- Find a function f : X → Y map input space X to output space Y.
- Predicting a quantitative response Y on the basis of a single predictor variable X.
- ▶ $Y \approx \beta_0 + \beta_1 X$, β_0 and β_1 are model coefficients or parameters.
- Once we have used training data to produce estimates $\hat{\beta}_0$ and $\hat{\beta}_1$, we can make prediction by: $\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 X$.

- ▶ Define $e_i = y_i \hat{y}_i$ be the ith residual (difference between the observed response value and the response value predicted by our model).
- ▶ Residual sum of squares (RSS) ass: $RSS = \sum_{i=1}^{n} e_i$ or

$$RSS = \sum_{i=1}^{n} (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i)^2.$$

Need to choose β_0 and β_1 that minimize RSS. Remember optimality conditions? (convex, differentiable function...)

$$\frac{\partial RSS}{\partial \beta_0} = \sum_{i=1}^n -2(y_i - \beta_0 - \beta_1 x_i)$$

Set this quantity to 0, we get $\sum_{i=1}^{n} y_i = n\beta_0 + \beta_1 \sum_{i=1}^{n} x_i.$

- ▶ Let $\bar{y} = \frac{\sum_{i=1}^{n} y_i}{n}$ and $\bar{x} = \frac{\sum_{i=1}^{n} x_i}{n}$ so $\beta_0 = \bar{y} \beta_1 \bar{x}$
- ▶ Do the same thing with β_1 :

$$\frac{\partial RSS}{\partial \beta_1} = \sum_{i=1}^n 2(y_i - \beta_0 - \beta_1 x_i)(-x_i).$$

We get

$$\sum_{i=1}^{n} y_i x_i = \beta_0 \sum_{i=1}^{n} x_i + \beta_1 \sum_{i=1}^{n} x_i^2$$
$$= (\bar{y} - \beta_1 \bar{x}) \sum_{i=1}^{n} x_i + \beta_1 \sum_{i=1}^{n} x_i^2$$

$$\beta_1 = \frac{\sum_{i=1}^{n} x_i y_i - n\bar{y}\bar{x}}{\sum_{i=1}^{n} x_i^2 - n\bar{x}^2}$$

Notice that: $\sum_{i=1}^{n} (\bar{x}^2 - \bar{x}x_i) = 0$ and $\sum_{i=1}^{n} \bar{x}\bar{y} - y_i\bar{x} = 0$

$$\sum_{i=1}^{n} x_i^2 - n\bar{x}^2 = \sum_{i=1}^{n} (x_i^2 - x_i\bar{x}) + \sum_{i=1}^{n} (\bar{x}^2 - \bar{x}x_i)$$
$$= \sum_{i=1}^{n} (x_i - \bar{x})^2.$$

$$\sum_{i=1}^{n} x_i y_i - n \bar{y} \bar{x} = \sum_{i=1}^{n} (x_i y_i - y_i \bar{x}) + \sum_{i=1}^{n} (\bar{x} \bar{y} - y_i \bar{x}).$$

$$= \sum_{i=1}^{n} (x_i - \bar{x})(y_i - \bar{y}.)$$

$$\beta_1 = \frac{\sum_{i=1}^{n} (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^{n} (x_i - \bar{x})^2}$$

▶ These solutions are least square estimates of parameters β_0 and β_1 .

- ▶ $Y = f(X) + \epsilon$. If f is approximated by a linear relationship: $Y = \beta_0 + \beta_1 X + \epsilon$.
- Assuming the error term is independent of X.
- Least square estimates are unbiased. That means on average, if we do the estimates on many data sets, the average of least square estimates will good.
- ► $SE(\hat{\beta}_0) = \sigma^2 \left[\frac{1}{n} + \frac{\bar{x}}{\sum_{i=1}^n (x_i \bar{x})}\right]$, $SE(\hat{\beta}_1) = \frac{\sigma^2}{\sum_{i=1}^n (x_i \bar{x})^2}$, where $\sigma^2 = Var(\epsilon)$, estimated by $RSE = \sqrt{RSS/(n-2)}$

