Report

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Introduction

In this report an MLE estimation of the constant rates model is calculated. The general model, with non constant birth-death rates, is also discussed (I could not find an analytical MLE solution for that case).

In addition, simulated data and an intent of finding a numerical solution were performed.

Analitycal analysis

We define,

- β : Speciation rate.
- δ : Extinction rate.
- t_i : Time when the event i (speciation or extinction) occurs.
- x_i : Number of species on time t_i .

The probability of a speciation ocurrs on time t is given by the exponential probability distribution

$$P_{\beta}(T=t;\beta) = -\beta e^{-\beta t}$$

whereas the probabilty of an extinction occurs on time t is

$$P_{\delta}(T=t;\delta) = -\delta e^{-\delta t}$$

Moreover, given the waiting time t_1 the probability of an speciation event occurs is $\beta/(\beta+\delta)$ and the probability an expeciation happens on t_i is $\delta/(\beta+\delta)$

Thus, a possible likelihood function corrisponding to the phylogenic tree would be

$$L = \prod_{i=1}^{n} (\beta + \delta) x_i e^{-(\beta + \delta) x_i t_i} \left(\frac{\beta}{\beta + \delta} \right)^{b_i} \left(\frac{\delta}{\beta + \delta} \right)^{d_i}$$

$$P(Birth; waiting time) = \frac{\beta}{\beta + \delta}$$

$$P(Death; waiting time) = \frac{\delta}{\beta + \delta}$$

Figure 1: Probability of speciation and extintions event.

Thus, the loglikelihood function is

$$lL = \sum_{i=1}^{n} ln(\beta + \delta) + ln(x_i) - (\beta + \delta)x_it_i + b_i(ln(\beta) - ln(\beta + \delta)) + d_i(ln(\delta) - ln(\beta + \delta))$$

Finally, to maximize the likelihood we solve the equations

$$\frac{\partial lL}{\partial \beta} = 0; \qquad \frac{\partial lL}{\partial \delta} = 0 \tag{1}$$

The analytical solution corresponding to equations (1) are

$$\beta = \frac{N \sum b_i}{\sum x_i t_i (\sum b_i + \sum d_i)}$$
$$\delta = \frac{N \sum d_i}{\sum x_i t_i (\sum b_i + \sum d_i)}$$

This is an useful result as a first step, however, since we assume β and δ are constant numbers, this is not yet a realistic case. We would be interested in the case when the speciation and extinction parameters depends on covariates, for instance

$$log(\beta) = \theta_0 + \theta_1 body \ size + \theta_2 ave. \ temp.$$

We consider the simple case

$$\beta = e^{\theta_0 + \theta_1 a}$$
$$\delta = e^{\phi_0 + \phi_1 a}$$

and we call it the model M2.

For MLE calculation, by the chain rule we have

$$\frac{\partial lL}{\partial \theta_0} = \frac{\partial lL}{\partial \beta} \frac{\partial \beta}{\partial \theta_0} = e^{\theta_0 + \theta_1 a} \frac{\partial lL}{\partial \beta};$$
$$\frac{\partial lL}{\partial \theta_1} = \frac{\partial lL}{\partial \beta} \frac{\partial \beta}{\partial \theta_0} = ae^{\theta_0 + \theta_1 a} \frac{\partial lL}{\partial \beta};$$

and similarly

$$\frac{\partial lL}{\partial \phi_0} = e^{\phi_0 + \phi_1 a} \frac{\partial lL}{\partial \delta};$$
$$\frac{\partial lL}{\partial \phi_1} = a e^{\phi_0 + \phi_1 a} \frac{\partial lL}{\partial \delta}.$$

$$log(\beta) = A\theta$$

$$log(\delta) = A\phi$$

Figure 2: An option is to model the parameters as a GLM

$$\begin{split} \frac{\partial \beta}{\partial \theta_0} &= e^{\theta_0 + \theta_1 a} \\ \frac{\partial \beta}{\partial \theta_1} &= a e^{\theta_0 + \theta_1 a} \end{split}$$

$$\frac{\partial \delta}{\partial \theta_0} = e^{\phi_0 + \phi_1 a}$$

$$\frac{\partial \delta}{\partial \theta_1} = a e^{\phi_0 + \phi_1 a}$$

$$\frac{\partial lL}{\partial \theta_0} = 0; \quad \frac{\partial lL}{\partial \theta_1} = 0; \quad \frac{\partial lL}{\partial \phi_0} = 0; \quad \frac{\partial lL}{\partial \phi_1} = 0$$
 (2)

is equivalent to (1) and we have two equations but four parameters.

