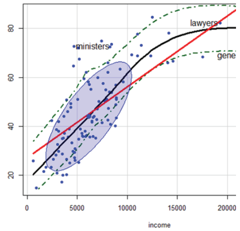
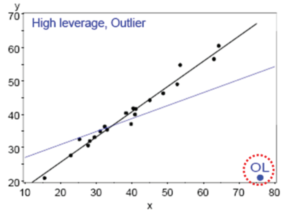




Visualizing Linear Models: An R Bag of Tricks
Session 1: Getting Started

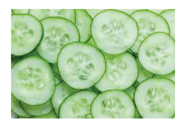
Michael Friendly
SCS Short Course
March, 2021
https://friendly.github.io/VisMLM-course/

Today's topics

- What you need for this course
- Why plot your data?
- Data plots
- Model (effect) plots
- Diagnostic plots

What you need

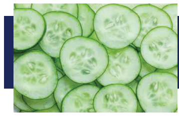
- R, version >=3.6
  - Download from https://cran.r-project.org/
- RStudio IDE, highly recommended
  - https://www.rstudio.com/products/rstudio/
- R packages: see course web page
  - car
  - effects
  - heplots
  - candisc
  - visreg



Why plot your data?

Getting information from a table is like extracting sunlight from a cucumber. --- Farquhar & Farquhar, 1891

Information that is imperfectly acquired, is generally as imperfectly retained; and a man who has carefully investigated a printed table, finds, when done, that he has only a very faint and partial idea of what he has read; and that like a figure imprinted on sand, is soon totally erased and defaced.
--- William Playfair, *The Commercial and Political Atlas* (p. 3), 1786



Cucumbers

Table 7
Stevens et al. 2006, table 2: Determinants of authoritarian aggression

Variable	Coefficient (Standard Error)
Constant	.41 (.93)
Countries	
Argentina	1.31 (.33)**[B,M]
Chile	.93 (.32)**[B,M]
Colombia	1.46 (.32)**[B,M]
Mexico	.07 (.32)[A,CH,CO,V]
Venezuela	.96 (.37)**[B,M]
Threat	
Retrospective egocentric economic perceptions	.20 (.13)
Prospective egocentric economic perceptions	.22 (.12)[#]
Retrospective sociotropic economic perceptions	–.21 (.12)[#]
Prospective sociotropic economic perceptions	–.32 (.12)*
Ideological distance from president	–.27 (.07)**
Ideology	
Ideology	.23 (.07)**
Individual Differences	
Age	.00 (.01)
Female	–.03 (.21)
Education	.13 (.14)
Academic Sector	.15 (.29)
Business Sector	.31 (.25)
Government Sector	–.10 (.27)
R^2	.15
Adjusted R^2	.12
N	500

Results of a one model for authoritarian aggression

The information is overwhelmed by footnotes & significance **stars**

**$p < .01$, *$p < .05$, [#]$p < .10$ (twotailed)

[A]Coefficient is significantly different from Argentina's at $p < .05$;

[B]Coefficient is significantly different from Brazil's at $p < .05$;

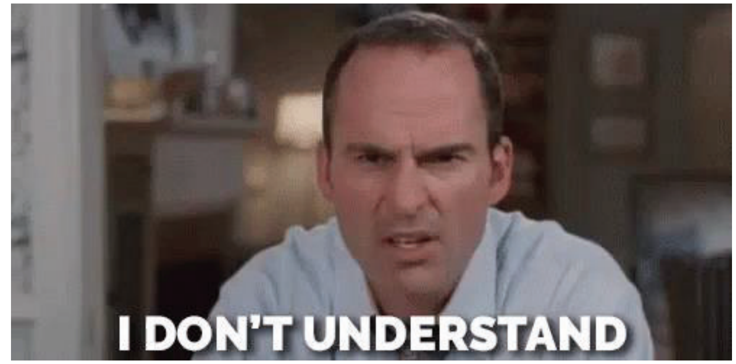
[CH]Coefficient is significantly different from Chile's at $p < .05$;

[CO]Coefficient is significantly different from Colombia's at $p < .05$;

[M]Coefficient is significantly different from Mexico's at $p < .05$;

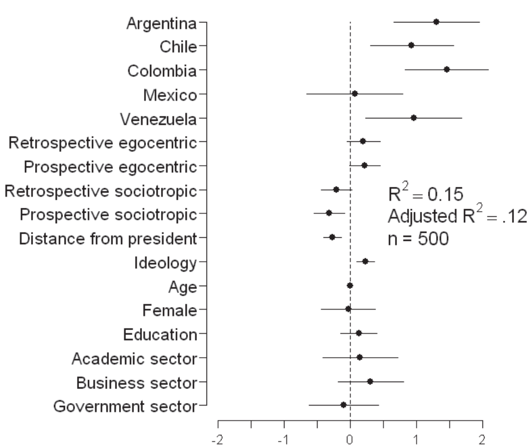
[V]Coefficient is significantly different from Venezuela's at $p < .05$.

What's wrong with this picture?



Sunlight

coefplot(model)



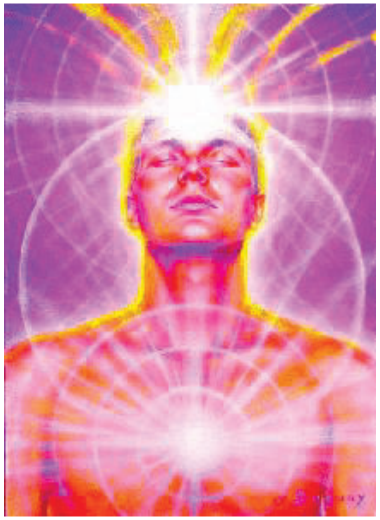
Why didn't they say this in the first place?

