#### Unit 24: Lagged Regression

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## Readings for Unit 24

Textbook chapter 1.4, 5.7.

#### Last Unit

1 Linear Regression with AR errors.

#### Motivation

We'll explore the lagged regression model: used to identify a relationship between two time series.

Bivariate Processes

2 Lagged Regression Model

Worked Example

#### Bivariate Processes

Consider the bivariate time series  $(x_1, y_1), (x_2, y_2), \dots (x_n, y_n)$ . Define the following:

- $\bullet \ \mathsf{E}(\mathsf{x}_t) = \mu_\mathsf{x}, \mathsf{E}(\mathsf{y}_t) = \mu_\mathsf{y}.$
- $\gamma_x(h) = \operatorname{Cov}(x_t, x_{t+h}), \gamma_y(h) = \operatorname{Cov}(y_t, y_{t+h}).$

#### **Cross-Covariance**

The cross-covariance function of two jointly stationary processes  $\{x_t\}$  and  $\{y_t\}$  is

$$\gamma_{xy}(h) = E[(x_{t+h} - \mu_x)(y_t - \mu_y)].$$
 (1)

### Joint Stationarity

**Jointly stationary**: constant means, autocovariances depending only on lag h, cross-covariance depends only on h.

Recall that the autocovariance function is symmetric. The cross-covariance function,  $\gamma_{xy}(h)$ , is not symmetric, i.e.  $\gamma_{xy}(h) \neq \gamma_{xy}(-h)$ . However,  $\gamma_{xy}(h) = \gamma_{yx}(-h)$ .

#### **Cross-Covariance**

- $\gamma_{xy}(h)$ :  $y_t$  is leading  $x_t$ .
- $\gamma_{xy}(-h)$ :  $x_t$  is leading  $y_t$ .

Consider  $x_t$  being the gas input and  $y_t$  the CO2 output of a furnace. The fluctuations of  $y_t$  is delayed with respect to the fluctuations of  $x_t$  due to chemical reaction time for gas to produce CO2.

#### **Cross-Correlation**

The cross-correlation function of jointly stationary  $\{x_t\}$  and  $\{y_t\}$  is

$$\rho_{xy}(h) = \frac{\gamma_{xy}(h)}{\sqrt{\gamma_x(0)\gamma_y(0)}}.$$
 (2)

#### Properties:

- $\bullet \ \rho_{XY}(h) = \rho_{YX}(-h).$
- $|\rho_{xy}(h)| \leq 1$ .

Consider the following processes:  $x_t = w_t + w_{t-1}$ ,  $y_t = x_t - x_{t-1}$ . Derive the cross-covariance function, cross-correlation function, and show that  $\{x_t\}$  and  $\{y_t\}$  are jointly stationary.

## Sample Cross-Covariance and Sample CCF

Sample cross-covariance

$$\hat{\gamma}_{xy}(h) = \frac{1}{n} \sum_{i=1}^{n-h} (x_{t+h} - \bar{x})(y_t - \bar{y})$$

for  $h \ge 0$ . The sample CCF is

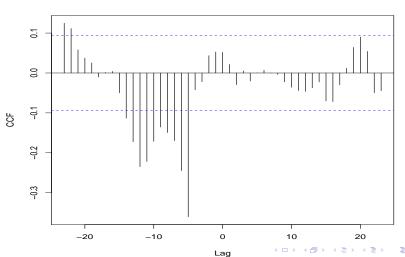
$$\hat{\rho}_{xy}(h) = \frac{\hat{\gamma}_{xy}(h)}{\sqrt{\hat{\gamma}_x(0)\hat{\gamma}_y(0)}}$$

If  $\{x_t\}$  or  $\{y_t\}$  is \_\_\_\_\_, then  $\hat{\rho}_{xy}(h) \sim N(0, 1/n)$ .

# Sample Cross-Covariance and Sample CCF

**Example**: CCF of SOI and recruit data.

#### **CCF** with Prewhitened Data



## Sample Cross-Covariance and Sample CCF

Peak appears at h=-5, this indicates that SOI at time t-5 has strongest correlation with recruitment at time t. SOI leads recruitment by 5 months. The CCF is negative, which tells us that the two time series move in opposite directions: increase in SOI is associated with a decrease in recruitment.

Bivariate Processes

2 Lagged Regression Model

Worked Example

# Lagged Regression Model in Time Domain

We typically consider lagged regression models of the form

$$y_{t} = \sum_{k=1}^{r} \omega_{k} y_{t-k} + \sum_{k=0}^{s} \delta_{k} x_{t-d-k} + u_{t}.$$
 (3)

where  $u_t$  is a stationary ARMA noise process. So we perform a regression on the lagged versions of both the input and output series to obtain the estimates of  $\boldsymbol{\beta} = (\omega_1, \cdots, \omega_r, \delta_0, \delta_1, \cdots, \delta_s)$ .

