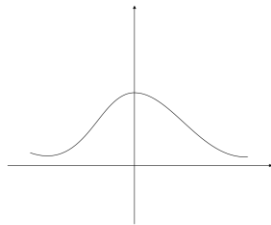
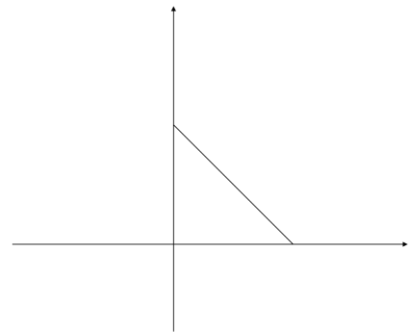


CONTINUOUS DISTRIBUTIONS

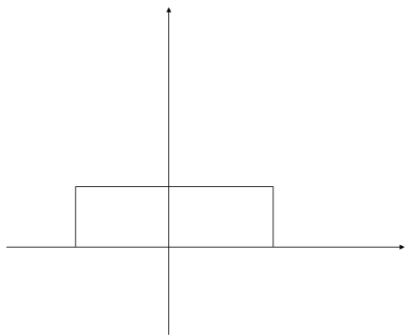
Common shapes of graph for a continuous distribution.



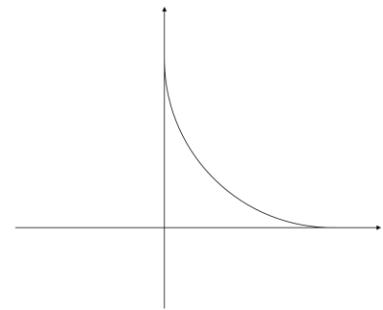
Normal



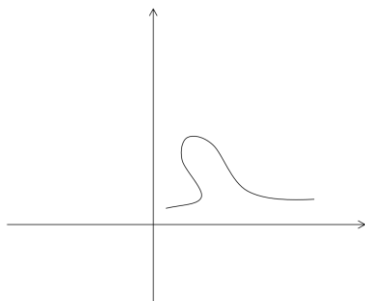
Triangular



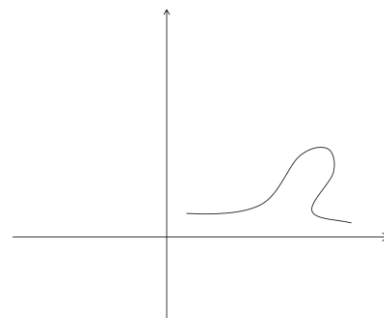
Uniform



Reverse

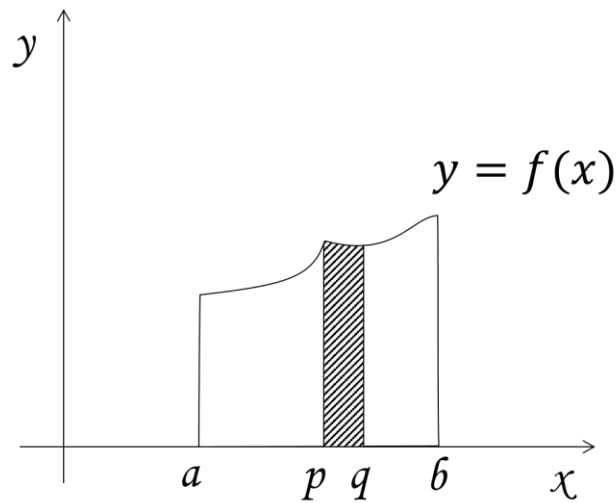


Positively Skewed



Negatively Skewed

Each continuous distribution can be defined by a curve with equation $y = f(x)$ for a domain say $a < x < b$. $f(x)$ is known as the **Probability Density Function** (PDF). The area under the graph will give the total of the probability, i.e. 1



i.e. $\int_a^b f(x) dx = 1$

What is $P(p < x < q)$?

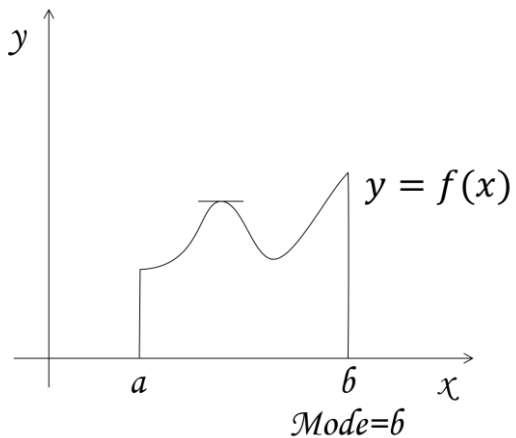
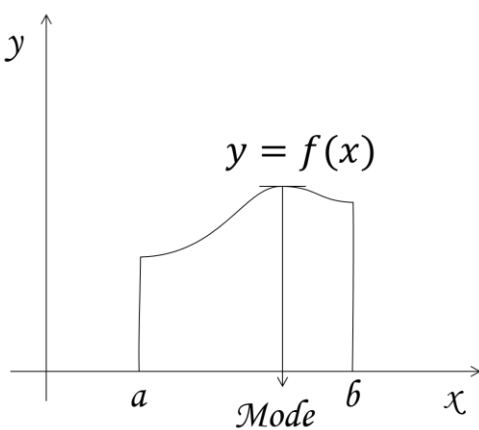
$$= \int_p^q f(x) dx$$