- Confidence interval for $\hat{\beta}_0$ and $\hat{\beta}_1$ at $1 \alpha\%$: $\hat{\beta}_0 \pm t_{n-2}^{1-\alpha/2} SE(\hat{\beta}_0)$, $\hat{\beta}_1 \pm t_{n-2}^{1-\alpha/2} SE(\hat{\beta}_1)$
- Hypothesis testing on the parameters:

$$H_0: \beta_1 = 0$$

 $H_a: \beta_1 \neq 0$.

$$t-statistic = rac{\hat{eta}_1 - 0}{SE(\hat{eta}_1)}$$

▶ Multiple linear regression model takes the form:

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_p X_p + \epsilon$$

= $X\beta + \epsilon$,

where X is the matrix of predictors X_1, X_2, \dots, X_p , augmented by a column of 1.

$$X = \begin{pmatrix} 1 & X_{11} & X_{12} & \cdots & X_{1p} \\ 1 & X_{21} & X_{22} & \cdots & X_{2p} \\ \cdots & \cdots & \cdots & \cdots \\ 1 & X_{n1} & X_{n2} & \cdots & X_{np} \end{pmatrix}$$

Least squares approach, choose β that minimize the sum of squared residuals:

$$RSS = \frac{1}{2} \sum_{i=1}^{n} (y_i - \hat{y}_i)^2.$$

$$= \frac{1}{2} \sum_{i=1}^{n} (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_{i1} - \hat{\beta}_2 x_{i2} - \dots - \hat{\beta}_p x_{ip})^2.$$

$$= \frac{1}{2} ||Y - X\beta||^2 = \frac{1}{2} \langle Y - X\beta, Y - X\beta \rangle$$

$$= \frac{1}{2} \beta^T X^T X \beta - \langle X^T Y, \beta \rangle.$$

- ▶ Optimality condition: $X^T X \beta X^T Y = 0$ or $X^T X \beta = X^T Y$. (Normal equation)
- ▶ Least square solution: $\beta^* = (X^T X)^{-1} X^T Y$.

- $\beta^* = \min_{\beta} \frac{1}{2} || Y X\beta ||^2.$
- ▶ Think of $X\beta$ as a combination of columns of X: $\beta_1 X^1 + \beta_2 X^2 + \cdots + \beta_n X^p$.
- ▶ Looking for a "combination" of columns of X that is very close to Y.
- \triangleright $X\beta$ is the orthogonal projection of y onto the column space of X.

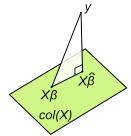


Figure: commons.wikimedia.org

- Simulate a data set with given β , n: number of observations, p: number predictors.
- Gradient descent method for Least square MLR.
- Matlab demonstration.

- Is there a relathionship between response and predictors?
- Hypothesis testing:

$$H_0: \beta_1 = \beta_2 = \cdots = \beta_p$$

 $H_1:$ at least one β_j is nonzero.

F-statistics=
$$\frac{(TSS-RSS)/p}{RSS/(n-p-1)}$$
, where $TSS = \sum_{i=1}^{n} (y_i - \bar{y})^2$. and $RSS = \sum_{i=1}^{n} (y_i - \hat{y}_i)^2$

▶ When there is no relationship between the response and the predictors, F-statistics will take on a value close to 1. Otherwise, F is expected to be greater than 1.

- If conclude that at least one of the parameters is nonzero, which one is it?
- When p is large, looking at individual hypothesis testing can lead to false discoveries.
- ► The task of determining which predictors are associated with the response (we need to fit a model with only these good predictors), is called variable selection.
- Naive approach: Try out all possible combinations of predictors. There are 2^p of them.

- Forward selection: Begin with null model (no predictors). Fit p SLR models and add to the null model that variable with lowest RSS. Repeat.
- Backward selection: Begin with full model with all predictors. Remove the variable with highest p value (least significant). Fit the model with p-1 variables. Repeat.
- Mixed selection: Combine forward and backward selection. Start with null modell, add variables as with forward selection. Remove variables with high p values.

- Computation: Solving normal equation vs. gradient descent method.
- ▶ When p is big, inverting X^TX is not a good idea.
- ▶ When $p \gg n$, X^TX is VERY close to be singular. Finding inverse matrix is hopeless. Gradient descent method is very slow (remember the steplength?).

- model=fitlm(X,y)
- ANOVA: anova(model)
- Confidence intervals: coefCI(model,alpha)