

Robust and scalable simulation-based inference

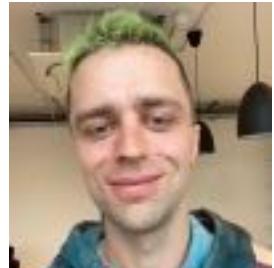
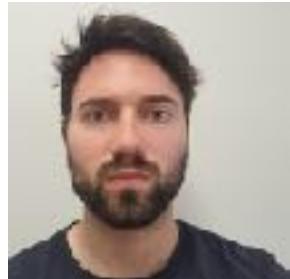
François-Xavier Briol
Department of Statistical Science
University College London
<https://fx briol.github.io/>
<https://fsml-ucl.github.io/>



Greek Stochastics - Folegandros 2025



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A (slightly biased) introduction to simulation-based inference



Intractable likelihoods

Our data: $y_1, \dots, y_n \sim \mathbb{Q}$

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Unknown data-generating
process defined on the
data-space \mathcal{X} .

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Our model: $\{\mathbb{P}_\theta\}_{\theta \in \Theta}$

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Our job is to recover θ^*

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Maximum likelihood:

$$\hat{\theta}_n := \arg \max_{\theta \in \Theta} \prod_{i=1}^n p(y_i | \theta)$$

Bayesian inference:

$$p(\theta | y_1, \dots, y_n) \propto \prod_{i=1}^n p(y_i | \theta) p(\theta)$$

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Simulation-based inference (SBI)

- A simulator (\mathbb{U}, G_θ) such that a draw from \mathbb{P}_θ can be obtained as:

$$x_i = G_\theta(u_i)$$

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\uparrow

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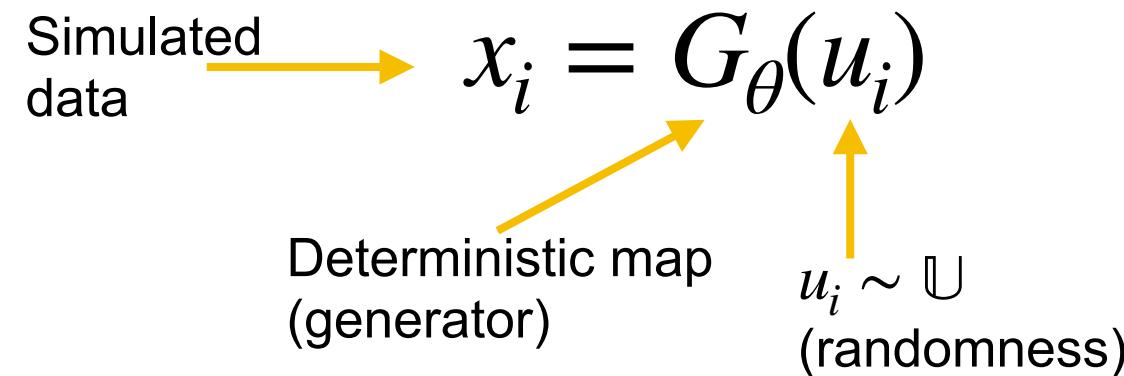
Deterministic map
(generator)

$u_i \sim \mathbb{U}$
(randomness)

The diagram illustrates the process of generating simulated data. At the top, the equation $x_i = G_\theta(u_i)$ is shown. Two arrows point downwards from this equation to the labels below. The left arrow is labeled "Deterministic map (generator)". The right arrow is labeled " $u_i \sim \mathbb{U}$ (randomness)".

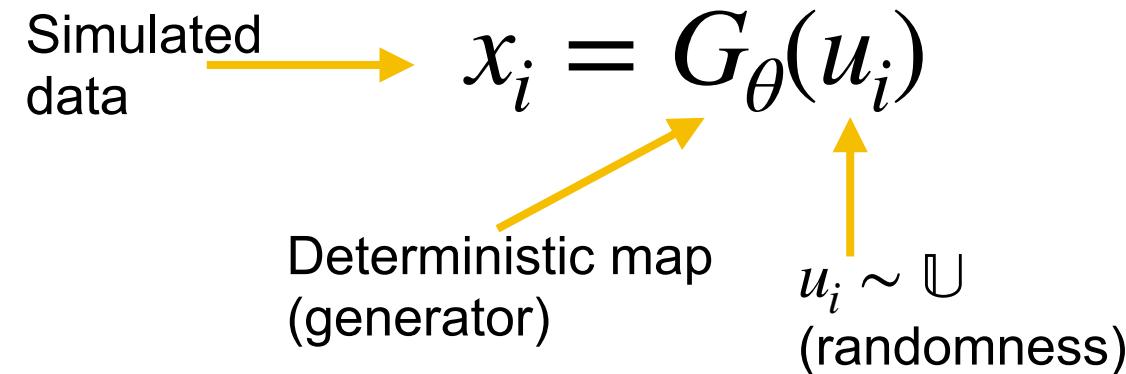
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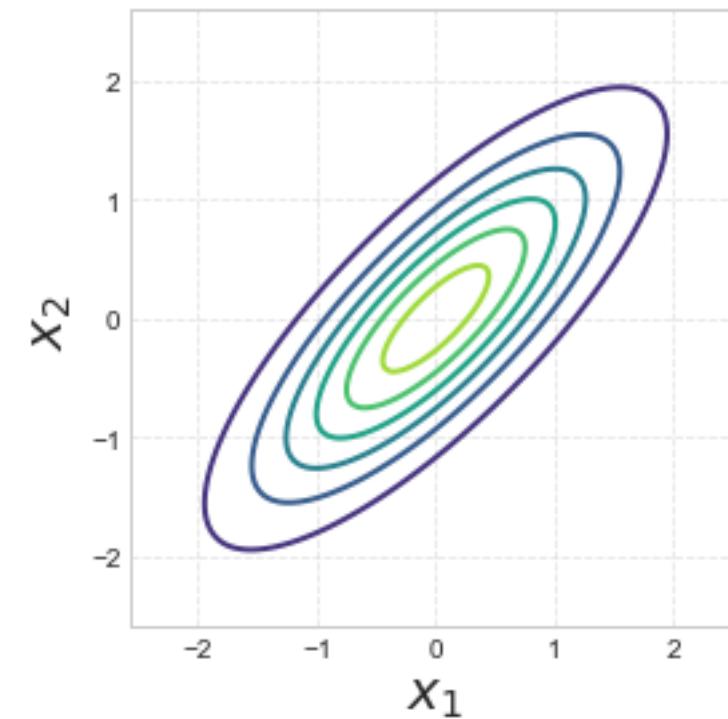
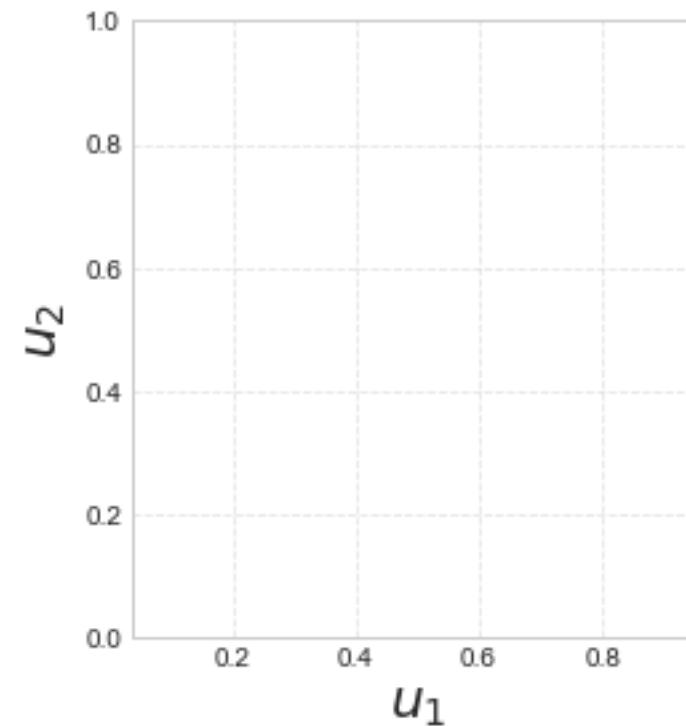
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Simulation-based inference: Inference using simulated data to replace evaluations of the likelihood!

