## Analysis of Rain Behaviour in Australia

Statistical Learning's Project

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### **Obtaining Data**

■ The weatherAUS dataset contains over 140,000 daily weather observations from 49 Australian weather stations collected in about 10 years. All the information has been collected by the Australian Bureau Of Meteorology.



Figure: Bureau of Meteorology, Copyright Commonwealth of Australia 2010.

- The specific dataset used in this project has been downloaded from Kaggle at the link: https://www.kaggle.com/jsphyg/weather-dataset-rattle-package.
- It is also available in a slightly different version via the R package rattle.data and at the link: https://rattle.togaware.com/weatherAUS.csv.



#### Characteristics of the Dataset

The dataset in organized in *blocks*, where each block contains the recordings of a given weather station. These blocks are ordered by the location of the station, and each block is in chronological order. Each row has some climatic data regarding a specific day and location. To be precise, two variables regards the rainfall of the following day: they will be the response variable (one for regression, one for classification). All other variables are relative to that day.

- Dimensions: 142193 observations and 24 variables;
- No duplicate observation;
- 56420 observations with at least one NA value.

Four variables (Evaporation, Sunshine, Cloud9am, Cloud3pm) has a high percentage of NAs (between 35 and 50%). The remaining variables have less than 10% of NAs.

Explanatory Variables, X

### OII LO



- Date: The date of observation (a date object).
- Location: The common name of the location of the weather station
- MinTemp: The minimum temperature in degrees centigrade
- MaxTemp: The maximum temperature in degrees centigrade
- Temp9am (Temp3pm): Temperature (degrees C) at 9 a.m.(3 p.m.)
- Evaporation: Class A pan evaporation (in millimeters) during 24 h
- Sunshine: The number of hours of bright sunshine in the day
- Rainfall: The amount of rainfall recorded for the day in millimeters.
- quarters: categorical variable that encodes the division of the year in to 4 groups of consecutive months
- Coast: 0-1 binary variable that tells whether a city is on the coast [1] (in a range of 50km from the sea) or not [0]

- WindGustDir: The direction of the strongest wind gust in the 24 h to midnight
- WindGustSpeed: The speed (in kilometers per hour) of the strongest wind gust in the 24 h to midnight
- WindDir9am (WindDir3pm): The direction of the wind gust at 9 a.m. (3p.m.)
- WindSpeed9am (WindSpeed3pm): Wind speed (in kilometers per hour) averaged over 10 min before 9 a.m. (3 p.m.)
- Humidity9am (Humidity3pm): Relative humidity (in percent) at 9 am (3 p.m.)
- Pressure9am (Pressure3pm): Atmospheric pressure (hpa) reduced to mean sea level at 9 a.m. (3 p.m.)
- Cloud9am (Cloud3pm): Fraction of sky obscured by cloud at 9 a.m.(3 p.m.). This is measured in "oktas," which are a unit of eighths.



Response(s),  $\mathbb{Y}$ 

### Rainfall variables

- Rainfall: amount of rainfall recorded for the day in mm;
- RainToday: "Yes" if precipitation in the 24 hours to 9am exceeds 1mm, otherwise "No";
- RISK\_MM: same as Rainfall, shifted by one day (in the future);
- RainTomorrow: same as RainToday, shifted by one day.
- Check correctness of these variables
- Restore some missing values.



Missing Values - NAs

**Missing data** are also referred to as *non-response* or *unobserved* data, and occur in most types of studies. Missing values can occur due to:

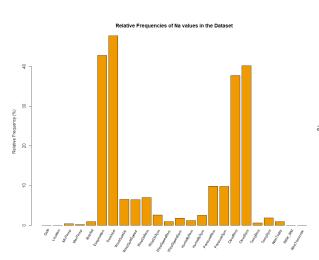
- Failure of measurement
- Data loss
- Out-of-range data and data loading issues
- Units fail to answer all questions
- Loss of follow-up or other plausible reasons.

David J. Hand

<sup>&</sup>quot;We should be suspicious of any dataset (large or small) which appears perfect."



