01NAEX - Lecture 01 Introduction to Design of Experiments (DOE)

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What does Statistical design of experiments (DOE) means and brings

Experiment: A test or a series of tests where purposeful changes are made to input variables to observe and identify reasons for changes in the output response.

DOE: The process of planning experiments so that data can be analyzed by statistical methods, resulting in valid, objective, and meaningful conclusions.

All experiments are designed experiments, some are poorly designed, some are well-designed.

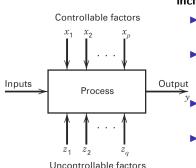
What does DOE means and brings

DOE

- is a methodology for systematically applying statistics to experimentation.
- lets experimenters develop a mathematical model that predicts how input variables interact to create output variables or responses in a process or system.
- can be used for a wide range of experiments for various purposes including nearly all fields of engineering and even in business marketing.

General model of a process or system

The process or system can be represented by the model, where we can visualize the process as a combination of operations, machines, methods, people, and other resources that transforms some input into an output that has one ore more observable response variables.



The objectives of the experiment may include the following:

- Determining which variables are most influential on the response y.
- ▶ Determining where to set the influential x's so that y is almost always near the desired nominal value.
 - Determining where to set the influential x's so that variability in y is small.
- Determining where to set the influential x's so that the effects of the variables z's are minimized.

Short motivation example

Imagine you're trying to weigh your luggage before a flight with moody scale to avoid extra fees.

Weighting problem: we want to weight 3 different items A, B, and C.

Standard weighting procedure - 1				
No.	measurement		_	Model for responses
	1	(1)	No item	$(1) = \mu + \varepsilon_1$

Canadand:

1	(1)	No item	$(1) = \mu + \varepsilon_1$
2	a	item A	
3	b	item B	$b = \mu + B + \varepsilon_3$
4	С	item C	$c = \mu + C + \varepsilon_4$

Where μ is offset of weighting device, A,B,C are weights of items A,B,C, and $\varepsilon_1,\varepsilon_2,\varepsilon_3,\varepsilon_4$ are measurement errors with standard deviation σ_ε .

The estimate of A is given by $\hat{A}^{(1)}=a-(1)=A+\varepsilon_2-\varepsilon_1$ and the corresponding for B and C.

Short motivation example

Weighting problem: we want to weight 3 different items A, B, and C.

Alternative weighting procedure - 2

		0 0.	
No. measurement	Notation	Meaning	Model for responses
1	(1)	No item	$(1)=\mu+\varepsilon_5$
2	ac	item A and C	$\mathit{ac} = \mu + \mathit{A} + \mathit{C} + arepsilon_{6}$
3	bc	item B and C	$bc = \mu + B + C + \varepsilon_7$
4	ab	item A and B	$ab = \mu + A + B + arepsilon_8$

The estimate of A is now given by $\hat{A}^{(2)} = \frac{-(1)+ac-bc+ab}{2} = A + \frac{-\varepsilon_5 + \varepsilon_6 - \varepsilon_7 + \varepsilon_8}{2}$ and the corresponding for B and C.

$$Var\left[\hat{A}^{(1)}\right] = 2\sigma_{\varepsilon}^{2}$$

 $Var\left[\hat{A}^{(2)}\right] = \sigma_{\varepsilon}^{2}$

The second design is more precise than the first one, since in the first design not all measurements are used to estimate all parameters, which is the case in the second design.

How does beer and brewery relate to DOE?

Did you know that the origins of modern DOE are linked to beer?

- Back in the early 1900s, William Sealy Gosset, an employee at the Guinness Brewery in Dublin, was trying to improve the quality of stout beer.
- Due to limited resources, he developed small sample statistical methods.
- ▶ What was the name of the test he used and why?