A trivial simulator for Gaussians

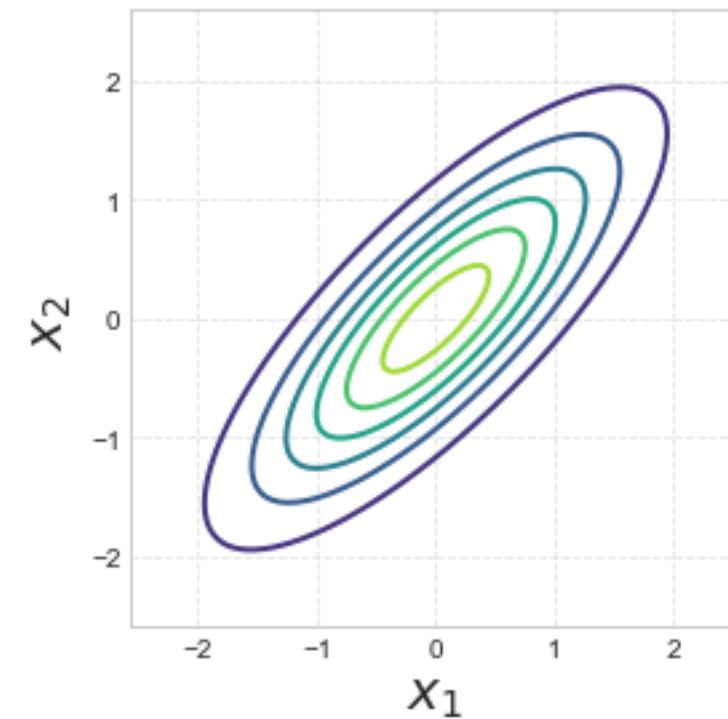
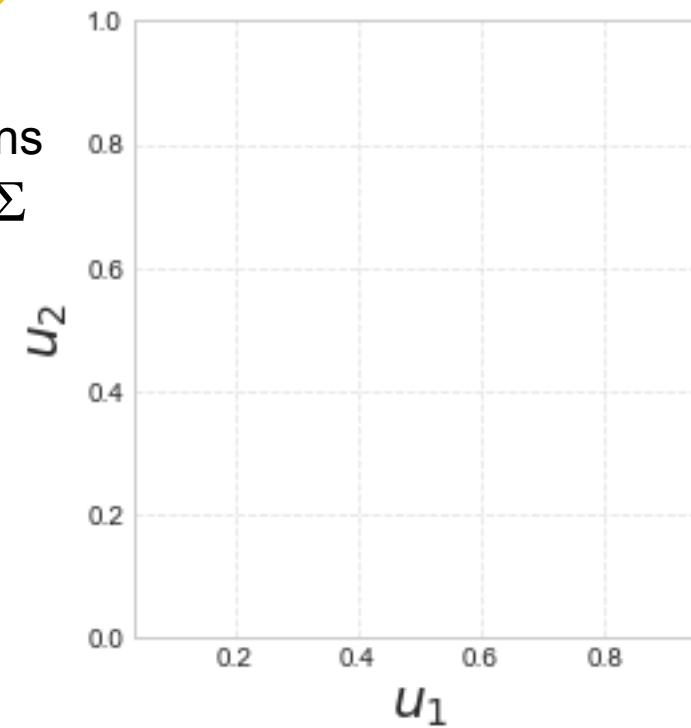
- $\mathbb{P}_\theta := \mathcal{N}(\mu, \Sigma), \quad u_i = (u_{i1}, u_{i2})^\top, \quad u_{i1}, u_{i2} \sim \text{Unif}(0,1) \quad G_\theta(u) = \mu + L \begin{pmatrix} \Phi^{-1}(u_{i1}) \\ \Phi^{-1}(u_{i2}) \end{pmatrix}$



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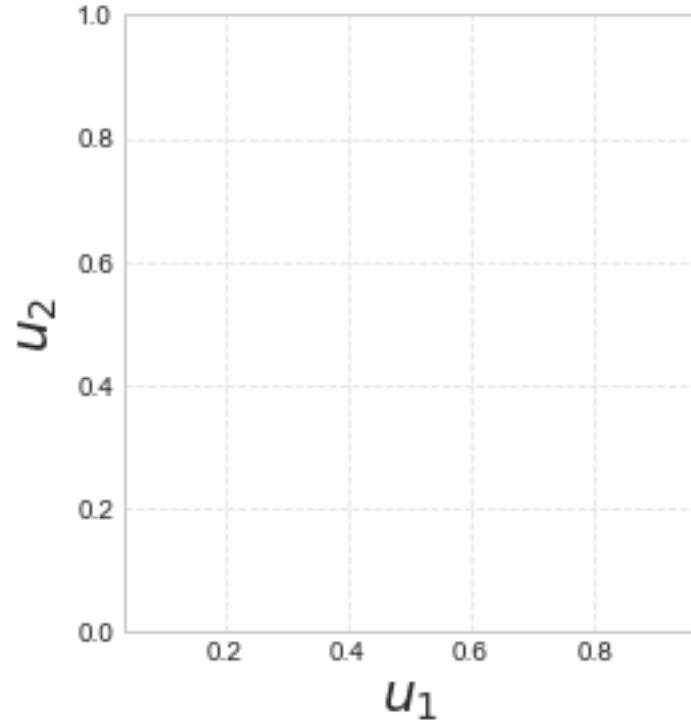
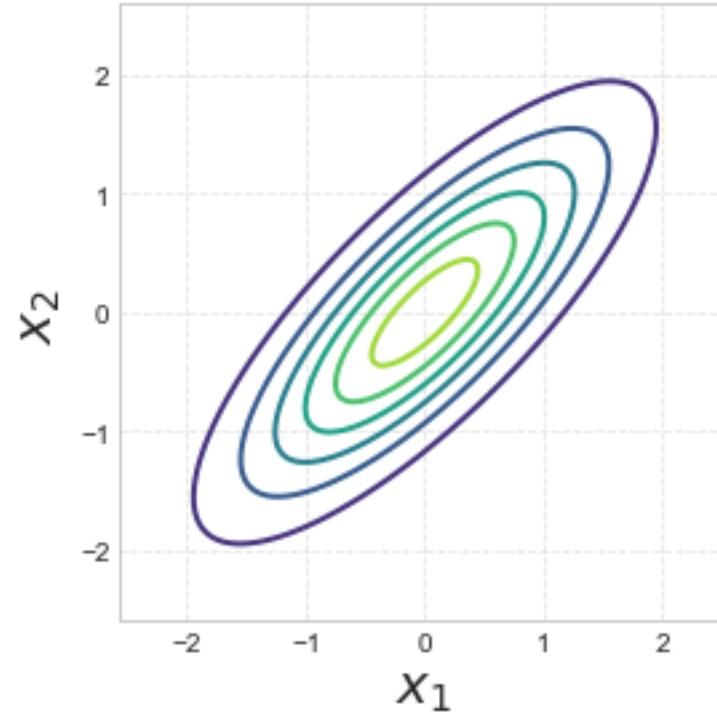
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i.e. θ contains both μ and Σ



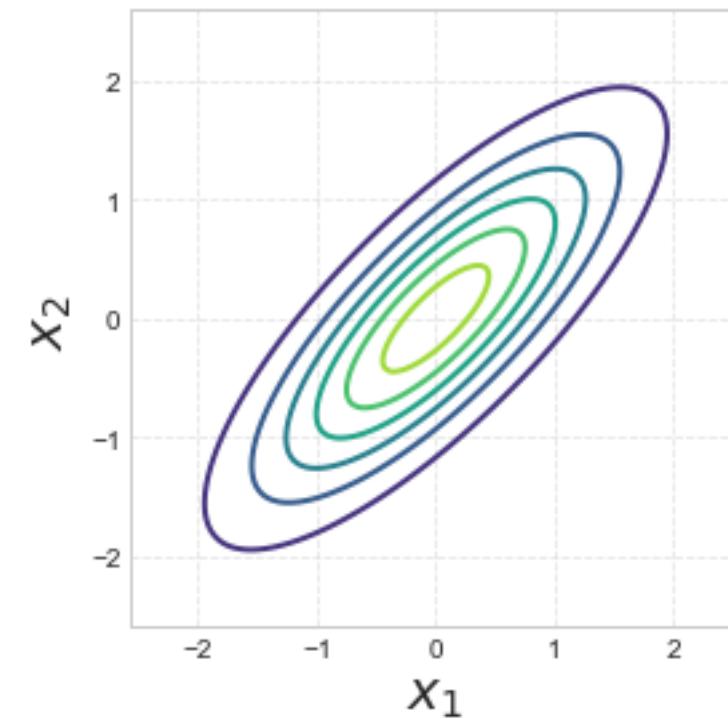
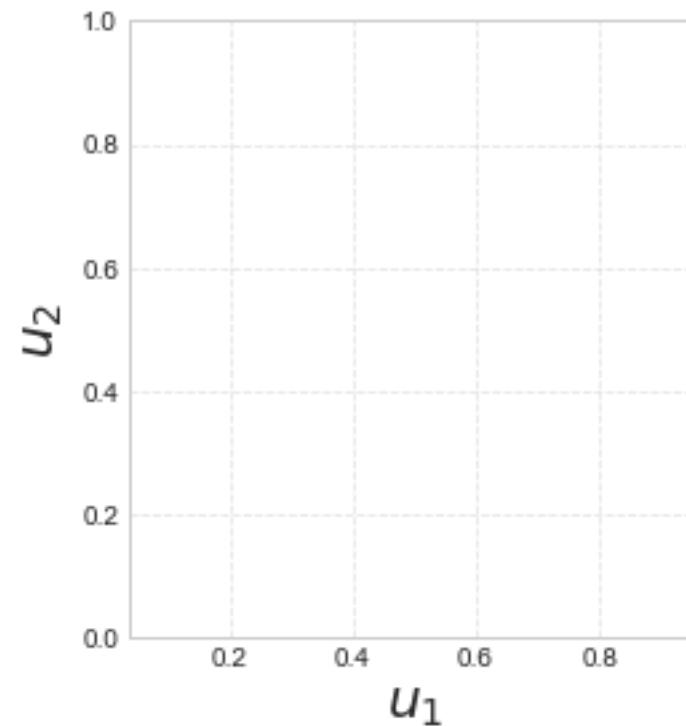
$$\Sigma = LL^\top$$

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- 
- $\Sigma = LL^\top$
- Inverse CDF of standard Gaussian!

