Lecture 3: Thresholds for general montone sets of graphs.

We are starting to get a picture of what the random graph G(n,p) looks like for different edge probabilities. Write $p \ll p^*$ as shorthand for $p/p^* \to 0$ and $p \gg p^*$ for $p/p^* \to \infty$. In Theorem 3.6 we saw that $p \ll \frac{1}{n}$ implies whp G(n,p) has no cycles and $p \gg \frac{1}{n}$ implies whp G(n,p) contains a cycle (or multiple cycles) as a subgraph. We are now interested in the threshold for containing as a subgraph. Note that \mathbf{x} contains a cycle itself, so containing a \mathbf{x} implies containing a cycle. Hence, the threshold function (if it exists) for containing \mathbf{x} we would expect it to be at least as big as $\frac{1}{n}$. It turns out to be $\frac{1}{n^{2/3}}$. We prove this in Theorem 3.7.

Threshold for containing \overline{\text{M}} as a subgraph

Theorem 3.7. The function $p^*(n) = \frac{1}{n^{2/3}}$ is a threshold for G(n,p) containing \mathbb{Z} as a subgraph.

Proof. Let p be such that $p/p^* \to 0$, i.e. $n^{2/3}p \to 0$ and let $G_n \sim G(n,p)$. Let $Y_n = Y_n(G_n)$ count the number of X in G_n . For each set S of 4 vertices from [n], let A_S be the event that G_n restricted to the vertices S is a X. We can write Y_n in terms of these indicator random variables.

$$Y_n = \sum_{S \in \binom{[n]}{4}} 1_{A_S}$$

As expectation is linear, the expected number of Σ in G_n is

$$\mathbb{E}(Y_n) = \sum_{S \in \binom{[n]}{4}} \mathbb{E}(1_{A_S}) \le n^4 p^6 = (n^{2/3} p)^6.$$

Hence $\mathbb{E}(Y_n) \to 0$ for $n^{2/3}p \to 0$. Observe $\mathbb{P}(G_n \text{ contains a } \mathbb{K}) = \mathbb{P}(Y_n > 0) \leq \mathbb{E}(Y_n)$ and so for $n^{2/3}p \to 0$ whp G_n does not contain \mathbb{K} as a subgraph.

Now it remains to show that for $p/p^* \to \infty$, i.e. for $n^{2/3}p \to \infty$ that whp $G_n \sim G(n,p)$ contains a \maltese . For this part of the proof we calculate the variance of Y_n by writing $Y_n = \sum_S 1_{A_S}$ and expanding. Write \sum_S for $\sum_{S \in \binom{[n]}{s}}$.

$$\mathbb{V}(Y_n) = \mathbb{E}(Y_n^2) - \mathbb{E}(Y_n)^2 = \mathbb{E}\left(\left(\sum_{S} 1_{A_S}\right)^2\right) - \left(\sum_{S} \mathbb{E}(1_{A_S})\right)^2.$$

We can rearrange a little to get an expression for the variance in terms of the probabilities of the events A_S and A_T

$$\mathbb{V}(Y_n) = \mathbb{E}\left(\sum_{S} 1_{A_S} \sum_{T \in [n]^3} 1_{A_T}\right) - \sum_{S} \mathbb{E}(1_{A_S}) \sum_{T} \mathbb{E}(1_{A_T})$$

$$= \sum_{S,T} \left(\mathbb{E}(1_{A_S} 1_{A_T}) - \mathbb{E}(1_{A_S}) \mathbb{E}(1_{A_T})\right)$$

$$= \sum_{S,T} \left(\mathbb{P}(A_S \& A_T) - \mathbb{P}(A_S) \mathbb{P}(A_T)\right). \tag{2}$$

If $S \cap T = \emptyset$, i.e. the vertex subsets S and T are disjoint then the events A_S and A_T are independent. Notice this is also true if S and T intersect in one vertex because they still share no edges in common. Hence if $|S \cap T| \leq 1$ then $\mathbb{P}(A_S \& A_T) = \mathbb{P}(A_S)\mathbb{P}(A_T)$ and these terms cancel in the expression for the variance (2) above.