Short history of DOE

- ► The agricultural origins, 1920s 1940s:
 - Sir R. A. Fisher & his co-workers.
 - Lady tasting tea, The Design of Experiments book (1935)
 - Factorial designs, ANOVA.
- ► The first industrial era, 1950s 1970s:
 - Box & Wilson, Response surface methodology (RSM).
 - ▶ Applications in chemical and process industries, finance, etc.
 - Total quality control (TQC), Statistical process control (SPC).
- ► The second industrial era, 1970s 1990s:
 - Quality control, Japanese style of TQC.
 - ► Taguchi and robust parameter design, process robustness.
 - ► Total Quality Management (TQM), Continuous Quality Improvement (CQI).
- ► The modern era, 1990s till today:
 - Wide use of computer technology and experiments in DOE.
 - Expanded use of DOE in Six-Sigma and in business.
 - Globalization is driving all sectors of the economy to be more competitive.

The basic principles of design of experiments

Think of randomization like shuffling a deck of cards before dealing. If you don't shuffle well, someone might end up with all the aces, and that's no fun. Unless you're the one getting the aces.

Randomization:

- Cornerstone underlying the use of statistical methods.
- Running trials in random order to balance out effects of hidden variables.
- Allocation of the experimental material is random.
- Reduce bias and systematic errors.

The basic principles of design of experiments

Replication is like trying a new flavor of ice cream multiple times. You know, just to be sure it's really your new favorite!

- An independent repeat of each factor combination.
- It allows us to estimate experimental error and improves the precision of our results.
- More data (sample size) gives a better estimate, reducing the impact of any one "odd" result.

The basic principles of design of experiments

Blocking is like organizing your wardrobe by seasons. You can't control the weather, but at least you won't end up wearing a swimsuit in a snowstorm!

Blocking:

- It is a design technique used to improve the precision.
- lt reduces or eliminates the variability from nuisance factors.
- Increases the precision of the experiment by accounting for known sources of variability.

Strategy of Experimentation

Best-guess approach:

- Wide-spread, trial and error experiments.
- More successful than you might suspect.
- Disadvantage:
 - cannot guarantee best solution has been found.

One-factor-at-a-time (OFAT) approach:

- Sometimes associated with the scientific or engineering method.
- Disadvantage:
 - fails to consider any possible interaction between factors.
 - very inefficient, requires many test runs.

Statistically designed approach:

- Based on Fisher's factorial concept.
- Modern and most efficient approach.
- Can determine how factors interact.

Factorial Designs

Treatment x Nuisance factors

- Treatment factor factor of primary interest to us.
- Nuisance factor factor which are not our primary focus, but we have to deal with them.

Experimental x Classification factors

- Experimental factor we can specify them and set the levels.
- ▶ Classification factor we can not change or randomly assign them.

Quantitative x Qualitative factors

- Quantitative factor we can assign any specified level of them.
- Qualitative factor have categories which are different types.

Guidelines for Designing Experiments

1. Recognition and statement of problem:

- Obvious, but in practice not so simple point.
- It is necessary to to develop all ideas about the objectives of the experiment, get input from all concerned parties, the team approach.

2. Choice of factors, levels, and ranges:

- ► The potential design factors are those factors that we wish to vary in the experiment .
- Classify factors as: design factor, held-constant factor, allow-to-vary factors, nuisance factors.
- Nuisance factors may have large effects that must be accounted for (controllable, uncontrollable, noise).
- ▶ Need to use engineering judgment or prior test results.

3. Selection of the response variable(s):

- Be sure that the variable really provides useful information about the process under study.
- Multiple responses are not unusual.

Guidelines for Designing Experiments

4. Choice of experimental design:

Involves consideration of sample size (no. of replicates), selection of run order, whether blocking or randomization restriction, software to use etc

5. Performing the experiment:

- Vital to monitor the process carefully
- Up/for planning is crucial to success
- Easy to underestimate logistical and planning aspects of running and design experiment in development environment

6. Statistical analysis of the data:

- Should be used to analyze the data so that results and conclusions are objective rather than judgmental in nature
- Use simple graphics whenever possible

7. Drawing conclusions and recommendations:

Follow-up test runs and confirmation testing to validate the conclusions from the experiment

Useful advices in Planning, Conducting & Analyzing an Experiment

Use your **non-statistical knowledge** of the problem, physical laws, background knowledge, etc. It is crucial to success.

