

Active Inference in continuous time notes

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Summary notation

- $\mathbf{x} = \{x_i\}_{i=1}^D$ environmental variables of the D -dimensional space constituting latent or hidden states;
- $\mathbf{s} = \{s_i\}_{i=1}^S$ body sensors input;
- $\boldsymbol{\mu} = \{\mu_i\}_{i=1}^D$ inner brain state representing the hidden environmental variables \mathbf{x}
- $P(\mathbf{x}, \mathbf{s})$ *Joint density*;
- $P(\mathbf{x}|\mathbf{s})$ *Posterior*;
- $P(\mathbf{s}|\mathbf{x})$ *Likelihood*;
- $P(\mathbf{x})$ *Prior*;
- $P(\mathbf{s}) = \int P(\mathbf{s}|\mathbf{x})P(\mathbf{x})d\mathbf{x}$ *marginal likelihood*;
- $Q(\mathbf{x})$ *R-density*;
- $F \equiv \int Q(\mathbf{x}) \ln \frac{Q(\mathbf{x})}{P(\mathbf{x}, \mathbf{s})} d\mathbf{x}$ *Variational Free Energy*;
- $L(\boldsymbol{\mu}, s) \equiv -\ln P(\boldsymbol{\mu}, s)$ *Laplace-encoded energy*;
- ε *Prediction error*;
- Σ^{-1} *Precision*

References: Baltieri and C. Buckley 2019, C. L. Buckley et al. 2017, Bogacz 2017,

1 Free Energy Principle (FEP)

The goal of an agent is to determine the probability of the hidden states given some sensory inputs:

$$P(\mathbf{x}|\mathbf{s}) = \frac{P(\mathbf{x}, \mathbf{s})}{P(\mathbf{s})} = \frac{P(\mathbf{s}|\mathbf{x})P(\mathbf{x})}{P(\mathbf{s})} \quad (1)$$

with

- $P(\mathbf{x}, \mathbf{s})$ *joint density*, beliefs about the states assumed to be encoded by the agent;
- $P(\mathbf{x}|\mathbf{s})$ *Posterior*, i.e. probability of hidden causes x given observed sensory data;
- $P(\mathbf{s}|\mathbf{x})$ *Likelihood*, i.e. organism's assumptions about sensory input \mathbf{s} given the hidden causes \mathbf{x} ;
- $P(\mathbf{x})$ *Prior*, i.e. agent's beliefs about hidden causes **before** that \mathbf{s} are received;
- $P(\mathbf{s}) = \int P(\mathbf{s}|\mathbf{x})P(\mathbf{x})d\mathbf{x}$ *marginal likelihood*, i.e. normalization factor.

For the agent it's not necessary to compute the complete posterior distribution, it has only to find the hidden state -or at least a good approximation- that maximize the posterior, i.e. $\arg \max_{\mathbf{x}} P(\mathbf{x}|\mathbf{s})$. The problem with the exact Bayesian scheme, is that $P(\mathbf{s})$ is often impossible to calculate, and moreover $P(\mathbf{x}|\mathbf{s})$ may not take a standard shape and could not have a summary statistics.

A biologically plausible technique consist in using an auxiliary distribution $Q(\mathbf{x})$ called *recognition density* (*R-density*) that has to be optimized to become a good approximation of the posterior.

In order to do this the Kullback-Leibler divergence is minimized:

$$\begin{aligned} D_{KL}(Q(\mathbf{x}) || P(\mathbf{x}|\mathbf{s})) &= \int Q(\mathbf{x}) \ln \frac{Q(\mathbf{x})}{P(\mathbf{x}|\mathbf{s})} d\mathbf{x} \\ &= \int Q(\mathbf{x}) \ln \frac{Q(\mathbf{x})P(\mathbf{s})}{P(\mathbf{x}, \mathbf{s})} d\mathbf{x} \\ &= \int Q(\mathbf{x}) \ln \frac{Q(\mathbf{x})}{P(\mathbf{x}, \mathbf{s})} d\mathbf{x} + \ln P(\mathbf{s}) \int Q(\mathbf{x}) d\mathbf{x} \\ &= F + \ln P(\mathbf{s}) \end{aligned} \tag{2}$$