Missing Values' Pattern

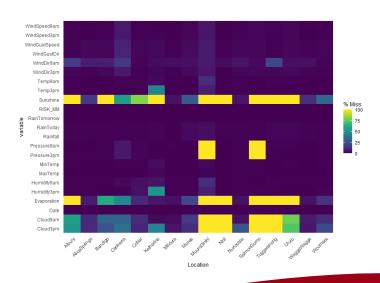


The subset of variables  $S = \{Evaporation, Sunshine, Cloud9am, Cloud3pm\}$  have approximately 40% of the NAs.

But, first of all, how are these values missing? And why?



Missing Values' Pattern





Missing Values' Pattern

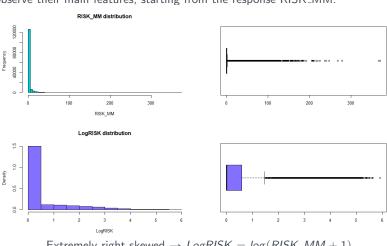
This leads to another fact, that the total absence of observations seems to be **related to the absence of the equipment needed to acquire the measurements** rather than a missingness due to the influence of external factor(s), or our predictors. So, we might not have a pattern of missings related to other variables in the overall analysis.

Our main strategy is to remove the Locations with total absence of measurements wrt the variables with the highest portion of NAs. Another possible strategy should consist in removing the predictors, but it may have bad consequences since we're deleting degrees of freedom instead of a portion of observations. Imputation in our case would be quite ineffective, especially with mean-median imputation of the missing data. So, our aim is to choose the best strategy without altering the structure of the dataset, in particular, relationships among explanatory variables and the response.



#### Rainfall variables

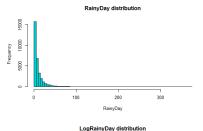
In this section we are going to analyze the distribution of the variables and observe their main features, starting from the response  $RISK\_MM$ .

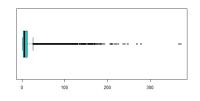


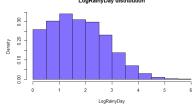


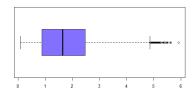
#### Rainfall variables

Class Frequency: "No" = 110314, "Yes" = 31877  $\rightarrow$  focus on rainy day





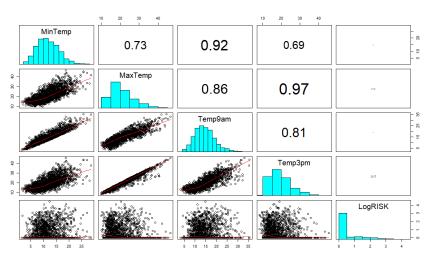




LogRainyDay is approx. a truncated normal



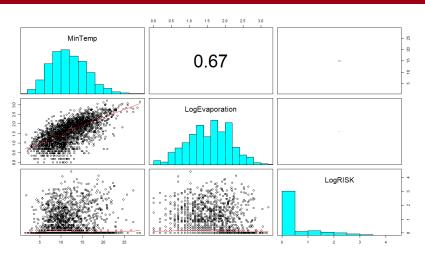
Temperature variables



Approx. bell-shaped. Highly correlated → keep MinTemp



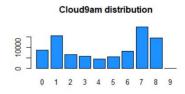
Evaporation

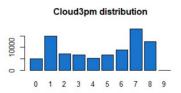


Right skewed and with lots of NAs. Correlated with Temperature variables

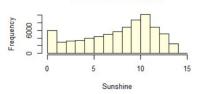


Cloud and Sunshine





#### Sunshine distribution

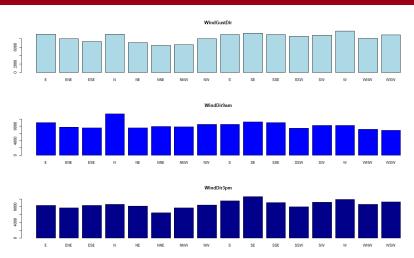


	Cloud9am	Cloud3pm	Sunshine	LogRISK
Cloud9am	1.00	0.41	-0.61	0.22
Cloud3pm	0.41	1.00	-0.69	0.36
Sunshine	-0.61	-0.69	1.00	-0.40
LogRISK	0.22	0.36	-0.40	1.00

Cloud9am and Cloud3pm are discrete. High correlation with Sunshine.