Simulations

We simulate data from the M2 model:

```
nT = 15
theta = c(1, 2)
phi = c(0, 1)
set.seed(123)
dat <- NULL
AA <- NULL
AA <- list()
A = rexp(1, 4)
dat[[1]] \leftarrow list(tm = 0, x = 1, A = A)
X = 0
X[1] = 1
tm = 0
D = 0
B = 0
Tm = tm
AA[[1]] = A
i <- 2
stm <- 0
while (i <= nT \& dat[[i - 1]]$x > 0) {
    beta <- theta[1] + theta[2] * dat[[i - 1]]$A
    delta <- phi[1] + phi[2] * dat[[i - 1]]$A</pre>
    prev.x = dat[[i - 1]]$x
    tm <- rexp(1, (sum(beta) + sum(delta)))</pre>
    stm <- stm + tm
    prob <- c(delta, beta)/(sum(beta) + sum(delta))</pre>
    BD <- sample(2 * prev.x, 1, prob = prob)
    if (BD <= prev.x) {</pre>
        x = prev.x - 1
        A = dat[[i - 1]] A[-BD]
        D[i] = 1
        B[i] = 0
    } else {
        x = prev.x + 1
        A = c(dat[[i - 1]] A, dat[[i - 1]] A[BD -
```

```
prev.x] * rgamma(1, 100, 100))
        D[i] = 0
        B[i] = 1
    }
    dat[[i]] \leftarrow list(tm = stm, x = x, A = A)
    X[i] = x
    Tm[i] = tm
    AA[[i]] = A
    i <- i + 1
}
  To see the iterations we type
dat
## [[1]]
## [[1]]$tm
## [1] O
##
## [[1]]$x
## [1] 1
##
## [[1]]$A
## [1] 0.2108643
##
##
## [[2]]
## [[2]]$tm
## [1] 0.3531868
##
## [[2]]$x
## [1] 2
##
## [[2]]$A
## [1] 0.2108643 0.2355912
##
##
## [[3]]
## [[3]]$tm
## [1] 1.320555
##
## [[3]]$x
## [1] 3
```