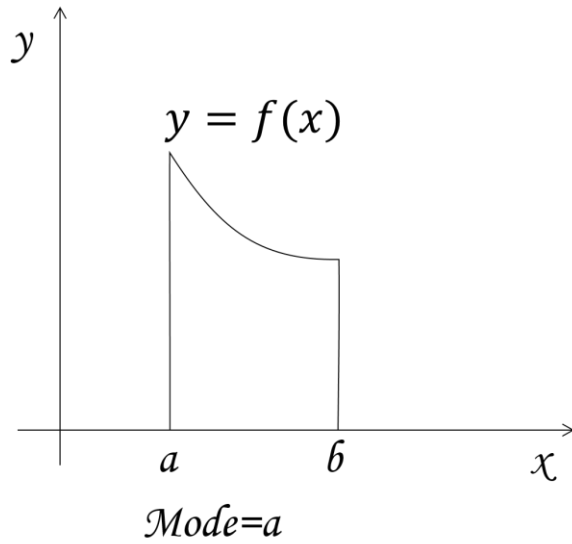
Often we write $P(p < x < q)$ where the variable X to be a continuous random variable with particular values x

$$P(X = p) = 0$$

Measures of central tendency

Mode

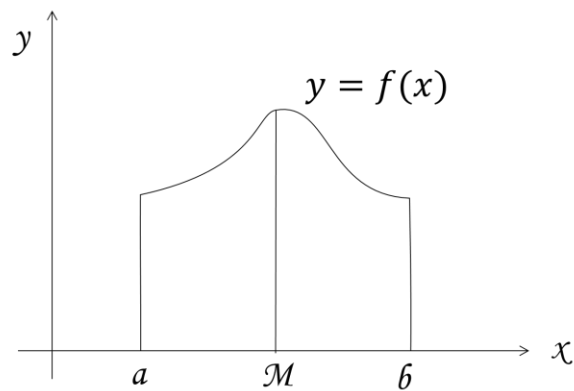




This is the value of x for which $f(x)$ is a maximum in the interval. This may be a calculus maximum in the interval $a < x < b$ or it may be at one of the points of the boundary $x = a$ or $x = b$.

i.e. $f'(x) = 0$, then solve the equation.

Median.



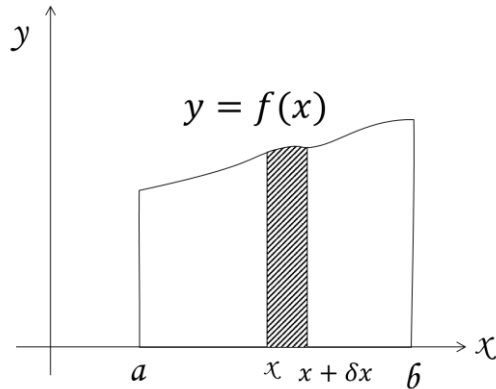
Median (M) is that value of x for which $\int_a^M f(x) dx = \frac{1}{2}$

Mean : μ or $E(X)$

Discrete random variable: $\mu = E(X) = \sum_{all\ x} x P(x)$

Continuous random variable: $\mu = E(X) = \sum_{x=a}^{x=b} x f(x) dx$

$$\rightarrow \int_a^b x f(x) dx \text{ as } dx \rightarrow 0$$



$$P(x < X < x + \delta x) \approx f(x) dx$$

For continuous:

$$E(X) = \mu = \int_a^b x f(x) dx \quad \text{also} \quad \int_a^b f(x) dx = 1$$

MEASURES OF VARIABILITY

Semi-interquartile range (or Quartile Deviation)

$$= \frac{1}{2}(Q_3 - Q_1)$$

$$\text{Where } \int_{Q_1}^{Q_3} f(x) dx = \frac{3}{4} \quad \text{or} \quad 0.75$$

$$\int_a^{Q_1} f(x) dx = \frac{1}{4} \quad \text{or} \quad 0.25$$

VARIANCE

Discrete:

$$Var(X) = \sum_{\text{all } x} (x - \mu)^2 \quad \text{also} \quad Var(X) = E[(x - \mu)^2]$$

$$\text{Continuous: } Var(X) = \sum_{\text{all } x} (x - \mu)^2 f(x) dx \quad \rightarrow Var(X) = \int_a^b (x - \mu)^2 f(x) dx$$

$$\begin{aligned}
Var(X) &= \int_a^b (x - \mu)^2 f(x) dx = \int_a^b (x^2 - 2x\mu + \mu^2) f(x) dx \\
&= \int_a^b x^2 f(x) dx - \int_a^b 2x\mu f(x) dx + \int_a^b \mu^2 f(x) dx \\
&= \int_a^b x^2 f(x) dx - 2\mu \int_a^b x f(x) dx + \mu^2 \int_a^b f(x) dx \\
&= \int_a^b x^2 f(x) dx - 2\mu(\mu) + \mu^2(1) \\
&= \int_a^b x^2 f(x) dx - 2\mu^2 + \mu^2 \\
&= \int_a^b x^2 f(x) dx - \mu^2 \\
\therefore Var(X) &= \int_a^b x^2 f(x) dx - \mu^2 \quad \text{or} \quad E(x^2) - [E(x)]^2
\end{aligned}$$

Remark:

Let $f(x)$ be a continuous function, then $E(X) = \mu = \int_{-\infty}^{\infty} x f(x) dx$ and

$$Var(X) = \int_{-\infty}^{\infty} x^2 f(x) dx - \mu^2$$

NOTE: $f(x)$ is called a probability density function if

- (i) $f(x) \geq 0$ for every value of x .
- (ii) $\int_{-\infty}^{\infty} f(x) dx = 1$
- (iii) $\int_a^b f(x) dx = P(a < x < b)$

Examples

1. A function $f(x)$ is defined as follows

$$f(x) = \begin{cases} 0, & x < 2 \\ \frac{1}{18}(2x + 3), & 2 \leq x \leq 4 \\ 0, & x > 4 \end{cases}$$

Show that it is a probability density function.

