Recent advances in approximating f-divergences between two Ising models

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This talk is based on two joint works with

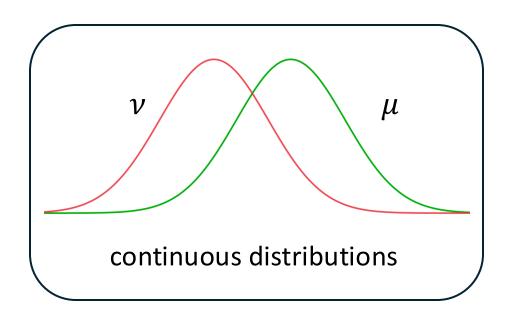
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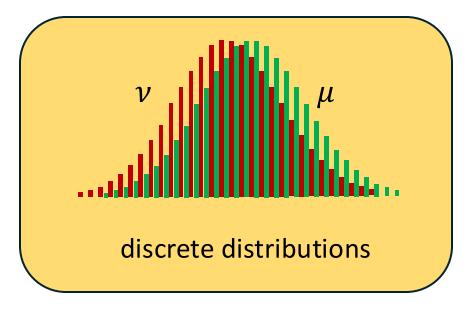
Seminar Talk
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Difference between two distributions

Input: two distributions ν and μ over state space Ω

Question: how to measure the difference between ν and μ ?

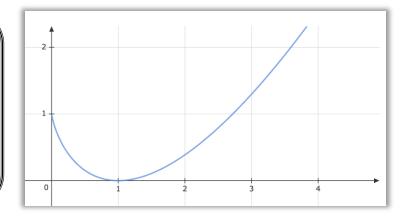




The f-divergence between two distributions

Let $f: \mathbb{R}_+ \to \mathbb{R}_{\geq 0}$ be a *convex* function s.t. f(1) = 0

$$f$$
-divergence: $D_f(\nu \parallel \mu) = \mathbb{E}_{X \sim \mu} \left[f\left(\frac{\nu(X)}{\mu(X)}\right) \right]$



- > χ^{α} divergence $f(x) = \frac{1}{2}|x 1|^{\alpha}$ for $\alpha \ge 1$ $\alpha = 1$ gives total variation (TV) distance $D_{TV}(v \parallel \mu) = \frac{1}{2} \sum_{x \in \Omega} |v(x) - \mu(x)|$
- $\alpha \text{ divergence } f(x) = \frac{x^{\alpha} \alpha x (1 \alpha)}{\alpha(\alpha 1)} \text{ for } \alpha \in \mathbb{R}$ $\alpha = 1 \text{ gives } \text{Kullback-Leibler (KL) } \text{ divergence } D_{KL}(\nu \parallel \mu) = \frac{1}{2} \sum_{x \in \Omega} \nu(x) \ln \frac{\nu(x)}{\mu(x)}$ $\alpha = 0 \text{ gives } \text{R\'enyi divergence } D_R(\nu \parallel \mu) = D_{KL}(\mu \parallel \nu)$
- > Squared Hellinger distance $f(x) = \frac{1}{2}(\sqrt{x} 1)^2$

Compute value the *f*-divergence

- Input: descriptions of two distributions ν , μ over Ω and a function f
- Output: the f-divergence $D_f(\nu \parallel \mu)$ between ν and μ for instance, TV distance: $D_{TV}(\nu \parallel \mu) = \frac{1}{2} \sum_{x \in \Omega} |\nu(x) \mu(x)|$

Trivial algorithm: enumerate all $x \in \Omega$ and add $\frac{1}{2}|\nu(x) - \mu(x)|$ together

Challenge: ν and μ have succinct descriptions (structured distribution)

- $|\Omega|$ can be **exponentially large** w.r.t. the size of input
- It can be *challenging* to evaluate the value of v(x) and $\mu(x)$

Examples: probabilistic graphical models, probabilistic circuits

Warm-up: Product distributions

Product distribution μ over $\{-, +\}^n$

$$\mu = \mu_1 \times \mu_2 \times \dots \times \mu_n$$

 μ_i is a distribution over $\{-, +\}$.

- μ can be described by n marginals
- Size the *input* 2n
 - Size of *sample space* $|\Omega| = 2^n$

Random sample
$$X = (X_1, X_2, ..., X_n) \sim \mu$$



 $X \in \{-, +\}^n$: n-dimensional random vector $X_i \in \{-, +\}$: independent sample from μ_i

Compute TV distance between product distributions

[Bhattacharyya, Gayen, Meel, Myrisiotis, Pavan, Vinodchandran, 2022]

- **Input**: distributions $\{\nu_i, \mu_i | 1 \le i \le n\}$ specifying ν and μ over $\{\pm\}^n$
- **Output**: the total variation distance between ν and μ

Results for computing TV distance between product distributions

Theorem [BGMMPV22]: the exact computing is **#P-complete**.

