

Chinese Development Projects and Subnational Poverty

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Abstract

China's emergence as a development financier presents renewed opportunity to study the impact of investment in infrastructure on global poverty and inequality. With a preference for investing in connective infrastructure and willingness to implement projects globally, China's policies on aid delivery are distinct. This is evident from the sectoral composition and geographical location of their development finance. This thesis employs a 2SLS strategy, first developed by Dreher et al. (2019), that exploits differences in local exposures to a common overproduction shock originating in China's steel industry to determine the local average treatment effect of Chinese infrastructure projects. Leveraging two sources of variation to form a shift-share style instrument to examine the short-term effects of Chinese development assistance on poverty levels as measured by the International Wealth Index. Results suggest that infrastructure projects within the transport sector have a significant short-term poverty reducing effect of as high as a 2.8%. Which translates roughly to 18 thousand households lifted out of poverty in an average subnational region. A significant effect that is in line with theoretical considerations and previous empirical studies. This thesis encourages traditional OECD donors to reevaluate the importance of their aid budgets to further global infrastructure development to reduce geographic inequalities.

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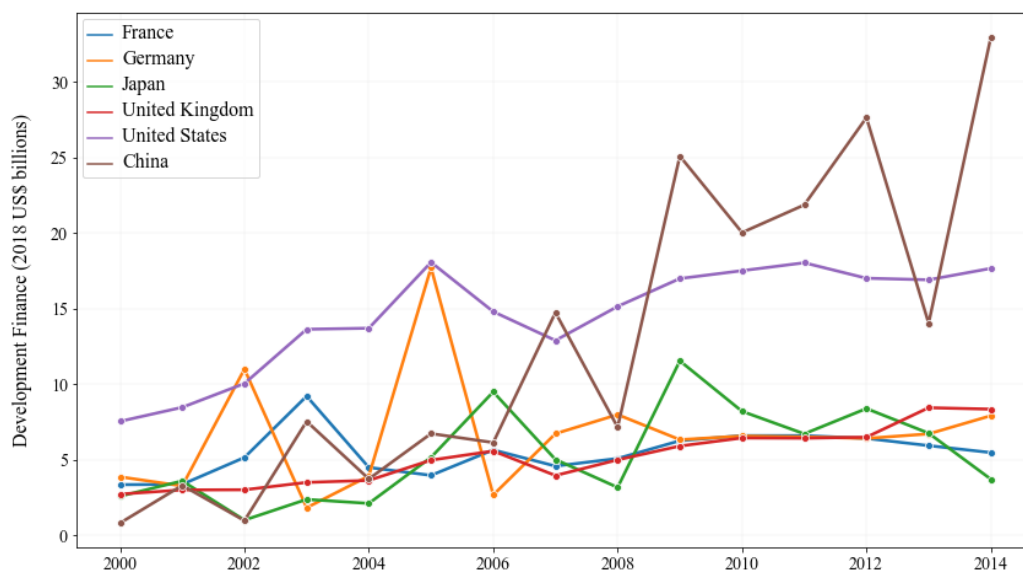
The analysis detailed in this thesis has been made available for reproduction. The repository with all requirements can be found here:

https://github.com/frankstevens1/china_devprojects_poverty#readme

1.0 Introduction

Following the announcement of a ‘Go Out’ policy designed to encourage Chinese firms to invest abroad, the world has seen a rapid expansion in financial transfers from China. The policy, adopted in 2000, fundamentally altered the scale and scope of all forms of state sponsored financial transfers (Djankov, 2016; Dreher et al., 2017) and has positioned China among the largest international development financiers. *Figure 1* shows that development finance from OECD countries flattened out between 2000 and 2014 while China continued to broaden its bilateral financing, marking a shift in the political and economic landscape of international cooperation. A shift that has been met with widespread international attention and has led to the emergence of various examinations of non-traditional donors amongst aid literature that had previously been centred around the investigation of western aid programmes.

Figure 1: Development Finance (ODA & OOF) between 2000-2014



Data sources: OECD & AidData

Early literature shows that China’s aid programme was poorly understood, largely due to the Chinese government’s lack of transparency regarding official flows of development finance (Bräutigam, 2011). A characteristic that led many critics to believe that China’s primary intentions were to smooth access to natural resources and that their aid practices threaten hard-won political reforms by providing support to authoritarian regimes. However, more recently, researchers have challenged the various ill-founded criticisms by making use of the considerable amounts of publicly available information that *do* exist. Revealing that China’s motives, determined by aid allocation patterns, are found to be no different from those of western donors (Dreher & Fuchs, 2015), though policies regarding aid disbursements are clearly distinct. Unlike traditional donors, China’s development assistance is generally a government to government relationship and rarely channelled through non-government organizations nor disbursed in

cash (Brautigam, 2009; Zhang et al., 2015). Additionally, the key focal point of China's aid programme is infrastructure development, and the principal forms of development assistance include complete projects, goods & materials, and technical cooperation (Hwang et al., 2016; State Council, 2014). A priority that was recently echoed by President Xi Jinping during his address at the 2017 Belt and Road Forum for International Cooperation, where he stated that "*infrastructure connectivity is the foundation of development through cooperation*" (Xinhua, 2017). An area that has been neglected by OECD development cooperation in recent years and the reason why Chinese development finance may provide an important counterpoint to finance from traditional donors. Allowing countries to develop much needed infrastructure and to invest in productive activities (Bräutigam, 2011).

To date, the empirical studies on the impacts of Chinese development assistance are mainly centred around economic growth, governance, and trade. However, how poverty and inequality – leading issues in development economics – are affected is not well documented (NguyenHuu & Schwiebert, 2019). The study detailed in this thesis makes use of AidData's global dataset of geocoded Chinese development projects, in combination with the Global Data Lab's International Wealth Index (IWI) to empirically examine the impact of Chinese development projects on global poverty. The methodology follows broadly a strategy developed by Dreher et al. (2019) to examine to what extent Chinese development projects affect economic growth in recipient countries¹. The strategy employs 2SLS instrumental variable regression to estimate the local average treatment effect (LATE) of receiving a Chinese development project. An instrumental variable strategy is used to overcome the expected simultaneity bias that exists because development finance is known to be allocated to less developed – and poorer – regions (Dreher & Fuchs, 2015). The instrument used exploits variation in Chinese steel production, a major input commodity used for infrastructure development, leaning on China's avowed preference to finance infrastructure as a key characteristic of their aid programme. For decades following Mao's Great Leap Forward, steel has been considered of strategic importance to China and excess supply is consistently maintained, much of which is exported and used to supply aid projects (Zhang & Singh, 2020). This results in increases in the provision of Chinese aid projects being observed in years following high steel production. Meanwhile, high correlations between steel production and other input commodities, such as timber or cement, allows for steel to serve as a reasonable proxy for the combined physical inputs for infrastructure projects (Bluhm et al., 2018).

Another constraint faced by researchers dealing with the impacts of development assistance is a lack of disaggregated data, which has largely limited existing empirical studies to the national level. However, large regional disparities raise concerns about discernability of the effects of development projects relative to other economic variations. In other words, at any given time the presence of an

¹ This instrument was originally introduced by (Dreher et al., 2019) in a 2016 working paper, and has since been adopted by subsequent work for varying research questions. Examples include; Bluhm et al., (2018); Gehring & Lang (2018); Humphrey & Michaelowa (2019).

infrastructure project in one part of the country is likely to be a significant source of economic variation within that subnational region while at the national level its significance is less evident. Motivating the use of subnational level analysis where national development indicators may be misleading. Moreover, to examine the effectiveness of development interventions intended to promote the income growth of the poorest implies understanding how successfully these interventions target poor regions *within* recipient countries. The findings of this study suggest that Chinese infrastructure projects within the transport sector have a significant effect in the short term of as high as a 2.8% reduction in subnational poverty levels.

The remainder of this thesis is setup as follows. First, a review of the literature is provided with respect to development finance, unequal spatial development outcomes and how infrastructure projects can diffuse economic activity to less developed regions. What follows is an explanation of the data and methodology, a discussion of the results and finally the conclusions. A separate section following the references contains all relevant tables, including summary statistics and regression tables as well as visual overviews of the global distribution of Chinese development projects. Minor technical issues regarding the regressions are discussed in the appendices based on a residual analysis and a fitted plot.

2.0 Literature Review

This section begins with a discussion of the relevance of subnational investigation when considering global development goals, such as poverty alleviation. Followed by a brief review of literature that examines the relationships between development finance and subnational development outcomes, which highlights the significance of spatial inequalities. Subsequently, theories of unequal spatial development are discussed, which affirms the pivotal role assumed by investment in infrastructure for national and subnational development. Followed by a section that reviews empirical literature relating to infrastructure investments and development outcomes. Finally, China's role as a development financier with a comparative advantage in infrastructure development is discussed, along with the current state of empirical evidence regarding the impacts of Chinese development projects to clarify the gap in the literature this thesis addresses.