Solution

If $f(x)$ is a probability density function, then

$$(i) \quad \int_{-\infty}^{\infty} f(x) dx = 1$$

$$\text{Here } \int_2^4 \frac{1}{18}(2x + 3) dx = \frac{1}{18} [x^2 + 3x]_2^4 = \frac{1}{18} [16 + 12 - 4 - 6] = 1$$

- (ii) $f(x) > 0$ for $2 \leq x \leq 4$
2. The diameter of an electric cable is assumed to be continuous random variate with probability density function:

$$f(x) = \begin{cases} 6x(1-x), & 0 \leq x \leq 1 \\ 0, & \text{elsewhere} \end{cases}$$

- (i) Verify that above is a probability density function.
- (ii) Find the mean and variance

Solution

(i) $\int_{-\infty}^{\infty} f(x) dx = 1$

$$\int_0^1 6x(1-x) dx = \int_0^1 (6x - 6x^2) dx = \left[3x^2 - 2x^3 \right]_0^1 = 3 - 2 = 1$$

Secondly, $f(x) > 0$, for $0 \leq x \leq 1$

Hence the given function is a probability density function.

(ii) $\mu = \int_{-\infty}^{\infty} xf(x) dx = \int_0^1 x \cdot 6x(1-x) dx = \int_0^1 (6x^2 - 6x^3) dx = \left[2x^3 - \frac{3}{2}x^4 \right]_0^1 = 2 - \frac{3}{2} = \frac{1}{2}$

$$\begin{aligned} \text{Var}(x) &= \int_0^1 x^2 f(x) dx - \mu^2 = \int_0^1 x^2 \cdot 6x(1-x) dx = \int_0^1 (6x^3 - 6x^4) dx - \left(\frac{1}{2} \right)^2 \\ &= \left[\frac{3}{2}x^4 - \frac{6}{5}x^5 \right]_0^1 - \frac{1}{4} = \frac{3}{2} - \frac{6}{5} - \frac{1}{4} = \frac{1}{20} \end{aligned}$$

3. The probability density function $f(x)$ of a continuous random variable x is defined by

$$f(x) = \begin{cases} \frac{A}{x^3}, & 5 \leq x \leq 10 \\ 0, & \text{elsewhere} \end{cases}$$

Determine the value of A .

Solution

$$\int_5^{10} \frac{A}{x^3} dx = 1,$$

$$\int_5^{10} A x^{-3} dx = \left[\frac{Ax^{-2}}{-2} \right]_5^{10} = \left[\frac{A}{-2x^2} \right]_5^{10} = \frac{A}{-2(10)^2} - \frac{A}{-2(5)^2} = \frac{A}{-200} - \frac{A}{-50} = \frac{3A}{200} = 1$$

$$\frac{3A}{200} = 1$$

$$A = \frac{200}{3}$$

4. A number of no-defective hours in a factory can be modelled by the continuous random variable X with p.d.f $f(x)$ given by

$$f(x) = \begin{cases} k(4x - x^2) & \text{for } 0 \leq x \leq 4 \\ 0, & \text{elsewhere} \end{cases}$$

Find:

- (i) The value of constant k ;
- (ii) The expected value, $E(x)$;
- (iii) The variance, $Var(x)$;
- (iv) The mode of x .
- (v) The median of x
- (vi) $P(1 \leq x \leq 3)$

Solution

(i) $\int_{-\infty}^{\infty} f(x) dx = 1$

$$\int_0^4 k(4x - x^2) dx = k \left[2x^2 - \frac{x^3}{3} \right]_0^4 = k \left[32 - \frac{64}{3} \right] = \frac{32}{3} k = 1$$

$$\frac{32}{3} k = 1$$

$$k = \frac{3}{32}$$

(ii)

$$\begin{aligned} \mu &= \int_{-\infty}^{\infty} x f(x) dx = \int_0^4 x k(4x - x^2) dx = \int_0^4 \frac{3}{32} x(4x - x^2) dx = \int_0^4 \left(\frac{3}{8} x^2 - \frac{3}{32} x^3 \right) dx \\ &= \left[\frac{1}{8} x^3 - \frac{3}{128} x^4 \right]_0^4 = \frac{1}{8} \times 64 - \frac{3}{128} \times 256 = 8 - 6 = 2 \end{aligned}$$

(iii)

$$Var(x) = \int_{-\infty}^{\infty} x^2 f(x) dx$$

$$\begin{aligned} Var(x) &= \int_0^4 x^2 \cdot \frac{3}{32} (4x - x^2) dx - \mu^2 = \int_0^4 \left(\frac{3}{8} x^3 - \frac{3}{32} x^4 \right) dx - 2^2 = \left[\frac{3}{32} x^4 - \frac{3}{160} x^5 \right]_0^4 - 4 \\ &= \frac{3}{32} \times 256 - \frac{3}{160} \times 1024 - 4 \\ &= 24 - 19.5 - 4 \\ &= 0.5 \end{aligned}$$

