

# ARE213 Problem Set #2B

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## Part A: Preliminaries

### (i) Comparison between TU and control States

**Starting with simple comparisons** We begin with simple comparisons between the dependent outcome of interest, the natural logarithm of traffic fatalities per capita ( $\log(\text{fatalities per capita})$ ), between a predefined composite treatment state “TU” (or, state #99), and all of the potential control states. The mean over the period before primary seatbelt laws were adopted in the treatment state is -1.4 and the mean for the control states is -1.7, indicating approximately a 30% lower typical fatalities rate in the treatment state than the average control state (even before the primary seatbelt law “treatment”). The trends for both shown in Figure 1 show that overall the fatalities were on the decline in both places before the treatment period.

**Roadmap** Extracting meaningful conclusions from these data is the goal of our analysis, which will require identifying the variation in traffic fatalities that can be attributed to seat belt laws. Confounding our analysis is the fact that these data are not in the context of an RCT but are from the “real world” with messy trends and linked systems that determine outcomes. We will be applying the synthetic controls method to identify a fleet of control states as a meaningful counterfactual to measure against for our composite treatment state.

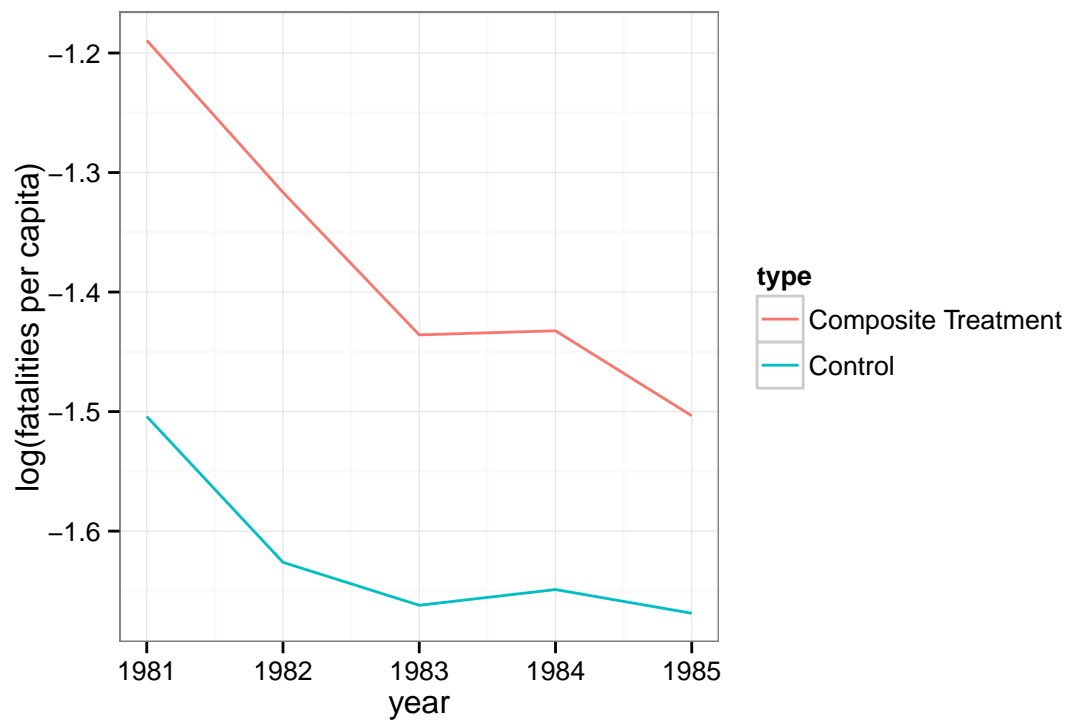


Figure 1: Trend in the dependent variable ( $\log(\text{fatalities per capita})$ ) for the composite treatment state and the average of the control states.

## (ii) “Best” control state comparison

**Sweet Home Alabama** We observe that Alabama is the best match for the composite treatment state based on a simple comparison of  $\log(\text{fatalities per capita})$  in the year before treatment in the composite state (1985). Figure 2 below shows the distribution in the dependent variable

**Fried green covariates and other stereotypes confirmed** Tables 1 and 2 compare the covariates for the composite treatment state and Alabama. There are broad differences between the states. Alabama has higher precipitation, lower college achievement, lower alcohol consumption, higher unemployment, etc. Additionally, the mean value for the dependent variable of interest,  $\log(\text{fatalities per capita})$ , is quite different for the two states. Examining the trends in the covariates (and dependent variable) for the two states (see Figure 3) shows that it could be construed as a coincidence that Alabama is the best match for the value of the dependent variable, since the trajectory in fatalities for both states are following opposite trends in that time and 1985 happens to be the time when they intersect. There are also important and long-term differences in precipitation and alcohol consumption.

Overall Alabama does not appear to be a particularly good match for the composite treatment state, motivating an application of synthetic controls methods to produce a better match.

## Part B: Synthetic Controls

### (i) Why synthetic controls?

**Unsweet Home Alabama:** We saw earlier the difficulties in selecting an exact counterfactual match for implementing differences in differences type selection on unobservables techniques. While Alabama would appear on face value to be a good match (based on having similar outcomes in the year prior to treatment) we saw that this was coincidental and that the covariates are not a good match to the composite treatment state. Synthetic control methods are motivated by producing a “better” match by combining (synthesizing) multiple control states in a weighting scheme to create a composite control state with better match of the important covariates and dependent variable than any particular control state.

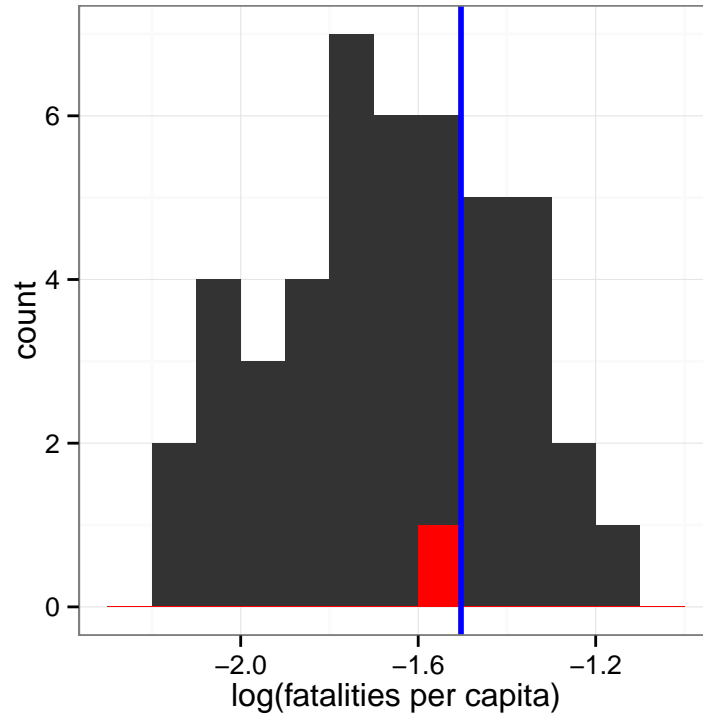


Figure 2: Distribution in traffic fatalities metric from 1985 for all control states with a vertical blue line indicating the value of the metric for the composite treatment state. The red block highlights the position of Alabama in the distribution. Alabama is the closest match to the composite treatment state for 1985, but as is shown here is one of about 11 states that is within 10% of the target value.

Table 1: Composite Treatment Group Summary

| Statistic   | N  | Mean          | St. Dev.   | Min        | Max         |
|-------------|----|---------------|------------|------------|-------------|
| state       | 23 | 99.000        | 0.000      | 99         | 99          |
| year        | 23 | 1,992.000     | 6.782      | 1,981      | 2,003       |
| college     | 23 | 0.234         | 0.014      | 0.209      | 0.259       |
| beer        | 23 | 1.507         | 0.074      | 1.394      | 1.670       |
| primary     | 23 | 0.783         | 0.422      | 0          | 1           |
| secondary   | 23 | 0.000         | 0.000      | 0          | 0           |
| population  | 23 | 13,597.660    | 1,813.520  | 10,737.810 | 16,862.220  |
| unemploy    | 23 | 6.085         | 1.124      | 3.855      | 8.014       |
| fatalities  | 23 | 2,619.014     | 258.667    | 2,246.977  | 3,268.613   |
| totalvmt    | 23 | 128,099.600   | 26,447.260 | 86,013.140 | 170,407.300 |
| precip      | 23 | 2.502         | 0.289      | 1.990      | 3.104       |
| snow32      | 23 | 0.143         | 0.058      | 0.013      | 0.270       |
| rural_speed | 23 | 63.443        | 6.568      | 55.000     | 72.886      |
| urban_speed | 23 | 59.184        | 5.858      | 55.000     | 67.138      |
| logfatalpc  | 23 | -1.643        | 0.168      | -1.805     | -1.189      |
| sqyears     | 23 | 3,968,108.000 | 27,020.830 | 3,924,361  | 4,012,009   |

Table 2: Closest match for pre-policy fatalities: Alabama

| Statistic   | N  | Mean          | St. Dev.   | Min       | Max       |
|-------------|----|---------------|------------|-----------|-----------|
| state       | 23 | 1.000         | 0.000      | 1         | 1         |
| year        | 23 | 1,992.000     | 6.782      | 1,981     | 2,003     |
| college     | 23 | 0.170         | 0.029      | 0.131     | 0.220     |
| beer        | 23 | 1.105         | 0.067      | 1.000     | 1.190     |
| primary     | 23 | 0.174         | 0.388      | 0         | 1         |
| secondary   | 23 | 0.304         | 0.470      | 0         | 1         |
| population  | 23 | 4,185.794     | 209.389    | 3,918.533 | 4,501.862 |
| unemploy    | 23 | 7.509         | 2.780      | 4.200     | 14.400    |
| fatalities  | 23 | 1,036.957     | 88.042     | 839       | 1,189     |
| totalvmt    | 23 | 44,826.090    | 10,109.350 | 27,852    | 58,637    |
| precip      | 23 | 4.944         | 0.701      | 3.737     | 6.342     |
| snow32      | 23 | 0.000         | 0.000      | 0         | 0         |
| rural_speed | 23 | 63.696        | 6.255      | 55        | 70        |
| urban_speed | 23 | 58.478        | 4.870      | 55        | 65        |
| logfatalpc  | 23 | -1.398        | 0.079      | -1.543    | -1.286    |
| sqyears     | 23 | 3,968,108.000 | 27,020.830 | 3,924,361 | 4,012,009 |