FPTAS (Full Poly-time Approximation Scheme)

A *deterministic* algorithm outputs a \hat{d} in time $poly(n, 1/\epsilon)$

$$(1 - \epsilon)D_{TV}(\nu \parallel \mu) \le \hat{d} \le (1 + \epsilon)D_{TV}(\nu \parallel \mu)$$

FPRAS (Full Poly-time Randomised Approximation Scheme)

A *randomized* algorithm outputs a random \hat{d} in time $\operatorname{poly}(n,1/\epsilon)$

$$\Pr[(1 - \epsilon)D_{TV}(\nu \| \mu) \le \hat{d} \le (1 + \epsilon)D_{TV}(\nu \| \mu)] \ge 2/3$$

Results for computing TV distance between product distributions

Theorem [BGMMPV22]: the exact computing is **#P-complete**.

Theorem [BGMMPV22] *FPTAS/FPRAS* exists *one of* the following condition holds

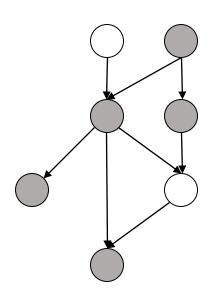
- μ has constant number of distinct marginals (e.g. uniform distribution over $\{-,+\}^n$)
- $\forall i \in [n], \nu_i(1) \geq \mu_i(1) \text{ and } \nu_i(1) \geq 1/2$ break symmetry lower bound

Theorem [FGJW23 and FLL23]

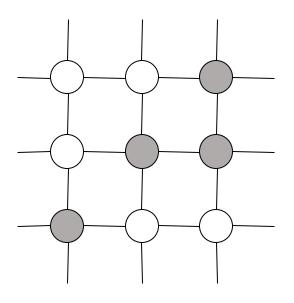
General product distributions ν , μ and error bound $0 < \epsilon < 1$

- FPTAS running time: $\tilde{O}\left(\frac{n^2}{\epsilon}\log\frac{1}{D_{TV}(\nu \parallel \mu)}\right)$
- FPRAS running time : $\tilde{O}\left(\frac{n^2}{\epsilon^2}\right)$

Beyond product distribution: graphical models



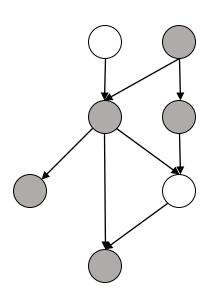
Bayesian network



Ising models

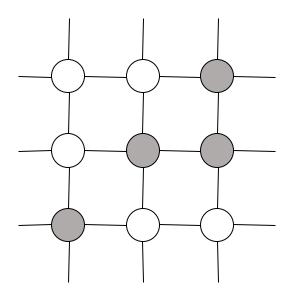
- Vertices are random variables and edges model local interactions
- Graphical models define joint distributions with complex correlations

Beyond product distribution: graphical models



Bayesian network

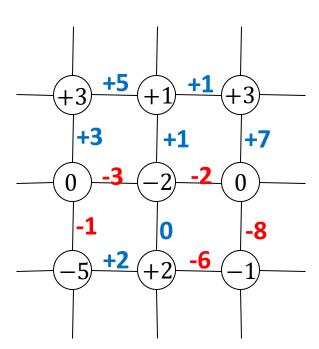
FPRAS for two *Bayesian networks*with bounded treewidth
[Bhattacharyya, Gayen, Meel, Myrisiotis,
Pavan, Vinodchandran 2025]



Ising models



Graph G = (V, E), weighted adjacent matrix $J \in \mathbb{R}^{V \times V}$, and external fields $h \in \mathbb{R}^{V}$

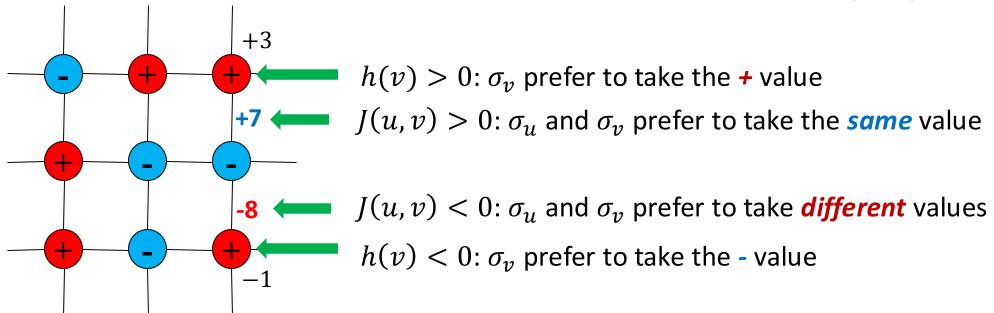


Graph G = (V, E), weighted adjacent matrix $I \in \mathbb{R}^{V \times V}$, and external fields $h \in \mathbb{R}^{V}$

∀ configurtation $\sigma \in \{-,+\}^V$

Weight
$$w(\sigma) = \exp\left(\frac{\sigma^T J \sigma}{2}\right) = \exp\left(\sum_{\{u,v\}\in E} \sigma_u \sigma_v J_{uv} + \sum_{v\in V} \sigma_v h_v\right)$$