NB: This is a presentation graph equivalent of the table

Shows coefficient with 95% CI

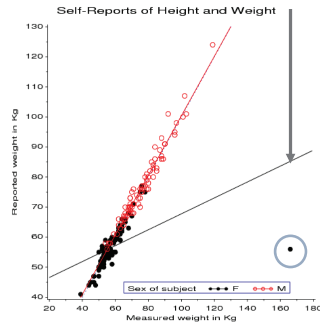
Run, don't walk toward the sunlight

Graphs can give enlightenment



The greatest value of a picture is when it forces us to notice what we never expected to see.
-- John W. Tukey

Effect of one rotten point on regression



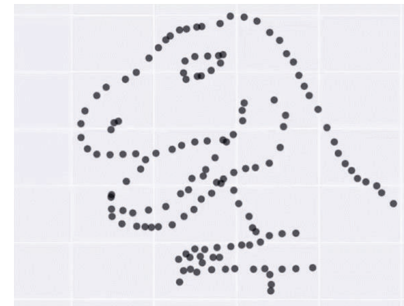
Dangers of numbers-only output

Student: You said to run descriptives and compute the correlation. What next?

Consultant: Did you plot your data?

```
X Mean: 54.26
Y Mean: 47.83
X SD  : 16.76
Y SD  : 26.93
Corr. : -0.06
```

With exactly the same stats, the data could be *any* of these plots



See how this in done in R: https://cran.r-project.org/web/packages/datasauRus/

Sometimes, don't need numbers at all

COVID transmission risk ~ Occupancy * Ventilation * Activity * Mask? * Contact.time

A complex 5-way table, whose message is clearly shown w/o numbers

There are 1+ unusual cells here. Can you see them?

From: N.R. Jones et-al (2020). Two metres or one: what is the evidence for physical distancing in covid-19? *BMJ* 2020;370:m3223, *doi: https://doi.org/10.1136/bmj.m3223*

If you do need tables– make them pretty

Several R packages make it easier to construct informative & pretty semi-graphic tables

Presentation graph

Perhaps too cute!

Distribution of variables shown

Flipper lengths (mm) of the famous penguins of Palmer Station, Antarctica.

		Female		Male	
Species	Distribution	Avg.	Std. Dev.	Avg.	Std. Dev.
		188	5.6	192	6.6
		192	5.8	200	6.0
		213	3.9	222	5.7

Artwork by @allison_horst

Visual table ideas: Heatmap shading

Heatmap shading: Shade the background of each cell according to some criterion

The trends in the US and Canada are made obvious

NB: Table rows are sorted by Jan. value, lending coherence

Unemployment rate in selected countries

January-August 2020, sorted by the unemployment rate in January.

country	Jan ▲	Feb	Mar	Apr	May	Jun	Jul	Aug
Japan	2.4%	2.4%	2.5%	2.6%	2.9%	2.8%	2.9%	3.0%
Netherlands	3.0%	2.9%	2.9%	3.4%	3.6%	4.3%	4.5%	4.6%
Germany	3.4%	3.6%	3.8%	4.0%	4.2%	4.3%	4.4%	4.4%
Mexico	3.6%	3.6%	3.2%	4.8%	4.3%	5.4%	5.2%	5.0%
US	**3.6%**	**3.5%**	**4.4%**	**14.7%**	**13.3%**	**11.1%**	**10.2%**	**8.4%**
South Korea	4.0%	3.3%	3.8%	3.8%	4.5%	4.3%	4.2%	3.2%
Denmark	4.9%	4.9%	4.8%	4.9%	5.5%	6.0%	6.3%	6.1%
Belgium	5.1%	5.0%	5.0%	5.1%	5.0%	5.0%	5.0%	5.1%
Australia	5.3%	5.1%	5.2%	6.4%	7.1%	7.4%	7.5%	6.8%
Canada	**5.5%**	**5.6%**	**7.8%**	**13.0%**	**13.7%**	**12.3%**	**10.9%**	**10.2%**
Finland	6.8%	6.9%	7.0%	7.3%	7.5%	7.8%	8.0%	8.1%

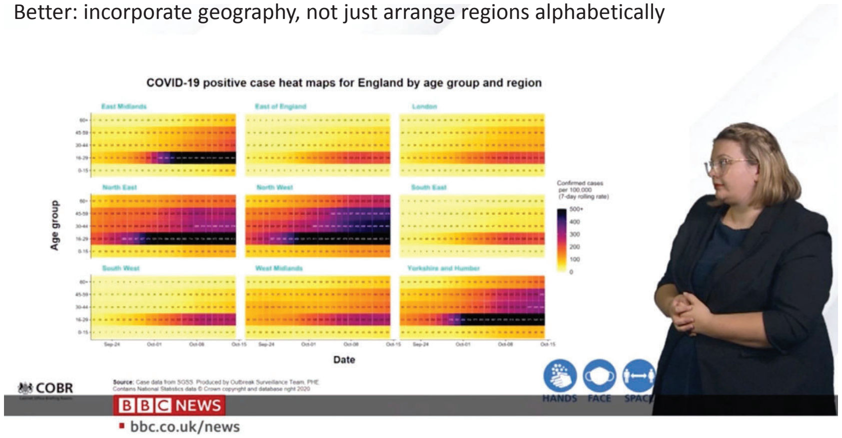
Source: OECD • Get the data • Created with Datawrapper

Visual table ideas: Heatmap shading

As seen on TV …

Covid rate ~ Age x Date x UK region

Better: incorporate geography, not just arrange regions alphabetically



Visual table ideas: Sparklines

Sparklines: Mini graphics inserted into table cells or text

Chicago's Taxi Trips Have Been on a Steady Decline

A closer look at the top 10 areas confirms taxis are suffering across the board

	Monthly Trips	Latest Month	vs. Same Month PY	Latest Yr Trips	vs. PY	YTD vs. PY
NEAR NORTH SIDE		418,986	-30.0%	5,908,758	-24.2%	75.8%
LOOP		305,562	-23.7%	4,288,823	-17.5%	82.5%
NEAR WEST SIDE		119,817	-26.5%	1,580,838	-23.7%	76.3%
OHARE		64,553	-28.0%	1,230,250	-20.2%	79.8%
LAKE VIEW		47,110	-48.4%	951,807	-38.5%	61.5%
LINCOLN PARK		44,191	-45.0%	733,344	-39.9%	60.1%
WEST TOWN		24,173	-50.9%	457,412	-42.9%	57.1%
GARFIELD RIDGE		16,070	-40.9%	320,152	-28.9%	71.1%
NEAR SOUTH SIDE		16,037	-46.9%	398,074	-26.9%	73.1%
UPTOWN		15,554	-43.3%	243,516	-37.8%	62.2%
	2014 2015 2016					0% 50% 100%