2.1 Development Finance & Subnational Development Outcomes

In 2015, the United Nations General Assembly (UNGA) set the Sustainable Development Goals (SDG) to be achieved by 2030 for a better and more sustainable future for all. Among the 17 goals are the eradication of poverty and the reduction of inequalities. Although, reducing inequality is a new addition to UNGA's development goals, the World Bank has been pursuing pro-poor growth for a long time (Coady et al., 2004). The bank's two overarching goals are to end extreme poverty and to promote shared prosperity. Explicitly referring to a reduction of the people living under \$1.90 a day to less than 3% by 2030 and promoting the income growth of the bottom 40% of the population in all countries (Chandra, 2013). It has long been a general agreement among traditional donors that recipient need should be factored into aid allocation decisions to target poverty reduction. Confirmed by empirical findings that traditional donors allocate their aid to countries with lower GDP per capita, particularly multilateral institutions, who are uniquely poverty sensitive (Dollar & Levin, 2006). However, promoting income growth of those at the bottom of the distribution implies strengthening the efforts of development interventions that successfully target the poor *within* recipient countries. Particularly in low- and middle-income countries, where considerable spatial disparities are widely present and regional variations in per capita income seem to increase with development (Kanbur & Venables, 2005b; Kim, 2008). Consequently, recent increases in the availability of disaggregated spatial data has led scholars within the aid allocation literature to analyse the subnational distribution of development projects – an understanding that dramatically improves poverty reducing efforts (Elbers et al., 2007).

Despite poverty reduction being at the top of donor's priorities, subnational aid allocation literature reveals that traditional development finance flows disproportionately to more prosperous regions (Öhler et al., 2017). Measured in terms of asset ownership and housing quality (Briggs, 2017) or proxied by infant mortality, maternal health and malnutrition (Öhler & Nunnenkamp, 2014). Nonetheless, assessing donor allocation of development finance requires consideration of several operational factors that may influence within-country allocation. One of the many factors likely to influence within-country geographical aid allocation is access. The costs of delivering aid to remote

areas is higher and therefore the expected return relative to more accessible areas will be lower. Additionally, while more accessible areas may generally perform better economically, they are also more populous and can therefore have a larger absolute poverty headcount. As a result, investing in accessible rather than remote areas may be deemed more efficient and lead to subnational patterns of aid allocation that differ from those solely based on poverty. Therefore, aid allocation literature is often assessed in parallel with aid effectiveness literature, which typically examines the effects of aid on growth with the most common result being that there is a minor positive effect (Clemens et al., 2012; Dalgaard et al., 2004), which is in turn expected to translate to poverty reduction. However, due to issues of causal inference there remains no clear consensus (Easterly, 2003; Roodman, 2007). Dreher & Lohmann (2015) argue that one reason for a lack of robust findings is the previous literature's focus on country level rather than regional level growth. Development projects may impact growth in targeted areas, though the magnitude of these impacts are less likely to be measurable at the country level while they may be detectable at the regional level.

Dreher & Lohmann (2015) conduct the first subnational study on the effects of aid on growth and highlight the importance of examining aid flows at the regional level. Although their study finds no robust evidence of any effects of World Bank projects on regional economic growth, they do find correlations that are significant and successively stronger at lower administrative levels. However, even if the relevance of studying development outcomes at the subnational level is evident, reliable measures are few and as a result empirical studies of development outcomes beyond economic growth are scarce. Nonetheless, regional growth is clearly relevant because rapid national development is often found to exacerbate rather than reduce spatial inequalities (Kim, 2008). An observation that is documented by burgeoning literature of spatial inequalities of various forms found across Africa, Asia, Latin America the Middle East and Central & Eastern Europe (Kanbur & Venables, 2005a; Kanbur, Venables, & Wan, 2006). This inequitable spatial distribution of development outcomes implies growth may in fact not translate equitable reductions in poverty or improvements in living standards, as is customarily expected. Therefore, the way in which development projects reduce – or reinforce – existing inequalities of spatial development should be considered when studying their impact. In the next section we turn to a theoretical overview of spatial inequality.

2.2 *Theories of Spatial Inequality*

Neoclassical economics suggest that regional disparities are equalized as a country develops, through a process of convergence occurring via factor mobility and diffusion. An idea that has heavily influenced development economics, as well as the study of regional inequality. Broadly, these theories view regional growth as a process of resource allocation, stressing the importance of mobility of supply side factors, i.e. capital and labour. They assume that labour moves to more developed regions in pursuit of higher wages, while capital moves to more labour-intensive sectors – found in less developed regions – in pursuit of higher profits. A process that ultimately leads to a convergence in regional development (Wei, 2015). In other words, interregional inequality is the result of temporary disequilibrium conditions

in the supply and demand of mobile factors, which equalize in the long run. A process infamously modelled by Kuznets' inverted-U model of income inequality and economic growth, which argues that regional inequality tends to rise before it reduces as economies mature over time (Kuznets, 1966).

As economics has become more plural over time, criticisms of neoclassical theories have gained more acceptance by the mainstream and the processes of development clearly remain disputed. Criticism varies, though it has often been directed at the neoclassical models' assumptions – particularly free factor mobility – and inability to account for short term changes, open economies or observed patterns of regional divergence (Armstrong & Taylor, 2000). Richardson (1978) criticizes the assumption of factor mobility and suggests that factors can be '*sticky*' because in many countries their mobility is impeded. Immobility can be the result of family and social ties, regional variations in the cost of living or the costs involved with moving. These causes of immobility can reasonably be expected to be particularly relevant for unskilled labour in the setting of low-income countries. Where a lack of employment opportunities is a national phenomenon, which increases the risk involved with migration. Therefore, a convergence in regional development may not occur naturally as suggested by neoclassical theory. Above all, additional factors – cultural, geographical and institutional – are also argued to influence factor mobility and regional development (Krugman, 1991; Porter, 1990).

Criticism of once conventional theories of development led to various evolutions in the field of development economics, one of which was the so-called '*geographical turn*'. An evolution that was characterized by a revitalized study of the role of location, physical geography and local conditions in determining regional inequality (Wei, 2015). For example, an important reason for the persistence of poverty and regional disparities in Africa is believed to be due to poor access and availability of local infrastructure (Christiaensen et al., 2003; Sachs, 2006). Theories of new economic geography suggest that investment in improving local conditions – including transportation, resources and labour market conditions – through localized infrastructure projects can reduce geographical impediments to development by improving regional linkages, providing access to markets and fostering employment opportunities (Gannon & Liu, 1997; Rodríguez-Pose & Hardy, 2015). It is through these mechanisms that donor funded infrastructure projects are expected to reduce regional poverty and foster convergence as countries develop. However, empirical studies of the relationship between localized infrastructure projects and subsequent regional development outcomes remain limited due to a lack of disaggregated data. Existing studies are either conducted using national level data or narrowed down to the study of a single country. The following section turns to the evidence presented by some of these studies.

2.3 *Infrastructure Development & Subnational Development Outcomes*

The importance of investing in transport infrastructure is particularly well documented by literature. Development of rural roads is found to increase access to markets (Mu & van de Walle, 2011; Parada, 2016) and to improve vulnerable households' productive capacity (Jacoby, 2000). Road infrastructure lowers transportation costs and in turn the cost of inputs which leads to higher production, higher

household income and increased regional employment (Khandker et al., 2009). Additionally, Khanna, (2014) point out the spillover effects of transit networks to geographically proximate regions which lead to convergence in incomes across regions. Baum-Snow et al. (2017) examine these effects in China and find that ring road investments diffused 50% of industrial GDP from urban areas to outlying areas. On the other hand, researchers have found that the type of road investments can define local gains. Though investment in local rural roads that join villages to towns can maximise rural potential, highways improving connectivity with the urban core are found to have no effect (Dercon & Hoddinott, 2011; Fan & Chan-Kang, 2005) or even reduce rural firms' competitiveness (Start, 2001).

To mitigate the negative externalities of increased urban connectivity, De Ferranti et al. (2005) advocate coordinating such investments with investments in education, health and the provision of credit. Which, is also shown to be an effective strategy for diffusing economic activity as it nurtures the development of local markets and reduces local unemployment (Isserman & Rephann, 1995; Meng, 2013). Therefore, development banks globally are scaling up efforts to improve financing to infrastructure as the deficit remains large in low- and middle-income countries, inhibiting global poverty reduction. The Asian Development Bank estimates that emerging Asian economies have an infrastructure deficit amounting to \$1.7 trillion annually, which is required to maintain growth and tackle poverty. Similarly, the African Development Bank estimates the African continent's current infrastructure needs lie between \$130 and \$170 billion per year. Infrastructure gaps that China is willing and able to finance as well as implement while traditional lenders are deterred by poor credit ratings or slowed down by excessive bureaucratic processes.

2.4 Chinese Development Finance & Infrastructure Projects

During the 1980s, China opened its economy to foreign investment and experienced rapid economic growth along with sharp increases in interregional inequality (Kanbur & Zhang, 2005). In response the government adapted policies of regional inclusion that redirected private and public investment to the economically challenged central and western China, a policy known as the "Develop the West Campaign" (Bluhm et al., 2018). Research found that the campaign helped to combat the trend of increasing inequality across regions (Huang & Wei, 2016). China's own experience with regional inequality are reflected by their foreign assistance policy's emphasis on investing in complementary social and productive sector projects in regions where transportation projects are implemented (Li et al., 2013). Taken together with previously discussed theoretical considerations, China may not only be the only willing financier of physical infrastructure, but also the most suitable to reduce regional inequality and poverty. This is based on their avowed preference for; (1) promoting connectivity through infrastructure and (2) complementary projects that promote agglomeration economies to reduce regional disparities. Additionally, their comparative advantage in rapidly implementing infrastructure projects is reflected by governments around the world preferring Chinese aid as it is perceived to materialize quicker than that of traditional donors (Swedlund, 2017). Therefore, considering infrastructure projects are effective investment for poverty reduction through the reduction spatial

inequality and regional spill overs, transport projects financed by China are expected to have discernible regional poverty reducing effects. Empirically supported by Dreher et al. (2017) who find that Chinese development projects have a positive effect on economic growth and Bluhm et al. (2018) who find that investments in transport infrastructure lead to a more equal distribution of economic activity.