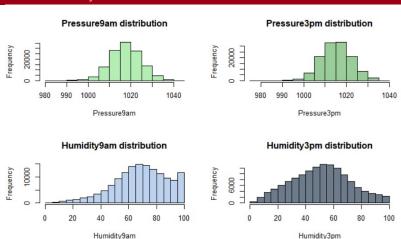
Wind variables



Categorical variables → from compass directions to Cartesian coordinates



Pressure and Humidity

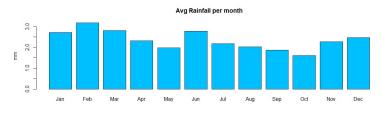


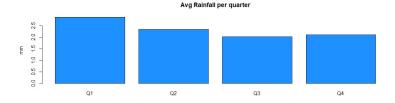
Range of pressure [977; 1041]. Humidity not much symmetric: pick in 100%. High correlation (0.53) between Humidity9am/3pm but might be both relevant. Very high correlation (0.96) between Pressure9am/3pm.



#### Date

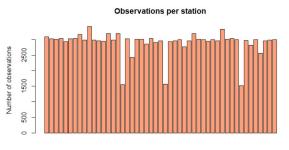
Date of observation, in the format Year-Month-Day  $\rightarrow$  extract the features month and quarter

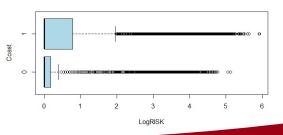






Location





- 49 different weather station;
- Most of them provided about 3000 days of meteorological recordings, but 3 of them (Katherine, Nhil and Uluru) provided approx one half;
- the weather recordings for these 3 stations start from March 2013, while the other stations provided data from February 2008;
- use this variable to create by hand the binary variable Coast.



Creating the datasets

#### **Problems**

- **1 Time Dependency**: the dataset basically is a sequence of time series, so the observations could not be *i.i.d.*;
- 2 NAs: some variables have a high percentage of missing values: certain locations have no recording for some features;
- **3** Tricky distribution shape of the response variable.

#### Possible counter-measures

- 1 Select one observation every 3 days;
- 2 Remove the locations with at least one variable of all NAs or impute NAs, e.g. with unsupervised methods (k-NN) using nearby stations; we might also consider a model based on only one weather station;
- 3 Consider a model regarding only rainy days.



### Model with all the explanatory variables - WeatherV11

We faced the Regression with two main strategies:

### Handcrafted and Hybrid\*Analyses

\*Automatic Procedure put beside with Handcrafted ones

Starting from the setting of the full model  $\mathcal{M}_{23}$ , we expected to encounter multicollinearity among the predictors from the preliminary phase of EDA.

Here we've confirmed our belief, using a statistic called  $\emph{VIF}$  - Inflation Variance Factor. From the literature, in particular with a relevant number of observations, an high VIF can lie in a range between 5 and 10, according to  $\emph{Hastie}$  et  $\emph{Al.}$  ( $\uppsi^1$ )

	GVIF	Df	GVIF(1/(2*Df))
MaxTemp	47.057766	1	6.859866
MinTemp	10.594707	1	3.254951
Temp3pm	55.884768	1	7.475612
Temp9am	24.198561	1	4.919203
Pressure3pm	25.280719	1	5.027994
Pressure9am	25.627725	1	5.062383
Humidity3pm	7.044856	1	2.654215

The square root of the variance inflation factor indicates how much larger the standard error is, compared with what it would be if that variable was uncorrelated with the other predictor variables in the model.



