

Visualizing Linear Models: An R Bag of Tricks Session 1: Getting Started

Michael Friendly SCS Short Course Oct-Nov, 2021

https://friendly.github.io/VisMLM-course/

Today's topics

- What you need for this course
- Why plot your data?
- Data plots
- Model (effect) plots
- Diagnostic plots

What you need

- R, version >= 3.6
 - Download from https://cran.r-project.org/
- RStudio IDE, highly recommended
 - https://www.rstudio.com/products/rstudio/
- R packages: see course web page
 - car
 - effects
 - heplots
 - candisc
 - visreg















Why plot your data?

Getting information from a table is like extracting sunlight from a cucumber. --- Farguhar & Farguhar, 1891

Information that is imperfectly acquired, is generally as imperfectly retained; and a man who has carefully investigated a printed table, finds, when done, that he has only a very faint and partial idea of what he has read; and that like a figure imprinted on sand, is soon totally erased and defaced.

--- William Playfair, The Commercial and Political Atlas (p. 3), 1786



Cucumbers

Table 7 Stevens et al. 2006, table 2: Determinants of authoritarian aggression

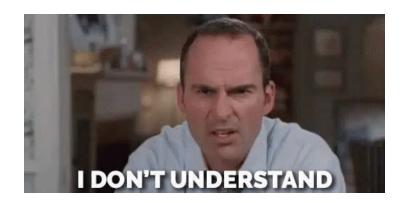
Coefficient
(Standard Error
.41 (.93)
1.31 (.33)**B,M
.93 (.32)**B,M
1.31 (.33)**B,M .93 (.32)**B,M 1.46 (.32)**B,M .07 (.32)A,CH,CO
.07 (.32)A,CH,CO
.96 (.37)**B,M
.20 (.13)
.22 (.12)#
21 (.12)#
32 (.12)*
27 (.07)**
.23 (.07)**
.00 (.01)
03 (.21)
.13 (.14)
.15 (.29)
.31 (.25)
10 (.27)
.15
.12
500

Results of a one model for authoritarian aggression

The information is overwhelmed by footnotes & significance **stars**

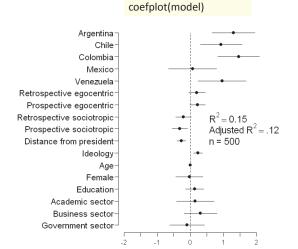
**p < .01, *p < .05, *p < .10 (twotailed)
A Coefficient is significantly different from Argentina's at $p < .05$;
^B Coefficient is significantly different from Brazil's at p < .05;
$^{\text{CH}}\text{Coefficient}$ is significantly different from Chile's at p < .05;
$^{\text{CO}}\text{Coefficient}$ is significantly different from Colombia's at $p<.05;$
M Coefficient is significantly different from Mexico's at p < .05;
VCoefficient is significantly different from Venezuela's at

What's wrong with this picture?



5

Sunlight



Why didn't they say this in the first place?

NB: This is a presentation graph equivalent of the table Shows coefficient

with 95% CI

Run, don't walk toward the sunlight



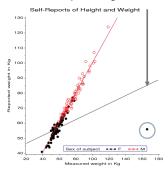
Graphs can give enlightenment



The greatest value of a picture is when it forces us to notice what we never expected to see.

-- John W. Tukey

Effect of one rotten point on regression



Dangers of numbers-only output

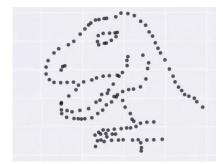
Student: You said to run descriptives and compute the correlation. What next?

Consultant: Did you plot your data?

X Mean: 54.26 Y Mean: 47.83 X SD : 16.76 Y SD : 26.93 Corr. : -0.06

With exactly the same stats, the data could be *any* of these plots

See how this in done in R: https://cran.r-project.org/web/packages/datasauRus/



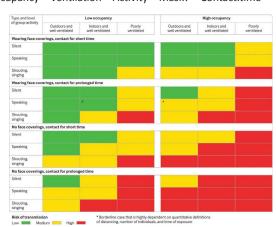
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Sometimes, don't need numbers at all

COVID transmission risk ~ Occupancy * Ventilation * Activity * Mask? * Contact.time

A complex 5-way table, whose message is clearly shown w/o numbers

There are 1+ unusual cells here. Can you see them?