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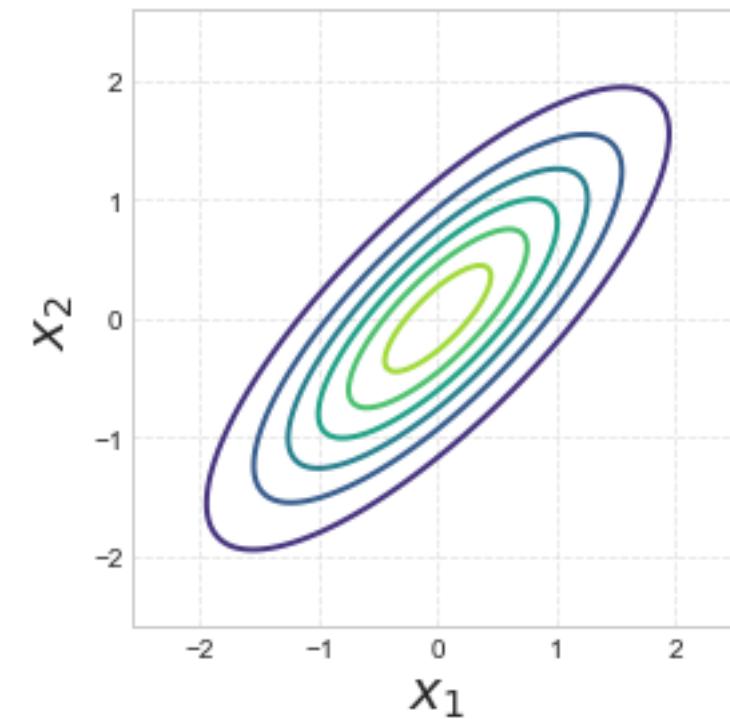
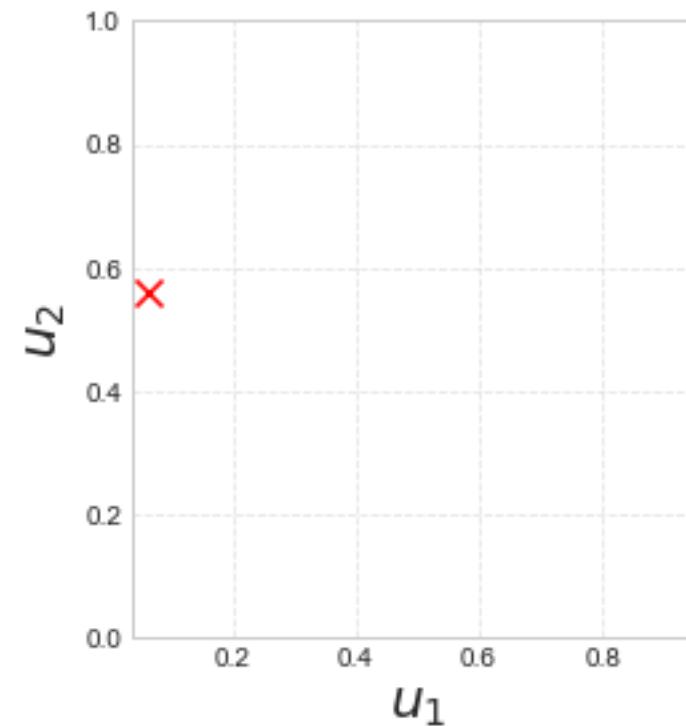


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Cholesky!