From: https://www.pluralsight.com/guides/tableau-playbook-sparklines

Linear models

- Model:

$$y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \cdots + \beta_p x_{ip} + \varepsilon_i$$

- Xs: quantitative predictors, factors, interactions, …
- Assumptions:
  - **Linearity**: Predictors (possibly transformed) are linearly related to the outcome, **y**. [This just means linear in the **parameters**.]
  - **Specification**: No important predictors have been omitted; only important ones included. [This is often key & overlooked.]
  - The "holy trinity":
    - **Independence**: the errors are uncorrelated
    - **Homogeneity of variance**: Var(ε_i) = σ^2 = constant
    - **Normality**: ε_i have a normal distribution

$$\left.\begin{array}{l}\text{Independence}\\ \text{Homogeneity of variance}\\ \text{Normality}\end{array}\right\} \varepsilon_i \sim_{iid} \mathcal{N}(0, \sigma^2)$$

The General Linear Model

- "linear" models can include:
  - transformed predictors: $\sqrt{age}$, log(*income*)
  - polynomial terms: age^2, age^3, `poly(age, n)`
  - categorical "factors", coded as dummy (0/1) variables
    - treated (Yes/No), Gender (M/F/non-binary)
  - interactions: effects of x_1 vary over levels of x_2
    - treated × age, treated × sex, (2 way)
    - treated × age × sex (3 way)
- Linear model means linear in the parameters (β_i),

$$y = \beta_0 + \beta_1 \text{age} + \beta_2 \text{age}^2 + \beta_3 \log(\text{income}) + \beta_4 (\text{sex}=\text{"F"}) + \beta_5 \text{age} \times (\text{sex}=\text{"F"}) + \epsilon$$

- In R, all handled by `lm(y ~ …)`

Fitting linear models in R: lm()

- In R, **`lm()`** for everything
  - Regression models (X1, ... quantitative)

```
lm(y ~ X1, data=dat)                  # simple linear regression
lm(y ~ X1+X2+X3, data=dat)            # multiple linear regression
lm(y ~ (X1+X2+X3)^2, data=dat)        # all two-way interactions
lm(log(y) ~ poly(X,3), data=dat)      # arbitrary transformations
```

  - ANOVA/ANCOVA models (A, B, ... factors)

```
lm(y ~ A)              # one way ANOVA
lm(y ~ A*B)            # two way: A + B + A:B
lm(y ~ X + A)          # one way ANCOVA
lm(y ~ (A+B+C)^2)      # 3-way ANOVA: A, B, C, A:B, A:C, B:C
```


Fitting linear models in R: lm()

- Multivariate models: **`lm()`** with 2+ y vars
  - Multivariate regression

```
lm(cbind(y1, y2) ~ X1 + X2 + X3)              # std MMreg: all linear
lm(cbind(y1, y2) ~ poly(X1,2) + poly(X2,2))   # response surface
```

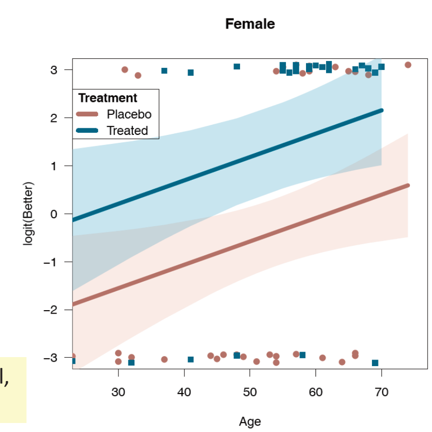
  - MANOVA/MANCOVA models

```
lm(cbind(y1, y2, y3) ~ A * B)     # 2-way MANOVA: A + B + A:B
lm(cbind(y1, y2, y3) ~ X + A)     # MANCOVA (equal slopes)
lm(cbind(y1, y2) ~ X + A + X:A)   # heterogeneous slopes
```


General*ized* Linear Models: glm()

Transformations of y & other error distributions

- y ∈ (0/1): lived/died; success/fail; …
- logit (log odds) model:
  - $\text{logit}(y) = \log \frac{\Pr(y=1)}{\Pr(y=0)}$
  - linear logit model: $\text{logit}(y) = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \ldots$



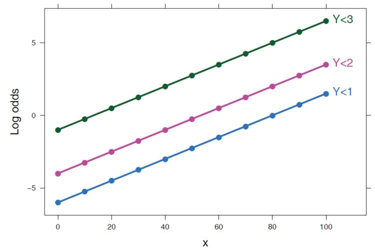
glm(better ~ age + treat, family=binomial, data=Arthritis)

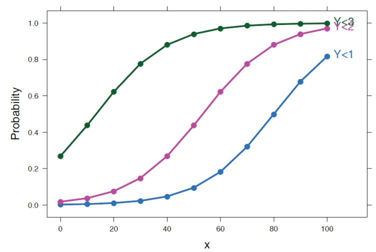
Generalized Linear Models

Ordinal responses

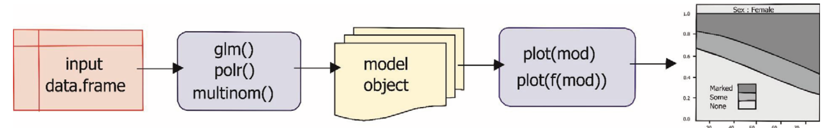
- Improved ∈ (“None” < “Some” < “Marked”)
- Models: Proportional odds, generalized logits, …

```
library(MASS)
polr(Improved ~ Sex + Treat + Age,
      data=Arthritis)

library(nnet)
multinom(Improved ~ Sex + Treat + Age,
          data=Arthritis)
```





Model-based methods: Overview



- models in R are specified by a symbolic model formula, applied to a data.frame
  - mod<-lm(prestige ~ income + educ, data=Prestige)
  - mod<-glm(better ~ age + sex + treat, data=Arthritis, family=binomial)
  - mod<-MASS:polr(improved ~ age + sex + treat, data=Arthritis)
- result (`mod`) is a “model object”, of class “`lm`”, “`glm`”, …
- method functions:
  - plot(mod), plot(f(mod)), …
  - summary(mod), coef(mod), predict(mod), …

Plots for linear models

- Data plots:
  - plot response (y) vs. predictors, with smooth summaries
  - scatterplot matrix --- all pairs
- Model (effect) plots
  - plot predicted response ($\hat{y}$) vs. predictors, controlling for variables not shown.
- Diagnostic plots
  - Influence plots: leverage & outliers
  - Spread-level plots (non-constant variance?)