From: N.R. Jones et-al (2020). Two metres or one: what is the evidence for physical distancing in covid-19? *BMJ* 2020;370:m3223, *doi: https://doi.org/10.1136/bmj.m3223*

If you do need tables – make them pretty

Several R packages make it easier to construct informative & pretty semi-graphic tables

Flipper lengths (mm) of the famous penguins of Palmer Station, Antarctica.

Presentation graph
Perhaps too cute!

Distribution of variables shown

Species Distribution
Avg. Std. Dev. Avg. Std. Dev.
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Artwork by @allison_horst

produced using modelsummary::datasummary, https://vincentarelbundock.github.io/modelsummary/articles/datasummary.html

Visual table ideas: Heatmap shading

Heatmap shading: Shade the background of each cell according to some criterion

The trends in the US and Canada are made obvious

NB: Table rows are sorted by Jan. value, lending coherence

Background shading ~ value:

US & Canada are made to stand out.

Tech note: use white text on a darker background

Unemployment rate in selected countries

January-August 2020, sorted by the unemployment rate in January.

country	Jan ^	Feb	Mar	Apr	May	Jun	Jul	Aug
Japan	2.4%	2.4%	2.5%	2.6%	2.9%	2.8%	2.9%	3.0%
Netherlands	3.0%	2.9%	2.9%	3.4%	3.6%	4.3%	4.5%	4.6%
Germany	3.4%	3.6%	3.8%	4.0%	4.2%	4.3%	4.4%	4.4%
Mexico	3.6%	3.6%	3.2%	4.8%	4.3%	5.4%	5.2%	5.0%
US	3.6%	3.5%	4.4%	14.7%	13.3%	11.1%		8.4%
South Korea	4.0%	3.3%	3.8%	3.8%	4.5%	4.3%	4.2%	3.2%
Denmark	4.9%	4.9%	4.8%	4.9%	5.5%	6.0%	6.3%	6.1%
Belgium	5.1%	5.0%	5.0%	5.1%	5.0%	5.0%	5.0%	5.1%
Australia	5.3%	5.1%	5.2%	6.4%	7.1%	7.4%	7.5%	6.8%
Canada	5.5%	5.6%	7.8%	13.0%	13.7%	12.3%	10.9%	10.2%
Finland	6.8%	6.9%	7.0%	7.3%	7.5%	7.8%	8.0%	8.1%

Source: OECD • Get the data • Created with Datawrapp

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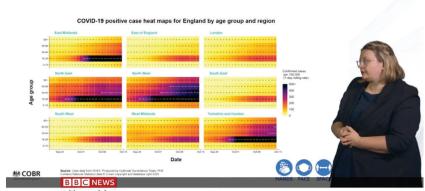
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Visual table ideas: Heatmap shading

As seen on TV ...

Covid rate ~ Age x Date x UK region

Better: incorporate geography, not just arrange regions alphabetically

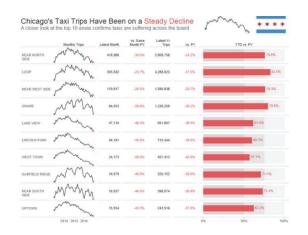


bbc.co.uk/news

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Visual table ideas: Sparklines

Sparklines: Mini graphics inserted into table cells or text



Linear models

Model:

$$\mathbf{y}_{i} = \beta_{0} + \beta_{1} \mathbf{X}_{i1} + \beta_{2} \mathbf{X}_{i2} + \dots + \beta_{p} \mathbf{X}_{ip} + \varepsilon_{i}$$