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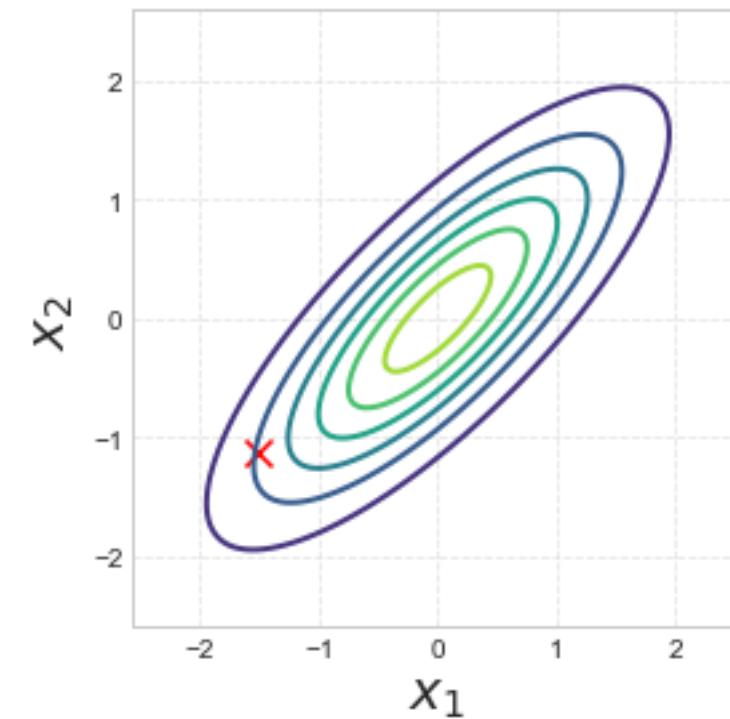
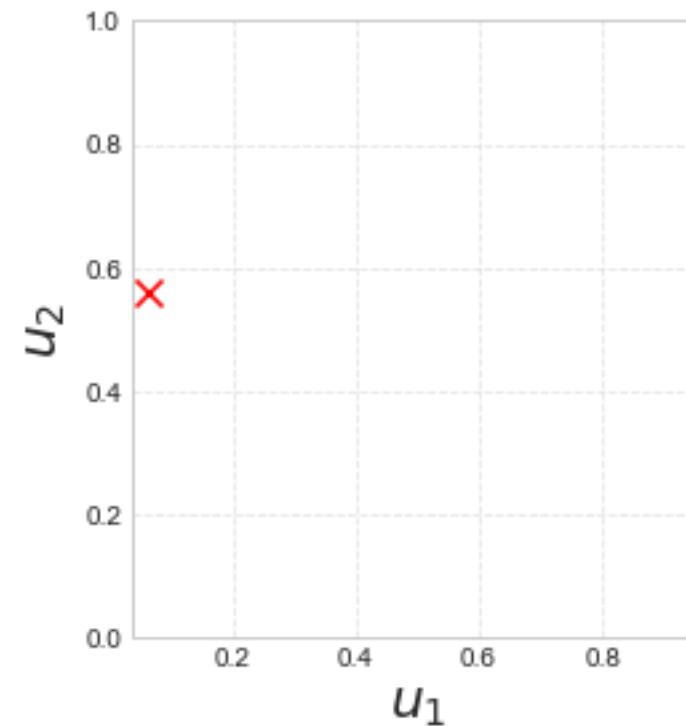
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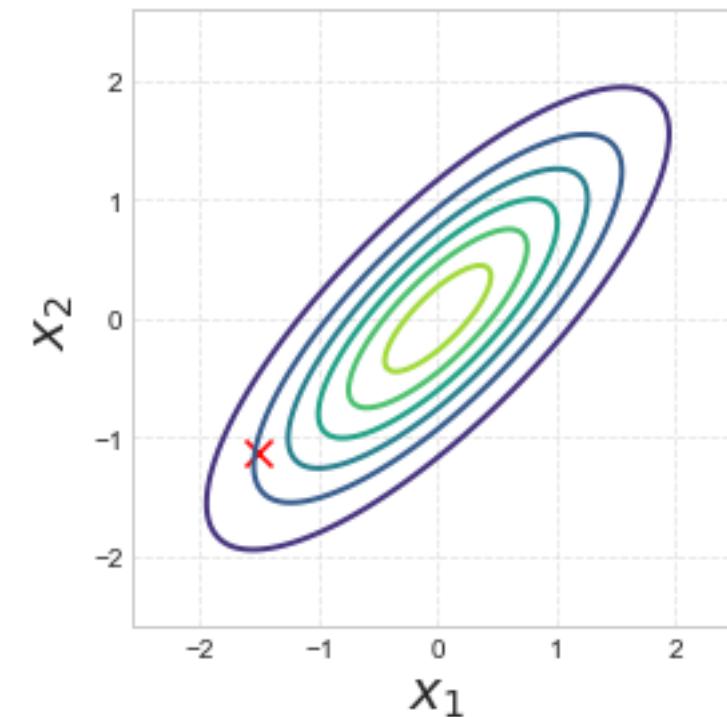
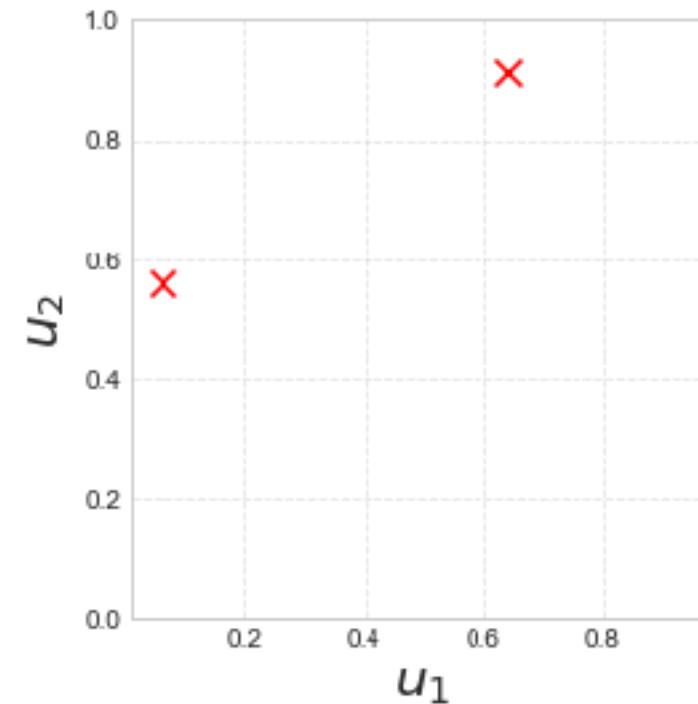
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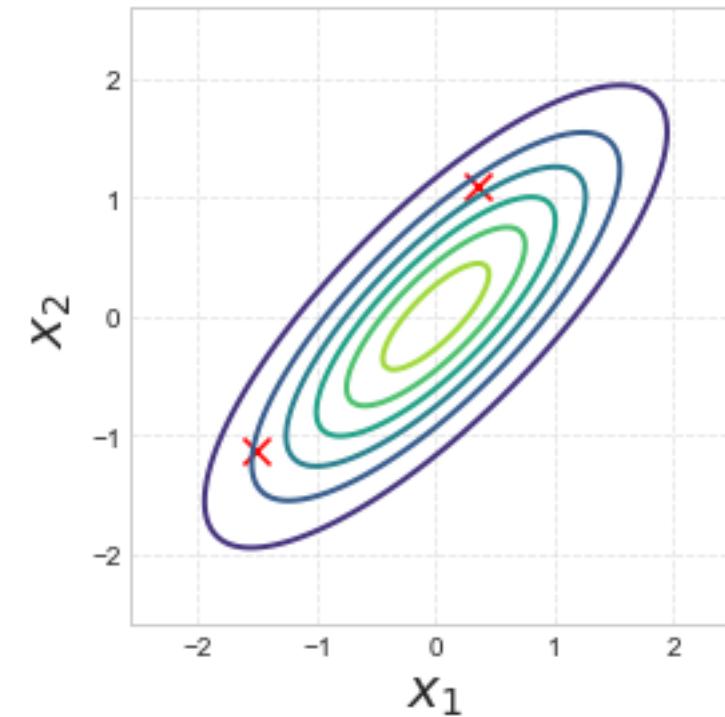
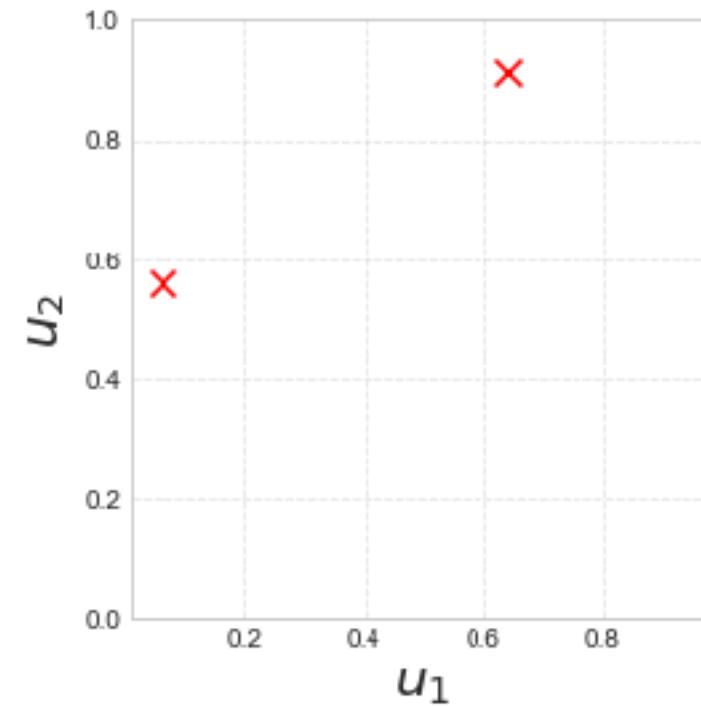
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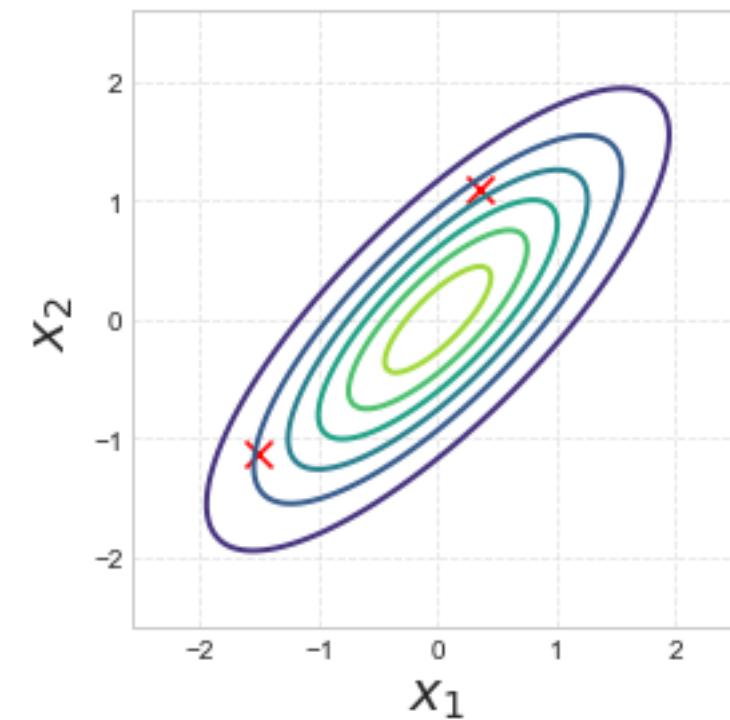
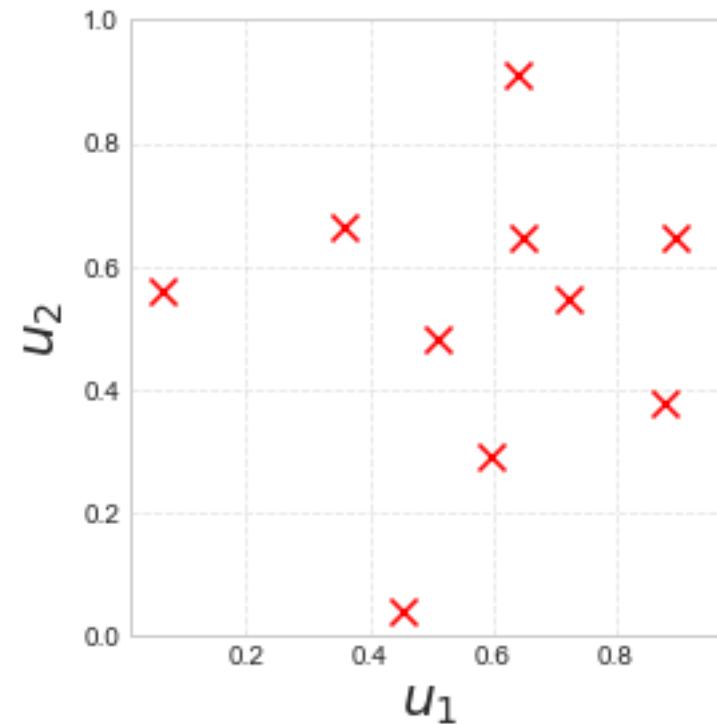
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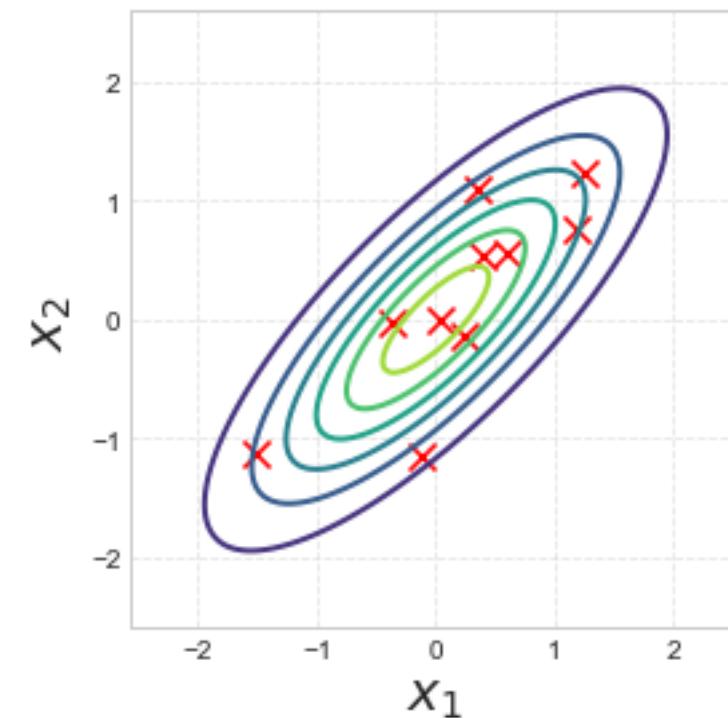
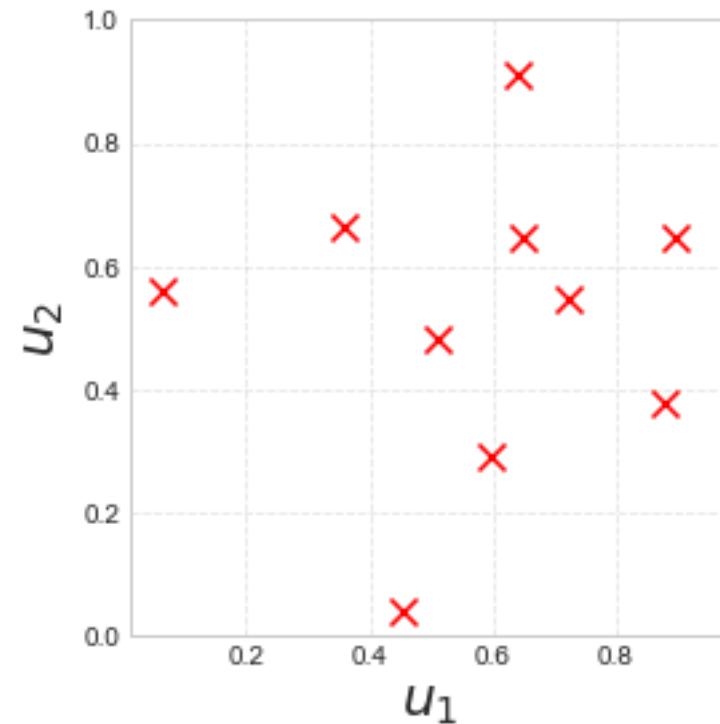
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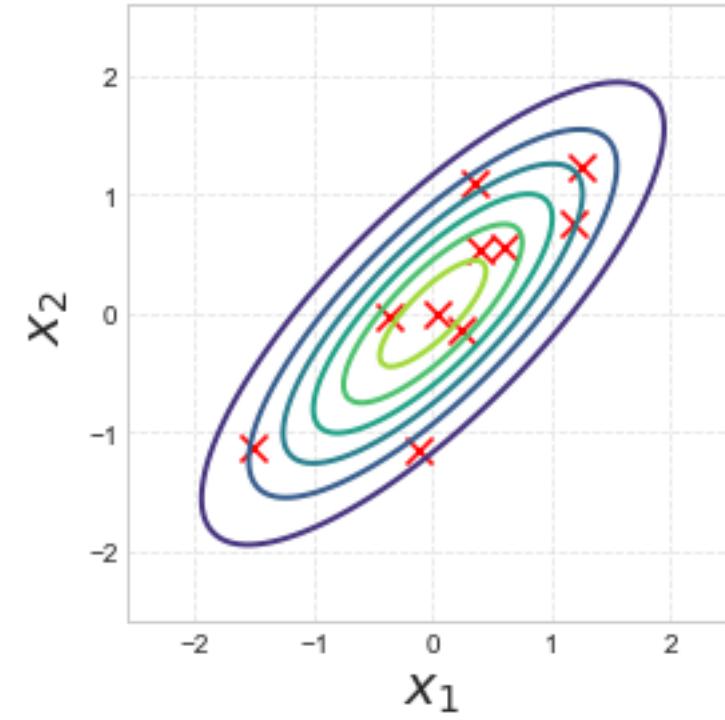
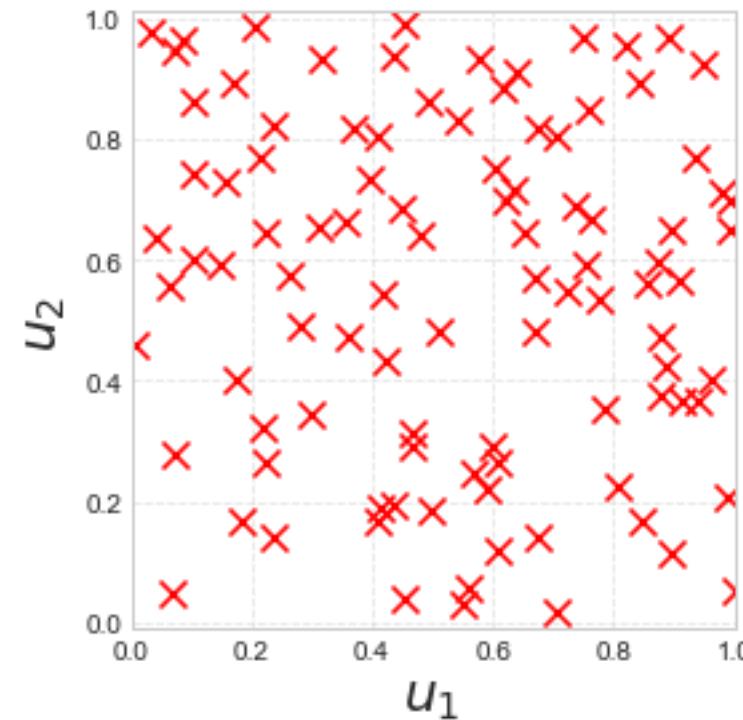
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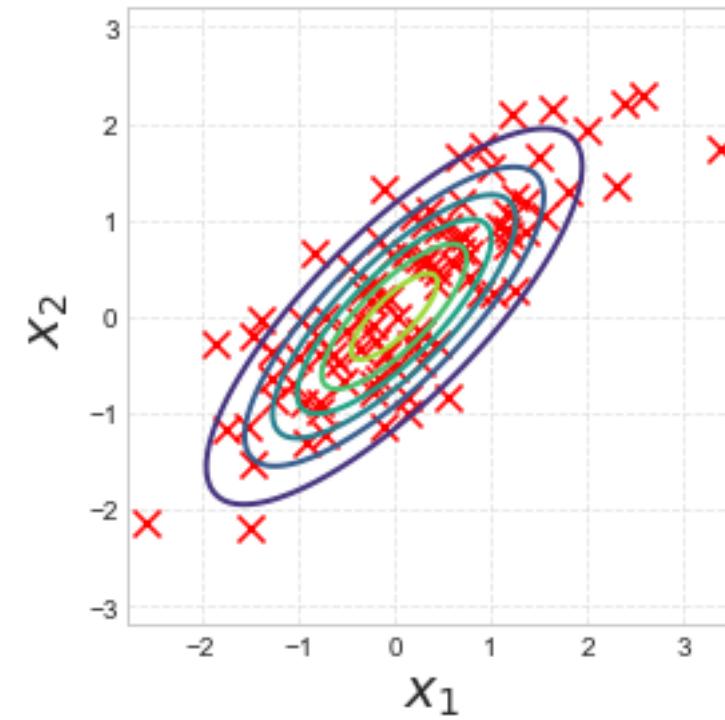
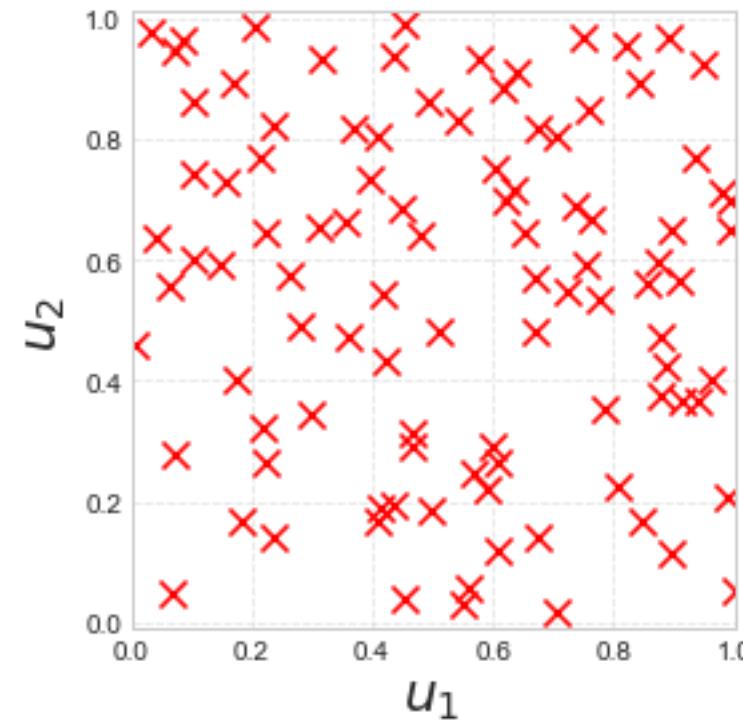
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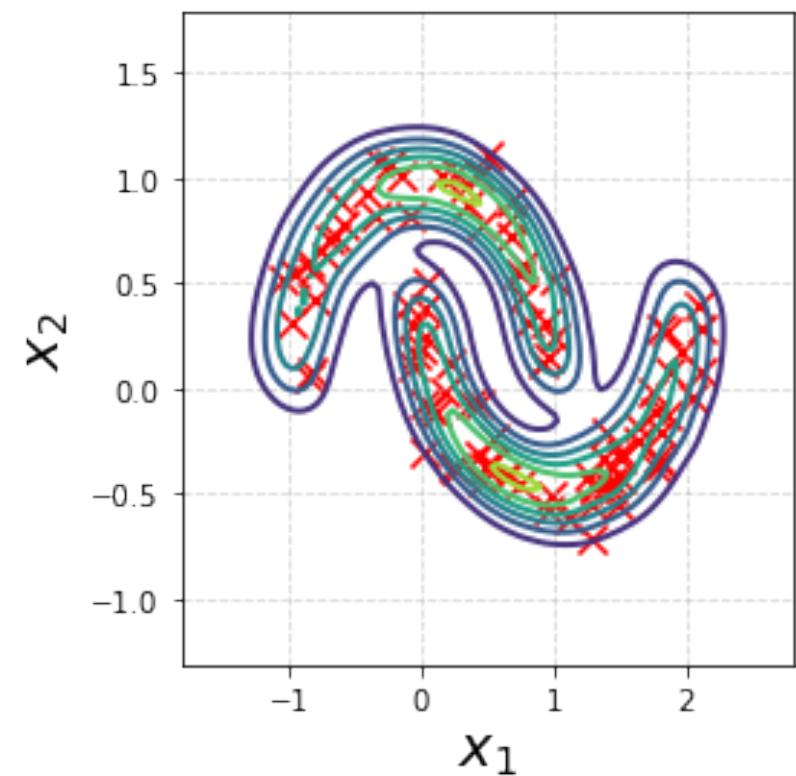
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Some slightly less trivial simulators....

