

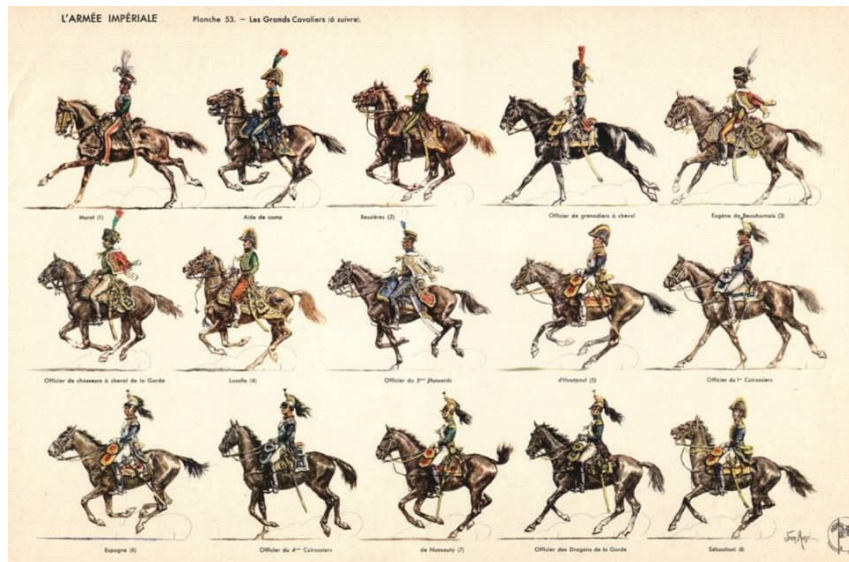
# Introduction to data science & artificial intelligence (INF7100)

Arthur Charpentier

#433 Memoryless Process (Geometric, Exponential)

été 2020

## ... and the Poisson distribution



# The Geometric Distribution

The **Geometric**  $\mathcal{G}(p)$ ,  $p \in (0, 1)$

$$\mathbb{P}(X = k) = p(1 - p)^{k-1} \text{ for } k = 1, 2, \dots$$

with cdf  $\mathbb{P}(X \leq k) = 1 - p^k$ .



Observe that this distribution satisfies the following relationship

$$\frac{\mathbb{P}(X = k + 1)}{\mathbb{P}(X = k)} = 1 - p \text{ (= constant) for } k \geq 1$$

First moments are here

$$\mathbb{E}(X) = \frac{1}{p} \text{ and } \text{Var}(X) = \frac{1 - p}{p^2}.$$

# The Period of Return & Memoryless

A return period, also known as a recurrence interval or repeat interval, is an average time or an estimated average time between events such as earthquakes, floods, landslides, or a river discharge flows to occur

$$\mathbb{E}(X) = \frac{1}{p} \text{ or } p = \frac{1}{\mathbb{E}(X)}$$

A 100-year flood is a flood event that has a 1 in 100 chance (1% probability) of being equaled or exceeded in any given year ([wikipedia](#))

Note that  $\mathbb{P}(X \geq h) = (1 - p)^h$ , then

$$\mathbb{P}(X \geq k + h | X \geq h) = \frac{\mathbb{P}(X \geq k + h)}{\mathbb{P}(X \geq h)} = \frac{(1 - p)^{k+h}}{(1 - p)^h} = (1 - p)^k$$

i.e.  $\mathbb{P}(X \geq k)$

# The Exponential Distribution

The **Exponential**  $\mathcal{E}(\lambda)$ ,  $\lambda > 0$

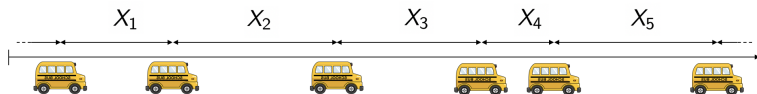
$$\mathbb{P}(X > x) = e^{-\lambda x} \text{ for } x \geq 0$$

with density  $\lambda e^{-\lambda x}$ .  $\mathbb{E}(X) = \frac{1}{\lambda}$  and  $\text{Var}(X) = \frac{1}{\lambda^2}$

Note: if  $X \sim \mathcal{E}(\lambda)$ ,  $Y = \lfloor X \rfloor \sim \mathcal{G}(p)$  with  $p = 1 - e^{-\lambda}$ .

The geometric and the exponential distribution are memoryless

$$\mathbb{P}(X > t + h | X > t) = \mathbb{P}(X > h).$$



# The Poisson Distribution

The **Poisson** distribution  $\mathcal{P}(\lambda)$ ,  $\lambda > 0$ , has distribution

$$\mathbb{P}(X = k) = e^{-\lambda} \frac{\lambda^k}{k!} \text{ where } k = 0, 1, \dots$$

Then  $\mathbb{E}(X) = \lambda$  and  $\text{Var}(X) = \lambda (= \mathbb{E}(X))$ .