(iv) Mode

$$f(x) = \frac{3}{32}(4x - x^2)$$

$$f(x) = \frac{3}{8}x - \frac{3}{32}x^2$$

$$f'(x) = 0, \quad f'(x) = \frac{3}{8} - \frac{3}{16}x, \quad \frac{3}{8} - \frac{3}{16}x = 0 \quad \therefore x = 2$$

(v) Median

$$\int_0^M f(x) dx = \frac{1}{2}$$

$$\int_0^M \left(\frac{3}{8}x - \frac{3}{32}x^2 \right) dx = \left[\frac{3}{16}x^2 - \frac{1}{32}x^3 \right]_0^M = \frac{3}{16}M^2 - \frac{1}{32}M^3 = \frac{1}{2}$$

$$\text{or} \quad \frac{1}{32}M^3 - \frac{3}{16}M^2 + \frac{1}{2} = 0$$

$$M = 5.4641 \text{ or } -1.4641 \text{ or } 2$$

(vi)

$$\begin{aligned} P(1 \leq x \leq 3) &= \int_1^3 \frac{3}{32}(4x - x^2) dx = \int_1^3 \left(\frac{3}{8}x - \frac{3}{32}x^2 \right) dx = \left[\frac{3}{16}x^2 - \frac{1}{32}x^3 \right]_1^3 \\ &= \left(\frac{3}{16} \times 3^2 - \frac{1}{32} \times 3^3 \right) - \left(\frac{3}{16} \times 1^2 - \frac{1}{32} \times 1^3 \right) \\ &= \frac{27}{16} - \frac{27}{32} - \frac{3}{16} + \frac{1}{32} \\ &= \frac{11}{16} \end{aligned}$$

5. A continuous random variable X is modelled by a probability density function.

$$f(x) = \begin{cases} kx, & 0 < x < 2 \\ k, & 2 < x < 5 \\ 0, & \text{elsewhere} \end{cases}$$

Where k is a constant.

Determine:

- (i) Value of constant k ;
- (ii) Mean value of x ;
- (iii) Median

Solution

$$(i) \quad \int_{-\infty}^{\infty} f(x) dx = 1$$

$$\int_0^2 kx \, dx + \int_2^5 k \, dx = 1$$

$$\left[\frac{kx^2}{2} \right]_0^2 + [kx]_2^5 = 1$$

$$2k + 5k + 2k = 1$$

$$9k = 1$$

$$k = \frac{1}{9}$$

(ii)

$$E(x) = \mu = \int_{-\infty}^{\infty} x f(x) \, dx$$

$$\begin{aligned} \mu &= \int_0^2 x \cdot k \, x \, dx + \int_2^5 x \cdot k \, dx \\ &= \int_0^2 \frac{1}{9} x^2 \, dx + \int_2^5 \frac{1}{9} x \, dx \\ &= \left[\frac{1}{27} x^3 \right]_0^2 + \left[\frac{1}{18} x^2 \right]_2^5 \\ &= \frac{1}{27} \times 2^3 + \frac{1}{18} \times 5^2 - \frac{1}{18} \times 2^2 \\ &= \frac{8}{27} + \frac{25}{18} - \frac{4}{18} \\ &= 1\frac{25}{54} \quad \text{or} \quad 1.4630 \end{aligned}$$

(iii)

$$\begin{aligned} \int_0^M f(x) \, dx &= \frac{1}{2} \\ \int_0^M \frac{1}{9} x \, dx + \int_M^5 \frac{1}{9} \, dx &= \frac{1}{2} \\ \left[\frac{1}{18} x^2 \right]_0^M + \left[\frac{1}{9} x \right]_M^5 &= \frac{1}{2} \\ \frac{1}{18} M^2 + \frac{5}{9} - \frac{1}{9} M &= \frac{1}{2} \\ \frac{1}{18} M^2 - \frac{1}{9} M + \frac{1}{18} &= 0 \quad \text{or} \quad M^2 - 2M + 1 = 0 \\ M &= 1 \quad \text{or} \quad 1 \\ \therefore M &= 1 \end{aligned}$$

6. The length of a circle is assumed to be a continuous random variable x with a probability density function $f(x)$ defined by:

$$f(x) = \begin{cases} kx^3, & 0 \leq x \leq 1 \\ k, & 1 \leq x \leq 2 \\ 0, & \text{elsewhere} \end{cases}$$

Determine the:

- (i) Value of the constant k ;
- (ii) Mean;
- (iii) $P\left(x \leq \frac{1}{2}\right)$