R packages

- `car`
  - Enhanced scatterplots
  - Diagnostic plots
- `effects`
  - Plot fitted effects of one predictor, controlling all others
- `visreg`
  - similar to effect plots, simpler syntax
- Both `effects` & `visreg` handle nearly all formula-based models
  - lm(), glm(), gam(), rlm(), nlme(), …

Occupational Prestige data

- Data on prestige of 102 occupations and
  - average education (years)
  - average income ($)
  - % women
  - type (Blue Collar, Professional, White Collar)

```
> car::some(Prestige, 6)
                     education income women prestige census type
architects               15.44  14163  2.69     78.1   2141 prof
physicians               15.96  25308 10.56     87.2   3111 prof
commercial.artists       11.09   6197 21.03     57.2   3314 prof
tellers.cashiers         10.64   2448 91.76     42.3   4133   wc
bakers                    7.54   4199 33.30     38.9   8213   bc
aircraft.workers          8.78   6573  5.78     43.7   8515   bc
```


Informative scatterplots

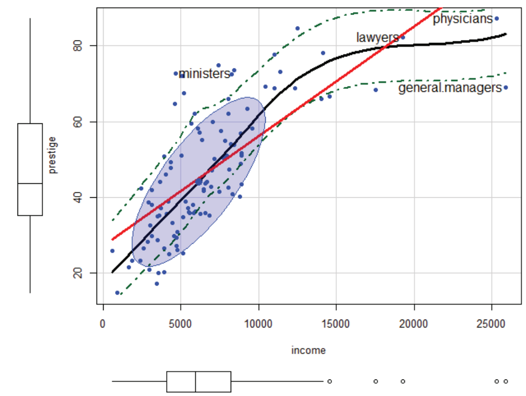
Scatterplots are most useful when enhanced with annotations & statistical summaries

Data ellipse and regression line show the linear model, prestige ~ income

Point labels show possible outliers

Smoothed (loess) curve and CI show the trend

Boxplots show marginal distributions



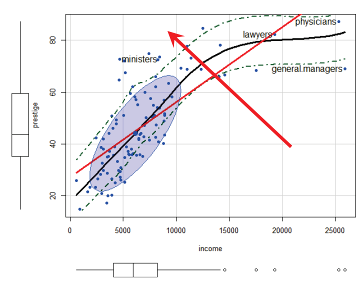
Informative scatterplots

`car::scatterplot()` provides all these enhancements

```
scatterplot(prestige ~ income, data=Prestige,
            pch = 16,
            regLine = list(col = "red", lwd=3),
            smooth = list(smoother=loessLine,
                          lty.smooth = 1, col.smooth = "black",
                          lwd.smooth=3, col.var = "darkgreen"),
            ellipse = list(levels = 0.68),
            id = list(n=4, col="black", cex=1.2))
```

Skewed distribution of income & non-linear relation suggest need for a transformation

Arrow rule: move on the scale of powers in direction of the bulge
e.g.: x → sqrt(income) or log(income)



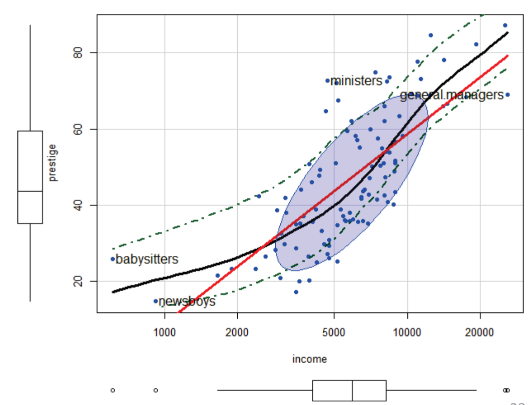
Try log(income)

```
scatterplot(prestige ~ income, data=Prestige,
            log = "x",
            pch = 16,
            regLine = list(col = "red", lwd=3),
            … )
```

Income now ~ symmetric

Relation closer to linear

log(income): interpret as effect of a multiple



Stratify by type?

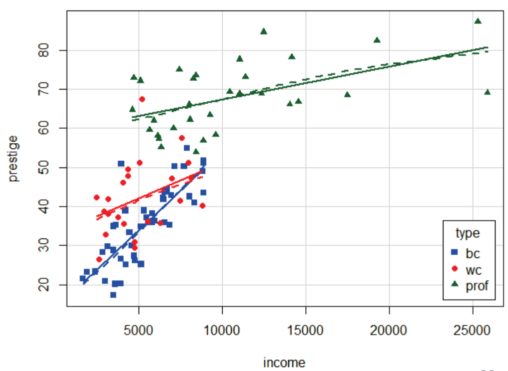
```
scatterplot(prestige ~ income | type, data=Prestige,
      col = c("blue", "red", "darkgreen"),
      pch = 15:17,
      legend = list(coords="bottomright"),
      smooth=list(smoother=loessLine, var=FALSE, span=1, lwd=4))
```

Formula: | type → “given type”

Different slopes: interaction of income * type

Provides another explanation of the non-linear relation

This may be a new finding!



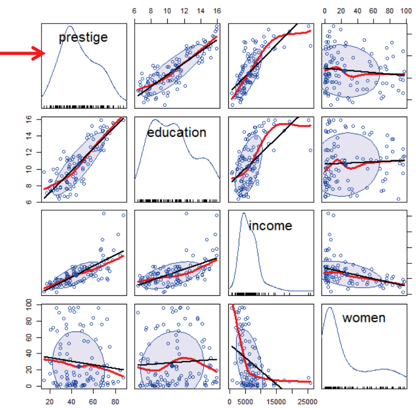
Scatterplot matrix

```
scatterplotMatrix(~ prestige + education + income + women ,
        data=Prestige,
        regLine = list(method=lm, lty=1, lwd=2, col="black"),
        smooth=list(smoother=loessLine, spread=FALSE,
                lty.smooth=1, lwd.smooth=3, col.smooth="red"),
        ellipse=list(levels=0.68, fill.alpha=0.1))
```

