013) and L is our simulated mean louse abundance. We report the 2.5%, 50%, and 97.5% quantiles of all three metrics from the 1000 migration paths for both early and normal migration timing. R (R Development Core Team 2018) code reproducing simulations is available online (see Data Accessibility section).

Results

Farm dynamics

The three salmon farms under study showed clear patterns of exponential growth of louse populations until treatment dates, after which louse populations declined (Figure 4). Growth rates and average lice per fish at time of treatment were in agreement with previous estimates (Krkošek et al. 2010, Table 1); slight differences were likely due to different assumptions about the statistical distribution of lice per fish. Krkošek et al. (2010) assumed lice were Poisson distributed, whereas we found lice were overdispersed on their hosts, and the negative binomial was a better fit to the data.

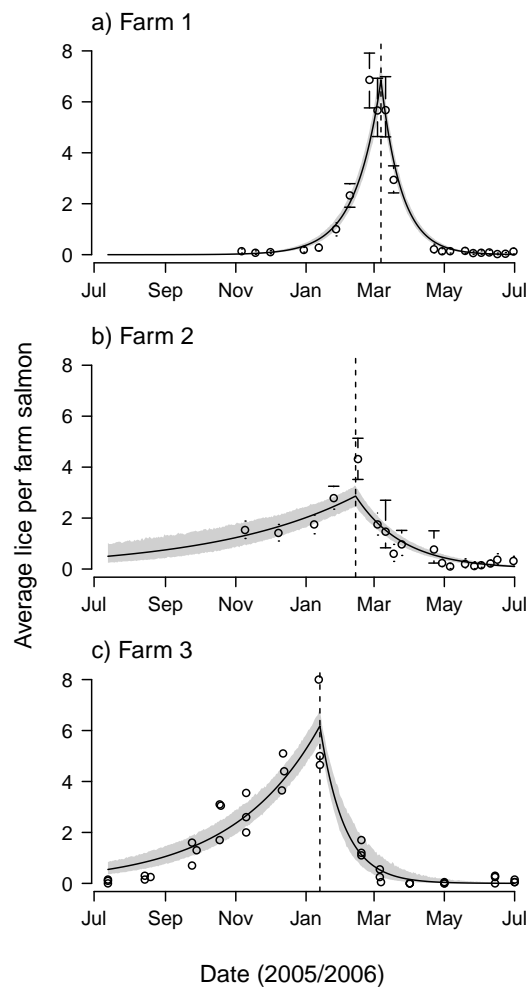


Figure 4: Growth and decay of sea-louse populations on three salmon farms before and after parasiticide treatments. (a) Farm 1 at $x = -3.7$ km, (b) Farm 2 at $x = 4.0$ km, and (c) Farm 3 at $x = 53.0$ km. Black lines are the model predictions for $f(t)$ from equation 1, with grey zones indicating the bootstrapped 95% confidence intervals on model predictions. Open points the average lice per farmed salmon \pm 95% bootstrapped confidence intervals. Vertical dashed lines indicate treatment dates. Note in (c) points are average lice per fish per pen, not counts of number of lice per fish. Corresponding parameter estimates are in Table 1.

Table 1: Parameter estimates (\pm 95% confidence intervals) for equation 1 fit to lice counts on farm salmon at three different salmon farms.

	Farm 1	Farm 2	Farm 3
r_1	0.045 (0.041, 0.048)	0.008 (0.005, 0.011)	0.013 (0.011, 0.016)
r_2	-0.056 (-0.06, -0.053)	-0.024 (-0.027, -0.022)	-0.048 (-0.062, -0.033)
f_0	6.848 (6.121, 7.574)	2.861 (2.444, 3.279)	6.177 (5.524, 6.83)
k or σ^*	1.547 (1.231, 1.863)	1.284 (0.986, 1.582)	0.638 (0.494, 0.781)

*For Farm 1 and Farm 2, k is the dispersion parameter of the negative binomial distribution, whereas for Farm 3, σ is the residual standard deviation of Gaussian errors.

Transmission to wild salmon

The production of sea-louse copepodites at salmon farms was around 2400 times greater (95% CI: 2311 - 2525) than background sources, assuming a farm footprint of 0.2 km (i.e., $\phi/(0.2 \times \kappa)$; Krkošek et al. 2005a). Dispersal and mortality resulted in relatively low densities of infectious sea lice along the migration route (i.e., $L_F(x, t) \ll 1$; Figure 6) and a maximum infestation pressure from farms that was 30 times greater than background sources ($\max_{x,t} \phi L_F(x, t)/\kappa$). This maximum occurred down-current from Farm 1 between 0.65 and 0.85 km along the migration on March 9, 2006, just two days after Farm 1 treated with EMB and prior to the treatment of Farm 2 (Figure 4). Overall, the infestation pressure from farms was greater than background sources for 12,500 km \times days, covering most of the migration corridor from January through May (Figure S6).

As well as identifying the sources of sea lice, parameter estimates give insight into both the life history of lice, survival rates, and differences in susceptibility of salmonid hosts to infestation (Table 2). We found that the background infestation pressure, $\kappa\beta_0$, was the same between pink and chum salmon, suggesting there was no difference in the susceptibility of those host species to initial infestation. Indeed, the data show similar numbers of copepodites on on both host species; however, the survival of both copepodite-stage and chalimus-stage sea lice was significantly higher on pink salmon than on chum salmon (Table

Table 2: Maximum likelihood parameter estimates (95% confidence intervals) for the model fit to lice counts on wild juvenile salmon from the Knight Inlet - Tribune Channel migration corridor of the Broughton Archipelago (Figure 2) in April and May, 2006.

