

Modern Difference-in-Difference Designs

Ryan T. Moore

American University

The Lab @ DC

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Calloway-Sant'Anna

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 - ▶ pre-tr outcome affected by (obs/unobs) predictors of post-tr outcome

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What's weak about this question? What's being ignored?

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This is **not**

$$ATT = \overline{[Y(1)|T = 1]} - \overline{[Y(0)|T = 1]}$$

Why not?

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- ▶ Change in Control: removes part of change in emp that would have occurred anyway, in absence of law change.
- ▶ (Economy growing, or fast food restaurants hit by health scares and lay off workers, ...)

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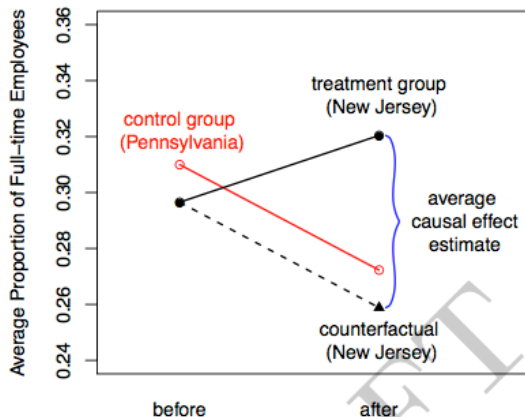
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- ▶ Key assumption: “parallel trends” – change in PA is what change in NJ *would have been* without new minimum wage.

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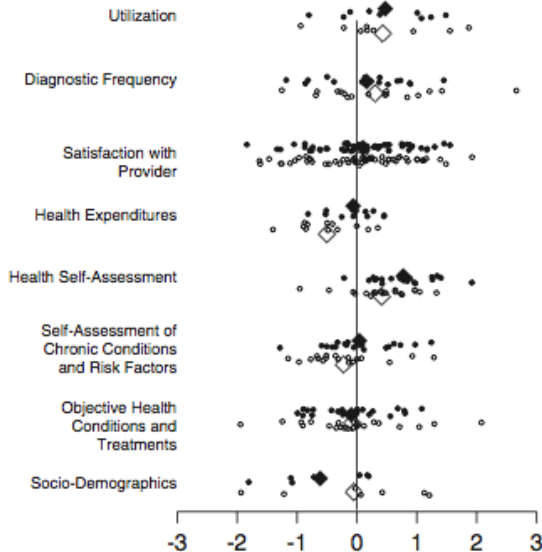
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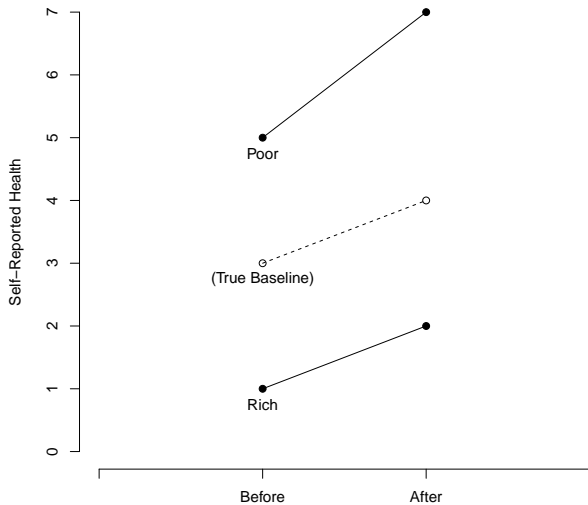


Self-Reported Health in *Seguro Popular* Evaluation

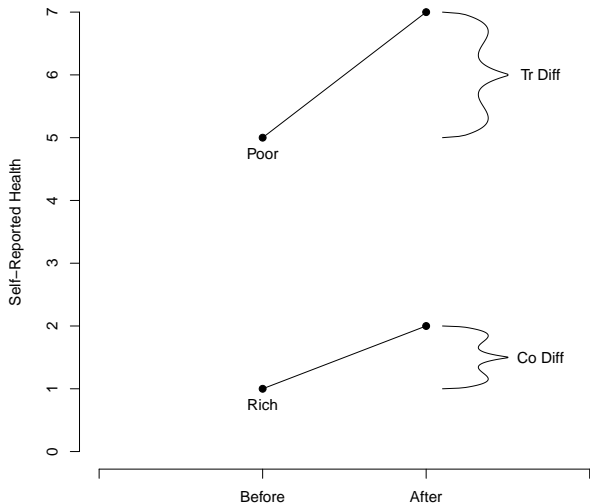
King, et al. (2007)