3.0 Data

The dataset used combines a poverty measure derived from an asset based wealth index with AidData's recently released geocoded dataset of Chinese government-financed development projects and steel production figures sourced from The World Steel Association's Statistical Yearbooks [2000, 2010, 2018, 2019]. The dataset covers a 15-year period from 2000 to 2014, this is the period for which project commitments are observed in AidData's dataset. *Table 9 on page 35* presents the summary statistics of the dataset.

The International Wealth Index (IWI) is an asset-based wealth index maintained by the Global Data Lab (GDL) and is the product of aggregating data gathered from various household surveys. These surveys are conducted in waves at varying intervals and administrative levels, which are not always the same official administrative subdivisions. Additionally, there can be changes of subdivisions over time, therefore for this study the subnational coding system developed by GDL, *gdlcodes*, is used as the unit of analysis. Fortunately, AidData's geocoding effort provides geolocation details for most Chinese development projects included in their dataset. The precision of the geolocation details vary from within a 25 km radius from the project site to first-order administrative regions. All the development projects can therefore be aggregated precisely to *gdlcodes* using GDL's publicly available shapefiles. A small number of project points lie marginally outside of polygons (*gdlcodes*), this is expected to be due to measurement error and/or offshore projects. These projects are assigned to their nearest polygon. The resulting sample includes 4,077 unique project locations over the 15-year period with a total estimated value of US\$162 billion. Projects are located within 617 unique subnational regions in 103 countries and annually there are a minimum of 86, average of 271 and maximum of 411 projects. The classification of commitments across CRS² sectors are shown on *page 26* in *table 1* and across CRS flow classes in *table 2*. Projects are located in Africa, Asia, Latin America, the Middle East, and Eastern Europe. *Figures 5 & 6* map project locations according to CRS sectors and flows, respectively.

3.1 The International Wealth Index

The IWI is the first asset based index of household material well-being that is comparable across place and time that is applicable to all low and middle income countries (Smits & Steendijk, 2015). The index was initially developed by employing data on 2.1 million household's possession of consumer durables, access to facilities and housing quality from 165 household surveys conducted between 1996 and 2011 in 97 low- and middle-income countries. Through principle component analysis, asset weights are derived to construct the IWI formula. The resulting index is a stable and easily interpretable index with which the economic situation of households can be evaluated and compared across all regions of the developing world. The position on the index indicates the extent to which households own basic assets

² Common reporting standard (CRS) used by participating donors to report their aid flows to the Development Assistance Committee (DAC) databases. China and many other non-DAC donors do not participate, but AidData researchers have classified Chinese financial flows according to these standards.

that are universally valued, including consumer durable such as a mobile phone, housing characteristics such as roofing quality and access to facilities such as running water. The index runs from 0 to 100, where 0 indicates the household owns none of the assets and 100 indicates all the assets are owned. Poverty measures for subnational regions can be derived by aggregating household IWI scores and determining the percentage of a region's population below a given IWI level. Smits & Steendijk (2015) tested the usefulness of the index as a poverty measure and found that the percentage of populations below an IWI of 50 (*iwipov50*) is strongly correlated with the commonly used poverty headcount ratio of people living below \$2.00 a day (Smits & Steendijk, 2015). Therefore, *iwipov50* can be interpreted as a reliable measure of poverty levels of subnational regions as it is based on an aggregation of household level wealth.

The Global Data Lab (GDL) at Nijmegen's Radboud University in the Netherlands updates a dataset of IWI scores as household surveys are released and makes IWI data publicly available alongside a variety of other subnational indicators. The dependant variable used in this thesis is *iwipov50_{irt}*. The percentage of households in region r in country i below an IWI value of 50 in year t . This indicator is only available for years in which a survey was conducted, which differs per country therefore the dataset contains gaps. To reduce gaps *iwipov50* linearly inter- and extrapolated to 3 years.

3.2 *AidData's Geocoded Global Dataset of Chinese Official Finance*

Details of the Chinese government's development financing program are not systematically made publicly available at the project-level. However, the AidData research lab, at the College of William & Mary in Williamsburg, Virginia, has developed a novel dataset that captures a total of 3,485 projects worth US\$ 274 billion. Most of the projects are geocoded with points at varying regional levels within countries, resulting in a total of 6,190 discrete project locations in 138 countries over the period of 2000-2014 in Africa, the Middle East, Asia and the Pacific, Latin America and the Caribbean, and Central and Eastern Europe. The underlying data is collected using an open source method called Tracking Underreported Financial Flows (TUFF), for details see Strange et al. (2017). Projects can be disaggregated by sector and flow class as projects are classified according to the common reporting standard developed by the Development Assistance Committee (DAC).

$CnAid_{irt}$ is the explanatory variable of interest derived from AidData's dataset, which indicates a project commitment made in year t destined for sub-national region r in country i and is classified by CRS sector & flow type. The variable can take 3 forms: as a dummy variable indicating the presence of 1 or more projects, a count indicating the number of projects, or the dollar value of project commitments. The preferred form to determine the local average treatment effect (LATE) is the binary form that equals 1 when a region received the treatment of a Chinese development project in year t or equals 0 when the region did not. This is because financial values of the commitments per project location are near impossible to estimate through open source methodology (Bluhm et al., 2018). Nonetheless, the effects of development projects are also expected to vary across the intensive margin

therefore additional regressions using project value and project count are also presented. However, these results should be interpreted with caution.

Additionally, the dummy is used to construct p_{ir} , the probability of region r in country i receiving a project. This is calculated as the fraction of years a region received 1 or more projects over the period 2000-2014. This probability varies only across regions and not over time.

3.3 *World Steel Association Statistical Yearbooks*

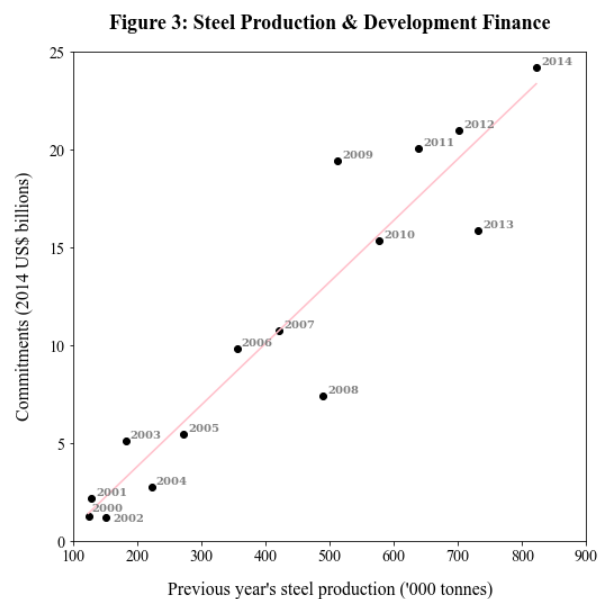
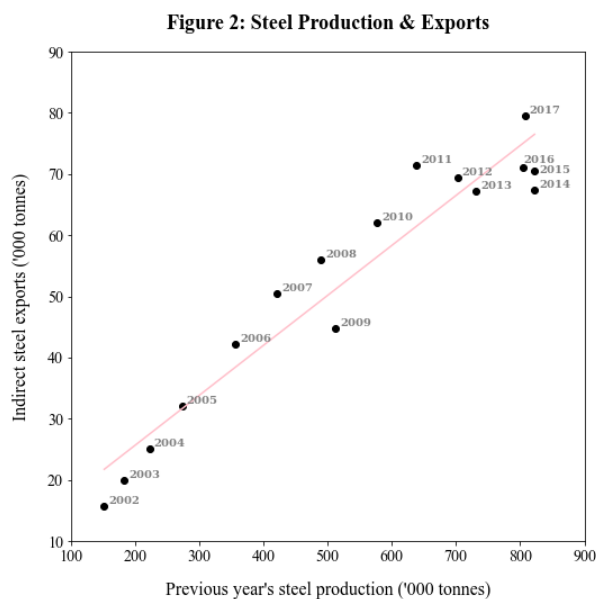
These yearbooks presents a cross-section of steel industry statistics including production and trade. Yearbooks 2000, 2010, 2018 & 2019 were used to compile a time-series of variable $steel_t$; the Chinese steel industry's total production of crude steel in year t measured in thousand tonnes. Indirect steel exports was also extracted to demonstrate the relationship between Chinese steel overproduction and steel exports in the following year, see *figure 2*.

4.0 Empirical Strategy

To determine the effect of Chinese development assistance on subnational poverty levels, the dependant variable of interest is the percentage of the region's population living under the poverty line. Smits & Steendijk (2015) have demonstrated the effectiveness of *iwipov50* in measuring poverty, with an impressive Pearson's correlation of 0.914 between national percentages of households with an IWI value below 50 and the World Bank's Poverty Headcount Ratio at \$2.00 a day. Therefore, as the poverty line we use IWI 50, and examine the effects of regions receiving a Chinese development project through the change in the percentage of households below this level. The main explanatory variable of interest would ideally be a dollar value committed per region per year, unfortunately comprehensive estimates of dollar values committed per project location are near impossible to estimate (Bluhm et al., 2018). Individual projects often fall under an umbrella investment and AidData then splits these data evenly across project locations. For this reason, a binary variable indicating the presence of development projects is preferred. The main results can be interpreted as the local average treatment effect (LATE) of an additional Chinese development projects on regional poverty. Estimated project values are used as the explanatory variable in a secondary regression, which can be interpreted as a robustness check of the significance, direction, and cautiously the magnitude of the effect.