prestige vs. all predictors

diagonal: univariate distributions
- income: + skewed
- %women: bimodal

off-diagonal: relations among predictors



Fit a model

```
> mod1 <- lm(prestige ~ education + poly(women, 2) +
+                    log(income)*type, data=Prestige)
> summary(mod1)

Coefficients:
                     Estimate Std. Error t value Pr(>|t|)
(Intercept)          -137.500     23.522   -5.85  8.2e-08 ***
education               2.959      0.582    5.09  2.0e-06 ***
poly(women, 2)1        28.339     10.190    2.78   0.0066 **
poly(women, 2)2        12.566      7.095    1.77   0.0800 .
log(income)            17.514      2.916    6.01  4.1e-08 ***
typeprof               74.276     30.736    2.42   0.0177 *
typewc                  0.969     39.495    0.02   0.9805
log(income):typeprof   -7.698      3.451   -2.23   0.0282 *
log(income):typewc     -0.466      4.620   -0.10   0.9199
---
Signif. codes:  0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’
1

Multiple R-squared:  0.879,  Adjusted R-squared:  0.868
F-statistic: 81.1 on 8 and 89 DF,  p-value: <2e-16
```

- allow women² term
- interaction of log(income) and type

Fits very well!

Model (effect) plots

- We’d like to see plots of the predicted value ($\hat{y}$) of the response against predictors (x_j)
  - Ordinary plot of y vs. x_j doesn’t allow for other correlations
  - → Must control for other predictors (x_{-j}) not shown in a given plot
  - Variables not shown (x_{-j}) are averaged over.
  - Slopes of lines reflect the partial coefficient in the model
  - Partial residuals can be shown also

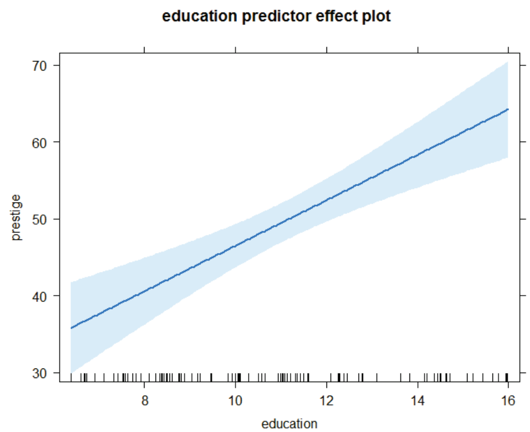
For details, see `vignette(“predictor-effects-gallery”, package=“effects”)`

Model (effect) plots: education

```
library("effects")
mod1.e1 <- predictorEffect("education", mod1)
plot(mod1.e1)
```

This graph shows the partial slope for education

For each ↑ year in education, fitted prestige ↑2.96 points, (other predictors held fixed)

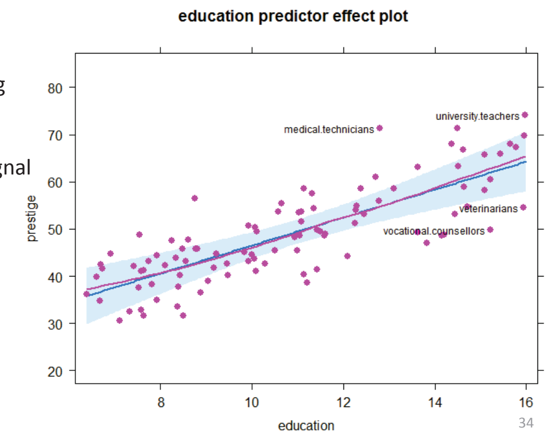


Model (effect) plots

```
mod1.e1a <- predictorEffect("education", mod1, residuals=TRUE)
plot(mod1.e1a,
     residuals.pch=16, id=list(n=4, col="black"))
```

Partial residuals show the residual of prestige controlling for other predictors

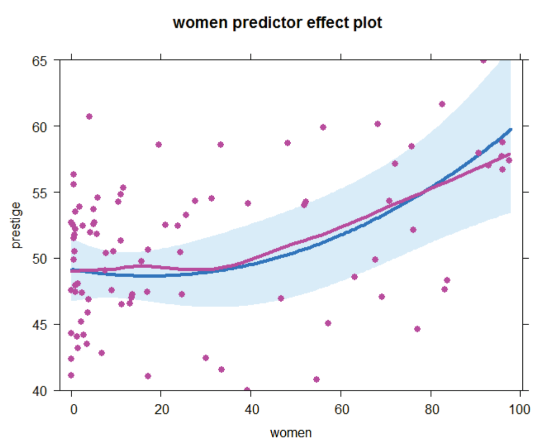
Unusual points here would signal undue influence



Model (effect) plots: women

```
mod1.e2 <- predictorEffect("women", mod1, residuals=TRUE)
plot(mod1.e2, ylim=c(40, 65), lwd=4,
     residuals.pch=16)
```

Surprise!
Prestige of occupations ↑ with % women (controlling for other variables)

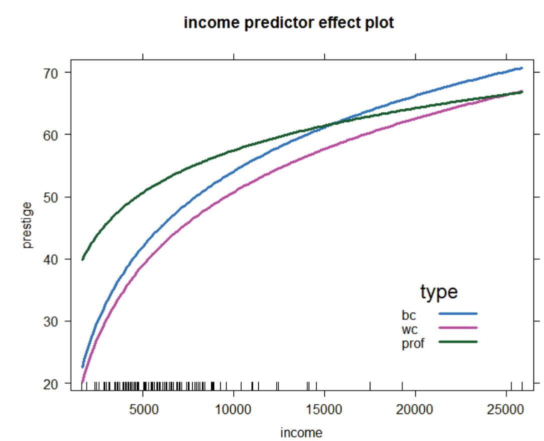


Model (effect) plots: income

```
plot(predictorEffect("income", mod1),
     lines=list(multiline=TRUE, lwd=3),
     key.args = list(x=.7, y=.35))
```