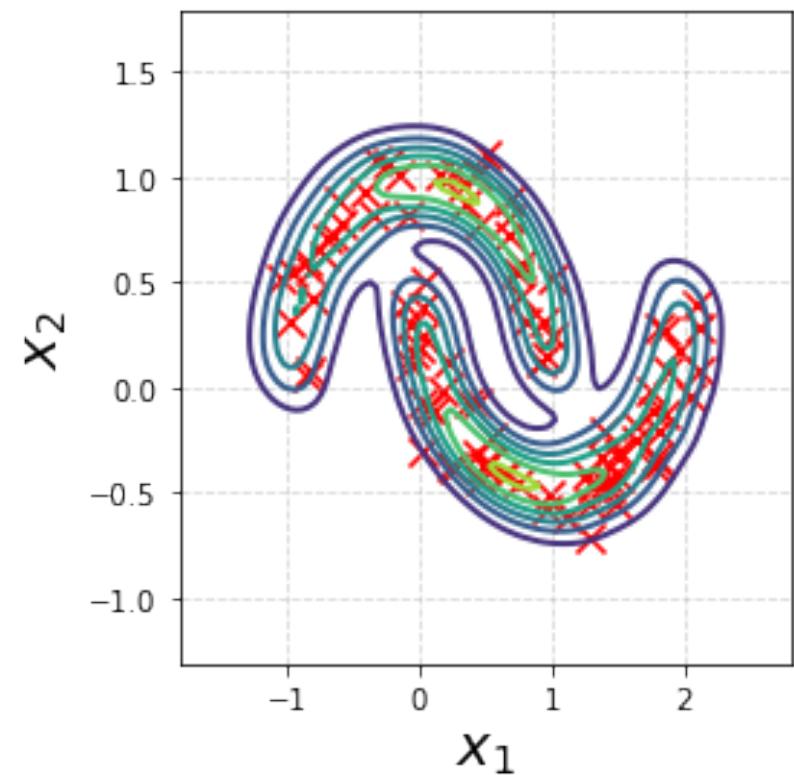
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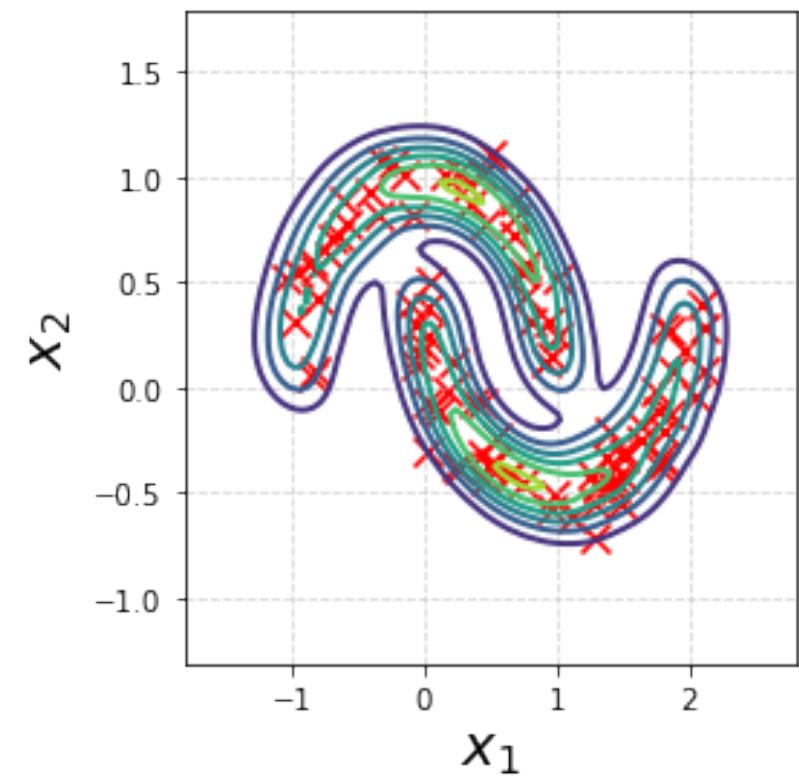