Full Model  $\mathcal{M}_{23}$  — Weather V11

### Characteristics of the full model $\mathcal{M}_{23}$

```
Call:
lm(formula = LogRISK ~ ... data = WeatherV11. na.action = na.omit)
Pesiduals.
   Min
             10 Median
-2.4293 -0.3913 -0.0846 0.2031 4.2671
Coefficients:
                Estimate Std. Error t value Pr(>|t|)
(Intercept)
               24.6322468 0.5441093 45.271 < 2e-16 ***
MaxTemp
               0.0422697 0.0028805 14.674 < 2e-16 ***
MinTemp
               -0.0101193 0.0014848 -6.815 9.50e-12
Temp3pm
               -0.0218684 0.0032007 -6.832 8.44e-12 ***
Temp9am
               -0.0136446 0.0021922 -6.224 4.87e-10 ***
LogEvaporation 0.1205677
                          0.0086129 13.998
Sunshine
               -0.0723552
Wind3pmX
Wind3pmY
               -0.0006827 0.0002637
Wind9amX
Wind9amY
               0.0033893
                          0.0002869 11.813
WindGustX
                          0.0001135 -1.326
WindGustY
               -0.0011459
                          0.0001303 -8.792
                         0.0003847 49.859
Humidity3pm
Humidity9am
               -0.0038992
                          0.0003216 -12.123
Pressure3pm
               -0.0771962
                          0.0021417 -36.045
Pressure9am
               0.0527827 0.0021445 24.613 < 2e-16
Cloud3nm
                         0.0016736
                                    6.369 1.91e-10
Cloud9am
               -0.0200855 0.0015936 -12.604
quarters2
               -0.0325618
                          0.0100942 -3.226
                                             0.00126 **
quarters3
               -0.0141130 0.0106589
                                    -1.324
quarters4
                                     1.811
                                             0.07012 .
Coast1
               -0.0681967 0.0076693 -8.892
LogRainfall
               0.2091049 0.0039697 52.675 < 2e-16 ***
Signif, codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.6955 on 56442 degrees of freedom
(30850 observations deleted due to missingness)
Multiple R-squared: 0.4027. Adjusted R-squared: 0.4025
F-statistic: 1655 on 23 and 56442 DF, p-value: < 2.2e-16
```

■ **BIC** of the model: 119485

■ **AIC** of the model: 119261.5

 $\blacksquare$   $R^2$  of the model: 0.4024949

### Reduced Model



Model with a subset the explanatory variables  $\mathcal{M}_{10,BSS}$  — Weather V11

# Characteristics of the reduced model $\mathcal{M}_{10,BSS}$

```
lm(formula = LogRISK ~ . - MinTemp - Temp9am - guarters - Temp3pm -
   Wind3pmX - Wind3pmY - Wind9amX - Wind9amY - WindGustX - Pressure3pm -
   WindGustY, data = WeatherV11, na.action = na.omit)
Residuals:
            10 Median
-2.3128 -0.4092 -0.0891 0.2034 4.2274
Coefficients:
                Estimate Std. Error t value Pr(>|t|)
(Intercept)
              20.7650934 0.5327769
MaxTemp
               0.0126529 0.0006517 19.415 < 2e-16
LogEvaporation 0.0434223 0.0080193
                                    5.415 6.16e-08
               -0.0763245 0.0014242 -53.590
Sunshine
Humidity3pm
               -0.0027297 0.0002568 -10.629
Humidity9am
Pressure9am
Cloud3pm
               0.0130019 0.0016845
                                     7,718 1,20e-14
Cloud9am
Coast1
               -0.1130630 0.0076763 -14.729
LogRainfall
               0.1726336 0.0039466 43.742
               0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.7117 on 56455 degrees of freedom
 (30850 observations deleted due to missingness)
Multiple R-squared: 0.3745.
                              Adjusted R-squared: 0.3744
```

F-statistic: 3380 on 10 and 56455 DF. p-value: < 2.2e-16

The Subset Selection was performed manually and through BSS. With respect to the full model  $\mathcal{M}_{23}$ , both procedures led to simpler models, preserving the fitting. The BSS model ( $\mathcal{M}_{11,BSS}$  updated to  $\mathcal{M}_{10,BSS}$ , VIF correction) was the least complex ( $\mathcal{M}_{18.M}$ ).