- Xs: quantitative predictors, factors, interactions, ...
- Assumptions:
 - Linearity: Predictors (possibly transformed) are linearly related to the outcome, y. [This just means linear in the parameters.]
 - Specification: No important predictors have been omitted; only important ones included. [This is often key & overlooked.]
 - The "holy trinity":
 - Independence: the errors are uncorrelated
 - Homogeneity of variance: $Var(\varepsilon_i) = \sigma^2 = constant$
 - Normality: ε, have a normal distribution

 $\mathcal{E}_{i} \sim_{iid} \mathcal{N}(0, \sigma^{2})$

From: https://www.pluralsight.com/guides/tableau-playbook-sparklines

The General Linear Model

- "linear" models can include:
 - transformed predictors: \sqrt{age} , log(income)
 - polynomial terms: age², age³, poly(age, n)
 - categorical "factors", coded as dummy (0/1) variables
 - treated (Yes/No), Gender (M/F/non-binary)
 - interactions: effects of x₁ vary over levels of x₂
 - treated × age, treated × sex, (2 way)
 - treated × age × sex (3 way)
- Linear model means linear in the parameters (β_i),

$$y = \beta_0 + \beta_1 \text{age} + \beta_2 \text{age}^2 + \beta_3 \log(\text{income}) + \beta_4 (\text{sex="F"}) + \beta_5 \text{age} \times (\text{sex="F"}) + \epsilon$$

In R, all handled by lm(y ~ ...)

Fitting linear models in R: Im()

- In R, lm() for everything
 - Regression models (X1, ... quantitative)

```
lm(y ~ X1, data=dat)  # simple linear regression
lm(y ~ X1+X2+X3, data=dat)  # multiple linear regression
lm(y ~ (X1+X2+X3)^2, data=dat)  # all two-way interactions
lm(log(y) ~ poly(X,3), data=dat)  # arbitrary transformations
```

ANOVA/ANCOVA models (A, B, ... factors)

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Fitting linear models in R: lm()

- Multivariate models: lm() with 2+ y vars
 - Multivariate regression

```
lm(cbind(y1, y2) ~ X1 + X2 + X3)  # std MMreg: all linear
lm(cbind(y1, y2) ~ poly(X1,2) + poly(X2,2))  # response surface
```

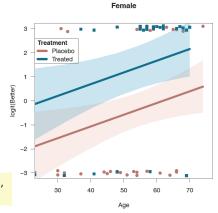
MANOVA/MANCOVA models

Generalized Linear Models: glm()

Transformations of y & other error distributions

- y ∈ (0/1): lived/died; success/fail; ...
- logit (log odds) model:
 - logit(y) = $log \frac{Pr(y=1)}{Pr(y=0)}$
 - linear logit model: logit(y) = $\beta_0 + \beta_1 x_1 + \beta_2 x_2 + ...$

glm(better ~ age + treat, family=binomial, data=Arthritis)



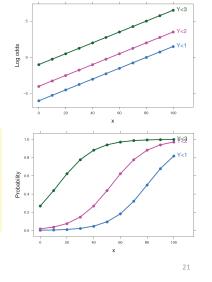
Generalized Linear Models

Ordinal responses

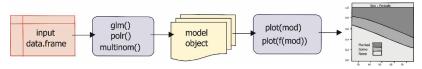
- Improved ∈ ("None" < "Some" < "Marked")
- Models: Proportional odds, generalized logits, ...

library(MASS) polr(Improved ~ Sex + Treat + Age, data=Arthritis)

library(nnet) multinom(Improved ~ Sex + Treat + Age, data=Arthritis)



Model-based methods: Overview



- models in R are specified by a symbolic model formula, applied to a data.frame
 - mod<-Im(prestige ~ income + educ, data=Prestige)</p>
 - mod<-glm(better ~ age + sex + treat, data=Arthritis, family=binomial)
 - mod<-MASS:polr(improved ~ age + sex + treat, data=Arthritis)</p>
- result (mod) is a "model object", of class "lm", "glm", ...
- method functions:
 - plot(mod), plot(f(mod)), ...
 - summary(mod), coef(mod), predict(mod), ...