where

- $F \equiv \int Q(\mathbf{x}) \ln \frac{Q(\mathbf{x})}{P(\mathbf{x}, \mathbf{s})} d\mathbf{x} = -\langle \ln P(\mathbf{x}, \mathbf{s}) \rangle_Q - \langle \ln Q(\mathbf{x}) \rangle_Q$ is the *Variational Free Energy* (VFE), a quantity that depends on the R-density and the knowledge about the environment i.e. the joint density $P(\mathbf{s}, \mathbf{x}) = P(\mathbf{s}|\mathbf{x})P(\mathbf{x})$ that we are assuming the agent has.
- $\ln P(\mathbf{s})$ is a term independent of the recognition density $Q(\mathbf{x})$ (\Rightarrow minimizing F with respect to $Q(\mathbf{x})$ will minimize the D_{KL})

2 Laplace approximation

Often optimizing F for arbitrary $Q(\mathbf{x})$ is particularly complex. Moreover, it is assumed that neural activity parametrise sufficient statistic. For these reasons, a common approximation is to assume that the R-density take a Gaussian form.

2.1 One dimensional case

Let us assume that the R-density $Q(x)$ has a peak at point μ . The Taylor-expansion of the logarithm around this peak is

$$\ln Q(x) \simeq \ln Q(\mu) - \frac{1}{2} \frac{(x - \mu)^2}{\Sigma} \tag{3}$$

with

$$\frac{1}{\Sigma} = - \frac{\partial^2}{\partial x^2} \ln Q(x) \Big|_{x=\mu} \tag{4}$$

Now it is possible to approximate the probability distribution $Q(x)$ with the distribution

$$\mathcal{N}(x; \mu, \Sigma) = \frac{1}{\sqrt{2\pi\Sigma}} e^{-\frac{(x-\mu)^2}{2\Sigma}} \tag{5}$$

i.e. a Gaussian distribution that has been normalized using the factor $Q(\mu)\sqrt{2\pi\Sigma}$.

Now the VFE can be written as follow

$$\begin{aligned} F &\approx \int \mathcal{N}(x; \mu, \Sigma) \left(-\frac{1}{2} \ln(2\pi\Sigma) - \frac{(x-\mu)^2}{2\Sigma} \right) dx - \int \mathcal{N}(x; \mu, \Sigma) \ln P(x, s) dx \\ &= -\frac{1}{2} \ln(2\pi\Sigma) - \frac{1}{2\Sigma} \int \mathcal{N}(x; \mu, \Sigma) (x-\mu)^2 dx - \int \mathcal{N}(x; \mu, \Sigma) \ln P(x, s) dx \\ &= -\frac{1}{2} \ln(2\pi\Sigma) - \frac{1}{2} - \int \mathcal{N}(x; \mu, \Sigma) \ln P(x, s) dx \end{aligned} \quad (6)$$

To end up with an analytic model of the FEP, further simplifications and assumptions are needed to evaluate the last term¹.

Let us first assume that the R-density is sharply peaked at its mean value and that $P(x, s)$ is a smooth function of x : under these assumptions is possible to consider the integrated function appreciably non-zero only near the peak, and is possible to use a second order Taylor expansion of the $L(x, s) \equiv -\ln P(x, s)$ around $x = \mu$.

$$L(x, s) \approx L(\mu, s) + \left[\frac{dL(x, s)}{dx} \right]_{x=\mu} (x - \mu) + \frac{1}{2} \left[\frac{d^2L(x, s)}{dx^2} \right]_{x=\mu} (x - \mu)^2 \quad (7)$$