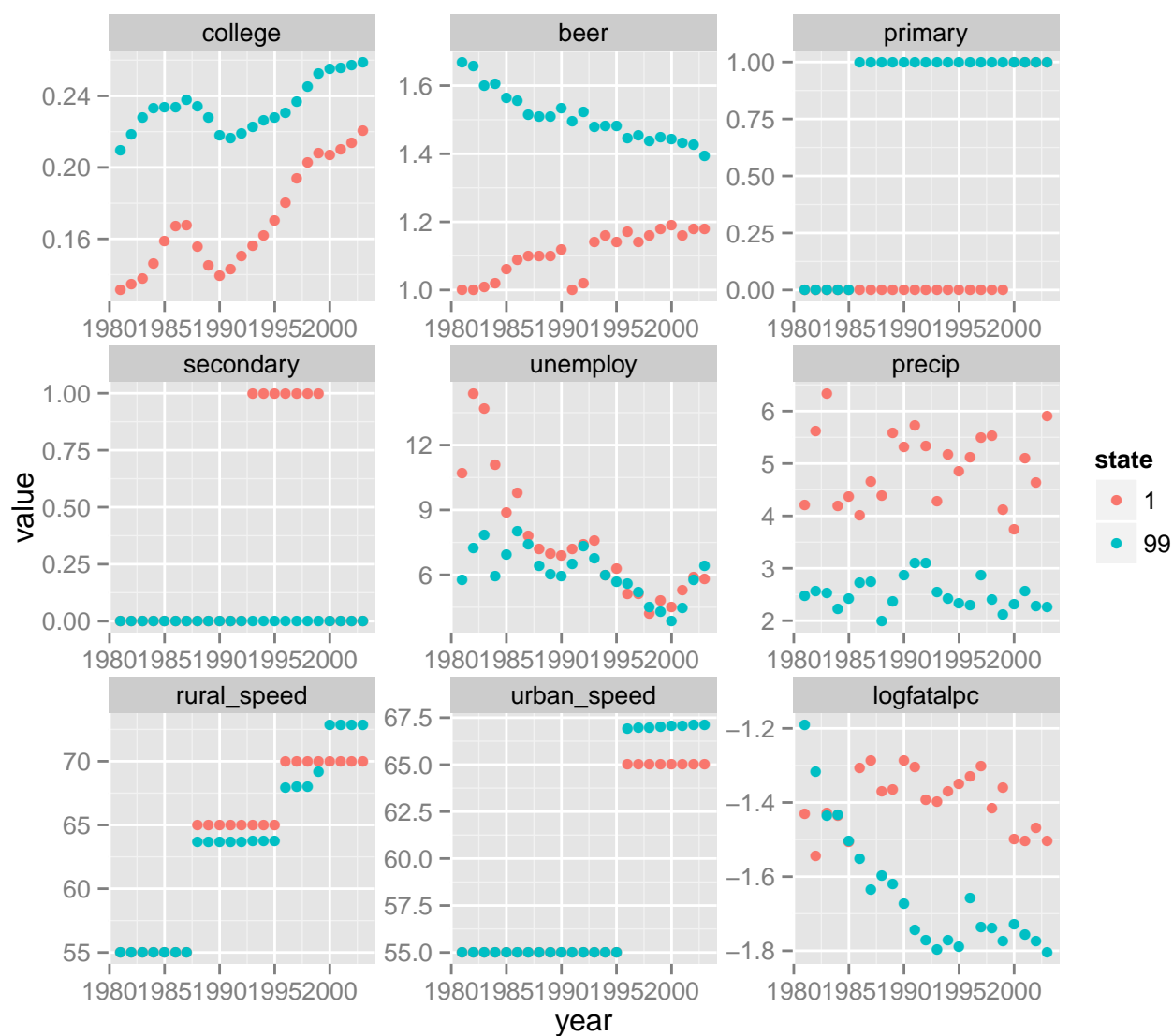


Figure 3: Trends in the covariate (and dependent) variables for the composite treatment state (99) and Alabama (1)

**Dr. Synth-love, or how I learned to stop worrying and love econometrics:** Synthetic controls have a multi-step, iterative process for developing weighting factors to apply to control states for construction of a composite control state. The goal is to identify a weighting matrix  $W$  that minimizes the distance between the treatment covariates (e.g. alcohol consumption, total VMT) and pre-intervention outcomes (log fatalities per capita) for the weighted control unit and the treatment unit. In particular, the following more formally defined criteria are sought (from Synth R package documentation):

- $\sum_{j=2}^{J+1} w_j^* \bar{Y}_j^{K_1} = \bar{Y}_1^{K_M}$ , where  $j$  refers to the state,  $w_j$  is state  $j$ 's weight, and  $\bar{Y}_j^{K_m}$  denotes the pre-treatment outcome in state  $j$  in year  $m$
- $\sum_{j=2}^{J+1} w_j^* U_j = U_1$ , where  $U_i$  is a vector of covariates for state  $i$

Pursuant to these criteria, the synthetic control method estimates the treatment effect as  $\hat{\alpha}_{1t} = Y_{1t} - \sum_{j=2}^{J+1} w_j^* Y_{jt}$

The steps taken in this estimation by the Synth package are as follows:

1. Define a  $(k \times 1)$  matrix (dubbed  $X_1$ ) of the characteristics (covariates  $U_1$  and pre-treatment outcomes  $\bar{Y}_1^{K_m}$ ) of the treatment unit and a similar  $(k \times J)$  matrix (dubbed  $X_J$ ) for the control units.
2. Weight the control characteristics matrix with weight vector  $W$ .
3. Minimize the distance between the treatment unit characteristic matrix and the weighted control characteristics matrix with respect to the weighting matrix. Formally, that's  $\min_W \sqrt{(X_1 - X_0 W)' V (X_1 - X_0 W)}$  where  $V$  is chosen by default to minimize the mean square error of the estimator.

**Pros and Cons:** The upsides to Synthetic control are that one can create a better match for the treatment unit than exists in reality and that it is a method that prevents issues of selection bias (i.e., it is possible to say, "I am using synthetic control" instead of needing to justify ex post the selection of particular units to match in classic differences in differences methods). Another nice feature of the method is the use of graphical placebo testing analysis for determination of the statistical power of results. It is an elegant and compelling way to approach error analysis. A potential methodological



downside is that the method approaches a black box estimate that does not provide much intuition compared to other methods. This may manifest as a lack of trust in results from this method compared to those that are more straightforward to understand.

## (ii) Synthesizing control

The process of creating a synthetic control unit involves 2 steps in the Synth package on [R]. First is specifying the form of the model in a “data prep” step. This is then passed to the synthetic control function to attempt implementing the algorithm described above. In practice we found that errors arise when predictors are included that do not have variation in the mean values among the control units. We used an additive process (adding more and more predictor covariates in the specification) to test whether there is variation. A sub-finding is that the computational intensity increases as covariates are added. This is a relatively small dataset but it is possible that this method could become computationally difficult with large datasets and many covariates. After the process of adding we found that there is variation in all the potentially meaningful covariates except rural and urban speed limits. Since speed limits were constant throughout the sample before 1986 they cannot be included in the synthetic controls specification. Additionally, the presence of secondary seatbelt laws does not vary in the pre-treatment period so is also left out of the potential covariates.

**Preferred specification:** We identified that the following specifications were best for synthetic control analysis of this data:

- Covariates to include in pre-treatment “training” period: Full set of tractable and reasonable covariates. This includes alcohol consumption, VMT per capita, college educational attainment, precipitation, snowfall, and unemployment rate. We tried other combinations of covariates and found little influence on the result. We hoped a VMT-only specification would provide clarity but the gap in the pre-treatment period was biased compared to using the full set of covariates.
- Pre-treatment period: We use the full set of years available in the data, from 1981 to 1985, for the pre-treatment period. We considered dropping 1985 to avoid anticipation effects but this is not done for two

reasons: first, it did not have noticeable impacts on the results (i.e., the divergence between treatment and synthetic control appears between 1985-1986 regardless of whether 1985 is included), and second, because there is a very short pre-treatment period available and we wished to maximize the support of the data.

## Part C: ...but does it work?

### (i) Gap between TU and synth control

We show a series of figures (4 - 9) for various specifications (including the preferred specification) below. Both the gaps (Synthetic unit - Treated unit) and the actual values have been plotted here. These plots appear to show a greater pre-treatment MSPE when only the VMT per capita covariate is used as compared with the full model. Ideally, the difference between the synthetic control unit and the treatment unit should be as close to zero as possible in the pre-treatment period, implying an MSPE approaching zero. For this reason, we select the full model using the aforementioned covariates.

The mean gap is about 0.15 on the log scale, which corresponds to approximately a 15% reduction in traffic fatalities per capita.