Simplified Example



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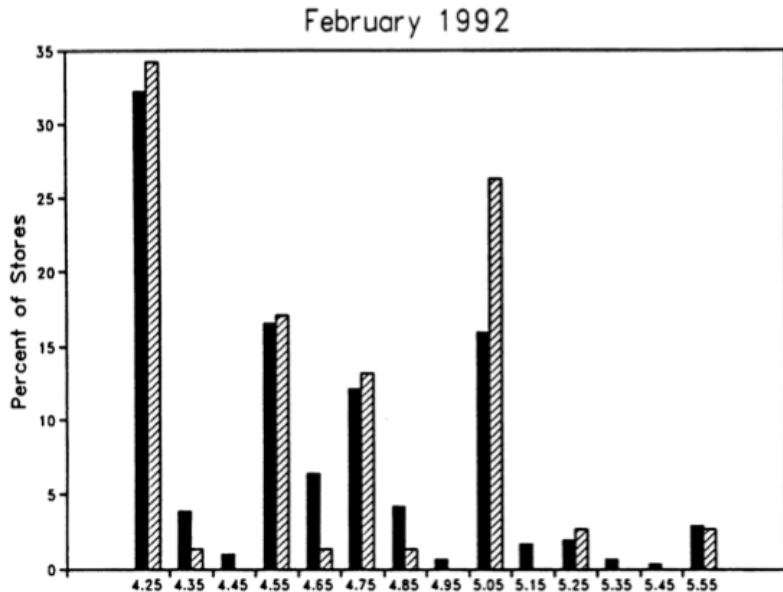


The Difference-in-Differences Estimator

Units	Outcome Before	Outcome After	Difference
Treated	\bar{Y}^{T_B}	\bar{Y}^{T_A}	$\bar{Y}^{T_A} - \bar{Y}^{T_B}$
Control	\bar{Y}^{C_B}	\bar{Y}^{C_A}	$\bar{Y}^{C_A} - \bar{Y}^{C_B}$
	$\bar{Y}^{T_B} - \bar{Y}^{C_B}$	$\bar{Y}^{T_A} - \bar{Y}^{C_A}$	$(\bar{Y}^{T_A} - \bar{Y}^{T_B}) - (\bar{Y}^{C_A} - \bar{Y}^{C_B})$

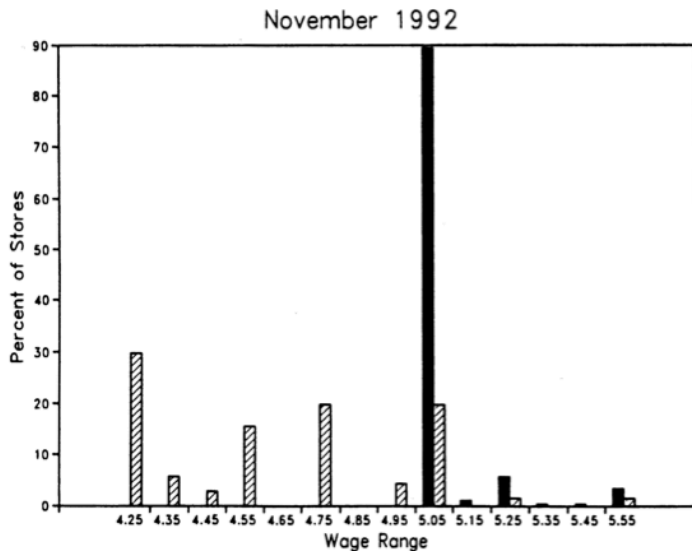
Example: Effect of Minimum Wage on Employment

The Baseline



Example: Effect of Minimum Wage on Employment

The Policy Change: Significant Shock



■ Minimum Wage ▨ Pennsylvania

Example: Effect of Minimum Wage on Employment

Variable	Difference, NJ – PA		
	PA (i)	NJ (ii)	NJ – PA (iii)
1. FTE employment before, all available observations	23.33 (1.35)	20.44 (0.51)	-2.89 (1.44)
2. FTE employment after, all available observations	21.17 (0.94)	21.03 (0.52)	-0.14 (1.07)
3. Change in mean FTE employment	-2.16 (1.25)	0.59 (0.54)	2.76 (1.36)