Due to the large number of parameters we are fitting, the following sequential methodology has been developed. **Step 1**: we fit an ARMA model for the input  $x_t$ , so we have estimates of  $\theta_x(B)$  and  $\phi_x(B)$ .

**Step 2**: prewhiten the input and output series by applying the inverse operator  $\frac{\phi_x(B)}{\theta_x(B)}$  to the input and output series

#### Prewhitening

Recall from slide 13 that we need either the input or the output series to be \_\_\_\_\_\_ so we know the theoretical variance of the sample CCF is 1/n. Thus we prewhiten the input series (and output) so we can study the CCF between the the prewhitened input and output series. Since prewhitening is a linear operation, any linear relationships will be preserved. Note that the operator  $\frac{\phi_x(B)}{\theta_x(B)}$  is tailor-made to transform the input to a white noise, not the output.

**Step 3**: Compute the cross-correlation of  $\widetilde{y}_t$ , the output series after prewhitening, with  $w_t$ ,  $\gamma_{\widetilde{y}w}(h)$  to estimate the time delay d and suggest a form for (3).

**Step 4**: Obtain  $\hat{\beta} = (\hat{\omega}_1, \dots, \hat{\omega}_r, \hat{\delta}_0, \hat{\delta}_1, \dots, \hat{\delta}_s)$  using a regression of the form in (3).

**Step 5**: Fit an ARMA model for the noise  $u_t$ .

Bivariate Processes

2 Lagged Regression Model

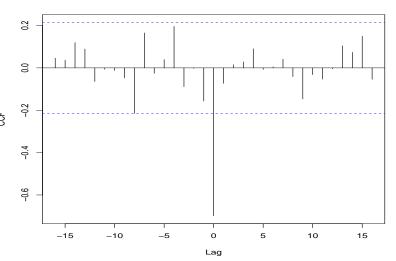
Worked Example

Some of these steps are worked out in some functions in R. What we still need to do is to examine the prewhitened CCF to determine the kind of lagged regression model we should fit, and examine residuals to determine their ARMA structure.

For this worked example, we examine the (log-transformed) sales and price of a certain potato chip from Bluebird Foods. The first step would be to transform the time series to obtain stationarity, and then examine the CCF for the prewhitened data. For this dataset, we take the first difference of both time series to obtain stationarity, and examine the CCF.

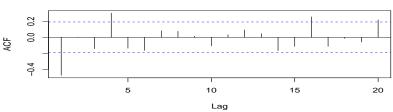
Some common patterns of CCF to look out for.

#### **CCF with Prewhitened Data**

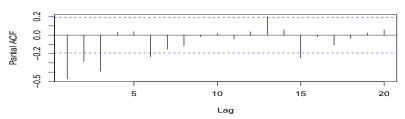


After deciding the appropriate (lagged) regression, fit the model, and examine the ACF and PACF of the residuals to decide their ARMA structure.





#### **PACF of Residuals**



Fit the (lagged) regression model and specify the ARMA structure of the residuals.

```
Call:
arima(x = dy, order = c(3, 0, 4), xreg = data.frame(dx))
Coefficients:
         ar1
                  ar2
                           ar3
                                  ma1
                                           ma2
                                                    ma3
                                                                 intercept
                                                            ma4
     -1.0465 -0.7252 -0.0315 0.2559 -0.0707 -0.7453 0.2096
                                                                   -0.0009
     0.3617
               0.4148
                      0.2862
                                0.3519
                                       0.1730
                                                 0.0933
                                                                    0.0037
s.e.
                                                         0.2977
          dx
     -2.5797
     0.1215
sigma^2 estimated as 0.02502: log likelihood = 40.93, aic = -63.85
```

Any comments?

```
Call:
```

```
arima(x = dy, order = c(2, 0, 3), xreg = data.frame(dx))
```

#### Coefficients:

```
ar1 ar2 ma1 ma2 ma3 intercept dx

-0.0125 -0.9565 -0.7926 0.8786 -0.6680 -0.0010 -2.4473

s.e. 0.0879 0.0687 0.1665 0.1144 0.1138 0.0036 0.1293
```

sigma^2 estimated as 0.02693: log likelihood = 39.12, aic = -64.25

Any comments?

#### Call:

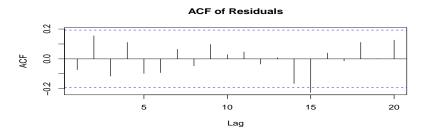
```
arima(x = dy, order = c(2, 0, 3), xreg = data.frame(dx), fixed = c(0, NA, NA, NA, NA, NA, NA, NA)
```

#### Coefficients:

```
ar1 ar2 ma1 ma2 ma3 intercept dx
0 -0.9488 -0.8129 0.8901 -0.6578 0 -2.4510
s.e. 0 0.0641 0.0855 0.0776 0.1109 0 0.1282
```

sigma^2 estimated as 0.02697: log likelihood = 39.07, aic = -68.15

When we think we want to choose a model, make sure to examine the residuals to ensure they appear to be white.



#### **PACF of Residuals**

