# Tests of Hypotheses Based on a Single Sample

**8.4** 

# P-Values

Using the rejection region method to test hypotheses entails first selecting a significance level  $\alpha$ .

Then after computing the value of the test statistic, the null hypothesis  $H_0$  is rejected if the value falls in the rejection region and is otherwise not rejected.

We now consider another way of reaching a conclusion in a hypothesis testing analysis.

This alternative approach is based on calculation of a certain probability called a *P-value*.

One advantage is that the P-value provides an intuitive measure of the strength of evidence in the data against  $H_0$ .

#### **Definition**

The *P***-value** is the probability, calculated assuming that the null hypothesis is true, of obtaining a value of the test statistic at least as contradictory to  $H_0$  as the value calculated from the available sample.

This definition is quite a mouthful. Here are some key points:

- The P-value is a probability.
- This probability is calculated assuming that the null hypothesis is true.
- Beware: The P-value is not the probability that  $H_0$  is true, nor is it an error probability!
- To determine the P-value, we must first decide which values of the test statistic are at least as contradictory to  $H_0$  as the value obtained from our sample.

Urban storm water can be contaminated by many sources, including discarded batteries. When ruptured, these batteries release metals of environmental significance.

The article "Urban Battery Litter" (*J. of Environ. Engr.*, 2009: 46–57) presented summary data for characteristics of a variety of batteries found in urban areas around Cleveland.

A sample of 51 Panasonic AAA batteries gave a sample mean zinc mass of 2.06g and a sample standard deviation of .141g.

Does this data provide compelling evidence for concluding that the population mean zinc mass exceeds 2.0g?

With  $\mu$  denoting the true average zinc mass for such batteries, the relevant hypotheses are  $H_0$ :  $\mu$  = 2.0 versus  $H_a$ :  $\mu$  > 2.0.

The sample size is large enough so that a z test can be used without making any specific assumption about the shape of the population distribution.

The test statistic value is

$$z = \frac{\overline{x} - 2.0}{s/\sqrt{n}} = \frac{2.06 - 2.0}{.141/\sqrt{51}} = 3.04$$

Now we must decide which values of z are at least as contradictory to  $H_0$ .

Let's first consider an easier task:

Which values of *x* are at least as contradictory to the null hypothesis as 2.06, the mean of the observations in our sample?

3.00

Ho = 2.0

Because > appears in  $H_a$ , it should be clear that 2.10 is at least as contradictory to  $H_0$  as is 2.06, and so in fact is *any*  $\bar{x}$  value that exceeds 2.06.

But an  $\overline{x}$  value that exceeds 2.06 corresponds to a value of z that exceeds 3.04. Thus the P-value is

*P*-value = 
$$P(Z \ge 3.04 \text{ when } \mu = 2.0)$$

Since the test statistic Z was created by subtracting the null value 2.0 in the numerator, when  $\mu = 2.0$ —i.e., when  $H_0$  is true—Z has approximately a standard normal distribution.

As a consequence,

*P*-value = 
$$P(Z \ge 3.04 \text{ when } \mu = 2.0)$$

 $\approx$  area under the z curve to the right of 3.04

$$=1-\Phi(3.04)$$

We will shortly illustrate how to determine the P-value for any z or t test—i.e., any test where the reference distribution is the standard normal distribution (and z curve) or some t distribution (and corresponding t curve).

For the moment, though, let's focus on reaching a conclusion once the *P*-value is available.

Because it is a probability, the *P*-value must be between 0 and 1.

What kinds of *P*-values provide evidence against the null hypothesis?

Consider two specific instances:

• P-value = .250: In this case, fully 25% of all possible test statistic values are at least as contradictory to  $H_0$  as the one that came out of our sample. So our data is not all that contradictory to the null hypothesis.

• P-value = .0018: Here, only .18% (much less than 1%) of all possible test statistic values are at least as contradictory to  $H_0$  as what we obtained. Thus the sample appears to be highly contradictory to the null hypothesis.