Income interacts with type in the model

The plot is curved because log(income) is in the model



visreg plots: Air quality data

Daily air quality measurements in New York, May - Sep 1973

How does Ozone concentration vary with solar radiation, wind speed & temperature?

```
> head(airquality)
  Ozone Solar.R Wind Temp Month Day
1    41     190  7.4   67     5   1
2    36     118  8.0   72     5   2
3    12     149 12.6   74     5   3
4    18     313 11.5   62     5   4
5    NA      NA 14.3   56     5   5
6    28      NA 14.9   66     5   6
```

see: https://pbreheny.github.io/visreg/ for examples & details

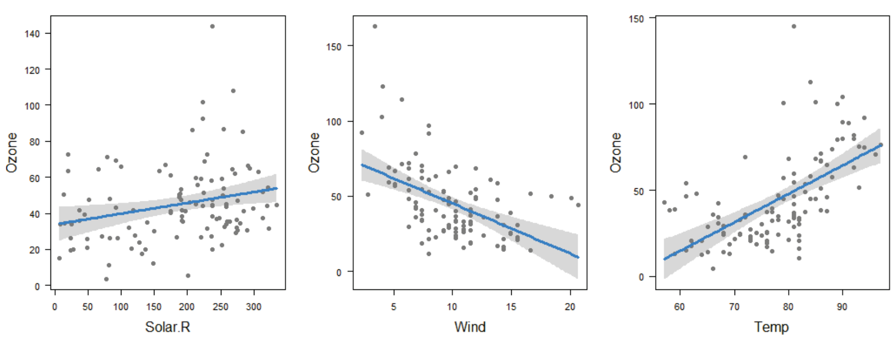
Air quality: main effects model

```
> fit1 <- lm(Ozone ~ Solar.R + Wind + Temp, data=airquality)
> summary(fit1)

Coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) -64.3421    23.0547   -2.79   0.0062 **
Solar.R       0.0598     0.0232    2.58   0.0112 *
Wind         -3.3336     0.6544   -5.09  1.5e-06 ***
Temp          1.6521     0.2535    6.52  2.4e-09 ***
---
Signif. codes:  0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 ‘.’ 0.1 ‘ ’ 1

Residual standard error: 21.18 on 107 degrees of freedom
  (42 observations deleted due to missingness)
Multiple R-squared:  0.6059,    Adjusted R-squared:  0.5948
F-statistic: 54.83 on 3 and 107 DF,  p-value: < 2.2e-16
```


visreg conditional plots

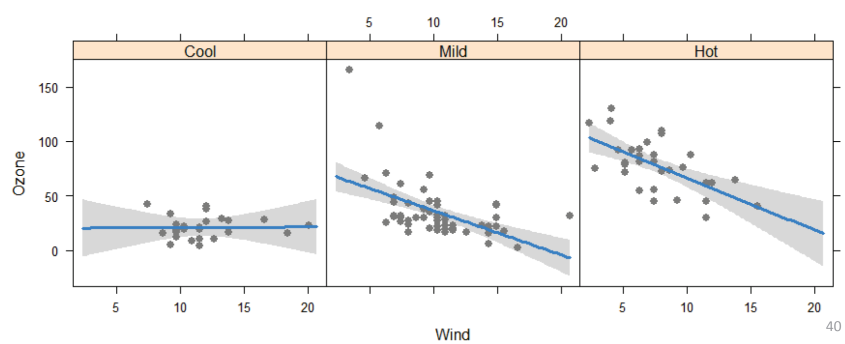
```
op <- par(mfrow=c(1,3), cex.lab=1.5)
visreg(fit1, "Solar.R“)
visreg(fit1, "Wind“)
visreg(fit1, "Temp")
par(op)
```

model summary =
- predicted values (line) +
- confidence band (uncertainty) +
- partial residuals (objections)



Factor variables & interactions

```
                                  # cut Temp into three ordered levels of equal range
airquality$Heat <- cut(airquality$Temp, 3,
                  labels=c("Cool","Mild","Hot"))

                                  # fit model with interaction  of Wind * Heat
fit2 <- lm(Ozone ~ Solar.R + Wind*Heat, data=airquality)
visreg(fit2, "Wind", by="Heat", layout=c(3,1), points=list(cex=1))
```



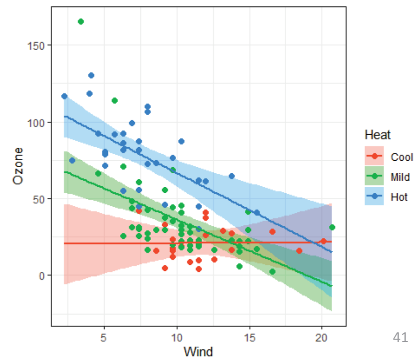
Factor variables & interactions

```
visreg(fit2, "Wind", by="Heat",
       overlay=TRUE,
          gg=TRUE,
       points=list(size=2)) +
  theme_bw()
```

overlay=TRUE → superpose panels
gg=TRUE → uses ggplot

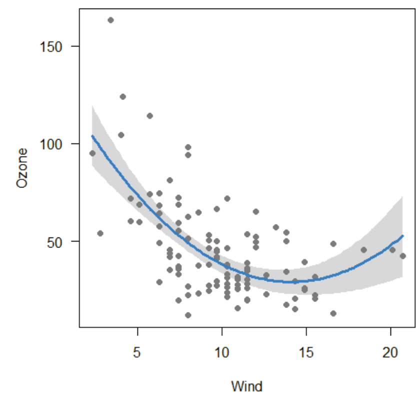
This allow slope for Wind to vary with Heat e.g., Wind has no effect when Cool

This model still assumes linear effects of Heat & Wind

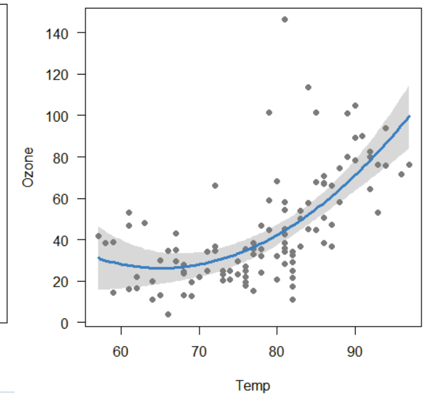


Non-linear effects

fit <- lm(Ozone ~ Solar.R + poly(Wind,2) +
Temp, data=airquality)
visreg(fit, "Wind")



