

Research



Cite this article: Klimek P, Kreuzbauer R, Thurner S. 2019 Fashion and art cycles are driven by counter-dominance signals of elite competition: quantitative evidence from music styles. *J. R. Soc. Interface* **16**: 20180731. <http://dx.doi.org/10.1098/rsif.2018.0731>

Received: 1 October 2018
Accepted: 10 January 2019

Subject Category:

Life Sciences—Mathematics interface

Subject Areas:

evolution

Keywords:

cultural evolution, evolutionary dynamics, network analysis, fashion cycle theory

Author for correspondence:

Peter Klimek
e-mail: peter.klimek@meduniwien.ac.at

Electronic supplementary material is available online at <https://dx.doi.org/10.6084/m9.figshare.c.4373018>.

Fashion and art cycles are driven by counter-dominance signals of elite competition: quantitative evidence from music styles

Peter Klimek^{1,2}, Robert Kreuzbauer³ and Stefan Thurner^{1,2,4,5}

¹Section for Science of Complex Systems, CeMSIS, Medical University of Vienna, Spitalgasse 23, 1090 Vienna, Austria

²Complexity Science Hub Vienna, Josefstädter Strasse 39, 1080 Vienna, Austria

³Department of Marketing and Retail Management, Surrey Business School, University of Surrey, Guildford GU2 7XH, UK

⁴Santa Fe Institute, 1399 Hyde Park Road, Santa Fe, NM 85701, USA

⁵IIASA, Schlossplatz 1, 2361 Laxenburg, Austria

PK, 0000-0003-1187-6713

Human symbol systems such as art and fashion styles emerge from complex social processes that govern the continuous re-organization of modern societies. They provide a signalling scheme that allows members of an elite to distinguish themselves from the rest of society. Efforts to understand the dynamics of art and fashion cycles have been placed on ‘bottom-up’ and ‘top-down’ theories. According to ‘top-down’ theories, elite members signal their superior status by introducing new symbols (e.g. fashion styles), which are adopted by low-status groups. In response to this adoption, elite members would need to introduce new symbols to signal their status. According to many ‘bottom-up’ theories, style cycles evolve from lower classes and follow an essentially random pattern. We propose an alternative explanation based on counter-dominance signalling (CDS). In CDS, elite members want others to imitate their symbols; changes only occur when outsider groups successfully challenge the elite by introducing signals that contrast those endorsed by the elite. We investigate these mechanisms using a dynamic network approach on data containing almost 8 million music albums released between 1956 and 2015. The network systematically quantifies artistic similarities of competing musical styles and their changes over time. We formulate empirical tests for whether new symbols are introduced by current elite members (top-down), randomness (bottom-up) or by peripheral groups through counter-dominance signals. We find clear evidence that CDS drives changes in musical styles. This provides a quantitative, completely data-driven answer to a century-old debate about the nature of the underlying social dynamics of fashion cycles.

1. Introduction

Of all species, only modern *Homo sapiens* has evolved the ability to use complex symbolic behaviour to organize and coordinate large anonymous societies [1–3]. Symbols facilitate the identification of group membership, social status and, consequently, the competition between elites—here defined as social groups, which have disproportionate access of control over economic, social, cultural, political or knowledge resources [4,5]. In recent times, ‘symbolic social coordination and competition’ are perhaps best captured in fashion and art style cycles (e.g. in music, literature, architecture) [6–10]. There, the social influence of an elite is represented by the number of people who adopt the elite’s core stylistic elements [11,12]. *Costly signalling theory* (ST) has been proposed as an explanation for the social and cultural evolutionary dynamics that determines the changes in fashion

and art styles [3,13]. According to ST, elite members introduce a style which is too costly to adopt for non-members (e.g. a chandelier) and hence provides an honest signal of the elite's superior status. Yet, non-elite members would sooner or later find ways to mimic the elite style (e.g. acquiring inexpensive chandeliers). Elite members would then abandon the style and replace it with a new one (e.g. a modern-style functionalistic pendant light), which again is too costly to adopt for non-members [6]; they are constantly 'on the run' from too many adopters and innovate in reaction to being imitated. Often the new style of the elite would be in stark contrast to the previous style, as shifting to a style highly distinctive from the previous one adds an additional obstacle for the adoption by lower classes [14]. This would explain why consecutive styles often occur as extreme opposites (e.g. ornamentation versus functionalism). However, this 'top-down' mechanism (stylistic elements spread from elite to non-members) has been questioned by others who propose a 'bottom-up' dynamic (elite members adopt stylistic elements from non-members to retain their status) to account for the fact that styles often evolve from lower classes or sub-cultures (e.g. Punk rock or ripped jeans) [8,15–17]. Following this logic, broad stylistic changes are driven by external factors such as symbolic elements chosen and promoted by cultural industries (e.g. trend scouts from fashion houses or music labels who may just randomly pick up symbols from a sub-culture) [18–20]. According to this view, fashion cycles should follow an essentially random pattern. This has been formalized in the so-called random pattern theories (RPTs). For instance, it has been shown that models based on the random copying of cultural traits can explain cyclical patterns of fashion change in a quite robust way; for instance, in the popularity of first names [21–23]. The central assumption in these models is that some styles will become highly popular simply because of imitation but not because they are in some way superior to other styles. However, there exist problems with both the ST and the RPT approach. For example, various 'elite styles' enjoy broad adoption by the middle and lower classes (e.g. ties, black suits, diamond rings) but elite members did not abandon them (as ST would predict). RPT, on the other hand, is unable to account for the fact that two subsequent styles often endorse completely opposite sets of symbols, such as the change from modernism to postmodernism. Instead, we argue that such prominent and enduring cyclical symbolic patterns of mass culture can only be understood by considering the structure of the networks that underlie the processes of social coordination, such as elite competition in the form of counter-dominance signalling (CDS).

To resolve the puzzle of conflicting predictions suggested by ST and RPT, we propose a third mechanism, that of CDS. This mechanism is rooted in cultural evolutionary theories of counter-dominance as well as theories of collective action and social movement [24–27]. Rather than assuming that members of an elite are constantly on the run from too many adopters, in the CDS framework elite members actually *like* to see others adopt their symbols and styles. This signals the cultural influence of elite members [11,24]. Stylistic changes occur when a new upcoming elite successfully challenges the dominance of the established elite by introducing a novel style, which is subsequently adopted by a sufficiently large number of followers. In other words, followers stop adopting the style from the established elite A and start to adopt the style from a new emerging elite B. The new style of elite B should starkly deviate from the previous style, because it

signals and emphasizes its disagreement with the previous hegemony. Examples for such CDS include ostentatiously decorative and non-functional ornaments of postmodern architecture that have the potential to provocatively signal a protest against the dominant elite group such as the 'form-follows-function' doctrine of modernism [24]. CDS serves as a focal point to facilitate the coordination of people who share a reservation against the current dominant elite group [11,24,26].