\Downarrow

$$\begin{aligned} - \int \mathcal{N}(x; \mu, \Sigma) \ln P(x, s) dx &\approx \int \mathcal{N}(x; \mu, \Sigma) \left\{ L(\mu, s) + \left[\frac{dL(x, s)}{dx} \right]_{x=\mu} (x - \mu) + \right. \\ &\quad \left. + \frac{1}{2} \left[\frac{d^2L(x, s)}{dx^2} \right]_{x=\mu} (x - \mu)^2 \right\} dx \\ &= L(\mu, s) + \left[\frac{dL(x, s)}{dx} \right]_{x=\mu} \left(\int \mathcal{N}(x; \mu, \Sigma) x dx - \mu \right) + \\ &\quad + \frac{1}{2} \left[\frac{d^2L(x, s)}{dx^2} \right]_{x=\mu} \int \mathcal{N}(x; \mu, \Sigma) (x - \mu)^2 dx \\ &= L(\mu, s) + \frac{1}{2} \left[\frac{d^2L(x, s)}{dx^2} \right]_{x=\mu} \Sigma \end{aligned} \quad (8)$$

Now is possible to rewrite the variational free energy as

$$F(\mu, \Sigma, s) \approx L(\mu, s) + \frac{1}{2} \left(\left[\frac{d^2L(x, s)}{dx^2} \right]_{x=\mu} \Sigma - \ln(2\pi\Sigma) - 1 \right) \quad (9)$$

with $L(\mu, s)$ said *Laplace-encoded energy*, and the variational free energy written as a function and not anymore as a functional.

Since the goal is to minimize the Kullback-Leibler divergence trough the minimization of the VFE, is possible to simplify further removing the Σ dependency taking the derivative with respect this and imposing $\frac{dF}{d\Sigma} = 0$

$$\frac{dF}{d\Sigma} = \frac{1}{2} \left(\left[\frac{d^2L(x, s)}{dx^2} \right]_{x=\mu} - \frac{1}{\Sigma} \right) = 0 \quad (10)$$

¹At the end we expand the implications for the interpretation of brain functions due to this issue.

$$\Rightarrow \Sigma = \left[\frac{d^2 L(x, s)}{dx^2} \right]_{x=\mu}^{-1} \equiv \Sigma^* \quad (11)$$

The final form of the VFE is then

$$F \approx L(\mu, s) - \frac{1}{2} \ln(2\pi\Sigma^*) , \quad (12)$$

that can be further simplified if the L function has a second order polynomial form², that implies that the second-order derivative of L with respect to x results in a constant factor that is useless and can be ignored³, leading to

$$F \approx L(\mu, s) \quad (13)$$

2.2 Multivariate case

Generalizing for a density $Q(\mathbf{x})$ over a D -dimensional space \mathbf{x} with peak at $\boldsymbol{\mu}$, let us go through the same procedure:

$$\ln Q(\mathbf{x}) \simeq \ln Q(\boldsymbol{\mu}) - \frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \hat{\boldsymbol{\Sigma}}^{-1}(\mathbf{x} - \boldsymbol{\mu}) \quad (14)$$

with

$$\left[\hat{\boldsymbol{\Sigma}}^{-1} \right]_{i,j} = - \frac{\partial^2}{\partial x_i \partial x_j} \ln Q(\mathbf{x}) \Big|_{\mathbf{x}=\boldsymbol{\mu}} \quad (15)$$

Now let us approximate $Q(\mathbf{x})$ with the multivariate Gaussian distribution

$$\mathcal{N}(\mathbf{x}; \boldsymbol{\mu}, \hat{\boldsymbol{\Sigma}}) = \frac{1}{\sqrt{(2\pi)^D \det \hat{\boldsymbol{\Sigma}}}} e^{-\frac{1}{2}(\mathbf{x}-\boldsymbol{\mu})^T \hat{\boldsymbol{\Sigma}}^{-1}(\mathbf{x}-\boldsymbol{\mu})} \quad (16)$$

and rewrite the VFE as follow:

$$\begin{aligned} F &\approx \int \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}, \hat{\boldsymbol{\Sigma}}) \left[-\frac{1}{2} \ln \left((2\pi)^D \det \hat{\boldsymbol{\Sigma}} \right) - \frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \hat{\boldsymbol{\Sigma}}^{-1}(\mathbf{x} - \boldsymbol{\mu}) \right] d\mathbf{x} + \\ &\quad - \int \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}, \hat{\boldsymbol{\Sigma}}) \ln P(\mathbf{x}, s) d\mathbf{x} \\ &= -\frac{1}{2} \ln \left((2\pi)^D \det \hat{\boldsymbol{\Sigma}} \right) - \int \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}, \hat{\boldsymbol{\Sigma}}) \left[\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \hat{\boldsymbol{\Sigma}}^{-1}(\mathbf{x} - \boldsymbol{\mu}) \right] d\mathbf{x} + \\ &\quad - \int \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}, \hat{\boldsymbol{\Sigma}}) \ln P(\mathbf{x}, s) d\mathbf{x} \end{aligned} \quad (17)$$