Parameter	Description	Estimate (95% CI)
$\kappa\beta_0$	background infestation pressure	0.010 (0.009, 0.011)
$\phi\beta_0$	Farm infestation pressure	4.79 (4.41, 5.21)
v	Migration speed	4.09 (3.79, 4.40)
s_c^{pink}	Survival of copepodites on pinks	0.95 (0.70, 0.99)
s_h^{pink}	Survival of chalimus on pinks	0.46 (0.41, 0.52)
s_c^{chum}	Survival of copepodites on chums	0.78 (0.69, 0.85)
s_h^{chum}	Survival of chalimus on chums	0.29 (0.25, 0.33)
r	Overdispersion parameter for negative binomial	0.59 (0.54, 0.65)

2). Previous studies have assumed that juvenile salmon migrate at $v = 1 \text{ km} \cdot \text{day}^{-1}$ (Krkošek et al. 2005a, 2006a), but we found this parameter was much higher at $v = 4.09 (3.79 - 4.40) \text{ km} \cdot \text{day}^{-1}$.

Sea lice were overdispersed on their juvenile salmon hosts, distributed according to the negative binomial with an overdispersion parameter of $r = 0.59 (0.54 - 0.65)$, significantly less than one (Table 2). Previous studies using similar models assumed a Poisson distribution (Krkošek et al. 2005a, 2006a), but we found the negative binomial fit the data much better despite adding an extra parameter (Likelihood ratio test, $\chi^2 = 1249$, $df = 1$, $p < 0.001$).

Model predictions captured the main peaks in infestations of juvenile salmon, especially for chalimus-stage lice (Figure 5b). However, the model under-predicted sea-louse abundance on juvenile salmon towards the end of the migration route, particularly for motile-stage lice (Figure 5c). Data cloning showed that the parameters in both the farm model and the spatiotemporal infestation model for wild salmon were clearly estimable given the available data (Figure S2 & Figure S5).

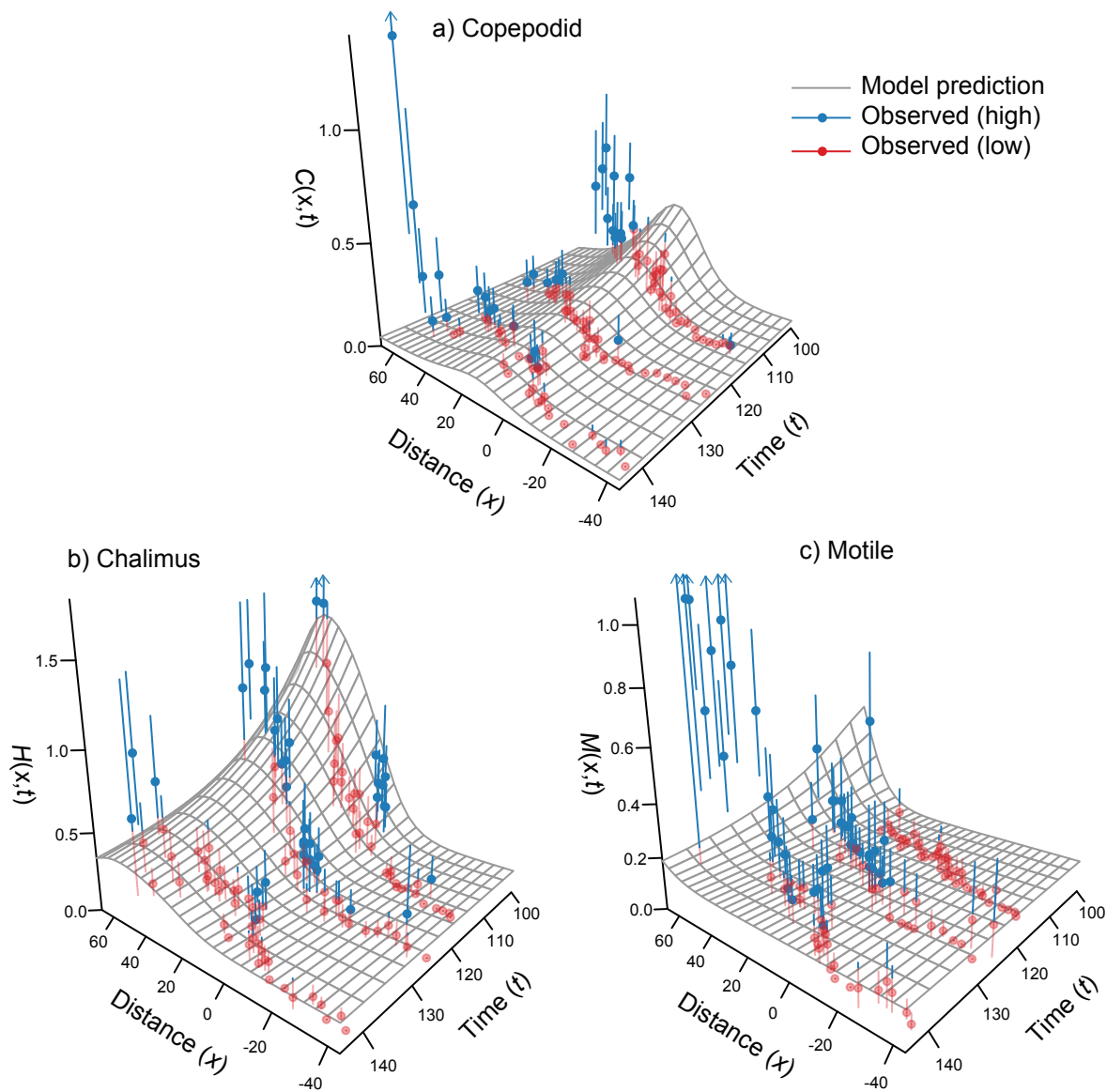


Figure 5: The abundance of (a) copepodid, (b) chalimus, and (c) motile sea lice per wild juvenile pink salmon along their migration corridor (x , km) from April 10, 2006 (day $t = 100$) to May 24, 2006 (day $t = 144$). The grey surface is the model prediction using fitted parameter estimates (Table 2). Points are the observed mean number of copepodid, chalimus, or motile sea lice on juvenile salmon ($\pm 95\%$ bootstrapped confidence intervals). Dark blue points indicate observations that are higher than model predictions and light red points indicate observations that are lower than model predictions. Arrows indicate 95% confidence intervals that extend beyond the plotting region. See Figure S7 in Online Supplement for fits to chum-salmon data. [25]our online.]