More generally, the smaller the P-value, the more evidence there is in the sample data against the null hypothesis and for the alternative hypothesis. That is,  $H_0$  should be rejected in favor of  $H_a$  when the P-value is sufficiently small. So what constitutes "sufficiently small"?

#### Decision rule based on the P-value

Select a significance level  $\alpha$  (as before, the desired type I error probability).

Then

reject  $H_0$  if P-value  $\leq \alpha$ 

do not reject  $H_0$  if P-value >  $\alpha$ 

Thus if the *P*-value exceeds the chosen significance level, the null hypothesis cannot be rejected at that level.

But if the P-value is equal to or less than  $\alpha$ , then there is enough evidence to justify rejecting  $H_0$ .

In Example 14, we calculated P-value = .0012. Then using a significance level of .01, we would reject the null hypothesis in favor of the alternative hypothesis because .0012  $\leq$  .01.

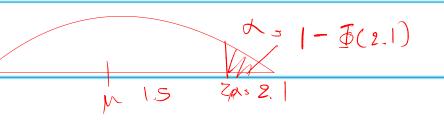
However, suppose we select a significance level of only .001, which requires more substantial evidence from the data before  $H_0$  can be rejected.

In this case we would not reject  $H_0$  because .0012  $\leq$  .001.

How does the decision rule based on the *P*-value compare to the decision rule employed in the rejection region approach?

The two procedures—the rejection region method and the P-value method—are in fact identical.

Whatever the conclusion reached by employing the rejection region approach with a particular  $\alpha$ , the same conclusion will be reached via the P-value approach using that same  $\alpha$ .



The nicotine content problem involved testing  $H_0$ :  $\mu = 1.5$  versus  $H_a$ :  $\mu > 1.5$  using a z test (i.e., a test which utilizes the z curve as the reference distribution).

The inequality in  $H_a$  implies that the upper-tailed rejection region  $z \ge z_{\alpha}$  is appropriate.

$$\sum$$

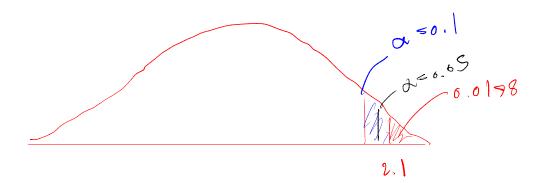
Suppose z=2.10. Then using exactly the same reasoning as in Example 14 gives P-value =  $1-\Phi(2.10)=.0179$ .

Consider now testing with several different significance levels:

$$\alpha = .10 \Rightarrow z_{\alpha} = z_{10} = 1.28 \Rightarrow 2.10 \ge 1.28 \Rightarrow \text{ reject } H_0$$

$$\alpha = .05 \Rightarrow z_{\alpha} = z_{.05} = 1.645 \Rightarrow 2.10 \ge 1.645 \Rightarrow \text{reject } H_0$$

$$\alpha = .01 \Rightarrow z_{\alpha} = z_{.01} = 2.33 \Rightarrow 2.10 < 2.33 \Rightarrow \text{do not reject } H_0$$



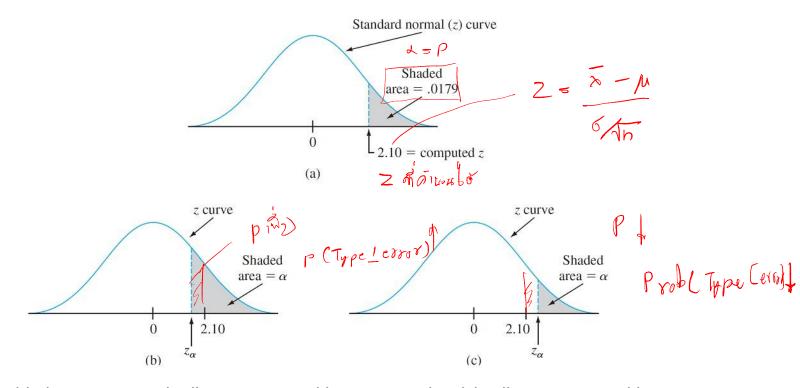
Because P-value = .0179  $\leq$  .10 and also .0179  $\leq$  .05, using the P-value approach results in rejection of  $H_0$  for the first two significance levels.