■ **BIC** of  $\mathcal{M}_{10,BSS}$ : 121953.7 (120691.4)

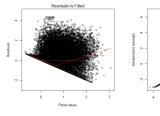
■ AIC of  $\mathcal{M}_{10,BSS}$ : 121846.5 (120530.5)

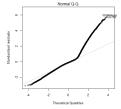
 $\blacksquare$   $R^2$  of  $\mathcal{M}_{10,BSS}$ : 0.3743616 (0.3888385)

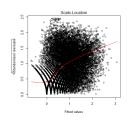
	GVIF
MaxTemp	2.300351
LogEvaporation	2.346201
Sunshine	3.193609
Humidity3pm	2.991694
Humidity9am	2.519413

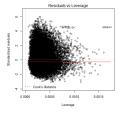


Assumptions and Diagnostics: Model(s) with all (subset) the explanatory variables - Weather V11









- Strong evidence of a systematic non-linear behaviour in the residuals in a neighborhood of 1;
- Assumption of Normality of the residuals not satisfied;
- Heteroskedasticity of the residuals high non-linearity in a neighborhood of zero. The outliers and the leverage point do not represent influencial points (wrt to the Cook's Distance)
- Presence of serious Autocorrelation of the residuals, according to the Durbin-Watson test (\*2):

**DW** = 1.8833, *p*-value <  $2.2 \times 10^{-16}$ .





Focusing on  $\mathcal{M}_{10}$ , we need to make some reflections about the problems encountered during the analysis.

In particular, treating the main sources of multicollinearity in the initial model the procedure led anyway to a model affected by both heteroskedasticity and autocorrelation of the residuals

The **OLS** estimates,  $\hat{\beta}_{OLS}$ , in presence of only one of the encountered issues, still lead to unbiased estimates, but not BLUE (i.e. not efficient), reason why reducing the multicollinearity was so important for our analysis. Furthermore, in absence of the satisfiability of such assumptions, inferential procedures for hypothesis testing and confidence interval building are no longer valid.

So, what's next?



A new perspective: Modeling according to WeatherV11[LogRisk>0,]

### Characteristics of the reduced model $\mathcal{M}_{10}$

```
Call:
lm(formula = LogRISK ~ . - MinTemp - Temp9am - quarters - Temp3pm -
    Wind3pmX - Wind3pmY - Wind9amX - Wind9amY - WindGustX - Pressure3pm -
   WindGustY, data = WeatherV11[LogRISK > 0, ], na.action = na.omit)
Residuals:
   Min
            10 Median
-2.4126 -0.6723 -0.1480 0.5962 3.5599
Coefficients:
                Estimate Std. Error t value Pr(>|t|)
(Intercept)
              22.5404856 1.0004229 22.531 < 2e-16 ***
MaxTemp
               0.0340589 0.0012734 26.746 < 2e-16
LogEvaporation 0.0345192 0.0157145
                                      2.197 0.028057 *
Sunshine
               -0.0655008 0.0030724 -21.319 < 2e-16 ***
               0.0167325 0.0005055 33.099 < 2e-16 ***
Humidity3pm
Humidity9am
              -0.0028149 0.0005764 -4.884 1.05e-06 ***
Pressure9am
              -0.0221332 0.0009749 -22.704 < 2e-16 ***
Cloud3pm
               0.0210094 0.0043395
                                     4.841 1.30e-06 ***
Cloud9am
              -0.0332228 0.0038392 -8.654 < 2e-16 ***
Coast1
              -0.0601403 0.0181843 -3.307 0.000944 ***
LogRainfall
               0.0998348 0.0067130 14.872 < 2e-16 ***
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.9153 on 19680 degrees of freedom
Multiple R-squared: 0.2409.
                               Adjusted R-squared: 0.2405
```

F-statistic: 624.5 on 10 and 19680 DF. p-value: < 2.2e-16

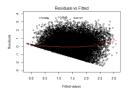
A particular aspect that we already introduced was the atypical shape of the response,  $\ensuremath{\mathbb{Y}}.$ 

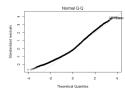
In fact, we noticed how the prevalent presence of zeros induced a non-linearity on its neighborhood, so we decide to restrict the analysis on the "rainy" observations, according to the threshold ( $\mathbb{Y}>0$  mm), modeling  $\mathbb{Y}_{>0}$ .