438 Simulations

439 The infestation pressure, maximum number of sea lice, and estimated mortality of juvenile
440 salmon due to sea lice were all lower for treatment scenarios that coordinated among farms
441 (i.e., scenarios C and D; Figure 7). For Scenarios A, B, and D, early migration timing
442 resulted in higher numbers of lice and higher mortality (Figure 7), as salmon migrated
443 closer in time to the peak infestation pressure at treatment (Figure 6). Mortality did not
444 increase with early migration timing for Scenario C where treatment was administered in
445 November, far in advance of the juvenile salmon migration. Even though our simulations
446 incorporated a recovery in sea-louse population growth rates on farms three months after
447 treatment, and louse populations in Scenario C had started to recover by the time the
448 juvenile salmon migrated past, lice did not reach high enough numbers on farms to result
449 in significant infestations of wild salmon (Figure 6).

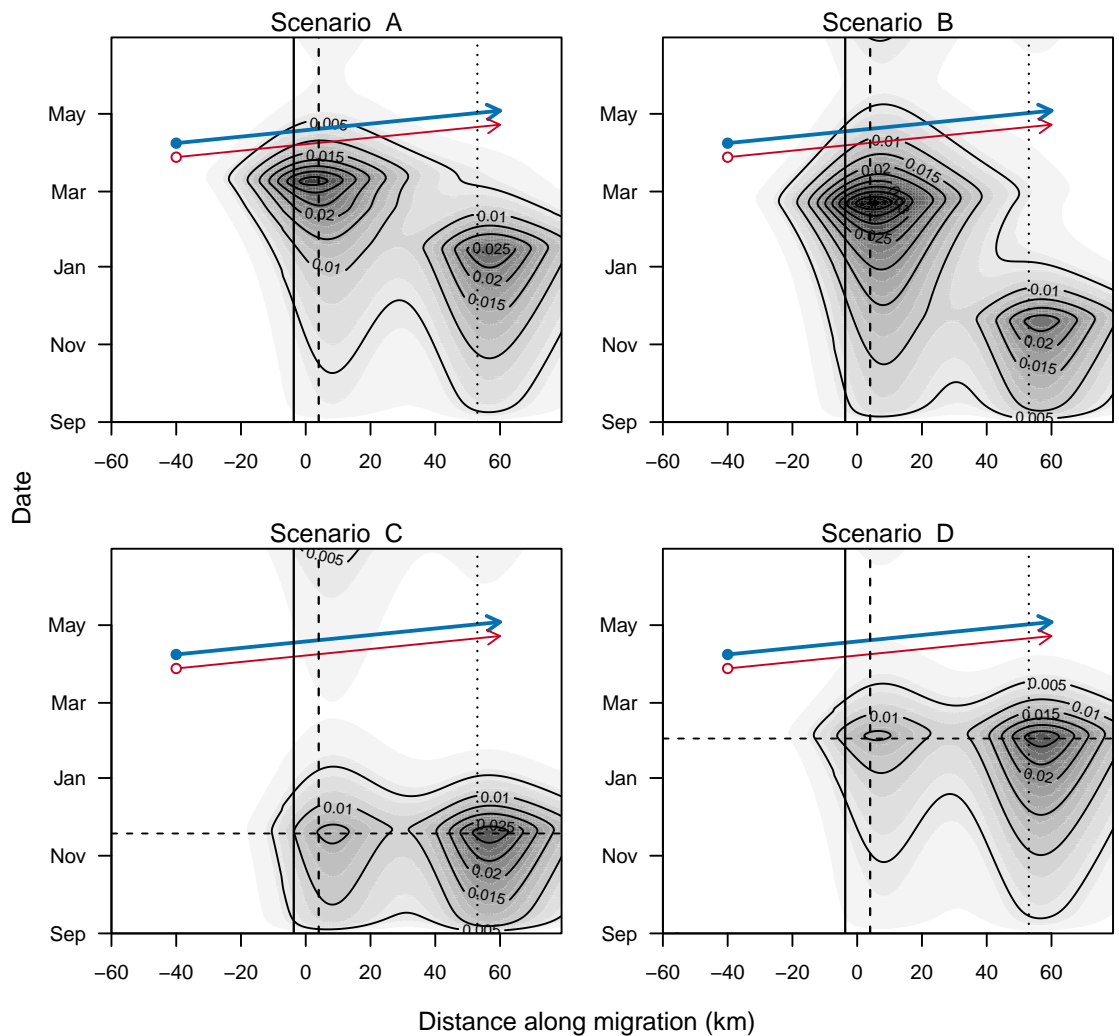


Figure 6: The simulated densities of infectious copepodites (darker = higher density) along the 1D migration corridor during 2005/2006 under four treatment scenarios. The x-axis is the Knight Inlet - Tribune Channel migration corridor (Figure 2) from -60 km to 80 km. The y-axis is time from September 1, 2005 to July 1, 2006. The locations of Farm 1, Farm 2, and Farm 3 (Figure 2) are indicated by vertical solid, dashed, and dotted lines, respectively. For scenarios C and D, the treatments on farms were coordinated, with the single treatment date indicated by the horizontal dashed line. The thick (blue) and thin (red) arrows show wild juvenile salmon migration routes under normal (closed blue point) and early (open red point) migration timing, respectively. When calculating metrics, we used 1000 such migration paths starting at different points in time to capture the uncertainty emergence time of juvenile salmon migrating from Glendale River. [Colour online.]