Let us focus on the second term making as first thing the change of variables with unitary Jacobian, $\mathbf{y} = \mathbf{x} - \boldsymbol{\mu}$.

$$\int \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}, \hat{\boldsymbol{\Sigma}}) \left[\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \hat{\boldsymbol{\Sigma}}^{-1}(\mathbf{x} - \boldsymbol{\mu}) \right] d\mathbf{x} = \int \mathcal{N}(\mathbf{y}; \mathbf{0}, \hat{\boldsymbol{\Sigma}}) \left[\frac{1}{2}\mathbf{y}^T \hat{\boldsymbol{\Sigma}}^{-1}\mathbf{y} \right] d\mathbf{y} \quad (18)$$

²as we will see in Sec.(3)

³let us remind that the final goal is to minimize F with respect to x

After that, since $\hat{\Sigma}^{-1}$ is a symmetric and real matrix, the spectral theorem guarantees the existence of an orthonormal matrix \mathbf{U} such that $\mathbf{U}^T \hat{\Sigma}^{-1} \mathbf{U} = \mathbf{\Lambda}$, with $\mathbf{\Lambda}$ diagonal matrix containing the eigenvalues $\{\lambda_i\}_{i=1}^D$ of $\hat{\Sigma}^{-1}$ and \mathbf{U} containing as columns the eigenvectors of $\hat{\Sigma}^{-1}$, so adding the identity matrix $\mathbf{I} = \mathbf{U}\mathbf{U}^T$ we get

$$\mathbf{y}^T \hat{\Sigma}^{-1} \mathbf{y} = \mathbf{y}^T \mathbf{U} \mathbf{U}^T \hat{\Sigma}^{-1} \mathbf{U} \mathbf{U}^T \mathbf{y} = \mathbf{z}^T \mathbf{\Lambda} \mathbf{z} = \sum_{i=1}^D \lambda_i z_i^2, \quad (19)$$

where $\mathbf{z} = \mathbf{U}^T \mathbf{y}$ is the \mathbf{y} representation in the orthonormal basis given by the eigenvectors of Σ^{-1} . Therefore moving to the variable \mathbf{z} (the Jacobian of this change of variables is unitary too), we obtain

$$\frac{1}{\sqrt{(2\pi)^D \det \hat{\Sigma}}} \int \frac{1}{2} \left[\sum_{i=1}^D \lambda_i z_i^2 \right] e^{-\frac{1}{2} \sum_{i=1}^D \lambda_i z_i^2} d\mathbf{z} = \frac{1}{2} \sqrt{\frac{\prod_{i=1}^D \lambda_i}{(2\pi)^D}} \sum_{i=1}^D \int \lambda_i z_i^2 e^{-\frac{1}{2} \sum_{i=1}^D \lambda_i z_i^2} dz_i \quad (20)$$

indicating with $\mathbf{z}_{\neq i} \equiv \{z_j\}_{j \neq i}$

$$\begin{aligned} \frac{1}{2} \sqrt{\frac{\prod_{i=1}^D \lambda_i}{(2\pi)^D}} \sum_{i=1}^D \int \lambda_i z_i^2 e^{-\frac{1}{2} \sum_{i=1}^D \lambda_i z_i^2} dz_i \int e^{\frac{1}{2} \sum_{j \neq i} \lambda_j z_j^2} d\mathbf{z}_{\neq i} = \\ = \frac{1}{2} \sqrt{\frac{\prod_{i=1}^D \lambda_i}{(2\pi)^D}} \sum_{i=1}^D \lambda_i \sqrt{\frac{2\pi}{\lambda_i}} \frac{1}{\lambda_i} \sqrt{\frac{(2\pi)^{D-1}}{\prod_{j \neq i} \lambda_j}} = \frac{1}{2} \sqrt{\frac{\prod_{i=1}^D \lambda_i}{(2\pi)^D}} \sum_{i=1}^D \sqrt{\frac{(2\pi)^D}{\prod_{i=1}^D \lambda_i}} = \frac{D}{2} \end{aligned} \quad (21)$$