Probability
$$\mu(\sigma) = \frac{w(\sigma)}{Z}$$
 Partition $Z = \sum_{x \in \{-1, +1\}^V} w(x)$



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Empty Graph J = 0

$$J=0$$



Product Distribution

Graph G = (V, E), weighted adjacent matrix $J \in \mathbb{R}^{V \times V}$, and external fields $h \in \mathbb{R}^{V}$

$$\forall \ \text{configurtation} \\ \sigma \in \{-, +\}^V \\ \hline Probability \ \mu(\sigma) = \exp\left(\frac{\sigma^T J \sigma}{2}\right) = \exp\left(\sum_{\{u,v\} \in E} \sigma_u \sigma_v J_{uv} + \sum_{v \in V} \sigma_v h_v\right) \\ \hline Probability \ \mu(\sigma) = \frac{w(\sigma)}{Z} \\ \hline Partition \\ Function \\ \hline Z = \sum_{x \in \{-1, +1\}^V} w(x)$$

Simplified Ising model (G, β)

- All edges have a *unified value* in interaction matrix $J(u,v)=\beta$ for all $\{u,v\}\in E$
- All vertices have **zero** external field h(v) = 0 for all $v \in V$

$$\mu(\sigma) \propto \prod_{\{u,v\} \in E} \exp(\sigma_u \sigma_v \beta) \propto \prod_{\{u,v\} \in E: \sigma_u = \sigma_v} \exp(2\beta) = \exp(2\beta \cdot \#\text{monochromatic edges})$$

Approximating the χ^{α} -divergence between two Ising models

- Input: two Ising models (G, J^{ν}, h^{ν}) and (G, J^{μ}, h^{μ}) defining ν and μ integer parameter α and error bound $\epsilon > 0$
- Output: $D \in (1 + \varepsilon) \cdot D_{\chi^{\alpha}} (\nu \parallel \mu)$ for χ^{α} -divergence $D_{\chi^{\alpha}}(\nu \parallel \mu)$

$$D_{\chi^{\alpha}}(\nu \parallel \mu) = \frac{1}{2} \sum_{x \in \Omega} \mu(x) \cdot \left| 1 - \frac{\nu(x)}{\mu(x)} \right|^{\alpha}$$



Sampling: draw **random sample** $X \sim \mu$ from the law of $\mu = \text{Ising}(G, J, h)$

Approximate Counting: compute an estimate \hat{Z} the **partition function** of Ising(G, J, h)

$$(1 - \varepsilon)Z \le \hat{Z} \le (1 + \varepsilon)Z$$

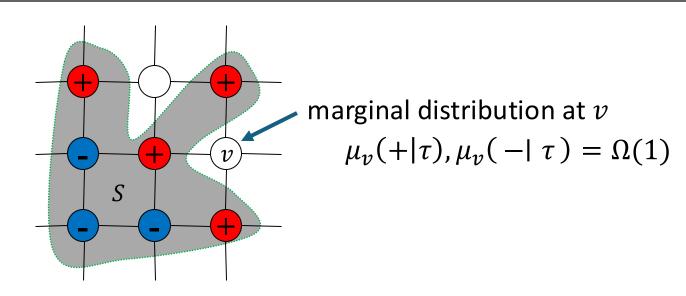
Our result: total variation distance ($\alpha = 1$)

Definition Marginal lower bound for Ising model

For any subset $S \subseteq V$, any vertex $v \in V \setminus S$, any pinning $\tau \in \{-1, +1\}^S$,

$$\forall c \in \{-1, +1\}, \qquad \mu_v(c \mid \tau) = \Omega(1)$$

Under any conditional, the marginal distribution on one vertex cannot be too biased



The assumption also appeared in *learning* [Bresler15], sampling and counting [CLV21]

Our result: total variation distance ($\alpha = 1$)

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Under any conditional, the marginal distribution on one vertex cannot be too biased

Theorem: total variation distance ($\alpha = 1$) [F, Liu, Yang, 2025]

Two Ising models $\nu = \text{Ising}(G, J^{\nu}, h^{\nu})$ and $\mu = \text{Ising}(G, J^{\mu}, h^{\mu})$ with marginal lower bound

Two models both admit $poly(n/\epsilon)$ -time algos for

- sampling
- approximate counting



poly (n/ϵ) -time algorithms for approximate $D_{TV}(\nu \parallel \mu)$

- Our result also work for *other graphical models*
- The marginal lower bound can be removed in some graphical models

Simplified Ising model (G, β)

- All edges have a *unified value* in interaction matrix $J(u,v)=\beta$ for all $\{u,v\}\in E$
- All vertices have **zero** external field h(v) = 0 for all $v \in V$

$$\mu(\sigma) \propto \prod_{\{u,v\} \in E} \exp(\sigma_u \sigma_v \beta) \propto \prod_{\{u,v\} \in E: \sigma_u = \sigma_v} \exp(2\beta)$$