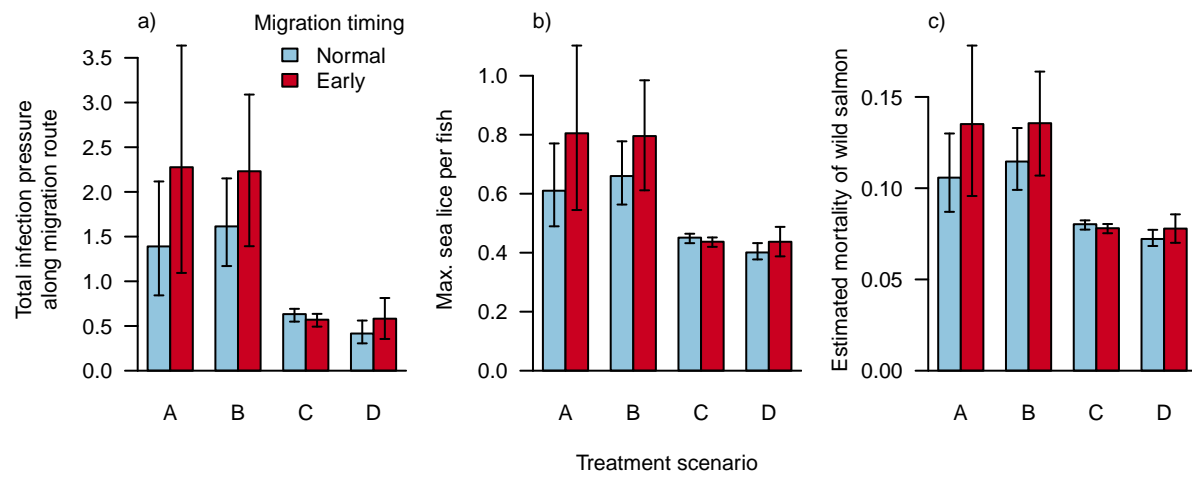


Figure 7: Three metrics of wild salmon health calculated over four treatment scenarios (A-D) under normal (light blue) and early (dark red) migration timing of juvenile salmon. Metrics are (a) total infestation pressure encountered by a juvenile salmon along their migration, calculated as the line integral over the distributions in Figure 6, (b) maximum number of lice (copepodite, chalimus, and motile stages) attached to a juvenile salmon during their migration, and (c) the mortality of wild salmon due to sea lice, calculated as $1 - e^{-cL}$, where c is the estimated per-sea-louse mortality rate for the generation of salmon (Peacock et al. 2013) and L is the mean sea-louse abundance on out-migrating juvenile salmon, calculated here from simulations. See main text for details. [Colour online.]

450 **Discussion**

451 The temporal dynamics of sea lice on salmon farms have been well studied in relation to
452 management and environmental variables (e.g., Revie et al. 2003, 2005; Rogers et al. 2013),
453 but there have been few attempts to connect those dynamics with empirical data on wild
454 juvenile salmon migrating through those areas using mechanistic models. In this study,
455 we empirically modelled the temporal dynamics of sea-louse populations on farmed salmon
456 and the consequences for the infestation and development of sea lice on juvenile wild pink
457 and chum salmon in the Broughton Archipelago, which has been a major salmon farming
458 region in Pacific Canada. Since 2006, salmon farms in the Broughton Archipelago have been
459 more proactive in treating farmed salmon for lice in order to ensure low prevalence during
460 the juvenile salmon out-migration (Peacock et al. 2013), but recent outbreaks highlight
461 the challenges to successful proactive treatments given variability in sea-louse population
462 growth rates and in the migration timing of juvenile salmon (Bateman et al. 2016). Our
463 fitted model suggested that, in 2006, the infestation pressure from salmon farms was 24
464 times greater than background sources of sea lice, such as returning wild adult salmon.
465 Although this result is not new - many studies have identified salmon farms as the main
466 source of sea lice on wild juvenile salmon in regions of salmon farming (e.g., Krkošek et al.
467 2005a, 2006b; Marty et al. 2010)) - there are several aspects of this study that advance our
468 understanding of the system in a significant way.

469 Previous models describing observations of sea lice on juvenile salmon assumed that sea-
470 louse production at salmon farms was constant (Krkošek et al. 2005a, 2006a). Sea-louse
471 populations on three salmon farms that were active in 2005-2006 in the Broughton Archipelago
472 showed clear patterns of exponential growth and decay, with louse abundances peaking just
473 prior to treatments of farmed salmon. The temporal dynamics of louse populations on these

474 farms in 2006 suggest that treatments were effective at reducing sea-louse abundances on
 475 farmed salmon; however the timing of the treatment could have been more precautionary
 476 to prevent peaks in louse numbers as juvenile salmon migrate past farms (Rogers et al.
 477 2013). For example, the peak in louse production occurred on March 9, 2006, close to when
 478 juvenile salmon began their migration. The two farms that were treated in January were
 479 minor sources of lice at that time. Including the temporal dynamics of sea-louse infestation
 480 pressure from farms allowed us to more accurately assess the importance of salmon farms
 481 as sources of infestation for wild salmon, but also to investigate other potential treatment
 482 strategies using a simulation approach.

483 Our estimates of louse-induced mortality from the true treatment schedule were on par with
 484 previous estimates of 8.3-22.3% (mean 15.9%) from an independent data set (c.f. Figure 6
 485 of Peacock et al. (2013)). Using our parameterized model, we were able to explore how that
 486 mortality might have changed if farms had treated differently. We found that coordinated
 487 treatment among salmon farms and early precautionary treatments would have reduced sea-
 488 louse abundance and minimized sea-louse-induced mortality of wild salmon. Our results
 489 likely underestimated the importance of coordinated treatments, as we did not account
 490 for transmission and infestation among salmon farms that could hinder area-wide control
 491 if management is not coordinated (Bateman et al. 2016). Early treatment was especially
 492 important when the migration timing of juvenile salmon was earlier than usual. Emergence
 493 of juvenile salmon has been shown to advance with warmer temperatures (Holtby et al.
 494 1989), and thus earlier migration timing may become the norm under climate change. In
 495 Pacific Canada, current license conditions require that salmon farms take management
 496 action when louse populations exceed an average of three motile lice per farmed salmon
 497 (Fisheries and Oceans Canada 2018), but there is no formal area-wide management plan
 498 that requires coordinated treatment among farms. Early treatment would seem to be

499 optimal after the influx of lice with returning wild salmon in the fall, and only if farms
500 coordinate to avoid reinfestation prior to the juvenile salmon migration. This coordination
501 is particularly important to minimize infestation of juvenile salmon at the start of their
502 migration, so that the migrating juvenile salmon themselves don't spread sea lice among
503 farms (Krkošek et al. 2006a). These benefits of early intervention would have to be weighed
504 against drawbacks of potentially more frequent treatments throughout a production cycle,
505 such as cost and increased opportunity for sea lice to develop resistance to treatments.