Extension 1: Regression Difference-in-Differences

Let

- ▶ $T_i \in \{0, 1\}$ be treatment status (“ever-treated”)
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Model

$$y_i = \beta_0 + \beta_1 T_i + \beta_2 A_i + \beta_3 T_i A_i + \epsilon_i$$

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Then

$$\begin{aligned}\bar{Y}^{T_A} &= \beta_0 + \beta_1 + \beta_2 + \beta_3 \\ \bar{Y}^{T_B} &= \beta_0 + \beta_1 \\ \bar{Y}^{C_A} &= \beta_0 + \beta_2 \\ \bar{Y}^{C_B} &= \beta_0\end{aligned}$$

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And

$$\begin{aligned}(\bar{Y}^{T_A} - \bar{Y}^{T_B}) - (\bar{Y}^{C_A} - \bar{Y}^{C_B}) &= ([\beta_0 + \beta_1 + \beta_2 + \beta_3] - [\beta_0 + \beta_1]) \\ &\quad - ([\beta_0 + \beta_2] - \beta_0) \\ &= (\beta_2 + \beta_3) - (\beta_2) \\ &= \beta_3\end{aligned}$$

Extension 1: Regression Difference-in-Differences

Note: $T_i \in \{0, 1\}$ needs to be “ever-treated”
(not “currently treated”)

Simulation:

- ▶ 40 units (20 Tr, 20 Co)
- ▶ 2 time periods (2018, 2022)
- ▶ Baseline outcome difference $\text{Tr} - \text{Co} = 1$
- ▶ True TE = 2
- ▶ (TE occurs in 2022, only for Tr)

Extension 1: Regression Difference-in-Differences

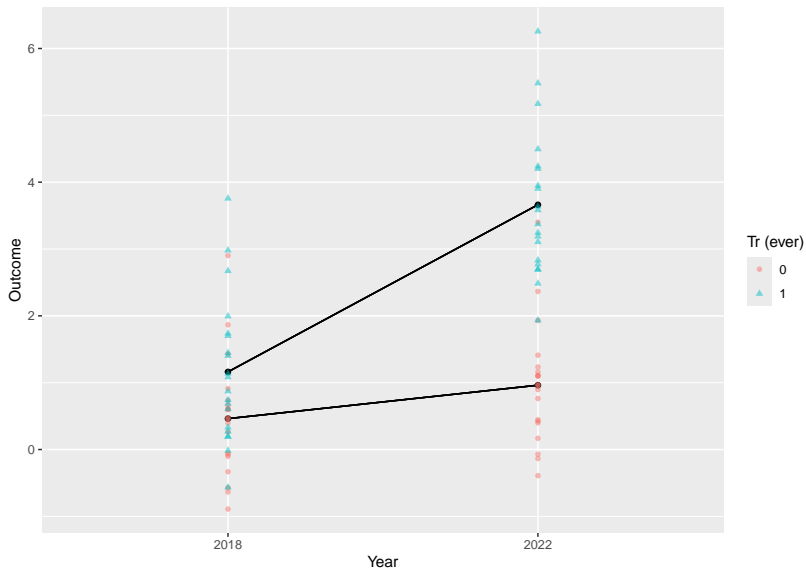


Figure 2: Simulated Data DiD

Extension 1: Regression Difference-in-Differences

Two possible regressions:

$$Y = \beta_0 + \beta_1(2022) + \beta_2(\text{TrEver}) + \beta_3(2022 \times \text{TrEver})$$

Extension 1: Regression Difference-in-Differences

Two possible regressions:

$$Y = \beta_0 + \beta_1(2022) + \beta_2(\text{TrEver}) + \beta_3(2022 \times \text{TrEver})$$

versus

$$Y = \beta_0 + \beta_1(2022) + \beta_2(\text{TrNow}) + \beta_3(2022 \times \text{TrNow})$$

Extension 1: Regression Difference-in-Differences

Table 1: Ever-treated vs. Currently-treated DiD Regressions

	Outcome	
	y	
	(1)	(2)
2022	0.500 (0.313)	0.150 (0.278)
Tr (Ever)	0.700** (0.313)	
2022 x Tr (Ever)	2.000*** (0.442)	
Tr (Now)		2.700*** (0.321)
2022 x Tr (Now)		
(Intercept)	0.462** (0.221)	0.812*** (0.160)
Observations	80	80
R ²	0.623	0.598
Adjusted R ²	0.608	0.588
Residual Std. Error	0.989 (df = 76)	1.015 (df = 77)
F Statistic	41.827*** (df = 3; 76)	57.261*** (df = 2; 77)