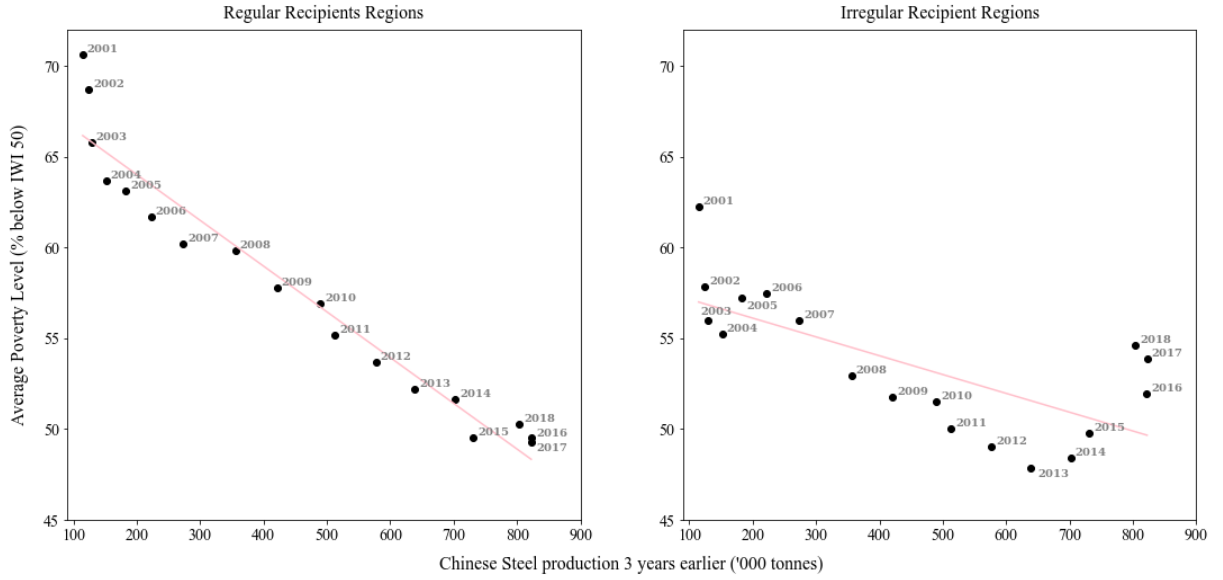
Unfortunately, the explanatory variable suffers from a simultaneity bias, because in order for the subnational allocation of development projects to "*make great efforts to ensure benefits for as many needy people as possible*" (State Council, 2014) it is bound to be based on regional poverty levels. The empirical strategy addresses this issue, following Bluhm et al. (2018) who employ the two stage least squares (2SLS) approach, initially developed by Dreher et al. (2019), with a shift-share style instrument to study the impact of Chinese infrastructure projects on the diffusion of economic activity in developing countries. The strategy exploits two sources of variations. First, the plausibly exogenous time variation in Chinese steel production, which is a supply-driven component of the explanatory variable of interest – Chinese infrastructure development assistance (*CnAid*). China's steel production has long been considered as a strategic industry and drove their rapid economic growth and as domestic demand slows the industry is producing surpluses which are driving exports, see *figure 2*. Chinese overproduction of steel leads to increased exports in the following year, much of which is destined to foreign infrastructure projects. Some of which are financed by China, whose development assistance does not usually involve bilateral financial transfer but are commonly delivered in kind, as exports of Chinese goods & materials, complete projects, export credits and technical cooperation (Bräutigam, 2011; State Council, 2014; Y. Zhang et al., 2015). Therefore, steel is expected to be a relevant instrument for Chinese development assistance, particularly in the form of infrastructure projects. To demonstrate this relationship *Figure 3* shows the correlation between Chinese steel production and total dollar value of development assistance committed in the following year. The 1-year lag of steel production allows for surpluses to translate to international development projects. Together, *Figures 1*

& 2, show that overproduction in steel leads to greater exports, which is argued to lead to more commitments to large scale infrastructure projects requiring physical inputs.



Chinese steel production only varies over time; however, the second source of variation is the cross-sectional variation in a regions' probability of being a recipient of Chinese development assistance (p_{ir}). Which is measured as the fraction of the number of years over the period 2000 and 2014 in which the region had an active project. Taken together, the interaction of these two sources of variation form a shift-share style instrument that exploits differences in local exposure to a common overproduction shock originating in China. The logic behind this approach is similar to that of a difference-in-difference estimation. In other words, the effects of Chinese development assistance on regional poverty levels induced by domestic overproduction of steel is compared across two groups: regular recipient regions and irregular recipient regions. To see this, we turn to the reduced form shown in *figure 4*. The regions in the sample are divided into two groups based on the frequency of receiving Chinese development projects during the sample period. The sample's median value is used to create 2 similarly sized groups, the regions below the median are considered irregular recipients and those above are regular recipients. The reduced form demonstrates the relationship between Chinese steel production, lagged by an additional 2 years to allow for the completion of development projects, and the average poverty level in regular and irregular recipient regions. There is a strong negative correlation among regular recipients, while that correlation is weak amongst irregular recipients. Naturally, these correlations do not imply causation. To distinguish the effects that are attributable to the presence of Chinese development projects our estimation strategy makes use of a strict set of fixed effects and controls for initial poverty levels. The expectation based on theoretical considerations and the relationship demonstrated by the reduced form in *figure 4* is that there is a significant and negative effect on poverty levels in regions that regularly host a Chinese development project.

Figure 4: Chinese Steel Production & Regional Poverty



This identification strategy relies on the interaction term being exogenous and assumes that a change in Chinese steel production does not lead to a change in the probability of receiving development projects between regular regions and irregular regions. An assumption that appears to be realistic. Additionally, Bluhm et al. (2018) examine this assumption in detail by examining the variation in steel production along variations in the location of projects for different quartiles of probabilities to receive projects and conclude that there is no reason to believe the assumption is violated.

A first limitation of this strategy is that it relies on a large sample for instrument strength, which is tested using F statistics tests of whether the endogenous regressors is under- or weakly identified. The regressions call for robust standard errors therefore the usual approach in the applied literature is to use Kleibergen-Paap F-statistic compared to critical values tabulated by Stock & Yogo (2005) and conclude that instruments are weak if the statistic is below those critical values. Another limitation is that the strategy relies on annual variation and with a dataset that spans only 15 years the analysis is limited to the use of annual data instead of data averaged over several periods. This means that the results will be interpreted as short term effects, unlike typical aid effectiveness literature. A final limitation is the lack of control variables available at subnational level, such as institutional quality. Therefore, the specification relies on multiple fixed effects to absorb a wide variety of confounding sources of variation and isolate the effect of development projects. REGHDFE, a Stata module developed by Correia (2016) is employed to fit regressions while absorbing multiple fixed effects and clustering standard errors at the country level to allow for within country spillover effects. Regional fixed effects absorb variation at unit of analysis to account for a lack of control variables available at the regional level. While year fixed effects absorb shocks that affect all regions of a country similarly in a particular year. Taken together the 1st and 2nd stage equations that are jointly estimated through 2SLS are:

$$CnAid_{ir(t-2)} = \beta_1(\ln(Steel_{t-3}) \times p_{ir}) + iwipov50_{ir(t-1)} + u_{ir} + \lambda_{it} + v_{irt} \quad (1)$$

$$iwipov50_{irt} = \beta_2 CnAid_{ir(t-2)} + iwipov50_{ir(t-1)} + \mu_{ir} + \lambda_{it} + v_{irt} \quad (2)$$

Where equation 2 is the main equation of interest and $iwipov50_{irt}$ is the percentage of households in poverty in region r in country i in year t . $CnAid_{ir(t-2)}$ indicates the presence of a Chinese development project, lagged by 2 years to allow for project completion – based on the average time that projects in the dataset are completed. $\ln(Steel_{t-3})$ is measured as the natural logarithm of metric tonnes of steel produced in China lagged by an additional year to allow for domestic over production to translate into the commitment of international development projects. p_{ir} is the probability that a region receives a project. The initial level of poverty, $iwipov50_{ir(t-1)}$, is included to control for an overall trend in poverty levels. The control is also included in the first stage equation.

The coefficient of the instrument in the first stage equation (β_1) is interpreted as the elasticity of the probability to receive a project induced by changes in Chinese steel production. Consider that the average annual change in steel production between 2000 to 2014 was 13.7% and take the coefficient of the first stage equation (β_1) = 0.349 (see table 4 column 4 on page 30). This average production increase raises the probability of a receiving a project by about 4.7% (0.137×0.349) in regions that always receive projects. While in regions that only receive a project in 15% of all years, the average production increase raises the probability of receiving a project by 0.7% ($0.137 \times (0.349 \times 0.15)$).

The main result of the study lies in the coefficient of the second stage equation (β_2) which is interpreted as the short term local average treatment effect (LATE) on poverty levels of Chinese development projects induced by domestic overproduction of steel. This LATE is for regions that host a project in all years and is therefore considered the upper boundary of the effects on poverty that regions can expect from hosting a Chinese development project.

5.0 Results & Analysis

This section will refer to the main results of the study, which are presented in the tables and figures that follow the references on *pages 26-35*. The following section begins with an analysis of the classification of the Chinese development projects by CRS sectors and flows. Followed by an analysis of the regression results.

Table 1 shows that the sample used covers 4077 unique project locations across a variety of CRS sectors amounting to an estimated total value of US\$ 162 billion, this is 65% of the entire AidData dataset that could be matched to *gdlcodes* (regions as classified by the Global Data Lab) using geolocation details. The largest sector, in terms of project count, is by far transport and storage, which includes the construction and maintenance of road, rail, water and air infrastructure as well as public transport services and transport policy administrative and planning assistance. However, in terms of estimated project value, the energy generation and supply sector makes up the largest sector, followed closely by transport. *Figure 5* shows that projects within the transport sector are most widely distributed across the globe, while the energy sector is more concentrated in Africa.