506 Although we have addressed several shortcomings of previous modelling efforts, there re-
507 main assumptions and caveats to our model. The estimability of parameters in previous
508 models of sea-louse transmission (Krkošek et al. 2005a, 2006a) was not clear from published
509 point estimates. The original models attempted to estimate 14 free parameters, and we
510 found many of these parameters were not estimable given the available data. Using data
511 cloning (Lele et al. 2010), we identified which parameters were estimable and simplified
512 the model to 7 free parameters. This involved making some assumptions and fixing some
513 parameters (e.g., the developmental times of louse stages) and assuming other parameters
514 did not differ between pinks and chum salmon (e.g., the background infestation pressure).
515 Although these added assumptions may weaken the generality of the model, the simplified
516 model was a better fit to the data. In this particular case, the main conclusion that salmon
517 farms are a main source of sea lice on juvenile salmon was unchanged when the model was
518 improved to ensure estimability. This may not always be the case, and we highlight that
519 considering parameter estimability is important when fitting mechanistic models to data
520 (Peacock et al. 2017), especially if the results have implications for the conservation and
521 management of endangered or keystone species.

522 Our model greatly simplifies the processes driving the dispersal of sea lice from salmon
523 farms. More complex hydrodynamic models have considered ocean currents and sea-louse

524 dispersal in three-dimensions (Foreman et al. 2009; Stucchi et al. 2011; Cantrell et al. 2018),
525 which undoubtedly yields more precise predictions of infestation pressure but is dependent
526 on the wind, tide, and freshwater forcing of the specific time period being modelled. Our
527 approach, although simpler, yields more generalizable insights into infestation patterns
528 and has been shown to capture the major spatial patterns in infestation (Krkošek et al.
529 2005a, 2006a). Further, we were able to fit the model to data and infer parameters. One
530 exciting area for future research would be to bridge these two approaches and confront the
531 more complex, simulation models with spatiotemporal data of infestations on wild salmon,
532 although there are significant computational challenges involved.

533 Our model appears to do a poor job of predicting lice on juvenile salmon towards the
534 end of their migration route, failing to capture rises in all lice stages, but particularly in
535 motiles, in the 40-60 km range (Figure 5). The more complex, hydrodynamic models of
536 the Broughton Archipelago have suggested that salmon farms north of Farm 3 (Figure 2),
537 which were not modelled in this study, may be major contributors of infestation pressure
538 in the area of Farm 3 (Cantrell et al. 2018). Sea lice emitted from these more northern
539 farms may be driving the higher-than-predicted infestations of wild salmon later in the
540 migration.

541 Several other simplifying assumptions in our model may explain the discrepancy towards
542 the end of the migration. For example, it is possible that sea lice develop to sexual maturity
543 on juvenile salmon and this may act as a third source of infestation pressure that we
544 did not account for. The generation time of *L. salmonis* is 4-8 weeks (depending on
545 temperature) (Costello 2006) and juvenile salmon may take up to three months to complete
546 their migrations (Krkošek et al. 2009). Reproduction of sea lice on wild salmon has been
547 accounted for before (Krkošek et al. 2006a), but given the complexity of this model already,
548 we could not include it. Once again, a simplification had to be made to ensure that the

549 model did not outstretch the data.

550 Our assumption of a constant migration speed along the course of the juvenile salmon
551 migration may also have contributed to the poor fit of the model towards the end of
552 the migration. If salmon slowed their migration around certain farms, the infestation
553 pressure may be effectively higher because of increased exposure time. Previous estimates
554 for the migration speed of juvenile salmon were $\sim 1 \text{ km} \cdot \text{day}^{-1}$ (Krkošek et al. 2006a;
555 Morton et al. 2010), but our results suggest juvenile salmon migrate four times as fast.
556 There are several reasons to believe that migration speed is not constant. In the simplest
557 case, migration speed may increase as the salmon grow from $\sim 20 \text{ mm}$ at ocean entry to
558 over 100 mm towards the end of our study period. Acoustic tagging studies have shown
559 considerable variability in the migration speed of juvenile salmon. For example, Welch
560 et al. (2011) report a standard deviation in migration speed of juvenile coho of ± 4.93
561 bodylengths per second, or roughly $10\text{-}40 \text{ km d}^{-1}$ for fish between 30 mm and 100 mm .
562 More complicated dynamics may arise if migration speed depends on sea-louse infestation
563 intensity; juvenile salmon may slow their migration as they acquire more lice (Nendick
564 et al. 2011; Brauner et al. 2012). In such a scenario, juvenile salmon may get caught
565 in infestation hotspots (Altizer et al. 2011; Peacock et al. 2018), and infestation pressure
566 in those areas may increase out of proportion to the densities of sea-louse larvae. There
567 have been no direct studies of the routes and timing of migration for juvenile salmon in
568 the Broughton Archipelago to assess support for these hypotheses, but such biological
569 complexity may mean that our assumption of constant migration speed misses key aspects
570 of the host-parasite dynamics.

571 We assumed that all parameters, not just migration speed, were constant over the period
572 of juvenile salmon outmigration. It is well established that many of these parameters can
573 vary - in particular development times and survival rates of sea lice are known to depend

on temperature and salinity, respectively (e.g., Figure S8). Development is slower at colder temperatures (Stien et al. 2005; Costello 2006), and temperature was generally increasing throughout our study period (Figure S8a). However, in the spatial dynamics of the model, the development times of attached lice are multiplied by the migration speed of juvenile salmon (equations 4), and migration speed may also have been increasing throughout the study as fish grow. Thus, faster sea-louse development associated with higher temperatures in combination with faster migration speeds of hosts may lead to the same overall spatial dynamics of infestation. However, the development of pre-infective-stage larvae is predicted to have ranged from 2.6 to 5.3 days given the range in temperature observed during our sampling (Figure S8a), and this may have introduced errors into our calculation of the farm footprint and infestation pressure on wild salmon. Although it must be acknowledged, the uncertainty introduced by fluctuating temperatures within the range that we observed is not substantially greater than the uncertainty in the parameter estimates themselves (Stien et al. 2005).