KISS principle (Keep it simple - Stupid!, Keep it simple and straightforward, Keep it short and simple). Don't use complex, sophisticated statistical techniques. If design is good, If design is good, analysis is relatively straightforward. If design is bad - even the most complex and elegant statistics cannot save the situation.

Think and experiment sequentially, pre-experimental planning is vital.

Recognize the difference between $\mbox{{\bf practical}}$ and $\mbox{{\bf statistical}}$ $\mbox{{\bf significance}}.$

General definitions from probability - probability distribution

Assume, probability triplet (Ω, \mathcal{F}, p) and a random variable , say y.

Probability distribution: If y is discrete, we call the p(y) the probability mass function of y. If y is continuous, we call f(y) the probability density function for y.

Mean, expected value:

$$\mu = E(y) \left\{ \begin{array}{ll} \int_{-\infty}^{+\infty} y f(y) dy & y \text{ is continuous random variable} \\ \sum_{\text{all } y} y p(y) & y \text{ is discrete random variable} \end{array} \right.$$

Variance, standard deviation:

$$\sigma^2 = Var(y) \left\{ \begin{array}{ll} \int_{-\infty}^{+\infty} (y-\mu)^2 f(y) dy & y \text{ is continuous random variable} \\ \sum_{\text{all } y} (y-\mu)^2 p(y) & y \text{ is discrete random variable} \end{array} \right.$$

General definitions: notation that we will use

If the population contains N elements and a sample of n of them is to be selected, and if each of the possible samples has an equal probability of being chosen, then the procedure employed is called random sampling.

Suppose that $y_1, y_2, \dots y_n$ represents a sample. Than we define two statistics: **Sample mean**:

$$\bar{y} = \frac{1}{n} \sum_{i=1}^{n} y_i$$

The sample mean \bar{y} is a unbiased point estimator of the population mean μ Sample variance:

$$S^{2} = \frac{1}{n-1} \sum_{i=1}^{n} (y_{i} - \bar{y})^{2}$$

The sample variance S^2 is a unbiased point estimator of the population variance σ^2 . Sometimes $S=\sqrt{S^2}$ is called **sample standard deviation**. The expression

$$SS = \sum_{i=1}^{n} (y_i - \bar{y})^2$$

is called the corrected sum of squares of the observations y_i .

General definitions from probability - Degrees of Freedom

Number of degrees of freedom refers to number of independent elements in sum of squares SS:

$$E(SS) = E\left[\sum_{i=1}^{n} (y_i - \bar{y})^2\right] = (n-1)\sigma^2.$$

The degrees of freedom associated with a sum-of-squares is the degrees-of-freedom of the corresponding component vectors.

Only (n-1) elements from the set $(y_1 - \bar{y}, y_2 - \bar{y}, \dots, y_n - \bar{y})$ are independent.

General definitions from probability - Normal distribution, χ^2 distribution

Let y be a random variable with normal distribution then the probability density function of y is

$$f(y) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2}\frac{(y-\mu)^2}{\sigma^2}}$$

where $\mu \in R$ is the mean of the distribution, $\sigma^2 > 0$) is the variance.

If x_1, x_2, \ldots, x_k are *iid* random variables with standard normal distribution then $y = x_1^2 + x_2^2 + \ldots + x_k^2$ is a random variable with χ^2 distribution with k degrees of freedom, denoted χ^2_k and the probability density function of y is

$$f(y) = \frac{1}{2^{\frac{k}{2}} \Gamma(\frac{k}{2})} y^{\frac{k}{2} - 1} e^{-\frac{y}{2}}$$

The distribution is skewed with mean $\mu = k$ and variance $\sigma^2 = 2k$.

General definitions from probability - t distribution

If z and χ_k^2 are independent standard normal and chi-squared r.v. then

$$t_k = \frac{z}{\sqrt{\frac{\chi_k^2}{k}}}$$

is r.v. with t distribution with k degrees of freedom and the probability density function of t_k is

$$f(y) = \frac{\Gamma(\frac{k+1}{2})}{\sqrt{(k\pi)}\Gamma(\frac{k}{2})} \frac{1}{\left(\frac{t^2}{k} + 1\right)^{\frac{k+1}{2}}}$$

The distribution is symmetric with mean $\mu=0$ and variance $\sigma^2=\frac{k}{k-2}$ for k>2 If $k=\infty$, the t distribution becomes the standard normal distribution.