### (ii) Placebos are the best medicine

**Placebo test plots:** We developed a series of placebo test plots to investigate the statistical significance of the gap between the treatment and its synthetic control. The series of plots below (Figures 10, 11, 12, 13, and 14) show the implications of different thresholds for keeping control states in the placebo test. The criterion for inclusion is the degree of difference between the pre-treatment (1981-85) MSPE for the placebo control gap vs. the gap for the treatment state. We try a range from basically allowing all the states (50x) to a strict filter (2x). The best compromise between allowing states to participate in the model and keeping out those with bad synthetic control is 10x (Figure 12). In that formulation the treatment state is relatively consistently below nearly all the placebo control states until the end of the period. We discount the results towards the end of the treatment period because it is far beyond the support of the pre-treatment period. These results indicate that there is very likely a significant decrease in fatalities from primary

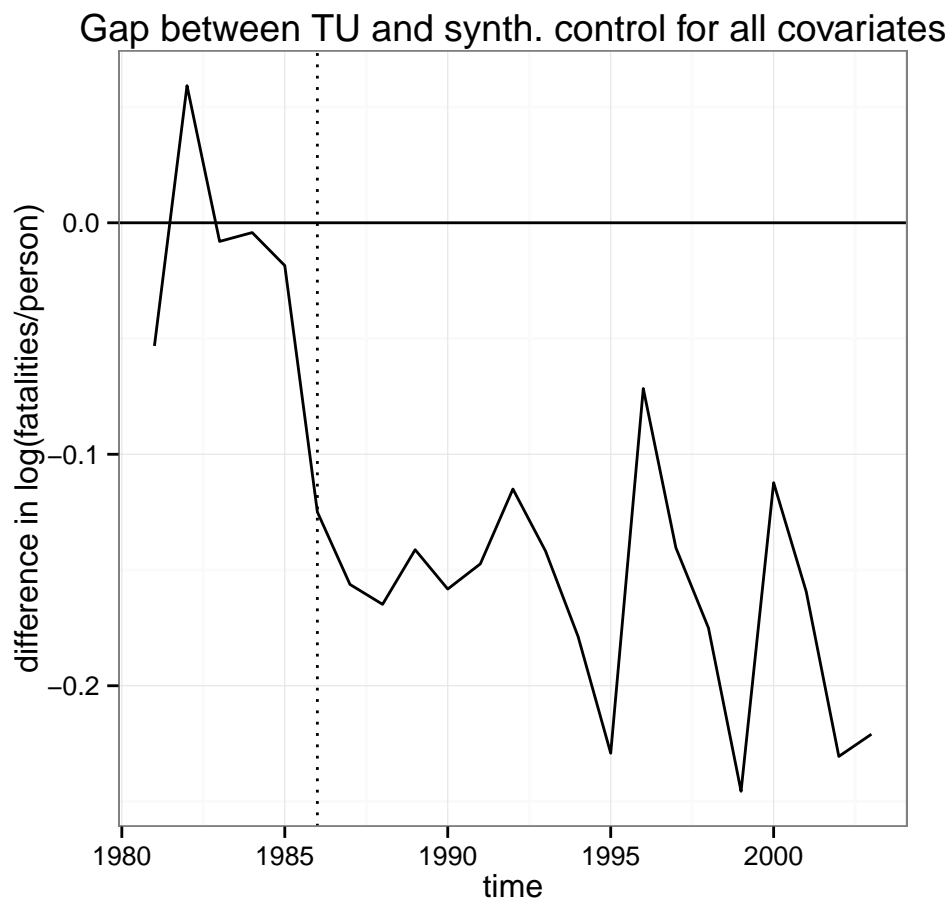


Figure 4: Gap between Treatment Unit and Synthetic Control developed using all previous time periods and all covariates.

Gap between TU and synth. control for all covariates, years 1981-1984

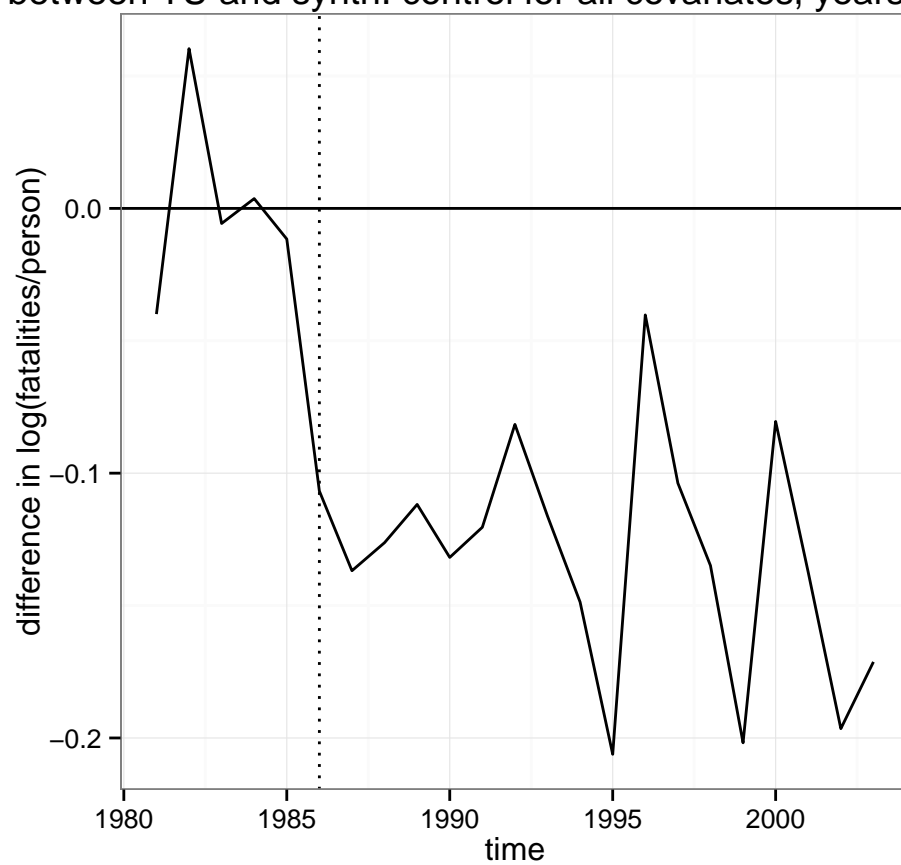


Figure 5: Gap between Treatment Unit and Synthetic Control developed using time periods 1981-1984 and all covariates.

Gap between TU and synth. control for VMT per capita cov

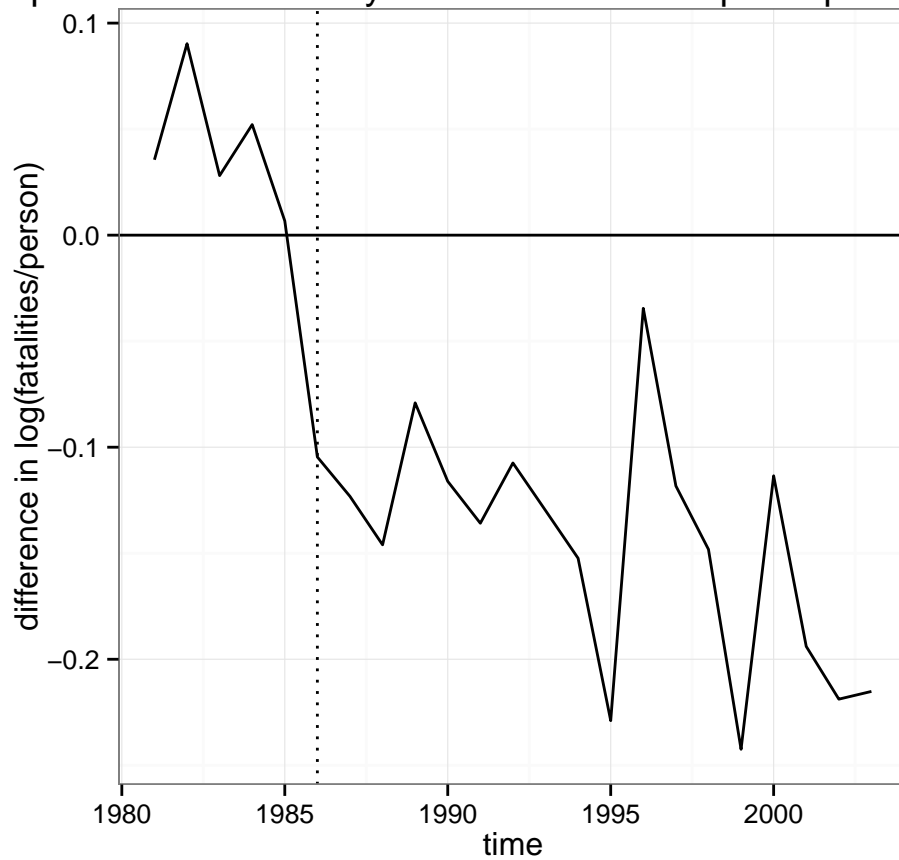


Figure 6: Gap between Treatment Unit and Synthetic Control developed using all previous time periods and VMT per capita as a covariate.

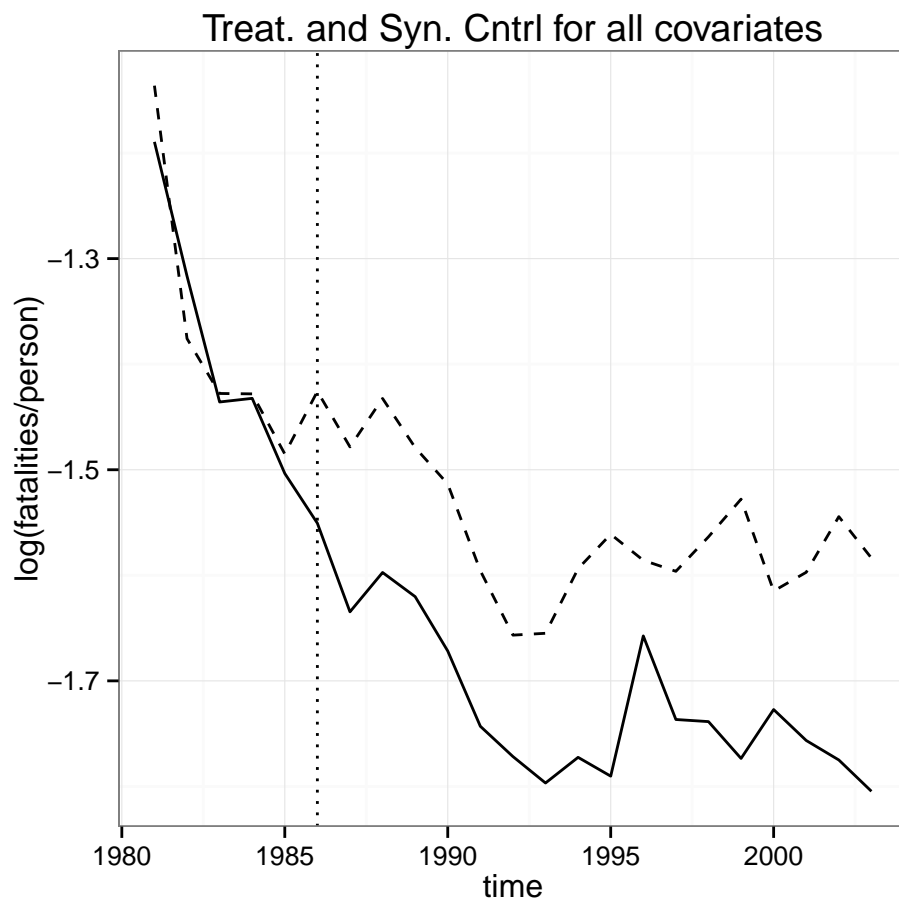


Figure 7: Treatment Unit and Synthetic Control log fatalities per capita developed using all previous time periods and all covariates.