Computational Phase Transition for Sampling and Approx. Counting

Max degree of graph Δ . Uniqueness threshold $\beta_c = \beta_c(\Delta) < 0$ s.t. $\exp(2\beta_c) = \frac{\Delta - 2}{\Delta}$

- *Polynomial time* sampling and approx. counting if $\beta \ge \beta_c$ [JS93, CCYZ25]
- Sampling and approx. counting is *hard* (unless NP=RP) if $\beta < \beta_c$ [SS14, GŠV16]

$$\beta_c = \frac{1}{2} \ln \left(\frac{\Delta - 2}{\Delta} \right)$$



Two Ising models $\nu = \text{Ising}(G, \beta_{\nu})$ and $\mu = \text{Ising}(G, \beta_{\mu})$

Two models **both above the threshold**

$$\min\{\beta_{\nu},\beta_{\mu}\} \geq \beta_{c}(\Delta)$$



FPRAS for the TV distance $D_{TV}(\nu \parallel \mu)$

Hardness result [Bhattacharyya, Gayen, Meel, Myrisiotis, Pavan, Vinodchandran, 2025]

Two Ising models $\nu = \text{Ising}(G, \beta_{\nu})$ and $\mu = \text{Ising}(G, \beta_{\mu})$

Two models **both below the threshold** $\max\{\beta_{\nu}, \beta_{\mu}\} < \beta_{c}(\Delta)$



No FPRAS for TV-distance unless NP=RP

Our result: χ^{α} -divergence

- Input: an integer $\alpha \geq 1$, two Ising models (G, J^{ν}, h^{ν}) and (G, J^{μ}, h^{μ}) , an error bound $\epsilon > 0$
- Output: $D \in (1 \pm \epsilon)D_{\chi^{\alpha}}(\nu \parallel \mu)$ for χ^{α} -divergence

$$D_{\chi^{\alpha}}(\nu \parallel \mu) = \sum_{x \in \{\pm\}^{V}} \mu(x) \left| 1 - \frac{\nu(x)}{\mu(x)} \right|^{\alpha}$$

Our result: χ^{α} -divergence

Theorem: approximation algorithm [F and Fu, 2025]

Two Ising models $\nu = \text{Ising}(G, J^{\nu}, h^{\nu})$ and $\mu = \text{Ising}(G, J^{\mu}, h^{\mu})$ with marginal lower bound

A *family* of Ising models
$$\mathcal{F} = \{(G, J^{(k)}, h^{(k)}) \mid \text{integer } 0 \leq k \leq \alpha\}$$
, where

$$J^{(k)} = kJ^{\nu} - (k-1)J^{\mu}$$

$$h^{(k)} = kh^{\nu} - (k-1)h^{\mu}$$

All Ising models in $\mathcal F$ admit $\operatorname{poly}(n/\epsilon)$ -time algos for

- sampling
- approximate counting



 $\operatorname{poly}(n/\epsilon)$ -time algorithms for approximate $D_{\chi^{\alpha}}(\nu \parallel \mu)$

$$\chi^{\alpha}$$
-divergence with $\alpha=1$



$$D_{TV} = D_{\chi}^{\alpha}$$

 ${\cal F}$ only contains **two** input Ising ν and μ



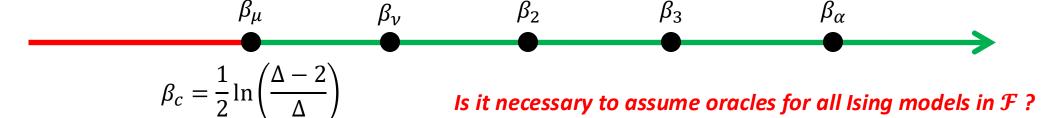
Recover TVdistance result

Corollary: simplified Ising model [F, Fu, 2025]

Two zero field Ising models (G, β_{ν}) and (G, β_{μ}) with unified non-zero values in interaction matrices

$$\mathcal{F} = \{(G, \beta_k) \mid \text{integer } 0 \le k \le \alpha\}, \text{ with } \beta_k = \beta_\mu + k(\beta_\nu - \beta_\mu) \text{ (note } \beta_0 = \beta_\mu \text{ and } \beta_1 = \beta_\nu)$$

Case $\beta_{\nu} \geq \beta_{\mu}$: poly-time algorithm for χ^{α} -divergence exist if $\beta_{\mu} \geq \beta_{c}$



Case $\beta_{\mu} > \beta_{\nu}$: poly-time algorithm for χ^{α} -divergence exist if $\beta_{\alpha} = \beta_{\mu} + k(\beta_{\nu} - \beta_{\mu}) \geq \beta_{c}$

$$\beta_{\alpha} \qquad \beta_{4} \qquad \beta_{3} \qquad \beta_{\nu} \qquad \beta_{\mu}$$