General definitions from probability - F distribution

If χ_k^2 and χ_l^2 are two independent chi-squared r.v. with k and l degrees of freedom then the ratio

$$F_{k,l} = \frac{\frac{\chi_k^2}{k}}{\frac{\chi_l^2}{l}}$$

follows the $F_{k,l}$ distribution.

F distribution is very important for us, suppose we have two independent normal populations with common variance σ .

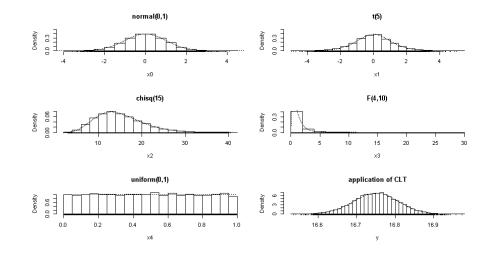
If $y_{11}, y_{12}, \ldots, y_{1n_1}$ is a random sample of n_1 observations from first population with sample variance S_1^2 and $y_{21}, y_{22}, \ldots, y_{2n_2}$ is a random sample of n_2 observations from second population with sample variance S_2^2 then

$$\frac{S_1^2}{S_2^2} \sim F_{n_1-1,n_2-2}.$$

The result comes from following

$$\frac{SS}{\sigma^2} = \frac{\sum_{i=1}^{n} (y_i - \bar{y})^2}{\sigma^2} \sim \chi_{n-1}^2.$$

General definitions from probability - CLT



General definitions from probability - Hypothesis testing

A statistical hypothesis is a statement about the parameters of one or more populations.

Suppose that we partition the parameter space Θ into two disjoint sets Θ_0 and Θ_1 and that we wish to test:

The null hypothesis $H_0: \theta \in \Theta_0$ versus The alternative hypothesis $H_1: \theta \in \Theta_1$

To test a hypothesis, we devise a procedure for taking a random sample, computing an appropriate test statistics, and then rejecting or failing to reject the null hypothesis H_0

Decisions in Hypothesis Testing

Decision	H₀ Is True	H_0 Is False	
Retain H_0 no error		type II error	
Reject H ₀	type I error	no error	

Significance level: $\alpha = P(\text{type I error})$

Power: $1 - \beta = 1 - P$ (type II error) = P (reject $H_0 \mid H_0$ is false)

General definitions from probability - p-value

The p-value is the smallest level of significance that would lead to rejection of the null hypothesis H_0 with the given data.

- ► The p-value is the most commonly reported statistic resulting from hypothesis testing.
- ▶ The significance level α is a threshold for deciding whether to reject H_0 . If the test is rejected at the significance level α , it will be also rejected at the level $\alpha' > \alpha$.
- The p-value is a measure of the evidence against H₀ (the smaller the p-value, the stronger evidence against H₀).
- ► The p-value in not the probability that H₀ is true (The p-value is calculated under the assumption that H₀ is true.)

p-value
$$\neq P(H_0 \mid Data)$$

p-value = $P(Data as extreme or more extreme than observed | <math>H_0$)

- ▶ 1 (p-value) is not the probability of the alternative hypothesis being true.
- ▶ The significance level of the test is not determined by the p-value.