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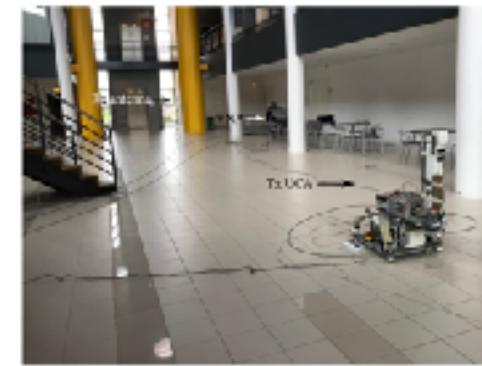
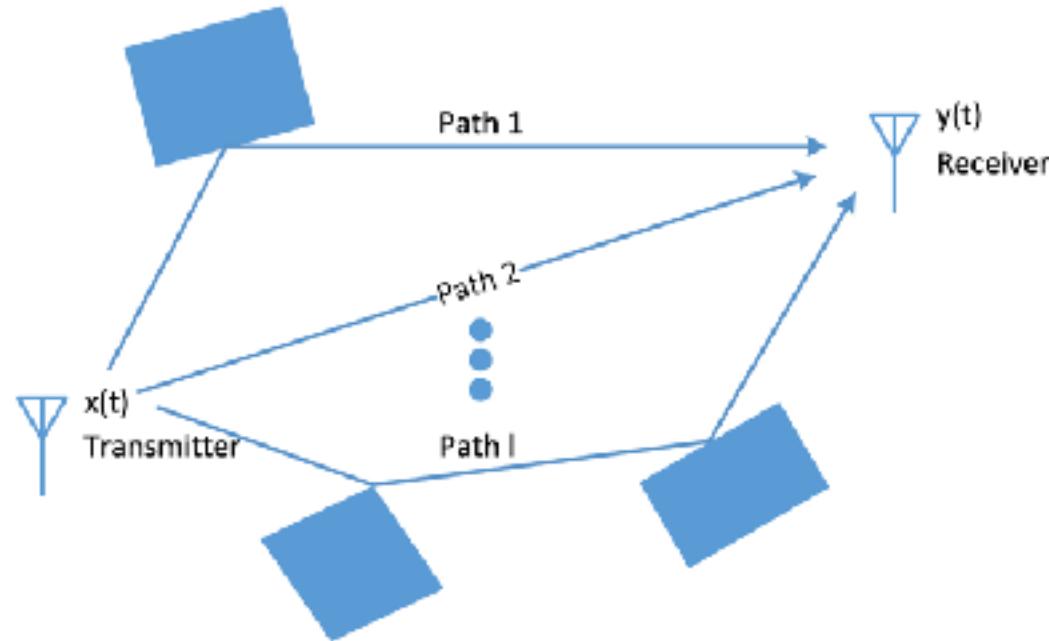
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- SBI often works with simulators **carefully crafted by scientists and engineers**. These simulators are hence implementations of complex mathematical models of the phenomena being studied.



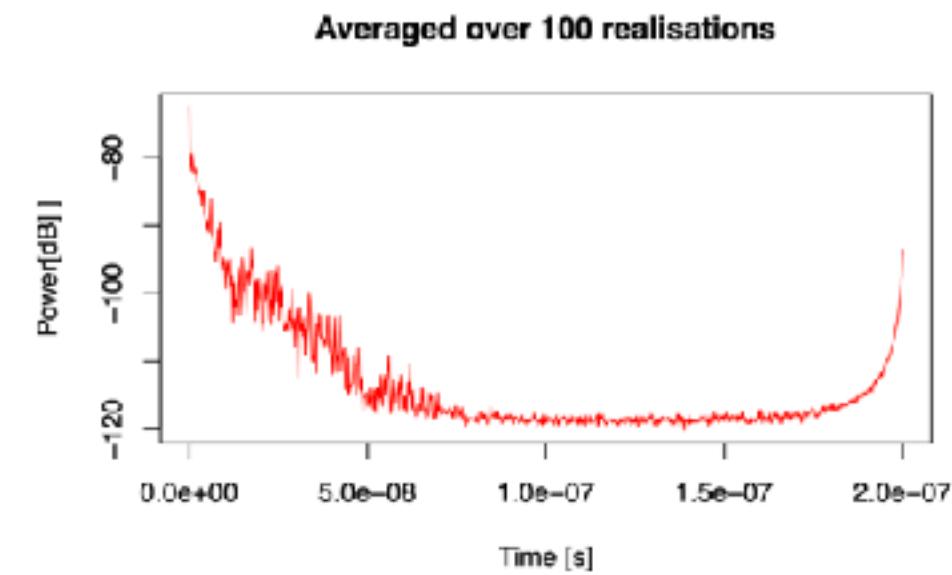
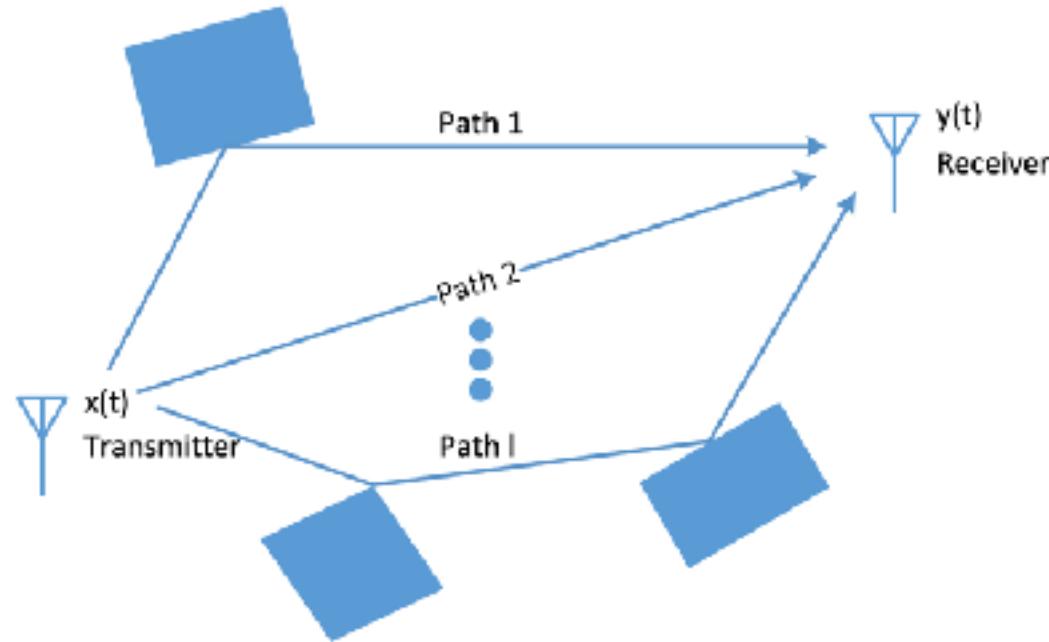
Simulators in telecommunications engineering