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Plots for linear models

- Data plots:
 - plot response (y) vs. predictors, with smooth summaries
 - scatterplot matrix --- all pairs
- Model (effect) plots
 - plot predicted response (\hat{y}) vs. predictors, controlling for variables not shown.
- Diagnostic plots
 - Influence plots: leverage & outliers
 - Spread-level plots (non-constant variance?)

R packages

- car
 - Enhanced scatterplots
 - Diagnostic plots
- effects
 - Plot fitted effects of one predictor, controlling all others
- visreg
 - similar to effect plots, simpler syntax
- Both effects & visreg handle nearly all formula-based models
 - Im(), glm(), gam(), rlm, nlme(), ...

Occupational Prestige data

- Data on prestige of 102 occupations and
 - average education (years)
 - average income (\$)
 - % women
 - type (Blue Collar, Professional, White Collar)

> car::some(Presti	ge, 6)						
	education	income	women	prestige	census	type	
architects	15.44	14163	2.69	78.1	2141	prof	
physicians	15.96	25308	10.56	87.2	3111	prof	
commercial.artists	11.09	6197	21.03	57.2	3314	prof	
tellers.cashiers	10.64	2448	91.76	42.3	4133	WC	
bakers	7.54	4199	33.30	38.9	8213	bc	
aircraft.workers	8.78	6573	5.78	43.7	8515	bc	

Follow along

The R script (prestige-ex.R) for this example is linked on the course page. Download and open in R Studio to follow along.

```
    Prestige data prestige-ex.R | pestige-ex.html
    Penguin data penguins-Im-ex.R | penguins-Im-ex.html
```

The script was run with `knitr` (ctrl+shift+K) in R Studio to create the HTML output (prestige-ex.html)

The **Code** button there allows you to download the R code and comments

Linear models example: Occupational Prestine Code data

Michael Friendly

(These show a simple way to turn R scripts into finished documents)

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Informative scatterplots

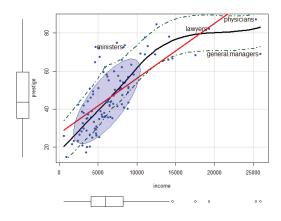
Scatterplots are most useful when enhanced with annotations & statistical summaries

Data ellipse and regression line show the linear model, prestige ~ income

Point labels show possible outliers

Smoothed (loess) curve and CI show the trend

Boxplots show marginal distributions

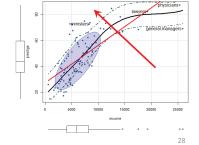


Informative scatterplots

car::scatterplot() provides all these enhancements

Skewed distribution of income & nonlinear relation suggest need for a transformation

Arrow rule: move on the scale of powers in direction of the bulge e.g.: x → sqrt(income) or log(income)

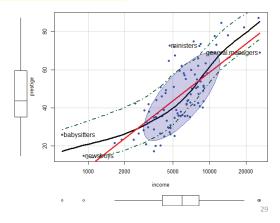


Try log(income)

Income now ~ symmetric

Relation closer to linear

log(income): interpret as effect of a multiple



Stratify by type?

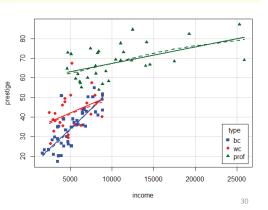
```
scatterplot(prestige ~ income | type, data=Prestige,
    col = c("blue", "red", "darkgreen"),
    pch = 15:17,
    legend = list(coords="bottomright"),
    smooth=list(smoother=loessLine, var=FALSE, span=1, lwd=4))
```

Formula: | type → "given type"

Different slopes: interaction of income * type

Provides another explanation of the non-linear relation

This may be a new finding!