General definitions from probability - Test on means with variance known

Tests on Means with Variance Known

Hypothesis	Test Statistic	Fixed Significance Level Criteria for Rejection	P-Value
H_0 : $\mu = \mu_0$			
$H_1: \mu \neq \mu_0$		$ Z_0 > Z_{lpha/2}$	$P = 2[1 - \Phi(Z_0)]$
$H_0: \mu = \mu_0$ $H_1: \mu < \mu_0$ $H_2: \mu = \mu_0$	$Z_0 = \frac{\bar{y} - \mu_0}{\sigma / \sqrt{n}}$	$Z_0 < -Z_{\alpha}$	$P = \Phi(Z_0)$
$H_0: \mu = \mu_0$ $H_1: \mu > \mu_0$		$Z_0 > Z_\alpha$	$P = 1 - \Phi(Z_0)$
H_0 : $\mu_1 = \mu_2$			
$H_1: \mu_1 \neq \mu_2$		$ Z_0 > Z_{lpha/2}$	$P = 2[1 - \Phi(Z_0)]$
$H_0: \mu_1 = \mu_2$	$\overline{y}_1 - \overline{y}_2$	$Z_0 < -Z_{cc}$	$P = \Phi(Z_0)$
$H_1: \mu_1 < \mu_2$	$Z_0 = \frac{y_1 - y_2}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$	$Z_0 \setminus -Z_{\alpha}$	$\Gamma = \Psi(Z_0)$
H_0 : $\mu_1 = \mu_2$			
$H_1: \mu_1 > \mu_2$		$Z_0 > Z_{\alpha}$	$P = 1 - \Phi(Z_0)$

General definitions from probability - Test on means with variance unknown

Tests on Means of Normal Distributions, Variance Unknown

Hypothesis	Test Statistic	Fixed Significance Level Criteria for Rejection	P-Value
$H_0: \mu = \mu_0$ $H_1: \mu \neq \mu_0$		$ t_0 > t_{\alpha/2, n-1}$	sum of the probability above t_0 and below $-t_0$
$H_0: \mu = \mu_0$ $H_1: \mu < \mu_0$ $H_0: \mu = \mu_0$	$t_0 = \frac{\overline{y} - \mu_0}{S/\sqrt{n}}$	$t_0 < -t_{\alpha,n-1}$	probability below t_0
$H_1: \mu > \mu_0$	$\frac{1}{1} \text{ if } \sigma_1^2 = \sigma_2^2$	$t_0 > t_{\alpha, n-1}$	probability above t_0
H_0 : $\mu_1 = \mu_2$			
H_1 : $\mu_1 \neq \mu_2$	$t_0 = \frac{\bar{y}_1 - \bar{y}_2}{S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$	$ t_0 > t_{lpha/2,v}$	sum of the probability above t_0 and below $-t_0$
	$v = n_1 + n_2 - 2$ if $\sigma_1^2 \neq \sigma_2^2$		
$H_0: \mu_1 = \mu_2$ $H_1: \mu_1 < \mu_2$	$t_0 = \frac{\overline{y}_1 - \overline{y}_2}{\sqrt{\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}}}$	$t_0 < -t_{\alpha,v}$	probability below t_0
$H_0: \mu_1 = \mu_2$ $H_1: \mu_1 > \mu_2$	$v = \frac{\left(\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}\right)^2}{\frac{\left(S_1^2/n_1\right)^2}{n_1 - 1} + \frac{\left(S_2^2/n_2\right)^2}{n_2 - 1}}$	$t_0 > t_{a,v}$	probability above t_0

General definitions from probability - Test on Variance of Normal Distributions

Tests on Variances of Normal Distributions

Hypothesis	Test Statistic	Fixed Significance Level Criteria for Rejection
$H_0: \sigma^2 = \sigma_0^2$ $H_1: \sigma^2 \neq \sigma_0^2$		$\chi_0^2 > \chi_{lpha/2, n-1}^2$ or $\chi_0^2 < \chi_{1-lpha/2, n-1}^2$
$H_0: \sigma^2 = \sigma_0^2$ $H_1: \sigma^2 < \sigma_0^2$	$\chi_0^2 = \frac{(n-1)S^2}{\sigma^2}$	$\chi_0^2 < \chi_{1-\alpha,n-1}^2$
$H_0: \sigma^2 = \sigma_0^2$ $H_1: \sigma^2 > \sigma_0^2$	00	$\chi_0^2 > \chi_{\alpha,n-1}^2$
$H_0: \sigma_1^2 = \sigma_2^2$ $H_1: \sigma_1^2 \neq \sigma_2^2$	$F_0 = \frac{S_1^2}{S_2^2}$	$F_0 > F_{\alpha/2, n_1 - 1, n_2 - 1}$ or
$H_0: \sigma_1^2 = \sigma_2^2$ $H_1: \sigma_1^2 < \sigma_2^2$	$F_0 = \frac{S_2^2}{S_1^2}$	$F_0 < F_{1-\alpha/2,n_1-1,n_2-1}$ $F_0 > F_{\alpha,n_2-1,n_1-1}$
$H_0: \sigma_1^2 = \sigma_2^2$ $H_1: \sigma_1^2 > \sigma_2^2$	$F_0 = \frac{S_1^2}{S_2^2}$	$F_0 > F_{\alpha, n_1 - 1, n_2 - 1}$