Our proposed mechanism of CDS allows us to solve a century-old debate on the driving mechanisms behind stylistic changes in areas of mass culture such as art and fashion [6]. On the one hand—consistent with RPT but inconsistent with ST—CDS is able to account for the fact that new styles are not necessarily introduced by elites, but often do evolve through a 'bottom-up' dynamic where elite members adopt specific stylistic elements from non-members. On the other hand—consistent with ST but inconsistent with RPT—CDS predicts patterns of starkly contrasting symbols in consecutive styles. The contrasts arise because outsiders use highly distinctive symbols to counter the dominance of existing elite groups and not because the elite introduced distinctive elements to prevent the masses from adopting them, as predicted by ST.

A fundamental challenge is to empirically test (i) if non-trivial mechanisms are indeed necessary to understand cultural change and, if so, (ii) which mechanism—ST, RPT or CDS—is actually realized in society. Therefore, we first formulate a null model for the evolution of musical styles that assumes that music producers choose the styles which they adopt completely at random; this provides a neutral model for the evolution of styles [28]. We show that such a model is not able to account for the existence of fashion cycles observed in the data of musical styles, which suggests that the evolution of styles is driven by mechanisms that depend on stylistic differences. The neutral model can therefore be modified to include the mechanisms of ST, RPT or CDS, which then indeed leads to the emergence of fashion cycles. Empirical tests are then needed to determine which model mechanism best describes the actual evolution of musical styles. To this end, we develop a method to quantify musical styles by determining each style's typical instrumentation. From a dataset containing almost 8 million albums that have been released since 1950, we extract information about a user-created taxonomy of 15 musical genres, 422 musical styles and 570 different instruments. The instruments that are typically associated with a given genre (or style) were shown to be a suitable approximation to formally describe the characteristics of a style [29]. Therefore, the similarity between styles can be quantified through the similarity of their instrumentation. For instance, in figure 1*a*, we show an example of four different musical styles (blue circles) that are linked to five instruments (green squares). Here a link indicates that the instrument is (typically) featured in a release belonging to that style. The higher the overlap in instruments between two styles, the higher is their similarity and the thicker is the line that connects the styles in figure 1*a*. Using this network representation, we are able to rephrase the question of which mechanism—ST, RPT or CDS—is realized to a question of detecting specific patterns of the network's evolution (figure 1*b*). There, we schematically show three different scenarios for the time evolution of a network (blue circles) that represent the three theories ST, RPT or CDS, respectively. The size of circles indicates the popularity of a given style, i.e. how many artists adopted it. Initially, there

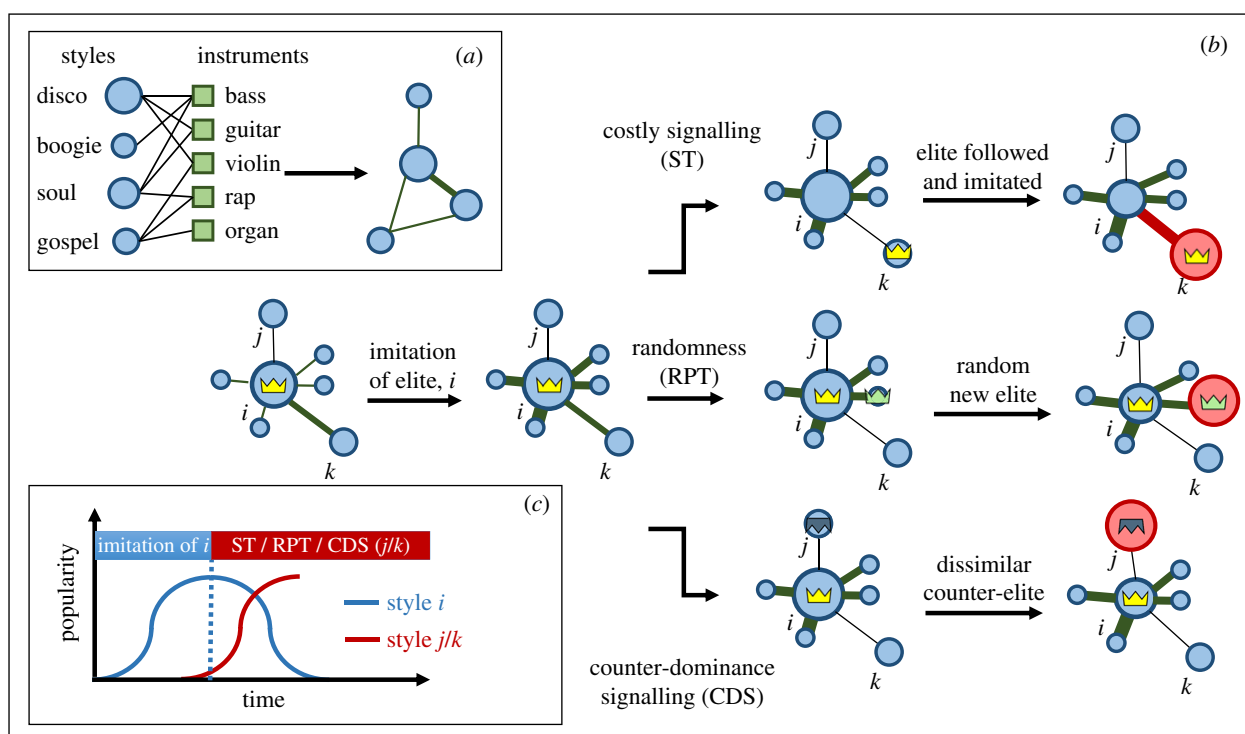


Figure 1. Network evolution for competing theories of cultural change. (a) The characteristics of each musical style (blue circles) are given by the instruments that are typically associated with this style (green squares). The similarities of two musical styles are measured by the number of instruments they share, leading to a style–style similarity network. The size of the circles is proportional to their popularity; the thickness of the link connecting two styles is proportional to their similarity. (b) Competing theories of cultural change imply different types of evolution of the network of musical styles. We consider a network with an elite (yellow crown) that initially adheres to style i . The popular style i will be imitated by other styles (links to i increase in thickness). Following costly signalling theory (ST), the elite seeks to differentiate itself from imitators and adopts a new style, k . Random pattern theory (RPT) suggests that a new elite (green crown) will emerge at a random position in the network. Counter-dominance signalling (CDS) predicts the emergence of a new counter-elite (blue, upside-down crown) that is highly dissimilar to the current elite, shown here for style j . (c) All three theories, ST, RPT and CDS, give rise to fashion cycles in which style i initially increases in popularity under imitation by other styles until a new style emerges through ST, RPT or CDS, and then dominates the next fashion cycle. (Online version in colour.)