Scatterplot matrix

```
prestige vs. all predictors
diagonal: univariate distributions
• income: + skewed
• %women: bimodal

off-diagonal: relations among predictors
```

Fit a simple model

```
> mod0 <- lm(prestige ~ education + income + women,</pre>
              data=Prestige)
> summary(mod0)
Coefficients:
               Estimate Std. Error t value Pr(>|t|)
(Intercept) -6.7943342 3.2390886 -2.098 0.0385
              4.1866373  0.3887013  10.771  < 2e-16 ***
income
              0.0013136 0.0002778
                                      4.729 7.58e-06 ***
women
             -0.0089052 0.0304071 -0.293 0.7702
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Multiple R-squared: 0.7982 Adjusted R-squared: 0.79
F-statistic: 129.2 on 3 and 98 DF, p-value: < 2.2e-16
                               Adjusted R-squared: 0.792
                                                                           Fits very well
```

But this ignores:

- nonlinear relation with income: should use log(income)
- occupation type
- · possible interaction of income*type

Fit a more complex model

```
> mod1 <- lm(prestige ~ education + women -
                                                                add interaction of log
               log(income)*type, data=Prestige)
> summary(mod1)
                                                                income by type
Coefficients:
                       Estimate Std. Error t value Pr(>|t|)
(Intercept)
                     -152.20589 23.24988
                                           -6.547 3.54e-09 ***
education
                        2.92817
                                  0.58828
                       0.08829
women
                                  0.03234
log(income)
                       18.98191
                                  2.82853
                       85.26415
                                  30.45819
                                                   0.00626 **
typeprof
                       29.41334
                                 36.50749
                       -9.01239
                                  3.41020
                                           -2.643
                                                   0.00970
log(income):typeprof
log(income):typewc
                       -3.83343
                                  4.26034
                                           -0.900
                                                   0.37063
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
                                                                   Fits even better!
Multiple R-squared: 0.8751, Adjusted R-squared: 0.8654
F-statistic: 90.07 on / and 90 DF, p-value: < 2.2e-16
                                                                   But how to understand?
```

Coefs for type compare mean "wc" and "prof" to "bc" Coefs for log(income) *type compare "wc" and "prof" slopes with that of "bc"

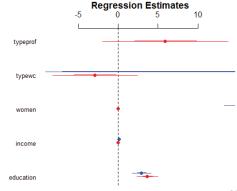
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Coefficient plots

Plots of coefficients with CI often more informative that tables

```
arm::coefplot(mod0, col.pts="red", cex.pts=1.5)
arm::coefplot(mod1, add=TRUE, col.pts="blue", cex.pts=1.5)
```

This plots raw coefficients, and the Xs are on different scales, so effect of income doesn't appear significant.



Model (effect) plots

- We'd like to see plots of the predicted value (\hat{y}) of the response against predictors (x_i)
 - Ordinary plot of y vs. x_i doesn't allow for other correlations
 - → Must control (adjust) for other predictors (x_{-j}) not shown in a given plot
- Effect plots
 - Variables not shown (x_{-i}) are averaged over.
 - Slopes of lines reflect the partial coefficient in the model
 - Partial residuals can be shown also

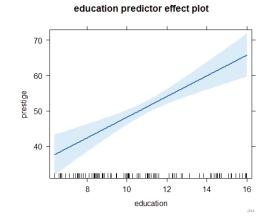
For details, see vignette("predictor-effects-gallery", package="effects)

Model (effect) plots: education

```
library("effects")
mod1.e1 <- predictorEffect("education", mod1)
plot(mod1.e1)</pre>
```

This graph shows the partial slope for education, controlling for all others

For each ↑ year in education, fitted prestige ↑2.93 points, (other predictors held fixed)

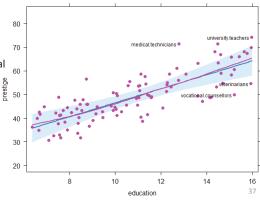


Model (effect) plots

mod1.ela <- predictorEffect("education", mod1, residuals=TRUE)
plot(mod1.ela,
 residuals.pch=16, id=list(n=4, col="black"))</pre>

Partial residuals show the residual of prestige controlling for other predictors

Unusual points here would signal undue influence



education predictor effect plot

Model (effect) plots: women

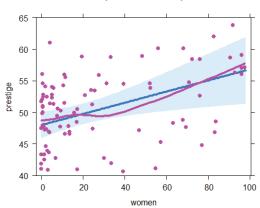


Surprise!