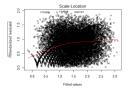
The BSS procedure lead to the following model.

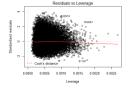


#### Assumptions and Diagnostics: $\mathcal{M}_{10}$ - Weather V11 [LogRisk>0,]









- We can appreciate a consistent reduction in the non-linearity of the residuals;
- Assumption of Normality of the residuals is satisfied (approximately);
- Heteroskedasticity in the residual, present, but relatively unimportant. The outliers and the leverage point do not represent influencial points (wrt to the Cook's Distance)

**DW**=1.8822, *p*-value= $2.2 \times 10^{-16}$ 

■ **BIC** of  $\mathcal{M}_{10,BSS}$ : 52501.8

■ **AIC** of  $\mathcal{M}_{10,BSS}$ : 52407.14

 $\blacksquare$   $R^2$  of  $\mathcal{M}_{10,BSS}$ : 0.2405082



Model with one City - Weather V17Positive

### Characteristics of the reduced model $\mathcal{M}_{10}$

```
lm(formula = LogRISK ~ . - MinTemp - MaxTemp - Temp9am - LogEvaporation -
   Cloud3pm - Cloud9am - Pressure9am - LogRainfall - Wind3pmX
   Wind9amX, data = WeatherV17Positive, na.action = na.omit)
Residuals:
               10 Median
    Min
-2.29815 -0.60932 -0.02128 0.54119 3.07768
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercent) 85,972849
                        5.480944
                                 15.686 < 2e-16
                                 -2.899 0.003839 **
Temp3pm
            -0.032153
                      0.011090
Sunshine
            -0.031739
                       0.011051
                                 -2.872 0.004183 **
Wind3pmY
            0.015598
                       0.003151
                                  4.951 8.94e-07 ***
Wind9amY
            0.012362
                       0.003392
                                  3 645 0 000284
WindGustX
            0.002363
                      0.001106
                                  2.136 0.032964
WindGustY
            0.002832
                      0.001302
                                   2 176 0 029830 *
Humidity3pm 0.007244
                      0.002409
                                  3.008 0.002712 **
Humidity9am -0.003385
                      0.001965
                                  -1.723 0.085248
Pressure3pm -0.083031
                      0.005363 - 15.484 < 2e - 16
                       0.119776
                                   3.803 0.000153 ***
guarters2
            0.455500
quarters3
            0.547042
                        0.127811
                                   4.280 2.08e-05 ***
guarters4
            0.188898
                        0.117084
                                   1.613 0.107042
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 0.7818 on 843 degrees of freedom
 (44 observations deleted due to missingness)
                               Adjusted R-squared: 0.3799
Multiple R-squared: 0.3886.
F-statistic: 44.64 on 12 and 843 DF. p-value: < 2.2e-16
```

The problem is that the models are quite inaccurate if they take into account all the various microclimates that are present in Australia. In fact, considering one city the performance of the Multiple Linear Regression Model improves

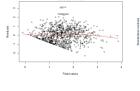
■ BIC of  $\mathcal{M}_{10,BSS}$ : 2089.25 ■ AIC of  $\mathcal{M}_{10,BSS}$ : 2022.718

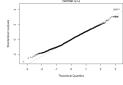
 $\blacksquare$   $R^2$  of  $\mathcal{M}_{10,BSS}$ : 0.37986

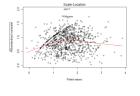
But still, residuals were autocorrelated.