Solution

$$(i) \quad \int_{-\infty}^{\infty} f(x) dx = 1$$

$$\int_0^1 kx^3 dx + \int_1^2 k dx = 1$$

$$\left[\frac{x^4}{4} k \right]_0^1 + [kx]_1^2 = 1$$

$$\frac{1}{4}k + 2k - k = 1$$

$$\frac{5}{4}k = 1$$

$$k = \frac{4}{5}$$

(ii)

$$\mu = \int_0^1 x \cdot \frac{4}{5} x^3 dx + \int_1^2 x \cdot \frac{4}{5} dx$$

$$= \int_0^1 \frac{4}{5} x^4 dx + \int_1^2 \frac{4}{5} x dx$$

$$= \left[\frac{4}{25} x^5 \right]_0^1 + \left[\frac{4}{10} x^2 \right]_1^2$$

$$= \frac{4}{25} + \frac{16}{10} - \frac{4}{10}$$

$$= 1 \frac{9}{25} \text{ or } 1.36$$

(iii)

$$P\left(x \leq \frac{1}{2}\right) = \int_0^{\frac{1}{2}} \frac{4}{5} x^3 dx = \left[\frac{4}{20} x^4 \right]_0^{\frac{1}{2}} = \frac{4}{20} \left(\frac{1}{2}\right)^4 = \frac{4}{320} = \frac{1}{80} \text{ or } 0.0125$$

7. A continuous random variable t has a probability density function defined by

$$f(t) = \begin{cases} c(1-t)^2, & 1 < t < 3 \\ 0, & \text{elsewhere} \end{cases}$$

Determine the:

- (i) Value of the constant c ;
- (ii) Mean;
- (iii) $P(1.2 \leq t \leq 2.2)$

Solution

$$(i) \quad \int_{-\infty}^{\infty} f(x) dx = 1$$

$$\int_1^3 c(1-t)^2 dt, \text{ let } u = 1-t, \frac{du}{dt} = -1 \text{ or } dt = -du$$

$$\begin{aligned} \int_1^3 cu^2(-du) &= \int_1^3 -cu^2 du = \left[\frac{-cu^3}{3} \right]_1^3 = \left[\frac{-c(1-t)^3}{3} \right]_1^3 = \left(\frac{-c(1-3)^3}{3} \right) - \left(\frac{-c(1-1)^3}{3} \right) \\ &= \left(\frac{-c(-2)^3}{3} \right) - \left(\frac{-c(0)^3}{3} \right) = \frac{-c(-8)}{3} - 0 = \frac{8c}{3} \end{aligned}$$

$$\frac{8c}{3} = 1$$

$$c = \frac{3}{8}$$

(ii)

$$\begin{aligned} \mu &= \int_1^3 t \cdot \frac{3}{8}(1-t)^2 dt = \int_1^3 \frac{3}{8}t(1-2t+t^2) dt = \int_1^3 (t-2t^2+t^3) dt \\ &= \left[\frac{t^2}{2} - \frac{2}{3}t^3 + \frac{t^4}{4} \right]_1^3 = \left(\frac{9}{2} - \frac{54}{3} + \frac{81}{4} \right) - \left(\frac{1}{2} - \frac{2}{3} + \frac{1}{4} \right) \\ &= \frac{27}{4} - \frac{1}{12} \\ &= \frac{20}{3} \text{ or } 6.667 \end{aligned}$$

(iii)

$$\begin{aligned}
P(1.2 \leq t \leq 2.2) &= \int_{1.2}^{2.2} \frac{3}{8}(1-t)^2 dt = \int_{1.2}^{2.2} \frac{3}{8}(1-2t+t^2) dt \\
&= \int_{1.2}^{2.2} \left(\frac{3}{8} - \frac{3}{4}t + \frac{3}{8}t^2 \right) dt = \left[\frac{3}{8}t - \frac{3}{8}t^2 + \frac{1}{8}t^3 \right]_{1.2}^{2.2} \\
&= \left(\frac{3}{8} \times 2.2 - \frac{3}{8} \times 2.2^2 + \frac{1}{8} \times 2.2^3 \right) - \left(\frac{3}{8} \times 1.2 - \frac{3}{8} \times 1.2^2 + \frac{1}{8} \times 1.2^3 \right) \\
&= 0.341 - 0.126 \\
&= 0.215
\end{aligned}$$

8. The continuous random variable X has expected value $E(X) = 0.6$ and density function

$$f(x) = \begin{cases} ax + bx^2 & \text{for } 0 < x < 1 \\ 0, & \text{elsewhere} \end{cases}$$

- (i) Find the constant a and b .
- (ii) Find the $P(x \leq 0.4)$.
- (iii) Calculate σ_x .

Solution

$$(i) \quad \int_0^1 f(x) dx = 1$$

$$\int_0^1 (ax + bx^2) dx = 1$$

$$\left[\frac{ax^2}{2} + \frac{bx^3}{3} \right]_0^1 = 1$$

$$\frac{a}{2} + \frac{b}{3} = 1 \quad \text{or} \quad 3a + 2b = 6 \dots (i)$$

$$E(x) = \int_0^1 xf(x) dx = 0.6$$

$$\int_0^1 x(ax + bx^2) dx = 0.6$$

$$\int_0^1 (ax^2 + bx^3) dx = 0.6$$

$$\left[\frac{ax^3}{3} + \frac{bx^4}{4} \right]_0^1 = 0.6$$

$$\frac{a}{3} + \frac{b}{4} = 0.6 \quad \text{or} \quad 4a + 3b = 7.2 \dots (ii)$$

Solving the two equations we get $a = 3.6$ and $b = -2.4$

(ii)

$$P(x \leq 0.4) = \int_0^{0.4} (3.6x - 2.4x^2) dx = \left[1.8x^2 - 0.8x^3 \right]_0^{0.4} = 1.8 \times 0.4^2 - 0.8 \times 0.4^3 = 0.2368$$