So by this observation and (2),

$$\mathbb{V}(Y_n) \leq \sum_{|S \cap T| = \{2,3,4\}} \mathbb{P}(A_S \& A_T). \tag{3}$$

We now consider the three options: $|S \cap T| = 2, 3, 4$. For each of these, for $S, T \in \binom{[n]}{4}$ with the given intersection we want to calculate $\mathbb{P}(A_S \& A_T)$. For $|S \cap T| = 2$, one edge is shared. There are 10 other edges that need to be present in order to have Σ on both S and on T. Hence $\mathbb{P}(A_S \& A_T) = p^{11}$ for $|S \cap T| = 2$. Similarly, for $|S \cap T| = 3$, we get $\mathbb{P}(A_S \& A_T) = p^9$ and for $|S \cap T| = 4$, we get $\mathbb{P}(A_S \& A_T) = \mathbb{P}(A_S) = p^6$.

The aim is to find an upper bound for the right hand side of (3). Hence we want to know how many $S, T \in \binom{[n]}{4}$ for each of the possible overlaps. When S and T overlap on 2 vertices, the number of ways to pick them is to first pick the set of vertices in S then pick the two vertices in S that will ovelap with S, and lastly pick the last two vertices in S (the ones that don't overlap with S). This makes $\binom{n}{4}\binom{4}{2}\binom{n}{2}$. Actually all we need is that the number of $S, T \in \binom{[n]}{4}$ which overlap on two vertices is at most S0. Similarly the number that overlap on three vertices is at most S1.

We can now that calculate an explicit upper bound on our variance. From (2),

$$V(Y_n) \le n^6 p^1 1 + n^5 p^9 + n^4 p^6. \tag{4}$$

Now we have a good upper bound on the variance. What we actually want to show is that whp G_n contains a \mathbb{Z} . In other words we want to show whp $Y_n > 0$.

We use the following non-obvious idea. I have some b for which I know b > 0 and I want to use this to show that a > 0. Notice it is enough to show that |b - a| < b.

Let's go. By some re-arraning and Chebyshev,

$$\mathbb{P}(Y_n > 0) \ge \mathbb{P}\Big(|Y_n - \mathbb{E}(Y_n)| < \mathbb{E}(Y_n)/2\Big) = 1 - \mathbb{P}\Big(|Y_n - \mathbb{E}(Y_n)| \ge \mathbb{E}(Y_n)/2\Big) \ge 1 - \frac{4\mathbb{V}(Y_n)}{\mathbb{E}(Y_n)^2}.$$

The problem is now reduced to terms we have already calculated. By (4),

$$\mathbb{P}(Y_n > 0) \ge 1 - \frac{n^6 p^1 1 + n^5 p^9 + n^4 p^6}{\binom{n}{4} p^6}.$$
 (5)

For $n^{2/3}p \to \infty$ the fraction in (5) goes to zero. Hence for $n^{2/3}p \to \infty$ whp G(n,p) contains a \boxtimes as a subgraph.

All 'nice' sets of graphs have a threshold function

In the following theorem, we say a graph property \mathcal{A} is non-trivial if for each large enough n it is neither always true or always false. This is equivalent to saying that $\exists n_0$ such that the complete graphs on $n > n_0$ vertices has the property $(K_n \in \mathcal{A})$ and the empty graph on $n > n_0$ vertices does not $(\overline{K}_n, \notin \mathcal{A})$.

Theorem 3.8 (Bollobás-Thomason). Any monotone non-trivial graph property has a threshold function p^* .

Proof. (non-examinable)

Let $p_0 = p_0(n)$ be such that $\mathbb{P}(G(n, p_0) \in \mathcal{A}) = 1/2$. Note this must exist by the intermediate value theorem because for each integer n, $f_n(p) = \mathbb{P}(G(n, p_0) \in \mathcal{A})$ is a polynomial in p (of degree at most $\binom{n}{2}$) such that $f_n(0) = 0$ and $f_n(1) = 1$.