Salinity is also known to affect sea-louse population dynamics with the survival of attached lice declining below 15 psu (Johnson and Albright 1991; Connors et al. 2008) and pre-infective stages being even more sensitive (Bricknell et al. 2006; Groner et al. 2016). During this study, only three out of 128 sampling events for juvenile salmon had salinities below 15 psu, and two of these occurred at adjacent sites on the same day and may be considered an anomaly (Figure S8b). Although salinities were frequently below 29 psu throughout our study, the concentration below which attachment success has been shown to decline (Bricknell et al. 2006), there were no obvious spatial or temporal patterns in salinity in our data (Figure S8b). Thus, the impact of salinity on attachment success would be averaged and included in the transmission coefficient (β_0) within the farm and background infestation pressures. Including relationships between temperature and devel-

599 opmental times or survival and salinity, such as those described by Groner et al. (2016), may
600 increase the realism of our model but are not likely to have changed the main results.

601 Some of our parameter estimates were somewhat surprising in light of laboratory studies
602 and previous models of sea-louse transmission in the Broughton. We found that sea lice
603 apparently survive better on pink salmon than on chum salmon, in contrast to laboratory
604 studies suggesting that pink salmon mount a more effective immune response (Jones et al.
605 2007; Sutherland et al. 2014). In the field, some studies have found higher apparent survival
606 on chum salmon (Morton et al. 2010), whereas others are inconsistent in which host species
607 sea lice survive better on a (Krkošek et al. 2006a). This uncertainty highlights the differences
608 between laboratory and field studies, and the need to consider ecological effects of sea-louse
609 infestation (e.g., Krkošek et al. 2011a), as well as physiological (Brauner et al. 2012). For
610 example, predation by coho salmon on both pink and chum salmon may alter the host-
611 parasite dynamics in the natural environment (Peacock et al. 2014).

612 The importance of salmon farms in driving sea-louse infestations of wild juvenile salmon
613 has by now been well established (e.g., Bjørn et al. 2001; Krkošek et al. 2006a; Marty
614 et al. 2010). The more critical problem from a conservation standpoint is determining
615 best practices for salmon farms in order to minimize impacts to wild salmon populations.
616 By parameterizing a spatiotemporal model that connects management actions on farms to
617 infestations of migrating wild salmon, we were able to explore different management sce-
618 narios in an empirically grounded framework. Our simulations suggest that precautionary
619 treatments in advance of the juvenile salmon migration that reduce sea-louse populations
620 on salmon farms before they grow to critical levels and siting salmon farms as far along
621 the migration routes as possible will minimize the impact of farm-origin sea lice on wild
622 pink and chum salmon. However, given recent reports of sea-louse resistance to chemical
623 treatments in the Pacific (Messmer et al. 2018), any management strategy will also have to

624 consider the evolutionary consequences for resistance to ensure long-term viability.

625 More generally, this study demonstrates the importance of considering spatial and tempo-
626 ral patterns in infection dynamics when attempting to understand and manage emerging
627 infectious diseases in wildlife populations. Mechanistic models are a powerful tool for un-
628 derstanding and predicting complex ecological and epidemiological processes (White et al.
629 2018), but there are few examples where such models are fit to data to estimate param-
630 eters of the system. The development of new analytical and statistical approaches, such
631 as the data cloning (Lele et al. 2010, 2007) we applied, have opened the door to fitting
632 mechanistic models to complex ecological datasets, with the potential to deepen our un-
633 derstanding of even well-studied systems such as sea-louse transmission between farmed
634 and wild salmon.

635 Data accessibility

636 All code and data for this study can be accessed at [https://github.com/sjpeacock/Spatiotemporal-](https://github.com/sjpeacock/Spatiotemporal-infection-model/)
637 [infection-model/](https://github.com/sjpeacock/Spatiotemporal-infection-model/).

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Figure captions

- 1 Schematic of the modelling framework for the sea-louse transmission model incorporating temporal dynamics of lice on salmon farms. The model was fit to data in two stages; (1) the farm dynamics were fit to counts of lice on farmed salmon (red), and (2) the final predictions of lice per wild juvenile salmon were confronted with data of sea-louse abundances on wild salmon throughout their migration route (blue). [Colour online.] 9
- 2 Models were fit to data from the Broughton Archipelago, on the south central coast of British Columbia, Canada. Sampling of farmed salmon took place on the three active salmon farms in the study region in 2006 (red squares). Approximately 100 juvenile wild pink and chum salmon were sampled every ~1 km (black stars) along their migration past these farms along the Knight Inlet - Tribune Channel migration corridor. The model was simulated from -60 km to 80 km along the migration corridor, with 0 km being a reference point at the confluence of Knight Inlet and Tribune Channel. The map was produced using the R package PBSmapping (Schnute et al. 2018) with shoreline data from Wessel and Smith (2016). [Colour online.] 12
- 3 The average number of motile *L. salmonis* per farmed salmon on three salmon farms under four different treatment scenarios: (a) independent treatments on the observed date, (b) independent but immediate treatment when threshold is reached, (c) coordinated treatment of all farms when the first farm reaches the threshold, and (d) coordinated treatments of all three farms prior to the juvenile salmon migration. The horizontal (red) line indicates the treatment threshold of three motile lice per farmed salmon (Fisheries and Oceans Canada 2018). [Colour online.] 19

