Statistics 191: Introduction to Applied Statistics

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Statistics 191: Introduction to Applied Statistics Correlated Errors, Whitening

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Correlated Erorrs

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Topics

- Autocorrelation.
- Whitening.

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Autocorrelation

- In the random effects model, outcomes within groups were correlated.
- Other regression applications also have correlated outcomes (i.e. errors).
- Common examples: time series data.
- Why worry? Can lead to underestimates of SE \rightarrow inflated t's \rightarrow false positives.

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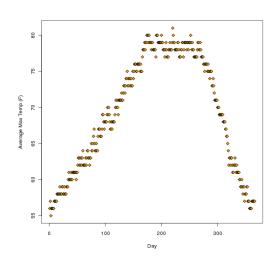
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Autocorrelation

- Suppose we plot Palo Alto's daily average temperature clearly we would see a pattern in the data.
- Sometimes, this pattern can be attributed to a deterministic phenomenon (i.e. predictable seasonal fluctuations).
- Other times, "patterns" are due to correlations in the noise, maybe small time fluctuations in the stock market, economy, etc.
- Example: financial time series, monthly bond return.
- Example: residuals regressing consumer expenditure on money stock.

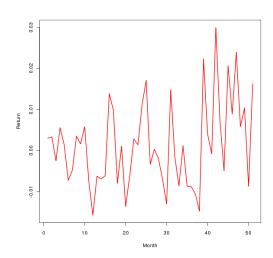
Average Max Temp in Palo Alto

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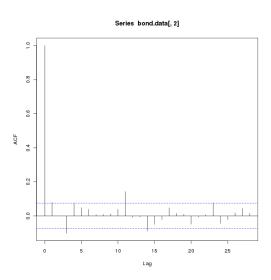
Monthly bond return

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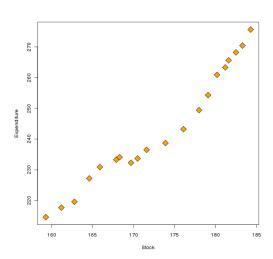
Monthly bond return, ACF

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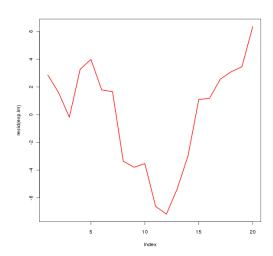
Expenditure vs. stock

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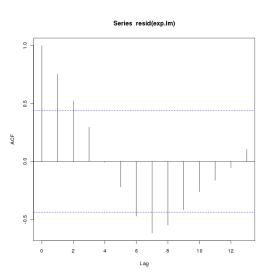
Expenditure vs. stock: residuals

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ACF of residuals

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AR(1) noise

 Suppose that, instead of being independent, the errors in our model were

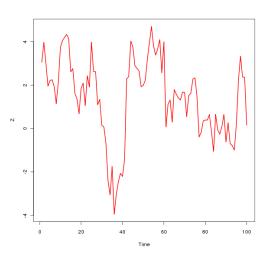
$$\varepsilon_t = \rho \cdot \varepsilon_{t-1} + \omega_t, \qquad -1 < \rho < 1$$

with $\omega_t \sim N(0, \sigma^2)$ independent.

- If ρ is close to 1, then errors are very correlated, $\rho=0$ is independence.
- This is "Auto-Regressive Order (1)" noise (AR(1)). Many other models of correlation exist: ARMA, ARIMA, ARCH, GARCH, etc.

AR(1) noise, $\rho = 0.9$

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Autocorrelation function

• For a "stationary" time series $(Z_t)_{1 \le t \le \infty}$ define

$$ACF(t) = Cor(Z_s, Z_{s+t}).$$

- Stationary means that correlation above does not depend on s.
- For AR(1) model,

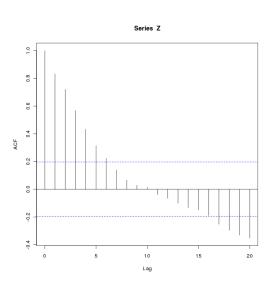
$$ACF(t) = \rho^t$$
.