In terms of project count transport is followed by 2 social sectors: the health and education sector. How infrastructure development within these sectors affects poverty in the short term is not well documented. However, in terms of estimated project value both sectors are relatively small, so effects are not expected to be large in magnitude. Additionally, the instrument of steel production, which is highly correlated with other commodities used in construction (Bluhm et al., 2018), would be driven primarily by the construction of schools and hospitals within these sectors. Which may have positive influences on non-monetary development outcomes such as child maternity and expected educational attainment but any effects on these outcomes is unlikely to materialize in the short term.

The 4th and 5th largest sectors in terms of project count are energy generation & supply and communications. Economic sectors that can plainly be expected to have short term effects on subnational poverty levels, considering that access to public facilities as well as ownership of a mobile phone are considered universally desirable assets and are amongst those that shape the International Wealth Index (IWI). Additionally, literature has shown that the impacts of improved infrastructure in these sectors are pro-poor (Puga et al., 2009; Urban, 2019) and the mechanisms through which they translates to poverty reduction, such as improved productivity, can intuitively be considered short term effects.

Table 2 shows how the sample is broken down by classifications as defined by the Development Assistance Committee (DAC). Official Development Assistance (ODA) is defined as “*flows of official financing administered with the promotion of the economic development and welfare of developing countries as the main objective*” (OECD, 2003). While Other Official Flows (OOF) are defined as flows that do not match the ODA definition and include grants for commercial purposes. The bulk of Chinese projects fall under ODA, however the estimated value of OOF is 3 times that of ODA. *Figure 6* shows

that ODA and OOF are spread equally across Asia, while Latin America receives primarily OOF projects and ODA is concentrated in the eastern region of Sub-Saharan Africa.

Table 3 & 4 present the main results of this thesis, the specifications estimating the effects of all Chinese development projects, those that are classified as ODA and those within the transport sector. First, the OLS models in *table 3 columns 1-3*, foreshadow a negative relationship although they are all insignificant and suffer from an expected simultaneity bias. *Columns 4-6* report the reduced form estimates, in which the instrument of lagged Chinese steel production interacted with the regional probability of receiving a project is regressed directly on the poverty indicator. Similarly, the direction of the effect is negative though the effect size is increased by an order of magnitude. However, the significance of the instruments in the reduced form relationship fall just short of the 10% level at 18.5%, 16% and 11.5% for all, ODA, and transport projects respectively. This is somewhat impressive given the restrictive two-way fixed effects while controlling for initial poverty levels with a lagged dependant variable. Tentatively suggesting that the OLS estimates are indeed upwardly biased due to simultaneity, contingent on the strength of our instrument.

Table 4 columns 1-3 present the results of the joint estimation of the first and second stage, which sees the effect of total and ODA projects remain insignificant. Transport projects, however, remains quantitatively similar to the reduced form estimate and significant at the 5% level. The Kleibergen-Paap F-statistic of 16.79 is just barely above the rule of thumb 10% critical value of 16.38. Indicating that we can reject the null hypothesis that the maximum bias relative to OLS due to a weak instrument is below 10% (Stock & Yogo, 2005) indicating 2SLS is an improved specification. Additionally, the first stage equation in *table 4 columns 4-5* shows that our instrument is significant at the 1% level and strongest for transport projects with a positive relationship of 0.802. Indicating that a 10% annual growth in Chinese steel production translates to an increase in a regular recipient region's probability of hosting a transport project by 8% (0.802×0.10) or by 1.6% in regions that only receive projects in 20% of all years ($(0.802 \times 0.20) \times 0.10$). The first stage coefficient of total projects and ODA projects are both about 0.3, still relevant but clearly less highly correlated. Which is expected because our instrument is driven by large infrastructure projects requiring steel and other physical inputs, such as those in the transport sector. Encouragingly, the first stage results are similar in magnitude and significance to those found by Bluhm et al., (2018) despite a smaller sample size. Considering the strength of the instrument, our results suggest a local average treatment effect of 2.8%. Indicating that subnational regions that receive a Chinese transportation project can expect a regional reduction in poverty levels of around 2.8% two years after the project is committed. Translating to approximately 16.8 thousand households for the average region with a population of 600 thousand households.

Table 5 & 6 present a similar analysis at the national level. The setup is the same, in terms of independent variables, fixed effects and controls used but observations are aggregated at the country level. Naturally implying that our sample size is significantly reduced, which was expected to pose

problems for instrument strength. *Table 5* shows that similar to the subnational estimation OLS is mostly insignificant and the effect sizes are an order of magnitude larger, however the reduced form estimates at the national level have improved significance. However, the first stage in *Table 6* shows that we can only rely on the strength of the instrument for transport projects as it remains positive and significant with a magnitude of 0.532. While the instrument for total projects turns negative and is only significant at the 10% level and the instrument of ODA projects remains only slightly positive but loses significance completely. Additionally, confirmed by the 2SLS joint estimates in *columns 1-2* which show that Kleibergen-Paap F-statistic of total and ODA projects are incredibly low indicating unacceptable maximum bias relative to OLS. Similarly, *table 6 column 3* shows that the instrument for transport has a Kleibergen-Paap F-statistic of 11.52 hovering below the acceptable 10% critical value, indicating that we cannot conclude that our instrument is sufficiently strong. Nonetheless, the effect size of transport projects is 2.6% which is almost identical to that of the subnational level and is significant at the 5% level. This provides confidence for our subnational results and suggesting that countries can expect a 2.6% reduction in household poverty from hosting a transport infrastructure project.

Finally, *table 7 & 8* present secondary regressions at the subnational level. *Table 7* finds that the instrument is relevant and strong for education projects. Suggesting that Chinese projects in this sector may indeed consist largely of construction within that sectors. The instrument strength is shown by a Kleibergen-Paap F-statistic of 69.08 (*table 7 column 2*) and a first stage effect of 0.823 (*table 7 column 6*). The effect of education projects on poverty however is not significant and of a small magnitude (*table 7 column 2*), a result that is expected because projects in education do not theoretically have income enhancing effects in the short term. The effect size of energy projects on poverty shown in *table 7 column 3* is significant and large at -6.4%, which can be explained by the fact that access to electricity is a universally valued commodity that is incorporated in the IWI. This result is encouraging but should be interpreted with caution given that the instrument is considered weak with a Kleibergen-Paap statistic of 8.926, which is well below the 16.38 critical value.

The binary explanatory variable indicating the presence of a Chinese project is replaced with project values and project counts in *Table 8*. The instrument is found to remain relevant with high F-statistics and the magnitude of the effect implies similar findings as the main regression results. A doubling of project value or number of projects in the transport sector is associated with a 0.7% and 0.9% reduction in poverty, respectively. While a 100% increase in ODA is associated with a 0.8% reduction in subnational poverty. However, the results presented in *table 8* can merely provide added confidence for instrument strength and significance within the transport sector.

6.0 *Conclusions*

The 2000s clearly marked a shift in the political and economic landscape of international development cooperation. Spurring an interest in improving the poor understanding of non-traditional donors' policies. A poor understanding which resulted in various unconstructive criticisms from sceptical scholars cautioning about the potential threats of the emergence of China as a major development financier. However, more recent studies revealed that China's motives cannot be distinguished from those of traditional donors as recipient need plays a central role in China's aid programme. More importantly, however, scholars have gained a better understanding of China's distinct foreign aid policies which in itself is a source of variation that can be leveraged to provide further understanding of China's role in international development cooperation.

An avowed preference for financing infrastructure development and willingness to provide expedited implementation sets China apart in its ability to reduce existing inequalities of spatial development, which are reinforced when subnational aid flows are disproportionately allocated to more prosperous regions. To promote the income growth of the poor implies strengthening the efforts of development interventions that successfully target the poor within recipient countries, which can be achieved by improved transport infrastructure. Theories of new economic geography suggest that factor mobility is not a guarantee in all countries and therefore there is an increased need for connective infrastructure to foster convergence in regional development and reduce regional poverty and inequality. These linkages are created through investment in transport infrastructure which improves factor mobility and inclusive development, as was experienced during China's develop the west campaign. The effects of a lack of access and impediments to mobility is experienced even by the workers in the development sector as costs of delivering aid to remote areas is increased and reduces the effectiveness of that aid. By investing in connective infrastructure Chinese development projects are reducing geographical impediments to development by improving regional linkages and diffusing economic activity.

The results of this study estimate that the local average treatment effect of Chinese transport projects is as high as 2.8%. This implies that the result found by Blum et al. (2018) that transport projects diffuse economic activity in regions where they are located is to the benefit of the region's poor. A conclusion that is in line with theoretical considerations that transport infrastructure can benefit the poor by improving regional linkages, access to markets and employment opportunities. This provides support for previous case studies that find that transport infrastructure lowers transportation costs and the cost of inputs which have the potential to increase production and household income and increase regional employment. This study has also highlighted the relevance of subnational investigation when considering global development goals such as poverty alleviation. Transport projects are found to have a poverty reducing effect at both the national and subnational level of up to 2.6% and 2.8% respectively, implying that these investments are fostering inclusive growth within and between countries.

The regressions conducted during this study were kept parsimonious at the subnational level due to few control variables being available at the regional level, however the results of our country level test suggest additional variables can be tested using the strategy proposed. *Table 6 column 7* presents a robustness check by including the logged FDI inflows to countries. One concern is that steel originating from China may not only be driving Chinese investments, but that other foreign investors purchase steel from China for foreign investment projects. However, *table 6 column 7* demonstrates that controlling for FDI inflows leaves the effect largely unchanged. Future research may investigate relaxing the fixed effects by including a larger array controls common in aid effectiveness literature, such as institutional quality, as subnational data becomes increasingly available.