General definitions from probability - Confidence interval

Confidence interval is a kind of interval estimate of a population parameter and is used to indicate the reliability of an estimate.

To obtain an interval estimate of an unknown parameter θ , we need to find two statistics L (lower limit) and U (upper limit) such that the probability statement

$$P(L \le \theta \le U) = 1 - \alpha$$

is true. The interval ($L \le \theta \le U$ is called a $100(1-\alpha)$ percent confidence interval for parameter θ .

Simple Comparative Experiment

We consider experiments to compare two conditions (also called treatments, levels, factors, etc).

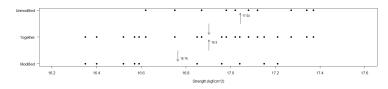
In the simple comparative experiment we use previous mentioned statistical concepts, such as random samples, sampling distributions, hypothesis testing etc.

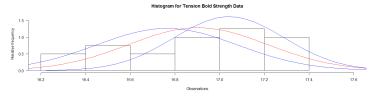
Simple Comparative Experiment - data

Tension Bond Strength Data Portland Cement Data

Observation	Modified Mortar	Unmodified Mortar			
j	y 1j	У2j			
1	16.57	18.15			
2	16.59	17.96			
3	17.15	17.90			
4	16.96	18.22			
5	17.04	17.75			
6	16.52	17.86			
7	16.35	18.00			
8	17.21	18.25			
9	16.40	17.63			
10	16.85	17.50			

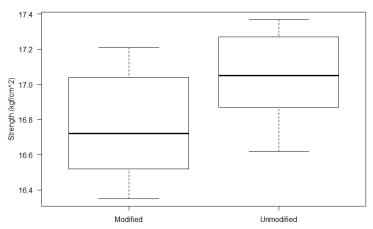
Simple Comparative Experiment - data plot



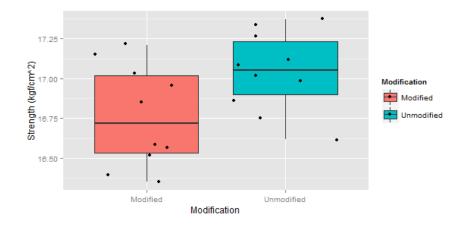


Simple Comparative Experiment - boxplot

Box plots for the portland cement tension bond strength data



Simple Comparative Experiment - boxplot (from gplot package)



Simple Comparative Experiment - F-test

If we do not know the variances of the data sets, we have to test, if they are identical or not.

```
var.test(y1, y2,ratio = 1,alternative="two.sided",
conf.level = 0.95)
F test to compare two variances
data: y1 and y2
F = 1.6293, num df = 9, denom df = 9, p-value = 0.4785
alternative hypothesis:
true ratio of variances is not equal to 1
90 percent confidence interval:
0.5125235 5.1792350
95 percent confidence interval:
0.4046845 6.5593806
99 percent confidence interval:
0.2490804 10.6571184
sample estimates:
ratio of variances 1.629257
```

Simple Comparative Experiment - The Two-Sample t-test

If we could assume that the variances of tension bond strengths are identical, or the sample sd are reasonably similar, we can use simple two-sample t-test to compare means.