• For a sample (Z_1, \ldots, Z_n) from a stationary time series

$$\widehat{ACF}(t) = \frac{\sum_{j=1}^{n-t} (Z_j - \overline{Z})(Z_{t+j} - \overline{Z})}{\sum_{j=1}^{n} (Z_j - \overline{Z})^2}.$$

ACF of AR(1) noise, $\rho = 0.9$

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Effects on inference

- So far, we have just mentioned that things may be correlated, but not thought about how it affects inference.
- Suppose we are in the "one sample problem" setting and we observe

$$W_i = Z_i + \mu, \qquad 1 \le i \le n$$

with the Z_i 's from an AR(1) time series. It is easy to see that

$$E(\overline{W}) = \mu$$

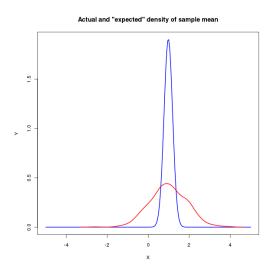
BUT, generally

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$$Var(\overline{W}) > \frac{Var(Z_1)}{r}$$

Misleading inference ignoring autocorrelation

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Model

item Observations:

$$Y_t = \beta_0 + \sum_{j=1}^{p} X_{tj}\beta_j + \varepsilon_t, \qquad 1 \le t \le n$$

Errors:

$$\varepsilon_t = \rho \cdot \varepsilon_{t-1} + \omega_t, \qquad -1 < \rho < 1$$

Question: who do we determine if autocorrelation is present? Question: what do we do to correct for autocorrelation?

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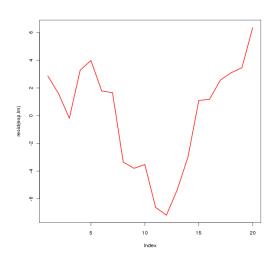
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Graphical checks for autocorrelation

- A plot of residuals vs. time is helpful.
- Residuals clustered above and below 0 line can indicate autocorrelation.
- Example: regressing consumer expenditure on money stock.

Expenditure vs. stock: residuals

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Durbin-Watson test

• In regression setting, if noise is AR(1), a simple estimate of ρ is obtained by (essentially) regressing e_t onto e_{t-1}

$$\widehat{\rho} = \frac{\sum_{t=2}^{n} (e_t e_{t-1})}{\sum_{t=1}^{n} e_t^2}.$$

• To formally test $H_0: \rho = 0$ (i.e. whether residuals are independent vs. they are AR(1)), use Durbin-Watson test, based on

$$d=2(1-\widehat{\rho}).$$

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Correcting for AR(1)

ullet Suppose we know ho, if we "pre-whiten" the data and regressors

$$\tilde{Y}_{t+1} = Y_{t+1} - \rho Y_t, t > 1$$

$$\tilde{X}_{(t+1)j} = X_{(t+1)j} - \rho X_{tj}, i > 1$$

for $1 \le t \le n-1$. This model satisfies "usual" assumptions, i.e. the errors

$$\tilde{\varepsilon}_t = \varepsilon_{t+1} - \rho \cdot \varepsilon_t$$

are independent $N(0, \sigma^2)$.

- For coefficients in new model $\tilde{\beta}$, $\beta_0 = \tilde{\beta}_0/(1-\rho)$, $\beta_j = \tilde{\beta}_j$.
- Problem: in general, we don't know ρ .

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Two-stage regression

- Step 1: Fit linear model to unwhitened data (OLS: ordinary least squares, i.e. no pre-whitening).
- Step 2: Estimate ρ with $\widehat{\rho}$.
- Step 3: Pre-whiten data using $\widehat{\rho}$ refit the model.

Other models of covariance

- Suppose we model covariance of ε 's differently, i.e. ARMA(p, q).
- As long as we can estimate parameters of covariance, from residuals of the OLS fit, we can use this two-stage procedure.
- This is very similar to weighted least squares.

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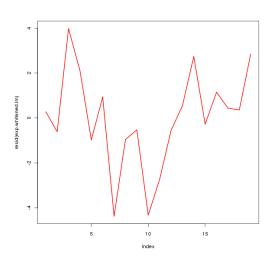
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Interpreting results of two-stage fit

- Basically, interpretation is unchanged, but the exact degrees of freedom in the error is not exactly clear.
- Common argument: "this works for large degrees of freedom, so we better hope we have enough degrees of freedom so this point is not important."
- Can treat *t*-statistics as *Z*-statistics, F's as χ^2 , appealing to asymptotics:
 - t_{ν} , with ν large is like N(0,1);
 - $F_{j,\nu}$, with ν large is like χ_j^2/j .

Expenditure vs. stock: whitened residuals

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Expenditure vs. stock: ACF of whitened residuals

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