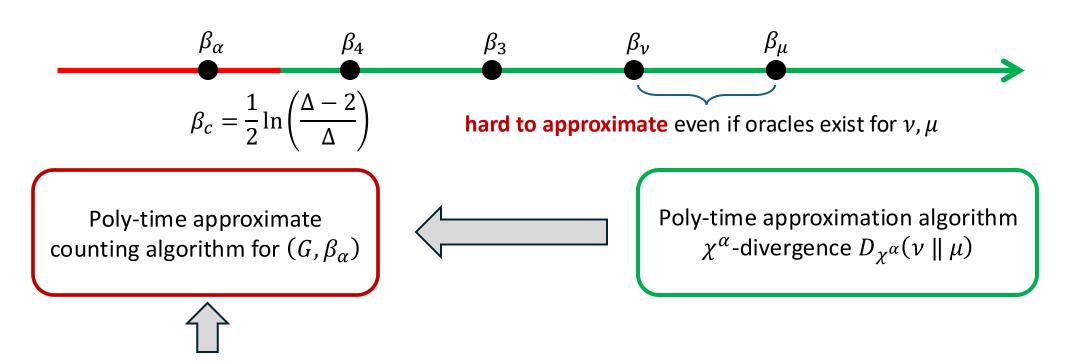
$$\beta_{c} = \frac{1}{2} \ln \left(\frac{\Delta - 2}{\Delta} \right)$$

Theorem: hardness of approximation [F. and Fu, 2025]

Fix integers $\alpha \geq 2$ and $\Delta \geq 3$. Fix $\beta_{\mu} > \beta_{\nu} \geq \beta_{c}(\Delta)$ such that

$$\beta_{\alpha} = \beta_{\mu} + k(\beta_{\nu} - \beta_{\mu}) < \beta_{c}(\Delta).$$

Unless NP=RP, no FPRAS for χ^{α} -divergence between (G, β_{ν}) and (G, β_{μ}) on Δ -regular graphs G



Approximate counting is *hard* for $\beta_{\alpha} < \beta_{c}$ [Sly and Sun 14, Galanis, Štefankovič and Vigoda 16]

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Computational Phase Transition

Corollary: approximation algorithms [F. and Fu, 2025]

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FPRAS exists for χ^{α} -divergence between (G, β_{ν}) and (G, β_{μ}) on graph G with max degree Δ

Summary of algorithmic results

Divergence	Function <i>f</i>	Existence of oracles for sampling / counting
χ^{α} for $\alpha \in \mathbb{N}$	$f(x) = \frac{1}{2} x - 1 ^{\alpha}$	$\left\{ \left(G, J^{(k)}, h^{(k)}\right) \mid 0 \le k \le \alpha \right\}$
α -divergence for $\alpha \neq 0,1$	$f(x) = \frac{t^{\alpha} - \alpha t - (1 - \alpha)}{\alpha(\alpha - 1)}$	$\{(G,J^{(k)},h^{(k)}) \mid k=0,1,\alpha\}$
Kullback–Leibler Rényi Jensen-Shannon	$f(x) = x \ln x - x + 1$ $f(x) = -\ln x + x - 1$ $f(x)$ $= \frac{1}{2} \left(x \ln x - (x+1) \ln \frac{x+1}{2} \right)$	$\left(G,J^{(u)},h^{(u)} ight)$ and $\left(G,J^{(\mu)},h^{(\mu)} ight)$
Squared Hellinger	$f(x) = \frac{1}{2}(\sqrt{x} - 1)^2$	$(G, J^{(\nu)}, h^{(\nu)}), (G, J^{(\mu)}, h^{(\mu)})$ and $(G, \frac{J^{(\nu)} + J^{(\mu)}}{2}, \frac{h^{(\nu)} + h^{(\mu)}}{2})$

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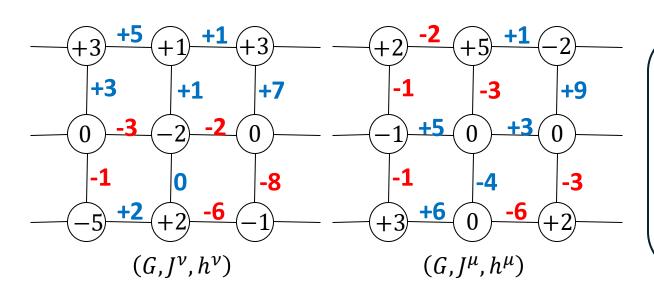
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 $\operatorname{poly}(n/\epsilon)$ -time algorithms for approximate $D_{\chi^{\alpha}}(\nu \parallel \mu)$

Parameter distance



 $d_{par}(\nu, \mu) = \max\{\text{edge diff, vertex diff}\}$

- Edge diff: $\max_{e \in E} |J^{\nu}(e) J^{\mu}(e)|$
- Vertex diff: $\max_{v \in V} \frac{|h^{\nu}(v) h^{\mu}(v)|}{\deg_G(v) + 1}$

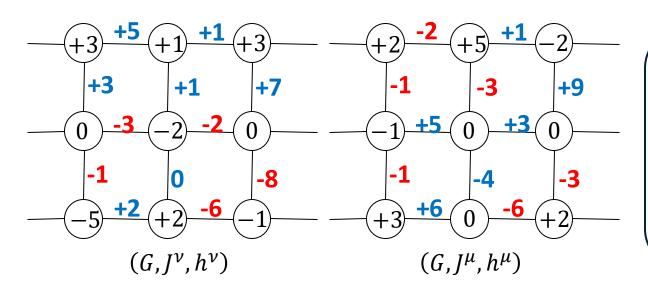
What is the relation between *parameter distance* and *TV-distance*/ χ^{α} -divergence?