917	4	Growth and decay of sea-louse populations on three salmon farms before	
918		and after parasiticide treatments. (a) Farm 1 at $x = -3.7$ km, (b) Farm	
919		2 at $x = 4.0$ km, and (c) Farm 3 at $x = 53.0$ km. Black lines are the	
920		model predictions for $f(t)$ from equation 1, with grey zones indicating the	
921		bootstrapped 95% confidence intervals on model predictions. Open points	
922		the average lice per farmed salmon \pm 95% bootstrapped confidence intervals.	
923		Vertical dashed lines indicate treatment dates. Note in (c) points are average	
924		lice per fish per pen, not counts of number of lice per fish. Corresponding	
925		parameter estimates are in Table 1.	22
926	5	The abundance of (a) copepodid, (b) chalimus, and (c) motile sea lice per	
927		wild juvenile pink salmon along their migration corridor (x , km) from April	
928		10, 2006 (day $t = 100$) to May 24, 2006 (day $t = 144$). The grey surface	
929		is the model prediction using fitted parameter estimates (Table 2). Points	
930		are the observed mean number of copepodid, chalimus, or motile sea lice	
931		on juvenile salmon (\pm 95% bootstrapped confidence intervals). Dark blue	
932		points indicate observations that are higher than model predictions and	
933		light red points indicate observations that are lower than model predictions.	
934		Arrows indicate 95% confidence intervals that extend beyond the plotting	
935		region. See Figure S7 in Online Supplement for fits to chum-salmon data.	
936		[Colour online.]	25

- 937 6 The simulated densities of infectious copepodites (darker = higher density)
 938 along the 1D migration corridor during 2005/2006 under four treatment sce-
 939 narios. The x-axis is the Knight Inlet - Tribune Channel migration corridor
 940 (Figure 2) from -60 km to 80 km. The y-axis is time from September 1,
 941 2005 to July 1, 2006. The locations of Farm 1, Farm 2, and Farm 3 (Figure
 942 2) are indicated by vertical solid, dashed, and dotted lines, respectively. For
 943 scenarios C and D, the treatments on farms were coordinated, with the sin-
 944 gle treatment date indicated by the horizontal dashed line. The thick (blue)
 945 and thin (red) arrows show wild juvenile salmon migration routes under un-
 946 der normal (closed blue point) and early (open red point) migration timing,
 947 respectively. When calculating metrics, we used 1000 such migration paths
 948 starting at different points in time to capture the uncertainty emergence
 949 time of juvenile salmon migrating from Glendale River. [Colour online.] . . . 27
- 950 7 Three metrics of wild salmon health calculated over four treatment scenarios
 951 (A-D) under normal (light blue) and early (dark red) migration timing of
 952 juvenile salmon. Metrics are (a) total infestation pressure encountered by a
 953 juvenile salmon along their migration, calculated as the line integral over the
 954 distributions in Figure 6, (b) maximum number of lice (copepodite, chalimus,
 955 and motile stages) attached to a juvenile salmon during their migration,
 956 and (c) the mortality of wild salmon due to sea lice, calculated as $1 - e^{-cL}$,
 957 where c is the estimated per-sea-lice mortality rate for the generation of
 958 salmon (Peacock et al. 2013) and L is the mean sea-lice abundance on
 959 out-migrating juvenile salmon, calculated here from simulations. See main
 960 text for details. [Colour online.] 28

Appendix: Estimation of spatiotemporal transmission dynamics
and analysis of management scenarios for sea lice of farmed and wild
salmon

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22 Appendix

23 Advection, diffusion, decay of farm-source sea lice

Sea-louse nauplii from salmon farms advect, diffuse, die and develop according to equation 2. The transient solution to the advection, diffusion, decay equation can be written in the form of a Green's function:

$$(A1) \quad G(x, t) = \frac{1}{\sqrt{4\pi Dt}} \exp \left[-(\mu_n + \theta)t - \frac{(x - \gamma t)^2}{4Dt} \right]$$

24 To account for the dynamic production of lice at farms, we calculated the convolution of
 25 equation A1 with the point source forcing function (equation 1) multiplied by the Dirac
 26 delta function, $\delta(y)$, for a farm located at y . This assumes that lice are produced at exactly
 27 the farm location, with production proportional to the number of motile adult lice at that
 28 farm, and then disperse according to the advection-diffusion-decay process. The resulting
 29 distribution kernel of nauplii in space and time is,

$$(A2) \quad \begin{aligned} k_n(x, t) &= \int_0^t \int_{-\infty}^{\infty} G(x - \xi, t - \tau) f(\tau) \delta(\xi) d\xi d\tau \\ &= \int_0^t G(x, t - \tau) f(\tau) d\tau \\ &= \int_0^t \frac{1}{\sqrt{4\pi D(t - \tau)}} \exp \left[-(\mu_n + \theta)(t - \tau) - \frac{[x - \gamma(t - \tau)]^2}{4D(t - \tau)} \right] f(\tau) d\tau \end{aligned}$$

30 These distributed nauplii then develop into infectious copepodites, which can attach to
 31 susceptible juvenile salmonids in the vicinity. To obtain the distribution of infectious
 32 copepodites, we consider the nauplii as a distributed source that develop and subsequently

diffuse and advect, leading to a second convolution for the distribution of infectious copepodites,

$$(A3) \quad k_c(x, t) = \int_0^t \int_{-\infty}^{\infty} G(x - \xi, t - \tau) k_n(\xi, \tau) d\xi d\tau$$

where the parameters in $G(x, t)$ are the same as before except μ_c , the mortality rate of copepodites, replaces $(\mu_n + \theta)$. In practice, we calculated the solutions given by equations (A3) and (A3) numerically, applying a fast Fourier transform (FFT) algorithm to ease the convolution step. We assume that the probability of successfully finding a host is low and so we ignore removal of planktonic copepodites through attachment. This solution, $k_c(x, t)$ is proportional to the infestation pressure on wild juvenile salmon from copepodites originating from salmon farms. We did not know the transmission coefficient (i.e., the proportion of infectious copepodites that attach to a host) and so we considered only relative densities of copepodites, and normalized $k_c(x, t)$ so the area under the highest peak of $k_c(x, t)$ in time equalled one. We also assumed the distributions originating from different farms were independent, and therefore the total farm-source copepodite density, $L(x, t)$ was the normalized sum of the copepodites originating from each of the three farms,

$$(A4) \quad L_1(x, t) = \frac{\sum_{i=1}^3 k_{c,i}(x, t)}{\max_{i,j} \left[\int_{-\infty}^{\infty} k_{c,i}(u, j) du \right]}$$

48 where $k_{c,i}(x, t)$ is the copepodite density from farm i .