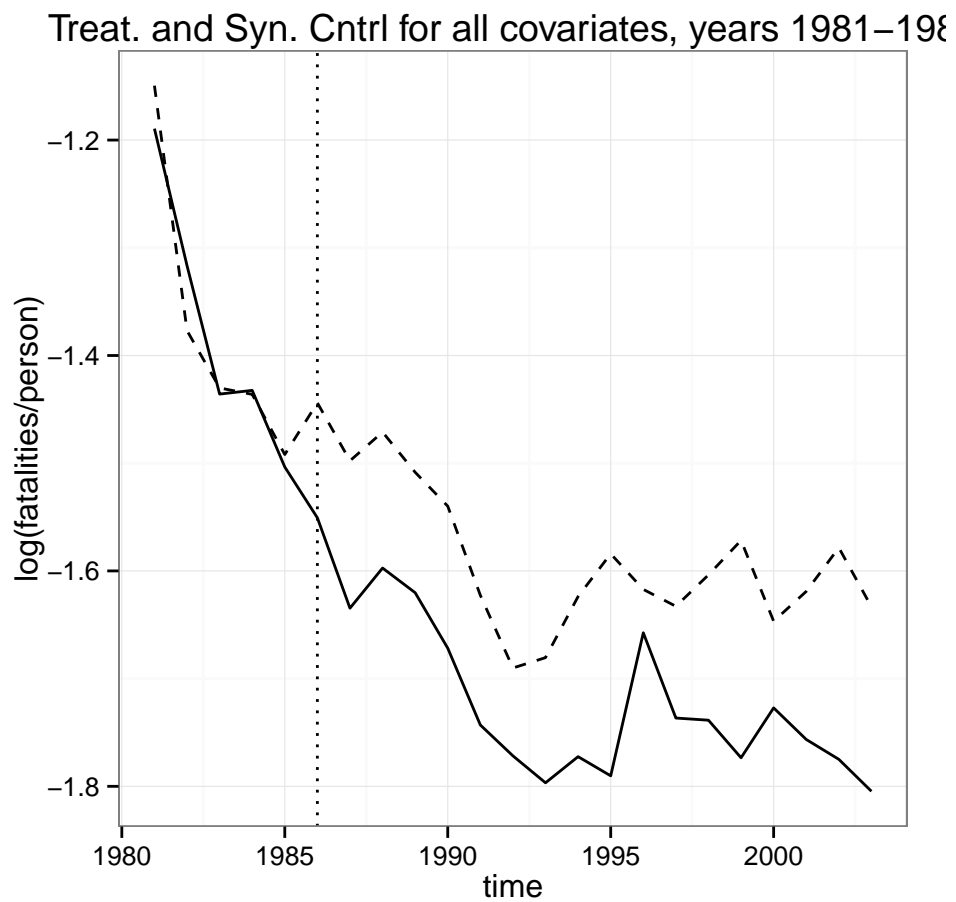


Figure 8: Treatment Unit and Synthetic Control log fatalities per capita developed using time periods 1981-1984 and all covariates.

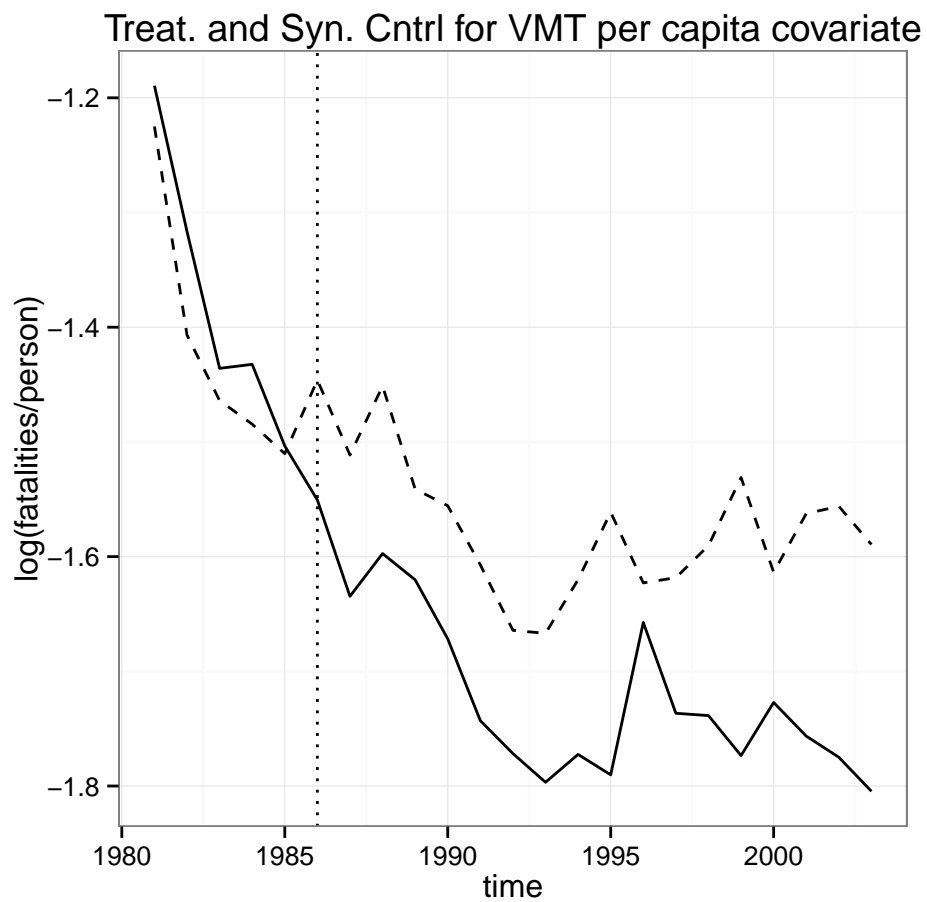


Figure 9: Treatment Unit and Synthetic Control log fatalities per capita developed using all previous time periods and VMT per capita as a covariate.



seatbelt laws.

Placebo test plot with  $-2$  control states excluded that have  
>50x divergence from TU pre-treatment MSPE  
(includes 31 control states)

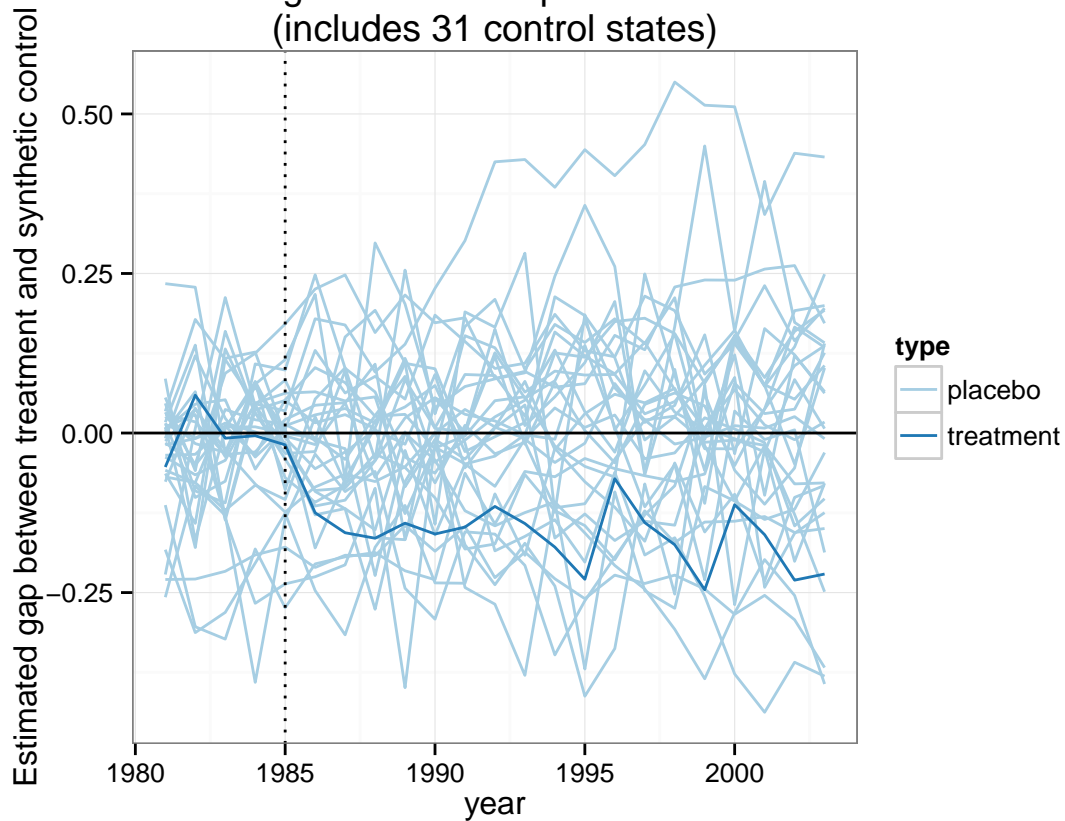


Figure 10: Placebo test results comparing treatment with placebo analysis on all control states. This plot does not exclude any states from the analysis.