The test statistic:

$$t_0 = \frac{\bar{y_1} - \bar{y_2}}{S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} = -2.186876,$$

where \mathcal{S}_{p} is an estimate of common variance $\sigma^{2}=\sigma_{1}^{2}=\sigma_{2}^{2}$

$$S_p = \frac{(n_1 - 1)S_1^2 + (n_2 - 1)S_2^2}{n_1 + n_2 - 2} = 0.2842534.$$

Because $t_0=-2.186876<-t_{0.025,18}=-2.100922$, we would reject the H_0 hypothesis at the significance level $\alpha=0.05$.

Simple Comparative Experiment - The Two-Sample t-test

Confidence interval on the true difference in means $\mu_1-\mu_2$ for the Portland cement problem.

The statistic:

$$t_0 = rac{ar{y_1} - ar{y_2} - (\mu_1 - \mu_2)}{S_p \sqrt{rac{1}{n_1} + rac{1}{n_2}}}$$

is distributed as $t_{n_1+n_2-2}$ and thus,

$$P(\bar{y}_1 - \bar{y}_2 - t_{\frac{\alpha}{2}, n_1 + n_2 - 2} S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}} \le \mu_1 - \mu_2 \le \bar{y}_1 - \bar{y}_2 + t_{\frac{\alpha}{2}, n_1 + n_2 - 2} S_p \sqrt{\frac{1}{n_1} + \frac{1}{n_2}})$$

is $100(1-\alpha)$ percent confidence interval for $\mu_1 - \mu_2$.

For $\alpha =$ 0.05 we obtain $-0.5450734 \le \mu_1 - \mu_2 \le -0.01092661$.

Simple Comparative Experiment - The Two-Sample t-test

If we could assume that the variances of tension bond strengths are identical, or the sample *sd* are reasonably similar, we can use simple two-sample t-test to compare means.

```
> t.test(y1,y2,alternative="two.sided",mu=0,
paired=F,var.equal=T,conf.level=0.95)
Two Sample t-test
data: y1 and y2
t = -2.1869, df = 18, p-value = 0.0422
alternative hypothesis:
true difference in means is not equal to 0
95 percent confidence interval:
-0.54507339 -0.01092661
99 percent confidence interval:
-0.64391308 0.08791308
sample estimates:
mean of x: 16.764, mean of y: 17.042
```

Simple Comparative Experiment - t-test

If the variances are not identical, the classical test statistic is not distributed exactly as t and some approximation of degrees of freedom must be used.

```
> t.test(y1,y2,alternative="two.sided",mu=0,
paired=FALSE, var.equal=FALSE, conf.level=0.95)
Welch Two Sample t-test
data: y1 and y2
t = -2.1869, df = 17.025, p-value = 0.043
alternative hypothesis:
true difference in means is not equal to 0
95 percent confidence interval:
-0.546174139 -0.009825861
99 percent confidence interval:
-0.64636235 0.09036235
sample estimates:
mean of x: 16.764, mean of y: 17.042
```

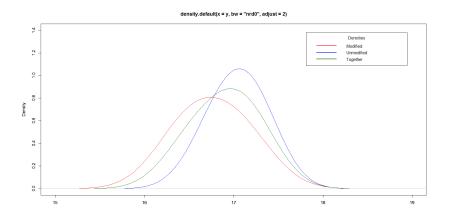
In using classical t-test we make following assumptions:

- observations are independent random variables.
- both samples are drawn from independent populations that are normal distributed.
- variances of both populations are equal.

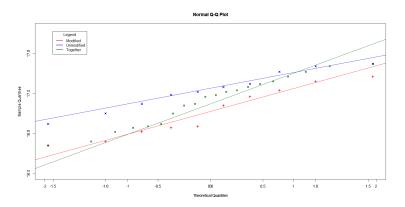
Strategy for Checking Model Assumptions:

- check the form of the model.
- check for outliers.
- check for independence.
- check for constant variance.
- check for normality.

Kernel density estimations of data sets



The equal variance and normality assumptions are easy to check using normal probability plot (QQ-plot).



Test for the hypothesis of normality.

One sample Kolmogorov - Smirnov test tests the null hypothesis that a probability distribution of given sample is the same as a reference probability distribution.

Shapiro - Wilk normality test and **Anderson - Darling normality test** test the null hypothesis that a given sample came from a normally distributed population.