In conclusion then

$$\int \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}, \hat{\Sigma}) \left[\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T \hat{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right] d\mathbf{x} = \frac{D}{2} \quad (22)$$

$$\Rightarrow F \approx -\frac{1}{2} \ln \left((2\pi)^D \det \hat{\Sigma} \right) - \frac{D}{2} - \int \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}, \hat{\Sigma}) \ln P(\mathbf{x}, \mathbf{s}) d\mathbf{x} \quad (23)$$

Let us now evaluate the last term. As done before let us assume that the R-density is sharply peaked at its mean value (i.e. both variance of and covariances between variables are small) and that $P(\mathbf{x}, \mathbf{s})$ is a smooth function of \mathbf{x} : under these assumptions is possible to consider the integrated function appreciably non-zero only near the peak, and is possible to use a second order Taylor expansion of the $L(\mathbf{x}, \mathbf{s}) \equiv -\ln P(\mathbf{x}, \mathbf{s})$ around $\mathbf{x} = \boldsymbol{\mu}$.

$$L(\mathbf{x}, \mathbf{s}) \approx L(\boldsymbol{\mu}, \mathbf{s}) + (\mathbf{x} - \boldsymbol{\mu})^T [\nabla L(\mathbf{x}, \mathbf{s})]_{\mathbf{x}=\boldsymbol{\mu}} + \frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T [\nabla^2 L(\mathbf{x}, \mathbf{s})]_{\mathbf{x}=\boldsymbol{\mu}} (\mathbf{x} - \boldsymbol{\mu}) \quad (24)$$

\Downarrow

$$\begin{aligned} - \int \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}, \hat{\Sigma}) \ln P(\mathbf{x}, \mathbf{s}) d\mathbf{x} &\approx \int \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}, \hat{\Sigma}) \left\{ L(\boldsymbol{\mu}, \mathbf{s}) + (\mathbf{x} - \boldsymbol{\mu})^T [\nabla L(\mathbf{x}, \mathbf{s})]_{\mathbf{x}=\boldsymbol{\mu}} + \right. \\ &\quad \left. + \frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^T [\nabla^2 L(\mathbf{x}, \mathbf{s})]_{\mathbf{x}=\boldsymbol{\mu}} (\mathbf{x} - \boldsymbol{\mu}) \right\} d\mathbf{x} \\ &= L(\boldsymbol{\mu}, \mathbf{s}) + \left(\int \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}, \hat{\Sigma}) \mathbf{x}^T d\mathbf{x} - \boldsymbol{\mu}^T \right) [\nabla L(\mathbf{x}, \mathbf{s})]_{\mathbf{x}=\boldsymbol{\mu}} \\ &\quad + \frac{1}{2} \int \mathcal{N}(\mathbf{x}; \boldsymbol{\mu}, \hat{\Sigma}) (\mathbf{x} - \boldsymbol{\mu})^T [\nabla^2 L(\mathbf{x}, \mathbf{s})]_{\mathbf{x}=\boldsymbol{\mu}} (\mathbf{x} - \boldsymbol{\mu}) d\mathbf{x} \\ &\stackrel{?}{=} L(\boldsymbol{\mu}, \mathbf{s}) + \frac{D}{2} [\nabla^2 L(\mathbf{x}, \mathbf{s})]_{\mathbf{x}=\boldsymbol{\mu}} \hat{\Sigma} \end{aligned} \quad (25)$$

\vdots

$$F \approx L(\boldsymbol{\mu}, \mathbf{s}) \quad (26)$$

2.3 About Laplace Approximation

Carrying on a theory based on an approximation that leads to Eq.(26), means to assume that the brain represents only the most likely environmental causes of sensory data and nothing else about the distribution. However, as we're going to see, the uncertainties are encoded directly in the form of the joint density.