If two Ising model are different, then we can blame it on a node or an edge

a similar observation (but with different definitions of distances) was made in "Test Ising Models" [Daskalakis, Dikkala and Kamath 19]

Parameter distance



 $d_{par}(\nu, \mu) = \max\{\text{edge diff, vertex diff}\}$

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What is the relation between *parameter distance* and *TV-distance*/ χ^{α} -divergence?

$$d_{
m par}(
u,\mu)$$
 is *large*

marginal lower bound

[Feng, Liu and Yang 2025]

 $D_{TV}(\nu \parallel \mu)$ is *large*

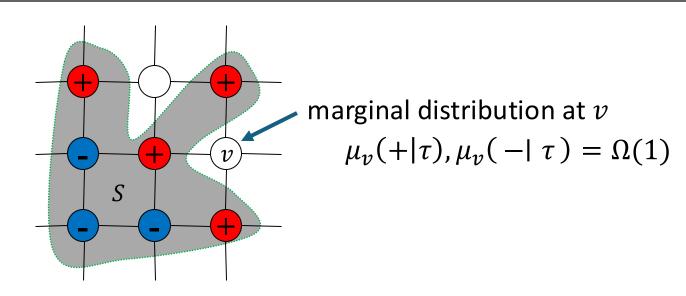
Our result: total variation distance ($\alpha = 1$)

Definition Marginal lower bound for Ising model

For any subset $S \subseteq V$, any vertex $v \in V \setminus S$, any pinning $\tau \in \{-1, +1\}^S$,

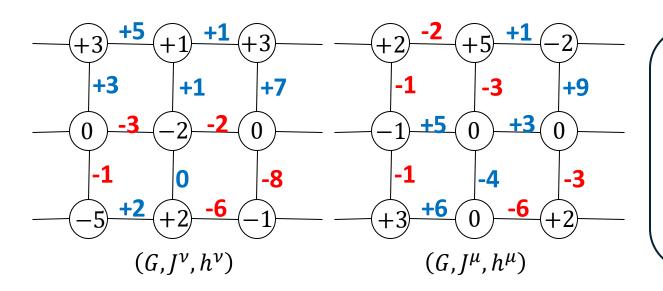
$$\forall c \in \{-1, +1\}, \qquad \mu_v(c \mid \tau) = \Omega(1)$$

Under any conditional, the marginal distribution on one vertex cannot be too biased



The assumption also appeared in *learning* [Bresler15], sampling and counting [CLV21]

Parameter distance



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What is the relation between *parameter distance* and *TV-distance*/ χ^{α} -divergence?



- Different edges/nodes may differ in different directions. Overall, all differences may be cancelled out
- The cancellation *cannot happen* for Ising models with marginal lower bounds

For two Ising models ν and μ both with marginal lower bound $b=\Omega(1)$

- The TV-distance: $D_{TV}(\nu \parallel \mu) \ge \frac{b^2}{2} d_{par}(\nu, \mu)$ [Feng, Liu and Yang 2025]
- The χ^{α} -divergence: $D_{\chi^{\alpha}}(\nu \parallel \mu) \geq \frac{b^{2\alpha}}{2} d_{\text{par}}^{\alpha}(\nu, \mu)$ [Feng and Fu 2025]
- A family of f-divergence can also be lower bounded in terms of b [Feng and Fu 2025]

Compute and check whether
$$d_{par}(\nu, \mu) \leq \frac{1}{poly(n)}$$
?

- Yes. All parameters in (J^{ν}, J^{μ}) and (h^{ν}, h^{μ}) are **similar** to each other use the similarity of parameters to design **well-concentrated** estimator
- No. Then the $D_{\chi^{\alpha}}(\nu\|\mu)$ is *large*, at least $\frac{1}{\operatorname{poly}(n)}$ relative-error approximation \implies we can tolerate certain large error