Prestige of occupations ↑ with % women (controlling for other variables)

Another 10% women ↑ prestige by 0.88 points

How to interpret this?



Model (effect) plots: income

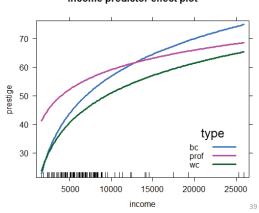
plot(predictorEffect("income", mod1),
 lines=list(multiline=TRUE, lwd=3),
 key.args = list(x=.7, y=.35))

income predictor effect plot

Income interacts with type in the model

The plot is curved because log(income) is in the model

Curvature reflects marginal effect of income for each occupation type



visreg plots: Air quality data

Daily air quality measurements in New York, May - Sep 1973

How does Ozone concentration vary with solar radiation, wind speed & temperature?

>	head(a	airqualit	ty)			
	Ozone	Solar.R	Wind	Temp	Month	Day
1	41	190	7.4	67	5	1
2	36	118	8.0	72	5	2
3	12	149	12.6	74	5	3
4	18	313	11.5	62	5	4
5	NA	NA	14.3	56	5	5
6	28	NA	14.9	66	5	6

see: https://pbreheny.github.io/visreg/ for examples & details

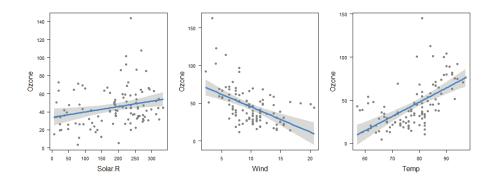
Air quality: main effects model

```
> fit1 <- lm(Ozone ~ Solar.R + Wind + Temp. data=airquality)</pre>
> summary(fit1)
Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) -64.3421
                       23.0547
                                -2.79
                                         0.0062 **
Solar.R
              0.0598
                        0.0232
                                         0.0112 *
             -3.3336
                         0.6544
                                  -5.09 1.5e-06 ***
Wind
              1.6521
                         0.2535
                                   6.52 2.4e-09 ***
Temp
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
Residual standard error: 21.18 on 107 degrees of freedom
  (42 observations deleted due to missingness)
Multiple R-squared: 0.6059, Adjusted R-squared: 0.5948
F-statistic: 54.83 on 3 and 107 DF, p-value: < 2.2e-16
```

visreg conditional plots

```
visreg(fit1, "Solar.R")
visreg(fit1, "Wind")
visreg(fit1, "Temp")

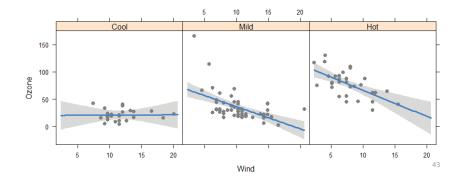
model summary =
predicted values (line) +
confidence band (uncertainty) +
partial residuals (objections)
```



Factor variables & interactions

cut Temp into three ordered levels of equal range airquality\$Heat <- cut(airquality\$Temp, 3, labels=c("Cool","Mild","Hot"))

fit model with interaction of **Wind * Heat** fit2 <- Im(Ozone ~ Solar.R + Wind*Heat, data=airquality) visreg(fit2, "Wind", by="Heat", layout=c(3,1), points=list(cex=1))

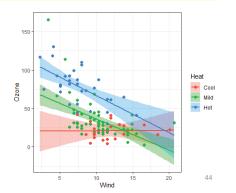


Factor variables & interactions

overlay=TRUE → superpose panels gg=TRUE → uses ggplot

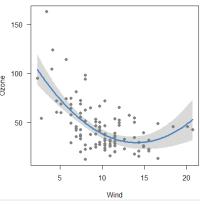
This allow slope for Wind to vary with Heat e.g., Wind has no effect when Cool

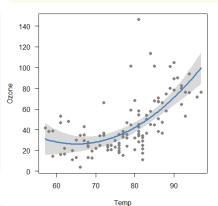
This model still assumes linear effects of Heat & Wind