(iii)

$$\begin{aligned} \text{Var}(x) &= \int_0^1 (3.6x^3 - 2.4x^4) dx - 0.6^2 \\ &= \left[0.9x^3 - 0.48x^4 \right]_0^1 - 0.36 \\ &= 0.9 - 0.48 - 0.36 \\ &= 0.06 \\ \therefore \sigma_x &= \sqrt{0.06} = 0.244948974 \end{aligned}$$

Exercises

1. The random variable x has the probability density function

$$f(x) = \begin{cases} kx & \text{if } 0 \leq x \leq 2 \\ 0, & \text{elsewhere} \end{cases}$$

Find k . Find x such that

- (i) $\Pr(X \leq x) = 0.1$
- (ii) $\Pr(X \leq x) = 0.95$

$$\text{Answer } [k = \frac{1}{2}, \text{ (i) } x = 0.632, \text{ (ii) } x = 1.949]$$

2. A continuous random variable X has a pdf $f(x)$ given by

$$f(x) = \begin{cases} k(3-x)(1+x) & \text{for } 0 \leq x \leq 3 \\ 0, & \text{elsewhere} \end{cases}$$

Determine:

- (i) Constant k ;
- (ii) $E(X)$;
- (iii) $\text{Var}(X)$

$$\text{Answer } [k = \frac{1}{9}, E(X) = 1.25, \text{Var}(X) = \frac{43}{80}]$$

3. A random variable X has a pdf $f(x)$ given by

$$f(x) = \begin{cases} \left[\frac{2+\sqrt{3}}{2} \right] [2+x]^{-1/2} & \text{for } 1 \leq x \leq 2 \\ 0, & \text{elsewhere} \end{cases}$$

Determine mode. **Answer** [1]

4. A random variable X has pdf $f(x)$ given by

$$f(x) = \begin{cases} k(2+x)^{-1/2} & \text{for } 1 \leq x \leq 2 \\ 0, & \text{elsewhere} \end{cases}$$

Determine:

- (i) Constant k ;
- (ii) Median of the distribution

$$\text{Answer } [k = \frac{2 + \sqrt{3}}{2}, M = \frac{4\sqrt{3} - 1}{4}]$$

5. Continuous random variable has pdf $f(x)$ given by

$$f(x) = \begin{cases} \frac{3}{64}x^2(4-x) & \text{for } 0 \leq x \leq 4 \\ 0, & \text{elsewhere} \end{cases}$$

Show that the median M is a root of equation $3M^4 - 16M^3 + 128 = 0$.

6. X is a continuous random variable with pdf $f(x)$ given by

$$f(x) = \begin{cases} \frac{x^2}{k} - \frac{kx}{18} + \frac{1}{3} & \text{for } 0 \leq x \leq 3 \\ 0, & \text{elsewhere} \end{cases}$$

Where k is a positive constant .

- (i) Determine k ;
- (ii) Determine the mode.

$$\text{Answer } [k = 6, Z = 3]$$

7. Find mean, median, mode for continuous distribution whose pdf $f(x)$ given by

$$f(x) = \begin{cases} 0.08(5-x) & \text{for } 0 \leq x \leq 5 \\ 0, & \text{elsewhere} \end{cases}$$

$$\text{Answer } [1\frac{2}{3}, 1.4645]$$

8. Determine $E(X)$ for the distribution with pdf $f(x)$ given by

$$f(x) = \begin{cases} \frac{3}{64}x^2(4-x) & \text{for } 0 \leq x \leq 4 \\ 0, & \text{elsewhere} \end{cases}$$

$$\text{Answer } [2.4]$$

9. For a random variable X with pdf

$$f(x) = \begin{cases} \frac{3}{2}x^{\frac{1}{2}} & \text{for } 0 \leq x \leq 1 \\ 0, & \text{elsewhere} \end{cases}$$

Determine median M correct to three decimal places.

$$\text{Answer } [0.630]$$

10. A continuous random variable X has the probability density function $f(x)$ defined by

$$f(x) = \begin{cases} \frac{c}{3}x, & 0 \leq x \leq 3 \\ c, & 3 \leq x \leq 4 \\ 0, & \text{otherwise} \end{cases}$$

Where c is a positive constant.

Determine the;

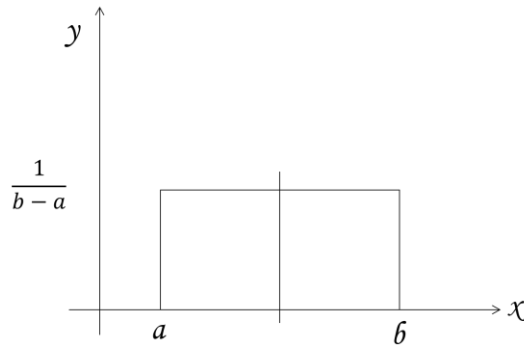
- (i) Value of c ;
- (ii) $E(5x - 4)$

Answer [$c = \frac{2}{5}$, 9]

CONTINUOUS UNIFORM DISTRIBUTION

A continuous random variable X has a uniform distribution over the interval (a, b) , with pdf $f(x)$ given by

$$f(x) = \begin{cases} \frac{1}{b-a} & \text{for } a < x < b \\ 0, & \text{elsewhere} \end{cases}$$