exists an elite, whose members have broadly adopted style i . Other styles with lower popularity (non-elite groups) will imitate the style of this elite group, resulting in an increased similarity of other styles with i (increase of thickness of the lines connecting them to i in figure 1b). According to ST, elite members will react to this imitation by adopting a new and different style j . Now it will be style j that the low-popularity groups seek to follow and imitate. This results in the increased popularity of j and increased similarity between j and the prior elite style i . In the RPT scenario, by contrast, it is a new elite that forms at a random position in the network that dominates the next fashion cycle. Finally, the mechanism of CDS suggests that the current elite i will be provocatively challenged by an outsider group, k , using signals (in our case sets of instruments) that are in stark opposition to those endorsed by i . Through the adoption of these signals that oppose i , a new counter-elite emerges with style k . Each of the three mechanisms leads to the same generic cyclic patterns of fashion styles (figure 1c). That is, initially style i increases in popularity as it is imitated by other styles. With the shift of the elite to a different style (ST) or the emergence of a random new (RPT) or counter elite (CDS), new styles gain popularity at the expense of style i and will therefore more frequently be imitated. This imitation triggers the next fashion cycle. By cycles we do not refer to patterns where a single style oscillates in popularity over time, but rather a pattern where certain styles rise sharply in popularity for some time, eventually level off in their growth,

and then fade back into oblivion as new styles emerge that trigger the next cycle. With the help of the proposed network formalism, we can specify concrete hypotheses to test whether the observed changes in similarities and popularities are best described by a network evolution mechanism that is compatible with ST, RPT or CDS.

2. Data and methods

2.1. Data

Discogs is a crowdsourced database of information about audio releases [30]. As of January 2016, it contained entries on 7911 789 albums released between 1952 and 2015. Each album is assigned one or several musical styles. In total, there are 422 musical styles grouped into 15 genres. For each album, Discogs provides a structured list of credits. These credits include a list of artists and the instruments the artists used in the recording. There are 570 different instruments in the database, ranging from lead vocals, finger snaps and countertenor over drums, electric guitars, keyboards and violins, to more exotic instruments such as hunting horns, Northumbrian pipes or MIDI controllers. This data structure allows us to characterize each style by a unique combination of instruments used in recordings associated with that style. Styles and instruments are introduced independently from each other in the data. Discogs employs a moderation system to categorize

styles into taxonomic hierarchies in a systematic way [30]. Styles cannot be freely introduced by users; instead this requires a certain number of releases, the justification for why they are different from currently existing styles and at least three trustworthy external citations of the style's use. It was statistically confirmed that genres indeed correspond to clusters of mutually similar styles [29]. This means that the way in which instruments are assigned to styles is far from random and highly correlated with the independently obtained folksonomic classification of styles. The emergence of a new style in the data therefore means that a stylistic change has occurred that convinced a large-enough group of experts in popular music to introduce a new category to label this kind of music.

2.2. Style popularity

Popularity of a style s is related to the number of albums released within a given interval of time t , in style s , $N_s(t)$. The total number of albums released per year increased from 3051 in 1952 to more than 200 000 after 2012. To account for these large variations, we define the popularity of a style as the relative frequency of releases within a specific time interval, $n_s(t) = N_s(t) / \sum_s N_s(t)$. This notion of popularity quantifies how likely a music producer is to adopt a given style, which is not necessarily the same as the popularity of the style among music consumers. Popularity change of a style, $\delta n_s(t)$, is measured as the per cent change in popularity between year t and year $t + \delta t$.

2.3. Highs, lows and newcomer

In each time interval, we consider the set of m styles with the lowest and highest popularities, which we refer to as *lows* and *highs*, respectively. *Newcomers* are defined as those *lows* that show the highest increases in popularity. That is, we rank styles with at least one release according to their popularity and obtain their popularity rank, $\text{Rank}(n_s(t))$. Low (high) rank values indicate high (low) popularity (rank 1 means the highest popularity). For each time interval t , we identify high-popularity styles as follows. *Highs* are all styles with a popularity rank, $\text{Rank}(n_s(t))$, below m ,

$$\text{Highs}(t, m) = \{s | \text{Rank}(n_s(t)) \leq m\}. \quad (2.1)$$

Lows include the remaining styles that have a popularity rank higher than m . In general, the number of such styles changes substantially from year to year, as the number of styles grows over time. We therefore define the *lows* as m randomly chosen styles with a popularity rank higher than m , i.e. with $\text{Rank}(n_s(t)) > m$. This definition ensures that none of our results are artefacts arising from finite-size fluctuations, i.e. they are driven by size variations of the data sample. All results involving *lows* have been averaged over 1000 random collections of low-popularity styles. For each year, we only consider styles with at least 10 releases in these groups. To ensure that our results are independent of the concrete choice of the threshold m , we carry out robustness tests by letting m vary over a wide range of choices. This test guarantees that our results are not driven by m , but reflect a general feature of style popularities. However, to increase the clarity of the presentation, we will present the main results for a choice of m .

Popularity change of a style, $\delta n_s(t)$, is measured as the per cent change in popularity,

$$\delta n_s(t) = \frac{n_s(t + \delta t) - n_s(t)}{n_s(t)}. \quad (2.2)$$

To identify successful styles among styles that do not yet belong to the *highs*, styles are ranked according to their $\delta n_s(t)$ values. In the following, we refer to the successful lows as *newcomers*, i.e. as those that have a rank of $\delta n_s(t)$ of no more than m but that do not yet belong to the *highs*,

$$\text{New}(t, m) = \{s \notin \text{Highs}(t, m) | \text{Rank}(\delta n_s(t)) \leq m\}. \quad (2.3)$$

If not specified otherwise, we fix $m = 10$ and $\delta t = 5$ years.