3 Building the generative model

Thanks to the Laplace approximation, we've been able to write the VFE in terms of the Laplace-encoded energy $L(\boldsymbol{\mu}, \mathbf{s})$, that in turn depends on the joint density $P(\mathbf{x}, \mathbf{s})$. In this function are encoded the brain beliefs about the environmental causes of the sensory input and the beliefs *a priori* about environmental states.

Therefore, what needs to be built is a *generative model*, that is a model in which is encoded how the brain believe the world works and where all the hypothesis about the agent's behaviour are formalized.

3.1 Static Model

Let us consider a simple case of an agent that believes in an environment with hidden state \mathbf{x} that stimulates a sensory channel \mathbf{s} . As we've seen in Sec.(2), the brain will represent the environment only through the inner state $\boldsymbol{\mu}$, and what remains to do is to explicit the mapping between brain states and sensory data that will allows to make explicit the joint density.

3.1.1 One dimensional case

Let us assume that the agent believes its sensory input are generated by

$$s = g(x) + \mathcal{N}(s; 0, \Sigma_s), \quad (27)$$

with g generic function that expresses the relation between states and sensory input, to which is summed a noise represented by the normal distribution with zero mean and variance Σ_s . This assumption means that we can write

$$P(s|x) = \mathcal{N}(s; g(x), \Sigma_s) = \frac{1}{\sqrt{2\pi\Sigma_s}} e^{-\frac{(s-g(x))^2}{2\Sigma_s}}. \quad (28)$$

Moreover let us also assume that the agent also has a prior knowledge regarding the environmental state given by $\bar{\mu}$ that is linked with the inner state through

$$x = \bar{\mu} + \mathcal{N}(x; 0, \Sigma_{\bar{\mu}}) \quad (29)$$

$$\Rightarrow P(x) = \mathcal{N}(x; \bar{\mu}, \Sigma_{\bar{\mu}}) = \frac{1}{\sqrt{2\pi\Sigma_{\bar{\mu}}}} e^{-\frac{(x-\bar{\mu})^2}{2\Sigma_{\bar{\mu}}}}. \quad (30)$$

Now that we have specified a likelihood and a prior, is possible to determine the joint density

$$P(x, s) = P(s|x)P(x) \quad (31)$$

and consequently the Laplace-encoded energy

$$\begin{aligned} L(\mu, s) &= -\ln P(s|\mu) - \ln P(\mu) \\ &= \frac{1}{2} \ln(2\pi\Sigma_s) + \frac{(s-\mu)^2}{2\Sigma_s} + \frac{1}{2} \ln(2\pi\Sigma_{\bar{\mu}}) + \frac{(\mu-\bar{\mu})^2}{2\Sigma_{\bar{\mu}}} \\ &= \frac{\varepsilon_s^2}{2\Sigma_s} + \frac{\varepsilon_{\mu}^2}{2\Sigma_{\bar{\mu}}} + \frac{1}{2} \ln(\Sigma_s\Sigma_{\bar{\mu}}) + \ln(2\pi), \end{aligned} \quad (32)$$

where the ε terms are said *prediction errors* and measure the discrepancy respectively between the actual sensory data s and the outcome of its prediction $g(\mu)$ and between μ itself and its prior expectation $\bar{\mu}$. Therefore the former ε_s describes sensory prediction errors, the latter ε_{μ} model prediction errors (i.e. how brain states deviate from their expectation) and each one is weighted with the the corresponding inverse of the variance Σ_s^{-1} and Σ_{μ}^{-1} (which are often said *precisions*). As said at the end of Sec.(2.1), since the L function has a quadratic form is possible to ignore all the terms apart from the following

$$F \approx \frac{\varepsilon_s^2}{2\Sigma_s} + \frac{\varepsilon_{\mu}^2}{2\Sigma_{\bar{\mu}}} + \frac{1}{2} \ln(\Sigma_s\Sigma_{\bar{\mu}}). \quad (33)$$

3.1.2 Multivariate case

Let's repeat the same procedure for the multivariate case, remembering that $\mathbf{x} = \{x_i\}_{i=1}^D$, $\mathbf{s} = \{s_i\}_{i=1}^S$ and $\boldsymbol{\mu} = \{\mu_i\}_{i=1}^D$.