Check: $\int_a^b \frac{1}{b-a} dx = 1$

$$\int_a^b \frac{1}{b-a} dx = \left[\frac{x}{b-a} \right]_a^b = \frac{b}{b-a} - \frac{a}{b-a} = \frac{b-a}{b-a} = 1$$

By symmetry $Mean = Median = \frac{1}{2}(a + b)$

$$\mu = \int_a^b x f(x) dx = \int_a^b \frac{x}{b-a} dx = \left[\frac{x^2}{2(b-a)} \right]_a^b = \frac{b^2}{2(b-a)} - \frac{a^2}{2(b-a)} = \frac{b^2 - a^2}{2(b-a)} = \frac{(b-a)(b+a)}{2(b-a)} = \frac{b+a}{2}$$

$$\begin{aligned} \text{Var}(x) &= \int_a^b x^2 f(x) dx - \mu^2 = \int_a^b \frac{x^2}{b-a} dx - \left(\frac{b+a}{2} \right)^2 \\ &= \left[\frac{x^3}{3(b-a)} \right]_a^b - \frac{1}{4}(b+a)^2 \\ &= \frac{b^3}{3(b-a)} - \frac{a^3}{3(b-a)} - \frac{(b+a)^2}{4} \\ &= \frac{b^3 - a^3}{3(b-a)} - \frac{(b+a)^2}{4} \\ &= \frac{(b-a)(b^2 + ab + a^2)}{3(b-a)} - \frac{(b+a)^2}{4} \\ &= \frac{b^2 + ab + a^2}{3} - \frac{b^2 + 2ab + a^2}{4} \\ &= \frac{4(b^2 + ab + a^2) - 3(b^2 + 2ab + a^2)}{12} \\ &= \frac{4b^2 + 4ab + 4a^2 - 3b^2 - 6ab - 3a^2}{12} \\ &= \frac{b^2 - 2ab + a^2}{12} \\ \therefore \text{Var}(x) &= \frac{(b-a)^2}{12} \end{aligned}$$

Examples

1. Assume the time of arrival is uniformly distributed on the interval from 12:00 noon to 12:30 p.m.
 - (i) Find the density and distribution function for T .
 - (ii) Find the mean and standard deviation of T .
 - (iii) Joseph arrives at the bus stop at precisely noon. Calculate the probability he waits at least 15 minutes for the bus to arrive.

Solution

$$(i) \quad f(t) = \begin{cases} \frac{1}{30} & \text{for } 0 < t < 30 \\ 0, & \text{elsewhere} \end{cases}$$

$$(ii) \quad \mu = \frac{30+0}{2} = 15 \text{ minutes} \quad \text{and} \quad \sigma = \sqrt{\frac{(30-0)^2}{12}} = \sqrt{75} = 8.66025$$

$$(iii) \quad P(t > 15) = \int_{15}^{30} \frac{1}{30} dt = \left[\frac{t}{30} \right]_{15}^{30} = \frac{30}{30} - \frac{15}{30} = \frac{15}{30} = \frac{1}{2}$$

2. Dennis travels regularly by Kenya Airways, when he travels by plane. From past records he found that it is equally likely that his take-off time will be between 80 and 120 minutes after check in at the airport which he departs. Determine the probability that

- (i) He wants more than 105 minutes for take-off after checking in.
- (ii) His waiting time for take-off after checking in will be within 1.5 standard deviation of the mean waiting time.

Solution

$$(i) \quad \text{PDF } f(t) = \frac{1}{40}$$

$$P(T > 105) = \int_{105}^{120} \frac{1}{40} dt = \left[\frac{t}{40} \right]_{105}^{120} = \frac{120}{40} - \frac{105}{40} = \frac{15}{40} = \frac{3}{8}$$

$$\text{Or Area of the rectangle} = L \times W = 15 \times \frac{1}{40} = \frac{3}{8}$$

(ii)

$$\text{Or } \mu = \int_{80}^{120} t \times \frac{1}{40} dt = \int_{80}^{120} \frac{t}{40} dt = \left[\frac{t^2}{80} \right]_{80}^{120} = \frac{120^2}{80} - \frac{80^2}{80} = \frac{14400 - 6400}{80} = 100$$

$$\text{Or } \mu = \frac{120 + 80}{2} = 100$$

$$\begin{aligned} \text{Var}(T) &= \int_{80}^{120} t^2 f(t) dt - \mu^2 = \int_{80}^{120} \frac{t^2}{40} dt - 100^2 = \left[\frac{t^3}{120} \right]_{80}^{120} - 10,000 \\ &= \frac{120^3}{120} - \frac{80^3}{120} - 10,000 \\ &= \frac{400}{3} \text{ or } 133\frac{1}{3} \end{aligned}$$

$$\text{Standard deviation} = \sqrt{133\frac{1}{3}} = 11.547 \text{ minutes}$$

P(T within 1.5 standard deviation of the mean waiting time)

$$= \int_{100-1.5 \times 11.547}^{100+1.5 \times 11.547} \frac{t}{40} dt = \left[\frac{t}{80} \right]_{100-1.5 \times 11.547}^{100+1.5 \times 11.547} = \frac{117.3205}{80} - \frac{82.6795}{80} = 0.866025$$

NEGATIVE EXPONENTIAL DISTRIBUTION

Has a pdf $f(x)$ given by

$$f(x) = \begin{cases} \lambda e^{-\lambda x} & \text{for } 0 < x < \infty \\ 0, & \text{elsewhere} \end{cases}$$

λ is said to be the parameter of the distribution as $mean = \frac{1}{\lambda}$ and $variance = \frac{1}{\lambda^2}$.