2.4. Characterizing styles by instrumentation

The relations between styles and their instruments can be encoded in a time-dependent, bipartite adjacency matrix A , which consists of two types of node, styles and instruments. If there is at least one album released at time t in style s and recorded using instrument i , we set $A_{si}(t) = 1$; otherwise, $A_{si}(t) = 0$. A detailed analysis of this bipartite network can be found in [29]. Some instruments, such as vocals or guitars, are highly ubiquitous and appear in many styles, whereas other instruments are highly specific for a given style, e.g. turntables in hip hop. To suppress the influence of highly ubiquitous instruments, we rescale the contributions of each instrument by its inverse frequency across all styles. This procedure is highly reminiscent of the common use of inverse document frequencies as a weighting factor to increase specificity in information retrieval tasks. This gives us the weighted, time-dependent network, $M(t)$, with entries $M_{si}(t) = A_{si}(t) / \sum_i A_{si}(t)$. The instrumentational similarity (style similarity) between two styles i at t_1 and j at t_2 is defined as the cosine similarity of their instrumentation vectors,

$$\varphi_{ij}(t_1, t_2) = \frac{\sum_s M_{si}(t_1) M_{sj}(t_2)}{|\mathbf{M}_{si}(t_1)| |\mathbf{M}_{sj}(t_2)|}, \quad (2.4)$$

where $|\mathbf{x}|$ denotes the Euclidean norm of vector \mathbf{x} . Note that there are alternative ways to define the style similarity φ_{ij} . Particularly, we will discuss the robustness of our results with respect to the choice of a similarity measure by replacing the cosine similarity in equation (2.4) by the Jaccard coefficient, by the inverse Euclidean distance between the instrumentation vectors and by their similarity as defined by the ProbS algorithm [31,32]. The style–style similarity network for a given time, $\varphi_{ij}(t) = \varphi_{ij}(t_1 = t, t_2 = t)$, is typically fully connected with most of the links having a relatively small weight. Such networks can efficiently be visualized by their maximum spanning tree (MST), the so-called backbone. For a connected network (all nodes are in the giant component) with N links, the MST is the set of $N - 1$ links that have the largest possible weights and that span each node of the original network.

For two sets of styles, $S_1(t_1, m)$ and $S_2(t_2, m)$, the similarity $\Phi(S_1(t_1, m), S_2(t_2, m))$ can be defined as the average similarity of each pair of styles, where one style is chosen from $S_1(t_1, m)$ and the other from $S_2(t_2, m)$. The set similarity $\Phi(S_1(t_1, m), S_2(t_2, m))$ is defined as

$$\Phi(S_1(t_1, m), S_2(t_2, m)) = \frac{1}{m^2} \sum_{i \in S_1(t_1, m), j \in S_2(t_2, m)} \varphi_{ij}(t_1, t_2). \quad (2.5)$$

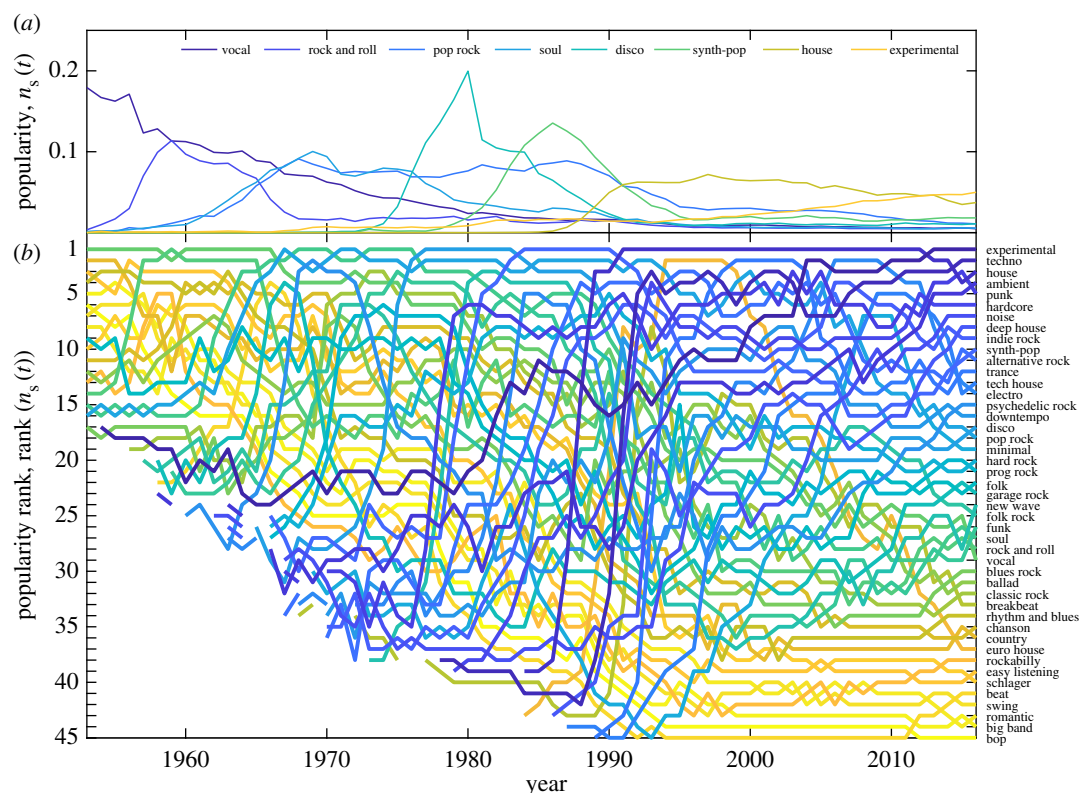


Figure 2. Fashion cycles in musical styles as ranked by their popularity $n_s(t)$. (a) For each year, we identify the most popular style and show its popularity $n_s(t)$ over the entire observation period. The first cycle is given by vocal (music strong focus on voice), followed by rock and roll, pop rock, soul, disco, synth-pop, house and finally experimental music. (b) Most styles enter with a rather low popularity (high rank) that they maintain over a number of years. These phases are eventually followed by a rapid increase in popularity (decrease in rank). Here we show only styles that were among the five most popular ones in at least one year. (Online version in colour.)

2.5. Models for cultural change dynamics

We first consider a neutral model for the evolution of musical styles. This approach is similar in spirit to the unified neutral theory of biodiversity in ecology [28]. There one seeks to explain the abundances of species (in our case: popularities of styles) using the neutrality assumption that the probability of a given species to produce offspring (i.e. release of a new album in a particular style) is proportional only to its abundance. For the evolution of musical styles, neutrality means that the popularity of a style does not depend on its structural characteristics, such as instrumentation. Such models can be formulated as Pólya urn-like models [33]. Imagine an urn with balls of different colours. Each ball is a musical release. The colour of the ball represents its musical style. At each time step, there are two possible actions—a copying step and an innovation step. With probability $1 - p$, $0 \leq p \leq 1$, one draws a ball from the urn, notes its colour and puts the ball back into the urn, together with a new ball of the same colour (copying step). With probability p , however, one introduces a ball with a new colour (innovation step). The style popularity $n_s(t)$ is the number of balls with colour s added to the urn in time interval t . Such Pólya urns are paradigmatic models for reinforcement processes [34]. In the long run, there is either a coexistence of different colours in the urn (different styles) or a winner-takes-all dynamics leads to a situation where almost all balls have the same colour—but no cycles emerge. In the following, we consider modifications to this neutral model that implement the mechanisms of ST, RPT and CDS, respectively. Each style s is associated with an angle $\theta_s \in [0, 2\pi]$ that describes the orientation of the style's instrument vector, A_{si} . Taking the average over the angles of all balls in the urn