Now the prior is built from the expectations

$$\mathbf{x} = \bar{\boldsymbol{\mu}} + \mathbf{z}_{\bar{\boldsymbol{\mu}}}, \quad (34)$$

where $\mathbf{z}_{\bar{\boldsymbol{\mu}}}$ is a D -dimensional vector describing correlated noise with zero mean and covariance matrix $\hat{\boldsymbol{\Sigma}}_{\bar{\boldsymbol{\mu}}}$ that allows us to write

$$P(\mathbf{x}) = \mathcal{N}(\mathbf{x}; \bar{\boldsymbol{\mu}}, \hat{\boldsymbol{\Sigma}}_{\bar{\boldsymbol{\mu}}}) = \frac{1}{\sqrt{(2\pi)^D \det \hat{\boldsymbol{\Sigma}}_{\bar{\boldsymbol{\mu}}}}} e^{-\frac{1}{2}(\mathbf{x}-\bar{\boldsymbol{\mu}})^T \hat{\boldsymbol{\Sigma}}_{\bar{\boldsymbol{\mu}}}^{-1}(\mathbf{x}-\bar{\boldsymbol{\mu}})}. \quad (35)$$

Similarly the sensory inputs are assumed to be generated by

$$\mathbf{s} = \mathbf{g}(\mathbf{x}) + \mathbf{z}_{\mathbf{s}}, \quad (36)$$

with again $\mathbf{z}_{\mathbf{s}}$ S -dimensional vector describing correlated noise with zero mean and covariance matrix $\hat{\boldsymbol{\Sigma}}_{\mathbf{s}}$ that brings to the likelihood

$$P(\mathbf{s}|\mathbf{x}) = \mathcal{N}(\mathbf{s}; \mathbf{g}(\mathbf{x}), \hat{\boldsymbol{\Sigma}}_{\mathbf{s}}) = \frac{1}{\sqrt{(2\pi)^S \det \hat{\boldsymbol{\Sigma}}_{\mathbf{s}}}} e^{-\frac{1}{2}(\mathbf{s}-\mathbf{g}(\mathbf{x}))^T \hat{\boldsymbol{\Sigma}}_{\mathbf{s}}^{-1}(\mathbf{s}-\mathbf{g}(\mathbf{x}))}. \quad (37)$$

As done before, now is possible to write the Laplace-encoded energy

$$\begin{aligned}
L(\boldsymbol{\mu}, \mathbf{s}) &= \frac{1}{2} \ln \left((2\pi)^S \det \hat{\boldsymbol{\Sigma}}_s \right) + \frac{1}{2} (\mathbf{s} - \mathbf{g}(\boldsymbol{\mu}))^T \hat{\boldsymbol{\Sigma}}_s^{-1} (\mathbf{s} - \mathbf{g}(\boldsymbol{\mu})) \\
&\quad + \frac{1}{2} \ln \left((2\pi)^D \det \hat{\boldsymbol{\Sigma}}_{\bar{\boldsymbol{\mu}}} \right) + \frac{1}{2} (\boldsymbol{\mu} - \bar{\boldsymbol{\mu}})^T \hat{\boldsymbol{\Sigma}}_{\bar{\boldsymbol{\mu}}}^{-1} (\boldsymbol{\mu} - \bar{\boldsymbol{\mu}}) \\
&= \frac{1}{2} \boldsymbol{\varepsilon}_s^T \hat{\boldsymbol{\Sigma}}_s^{-1} \boldsymbol{\varepsilon}_s + \frac{1}{2} \boldsymbol{\varepsilon}_{\bar{\boldsymbol{\mu}}}^T \hat{\boldsymbol{\Sigma}}_{\bar{\boldsymbol{\mu}}}^{-1} \boldsymbol{\varepsilon}_{\bar{\boldsymbol{\mu}}} \\
&\quad + \frac{1}{2} \ln \left(\det \hat{\boldsymbol{\Sigma}}_{\bar{\boldsymbol{\mu}}} \det \hat{\boldsymbol{\Sigma}}_s \right) + \frac{1}{2} \ln \left((2\pi)^D (2\pi)^S \right),
\end{aligned} \tag{38}$$

where again we have introduced the prediction errors $\boldsymbol{\varepsilon}_s \equiv \mathbf{s} - \mathbf{g}(\boldsymbol{\mu})$ and $\boldsymbol{\varepsilon}_{\bar{\boldsymbol{\mu}}} \equiv \boldsymbol{\mu} - \bar{\boldsymbol{\mu}}$.