Note:

*p<0.1; **p<0.05; ***p<0.01

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$$Y = \underbrace{0.46}_{\text{Co Mean 2018}} + \underbrace{0.5}_{\text{Time Trend}} (2022) + \underbrace{0.7}_{\text{Baseline Diff}} (\text{TrEver}) + \underbrace{2}_{\widehat{ATT}} (2022 \times \text{TrEver})$$

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versus

$$Y = \underbrace{0.81}_{\text{Part of Co Mean 2018} + \text{Time Trend??}} + \underbrace{0.15}_{\text{Rest of Time Trend??}} (2022) + \underbrace{2.7}_{\text{Baseline Diff} + \widehat{ATT}??} (\text{TrNow}) + \underbrace{NA}_{\text{Unest Interaction for } \widehat{ATT}??} (2022 \times \text{TrNow})$$

Extension 2: Regression Difference-in-Differences

Related: Use change scores or regressor variable method?

(Allison 90)

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$$y_{i2} = \alpha + \beta_1 y_{i1} + \beta_2 T_i + \epsilon_i \quad (1)$$

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Related: Use change scores or regressor variable method?

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$$y_{i1} = \alpha + \gamma G_i + \epsilon_{i1} \quad (2)$$

$$y_{i2} = \alpha + \underbrace{\tau}_{\text{secular trend}} + \gamma G_i + \delta T_i + \epsilon_{i2} \quad (3)$$

$$y_{i2} - y_{i1} = \tau + \delta T_i + (\epsilon_{i2} - \epsilon_{i1}) \quad (4)$$

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$$y_{i2} - y_{i1} = \tau + \delta T_i + (\epsilon_{i2} - \epsilon_{i1}) \quad (4)$$

- ▶ Change score is a DiD estimate
- ▶ Worse if pretest not very strong, and restriction of $\beta_1 = 1$ creates bias.
- ▶ Change score is more restrictive on parameter, but makes different assumption about $Cor(y_1, \epsilon)$.

Extension 3: Regression Difference-in-Differences

Add other covariates:

$$y_i = \beta_0 + \beta_1 T_i + \beta_2 A_i + \beta_3 T_i A_i + \gamma' \mathbf{x}_i + \epsilon_i$$

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(At $\alpha = 0.05$, false findings actually 0.45!)

Actually, be careful about bias with many time periods, too!

(See here.)

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- ▶ Useful for estimating effect across two subgroups
- ▶ (E.g., Empl vs. unempl'd; or placebo robustness check if one subgroup should *not* be affected by treatment)

Extension 4: Difference in (Differences-in-Differences)

Group 1:

Units	Outcome Before	Outcome After	Difference
Treated	T_{1B}	T_{1A}	$T_{1A} - T_{1B}$
Control	C_{1B}	C_{1A}	$C_{1A} - C_{1B}$
	$T_{1B} - C_{1B}$	$T_{1A} - C_{1A}$	$(T_{1A} - T_{1B}) - (C_{1A} - C_{1B})$

Group 2:

Units	Outcome Before	Outcome After	Difference
Treated	T_{2B}	T_{2A}	$T_{2A} - T_{2B}$
Control	C_{2B}	C_{2A}	$C_{2A} - C_{2B}$
	$T_{2B} - C_{2B}$	$T_{2A} - C_{2A}$	$(T_{2A} - T_{2B}) - (C_{2A} - C_{2B})$

DiDiD Estimate:

$$[(T_{2A} - T_{2B}) - (C_{2A} - C_{2B})] - [(T_{1A} - T_{1B}) - (C_{1A} - C_{1B})]$$

Extension 5: Multiple Time Periods, Staggered Designs

- ▶ Callaway and Sant'Anna (2021b) on multiple time periods

Extension 5: Multiple Time Periods, Staggered Designs

- ▶ Callaway and Sant'Anna (2021b) on multiple time periods
- ▶ Goodman-Bacon (2021) on staggered treatment assignment

Extension 5: Multiple Time Periods, Staggered Designs

- ▶ Callaway and Sant'Anna (2021b) on multiple time periods
- ▶ Goodman-Bacon (2021) on staggered treatment assignment
- ▶ Package `did`

The Stepped Wedge Design

Calloway-Sant'anna Approach: Multiple Time Periods

With *many* time periods, challenging to defining **the** estimand (like ATT).