yields the average angle, $\bar{\theta}$. There is a unique elite style e characterized by $\theta_e(t)$ with popularity $n_e(t)$. The modified urn models also have a copying and an innovation step, but we introduce two changes with respect to the neutral model. First, in the copying step, one copies either the current elite style or the style of a randomly chosen release. The second modification is in the innovation step. Each innovation introduces a new elite style that may or may not become successful. The parameter p can be a function of the state of the urn. For ST, p is proportional to the elite style popularity, $p \sim n_e(t)$ (the more people imitate an elite style, the higher are the chances for the elite to abandon that style). For RPT, p is constant (new elites form at random). For CDS, p is proportional to the similarity between the current elite style and the average styles of all releases, $p \sim \cos^2(\bar{\theta} - \theta_e)$ (the more uniform the current music releases are in style, the stronger the effect of CDS). In the case of RPT, the new elite style e' has a random style angle $\theta_{e'} \in [0, 2\pi]$. For ST and CDS, the new elite style is orthogonal to the current average style, $\theta_{e'} = \bar{\theta} - \pi/2$; see electronic supplementary material, text S1.

3. Results

3.1. Fashion cycles in musical styles

Figure 2a shows fashion cycles of selected musical styles as they appear in the empirical data. For each year, we identify the style with the highest popularity and show its $n_s(t)$ over the entire observation window. There is a clear pattern that styles increase in popularity until they reach a peak, after which they start to decrease while a new style takes over.

A more complete picture of the evolution of styles is shown in figure 2*b*. For each year, we rank styles according to their popularity, $\text{Rank}(n_s(t))$, measured by the number of releases in that style relative to the number of all releases. Figure 2*b* shows a selection of those styles that ranked among the five most popular styles in at least one year during the observations. Styles typically enter at high ranks (low popularity), where they remain over a certain period. At some point, they start to increase in popularity (decrease in rank). In some cases, this occurs within 2–3 years. Some styles manage to maintain high popularity (low ranks) over an extended period, whereas others fade back into oblivion rapidly (e.g. euro house after the 1990s).

3.2. Fashion cycles in models of cultural change

The neutral model is incapable of explaining the existence of fashion cycles as seen in figure 2*a*. The modified models ST, RPT and CDS do produce such cycles; see electronic supplementary material, figures S1 and S2. The RPT model assumes that two consecutive high-popularity styles are independent of each other in terms of their instrumentation, while in the ST and CDS models they tend to be opposites of each other. All the proposed theories, ST, RPT and CDS, can explain the observed fashion cycles. Further empirical tests are therefore necessary to determine which of the three competing mechanisms is at work.

3.3. Empirical tests for theories of fashion cycles

We now consider three different hypotheses to empirically test the theories of fashion cycles. We control the family-wise error rate (FWER) at a level of $\alpha < 0.05$ using the Bonferroni–Holm method [35]. We first test if styles with high popularity indeed tend to be imitated by other styles. For this, we assume that *lows* (styles with low popularity) have an overall tendency to imitate (or follow) the most popular styles by adopting the instruments used by the *highs* (indirectly representing the elites). Imitation can then be understood as a process that results in the similarity between *highs* at time t and *lows* at time t being lower than the similarity between *highs* at t and *lows* at $t + \delta t$ (*lows* follow *highs*). This means that *lows* tend to ‘move’ towards a higher similarity to the current *highs*. Using a similarity measure Φ to describe the overlap between two sets of musical styles, we can formulate the null hypothesis to reject the process of imitation in the data,

$$H_{\text{Imitation}}: \Phi(\text{Lows}(t + \delta t), \text{Highs}(t)) \leq \Phi(\text{Lows}(t), \text{Highs}(t)). \quad (3.1)$$

Figure 3*a* shows that the null hypothesis for imitation, $H_{\text{Imitation}}$, must be rejected given the data ($p < 0.001$, one-sided paired t -test using all observations). This means that there is a statistically highly significant effect in the data that low-popularity styles tend to imitate high-popularity styles in terms of instrumentation.

ST assumes that, as a result of this imitation, elite members will adopt a new style to differentiate themselves from their imitators. This new style will gain popularity and will be imitated by those that have followed the elite before. Now, the style formerly adopted by the elite members will belong to the group of *highs*, while the elite’s newly adopted style belongs to the group of *newcomers*. It follows that the similarity between *newcomers* at time t and *highs* at time $t + \delta t$, $\Phi(\text{New}(t), \text{Highs}(t + \delta t))$,

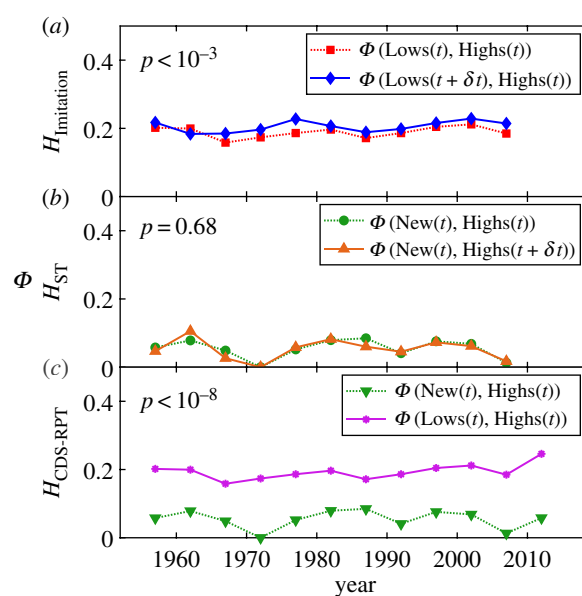


Figure 3. Three hypothesis tests for theories of cultural change. In each panel, a test represents where the null hypothesis is rejected if the similarities shown by solid lines are significantly larger than the similarities shown by dotted lines. (a) An imitation effect in the data implies that *lows* have a tendency to become more similar to the *highs* over time. The corresponding null hypothesis can be rejected, i.e. there is significant imitation in the data. (b) Under ST, one would expect that *newcomers* emerge from stylistic changes of the current *highs*. The corresponding null hypothesis cannot be rejected, i.e. there is no evidence for ST. (c) As opposed to RPT, CDS suggests that the next fashion cycle will be dominated by a counter-elite through its use of stylistic elements that are in direct opposition to the current elite. The corresponding null hypothesis can be rejected in the data. Cultural change occurs through counter-dominance signals and not by random choices. (Online version in colour.)