Non-linear effects

fit <- Im(Ozone ~ Solar.R + poly(Wind,2) + Temp, data=airquality) visreg(fit, "Wind")

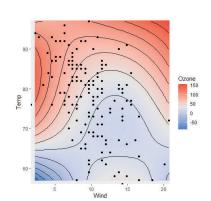


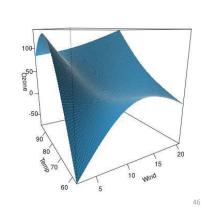


Response surface models (visreg2d)

Fit quadratics in both Wind & Temp and interaction Wind * Temp fitp <- Im(Ozone ~ Solar.R + poly(Wind,2) * poly(Temp,2), data=airquality)

visreg2d(fitp, "Wind", "Temp", plot.type="gg") + geom_contour(aes(z=z), color="black") visreg2d(fitp, "Wind", "Temp", plot.type="persp")

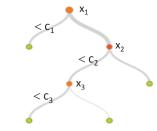


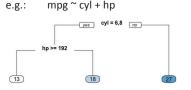


Regression trees

Regression trees are a non-parametric alternative to linear models

- Essential ideas:
 - Find predictor and split value which minimizes SSE
 - fitted value in each subgroupmean
 - repeat, recursively, splitting by next best predictor
- Large literature
 - cost, complexity tradeoff
 - pruning methods
 - boosting, cross-validation
 - tree averaging



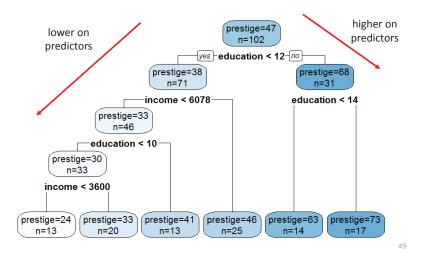


Prestige data: rpart tree

```
> library(rpart)
                      # calculating regression trees
                      # plotting regression trees
> library(rpart.plot)
> rmod <- rpart(prestige ~ education + income + women + type,
       data=Prestige,
       method = "anova")
> rpart.rules(rmod)
                       # print prediction rules
prestige
   24 when education < 10
                                 & income < 3600
   33 when education < 10
                                 & income is 3600 to 6078
   41 when education is 10 to 12 & income < 6078
   46 when education < 12
                                 & income >=
   63 when education is 12 to 14
   73 when education >= 14
```

Prestige data: rpart tree

rpart.plot(rmod, prefix="prestige=")



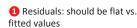
Diagnostic plots

- The linear model, $y=X\beta+\epsilon$ assumes:
 - Residuals, ε_i are normally distributed, $\varepsilon_i \sim N(0,\sigma^2)$
 - (Normality not required for Xs)
 - Constant variance, $Var(\varepsilon_i) = \sigma^2$
 - Observations y_i are statistically independent
- Violations → inferences may not be valid
- A variety of plots can diagnose all these problems
- Other methods (boxCox, boxTidwell) diagnose the need for transformations of y or Xs.

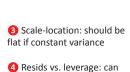
The "regression quartet"

In R, plotting a 1m model object → the "regression quartet" of plots

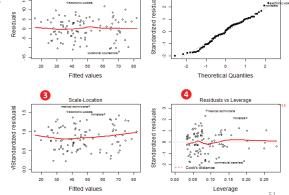
plot(mod1, lwd=2, cex.lab=1.4)



2 Q-Q plot: should follow the 45° line



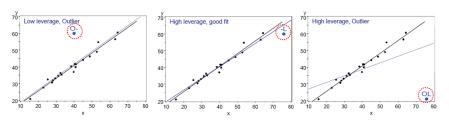
show influential observations



Unusual data: Leverage & Influence

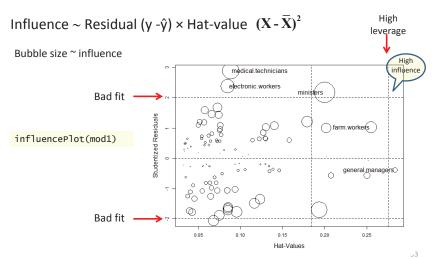
- "Unusual" observations can have dramatic effects on least-squares estimates in linear models
- Three archetypal cases:
 - Typical X (low leverage), bad fit -- Not much harm
 - Unusual X (high leverage), good fit -- Not much harm
 - Unusual X (high leverage), bad fit -- BAD, BAD, BAD
- Influential observations: unusual in both X & Y
- Heuristic formula:

