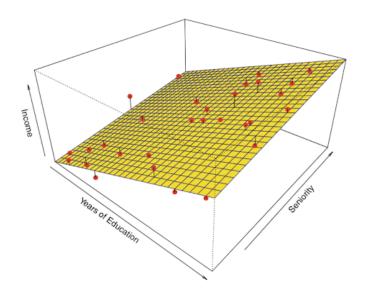
Introduction to data science & artificial intelligence (IF7100)

Arthur Charpentier

#322 Multiple Regression

été 2020

Linear Regression



From 1 to k features

$$y_i = \beta_0 + \beta_1 x_{i,1} + \dots + \beta_k x_{i,k} + \varepsilon_i \text{ or } y_i = \mathbf{x}_i^{\top} \boldsymbol{\beta} + \varepsilon_i,$$

From a mathematical perspective, use matrix notations

$$\underbrace{\begin{pmatrix} y_1 \\ \vdots \\ y_n \end{pmatrix}}_{\mathbf{y}, n \times 1} = \underbrace{\begin{pmatrix} 1 & x_{1,1} & \cdots & x_{1,k} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & x_{n,1} & \cdots & x_{n,k} \end{pmatrix}}_{\mathbf{X}, n \times (k+1)} \underbrace{\begin{pmatrix} \beta_0 \\ \beta_1 \\ \vdots \\ \beta_k \end{pmatrix}}_{\boldsymbol{\beta}, (k+1) \times 1} + \underbrace{\begin{pmatrix} \varepsilon_1 \\ \vdots \\ \varepsilon_n \end{pmatrix}}_{\boldsymbol{\varepsilon}, n \times 1}.$$

i.e. $\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\varepsilon}$. Assume that $\mathbb{E}(\varepsilon_i) = 0$ and $\text{Var}[\varepsilon_i] = \sigma^2$. The OLS estimator is $\widehat{\beta} \in \operatorname{argmin} \sum_{i=1}^{n} (y_i - \mathbf{x}_i^{\top} \widehat{\beta})^2$, i.e.

$$\widehat{oldsymbol{eta}} = \operatorname{argmin}ig(oldsymbol{y} - oldsymbol{X}oldsymbol{eta}ig)^ op ig(oldsymbol{y} - oldsymbol{X}oldsymbol{eta}ig)$$
 $\widehat{oldsymbol{eta}} = ig(oldsymbol{X}^ op oldsymbol{X}ig)^{-1}oldsymbol{X}^ op oldsymbol{y}$

Linear Regression

Then $\mathbb{E}(\widehat{\boldsymbol{\beta}}) = \boldsymbol{\beta}$ and $\text{Var}(\widehat{\boldsymbol{\beta}}) = \sigma^2(\boldsymbol{X}^{\top}\boldsymbol{X})^{-1}$. Assuming either that $\varepsilon_i \sim \mathcal{N}(0, \sigma^2)$ or *n* large,

$$\widehat{\boldsymbol{\beta}} pprox \mathcal{N}(\boldsymbol{\beta}, \sigma^2 oldsymbol{ig(oldsymbol{X}^{ op} oldsymbol{X}ig)^{-1}})$$

As previously (#321) we can test for significance H_0 : $\beta_i = 0$ and we can derive some confidence intervals, for β_i (for any i)

For prediction, $\hat{\mathbf{v}} = \hat{\beta}_0 + \hat{\beta}_1 \mathbf{x}_1 + \cdots + \hat{\beta}_{\nu} \mathbf{x}_{\nu}$.

Ceteris paribus can be translated into "all other things being equal" or "holding other factors constant"

Mutatis mutandis approximately translates as "allowing other things to change accordingly" or "the necessary changes having been made"



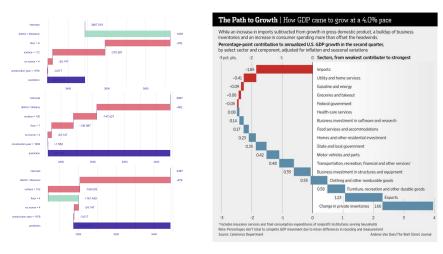
Regression

```
1 > import numpy as np
2 > import statsmodels.api as sm
x = \text{np.array}([[5,1], [15,4], [25,-5], [35,4],
     [45,-2], [55,2]
4 > x = sm.add_constant(x)
y = \text{np.array}([5, 20, 14, 32, 22, 38])
6 > model = sm.OLS(y, x)
7 > results = model.fit()
8 > print(results.summary())
coef. std err t P>|t| [0.025 0.975]
12 const 4.0581 3.370 1.204 0.315 -6.668 14.785
13 x1 0.5578 0.097 5.770 0.010 0.250 0.865
14 x2 1.5604 0.508 3.071 0.055 -0.057 3.178
16 Dep. Variable: y R-squared: 0.931
17 Model:
                OLS Adj. R-squared: 0.886
                       F-statistic: 20.37
18
```

Regression

```
1 > df = data.frame(x1 = c(5, 15, 25, 35, 45, 55),
                  x2 = c(1, 4, -5, 4, -2, 2),
2
                  y = c(5, 20, 14, 32, 22, 38))
3
4 > model = lm(y~x1+x2, data=df)
5 > summary(model)
6
7 Coefficients:
             Estimate Std. Error t value Pr(>|t|)
8
9 (Intercept) 4.05810 3.37049 1.204 0.3149
            10 x 1
            1.56037 0.50818 3.071 0.0545 .
11 x2
12 ---
13 Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.'
14
15 Residual standard error: 4.037 on 3 degrees of freedom
16 Multiple R-squared: 0.9314, Adjusted R-squared: 0.8857
17 F-statistic: 20.37 on 2 and 3 DF, p-value: 0.01796
```

Break-down for Additive Models



See The Wall Street Journal, on accounting equations

Partial Dependence Plot

Introduced in Greedy function approximation: A gradient boosting machine, Friedman (2001)

Let x be splited in two parts : x_s (variable(s) of interest) and x_s the complementary, $\mathbf{x} = (\mathbf{x}_s, \mathbf{x}_c)$. Partial dependence of \mathbf{x}_s is

$$p(\boldsymbol{x}_s) = \mathbb{E}[m(\boldsymbol{x}_s, \boldsymbol{S}_c)] \text{ and } \widehat{p}(\boldsymbol{x}_s) = \frac{1}{n} \sum_{i=1}^n \widehat{m}(\boldsymbol{x}_s, \boldsymbol{x}_{i,c})$$