$\Phi(\text{Highs}(t + \delta t))$, should be larger than their similarity with the *highs* at time t , $\Phi(\text{New}(t), \text{Highs}(t))$. Otherwise, this would mean that the formation of *newcomers* is not related to concurrent changes in instrumentation of the *highs*—in this sense *newcomers* and *highs* would be independent of each other. However, ST posits that the *highs* would seek to differentiate themselves from the *newcomers* in order to signal their distinguished status. We test the null hypothesis to reject ST,

$$H_{\text{ST}}: \Phi(\text{New}(t), \text{Highs}(t + \delta t)) \leq \Phi(\text{New}(t), \text{Highs}(t)). \quad (3.2)$$

Figure 3*b* shows that the null hypothesis for ST, H_{ST} , cannot be rejected in the data ($p = 0.68$). This result suggests that styles do not become popular because of the adoption by an elite group that is then followed by a large number of imitators. The dynamics of *newcomers* and of the current *highs* seem to be independent of each other, meaning that they represent different elite groups in competition with each other.

In contrast to ST, the mechanisms of RPT and CDS suggest that styles become popular because of a new elite group that adopts them. In particular, if RPT were correct, one would expect that it is a randomly chosen style from the set of *lows* that will dominate the next fashion cycle. CDS in turn would imply that the next cycle is characterized by a counter-elite that adopts an instrumentation that is in stark contrast to the current *highs*. In other words, CDS suggests that the emerging *newcomers* follow styles that are more dissimilar to the current *highs* than randomly chosen *lows*. A null hypothesis that

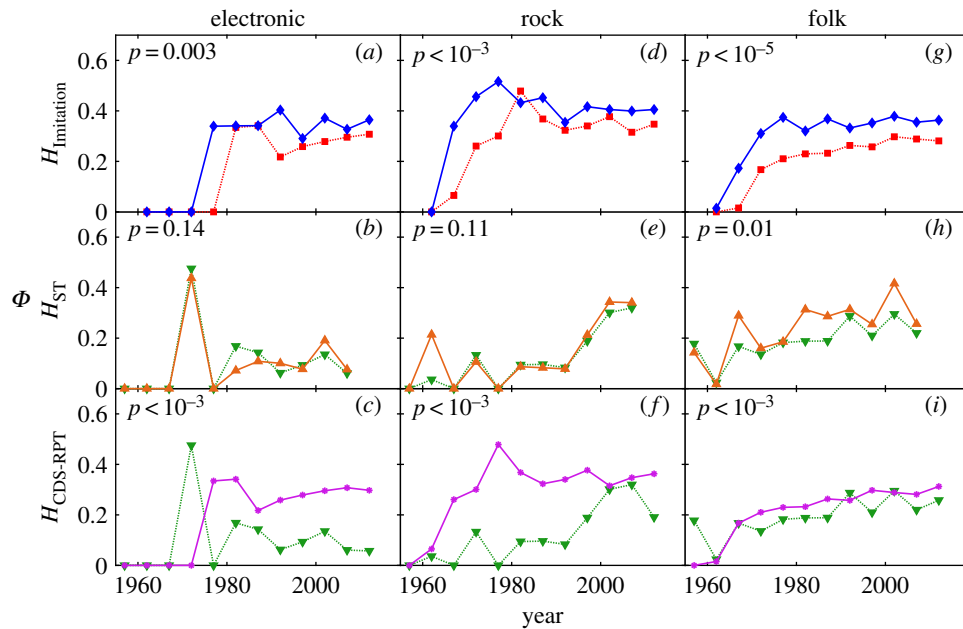


Figure 4. Genre-specific tests for three theories of cultural change. For electronic and rock music, we observe a significant imitation effect, no evidence for ST and a preference of CDS over RPT. For folk music, we observe a particularly strong imitation effect, and evidence for ST and CDS. (Online version in colour.)

fashion cycles can be explained on the basis of RPT instead of CDS can then be formulated as follows,

$$H_{\text{CDS-RPT}}: \Phi(\text{New}(t), \text{Highs}(t)) \geq \Phi(\text{Lows}(t), \text{Highs}(t)). \quad (3.3)$$

Figure 3c shows that the null hypothesis for RPT, $H_{\text{CDS-RPT}}$, can be firmly rejected in the data ($p < 10^{-8}$). This result shows with very high significance that *newcomers* typically are from an outsider group that tends to be in ‘opposition’ to the current *highs* in the use of instrumentation.

3.4. Genre-specific empirical tests for theories of cultural change

The three hypotheses, $H_{\text{Imitation}}$, H_{ST} and $H_{\text{CDS-RPT}}$, were tested for several musical genres. We considered the three genres with the highest numbers of related styles, namely electronic, rock and folk music; figure 4. For electronic (figure 4a–c) and rock music (figure 4d–f), again we find that imitation is a significant feature in the data, that there is no evidence for ST and that RPT must be rejected in the statistical test against expectations from CDS. For folk music, we find evidence for a particularly strong imitation effect (figure 4g) and a preference of CDS over RPT (figure 4i). However, in contrast to all other tests, we reject the null hypothesis for ST, H_{ST} ($p = 0.01$) (figure 4h). This means that, in folk music, we cannot rule out ST as a relevant mechanism, though, statistically, it is a much weaker feature of the data than CDS.

3.5. Robustness

There is a dependence on the parameters δt and m . To ensure that our results are independent of the particular choice of threshold m , we carry out a robustness test by letting m vary over a wide range. This test guarantees that our results are not driven by m , but reflect a general feature of style popularities. There is significant evidence for an imitation effect and a preference of CDS over RPT in every test. The evidence for CDS is particularly strong over short time intervals, δt . There

is no evidence for ST in almost all parameter settings, except for very small style numbers, m , and for long time intervals, δt ; see electronic supplementary material, figure S3.

We also study the robustness of our main results with respect to the choice of the similarity measure for styles in equation (2.4). Referring to the p -values of the three hypothesis tests shown in figure 3, for all considered measures, we find a significant imitation effect ($p = 0.014$ for the Jaccard coefficient, $p = 0.024$ when using the inverse Euclidean distance and $p = 0.019$ for ProbS), no evidence for ST (Jaccard: $p = 0.95$, Euclid: $p = 0.90$, ProbS: $p = 0.35$) and a preference of CDS over RPT (Jaccard: $p < 10^{-3}$, Euclid: $p = 0.0017$, ProbS: $p < 10^{-3}$). By comparing the results across different similarity measures, we see that the imitation effect is a less significant feature of the data than the CDS-RPT effect. For instance, if we would control the FWER by the more conservative Bonferroni procedure (instead of the uniformly more powerful Bonferroni–Holm method), the imitation effect would disappear when using the inverse Euclidean distance or ProbS as the similarity measure.