Influence = X leverage x Y residual



Influence plots

Influence (Cook's D) measures impact of individual obs. on coefficients, fitted values

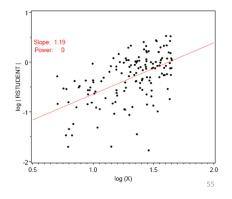


Spread-level plots

- To diagnose non-constant variance, plot:
 - log |Std. residual| vs. log (x)
 - log (IQR) vs log (median) [for grouped data]
- If \approx linear w/ slope b, transform y \rightarrow y (1-b)

Artificial data, generated so $\sigma \sim x$

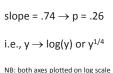
- $b \approx 1 \rightarrow power = 0$
- → analyze log(y)

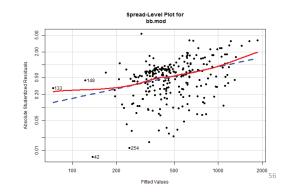


Spread-level plot: baseball data

Data on salary and batter performance from 1987 season

Suggested power transformation: 0.2609





Box Cox transformation

- Box & Cox proposed to transform y to a power, $y \to y^{(\lambda)}$ to minimize the residual SS (or maximize the likelihood)
 - Makes y^(λ) more nearly normal
 - Makes y^(λ) more nearly linear in with X

Formula for $y^{(\lambda)}$

- y⁽⁰⁾: log_e(y)
- λ < 0: flip sign to keep same order

$$y_i^{(\lambda)} = egin{cases} rac{y_i^{\lambda} - 1}{\lambda} & ext{if } \lambda
eq 0, \ \ln{(y_i)} & ext{if } \lambda = 0, \end{cases}$$

Power(p)	Transformation	Name
2	Y^2	Square
1	Y (No transformation)	Original Data
1/2	√ Y	Square root
"0"	log Y or log 10 (Y)	Logarithm
-1/2	-1 / √ Y	Reciprocal Root
-1	-1 / Y	Reciprocal
-2	-1 / Y^2	Reciprocal Square

Example: Cars93 data

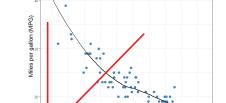
How does gas mileage (MPG.city) depend on vehicle weight?

```
> cars.mod <- lm(MPG.city ~ Weight, Cars93)
> coef(cars.mod)
(Intercept) Weight
    47.04835    -0.00803
```

Relationship clearly non-linear

Tukey arrow rule: transform Y (or X) as arrow thru the curve bulges

 $y \rightarrow \sqrt{y}$, log(y), 1/y $x \rightarrow \sqrt{x}$, log(x), 1/x

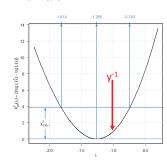


Untransformed response

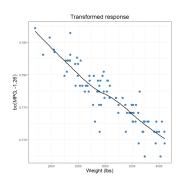
MASSextra package

- > library(MASSExtra)
- > box_cox(cars.mod) # pl
- # plot log likelihood vs. lambda
- > lamba(cars.mod)
- [1] -1.26

The plot of $-\log(L) \sim RSS$ shows the minimum & CI



plot (bc(MPG.city, lamba(cars.mod))



Summary

- Tables are for look-up; graphs can give insight
- "Linear" models include so much more than ANOVA & regression
- Data plots are more effective when enhanced
 - data ellipses → strength & precision of correlation
 - regression lines and smoothed curves
 - \blacksquare point identification \rightarrow noteworthy observations
- Effect plots show informative views of models
 - Visualize conditional effects, holding others constant
- Diagnostic plots can reveal influential observations and need for transformations.