The aim of this study was to fill the gap in literature regarding the effects of China as an emerging development financier on global poverty through a focus on localized infrastructure investments. Recent work has improved our understanding of Chinese aid and highlighted infrastructure as a key characteristic of their aid programme. This presents renewed opportunity to study the impact of infrastructure development, which is an area that has been neglected by OECD donors in recent years while major infrastructure deficits in low- and middle-income countries around the world persist. Therefore, scaling up aid for infrastructure is still a noteworthy policy goal as research is making it overwhelmingly clear that localized infrastructure projects can benefit vulnerable households' welfare through improved productive capacity, employment opportunities, access to markets and a general reduction in the costs of living. Massive infrastructure deficits need to be financed in order to maintain growth and tackle poverty around the world. Traditional OECD donors are encouraged to assume global responsibility in understanding that a foreign aid budget of 0.7% is an insufficient target to reach this major impediment to global advancement.

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Table 1: Project Commitments and Estimated Value by CRS Sector

<i>CRS Sector</i>	<i>Projects</i>	<i>Commitments (2014 US\$ millions)</i>
<i>Transport and Storage</i>	1042	57448
<i>Health</i>	661	1760
<i>Education</i>	456	1194
<i>Energy Generation and Supply</i>	348	62770
<i>Communications</i>	335	3916
<i>Government and Civil Society</i>	244	1995
<i>Emergency Response</i>	230	250
<i>Agriculture, Forestry and Fishing</i>	188	2570
<i>Other Social infrastructure and services</i>	183	4943
<i>Water Supply and Sanitation</i>	157	3967
<i>Industry, Mining, Construction</i>	101	15787
<i>Other Multisector</i>	53	556
<i>Developmental Food Aid/Food Security Assistance</i>	18	7
<i>General Environmental Protection</i>	13	79
<i>Action Relating to Debt</i>	9	147
<i>Support to NGOs and Government Organizations</i>	8	376
<i>Women in Development</i>	8	12
<i>General Budget Support</i>	6	13
<i>Non-food commodity assistance</i>	4	41
<i>Trade and Tourism</i>	4	908
<i>Business and Other Services</i>	4	3056
<i>Unallocated / Unspecified</i>	3	30
<i>Population Policies / Programmes and Reproductive Health</i>	1	< 1
<i>Banking and Financial Services</i>	1	10
Total	4077	162045

Table 2: Project Commitments and Estimated Value by Flow Class

<i>Flow Class</i>	<i>Projects</i>	<i>Commitments (2014 US\$ millions)</i>
<i>ODA-like</i>	2670	37751
<i>OOB-like</i>	784	90896
<i>Vague (Official Finance)</i>	623	33189
Total	4077	161836

Figure 5: *Global Distribution of Development Projects by Top 5 CRS Sectors*

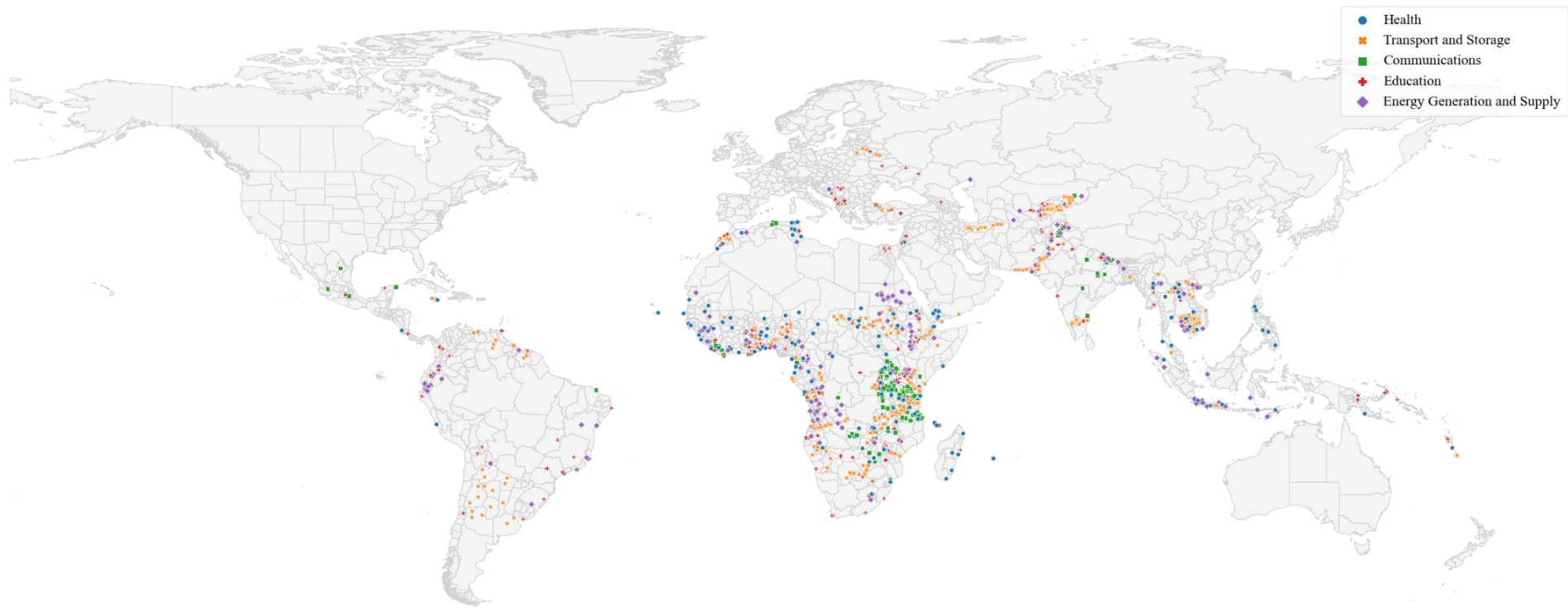


Figure 6: *Global Distribution of Chinese Development Projects by Flow Class*

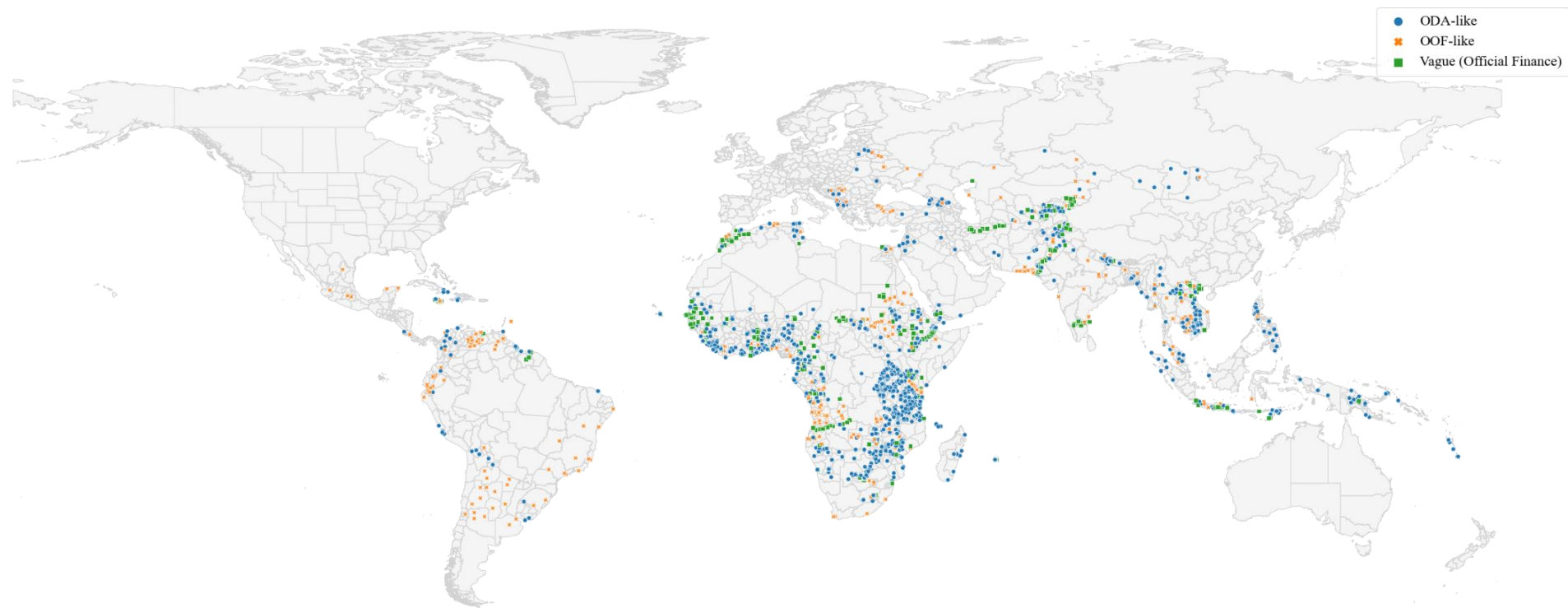


Table 3: OLS & Reduced Form Regressions – Subnational Level

	(1) <i>iwipov50</i>	(2) <i>iwipov50</i>	(3) <i>iwipov50</i>	(4) <i>iwipov50</i>	(5) <i>iwipov50</i>	(6) <i>iwipov50</i>
<i>iwipov50</i> _{<i>t-1</i>}	0.968**** (0.0108)	0.968**** (0.0108)	0.968**** (0.0107)	0.965**** (0.0111)	0.965**** (0.0112)	0.963**** (0.0104)
<i>CnAid</i> _{<i>t-2</i>} (<i>Total</i>)	-0.000260 (0.000860)					
<i>CnAid</i> _{<i>t-2</i>} (<i>ODA</i>)		-0.000999 (0.000877)				
<i>CnAid</i> _{<i>t-2</i>} (<i>Transport</i>)			-0.000840 (0.00293)			
<i>ln(steel)</i> _{<i>t-3</i>} × <i>p</i> _{<i>total</i>}				-0.00433 (0.00330)		
<i>ln(steel)</i> _{<i>t-3</i>} × <i>p</i> _{<i>ODA</i>}					-0.00505 (0.00357)	
<i>ln(steel)</i> _{<i>t-3</i>} × <i>p</i> _{<i>transport</i>}						-0.0227 (0.0142)
<i>Region FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Adj. R</i> ² (<i>within</i>)	0.934	0.934	0.934	0.932	0.932	0.932
<i>Clusters (Country)</i>	102	102	102	101	101	101
<i>Regions</i>	1162	1162	1162	1142	1142	1142
<i>Observations</i>	12146	12146	12146	11422	11422	11422
<i>Estimation Method</i>	OLS	OLS	OLS	OLS	OLS	OLS