### (iii) MSPE Ratios

... **Was it significant??** As shown in Figure 15, we find that the MSPE ratio for the TU is about 20, which is the second highest among the combined set of the TU and the control states. Florida (as a placebo) had a higher

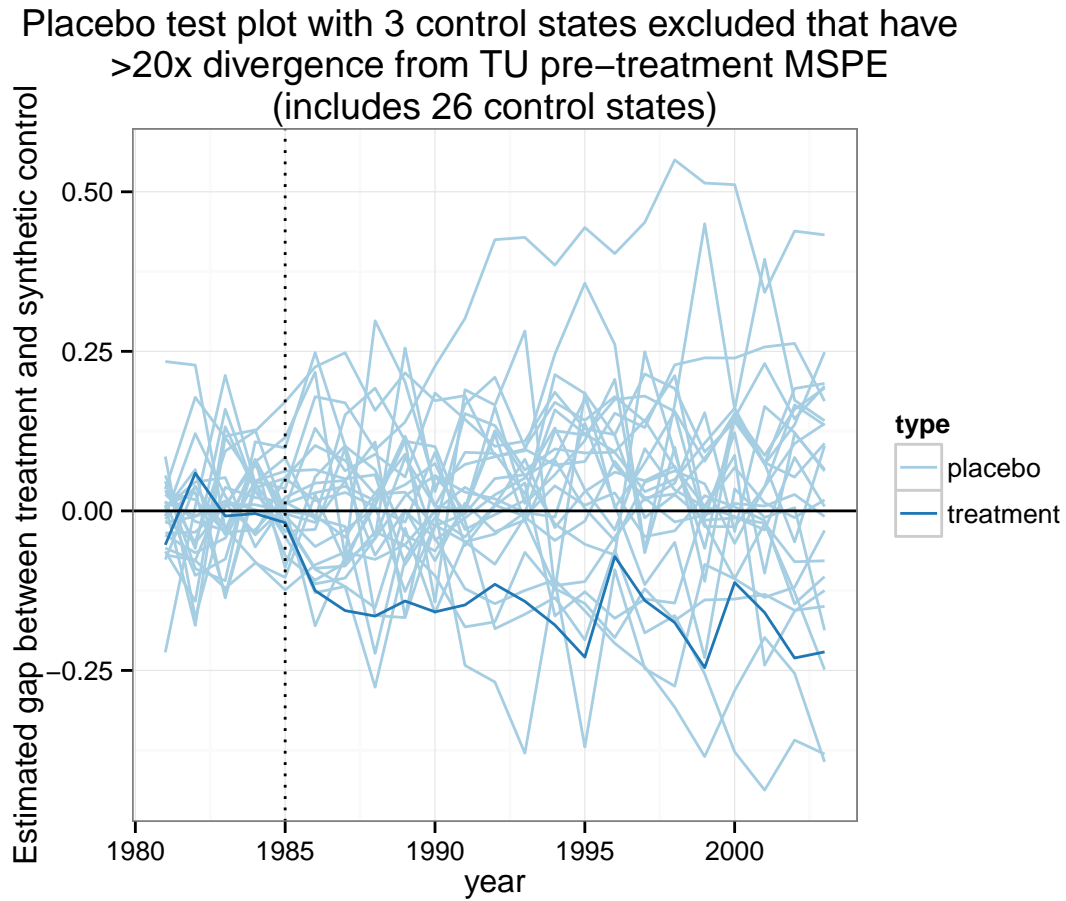


Figure 11: Placebo test results comparing treatment with placebo analysis on all control states. This plot excludes control states with MSPE in the pre-treatment period greater than 20x the MSPE for the treatment state.

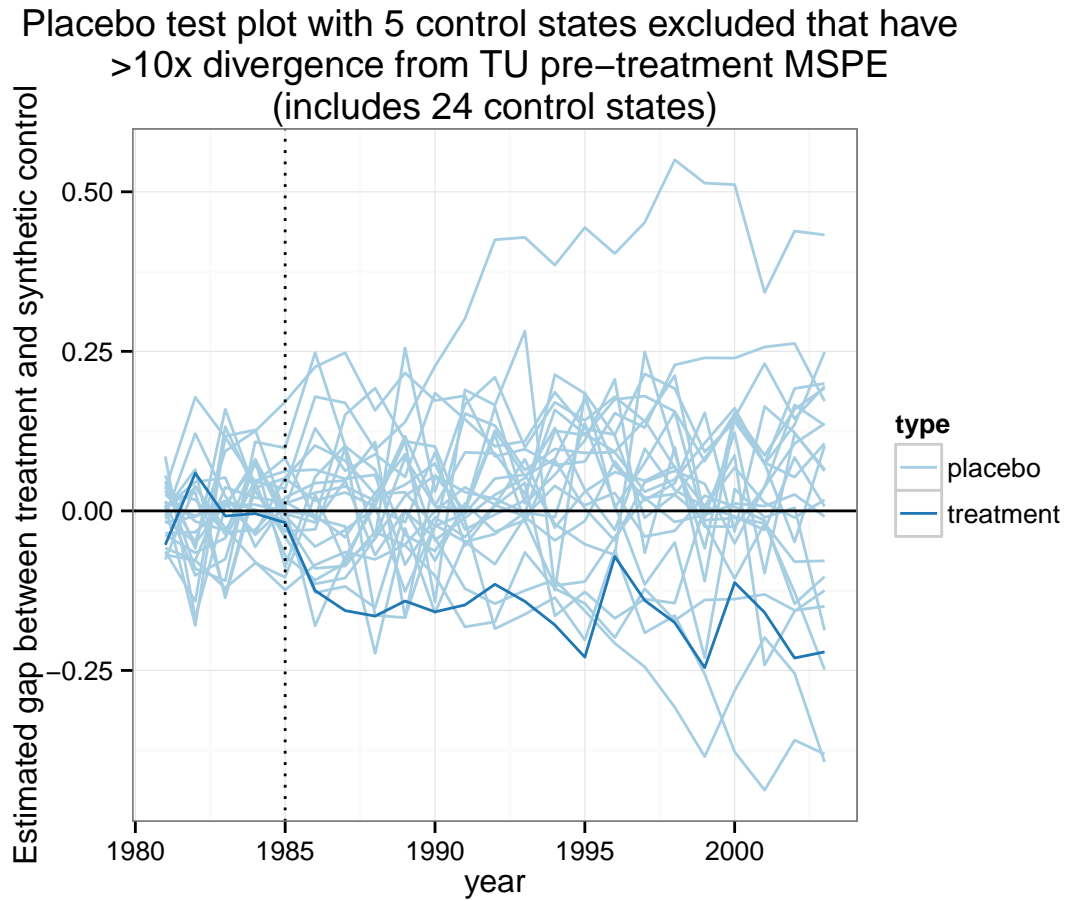


Figure 12: Placebo test results comparing treatment with placebo analysis on all control states. This plot excludes control states with MSPE in the pre-treatment period greater than 10x the MSPE for the treatment state

Placebo test plot with 13 control states excluded that have  
 $>5x$  divergence from TU pre-treatment MSPE  
 (includes 16 control states)

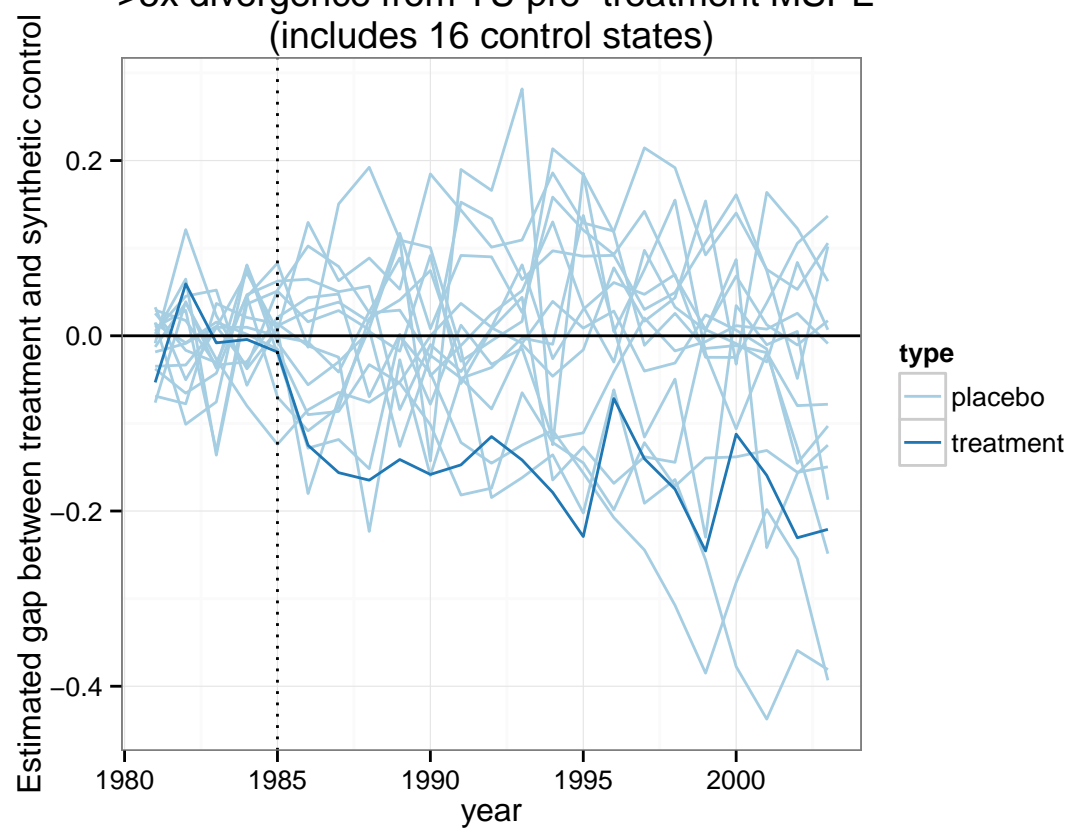


Figure 13: Placebo test results comparing treatment with placebo analysis on all control states. This plot excludes control states with MSPE in the pre-treatment period greater than 5x the MSPE for the treatment state