Hypothesis testing of normality

Data	Shapiro-Wilk test		Kolmogorov-Smirnov test	
	Statistics W	p-value	Statistics D	p-value
Modified	0.9186	0.3457	0.2088	0.7027
Unmodified	0.9626	0.8153	0.1211	0.9943
Together	0.9540	0.4326	0.1226	0.8891

Simple Comparative Experiment -Non-parametric alternative

For two-sample comparisons, two non-parametric tests can be used, depending on the way data are collected. If both sample are independent, we use Mann-Whitney-Wilcoxon rank sum test, while for paired sample the corresponding test is called Wilcoxon signed rank test.

Both are called using R function:

wilcox.test

and the option

paired={TRUE|FALSE}.

Choice of sample size

In hypothesis testing we eliminate the probability of type I error by setting the significance level α , but how to decrease the probability of type II error β ?

In t-test, the Power depends on the significance level α , on the true difference in means δ and the sample size n.

Two-sample t test power calculation

Sample size calculation	Power calculation	
n = 28.17	n = 10	
delta = 0.28	delta = 0.28	
sd = 0.28	sd = 0.28	
sig.level = 0.05	sig.level = 0.05	
power = 0.95	power = 0.54	
${\sf alternative} = {\sf two.sided}$	alternative = two.sided	

45

Choice of sample size

Calculation of Power for equal but unknown populations SD's:

We have four connected quantities:

- sample size;
- difference between means:
- significance level of the test;
- power of the test.

We can plot the graph of β (the probability of type II error) versus δ - the true difference in means. This curve is called Operating Characteristic curve (OC curve).

Choice of sample size

Operating characteristic curve (OC curve) for two-sided t-test, where $d=\frac{|\delta|}{2\sigma}$.

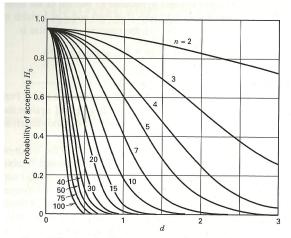
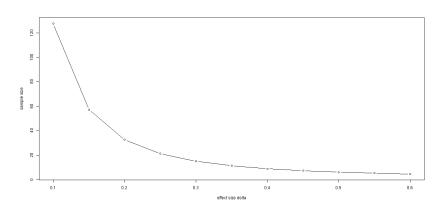


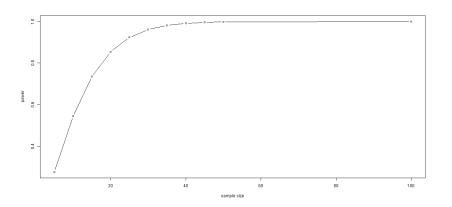
Figure 2-12 Operating characteristic curves for the two-sided t-test with $\alpha=0.05$. (Reproduced with permission from "Operating Characteristics for the Common Statistical Tests of Significance," C. L. Ferris, F. E. Grubbs, and C. L. Weaver, *Annals of Mathematical Statistics*, June 1946.)

Choice of sample size - Sample curve

```
for (i in c(.1,.15,.2,.25,.3,.35,.4,.45,.5,0.55,0.6))
{
    pwrt<-power.t.test(sd=0.284,delta=i,power=.8,sig.level=.05,
        type="two.sample",alternative="two.sided")
        ptab<-rbind(ptab, cbind(pwrt$d, pwrt$n))
}
plot(ptab[,1],ptab[,2],type="b",xlab="effect size",ylab="sample size")</pre>
```



Choice of sample size - Power curve



Next Lecture and Today Exercise

Next Lecture: Comparing several treatment means, linear regression

- Review basic statistical concepts and analysis of variance (ANOVA).
- Design and analyze single factor experiments using ANOVA.
- Analysis of the fixed effects model.
- Interpretation of results.
- Model diagnostic checking, residual analysis (normal probability plot, plot of residuals, residuals versus fitted values).
- Multiple linear regression model.
- Submodel, Stepwise regression.

Today Exercise:

- Run and familiarize with R.
- Solve problems from Montgomery Design and Analysis of Experiments Github link to exercise