

The Distribution of a Linear Combination

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The Distribution of a Linear Combination

The sample mean \overline{X} and sample total T_o are special cases of a type of random variable that arises very frequently in statistical applications.

Definition

Given a collection of n random variables X_1, \ldots, X_n and n numerical constants a_1, \ldots, a_n , the rv

$$Y = a_1 X_1 + \dots + a_n X_n = \sum_{i=1}^n a_i X_i$$
 (5.7)

is called a **linear combination** of the X's.

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For example, $4X_1 - 5X_2 + 8X_3$ is a linear combination of X_1 , X_2 , and X_3 with $a_1 = 4$, $a_2 = -5$, and $a_3 = 8$.

Taking $a_1 = a_2 = \ldots = a_n = 1$ gives $Y = X_1 + \ldots + X_n = T_0$, and $a_1 = a_2 = \ldots = a_n = \frac{1}{n}$ yields

$$Y = \frac{1}{n}X_1 + \dots + \frac{1}{n}X_n = \frac{1}{n}(X_1 + \dots + X_n) = \frac{1}{n}T_o = \overline{X}$$

Notice that we are not requiring the X_i 's to be independent or identically distributed. All the X_i 's could have different distributions and therefore different mean values and variances. We first consider the expected value and variance of a linear combination.

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Proposition

Let X_1, X_2, \ldots, X_n have mean values μ_1, \ldots, μ_n , respectively, and variances $\sigma_1^2, \ldots, \sigma_n^2$, respectively.

1. Whether or not the X_i 's are independent,

$$E(a_1X_1 + a_2X_2 + \dots + a_nX_n) = a_1E(X_1) + a_2E(X_2) + \dots + a_nE(X_n)$$

= $a_1\mu_1 + \dots + a_n\mu_n$ (5.8)

2. If X_1, \ldots, X_n are independent,

$$V(a_1X_1 + a_2X_2 + \dots + a_nX_n) = a_1^2V(X_1) + a_2^2V(X_2) + \dots + a_n^2V(X_n)$$

= $a_1^2\sigma_1^2 + \dots + a_n^2\sigma_n^2$ (5.9)

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And

$$\sigma_{a_1X_1+\dots+a_nX_n} = \sqrt{a_1^2\sigma_1^2 + \dots + a_n^2\sigma_n^2}$$
 (5.10)

3. For any $X_1, ..., X_n$,

$$V(a_1X_1 + \cdots + a_nX_n) = \sum_{i=1}^n \sum_{j=1}^n a_i a_j \text{Cov}(X_i, X_j)$$
 (5.11)

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Proofs are sketched out at the end of the section. A paraphrase of (5.8) is that the expected value of a linear combination is the same as the linear combination of the expected values—for example, $E(2X_1 + 5X_2) = 2\mu_1 + 5\mu_2$.

The result (5.9) in Statement 2 is a special case of (5.11) in Statement 3; when the X_i 's are independent, $Cov(X_i, X_j) = 0$ for $i \neq j$ and $= V(X_i)$ for i = j (this simplification actually occurs when the X_i 's are uncorrelated, a weaker condition than independence).

Specializing to the case of a random sample (X_i 's iid) with $a_i = 1/n$ for every i gives $E(\overline{X}) = \mu$ and $V(\overline{X}) = \sigma^2/n$. A similar comment applies to the rules for T_o .

Example 29

A gas station sells three grades of gasoline: regular, extra, and super.

These are priced at \$3.00, \$3.20, and \$3.40 per gallon, respectively.

Let X_1 , X_2 , and X_3 denote the amounts of these grades purchased (gallons) on a particular day.

Suppose the X_i 's are independent with μ_1 = 1000, μ_2 = 500, μ_3 = 300, σ_1 = 100, σ_2 = 80, and σ_3 = 50.

Example 29

cont'd

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The revenue from sales is $Y = 3.0X_1 + 3.2X_2 + 3.4X_3$, and

$$E(Y) = 3.0\mu_1 + 3.2\mu_2 + 3.4\mu_3$$

= \$5620

$$V(Y) = (3.0)^2 \sigma_1^2 + (3.2)^2 \sigma_2^2 + (3.4)^2 \sigma_3^2$$

= 184,436

$$\sigma_{V} = \sqrt{184,436}$$

$$= $429.46$$

The Difference Between Two Random Variables

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The Difference Between Two Random Variables

An important special case of a linear combination results from taking n = 2, $a_1 = 1$, and $a_2 = -1$:

$$Y = a_1 X_1 + a_2 X_2 = X_1 - X_2$$

We then have the following corollary to the proposition.

Corollary

$$E(X_1-X_2)=E(X_1)-E(X_2)$$
 for any two rv's X_1 and X_2 . $V(X_1-X_2)=V(X_1)+V(X_2)$ if X_1 and X_2 are independent rv's.

The Difference Between Two Random Variables

The expected value of a difference is the difference of the two expected values, but the variance of a difference between two independent variables is the *sum*, *not* the difference, of the two variances.

There is just as much variability in X_1-X_2 as in X_1+X_2 [writing $X_1-X_2=X_1+(-1)X_2$, $(-1)X_2$ has the same amount of variability as X_2 itself].

Example 30

A certain automobile manufacturer equips a particular model with either a six-cylinder engine or a four-cylinder engine.

Let X_1 and X_2 be fuel efficiencies for independently and randomly selected six-cylinder and four-cylinder cars, respectively. With μ_1 = 22, μ_2 = 26, σ_1 = 1.2, and σ_2 = 1.5,

$$E(X_1 - X_2) = \mu_1 - \mu_2$$

$$= 22 - 26$$

$$= -4$$

Example 30

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cont'd

$$V(X_1 - X_2) = \sigma_1^2 + \sigma_2^2$$

$$= (1.2)^2 + (1.5)^2$$

$$= 3.69$$

$$\sigma_{X_1 - X_2} = \sqrt{3.69}$$

$$= 1.92$$

If we relabel so that X_1 refers to the four-cylinder car, then $E(X_1 - X_2) = 4$, but the variance of the difference is still 3.69.

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The Case of Normal Random Variables

The Case of Normal Random Variables

When the X_i 's form a random sample from a normal distribution, \overline{X} and T_o are both normally distributed. Here is a more general result concerning linear combinations.

Proposition

If X_1, X_2, \ldots, X_n are independent, normally distributed rv's (with possibly different means and/or variances), then any linear combination of the X_i 's also has a normal distribution. In particular, the difference $X_1 - X_2$ between two independent, normally distributed variables is itself normally distributed.

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Example 31

The total revenue from the sale of the three grades of gasoline on a particular day was $Y=3.0X_1+3.2X_2+3.4X_3$, and we calculated $\mu_{r}=5620$ and (assuming independence) $\sigma_{r}=429.46$. If the $X_{\rm S}$ are normally distributed, the probability that revenue exceeds 4500 is

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$$P(Y > 4500) = P\left(Z > \frac{4500 - 5620}{429.46}\right)$$
$$= P(Z > -2.61)$$
$$= 1 - \Phi(-2.61)$$
$$= .9955$$

The Case of Normal Random Variables

The CLT can also be generalized so it applies to certain linear combinations. Roughly speaking, if *n* is large and no individual term is likely to contribute too much to the overall value, then Y has approximately a normal distribution.