3.6. Spreading of popularity in the network of musical styles

Our formalism enables us to visualize the network evolution of musical styles. They relate to each other in a complex network of style–style similarities. In figure 5, we show the backbone of the style similarity network for three time intervals of 20 years, from 1956 to 2015. Each node represents a style with a size proportional to its popularity, $n_s(t)$; the colours indicate genres. Similar styles are connected by links. Styles belonging to the same genre tend to be in close proximity to each other. For each time interval, the network consists of a core of styles with high degrees, i.e. a large number of similar styles. The periphery contains low-degree nodes, i.e. styles with a low number of similar styles. Initially, styles belonging to the genres electronic and hip hop correspond almost exclusively to low-popularity nodes in the periphery of the network. Over time, they show large gains in popularity at the expense of

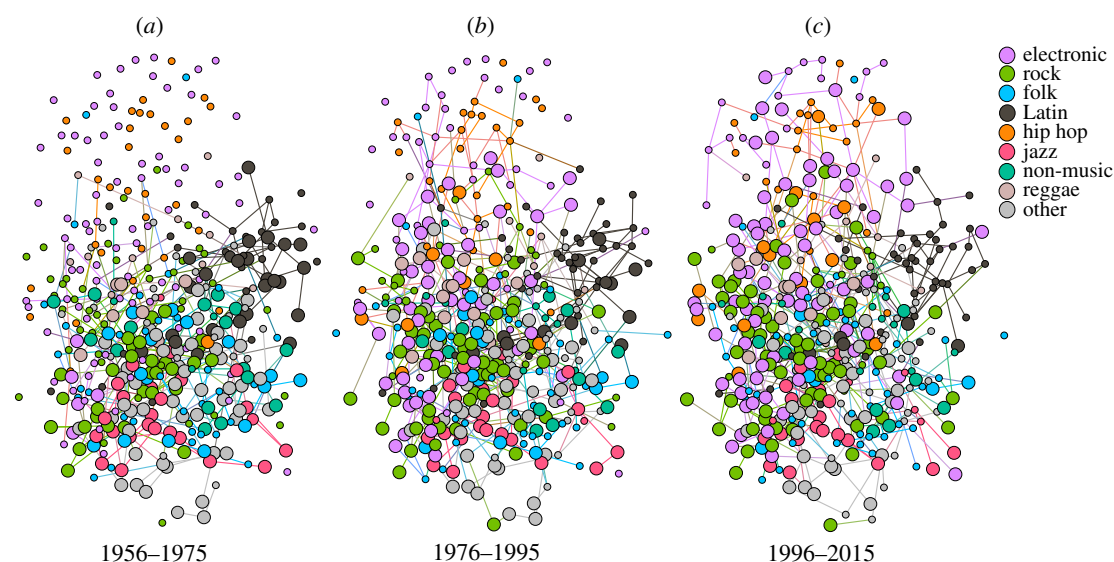


Figure 5. Dynamics of the style–style similarity network. We show the MST of the style–style similarity network for three time intervals. Nodes correspond to styles with colours given by their genre. The size of the nodes is proportional to their popularity, $n_s(t)$. There is a large number of styles in the periphery of the network with substantial gains in popularity, whereas some styles in the core of the network decrease in popularity, e.g. Latin music styles. (Online version in colour.)

former high-popularity styles lying in the core of the network, such as styles belonging to Latin music. Outsider styles in the periphery of the network can increase their chances of dominating the next fashion cycles by sending strong counter-dominance signals. Overall, this leads to a network evolution, where popularity diffuses from the periphery to the core, and then dissipates to the periphery again.

4. Discussion

By introducing a notion of cultural similarity, we formulated quantitative models for three possible mechanisms of cultural change, together with empirical, statistical tests to clarify which of these mechanisms best describe the actual data. We use a comprehensive dataset that contains almost all major album releases since the second half of the twentieth century, including the entire range of popularity of a genre or style (i.e. from least to most popular, as opposed to other studies that focused only on popular releases, such as the Billboard Hot 100 [36]). We found that low-popularity musical styles follow and imitate stylistic characteristics of high-popularity styles in a way that cannot be accounted for by neutral models of evolution. We found very little evidence for ST, i.e. of members of an elite ‘fleeing’ from their imitators. Instead, fashion cycles seem to emerge due to competition between members of different elite groups. Furthermore, in contrast to predictions of RPT, we found that styles with characteristics in strong opposition to styles of the current elite dominate the next fashion cycles. This supports our proposed mechanism of CDS according to which changes in art and fashion styles happen whenever a new elite successfully challenges the hegemony of a previous elite. We could confirm these findings in the genre-specific tests for electronic, rock and folk music, with the exception of some evidence for ST in folk music. This indicates that ST is only a feasible mechanism of cultural change in folk music (i.e. more traditional musical styles) over long time intervals. In conclusion, CDS drives short-term cultural change in all considered cases. A possible explanation for this finding is that

folk music constitutes a genre where music is frequently used as a means of cultural identification. Research in cross-cultural psychology applicable to folk music has shown that members of cultural groups tend to choose those stylistic members that are in stark contrast to the ones endorsed by other cultural groups [14,37]. Such costly signals could help to preserve a culture’s identity as they are less likely to be mimicked by members of other cultural groups.

Note that a limitation arises from our use of a partly folksonomic classification of musical styles [38]. That is, boundaries between styles emerge out of a collective and moderated effort of social tagging that might introduce some biases towards styles with a dedicated base of users. In addition, some care needs to be taken in interpreting the results of the hypothesis tests. In particular, our rejection of RPT over CDS in the third test does not necessarily imply that CDS is the *only* possible explanation for the observed fashion cycles, but rather that, of the mechanisms considered here, CDS is the least likely to be ruled out.

In 1905, German sociologist Georg Simmel [6] published his *Philosophie der Mode* [*Philosophy of Fashion*]. Still considered to be one of the most influential sociological theories of the early twentieth century, Simmel’s work ignited a century-long debate about the question of whether there are specific patterns in the social dynamics that underlie art and fashion cycles and—if yes—what the mechanisms driving them are. Here, we provide an entirely data-driven and quantitative answer to this question: cyclical changes of complex mass-cultural symbol systems manifested in art and fashion styles are driven by outside groups that successfully challenge the current elites.

Data accessibility. The data can be bulk downloaded under <https://data.discogs.com/>.

Authors’ contributions. All authors conceived of the study, designed the study, gave final approval for publication and agree to be held accountable for the work performed herein. P.K. analysed the data and drafted the manuscript. R.K. and S.T. critically revised the manuscript and participated in data analysis.

Competing interests. We declare we have no competing interests.