The algorithm for small parameter distance $d_{\text{par}}(\nu, \mu) \leq \frac{1}{\text{poly}(n)}$

$$D_{\chi^{\alpha}}(\nu,\mu) = \frac{1}{2} \sum_{x \in \{\pm\}^{V}} \mu(x) \left| 1 - \frac{\nu(x)}{\mu(x)} \right|^{\alpha} = \frac{1}{2} \sum_{x \in \{\pm\}^{V}} \mu(x) \left| 1 - \frac{w_{\nu}(x)}{w_{\mu}(x)} \cdot \frac{Z_{\mu}}{Z_{\nu}} \right|^{\alpha}$$

$$W = \frac{w_{\nu}(X)}{w_{\mu}(X)} = \exp\left(X^{T} \left(J^{\nu} - J^{\mu}\right)X + X^{T} \left(h^{\nu} - h^{\mu}\right)\right) \text{ for } X \sim \mu \qquad \Longrightarrow \qquad \mathbb{E}[W] = \frac{Z_{\nu}}{Z_{\mu}}$$

- Draw random samples of W to estimate $\frac{1}{\mathbb{E}[W]} = \frac{Z_{\mu}}{Z_{\nu}}$
- Draw random samples of W to estimate the expectation of $\frac{1}{2} \left| 1 W \cdot \frac{Z_{\mu}}{Z_{\nu}} \right|^{\alpha}$
- How many samples do we need?
- How well are W and $\left|1-W\cdot\frac{Z_{\mu}}{Z_{\nu}}\right|^{\alpha}$ concentrated around their mean?

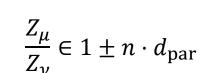
The algorithm for small parameter distance $d_{par}(\nu, \mu) \leq \frac{1}{poly(n)}$

$$D_{\chi^{\alpha}}(\nu,\mu) = \frac{1}{2} \sum_{x \in \{\pm\}^{V}} \mu(x) \left| 1 - \frac{\nu(x)}{\mu(x)} \right|^{\alpha} = \frac{1}{2} \sum_{x \in \{\pm\}^{V}} \mu(x) \left| 1 - \frac{w_{\nu}(x)}{w_{\mu}(x)} \cdot \frac{Z_{\mu}}{Z_{\nu}} \right|^{\alpha}$$

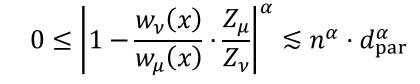
$$d_{\text{par}}(\nu,\mu) \le \frac{1}{\text{poly}(n)}$$

$$\frac{w_{\nu}(x)}{w_{\mu}(x)} \in 1 \pm n \cdot d_{\text{par}}$$









$$\leq \operatorname{poly}(n) \cdot D_{\chi^{\alpha}}(\nu \parallel \mu)$$

Divergence is at least $\Omega(d_{\mathrm{par}}^{lpha})$

$$\mathbf{Var}\left[\left|1 - W \cdot \frac{Z_{\mu}}{Z_{\nu}}\right|^{\alpha}\right] \le \operatorname{poly}(n) \cdot D_{\chi^{\alpha}}^{2}(\nu \parallel \mu)$$

 $\operatorname{poly}\left(\frac{n}{\varepsilon}\right)$ samples are enough to achieve $(1\pm\varepsilon)$ *relative error* approximation

Algorithm *sketch* for large parameter distance $d_{\mathrm{par}}(\nu,\mu) > \frac{1}{\mathrm{poly}(n)}$

 $d_{\sf par}(
u,\mu)$ is **large**

marginal lower bound

[Feng, Liu and Yang 2025]

 $D_{\chi^{\alpha}}(\nu,\mu)$ is *large*

If $D_{\chi^{\alpha}}(\nu \parallel \mu) > \frac{1}{\text{poly}(n)}$, then the following lower bound holds

$$D_{\chi^{\alpha}}(v \parallel \mu) = \frac{1}{2} \sum_{x \in \{\pm\}^{V}} \mu(x) \left| 1 - \frac{\nu(x)}{\mu(x)} \right|^{\alpha} \ge \frac{1}{\text{poly}(n)} \cdot \sum_{x \in \{\pm\}^{V}} \mu(x) \left(1 + \frac{\nu(x)}{\mu(x)} \right)^{\alpha}$$

$$\sum_{\sigma \in \{+\}^{V}} \frac{\nu^{k}(\sigma)}{\mu^{k-1}(\sigma)} = \frac{Z_{\mu}^{k-1}}{Z_{\nu}^{k}} \cdot \sum_{x \in \{+\}^{V}} \exp(x^{T}(kJ^{\nu} - (k-1)J^{\mu})x + x^{T}(kh^{\nu} - (k-1)h^{\mu}))$$

Algorithm for χ^{α} -divergence between two Ising models

Theorem: approximation algorithm [F and Fu, 2025]

Two Ising models $\nu = \text{Ising}(G, J^{\nu}, h^{\nu})$ and $\mu = \text{Ising}(G, J^{\mu}, h^{\mu})$ with marginal lower bound

A *family* of Ising models
$$\mathcal{F} = \{(G, J^{(k)}, h^{(k)}) \mid \text{integer } 0 \leq k \leq \alpha\}$$
, where

$$J^{(k)} = kJ^{\nu} - (k-1)J^{\mu}$$

$$h^{(k)} = kh^{\nu} - (k-1)h^{\mu}$$

All Ising models in \mathcal{F} admit $\operatorname{poly}(n/\epsilon)$ -time algos for