Further, if  $X_1 \sim \mathcal{P}(\lambda_1)$  and  $X_2 \sim \mathcal{P}(\lambda_2)$  are independent, then

$$X_1 + X_2 \sim \mathcal{P}(\lambda_1 + \lambda_2)$$

Observe that a recursive equation can be obtained

$$\frac{\mathbb{P}(X = k + 1)}{\mathbb{P}(X = k)} = \frac{\lambda}{k + 1} \text{ pour } k \geq 1$$

Note:  $\mathbb{P}(N = 0) = e^{-\lambda}$ , e.g. if  $\lambda = 1$ ,  $\mathbb{P}(N = 0) \simeq 36.788\%$   
(and  $\mathbb{P}(N > 0) \simeq 63.212\%$ )

# The Poisson Approximation

Let  $X_i \sim \mathcal{B}(p)$ ,

$$\mathbb{P}(X = 0) = 1 - p \text{ and } \mathbb{P}(X = 1) = p.$$

then  $X = X_1 + \dots + X_n \sim \mathcal{B}(n, p)$  (binomial distribution)

$$\mathbb{P}(X = k) = \binom{n}{k} p^k (1-p)^{n-k} \text{ where } k = 0, 1, \dots, n, \binom{n}{k} = \frac{n!}{k!(n-k)!}$$

If  $n \cdot p \simeq \lambda$ ,  $X \simeq \mathcal{P}(\lambda)$

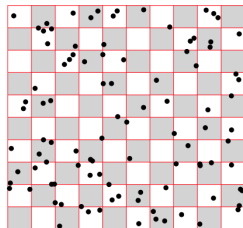
$$\begin{aligned}\mathbb{P}(X = 0) &= (1 - p)^n \\ &\simeq \left(1 - \frac{\lambda}{n}\right)^n \\ &\simeq e^{-\lambda}\end{aligned}$$

|     | Number of years without catastrophes |       |       |       |       |
|-----|--------------------------------------|-------|-------|-------|-------|
|     | 10                                   | 20    | 50    | 100   | 200   |
| 10  | 65.1%                                | 40.1% | 18.3% | 9.6%  | 4.9%  |
| 20  | 87.8%                                | 64.2% | 33.2% | 18.2% | 9.5%  |
| 50  | 99.5%                                | 92.3% | 63.6% | 39.5% | 22.5% |
| 100 | 99.9%                                | 99.4% | 86.7% | 63.4% | 39.5% |
| 200 | 99.9%                                | 99.9% | 98.2% | 86.6% | 63.3% |

# The Poisson Distribution

Consider some  $10 \times 10$  chess-board. Threw  $n = 100$  stones on it, and count the number of stones in each square.

```
1 > data.frame(N,F=table(nb_cell),P=c
2     (dpois(0:4,1),1-ppois(4,1)))
3     N   F     P
4 1    0 36 36.78
5 2    1 39 36.78
6 3    2 16 18.39
7 4    3  7  6.13
8 5    4  2  1.53
9 6 5+   0  0.37
```



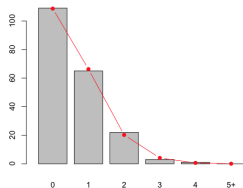
Consider some square, and let  $X$  denote the number of



# The Poisson Distribution

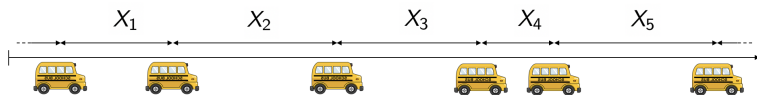
Nombre de soldats de cavaliers morts par ruade de cheval, entre 1875 et 1894, dans 10 corps (soit 200 corps annuels) **Bortkiewicz (1898)**

```
1 > data.frame(N,F=table(ruades),P=c(
    dpois(0:4,mean(ruades)),1-ppois
    (4,mean(ruades))))
2     N     F      P
3 1  0 109 108.67
4 2  1  65  66.21
5 3  2  22  20.22
6 4  3   3   4.11
7 5  4   1   0.63
8 6 5+   0   0.08
```



# The Poisson Process

Consider some bus arrivals, at times  $T_i$ , and let  $X_i = T_i - T_{i-1}$ , assume that  $X_i \sim \mathcal{E}(\lambda)$  are independent.



The number of buses in the time interval  $[0, t]$  is  $N_t \sim \mathcal{P}(\lambda t)$ .

See also [queuing theory](#)

# The Records Process

Let  $X_1, X_2, \dots$  denote some yearly observed value (maximum temperature, etc).

Let  $T_1 = \Delta_1 = 1$ ,  $T_n = \Delta_1 + \dots + \Delta_n$ , and

$$\Delta_{n+1} = \min_k \{X_{T_n+k} > X_{T_n}\}$$

If  $X_t$ 's are independent and identically distributed,

$$\frac{\log T_n - n}{\sqrt{n}} \rightarrow \mathcal{N}(0, 1)$$

hence  $\mathbb{E}(\log T_n) \simeq n$

Note: similarly  $\mathbb{E}(\log \Delta_n) \simeq n$