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- ▶ ID disaggregated parameters of interest

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The Callaway and Sant'Anna (2021b) approach:

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- ▶ Decide how to aggregate them

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With *many* time periods, challenging to defining **the** estimand (like ATT).

The Callaway and Sant'Anna (2021b) approach:

- ▶ ID disaggregated parameters of interest
- ▶ Decide how to aggregate them
- ▶ Estimate them

Calloway-Sant'anna Approach: Multiple Time Periods

Central disaggregated component:

the *group-time average treatment effect*.

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- ▶ $ATT(g, t)$

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("effect for those first treated in 2020")

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- ▶ Cumulative
("total effect up to 2020")
- ▶ Overall
("total effect up to 2024")

Calloway-Sant'anna Approach: Multiple Time Periods

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("effect for those w/ 2 years treatment")
- ▶ Group
("effect for those first treated in 2020")
- ▶ Cumulative
("total effect up to 2020")
- ▶ Overall
("total effect up to 2024")
- ▶ Overall, group as unit
("total effect up to 2024: get state avgs, avg those")

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Key assumptions:

1. No treated at $t = 1$. If treated at t^* , then treated at $t^* + 1$. Once treated, stay “treated”.

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Calloway-Sant'Anna Approach: Multiple Time Periods

Key assumptions:

1. No treated at $t = 1$. If treated at t^* , then treated at $t^* + 1$. Once treated, stay “treated”.
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3/4/5 define valid comparison group, thus analysis sample.

C-S Example

Effect of US state min wages on *teen* employment.

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- ▶ Annual data, 2003-2007

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Callaway and Sant'Anna (2021a): `did`

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Calculate all the $ATT(g, t)$ effects, for each (state \times year)

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cs_out <- att_gt(ymame = "lemp",  
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                 tname = "year",  
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                 data = mpdta,  
                 est_method = "reg"  
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```

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C-S Example

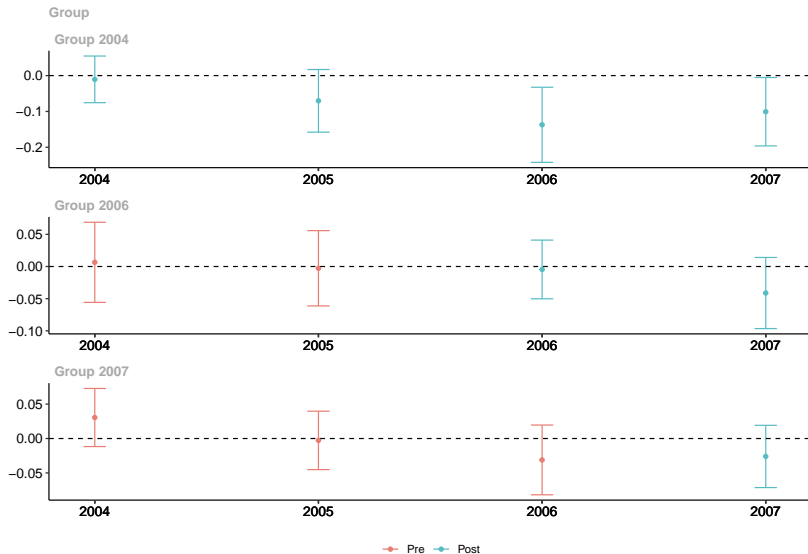
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- ▶ `est_method`: regression, IPW, doubly-robust, etc.
- ▶ `xformula`: formula for covariate adjustment
- ▶ `control_group`: "nevertreated", "notyettreated"