##

```
## [[3]]$A
## [1] 0.2108643 0.2355912 0.2366910
##
##
## [[4]]
## [[4]]$tm
## [1] 1.520057
##
## [[4]]$x
## [1] 4
##
## [[4]]$A
## [1] 0.2108643 0.2355912 0.2366910 0.2136786
##
##
## [[5]]
## [[5]]$tm
## [1] 1.524415
## [[5]]$x
## [1] 5
##
## [[5]]$A
## [1] 0.2108643 0.2355912 0.2366910 0.2136786
## [5] 0.2758153
##
##
## [[6]]
## [[6]]$tm
## [1] 1.569685
##
## [[6]]$x
## [1] 6
##
## [[6]]$A
## [1] 0.2108643 0.2355912 0.2366910 0.2136786
## [5] 0.2758153 0.2699149
##
##
## [[7]]
## [[7]]$tm
## [1] 1.610068
## [[7]]$x
```

```
## [1] 7
##
## [[7]]$A
## [1] 0.2108643 0.2355912 0.2366910 0.2136786
## [5] 0.2758153 0.2699149 0.2400447
##
##
## [[8]]
## [[8]]$tm
## [1] 1.739818
##
## [[8]]$x
## [1] 6
##
## [[8]]$A
## [1] 0.2108643 0.2366910 0.2136786 0.2758153
## [5] 0.2699149 0.2400447
##
##
## [[9]]
## [[9]]$tm
## [1] 1.786115
##
## [[9]]$x
## [1] 7
##
## [[9]]$A
## [1] 0.2108643 0.2366910 0.2136786 0.2758153
## [5] 0.2699149 0.2400447 0.2729063
##
##
## [[10]]
## [[10]]$tm
## [1] 1.865547
##
## [[10]]$x
## [1] 8
##
## [[10]]$A
## [1] 0.2108643 0.2366910 0.2136786 0.2758153
## [5] 0.2699149 0.2400447 0.2729063 0.2432511
##
##
## [[11]]
```

```
## [[11]]$tm
## [1] 1.949677
##
## [[11]]$x
## [1] 9
##
## [[11]]$A
## [1] 0.2108643 0.2366910 0.2136786 0.2758153
## [5] 0.2699149 0.2400447 0.2729063 0.2432511
## [9] 0.2427908
##
##
## [[12]]
## [[12]]$tm
## [1] 1.99815
## [[12]]$x
## [1] 10
##
## [[12]]$A
## [1] 0.2108643 0.2366910 0.2136786 0.2758153
## [5] 0.2699149 0.2400447 0.2729063 0.2432511
## [9] 0.2427908 0.1758324
##
##
## [[13]]
## [[13]]$tm
## [1] 2.124588
##
## [[13]]$x
## [1] 11
##
## [[13]]$A
## [1] 0.2108643 0.2366910 0.2136786 0.2758153
## [5] 0.2699149 0.2400447 0.2729063 0.2432511
## [9] 0.2427908 0.1758324 0.2370110
##
##
## [[14]]
## [[14]]$tm
## [1] 2.142101
##
## [[14]]$x
## [1] 12
```

```
##
## [[14]]$A
## [1] 0.2108643 0.2366910 0.2136786 0.2758153
## [5] 0.2699149 0.2400447 0.2729063 0.2432511
## [9] 0.2427908 0.1758324 0.2370110 0.2635687
##
##
## [[15]]
## [[15]]$tm
## [1] 2.202867
##
## [[15]]$x
## [1] 13
##
## [[15]]$A
## [1] 0.2108643 0.2366910 0.2136786 0.2758153
## [5] 0.2699149 0.2400447 0.2729063 0.2432511
## [9] 0.2427908 0.1758324 0.2370110 0.2635687
## [13] 0.2066377
```

That is

```
it = length(dat)
tree <- data.frame(time = it, x = 2)
for (i in 1:it) {
    tree[i, ] = c(t = dat[[i]]$tm, x = dat[[i]]$x)
}
library(xtable)
options(xtable.comment = FALSE)
options(xtable.booktabs = TRUE)
xtable(tree, caption = "Table with simulated waiting times and number of species")</pre>
```

	time	х
1	0.00	1.00
2	0.35	2.00
3	1.32	3.00
4	1.52	4.00
5	1.52	5.00
6	1.57	6.00
7	1.61	7.00
8	1.74	6.00
9	1.79	7.00
10	1.87	8.00
11	1.95	9.00
12	2.00	10.00
13	2.12	11.00
14	2.14	12.00
15	2.20	13.00

Table 1: Table with simulated waiting times and number of species

We try an MLE calculation maximizing the Likelihood function, but for different initial values the estimated parameters are very different and most of them does not make any sense.

```
fn <- function(theta) {
    beta = 0
    delta = 0
    for (i in 1:nT) {
        beta[i] = theta[1] + sum(theta[2] * AA[[i]])
        delta[i] = theta[3] + sum(theta[4] * AA[[i]])
    }
    sum(log(beta + delta) + log(x) - (beta + delta) *
        x * Tm + D * (log(beta) - log(beta + delta)) +
        B * (log(delta) - log(beta + delta)))
}</pre>
```

```
nlm(fn, theta <- c(0.5, 0.1, 0.7, 0.8), hessian = TRUE)
## $minimum
## [1] -362635.6
##
## $estimate
## [1] 2861.808 3155.901 2855.841 3147.169
##
## $gradient
## [1] -28.63700 -31.58295 -28.63543 -31.58056
##
## $hessian
                 [,1]
                               [,2]
## [1,] -3.908965e-08 -3.093552e-08
## [2,] -3.093552e-08 -3.915688e-08
## [3,] -1.994177e-08 -3.875021e-09
## [4,] -3.877670e-09 -5.860527e-10
##
               [,3]
                        [, 4]
## [1,] -1.994177e-08 -3.877670e-09
## [2,] -3.875021e-09 -5.860527e-10
## [3,] -2.697764e-07 -3.179857e-07
## [4,] -3.179857e-07 -5.359631e-07
##
## $code
## [1] 4
##
## $iterations
## [1] 100
```