

1. Conjugate priors

(a) Let

$$\text{Likelihood: } \mathbb{P}(x_1, \dots, x_N) = \prod_i^N \lambda \exp(-\lambda x_i)$$

$$\text{Prior: } \text{gamma}(\lambda|\alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda}$$

Then,

Posterior:

$$\begin{aligned} \mathbb{P}(\lambda|x_1, \dots, x_N) &= \prod_i^N \lambda \exp(-\lambda x_i) \cdot \frac{\beta^\alpha}{\Gamma(\alpha)} \lambda^{\alpha-1} e^{-\beta\lambda} \\ &= \lambda \exp\left(\sum_i^N -\lambda x_i\right) \lambda^{\alpha-1} e^{-\beta\lambda} \frac{\beta^\alpha}{\Gamma(\alpha)} \\ &= \lambda^{\alpha+N-1} \exp\left(\sum_i^N -\lambda x_i - \beta\lambda\right) \frac{\beta^\alpha}{\Gamma(\alpha)} \\ &= \lambda^{\alpha+N-1} \exp\left(-\lambda \sum_i^N x_i + \beta\right) \frac{\beta^\alpha}{\Gamma(\alpha)} \sim \text{gamma}(\alpha + N, \beta + \sum_i^N x_i) \end{aligned}$$

Since the posterior also has a gamma distribution, we find the updates parameters are of the form  $\alpha + N, \beta + \sum_i^N x_i$ . To find the prediction distribution,

$$\begin{aligned} \mathbb{P}(x_{N+1}|x_1, \dots, x_N) &= \frac{\beta^\alpha}{\Gamma(\alpha)} \int \lambda \exp(-\lambda x_{N+1}) \cdot \lambda^{\alpha+N-1} \exp\left(-\lambda(\beta + \sum_i^N x_i)\right) d\lambda \\ &= \frac{\beta^\alpha}{\Gamma(\alpha)} \int \lambda \cdot \lambda^{\alpha+N-1} \exp\left(-\lambda(\beta + \sum_i^{N+1} x_i)\right) d\lambda \\ &= \frac{\beta^\alpha}{\Gamma(\alpha)} \int \lambda \cdot P(\lambda|\alpha + N, \beta + \sum_i^{N+1} x_i) d\lambda \end{aligned}$$

We note that this describes the expectation for  $\lambda$  given a gamma function  $\sim \text{gamma}(\lambda|\alpha + N, \beta + \sum_i^{N+1} x_i)$ . Therefore,

$$\mathbb{P}(x_{N+1}|x_1, \dots, x_N) = \frac{\alpha + N}{\beta + \sum_i^{N+1} x_i}$$

(b) Given the geometric distribution

$$P(X_i = k|\theta) = (1 - \theta)^{k-1} \cdot \theta$$

and the beta distribution

$$\beta(\theta|a, b) = \alpha \theta^{a-1} (1 - \theta)^{b-1}$$

we prove that the beta distribution is the conjugate prior for a likelihood with a geometric distribution.

$$\begin{aligned}
P(\theta|X) &= P(\theta) \cdot P(X|\theta) \\
&= \alpha \cdot \theta^{a-1} \cdot (1-\theta)^{b-1} \cdot (1-\theta)^{k-1} \cdot \theta \\
&= \alpha \cdot \theta^a \cdot (1-\theta)^{b+k-2} \\
&= \beta(\theta|a+1, b+k-1)
\end{aligned}$$

The posterior has the form of a beta distribution so therefore the beta distribution is the conjugate prior for the geometric distribution.

The update procedure for a beta posterior simply involves updating the  $a$  and  $b$  parameters

$$\begin{aligned}
a_{N+1} &\leftarrow a_N + 1 \\
b_{N+1} &\leftarrow b_N + k - 1
\end{aligned}$$

2. Bayesian Naive Bayes

3. Logistic regression for credit scoring

- (a) The data structure we chose for logistic regression is simply a class that keeps a set of weights for each of the features, has an update method for updating the weights, and draws predictions using the logit function

$$\text{Probability} = \frac{1}{1 + e^{-w^T x}}$$

- (b) The likelihood is

$$\begin{aligned}
L(w) &= \frac{1}{1 + e^{-yw^T x}} \\
\log \text{likelihood} &= \log \frac{1}{1 + e^{-yw^T x}} \\
&= -\log(1 + e^{-yw^T x}) \\
\text{negative log likelihood} &= \log(1 + e^{-yw^T x})
\end{aligned}$$

Now we compute the gradient of the negative log likelihood

$$\begin{aligned}
\nabla \log(1 + e^{-yw^T x}) &= \nabla \log \left( \frac{e^{yw^T x} + 1}{e^{yw^T x}} \right) \\
&= \nabla \left( \log(e^{yw^T x} + 1) - \log(e^{yw^T x}) \right) \\
&= \left( \frac{1}{e^{yw^T x} + 1} \cdot e^{yw^T x} \cdot -yx_i \right) - \left( \frac{1}{e^{yw^T x}} \cdot e^{yw^T x} \cdot -yx_i \right) \\
&= yx_i - yx_i \cdot \frac{e^{yw^T x}}{e^{yw^T x} + 1} \\
&= yx_i - yx_i \cdot \left( \frac{e^{yw^T x} + 1}{e^{yw^T x}} \right)^{-1} \\
&= yx_i - yx_i \cdot (1 + e^{-yw^T x})^{-1} \\
&= yx_i \left( 1 - \frac{1}{1 + e^{-yw^T x}} \right)
\end{aligned}$$

Therefore our update rule is simply

$$\begin{aligned}w_{i+1} &= w_i + \alpha \cdot \nabla L \\ &= w_i + \alpha \cdot yx_i \cdot \left(1 - \frac{1}{1 + e^{-yw^Tx}}\right)\end{aligned}$$