As done in the one dimensional case, is now possible write the VFE approximated form getting rid of all constants and the term depending on $\hat{\boldsymbol{\Sigma}}^*$:

$$F \approx \frac{1}{2} \boldsymbol{\varepsilon}_s^T \hat{\boldsymbol{\Sigma}}_s^{-1} \boldsymbol{\varepsilon}_s + \frac{1}{2} \boldsymbol{\varepsilon}_{\bar{\boldsymbol{\mu}}}^T \hat{\boldsymbol{\Sigma}}_{\bar{\boldsymbol{\mu}}}^{-1} \boldsymbol{\varepsilon}_{\bar{\boldsymbol{\mu}}} + \frac{1}{2} \ln \left(\det \hat{\boldsymbol{\Sigma}}_{\bar{\boldsymbol{\mu}}} \det \hat{\boldsymbol{\Sigma}}_s \right). \tag{39}$$

Often, the VFE is further simplified making another assumption, that is the *mean field approximation*, which implies statistical independence between environmental variables and between sensory inputs⁴, leading to a likelihood and a prior factorised respectively

$$P(\mathbf{x}) = \prod_{j=1}^D \mathcal{N}(x_j; \bar{\mu}_j, \Sigma_{\bar{\boldsymbol{\mu}}}^{(j,j)}) \tag{40}$$

$$P(\mathbf{s}|\mathbf{x}) = \prod_{i=1}^S \mathcal{N}(s_i; g_i(\mathbf{x}), \Sigma_s^{(i,i)}) \tag{41}$$

and a VFE with form

$$\begin{aligned}
F &\approx \sum_{i=1}^S \left[\frac{(\varepsilon_s^{(i)})^2}{2\Sigma_s^{(i,i)}} + \frac{1}{2} \ln \Sigma_s^{(i,i)} \right] + \sum_{j=1}^D \left[\frac{(\varepsilon_{\bar{\boldsymbol{\mu}}}^{(j)})^2}{2\Sigma_{\bar{\boldsymbol{\mu}}}^{(j,j)}} + \frac{1}{2} \ln \Sigma_{\bar{\boldsymbol{\mu}}}^{(j,j)} \right] \\
&= \sum_{i=1}^S \left[\frac{(\varepsilon_s^{(i)})^2}{2\Sigma_s^{(i,i)}} \right] + \sum_{j=1}^D \left[\frac{(\varepsilon_{\bar{\boldsymbol{\mu}}}^{(j)})^2}{2\Sigma_{\bar{\boldsymbol{\mu}}}^{(j,j)}} \right] + \frac{1}{2} \ln \left(\det \hat{\boldsymbol{\Sigma}}_{\bar{\boldsymbol{\mu}}} \det \hat{\boldsymbol{\Sigma}}_s \right).
\end{aligned} \tag{42}$$

3.2 Dynamic Model

4 VFE minimisation

In Sec.(2) we have expressed, in an approximated form, the Variational Free Energy in terms of the Laplace-encoded Energy, which depends on joint density $P(\mathbf{x}, \mathbf{s})$, that we have seen how to build starting from the generative model in Sec.(3). Now we are going to see a possible implementation of a biologically plausible mechanism to minimise VFE.

In the Free Energy Principle framework, it is proposed that the innate dynamics of the neural activity evolves in such a way that it implement a gradient descent scheme on the VFE.

⁴ $\Rightarrow \forall i \neq j \Sigma_{\bar{\boldsymbol{\mu}}}^{(i,j)} = 0, \Sigma_s^{(i,j)} = 0$

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