C-S Example

```
ggdid(cs_out)
```



C-S Example

- ▶ “Pre” C-S $ATT(g, t)$ values: *pre*-test of parallel trends assumption

C-S Example

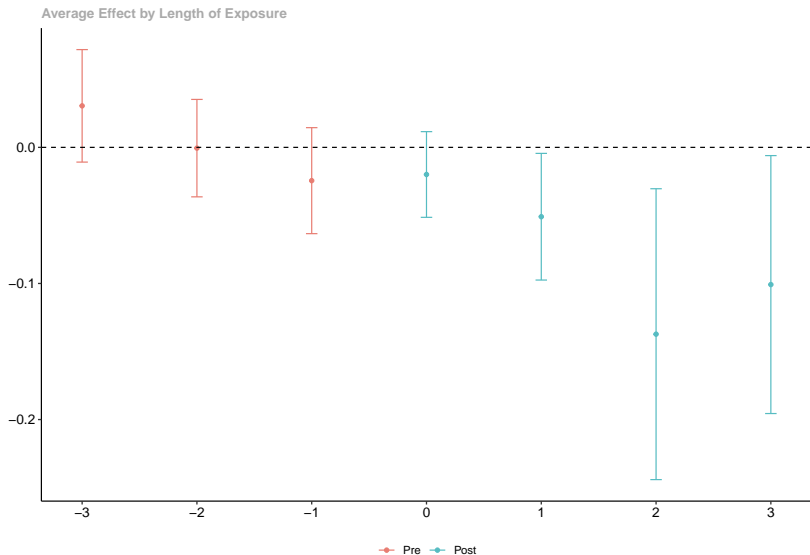
- ▶ “Pre” C-S $ATT(g, t)$ values: *pre*-test of parallel trends assumption
- ▶ “Post” C-S $ATT(g, t)$ values: effects of min wage policy

C-S Example

```
cs_es_out <- aggte(cs_out, type = "dynamic")
```

C-S Example

```
ggdid(cs_es_out)
```



C-S Example

```
summary(cs_es_out)
```

C-S Example

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Overall summary of ATT's based on event-study/dynamic aggregation:

ATT	Std. Error	[95% Conf. Int.]
-0.0772	0.0222	-0.1207 -0.0338 *

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```

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“Overall” effect of participating: get avg effect at each exposure length, avg those across different exposure lengths.

C-S Example

Alternative: get ATE for each group, avg across groups.

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group_effects <- aggte(cs_out, type = "group")
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```

Overall summary of ATT's based on group/cohort aggregation:

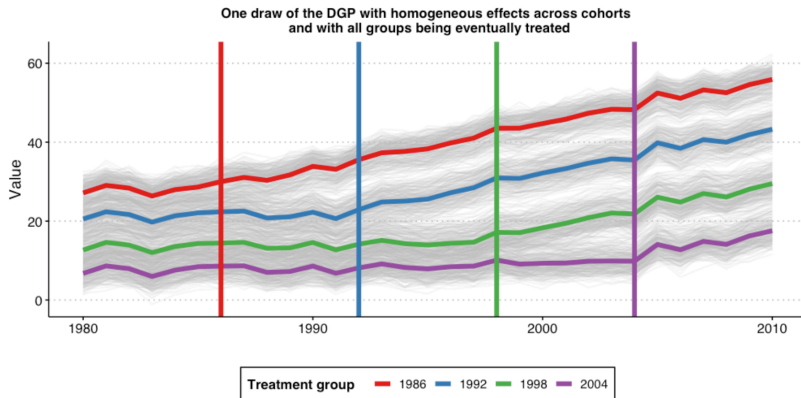
ATT	Std. Error	[95% Conf. Int.]
-0.031	0.0127	-0.056 -0.006 *

Group Effects:

Group	Estimate	Std. Error	[95% Simult. Conf. Band]
2004	-0.0797	0.0301	-0.1455 -0.0140 *
2006	-0.0229	0.0172	-0.0606 0.0148
2007	-0.0261	0.0169	-0.0631 0.0109

Linear Models and Two-Way Fixed-Effects

Callaway and Sant'Anna (2023)



The above plot shows the 1,000 individual values of $Y_{i,t}$, as well as the average by treatment-group (the thicker colorful lines), and the vertical lines show the period when treatment begins for each group. Because treatment effects here grows linearly with time elapsed since treatment started, we can see that the earliest treated units end up with the highest value of Y , and that the difference grows by the gap between treatment years.