Placebo test plot with 20 control states excluded that have  
 $>2x$  divergence from TU pre-treatment MSPE  
 (includes 9 control states)

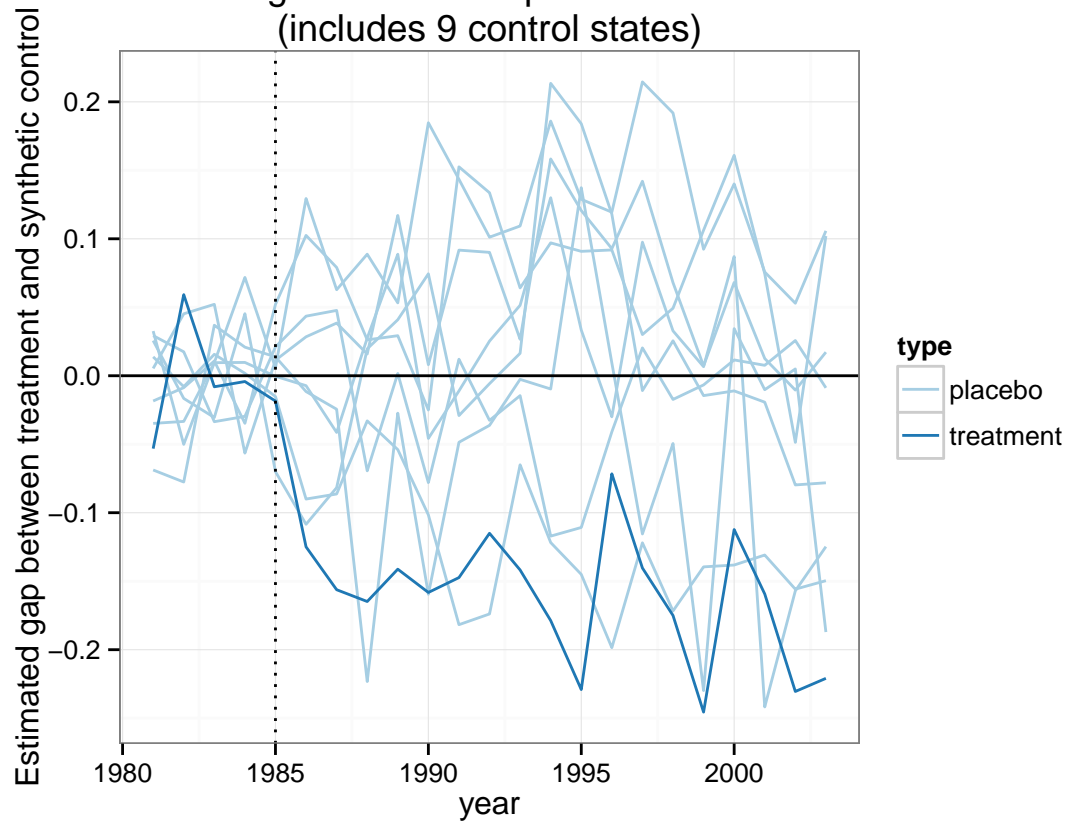


Figure 14: Placebo test results comparing treatment with placebo analysis on all control states. This plot excludes control states with MSPE in the pre-treatment period greater than  $2x$  the MSPE for the treatment state

MSPE ratio, about 120. This sows seeds of doubt but we still find that the TU has a higher apparent effect then 93% of units.

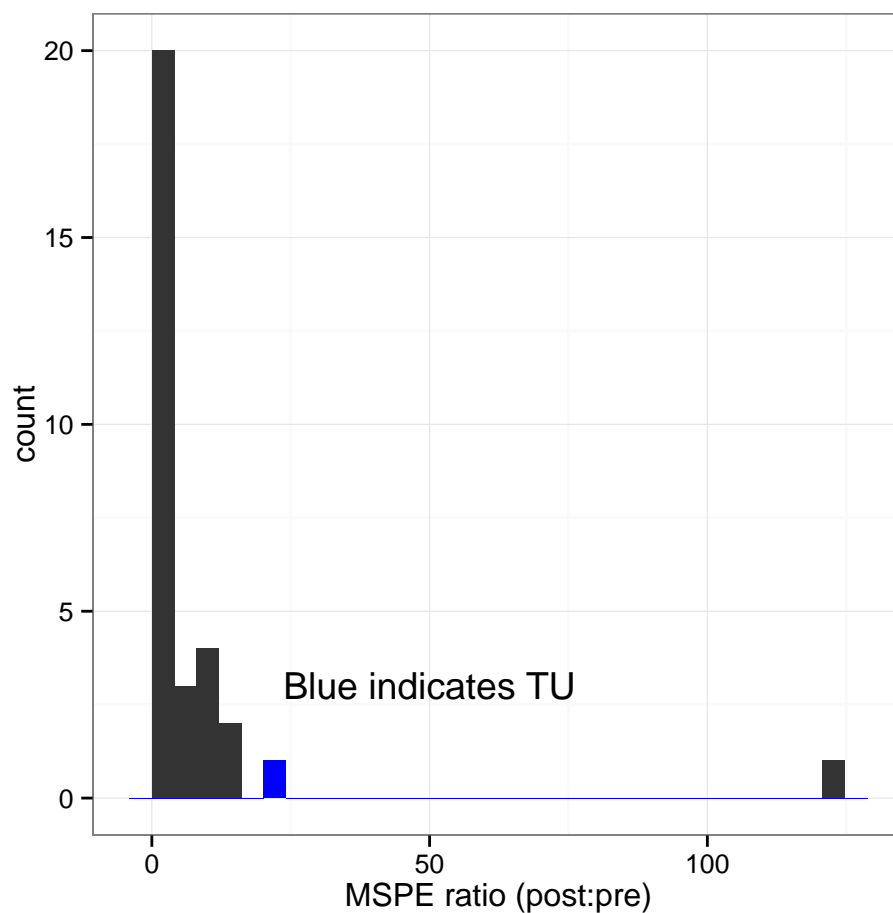


Figure 15: Ratio of post-treatment mean square percentage error (mspe) to pre-treatment mspe by state

## Part D: Compare with FE Model

Our central estimate for the impact of seatbelt laws using synthetic control methods is a 15% reduction in fatality rate. This is roughly double the

estimates we found using a fixed effects model (8% with significance at a 0.05 level) in the previous assignment. It is notable that many of the alternative estimates for the coefficient on primary seatbelt laws had higher (but not significant) results closer to those we find with the synthetic controls method.

The differences may stem from a reduced sample of treatment states in the synthetic control method. Since only four states with primary seatbelt laws were included (out of nearly 20 that eventually have the laws) it could be the case that the analysis is biased.

Another source of difference is in the structural way the estimates are made. Where the fixed effects model compares pre and post treatment, the synthetic controls method compares only in the post-treatment period. The fixed effects model includes first and second order corrections for year which should address some of these issues, but to the extent that the trend in national-level drivers for traffic fatalities does not obey the quadratic functional form there are potential errors that are not present in the synthetic control method, since synthetic controls does not impose a linear function in the same way on the covariates.

## Part E: Appendix: Code Listings

```
1 ## Frank's wd
2 ## setwd("/media/frank/Data/documents/school/berkeley/fall13/are213/are213/
  ps2")
3 ## Peter's wd
4 # setwd("~/Google Drive/ERG/Classes/ARE213/are213/ps2")
5
6 library(foreign) #this is to read in Stata data
7 library(Hmisc)
8 library(psych)
9 library(stargazer)
10 library(ggplot2) # for neato plotting tools
11 library(plyr) # for nice data tools like ddply
12 library(car) # "companion for applied regression" - recode fxn, etc.
13 library(gmodels) #for Crosstabs
14 library(plm) # for panel data
15 library(Synth) #for synthetic control
16 library(reshape)
17
18 source("../util/are213-func.R")
19
20 ps2a.data <- read.dta('traffic_safety2.dta')
21 ps2a.datakey <- data.frame(var.name=names(ps2a.data), var.labels = attr(ps2a
  .data, "var.labels"))
22 ps2a.data$logfatalpc <- with(ps2a.data, log(fatalities/population))
23 ps2a.data$sqyears <- with(ps2a.data, year^2)
24
```

```

25 # make state labels dataframe
26 state.labels <- attr(ps2a.data, "label.table")
27 state.labels$state.name <- attr(state.labels$state_number, "names")
28 state.labels <- as.data.frame(state.labels)
29 state.labels$state <- state.labels$state_number
30 state.labels$state.name <- as.character(state.labels$state.name)
31 state.labels$state_number <- NULL
32 state.labels <- rbind(state.labels, c("TU", 99))
33
34
35
36 tu.label <- data.frame(state_number = 99, state.name = "TU")
37
38
39 ## Problem 1 -----
40
41 ## Part a -----
42 ## Subpart i -----
43
44 pre.treatment.avg <- with(subset(ps2a.data, (state == 99 & primary == 0)),
45                           # Average pre-period log fatalities
46                           # per cap.
47                           mean(logfatalpc))
48 pre.control.avg <- with(subset(ps2a.data, (state != 99 & primary == 0)),
49                           mean(logfatalpc))
50
51
52 # Define state types
53 ps2a.data$type <- "Control"
54
55 for(i in 1:length(ps2a.data$year)){
56   if(ps2a.data$state[i] == 99){
57     ps2a.data$type[i] <- "Composite Treatment"}
58   if(ps2a.data$state[i] %in% c(4,10,30,41)){
59     ps2a.data$type[i] <- "Treatment State"}
60 }
61
62 use.rows <- which(ps2a.data$type != "Treatment State")
63
64 pre.period.data <- ddply(ps2a.data[use.rows,], .(year, type), summarize,
65                           logfatalpc = mean(logfatalpc))
66
67 # Plot of average fatality rates pre
68 # -1986.
69 preperiod.plot <- ggplot(data = subset(pre.period.data, year < 1986), aes(x
70   = year, y = logfatalpc, color = type))
71 preperiod.plot <- preperiod.plot +
72   geom_line() +
73   theme_bw() +
74   ylab("log(fatalities per capita)")
75
76 ggsave(filename = "img-p2b.logfatTrend.pdf", plot = preperiod.plot, width =
77   6, height = 4)
78
79 ## Subpart ii
80
81 # Compare logfatalpc in year before
82 # treatment