However, for  $\alpha$  = .01, 2.10 is not in the rejection region and .0179 is larger than .01.

More generally, whenever  $\alpha$  is smaller than the P-value .0179, the critical value  $z_{\alpha}$  will be beyond the calculated value of z and  $H_0$  cannot be rejected by either method.

This is illustrated in Figure 8.7.





Relationship between  $\alpha$  and tail area captured by computed z: (a) tail area captured by computed z; (b) when  $\alpha$  < .0179,  $z_{\alpha}$  < 2.10 and  $H_0$  is rejected; (c) when  $\alpha$  < .0179,  $z_{\alpha}$  < 2.10 and  $H_0$  is not rejected

Figure 8.7

#### **Proposition**

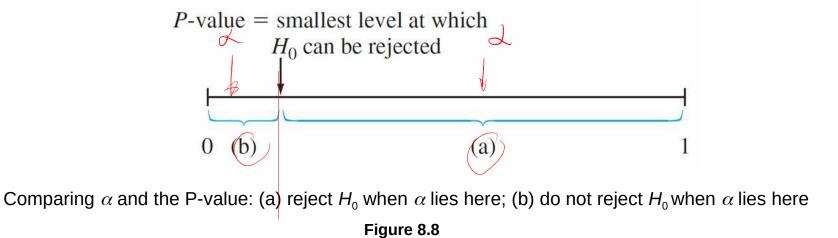
The P-value is the smallest significance level  $\alpha$  at which the null hypothesis can be rejected.

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Because of this, the P-value is alternatively referred to as the **observed significance level** (OSL) for the data. It is customary to call the data *significant* when  $H_0$  is rejected and *not significant* otherwise.

The *P*-value is then the smallest level at which the data is significant.

An easy way to visualize the comparison of the *P*-value with the chosen a is to draw a picture like that of Figure 8.8.



The calculation of the P-value depends on whether the test is upper-, lower-, or two-tailed. However, once it has been calculated, the comparison with  $\alpha$  does not depend on which type of test was used.

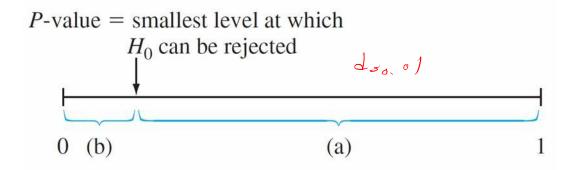
The true average time to initial relief of pain for a best-selling pain reliever is known to be 10 min.

Let  $\mu$  denote the true average time to relief for a company's newly developed reliever.

Suppose that when data from an experiment involving the new pain reliever is analyzed, the P-value for testing  $H_0$ :  $\mu = 10$  versus  $H_a$ :  $\mu < 10$  is calculated as .0384.

> 6.0384

Since  $\alpha$  = .05 is larger than the *P*-value [.05 lies in the interval (a) of Figure 8.8],  $H_0$  would be rejected by anyone carrying out the test at level .05.



Comparing  $\alpha$  and the P-value: (a) reject  $H_0$  when  $\alpha$  lies here; (b) do not reject  $H_0$  when  $\alpha$  lies here Figure 8.8

However, at level .01,  $H_0$  would not be rejected because .01 is smaller than the smallest level (.0384) at which  $H_0$  can be rejected.

The P-value for a z test (one based on a test statistic whose distribution when  $H_0$  is true is at least approximately standard normal) is easily determined from the information in Appendix Table A.3.

Consider an upper-tailed test and let z denote the computed value of the test statistic Z.

The null hypothesis is rejected if  $z \ge z_{\alpha}$ , and the *P*-value is the smallest a for which this is the case. Since  $z_{\alpha}$  increases as  $\alpha$  decreases, the *P*-value is the value of a for which  $z = z_{\alpha}$ .

That is, the *P*-value is just the area captured by the computed value *z* in the upper tail of the standard normal curve.