Linear Models and Two-Way Fixed-Effects Event Study

Model

$$Y_{i,t} = \alpha_i + \beta_t + \gamma(Pre) + \delta_j(Lag_j) + \zeta_k(Lead_k) + \epsilon_{i,t}$$

- ▶ Indicators for units
- ▶ Indicators for time periods
- ▶ Indicators for lags (“tr how long ago?”)
- ▶ Indicators for leads (“tr how far into future?”)
- ▶ (Cluster SE’s on “states” of assignment)

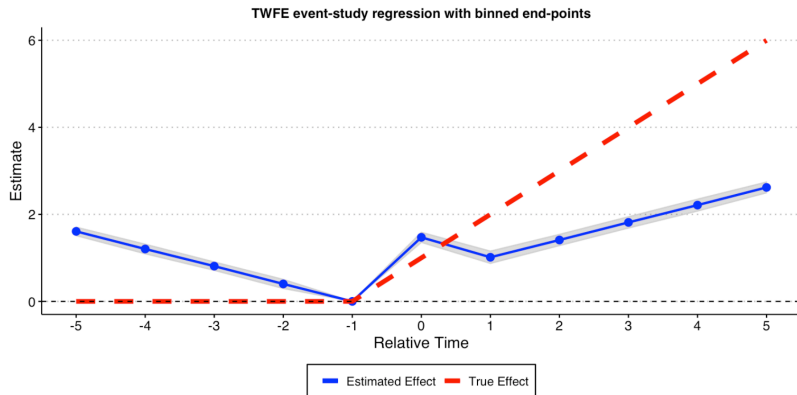
Linear Models and Two-Way Fixed-Effects Event Study

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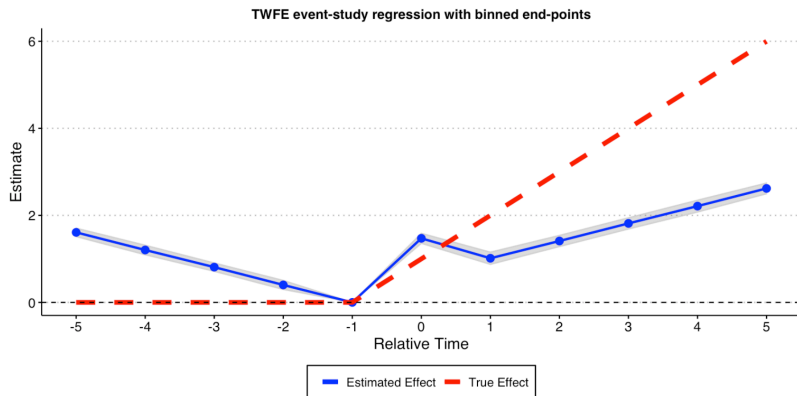
$$Y_{i,t} = \alpha_i + \beta_t + \gamma(Pre) + \delta_j(Lag_j) + \zeta_k(Lead_k) + \epsilon_{i,t}$$

- ▶ δ_j often interpreted as “ATE of j periods of treatment”
- ▶ ζ_j often interpreted as “pre-tr trends”

Linear Models and Two-Way Fixed-Effects Event Study



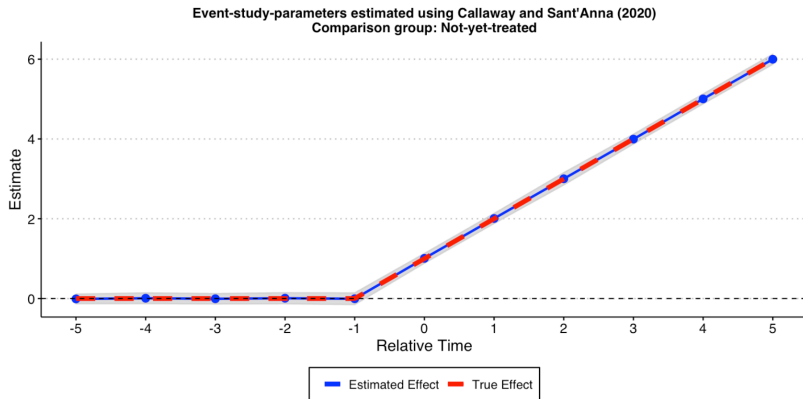
Linear Models and Two-Way Fixed-Effects Event Study



Severe bias for “effect of 3 years of Tr”, e.g.!

Linear Models and Two-Way Fixed-Effects Event Study

Instead, using C-S approach correctly ID's identical lead/pre-tr trends and cumulative effects of Tr!



Linear Models and Two-Way Fixed-Effects Event Study

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Linear Models and Two-Way Fixed-Effects Event Study

- ▶ If you include a never-treated group, TWFE-ES can recover truth (but so does Calloway-Sant'Anna)
- ▶ But, if there are heterogeneous treatment effects, TWFE-ES won't recover truth (even with never-treated group)
- ▶ Calloway-Sant'Anna does!

Thanks!

rtm@american.edu
www.ryantmoore.org

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