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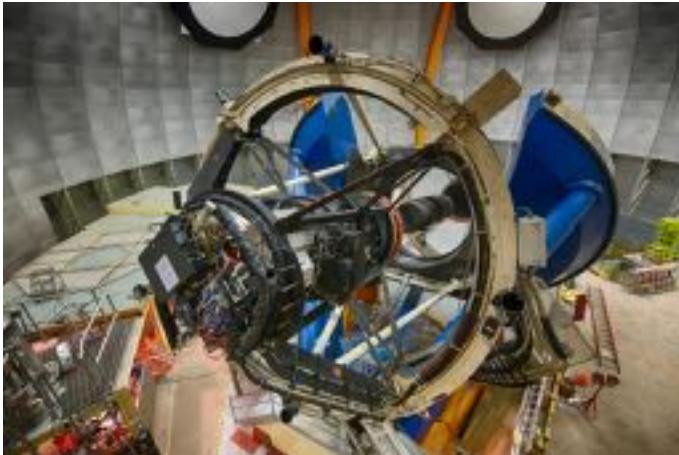
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Simulators in cosmology



(+ ≈ 400 scientists
from 25 institutions
in 7 countries)

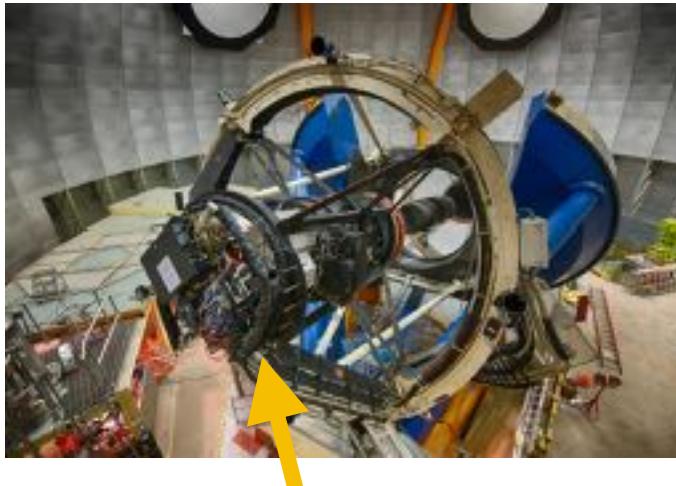


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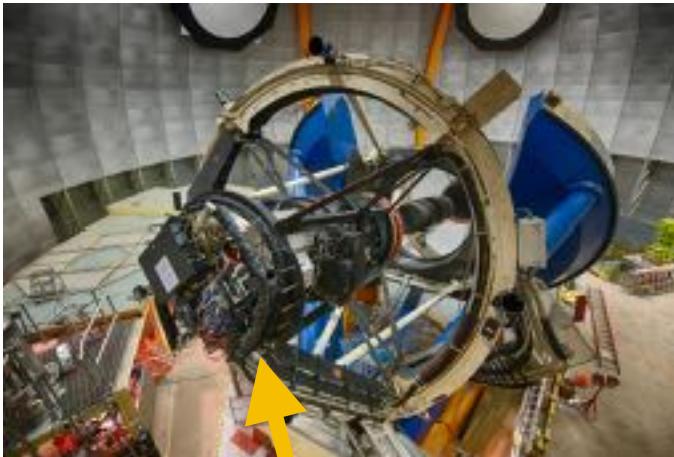
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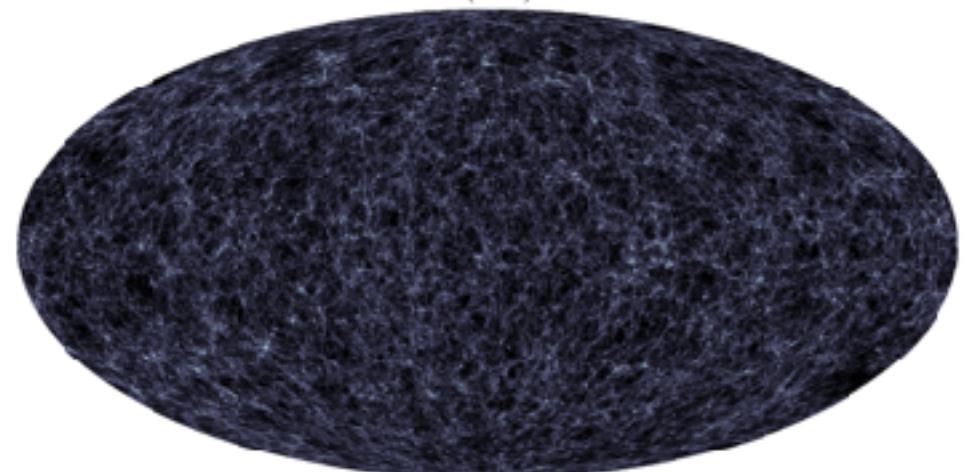
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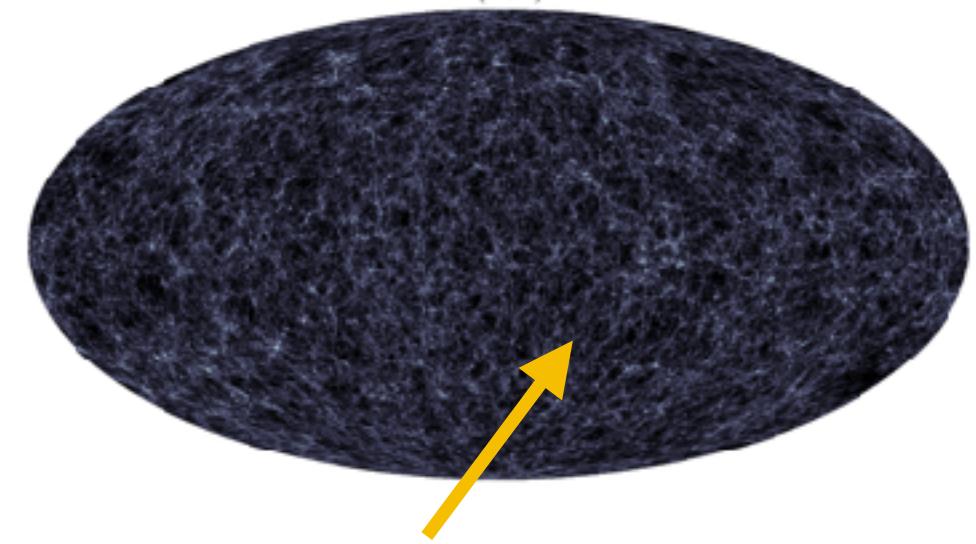
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'Gower Street simulation' run by Niall and colleagues at UCL Physics

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Simulators in the sciences and beyond

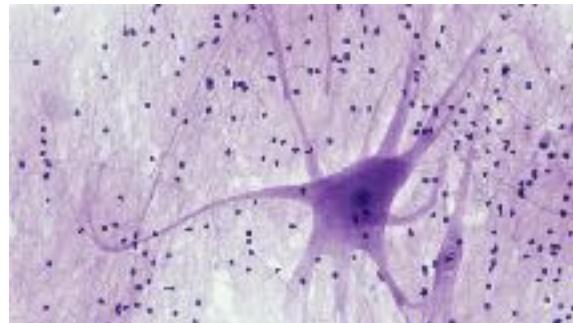


Particle Physics (CERN)

Simulators in the sciences and beyond



Particle Physics (CERN)

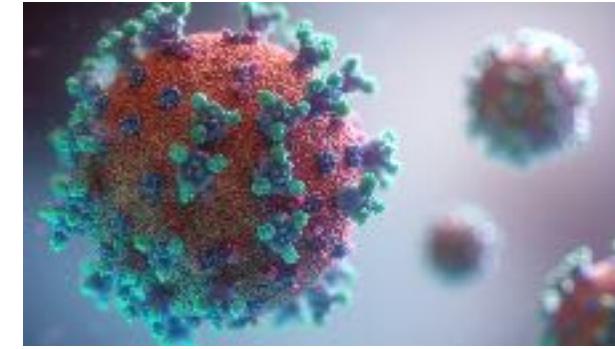


Neuroscience

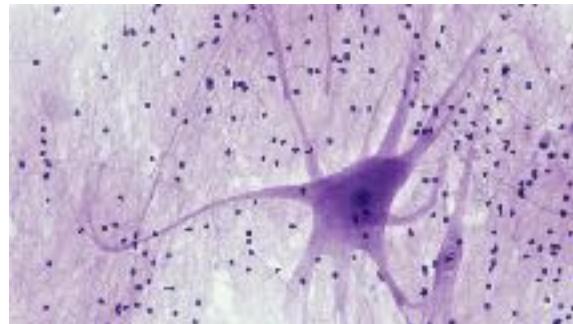
Simulators in the sciences and beyond



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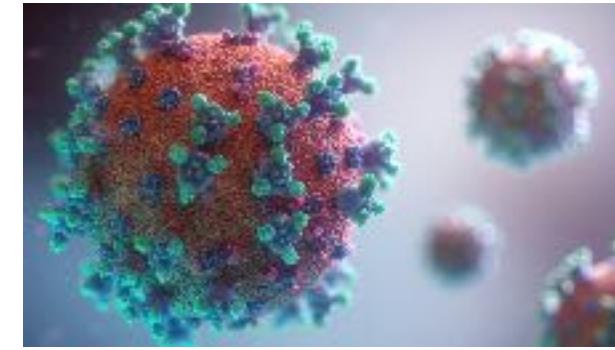