Notes: *CnAid* refers to a dummy variable indicating the presence of Chinese development projects within the sector or flow class indicated. Similarly, *p* refers to the probability of receiving a Chinese development project within the indicated sector or flow class. Standard errors clustered by country are in parentheses. Columns 4-6 are reduced form models using the proposed instrument. Significance levels: * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$, **** $p < 0.001$.

Table 4: 2SLS & 1st Stage Regressions – Subnational Level

	(1) <i>iwipov50</i>	(2) <i>iwipov50</i>	(3) <i>iwipov50</i>	(4) <i>CnAid_{t-2} (Total*)</i>	(5) <i>CnAid_{t-2} (ODA*)</i>	(6) <i>CnAid_{t-2} (Transport*)</i>
<i>iwipov50_{t-1}</i>	0.965**** (0.0108)	0.964**** (0.0110)	0.963**** (0.00964)	-0.00983 (0.0785)	-0.0421 (0.0607)	-0.0135 (0.0532)
<i>CnAid_{t-2} (Total)</i>	-0.0125 (0.00967)					
<i>CnAid_{t-2} (ODA)</i>		-0.0155 (0.0116)				
<i>CnAid_{t-2} (Transport)</i>			-0.0283** (0.0142)			
<i>ln(steel)_{t-3} × p*</i>				0.349**** (0.0480)	0.322**** (0.0543)	0.802**** (0.200)
<i>Region FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Kleibergen-Paap F-Statistic</i>	52.62	35.12	16.79			
<i>Clusters (Country)</i>	101	101	101	102	102	102
<i>Regions</i>	1141	1141	1141	1161	1161	1161
<i>Observations</i>	11405	11405	11405	11639	11639	11639
<i>Estimation Method</i>	2SLS	2SLS	2SLS	1 st Stage	1 st Stage	1 st Stage

Notes: *CnAid* refers to a dummy variable indicating the presence of Chinese development projects within the sector or flow class indicated. Similarly, *p* refers to the probability of receiving a Chinese development project within the sector or flow class indicated in the header. Standard errors clustered by country are in parentheses.

Significance levels: * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$, **** $p < 0.001$.

Table 5: OLS & Reduced Form Regressions – Country Level

	(1) <i>iwipov50</i>	(2) <i>iwipov50</i>	(3) <i>iwipov50</i>	(4) <i>iwipov50</i>	(5) <i>iwipov50</i>	(6) <i>iwipov50</i>
<i>iwipov50</i> _{<i>t-1</i>}	0.997*** (0.0157)	0.998*** (0.0157)	0.998*** (0.0154)	0.988*** (0.0171)	0.988*** (0.0170)	0.986*** (0.0169)
<i>CnAid</i> _{<i>t-2</i>} (<i>Total</i>)	0.000717* (0.000405)					
<i>CnAid</i> _{<i>t-2</i>} (<i>ODA</i>)		0.000396 (0.000455)				
<i>CnAid</i> _{<i>t-2</i>} (<i>Transport</i>)			-0.00117 (0.000979)			
<i>ln(steel)</i> _{<i>t-3</i>} × <i>p</i> _{<i>total</i>}				-0.00768*** (0.00288)		
<i>ln(steel)</i> _{<i>t-3</i>} × <i>p</i> _{<i>ODA</i>}					-0.00793*** (0.00272)	
<i>ln(steel)</i> _{<i>t-3</i>} × <i>p</i> _{<i>transport</i>}						-0.0147** (0.00643)
<i>Country FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Adj. R</i> ² (<i>within</i>)	0.970	0.970	0.970	0.970	0.971	0.971
<i>Clusters (Country)</i>	101	101	101	100	100	100
<i>Observations</i>	1103	1103	1103	1041	1041	1041
<i>Estimation Method</i>	OLS	OLS	OLS	OLS	OLS	OLS

Notes: *CnAid* refers to a dummy variable indicating the presence of Chinese development projects within the sector or flow class indicated. Similarly, *p* refers to the probability of receiving a Chinese development project within the indicated sector or flow class. Standard errors clustered by country are in parentheses. Columns 4-6 are reduced form models using the proposed instrument. Significance levels: * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$, **** $p < 0.001$.

Table 6: 2SLS & 1st Stage Regressions – Country Level

	(1) <i>iwipov50</i>	(2) <i>iwipov50</i>	(3) <i>iwipov50</i>	(4) <i>CnAid_{t-2} (Total*)</i>	(5) <i>CnAid_{t-2} (ODA*)</i>	(6) <i>CnAid_{t-2} (Transport*)</i>	(7) <i>iwipov50</i>
<i>iwipov50_{t-1}</i>	0.989**** (0.0232)	0.966**** (0.0842)	0.995**** (0.0122)	-0.287 (0.329)	-0.363 (0.311)	0.256 (0.437)	0.993**** (0.0133)
<i>CnAid_{t-2} (Total)</i>	0.0411* (0.0243)						
<i>CnAid_{t-2} (ODA)</i>		-0.200 (0.417)					
<i>CnAid_{t-2} (Transport)</i>			-0.0262*** (0.00880)				-0.0235*** (0.00885)
<i>ln(steel)_{t-3} × p*</i>				-0.190* (0.0976)	0.0174 (0.0777)	0.532*** (0.164)	
<i>ln(FDI)_{t-2}</i>							0.000838 (0.000691)
<i>Country FE</i>	Yes	Yes	Yes	Yes	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	Yes	Yes	Yes	Yes	Yes
<i>Kleibergen-Paap F-statistic</i>	3.546	0.246	11.04				12.68
<i>Clusters (Country)</i>	100	100	100	101	101	101	99
<i>Observations</i>	1039	1039	1039	1056	1056	1056	974
<i>Estimation Method</i>	2SLS	2SLS	2SLS	1 st Stage	1 st Stage	1 st Stage	2SLS

Notes: *CnAid* refers to a dummy variable indicating the presence of Chinese development projects within the sector or flow class indicated. Similarly, *p* refers to the probability of receiving a Chinese development project within the sector or flow class indicated in the header. Standard errors clustered by country are in parentheses.

Significance levels: * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$, **** $p < 0.001$.

Table 7: Largest Sectors Following Transport – Subnational Level

	(1) <i>iwipov50</i>	(2) <i>iwipov50</i>	(3) <i>iwipov50</i>	(4) <i>iwipov50</i>	(5) <i>CnAid</i> _{<i>t-2</i>} (<i>health</i> *)	(6) <i>CnAid</i> _{<i>t-2</i>} (<i>education</i> *)	(7) <i>CnAid</i> _{<i>t-2</i>} (<i>energy</i> *)	(8) <i>CnAid</i> _{<i>t-2</i>} (<i>communications</i> *)
<i>iwipov50</i> _{<i>t-1</i>}	0.966**** (0.0116)	0.966**** (0.0117)	0.961**** (0.0105)	0.966**** (0.0116)	0.0445 (0.0337)	-0.0397 (0.0274)	-0.0363 (0.0296)	0.0174 (0.0199)
<i>CnAid</i> _{<i>t-2</i>} (<i>health</i>)	-0.000675 (0.0128)							
<i>CnAid</i> _{<i>t-2</i>} (<i>education</i>)		-0.000790 (0.00909)						
<i>CnAid</i> _{<i>t-2</i>} (<i>energy</i>)			-0.0646** (0.0276)					
<i>CnAid</i> _{<i>t-2</i>} (<i>communications</i>)				-0.0354 (0.0397)				
<i>ln(steel)</i> _{<i>t-3</i>} × <i>p</i> *					0.448**** (0.114)	0.823**** (0.0990)	0.689*** (0.231)	0.637* (0.377)
<i>Region FE</i>	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
<i>Kleibergen-Paap F-Statistic</i>	15.68	69.08	8.926	2.844				
<i>Clusters (Country)</i>	101	101	101	101	102	102	102	102
<i>Regions</i>	1141	1141	1141	1141	1161	1161	1161	1161
<i>Observations</i>	11405	11405	11405	11405	11639	11639	11639	11639
<i>Estimation Method</i>	2SLS	2SLS	2SLS	2SLS	1 st Stage	1 st Stage	1 st Stage	1 st Stage

Notes: *CnAid* refers to a dummy variable indicating the presence of Chinese development projects within the sector or flow class indicated. Similarly, *p* refers to the probability of receiving a Chinese development project within the sector indicated in the column header. Standard errors clustered by country are in parentheses.

Significance levels: * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$, **** $p < 0.001$.