Let the threshold function be this p_0 , we set $p^*(n) = p_0(n)$. It will be sufficient to prove that $\forall \varepsilon > 0$, if $p/p_0 \to 0$ then $\mathbb{P}(G(n,p) \in A) < \varepsilon$ and if $p/p_0 \to \infty$ then $\mathbb{P}(G(n,p) \in A) > 1 - \varepsilon$.

Fix $\varepsilon > 0$. We first prove that $\exists p_b$ such that $\mathbb{P}(G(n, p_b) \in \mathcal{A}) > 1 - \varepsilon$.

Consider the union of k copies of $G(n, p_0)$, for some k which we will decide later. Let $H = ([n], E(G_1) \cup \ldots \cup E(G_k))$ where each $G_i \sim G(n, p_0)$. Here the graphs G_i are all defined on the same vertex set [n], and H is the random graph on this vertex set with edge set the union of the edge sets of the G_i . For any given $i \neq j \in [n]$ the probability ij is not in the edge set of H is exactly the probability that the edge ij does not appear in any of the G_i , which is $(1-p)^k$. This means H is the random graph where each edge is present independently with probability $1-(1-p)^k$. Thus $H \sim G(n, 1-(1-p)^k)$.

The next idea is to notice that \mathcal{A} monotone means that $H \in \mathcal{A}$ if $\exists i$ such that $G_i \in \mathcal{A}$. Thus,

$$\mathbb{P}(H \in \mathcal{A}) \le 1 - \mathbb{P}(\forall i, G_i \notin \mathcal{A}) = 1 - \mathbb{P}(G_1 \notin \mathcal{A})^k = 1 - \frac{1}{2^k}.$$
 (6)

Recall that we are still free to choose k, set $k = \lceil \log_2(1/\varepsilon) \rceil$ and let $p_1 = 1 - (1-p)^k$. Now by our choices of p_1 and k the equation on line (6) says precisely that $\mathbb{P}(G(n, p_1) \in \mathcal{A}) > 1 - \varepsilon$.

Claim 3.9. For all $\varepsilon > 0$ there exists function $p_{\ell}(n) = p_{\ell}(n, \varepsilon)$ such that $\mathbb{P}(G(n, p) \in \mathcal{A}) < \varepsilon$.

$$Proof.$$
 (exercise)

Claim 3.10. For all $\varepsilon > 0$, there exists $N = N(\varepsilon)$ such and if $p/p_0 \to 0$ then (for all n > N) $p(n) < p_{\ell}(n, \varepsilon)$ and if $p/p_0 \to \infty$ then (for all n > N) $p(n) > p_b(n, \varepsilon)$.

$$Proof.$$
 (exercise)

With the help of these claims we can now finish the proof. Suppose $p/p_0 \to 1$ (we want to show $P(G(n,p) \in \mathcal{A}) > 1 - \varepsilon$). By Claim 3.10 $p/p_0 \to 1$ implies $p(n) > p_b(n,\varepsilon)$. Therefore by Theorem 3.4 $P(G(n,p) \in \mathcal{A}) \ge \mathbb{P}(G(n,p_b) \in \mathcal{A})$. But by our choice of $p_b(n) = p_b(n,\varepsilon)$ we know that $\mathbb{P}(G(n,p_b) \in \mathcal{A}) > 1 - \varepsilon$ and so $P(G(n,p) \in \mathcal{A}) > 1 - \varepsilon$ as required.

Now suppose $p/p_0 \to 0$ (we want to show $P(G(n,p) \in \mathcal{A}) < \varepsilon$). By Claim 3.10 $p/p_0 \to 1$ implies $p(n) < p_{\ell}(n,\varepsilon)$. The monotonicity result, Theorem 3.4, implies $P(G(n,p) \in \mathcal{A}) \leq \mathbb{P}(G(n,p_b) \in \mathcal{A})$. But by Claim 3.9, $P(G(n,p_b) \in \mathcal{A}) < \varepsilon$ and so we are done.