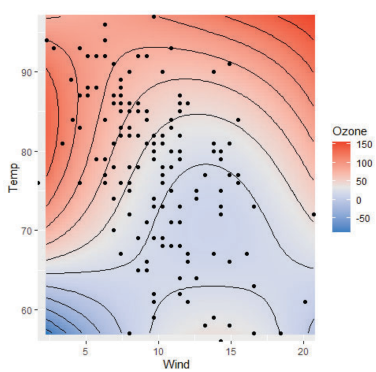
fit <- lm(Ozone ~ Solar.R + Wind +
poly(Temp,2), data=airquality)
visreg(fit, "Temp")



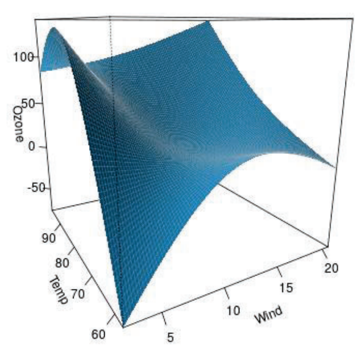
Response surface models (visreg2d)

# Fit quadratics in both Wind & Temp and interaction Wind * Temp
fitp <- lm(Ozone ~ Solar.R + poly(Wind,2) * poly(Temp,2), data=airquality)

visreg2d(fitp, "Wind", "Temp", plot.type="gg") +
geom_contour(aes(z=z), color="black")



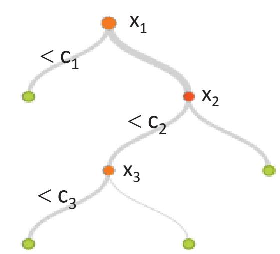
visreg2d(fitp, "Wind", "Temp", plot.type="persp")



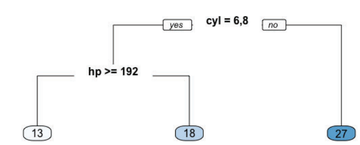
Regression trees

Regression trees are a non-parametric alternative to linear models

- Essential ideas:
  - Find predictor and split value which minimizes SSE
  - fitted value in each subgroup = mean
  - repeat, recursively, splitting by next best predictor
- Large literature
  - cost, complexity tradeoff
  - pruning methods
  - boosting, cross-validation
  - tree averaging



e.g.: mpg ~ cyl + hp

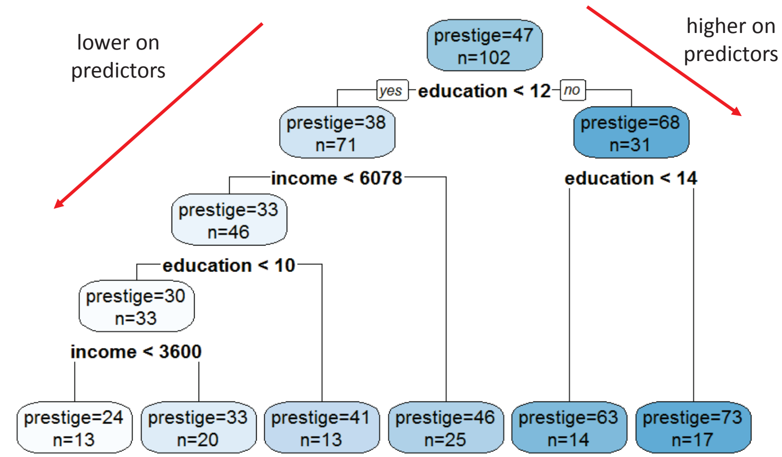


Prestige data: rpart tree

```
> library(rpart)          # calculating regression trees
> library(rpart.plot)     # plotting regression trees

> rmod <- rpart(prestige ~ education + income + women + type,
        data=Prestige,
        method = "anova")

> rpart.rules(rmod)       # print prediction rules
 prestige
    24 when education <  10         & income <  3600
    33 when education <  10         & income is 3600 to 6078
    41 when education is 10 to 12 & income <  6078
    46 when education <  12         & income >=        6078
    63 when education is 12 to 14
    73 when education >=       14
```


Prestige data: rpart tree

```
rpart.plot(rmod, prefix="prestige=")
```



Diagnostic plots

- The linear model, $\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \epsilon$ assumes:
  - Residuals, ε_i are normally distributed, $\varepsilon_i \sim N(0,\sigma^2)$
  - (Normality not required for **X**s)
  - Constant variance, $\mathrm{Var}(\varepsilon_i) = \sigma^2$
  - Observations y_i are statistically independent
- Violations → inferences may not be valid
- A variety of plots can diagnose all these problems
- Other methods (boxCox, boxTidwell) diagnose the need for transformations of **y** or **X**s.

The "regression quartet"

In R, plotting a `lm` model object → the "regression quartet" of plots

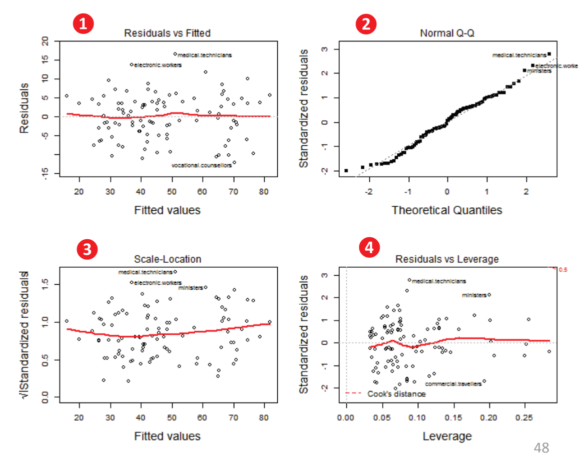
```
plot(mod1, lwd=2, cex.lab=1.4)
```

1 Residuals: should be flat vs. fitted values

2 Q-Q plot: should follow the 45° line

3 Scale-location: should be flat if constant variance

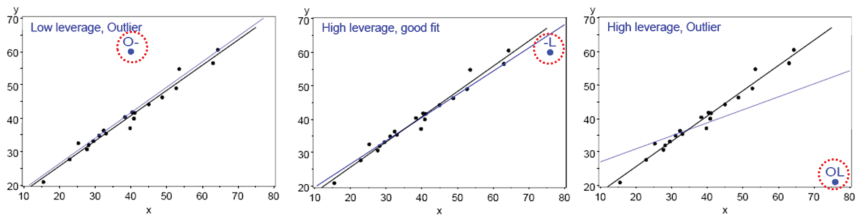
4 Resids vs. leverage: can show influential observations