Table 8: 2SLS with Project Values & Counts – Subnational Level

	(1) <i>iwipov50</i>	(2) <i>iwipov50</i>	(3) <i>iwipov50</i>	(4) <i>iwipov50</i>	(5) <i>iwipov50</i>	(6) <i>iwipov50</i>
<i>iwipov50_{t-1}</i>	0.964*** (0.0103)	0.963*** (0.00948)	0.963*** (0.00983)	0.965*** (0.0109)	0.965*** (0.0111)	0.966*** (0.0102)
<i>ln(\$CnAid_{t-2})_{total}</i>	-0.00406 (0.00303)					
<i>ln(\$CnAid_{t-2})_{ODA}</i>		-0.00819 (0.00507)				
<i>ln(\$CnAid_{t-2})_{transport}</i>			-0.00729* (0.00398)			
<i>CnAid_{t-2} (Total)</i>				-0.00287 (0.00210)		
<i>CnAid_{t-2} (ODA)</i>					-0.00437 (0.00309)	
<i>CnAid_{t-2} (Transport)</i>						-0.00907* (0.00462)
<i>Region FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Year FE</i>	Yes	Yes	Yes	Yes	Yes	Yes
<i>Kleibergen-Paap F-Statistic</i>	15.01	46.27	32.57	46.30	55.73	17.11
<i>Clusters (Country)</i>	101	101	101	101	101	101
<i>Regions</i>	1141	1141	1141	1141	1141	1141
<i>Observations</i>	11405	11405	11405	11405	11405	11405
<i>Estimation Method</i>	2SLS	2SLS	2SLS	2SLS	2SLS	2SLS

Notes: \$CnAid refers to a variable indicating the value of Chinese development projects within the sector or flow class indicated. CnAid refers to a count variable indicating the number of Chinese development projects. Similarly, *p* refers to the probability of receiving a Chinese development project within the indicated sector or flow class. Standard errors clustered by country are in parentheses. Significance levels: * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$, **** $p < 0.001$.

Table 9: Summary Statistics

<i>Subnational Level</i>	<i>count</i>	<i>mean</i>	<i>sd</i>	<i>min</i>	<i>max</i>
<i>iwipov50</i>	14038	.5135949	.3708452	0	1
<i>CnAid_{total} (dummy)</i>	17953	.1044394	.3058382	0	1
<i>CnAid_{ODA} (dummy)</i>	17953	.0773687	.267183	0	1
<i>CnAid_{transport} (dummy)</i>	17953	.0206094	.1420766	0	1
<i>CnAid_{total} (count)</i>	17953	.2259789	.9890838	0	32
<i>CnAid_{ODA} (count)</i>	17953	.1482761	.7107718	0	19
<i>CnAid_{transport} (count)</i>	17953	.0580404	.5789651	0	29
<i>steel</i>	17953	467758.1	234911.8	128500	822306
<i>ln(steel)</i>	17953	12.89472	.6072636	11.76368	13.61987
<i>p_{total}</i>	17953	.1043948	.1659717	0	1
<i>p_{ODA}</i>	17953	.0773018	.1493906	0	.9333333
<i>p_{transport}</i>	17953	.0205202	.0525048	0	.5333333
<i>p_{health}</i>	17953	.0276425	.0870494	0	.7333333
<i>p_{education}</i>	17953	.0182439	.0533083	0	.4
<i>p_{energy}</i>	17953	.0116712	.0367014	0	.3333333
<i>p_{communications}</i>	17953	.007605	.0256832	0	.2666667
<i>\$CnAid_{total}</i>	17953	8.757005	107.7993	0	8221.052
<i>\$CnAid_{ODA}</i>	17953	2.102352	26.31476	0	1075.026
<i>\$CnAid_{transport}</i>	17953	3.199897	51.74535	0	2541.69
<i>Country Level</i>	<i>count</i>	<i>mean</i>	<i>sd</i>	<i>min</i>	<i>max</i>
<i>iwipov50</i>	1269	.4832491	.3551277	.0001	.999
<i>CnAid_{total} (dummy)</i>	1540	.5318182	.4991487	0	1
<i>CnAid_{ODA} (dummy)</i>	1540	.4324675	.4955793	0	1
<i>CnAid_{transport} (dummy)</i>	1540	.1168831	.3213853	0	1
<i>steel</i>	1540	468189.2	235235.4	128500	822306
<i>ln(steel)</i>	1540	12.8955	.6077545	11.76368	13.61987
<i>p_{total}</i>	1540	.5317316	.2587079	.0666667	1
<i>p_{ODA}</i>	1540	.4325541	.2834281	0	1
<i>p_{transport}</i>	1540	.1178788	.1391439	0	.6666667
<i>FDI</i>	1492	2.51e+09	7.18e+09	-9.37e+09	9.93e+10
<i>ln(FDI)</i>	1422	19.88227	2.200889	10.43922	25.32134

Appendices

Figure 7 presents the distribution of the poverty variable; it has a bimodal distribution with peaks at the higher and lower end. The study included several regions with low poverty levels; however, these were included in the study as they were still recipients of Chinese projects. Conducting an analysis – not presented – without the regions with an *iwipov50* of 0 had no significant effect on the results.

Figure 8 presents a residual analysis and an actual versus predicted value plot. The residuals are within acceptable bounds and show no patterns indicating misspecification or nonlinearity. However, there are several predicted values (45) that fall below 0, which is realistically not possible. This results in a slight downward bias, which is expected to be due to the binary nature of the explanatory variable with many 0 values. However, removing the observations that yield predicted values below 0, shows that their influence on the results is negligible. Table 10 column 1 presents the results with the observations included and columns 2 with the observations removed. For future research the use of the change in IWI may be better suited for the identification strategy, however the changes in year to year IWI levels are so small that it would require grouping data by 5 year period in order to have sufficient variation. This would lead to a study of the local average treatment effect of Chinese development projects on poverty levels in the long run. However, currently neither IWI data nor Chinese project data is not available for longer periods to allow for this study to be done.

Figure 7: Histogram of Dependant Variable

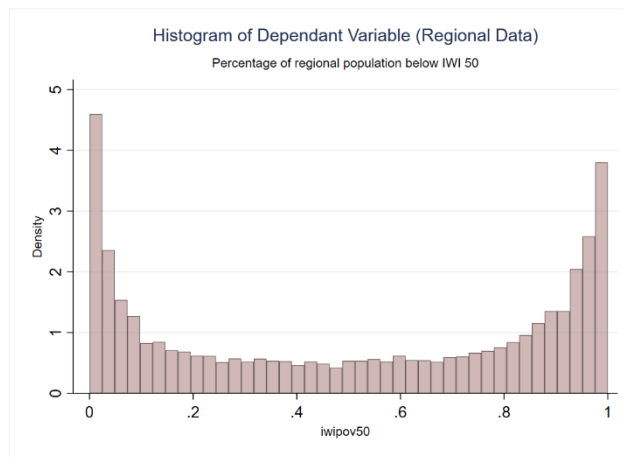


Figure 8: Residual Analysis & Regression Fit

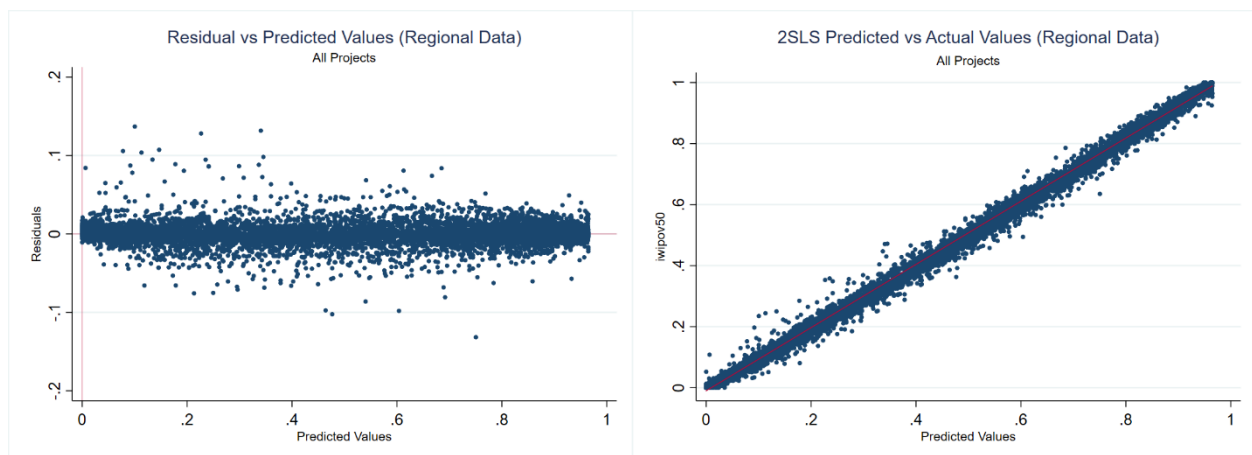


Table 10: Influence of $Y_{hat} < 0$

	(1)	(2)
	<i>iwipov50</i>	<i>iwipov50</i>
<i>iwipov50_{t-1}</i>	0.963**** (0.00964)	0.963**** (0.00963)
<i>CnAid_{t-2}(Transport)</i>	-0.0283** (0.0142)	-0.0282* (0.0142)
<i>Region FE</i>	Yes	Yes
<i>Year FE</i>	Yes	Yes
<i>Kleibergen-Paap F-Statistic</i>	16.79	16.63
<i>Clusters (Country)</i>	101	101
<i>Regions</i>	1188	1188
<i>Observations</i>	11405	11332
<i>Estimation Method</i>	2SLS	2SLS

Standard errors in parentheses

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$, **** $p < 0.001$