49 infestation model

We model infection as a Gamma-Poisson process, where the transmission coefficient β is a Gamma random variable with mean β_0 and dispersion parameter r . The expected number of copepodites per juvenile salmon $C(x, t)$ is therefore also a Gamma random variable with mean $\beta_0 L_0$ and dispersion parameter r :

$$(A5) \quad g(C; \beta_0 L_0, k) = C^{k-1} \left[\frac{k}{\beta_0 L_0} \right]^k \frac{\exp \left[- \left(\frac{k}{\beta_0 L_0} \right) C \right]}{\Gamma(k)}$$

where $L_0 = \int_0^{\tau_c} L(x - vu, t - u) du$ (equation 5a). The number of copepodid sea lice on an individual fish, N_c will be a random variable with probability density:

$$(A6) \quad Pr\{N_c = c\} = \int_0^\infty g(C; \beta_0 L_0, k) \frac{C^c}{c!} e^{-C} dC$$

$$(A7) \quad = \frac{(c+k-1)!}{c!(k-1)!} \left(\frac{k}{\beta_0 L_0 + k} \right)^k \left(1 - \frac{k}{\beta_0 L_0 + k} \right)^c$$

which is the negative binomial distribution with mean $C(x, t)$ and dispersion parameter r . A count of h chalimus sea lice on an individual fish can result from any of i attached copepodites surviving to the chalimus stage with probability s_c . We define N_i as the discrete random variable for the number of attached copes available for recruitment into the chal stage on an individual fish at point (x, t) . Therefore the probability of having a fish with h chalimus sea lice is:

$$(A8) \quad Pr\{N_h = h\} = \sum_{i=h}^{\infty} \left[\binom{i}{h} s_c^h (1 - s_c)^{i-h} Pr\{N_i = i\} \right]$$

The distribution for N_i is:

$$(A9) \quad Pr\{N_i = i\} = \int_0^\infty g(I; \beta_0 L_1, k) \frac{I^i}{i!} e^{-I} dI$$

$$(A10) \quad = \frac{(i+k-1)!}{i!(k-1)!} \left(\frac{k}{\beta_0 L_1 + k} \right)^k \left(1 - \frac{k}{\beta_0 L_1 + k} \right)^i$$

yielding the probability of observing h chalimus sea lice as

$$\begin{aligned} Pr\{N_h = h\} &= \sum_{i=h}^{\infty} \left[\frac{i!}{h!(i-h)!} s_c^h (1-s_c)^{i-h} \int_0^\infty g(I; \beta_0 L_1, k) \frac{I^i}{i!} e^{-I} dI \right] \\ &= \binom{h+k-1}{h} \left(\frac{k}{k+s_c \beta_0 L_1} \right)^k \left(1 - \frac{k}{k+s_c \beta_0 L_1} \right)^h \end{aligned}$$

which is the negative binomial distribution with mean value $s_c \beta_0 L_1$ and dispersion parameter r , where $L_1 = \int_{\tau_c}^{\tau_c + \tau_h} L(x - vu, t - u) du$. Similar logic can be followed to arrive at the distribution for the number of motile lice per fish, giving formulas for the probabilities of observing numbers of copepodite, chalimus and motile sea lice on individual fish. We assume that sea-louse infestations on an individual fish are independent and that observations of sea lice on different fish at the same sample site are independent. Therefore, the likelihood of the data given a certain parameter set θ is

$$(A11) \quad \mathcal{L}(\text{data} | \theta) = \prod_{i=1}^N Pr\{N_c(x_i, t_i) = c_i\} \cdot Pr\{N_h(x_i, t_i) = h_i\} \cdot Pr\{N_m(x_i, t_i) = m_i\}$$

50 where N is the total number of **fish** sampled, and c_i , h_i and m_i are the observed number
 51 of cope, chal and mot lice on fish i , and x_i and t_i are the place and time that fish i was
 52 sampled.

Table A1: List of parameters and variables referred to throughout the text.

Type*	Symbol	Description	Equation [†]	Value and/or units
Variable	$f(t)$	The expected number of sea lice per farmed salmon at time t	1	motiles · fish ⁻¹
Parameter	f_0	The average number of motile sea lice per farmed salmon at the time of treatment	1	motiles · fish ⁻¹
Parameter	r_1, r_2	Rates of exponential growth and decay of lice on farmed salmon	1	day ⁻¹
Parameter	k, σ	Dispersion parameter (k , for SP and HR farms) or residual standard deviation (σ , for BG farm) for modelling lice dynamics on farms	Table 1	
Variable	$n(x, t)$	The density of nauplii at km x and day t	2	
Constant	D	Diffusion coefficient	2	22.67 km ² · day ⁻¹
Constant	γ	Advection coefficient	2	1.56 km · day ⁻¹
Constant	μ_n	Mortality rate of nauplii	2	1/5 day ⁻¹
Constant	μ_c	Mortality rate of copepodites	A3	4/5 day ⁻¹
Variable	$L_1(x, t)$	Distribution of infectious copepodites from farm sources	A4 & 3	
Parameter	κ	Scale of background infestation pressure	3	copepodites · fish ⁻¹
Parameter	ϕ	Scale of farm infestation pressure	3	
Variable	$L(x, t)$	Distribution of infectious copepodites from all sources	3	
Variable	$C(x, t)$	Expected number of copepodites per wild juvenile salmon	4	
Constant	β	Transmission coefficient	4	1
Constant	τ_c	Developmental time of the copepodid stage	4	3.6 days
Parameter	v	Migration speed of juvenile salmon	4	km · day ⁻¹
Variable	$H(x, t)$	Expected number of chalimus per wild juvenile salmon	4	
Parameter	s_c	Survival of copepodites to chalimus stage	4	
Constant	τ_h	Developmental time of the chalimus stage	4	16.0 days
Variable	$M(x, t)$	Expected number of motiles per wild juvenile salmon	4	
Parameter	s_h	Survival of chachalimuslimi to motile stage	4	
Constant	τ_m	Developmental time of the motile stage	4	28.6 days
Parameter	r	Dispersion parameter in negative binomial distribution	A5	

*Type distinguishes variables from parameters, and specifies which parameters are constant (i.e., not free). [†]Equation is the equation in which the parameter or variable first appears.