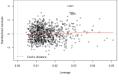


### Assumptions and Diagnostics: $\mathcal{M}_{10}$ - Weather V17Positive









- We can appreciate a consistent reduction in the non-linearity of the residuals;
- Assumption of Normality of the residuals is satisfied (approximately);
- Heteroskedasticity in the residual, present, but unimportant. The outliers and the leverage point do not represent influencial points (wrt to the Cook's Distance)

Presence of Autocorrelation of the residuals, according to the Durbin-Watson test
DW=1.897, p-value=0.05117

■ **BIC** of M<sub>10,BSS</sub>: 2089.25

■ AIC of  $\mathcal{M}_{10,BSS}$ : 2022.718

■  $R^2$  of  $\mathcal{M}_{10,BSS}$ : 0.37986



#### Residuals' Modeling: Using AR(1) according to Cochrane-Orcutt

Serially autocorrelated residuals could lead to inexact inferences, making the model pretty unuseful.

Still, we wanted to look for possible ways to model them, since they present a systematic pattern (in a specific interval of time), so we applied the **Cochrane-Orcutt**, removing the autocorrelation of the residuals, expecting a 1<sup>st</sup> order autocorrelation pattern (also more complex patterns are plausible).

 $\blacksquare$   $R^2$  of  $\mathcal{M}_{10,BSS}$ : 0.388176

Weather17	$\mathcal{M}_{10}$	$\mathcal{M}_{10,OC}$
Intercept	5.480944	5.5828056
Temp3pm	0.011090	0.0112755
Sunshine	0.011051	0.0110389
Wind3pmY	0.003151	0.0031483
Wind9amY	0.003392	0.0033935
WindGustX	0.001106	0.0011013
WindGustY	0.001302	0.0013022
Humidity3pm	0.002409	0.0024117
Humidity9am	0.001965	0.0019597
Pressure3pm	0.005363	0.0054621
quarters2	0.119776	0.0054621
quarters3	0.127811	0.1318719
quarters4	0.117084	0.1217643
DW	$\mathcal{M}_{10}$	$\mathcal{M}_{10,OC}$
	1.89696, 5×10 <sup>-2</sup>	1.99721, 4.41×10 <sup>-1</sup>

### Classification Model



#### Implementation and results

The aim of this section is to build a classifier able to predict whether tomorrow will rain or not. We performed three different procedures for this task.

### Procedure 1:

- selection of the dataset with the 13 most promising variables and one observation every 3 days;
- division in training and test set (80% and 20%);
- construction of different Logistic Regression models with different number of variables;
- selection of the best model (the one with 9 variables), according to the deviance difference test;
- selection of the threshold that provides the best Sensitivity and Specificity.

### Classification Model



Implementation and results

#### Results 1:

- List of the 9 variables of the final model: Sunshine, Wind3pmX, Wind9amX, Wind9amY, Humidity3pm, Humidity9am, Pressure3pm, quarters, LogRainfall;
- VIF values are quite close to 1. The largest are those corresponding to Humidity3pm/9am.

Performance measures on the training set:

Accuracy	Specificity	Sensitivity
0.780	0.774	0.8

Performance measures on the test set:

Accuracy	Specificity	Sensitivity
0.773	0.756	0.827

#### Procedure 2:

- selection of the dataset with all the variables and one observation every 3 days;
- under-sampling to balance the dataset;
- division in training and test set (80% and 20%);
- Principal Component Analysis and Logistic Regression;

Accuracy of the resulting model:

Training set accuracy	Test set accuracy
0.777	0.769

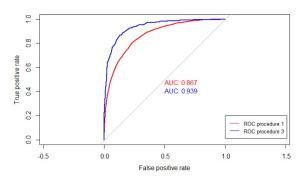
## Classification Model



Implementation and results

Procedure 3:
to conclude, we
considered the dataset
concerning just one city,
and we performed the
same steps described in
the first procedure.

In this case, the classifier is able to model the phenomenon more precisely, as can be seen from these results.