```



```

75 year.before <- ddply(subset(ps2a.data, year == 1985 & type != "Treatment
    State"), .(state), summarize,
76     type = type,
77     logfatalpc = logfatalpc)
78
79 target.fatal <- year.before$logfatalpc[which(year.before$type== "Composite
    Treatment")]
80 year.before$distance <- abs(year.before$logfatalpc - target.fatal)
81 year.before$distance[which(year.before$type== "Composite Treatment")] <- NA
82
83
84                                     # plot all states and TU
85 pdf("img-ps2b-compareStates.pdf", width = 4, height = 4)
86
87 ggplot(year.before, aes(logfatalpc)) +
88   geom_histogram(binwidth=0.1) +
89   geom_histogram(data=subset(year.before, state==1), aes(logfatalpc),
    binwidth=0.1, fill="red") +
90   theme_bw() +
91   xlab("log(fatalities per capita)") +
92   geom_vline(aes(xintercept=target.fatal), size = 1, color = "blue")
93
94 dev.off()
95
96
97 best.yb.match <- which(year.before$distance == min(year.before$distance, na.
    rm=T))
98
99 print(paste("The state number for the closest year before match is", year.
    before$state[best.yb.match]))
100
101                                     # The best match is Alabama.
102
103                                     # Tables comparing Alabama to the
    composite treatment group
104 stargazer(subset(ps2a.data, state==99),
105   out = "tab-ps2b-1a.tex",
106   title = "Composite Treatment Group Summary",
107   label = "tab:a21")
108
109 stargazer(subset(ps2a.data, state==1),
110   out = "tab-ps2b-1b.tex",
111   title = "Closest match for pre-policy fatalities: Alabama",
112   label = "tab:a22")
113
114
115 # graphical comparison between states
116
117 prep.gg.compare <- melt.data.frame(subset(ps2a.data, state==99 | state ==1),
    id.vars = c("state","year"))
118
119 prep.gg.compare$value <- as.numeric(prep.gg.compare$value)
120 prep.gg.compare$state <- as.factor(prep.gg.compare$state)
121
122 vars.we.care.about <- c("beer", "college", "primary", "secondary", "unemploy
    ", "logfatalpc", "rural_speed", "urban_speed", "precip")

```

```

123 rows.we.care.about <- which(prepare.compare$variable %in% vars.we.care.about
124 )
125 gg.compare.states <- ggplot(prepare.compare[rows.we.care.about,], aes(x=year
126 , y=value, color=state)) +
127   geom_point() +
128   facet_wrap("variable", scales = "free")
129 pdf("img-ps2b-compareStatesFacets.pdf", width=7, height = 6)
130 gg.compare.states
131 dev.off()
132
133 ## Part B - Synthetic control method
134
135
136                                     # Identify which states can be "
                                     control" (i.e., those that never
                                     have a primary seatbelt law)
137
138 state.policy <- ddply(ps2a.data, .(state), summarize, primary.max = max(
139   primary))
140 good.states <- which(state.policy$primary.max < 1)
141 control.states <- state.policy$state[good.states]
142 all.states <- c(control.states, 99)
143
144 ## subpart ii
145
146 # add to dataset
147 ps2a.data$vmt_percapita <- with(ps2a.data, totalvmt / population)
148 ps2a.data <- join(ps2a.data, state.labels)
149
150 #helper function: extract useful data from synth results
151 get.plot.data <- function(synth.res, dataprep.res, label="unlabeled", type =
152   "unknown"){
153   out <- data.frame(time=dataprep.res$tag$time.plot)
154
155   synthetic.trend <- dataprep.res$Y0plot %*% synth.res$solution.w
156   treatment.trend <- dataprep.res$Y1plot
157   gap <- treatment.trend - synthetic.trend
158
159   out$synth <- as.numeric(synthetic.trend)
160   out$treat <- as.numeric(treatment.trend)
161   out$gap <- as.numeric(gap)
162   out$spe <- out$gap^2
163   out$label <- label
164   out$type <- type
165
166   return(out)
167 }
168
169 #helper function: run dataprep, synth, and get.plot.data for a set of inputs
170
171 run.syn <- function(predictor.set, time.prior, treatment.state = std.
172   treatment.state, control.states, label = "unlabeled", type = "unknown"){

```

```

172 # synthetic controls dataprep with specified set of tractable predictors
173   dataprep.results <- dataprep(foo = ps2a.data,
174                                predictors = predictor.set,
175                                predictors.op = c("mean"),
176                                dependent = "logfatalpc",
177                                unit.variable = "state",
178                                time.variable = "year",
179                                treatment.identifier = treatment.state,
180                                controls.identifier = control.states,
181                                time.predictors.prior = time.prior,
182                                time.optimize.ssr = time.prior,
183                                time.plot = c(1981:2003),
184                                unit.names.variable = "state.name"
185                                )
186
187   synth.results <- synth(data.prep.obj = dataprep.results)
188
189   output <- get.plot.data(synth.results, dataprep.results, label, type)
190
191   return(output)
192 }
193
194 ## helper function: makes a gap plot so we can get them to look consistent
195 make.gap.plot <- function(syntheticresults, finame, desc){
196   gap.plot <- ggplot(data = syntheticresults, aes(y = gap, x = time))
197   gap.plot <- gap.plot +
198     geom_path() +
199     geom_vline(xintercept = 1986, linetype = 'dotted') +
200     theme_bw() + geom_hline(yintercept = 0) + ylab("difference
      in log(fatalities/person)") +
201     labs(title = paste("Gap between TU and synth. control
      for", desc))
202   ggsave(filename = paste0('img-gap-', finame, '.pdf'), plot = gap.plot,
203           width=5, height = 5)
204
205   split.plot <- ggplot(data = syntheticresults, aes(y = treat, x = time))
206   split.plot <- split.plot +
207     geom_path() +
208     geom_path(aes(y = synth), linetype = 'dashed') +
209     geom_vline(xintercept = 1986, linetype = 'dotted') +
210     theme_bw() + ylab('log(fatalities/person)') +
211     labs(title = paste("Treat. and Syn. Cntrl for",
212                        desc))
212   ggsave(filename = paste0('img-split-', finame, '.pdf'), plot = split.
213     plot, width=5, height = 5)
214
215   return(gap.plot)
216 }
217
218                                     # default entries to run.syn
219                                     function
220 std.treatment.state <- 99
221 full.predictor.set <- c("college", "precip", "snow32", "beer", "vmt_
      percapita", "unemploy")
222 full.time.prior <- c(1981:1985)

```

```

222 ## Part C: Pretty pictures
223 ## Part (i)
224 full.synthesis <- run.syn(full.predictor.set, full.time.prior, 99, control.
    states)
225 full.gap.plot <- make.gap.plot(full.synthesis, 'full', 'all covariates')
226
227 full.synthesis1984 <- run.syn(full.predictor.set, c(1981:1984), 99, control.
    states)
228 full.1984.gap.plot <- make.gap.plot(full.synthesis1984, 'full1984', 'all
    covariates, years 1981-1984')
229
230 vmt.synthesis <- run.syn("vmt_percapita", full.time.prior, 99, control.
    states)
231 vmt.gap.plot <- make.gap.plot(vmt.synthesis, 'vmt', 'VMT per capita
    covariate')
232
233 # Script to generate a placebo test plot
234 ## Part (ii)
235
236 placebo.test <- run.syn(full.predictor.set, full.time.prior, treatment.state
    = 99, control.states, label = 99, type = "treatment")
237
238 for(state in control.states){
239   updated.control <- control.states[which(control.states!=state)]
240   placebo.additional.result <- run.syn(full.predictor.set, full.time.prior
    , state, updated.control, label = state, type = "placebo")
241   placebo.test <- rbind(placebo.test, placebo.additional.result)
242   print(paste("Finished with state number", state, "and you should be
    patient for the rest to finish :"))
243 }
244
245 placebo.test$treated <- (placebo.test$time > 1985)
246
247 MPSEs <- ddply(placebo.test, .(label), summarize,
248   preMPSE = mean(spe[!treated]),
249   postMPSE = mean(spe[treated]))
250 MPSEs$ratio <- with(MPSEs, postMPSE/preMPSE)
251
252
253 # Exclusion threshold for placebo plot based on
254 preThreshold <- 50 #exclude states with pre-intervention MSPE greater than x
    times that for TU
255 tu.mspe <- MPSEs$preMPSE[which(MPSEs$label==99)]
256 include.states <- MPSEs$label[which(MPSEs$preMPSE < tu.mspe * preThreshold)]
257 include.states.rows <- which(placebo.test$label %in% include.states)
258
259 ex.number <- length(control.states) - length(include.states) - 1
260 in.number <- length(include.states)
261
262 # plot all the lines
263 placeboTest <- ggplot(placebo.test[include.states.rows,], aes(time, gap,
    group=label))
264 placeboTest <- placeboTest +
265   geom_line(aes(color=type))+
266   geom_vline(aes(xintercept = 1985), linetype="dotted") +
267   theme_bw() +
268   xlab("year") +