References

- Henshilwood C, d'Errico F. 2011 *Homo symbolicus: the dawn of language, imagination and spirituality*. Amsterdam, The Netherlands: John Benjamins Publishing Company.
- Kreuzbauer R, King D, Basu S. 2015 The mind in the object—psychological valuation of materialized human expression. *J. Exp. Psychol. Gen.* **144**, 764–787. (doi:10.1037/xge0000080)
- Bliege Bird R, Smith E. 2005 Signaling theory, strategic interaction, and symbolic capital. *Curr. Anthropol.* **46**, 221–248. (doi:10.1086/427115)
- Rahman Khan S. 2012 The sociology of elites. *Annu. Rev. Sociol.* **38**, 361–377. (doi:10.1146/annurev-soc-071811-145542)
- Cheng JT, Tracy JL, Foulsham T, Kingstone A, Henrich J. 2013 Two ways to the top: evidence that dominance and prestige are distinct yet viable avenues to social rank and influence. *J. Pers. Soc. Psychol.* **104**, 103–125. (doi:10.1037/a0030398)
- Simmel G. 1904 Fashion. *Int. Q.* **10**, 130–145. Reprinted in *Am. J. Sociol.* **62**, 541–558 (1957).
- Bourdieu P. 1984 *Distinction: a social critique of the judgement of taste*. Cambridge, MA: Harvard University Press.
- Aspers P, Godart F. 2013 Sociology of fashion: order and change. *Annu. Rev. Sociol.* **39**, 171–192. (doi:10.1146/annurev-soc-071811-145526)
- Acerbi A, Ghirlanda S, Enquist M. 2012 The logic of fashion cycles. *PLoS ONE* **7**, e32541. (doi:10.1371/journal.pone.0032541)
- Pesendorfer W. 1995 Design innovation and fashion cycles. *Am. Econ. Rev.* **85**, 771–779.
- Chwe MSY. 2013 *Rational ritual: culture, coordination, and common knowledge*. Princeton, NJ: Princeton University Press.
- Turchin P. 2016 *Ages of discord*. Chaplin, CT: Beresta Books.
- Veblen T. 1899/2009 *The theory of the leisure class*. Oxford, UK: Oxford University Press.
- Berger J, Ward M. 2010 Subtle signals of in-conspicuous consumption. *J. Consum. Res.* **37**, 555–569. (doi:10.1086/655445)
- Blumer H. 1969 Fashion: from class differentiation to collective selection. *Sociol. Q.* **10**, 275–291. (doi:10.1111/j.1533-8525.1969.tb01292.x)
- Kaiser S, Nagasawa R, Hutton S. 1995 Construction of an SI theory of fashion: Part 1. Ambivalence and change. *Cloth. Text. Res. J.* **13**, 172–183. (doi:10.1177/0887302X9501300304)
- Crane D. 1999 Diffusion models and fashion: a reassessment. *Ann. Am. Acad. Polit. Soc. Sci.* **566**, 13–24. (doi:10.1177/000271629956600102)
- Tschmuck P. 2012 *Creativity and innovation in the music industry*. Berlin, Germany: Springer.
- Stanley B. 2013 *Yeah yeah yeah: the story of modern pop*. London, UK: Faber & Faber.
- Peterson RA, Berger DG. 1975 Cycles in symbol production. *Am. Soc. Rev.* **40**, 158–173. (doi:10.2307/2094343)
- Bentley RA, Hahn MW, Shennan SJ. 2004 Random drift and culture change. *Proc. R. Soc. B* **271**, 1443–1450. (doi:10.1098/rspb.2004.2746)
- Bentley RA, Lipo CP, Herzog HA, Hahn MW. 2007 Regular rates of popular culture change reflect random copying. *Evol. Hum. Behav.* **28**, 151–158. (doi:10.1016/j.evolhumbehav.2006.10.002)
- Yoganarasimhan H. 2017 Identifying the presence and cause of fashion cycles in data. *J. Market. Res.* **54**, 5–26. (doi:10.1509/jmr.15.0119)
- Kreuzbauer R, Cheong B. 2015 Strategies of counterdominance – when luxury doesn't give you power. In *Proc. Society for Consumer Psychology Conf., Vienna*.
- Boehm C. 1993 Egalitarian behavior and reverse dominance hierarchy. *Curr. Anthropol.* **34**, 227–254. (doi:10.1086/204166)
- Schelling TC. 1980 *The strategy of conflict*. Cambridge, MA: Harvard University Press.
- Muthukrishna M, Henrich J. 2016 Innovation in the collective brain. *Phil. Trans. R. Soc. B* **371**, 20150192. (doi:10.1098/rstb.2015.0192)
- Hubbell SP. 2001 *The unified neutral theory of biodiversity and biogeography*. Princeton, NJ: Princeton University Press.
- Percino G, Klimek P, Thurner S. 2014 Instrumental complexity of music genres and why simplicity sells. *PLoS ONE* **9**, e115255. (doi:10.1371/journal.pone.0115255)
- Discogs. 2000 <http://www.discogs.com> [retrieved February 2016].
- Zhou T, Kuscsik Z, Liu J-G, Medo M, Wakeling JR, Zhang Y-C. 2010 Solving the apparent diversity-accuracy dilemma of recommender systems. *Proc. Natl Acad. Sci. USA* **107**, 4511–4515. (doi:10.1073/pnas.1000488107)
- Yildirim MA, Coscia M. 2014 Using random walks to generate associations between objects. *PLoS ONE* **9**, e104813. (doi:10.1371/journal.pone.0104813)
- Hoppe FM. 1984 Pólya-like urns and the Ewens' sampling formula. *J. Math. Biol.* **20**, 91–94. (doi:10.1007/BF00275863)
- Eggenberger F, Pólya G. 1923 Über die Statistik verketteter Vorgänge. *J. Appl. Math. Mech.* **3**, 279–289.
- Holm S. 1979 A simple sequentially rejective multiple test procedure. *Scand. J. Stat.* **6**, 65–70.
- Mauch M, MacCallum RM, Levy M, Leroi AM. 2015 The evolution of popular music: USA 1960–2010. *R. Soc. open sci.* **2**, 150081. (doi:10.1098/rsos.150081)
- Kreuzbauer R, Chiu CY, Lin S, Bae SH. 2014 When does life satisfaction accompany relational identity signaling: a cross-cultural analysis. *J. Cross-Cult. Psychol.* **45**, 646–659. (doi:10.1177/0022022113518369)
- Gruber T. 2007 Ontology of folksonomy: a mash-up of apples and oranges. *J. Semant. Web Inform. Syst. (IJSWIS)* **3**, 1–11. (doi:10.4018/jswis.2007010101)