The corresponding cumulative area is  $\Phi(z)$ , so in this case P-value =  $1 - \Phi(z)$ .

An analogous argument for a lower-tailed test shows that the P-value is the area captured by the computed value z in the lower tail of the standard normal curve.

More care must be exercised in the case of a two-tailed test. Suppose first that z is positive. Then the P-value is the value of  $\alpha$  satisfying  $z = z_{\alpha/2}$  (i.e., computed z = upper-tail critical value).

This says that the area captured in the upper tail is half the P-value, so that P-value =  $2[1 - \Phi(z)]$ .

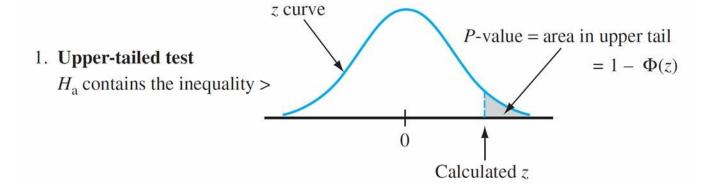
If z is negative, the P-value is the  $\alpha$  for which  $z = -z_{\alpha/2}$ , or, equivalently,  $-z = z_{\alpha/2}$  so that P-value =  $2[1 - \Phi(-z)]$ .

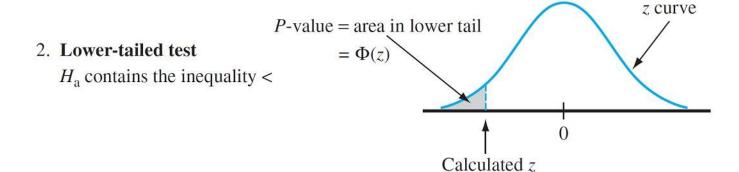
Since -z = |z| when z is negative, P-value =  $2[1 - \Phi(|z|)]$  for either positive or negative z.

$$P\text{-value: } P = \begin{cases} 1 - \Phi(z) & \text{for an upper-tailed } z \text{ test} \\ \Phi(z) & \text{or an lower-tailed } z \text{ test} \\ 2[1 - \Phi(|z|)] & \text{for a two-tailed } z \text{ test} \end{cases}$$

Each of these is the probability of getting a value at least as extreme as what was obtained (assuming  $H_0$  true).

The three cases are illustrated in Figure 8.9.





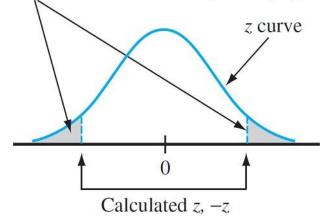
Determination of the P-value for a z test

Figure 8.9

cont'd

*P*-value = sum of area in two tails =  $2[1 - \Phi(|z|)]$ 

## 3. Two-tailed test $H_a$ contains the inequality $\neq$



Determination of the P-value for a z test Figure 8.9

The target thickness for silicon wafers used in a certain type of integrated circuit is 245  $\mu$ m.

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A sample of 50 wafers is obtained and the thickness of each one is determined, resulting in a sample mean thickness of 246.18  $\mu$ m and a sample standard deviation of 3.60  $\mu$ m.

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Does this data suggest that true average wafer thickness is something other than the target value?

- 1. Parameter of interest:  $\mu$  = true average wafer thickness
- 2. Null hypothesis:  $H_0$ :  $\mu$  = 245
- 3. Alternative hypothesis:  $H_a$ :  $\mu \neq 245$
- 4. Formula for test statistic value:  $z = \frac{\overline{x} 245}{s/\sqrt{n}}$
- 5. Calculation of test statistic value:  $z = \frac{246.18 245}{3.60/\sqrt{50}} = 2.32$

- 6. Determination of P-value: Because the test is two-tailed, P-value =  $2(1 \Phi(2.32)) = .0204$
- 7. Conclusion: Using a significance level of .01,  $H_0$  would not be rejected since .0204 > .01.

At this significance level, there is insufficient evidence to conclude that true average thickness differs from the target value.