Unusual data: Leverage & Influence

- "Unusual" observations can have dramatic effects on least-squares estimates in linear models
- Three archetypal cases:
  - Typical X (low leverage), bad fit -- Not much harm
  - Unusual X (high leverage), good fit -- Not much harm
  - Unusual X (high leverage), bad fit -- BAD, BAD, BAD
- Influential observations: unusual in *both* X & Y
- Heuristic formula:

Influence = X leverage x Y residual



Influence plots

Influence (Cook's D) measures impact of individual obs. on coefficients, fitted values

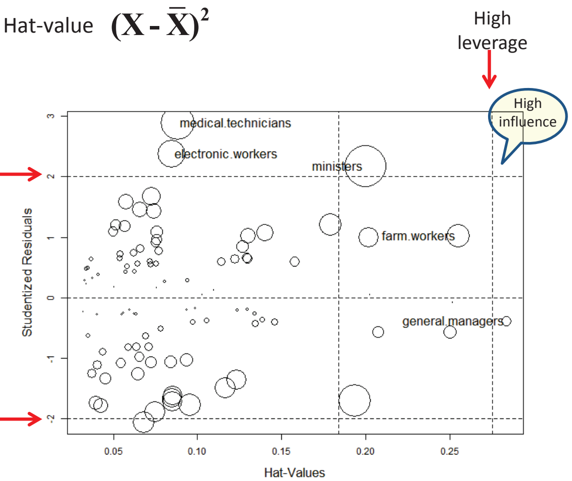
Influence ~ Residual (y -ŷ) × Hat-value $(\mathbf{X} - \bar{\mathbf{X}})^2$

Bubble size ~ influence

```
influencePlot(mod1)
```

Bad fit

Bad fit

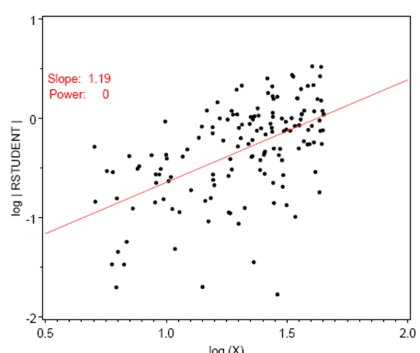


Spread-level plots

- To diagnose non-constant variance, plot:
  - log |Std. residual| vs. log (x)
  - log (IQR) vs log (median) [for grouped data]
- If ≈ linear w/ slope b, transform y → y $^{(1-b)}$

Artificial data, generated so σ ~ x

- b ≈ 1 → power = 0
- → analyze log(y)



Spread-level plot: baseball data

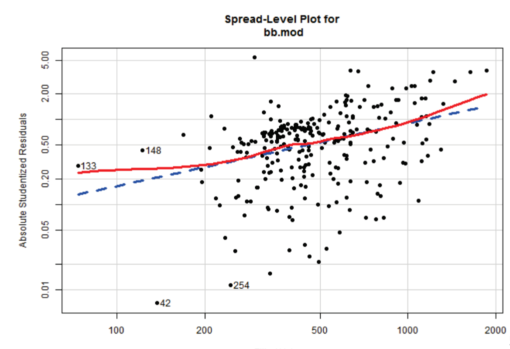
Data on salary and batter performance from 1987 season

```
data("Baseball", package="vcd")
bb.mod <- lm(sal87 ~ years + hits + runs + homeruns, data=Baseball)
spreadLevelPlot(bb.mod, pch=16, lwd=3,
                id=list(n=2))

## Suggested power transformation: 0.2609
```

slope = .74 → p = .26

i.e., y → log(y) or $y^{1/4}$

NB: both axes plotted on log scale



Box Cox transformation

- Box & Cox proposed to transform y to a power, $y \to y^{(\lambda)}$ to minimize the residual SS (or maximize the likelihood)
  - Makes $y^{(\lambda)}$ more nearly normal
  - Makes $y^{(\lambda)}$ more nearly linear in with X

Formula for $y^{(\lambda)}$
- $y^{(0)}$: $\log_e(y)$
- $\lambda < 0$: flip sign to keep same order

$$y_i^{(\lambda)} = \begin{cases} \dfrac{y_i^\lambda - 1}{\lambda} & \text{if } \lambda \neq 0, \\ \ln(y_i) & \text{if } \lambda = 0, \end{cases}$$

Power(p)	Transformation	Name
2	Y^2	Square
1	Y (No transformation)	Original Data
½	√ Y	Square root
"0"	log Y or $\log_{10}$ (Y)	Logarithm
-½	-1 / √ Y	Reciprocal Root
-1	-1 / Y	Reciprocal
-2	-1 / Y^2	Reciprocal Square

Example: Cars93 data

How does gas mileage (MPG.city) depend on vehicle weight?

```
> cars.mod <- lm(MPG.city ~ Weight, Cars93)
> coef(cars.mod)
(Intercept)      Weight
   47.04835    -0.00803
```

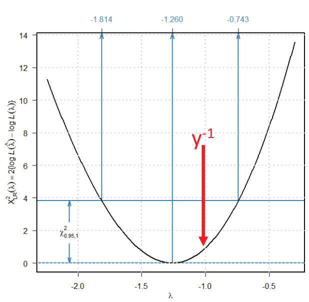
Relationship clearly non-linear

Tukey arrow rule: transform Y (or X) as arrow thru the curve bulges
$y \to \sqrt{y}$, log(y), 1/y
$x \to \sqrt{x}$, log(x), 1/x

MASSextra package

```
> library(MASSExtra)
> box_cox(cars.mod)      # plot log likelihood vs. lambda
> lamba(cars.mod)
[1] -1.26
```

The plot of – log(L) ~ RSS shows the minimum & CI



plot (bc(MPG.city, lamba(cars.mod))

Summary

- Tables are for look-up; graphs can give insight
- "Linear" models include so much more than ANOVA & regression
- Data plots are more effective when enhanced
  - data ellipses → strength & precision of correlation
  - regression lines and smoothed curves
  - point identification → noteworthy observations
- Effect plots show informative views of models
  - Visualize conditional effects, holding others constant
- Diagnostic plots can reveal influential observations and need for transformations.