- sampling
- approximate counting



 $\operatorname{poly}(n/\epsilon)\text{-time algorithms for}$ $\operatorname{approximate} D_{\chi^\alpha}(\nu \parallel \mu)$

If $D_{\chi^{\alpha}}(\nu \parallel \mu) > \frac{1}{\text{poly}(n)}$, then the following lower bound holds

$$D_{\chi^{\alpha}}(\nu \parallel \mu) = \frac{1}{2} \sum_{x \in \{\pm\}^{V}} \mu(x) \left| 1 - \frac{\nu(x)}{\mu(x)} \right|^{\alpha} \ge \frac{1}{\text{poly}(n)} \cdot \sum_{x \in \{\pm\}^{V}} \mu(x) \left(1 + \frac{\nu(x)}{\mu(x)} \right)^{\alpha}$$

$$\sum_{\sigma \in \{+\}^{V}} \frac{\nu^{k}(\sigma)}{\mu^{k-1}(\sigma)} = \frac{Z_{\mu}^{k-1}}{Z_{\nu}^{k}} \cdot \sum_{x \in \{+\}^{V}} \exp(x^{T} J^{(k)} x + x^{T} h^{(k)}) = \frac{Z_{\mu}^{k-1}}{Z_{\nu}^{k}} \cdot Z^{(k)}$$

If $D_{\chi^{\alpha}}(\nu \parallel \mu) > \frac{1}{\text{poly}(n)}$, then the following lower bound holds

$$D_{\chi^{\alpha}}(\nu \parallel \mu) = \frac{1}{2} \sum_{x \in \{\pm\}^{V}} \mu(x) \left| 1 - \frac{\nu(x)}{\mu(x)} \right|^{\alpha} \ge \frac{1}{\text{poly}(n)} \cdot \sum_{x \in \{\pm\}^{V}} \mu(x) \left(1 + \frac{\nu(x)}{\mu(x)} \right)^{\alpha}$$

- Approx. each $\sum_{x \in \{\pm\}^V} \frac{v^k(x)}{\mu^{k-1}(x)}$ with relative err $\frac{\varepsilon}{\operatorname{poly}(n)} = additive \ err \frac{\varepsilon}{\operatorname{poly}(n)} \cdot \sum_{x \in \{\pm\}^V} \frac{v^k(x)}{\mu^{k-1}(x)}$
- Approx. whole sum $D_{\chi^{\alpha}}(\nu \parallel \mu)$ with *additive error*

$$\frac{\varepsilon}{\operatorname{poly}(n)} \cdot \sum_{0 \le k \le \alpha} {\alpha \choose k} \sum_{x \in \{\pm\}^V} \frac{v^k(x)}{\mu^{k-1}(x)} = \frac{\varepsilon}{\operatorname{poly}(n)} \cdot \sum_{x \in \{\pm\}^V} \mu(x) \left(1 + \frac{v(x)}{\mu(x)}\right)^{\alpha} \le \varepsilon \cdot D_{\chi^{\alpha}}(v \parallel \mu)$$
relative err ε

$$D_{\chi^{\alpha}}(\nu,\mu) = \frac{1}{2} \sum_{x:\nu(x)>\mu(x)} \mu(x) \left(\frac{\nu(x)}{\mu(x)} - 1\right)^{\alpha} + \frac{1}{2} \sum_{x:\nu(x)<\mu(x)} \mu(x) \left(\frac{\nu(x)}{\mu(x)} - 1\right)^{\alpha}$$

$$= \frac{1}{2} \sum_{0 \le k \le \alpha} (-1)^{\alpha - k} {\alpha \choose k} \left(\sum_{x: \nu(x) > \mu(x)} \frac{\nu^k(x)}{\mu^{k - 1}(x)} + (-1)^{\alpha} \sum_{x: \nu(x) < \mu(x)} \frac{\nu^k(x)}{\mu^{k - 1}(x)} \right)$$

- Sample $X \sim \text{Ising}(G, J^{(k)}, h^{(k)})$ Estimator: $W_k = \mathbf{1}[\nu(X) > \mu(X)] \cdot \frac{Z_{\mu}^{k-1}}{Z_{\nu}^k} Z^{(k)}$ $\mathbb{E}[W_k] =$



$$\mathbb{E}[W_k] =$$

We can only approximate $\nu(X)$ and $\mu(X)$, but **cannot** exactly compute $\mathbf{1}[\nu(X) > \mu(X)]$

We only make mistake when $\nu(X)$ is very close to $\mu(X)$

We use random samples to *estimate* the expectation and we need to put all terms together

The concentration error can be bounded in a similar way as that for even α case

Open problems

- remove the marginal lower bound assumption.
- more general graphical models or general distributions
- χ^{α} -divergence for real number α or other divergences
- deterministic approximation algorithms
- faster algorithms (current algorithm require $\alpha = O(1)$ with running time $n^{O(\alpha)}$)
- connections or applications in learning and testing

Thank You