```

```

269 ylab("Estimated gap between treatment and synthetic control") +
270 scale_color_brewer(palette = "Paired") +
271 geom_hline(yintercept = 0) +
272 ggtitle(paste0("Placebo test plot with ",ex.number," control states
    excluded that have \n >", preThreshold, "x divergence from TU pre-
    treatment MSPE\n(includes ", in.number, " control states)"))
273 ggsave(paste0("img-placeboTest", preThreshold, ".pdf"), plot = placeboTest,
    width = 6, height = 5)
274
275
276 ## Part(iii)
277
278
279 MPSE.plot <- ggplot(MPSEs, aes(ratio))
280 MPSE.plot <- MPSE.plot +
281   geom_histogram() +
282   geom_histogram(data=subset(MPSEs, label==99), aes(ratio), fill="blue") +
283   theme_bw() +
284   annotate(geom="text", x=50, y=3, label="Blue indicates TU") +
285   xlab("MSPE ratio (post:pre)")
286
287 ggsave(filename = "img-mspeRatio.pdf", plot = MPSE.plot, width = 5, height =
    5)
288
289 # # Visual exploration of plots....only works if you have the native
    dataprep and synth objects. We would rather make our own plots :)
290 # gaps.plot(seatbelts.synth, syn.data.full)
291 # path.plot(seatbelts.synth, syn.data.full, Ylim = c(-1.3, -2))
292
293 # # Tables for synthetic controls results
294 # syn.table <- synth.tab(seatbelts.synth, syn.data.full,3)

```

./ps2b.r

```

1 # Econometrics helper functions for [R]
2 #
3 # Peter Alstone and Frank Proulx
4 # 2013
5 # version 1
6 # contact: peter.alstone AT gmail.com
7
8 # Category: Data Management -----
9
10
11 # Category: Data Analysis -----
12
13 # Function: Find adjusted R^2 for subset of data
14 # This requires a completed linear model...pull out the relevant y-values
    and residuals and feed them to function
15 # [TODO @Peter] Improve function so it can simply evaluate lm or glm object,
    add error handling, general clean up.
16 adjr2 <- function(y,resid){
17   r2 <- 1-sum(resid^2) / sum((y-mean(y))^2)
18   return(r2)
19 } #end adjr2
20
21

```

```

22 # Category: Plots and Graphics -----
23
24 ## Function for arranging ggplots. use png(); arrange(p1, p2, ncol=1); dev.
    off() to save.
25 require(grid)
26 vp.layout <- function(x, y) viewport(layout.pos.row=x, layout.pos.col=y)
27 arrange_ggplot2 <- function(..., nrow=NULL, ncol=NULL, as.table=FALSE) {
28   dots <- list(...)
29   n <- length(dots)
30   if(is.null(nrow) & is.null(ncol)) { nrow = floor(n/2) ; ncol = ceiling(n/
        nrow)}
31   if(is.null(nrow)) { nrow = ceiling(n/ncol)}
32   if(is.null(ncol)) { ncol = ceiling(n/nrow)}
33   ## NOTE see n2mfrow in grDevices for possible alternative
34   grid.newpage()
35   pushViewport(viewport(layout=grid.layout(nrow,ncol) ))
36   ii.p <- 1
37   for(ii.row in seq(1, nrow)){
38     ii.table.row <- ii.row
39     if(as.table) {ii.table.row <- nrow - ii.table.row + 1}
40     for(ii.col in seq(1, ncol)){
41       ii.table <- ii.p
42       if(ii.p > n) break
43       print(dots[[ii.table]], vp=vp.layout(ii.table.row, ii.col))
44       ii.p <- ii.p + 1
45     }
46   }
47 }
48
49 robust <- function(model){ #This calculates the Huber-White Robust standard
    errors -- code from http://thetarzan.wordpress.com/2011/05/28/heteroskedasticity-robust-and-clustered-standard-errors-in-r/
50   s <- summary(model)
51   X <- model.matrix(model)
52   u2 <- residuals(model)^2
53   XDX <- 0
54
55   for(i in 1:nrow(X)) {
56     XDX <- XDX +u2[i]*X[i,]%*%t(X[i,])
57   }
58
59 # inverse(X'X)
60   XX1 <- solve(t(X)%*%X)
61
62 #Compute variance/covariance matrix
63   varcovar <- XX1 %*% XDX %*% XX1
64
65 # Degrees of freedom adjustment
66   dfc <- sqrt(nrow(X))/sqrt(nrow(X)-ncol(X))
67
68   stdh <- dfc*sqrt(diag(varcovar))
69
70   t <- model$coefficients/stdh
71   p <- 2*pnorm(-abs(t))
72   results <- cbind(model$coefficients, stdh, t, p)
73   dimnames(results) <- dimnames(s$coefficients)
74   results

```

```

75 }
76
77 ## Two functions for clustered standard errors below from: http://people.su.se/~ma/clustering.pdf -----
78
79 clx <-
80   function(fm, dfcw, cluster){
81     # R-codes (www.r-project.org) for computing
82     # clustered-standard errors. Mahmood Arai, Jan 26, 2008.
83
84     # The arguments of the function are:
85     # fitted model, cluster1 and cluster2
86     # You need to install libraries 'sandwich' and 'lmtest'
87
88     # reweighting the var-cov matrix for the within model
89     library(sandwich);library(lmtest)
90     M <- length(unique(cluster))
91     N <- length(cluster)
92     K <- fm$rank
93     dfc <- (M/(M-1))*((N-1)/(N-K))
94     uj <- apply(estfun(fm),2, function(x) tapply(x, cluster, sum));
95     vcovCL <- dfc*sandwich(fm, meat=crossprod(uj)/N)*dfcw
96     coeftest(fm, vcovCL) }
97
98 mclx <-
99   function(fm, dfcw, cluster1, cluster2){
100     # R-codes (www.r-project.org) for computing multi-way
101     # clustered-standard errors. Mahmood Arai, Jan 26, 2008.
102     # See: Thompson (2006), Cameron, Gelbach and Miller (2006)
103     # and Petersen (2006).
104     # reweighting the var-cov matrix for the within model
105
106     # The arguments of the function are:
107     # fitted model, cluster1 and cluster2
108     # You need to install libraries 'sandwich' and 'lmtest'
109
110     library(sandwich);library(lmtest)
111     cluster12 = paste(cluster1,cluster2, sep=" ")
112     M1 <- length(unique(cluster1))
113     M2 <- length(unique(cluster2))
114     M12 <- length(unique(cluster12))
115     N <- length(cluster1)
116     K <- fm$rank
117     dfc1 <- (M1/(M1-1))*((N-1)/(N-K))
118     dfc2 <- (M2/(M2-1))*((N-1)/(N-K))
119     dfc12 <- (M12/(M12-1))*((N-1)/(N-K))
120     u1j <- apply(estfun(fm), 2, function(x) tapply(x, cluster1, sum))
121     u2j <- apply(estfun(fm), 2, function(x) tapply(x, cluster2, sum))
122     u12j <- apply(estfun(fm), 2, function(x) tapply(x, cluster12, sum))
123     vc1 <- dfc1*sandwich(fm, meat=crossprod(u1j)/N )
124     vc2 <- dfc2*sandwich(fm, meat=crossprod(u2j)/N )
125     vc12 <- dfc12*sandwich(fm, meat=crossprod(u12j)/N)
126     vcovMCL <- (vc1 + vc2 - vc12)*dfcw
127     coeftest(fm, vcovMCL)}
128
129 ## Function to compute ols standard errors , robust, clustered...
130 ## Based on http://diffuseprior.wordpress.com/2012/06/15/standard-robust-and

```

```

131   -clustered-standard-errors-computed-in-r/
132 ols.hetero <- function(form, data, robust=FALSE, cluster=NULL,digits=3){
133   r1 <- lm(form, data)
134   if(length(cluster)!=0){
135     data <- na.omit(data[,c(colnames(r1$model),cluster)])
136     r1 <- lm(form, data)
137   }
138   X <- model.matrix(r1)
139   n <- dim(X)[1]
140   k <- dim(X)[2]
141   if(robust==FALSE & length(cluster)==0){
142     se <- sqrt(diag(solve(crossprod(X)) * as.numeric(crossprod(resid(r1))/(n
143       -k))))
144     res <- cbind(coef(r1),se)
145   }
146   if(robust==TRUE){
147     u <- matrix(resid(r1))
148     meat1 <- t(X) %*% diag(diag(crossprod(t(u)))) %*% X
149     dfc <- n/(n-k)
150     se <- sqrt(dfc*diag(solve(crossprod(X)) %*% meat1 %*% solve(crossprod(X)
151       )))
152     res <- cbind(coef(r1),se)
153   }
154   if(length(cluster)!=0){
155     clus <- cbind(X,data[,cluster],resid(r1))
156     colnames(clus)[(dim(clus)[2]-1):dim(clus)[2]] <- c(cluster,"resid")
157     m <- dim(table(clus[,cluster]))
158     dfc <- (m/(m-1))*((n-1)/(n-k))
159     uclust <- apply(resid(r1)*X,2, function(x) tapply(x, clus[,cluster],
160       sum))
161     se <- sqrt(diag(solve(crossprod(X)) %*% (t(uclust) %*% uclust) %*% solve
162       (crossprod(X))*dfc))
163     res <- cbind(coef(r1),se)
164   }
165   res <- cbind(res,res[,1]/res[,2],(1-pnorm(abs(res[,1]/res[,2])))*2)
166   res1 <- matrix(as.numeric(sprintf(paste("%. ",paste(digits,"f",sep=""),sep=
167     " "),res)),nrow=dim(res)[1])
168   rownames(res1) <- rownames(res)
169   colnames(res1) <- c("Estimate","Std. Error","t value","Pr(>|t|)")
170   return(res1)
171 }

```

../util/are213-func.R