Reverse 'J' shaped graph.

$$\text{Check: } \int_0^{\infty} \lambda e^{-\lambda x} dx = \left[\frac{\lambda e^{-\lambda x}}{-\lambda} \right]_0^{\infty} = \left[-e^{-\lambda x} \right]_0^{\infty} = -e^{-\infty} - (-e^0) = 0 + 1 = 1$$

Cumulative Density Function (CDF) $F(x) = P(0 < X < x)$

$$\begin{aligned} &= \left[-e^{-\lambda x} \right]_0^x \\ &= -e^{-\lambda x} - (-e^0) \\ &= 1 - e^{-\lambda x} \end{aligned}$$

$$\text{Mean } \mu = E(x) = \int_0^{\infty} \lambda x e^{-\lambda x} dx$$

Applying integration by parts $\int u dv = uv - \int v du$

$$\text{Let } u = \lambda x, \quad \frac{du}{dx} = \lambda \quad \text{or} \quad du = \lambda dx$$

Let $dv = e^{-\lambda x} dx$, then on integrating both sides we get

$$\begin{aligned} \int dv &= \int e^{-\lambda x} dx \\ v &= -\frac{e^{-\lambda x}}{\lambda} \end{aligned}$$

$$\begin{aligned}
\mu &= \int_0^{\infty} \lambda x e^{-\lambda x} dx = \lambda x \left(\frac{-\lambda e^{-\lambda x}}{\lambda} \right) - \int_0^{\infty} \frac{-e^{-\lambda x}}{\lambda} (\lambda dx) \\
&= -x e^{-\lambda x} + \int_0^{\infty} e^{-\lambda x} dx \\
&= \left[-x e^{-\lambda x} - \frac{e^{-\lambda x}}{\lambda} \right]_0^{\infty} \\
&= \left(-\infty e^{-\infty} - \frac{e^{-\infty}}{\lambda} \right) - \left(0 - \frac{e^0}{\lambda} \right) \\
&= (0 - 0) - \left(-\frac{1}{\lambda} \right) \\
\therefore \mu &= \frac{1}{\lambda}
\end{aligned}$$

Variance: $Var(x) = \sigma^2 = \int_0^{\infty} \lambda x^2 e^{-\lambda x} dx - \left(\frac{1}{\lambda} \right)^2$

Applying integration by parts $\int u dv = uv - \int v du$

Let $u = \lambda x^2$, then $\frac{du}{dx} = 2\lambda x$ or $du = 2\lambda x dx$

Let $dv = e^{-\lambda x} dx$, then on integrating both sides we get

$$\begin{aligned}
\int dv &= \int e^{-\lambda x} dx \\
v &= -\frac{e^{-\lambda x}}{\lambda}
\end{aligned}$$

$$\begin{aligned}
Var(x) &= \int_0^{\infty} \lambda x^2 e^{-\lambda x} dx - \frac{1}{\lambda^2} \\
&= \lambda x^2 \left(\frac{-e^{-\lambda x}}{\lambda} \right) - \int_0^{\infty} \left(\frac{-e^{-\lambda x}}{\lambda} \right) (2\lambda x dx) - \frac{1}{\lambda^2} \\
&= -x^2 e^{-\lambda x} + \int_0^{\infty} 2x e^{-\lambda x} dx - \frac{1}{\lambda^2} \quad \dots\dots\dots(1)
\end{aligned}$$

Again applying integration by parts for $\int_0^{\infty} 2x e^{-\lambda x} dx$

Let $u = 2x$, then $\frac{du}{dx} = 2$ or $du = 2dx$

Let $dv = e^{-\lambda x} dx$, then on integrating both sides we get

$$\int dv = \int e^{-\lambda x} dx$$

$$v = -\frac{e^{-\lambda x}}{\lambda}$$

$$\begin{aligned}\int_0^{\infty} 2xe^{-\lambda x} dx &= 2x \left(\frac{-e^{-\lambda x}}{\lambda} \right) - \int_0^{\infty} \frac{-e^{-\lambda x}}{\lambda} 2dx \\ &= \frac{-2xe^{-\lambda x}}{\lambda} + \int_0^{\infty} \frac{2e^{-\lambda x}}{\lambda} dx \\ &= \frac{-2xe^{-\lambda x}}{\lambda} - \frac{2e^{-\lambda x}}{\lambda^2} \dots\dots\dots(2)\end{aligned}$$

Substituting the RHS of equation (2) into equation (1) gives

$$\begin{aligned}Var(x) &= -x^2 e^{-\lambda x} + \int_0^{\infty} 2xe^{-\lambda x} dx - \frac{1}{\lambda^2} \\ &= \left[-x^2 e^{-\lambda x} - \frac{2xe^{-\lambda x}}{\lambda} - \frac{2e^{-\lambda x}}{\lambda^2} \right]_0^{\infty} - \frac{1}{\lambda^2} \\ &= (0 - 0 - 0) - \left(0 - 0 - \frac{2}{\lambda^2} \right) - \frac{1}{\lambda^2} \\ &= \frac{2}{\lambda^2} - \frac{1}{\lambda^2} \\ \therefore Var(x) &= \frac{1}{\lambda^2}\end{aligned}$$