Performance measures on the test set:

Accuracy	Specificity	Sensitivity
0.848	0.851	0.836

## Next step



### **Regression Analysis**

- Fitting a Truncated Normal Model, incorporating the information of the Date in the model;
- Adapt a Semi-Parametrical apporach, with Regression Splines (according to the termplot() function's output);
- Apply the latter models considering Clusters of cities, taking into account their specific microclimate;

## That's all folks!



Thank you for the attention

## Technical appendix



**Counting Outliers** 

```
# Define a function that counts the number of outliers and
# their percentage.
NumberOfOutliers <- function(variable) {
  first_quartile = quantile(variable, na.rm = TRUE)[2]
  third_quartile = quantile(variable, na.rm = TRUE)[4]
  igr = IQR(variable, na.rm = TRUE)
  lower = first_quartile - 1.5*iqr
  upper = third_quartile + 1.5*iqr
  num_outliers = sum(variable < lower | variable > upper,
                                               na.rm = TRUE
  perc_outliers = 100*num_outliers/length(variable)
  return ( data . frame (num_outliers , perc_outliers ))
```

## Technical appendix



### Compass directions to cartesian coordinates

```
# Define a function that convert compass directions into cartesian coordinates,
# to keep track of the cyclic behaviour of the directions.
CardinalToNumbers <- function(variable) {
        l = length(variable)
        NS = rep(0.1)
        WE = rep(0, 1)
        cosine = c(0, 0.5, sqrt(2)/2, sqrt(3)/2, 1, sqrt(3)/2, sqrt(2)/2, 0.5,
                   0. -0.5. -sart(2)/2. -sart(3)/2. -1. -sart(3)/2. -sart(2)/2. -0.5
        sine = c(1, sart(3)/2, sart(2)/2, 0.5, 0, -0.5, -sart(2)/2, -sart(3)/2,
                 -1, -\operatorname{sqrt}(3)/2, -\operatorname{sqrt}(2)/2, -0.5, 0, 0.5, \operatorname{sqrt}(2)/2, \operatorname{sqrt}(3)/2)
        conversor = matrix(c(cosine, sine), nrow = 2, byrow = TRUE)
        colnames (conversor) = directions
        for (i in seq(1.1.length.out = 1)){}
                if (is.na(variable[i])){
                        NS[i] = NA
                        WE[i] = NA
                else -
                        NS[i] = conversor[2, variable[i]]
                        WE[i] = conversor[1. variable[i]]
        return (data . frame (NS. WE))
```



■ The VIF  $(*^1)$  coefficient is a multiple that determines the increase of the variance of  $\beta_j$  due to the correlation between  $X_j$  and other explanatory variables.

$$VIF_j = \frac{1}{1 - R_j^2}$$

More precisely, it is shown (e.g. Greene, 2011) that:

$$\mathbb{V}\left(\hat{\beta}_{j}\right) = VIF_{j} \frac{\sigma^{2}}{\sum_{i=1}^{n} \left(x_{ij} - \bar{x}_{j}\right)^{2}}$$

So, if  $R_j^2=0$  and therefore  $\mathit{VIF}_j=1$ , we have that  $\mathbb{V}(\beta_j)$  coincides with the variance that we would have if  $X_j$  was the only one explanatory variable in the model. As  $R_j^2$  increases,  $\mathbb{V}(\beta_j)$  will also increase: the *multicollinearity leads to instability in estimates*. "Generalized collinearity diagnostics", Monette et Al.

## Technical appendix

### Università degli Studi di Padova

#### **DW Test Statistic**

■ The DW test is used to test the hypothesis that the residuals are serially correlated at lag 1 , i.e. that in the following model:

$$\epsilon_i = \rho \epsilon_{i-1} + v_i$$

the hypothesis being tested is:

$$\left\{ \begin{array}{ll} H_0: \rho = 0 & \text{ (no serial correlation)} \\ H_1: \rho \neq 0 & \text{ (serial correlation at lag 1)} \end{array} \right.$$

since we do not observe the true error terms, we use the OLS residuals  $\hat{\epsilon}_i$  and calculate the Durbin-Watson statistic as:

$$DW = \frac{\sum_{i=2}^{N} (\hat{\epsilon}_{i} - \hat{\epsilon}_{i-1})^{2}}{\sum_{i=1}^{N} \hat{\epsilon}_{i}^{2}}$$

Furthermore, its distribution no longer holds, when the equation of  $Y_i$  contains a lagged dependent variable,  $Y_{i-1}$ . As a quick rule of thumb, if the DW statistic is near 2, then we do not reject the null hypothesis of no serial correlation.