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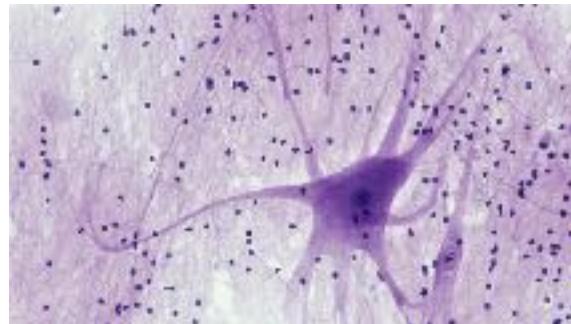
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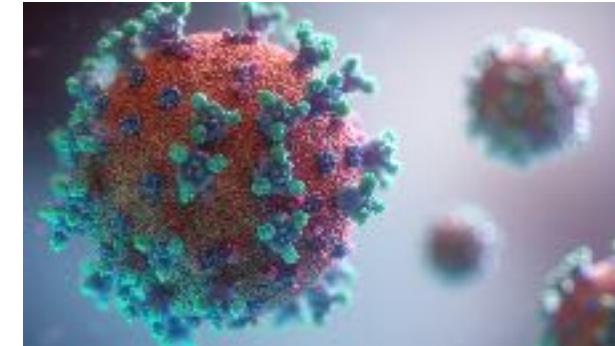


Genomics

Simulators in the sciences and beyond



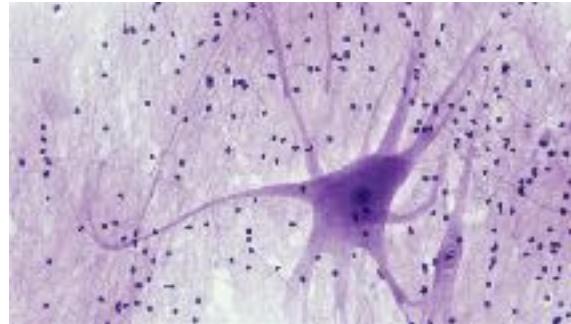
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Epidemiology



Health monitoring (Apple)



Neuroscience

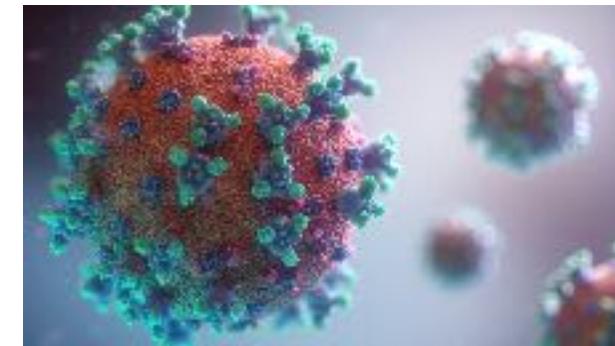


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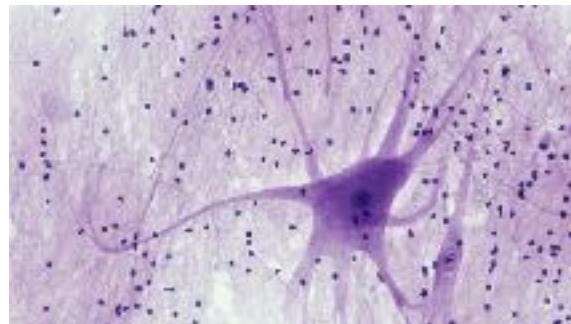
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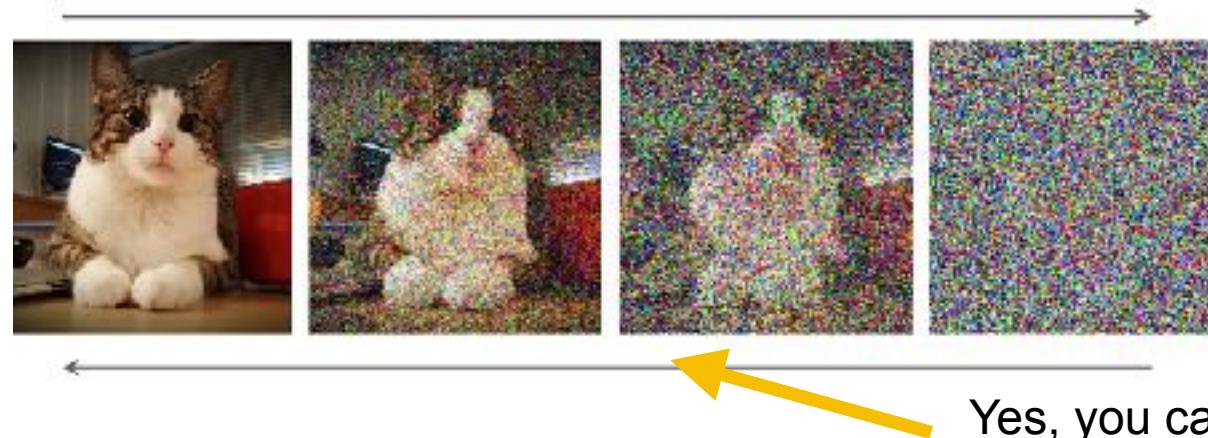
Clarifying terminology

- Are diffusion models simulators?



Clarifying terminology

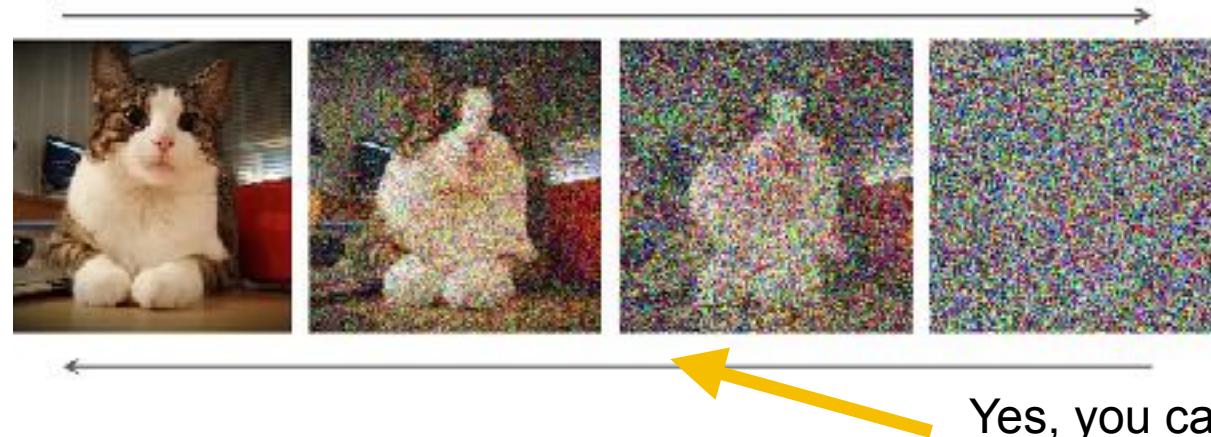
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Yes, you can think of this process as defining a simulator!

Clarifying terminology

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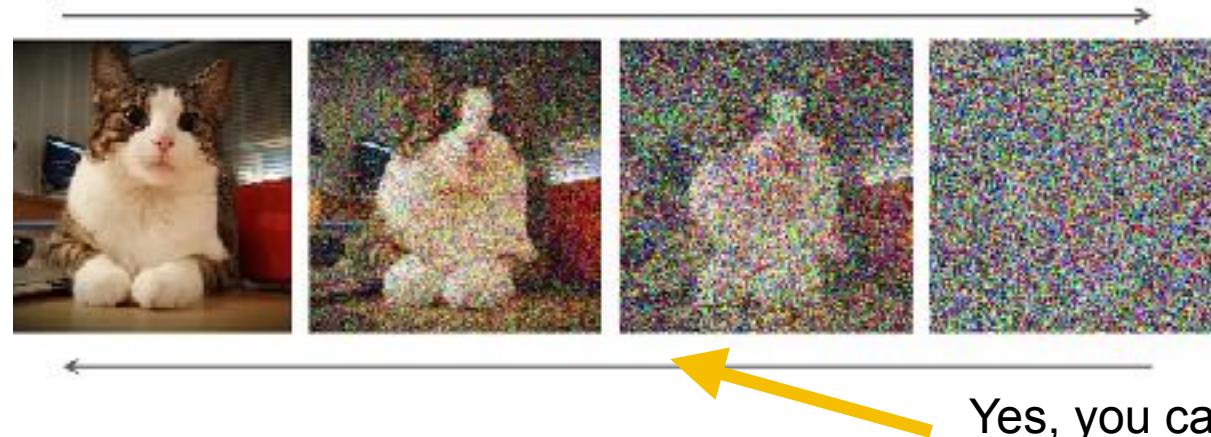


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Clarifying terminology

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No! In SBI, we typically have **scientifically meaningful simulators** where the parameter θ can be interpreted. We therefore really care about estimating it and providing **uncertainty estimates**!

Any Questions?

What is coming up

- Basic methods:

Minimum distance
estimation

Approximate Bayesian
Computation

Neural simulation-
based inference

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- Discussion of the main challenges in SBI.

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- Discussion of the main challenges in SBI.
- Some illustrations of recent advances:

Hikida, Y., Bharti, A., Jeffrey, N. & **Briol, F-X** (2025). Multilevel neural simulation-based inference. arXiv:2506.06087 (to appear at NeurIPS?).

Bharti, A., Huang, D., Kaski, S., & **Briol, F.-X.** (2025). Cost-aware simulation-based inference. International Conference on Artificial Intelligence and Statistics, 28–36.

Dellaporta, C., Knoblauch, J., Damoulas, T. & **Briol, F-X** (2022). Robust Bayesian inference for simulator-based models via the MMD posterior bootstrap. AISTATS, 943-970. Best paper award.

Minimum Distance Estimation



Minimum Distance Estimation



(i.e. how to be a frequentist in SBI...)

The method of moments

- Model: $\{\mathbb{P}_\theta\}_{\theta \in \Theta}$, Data-generating process: \mathbb{Q} , Data: y_1, \dots, y_n

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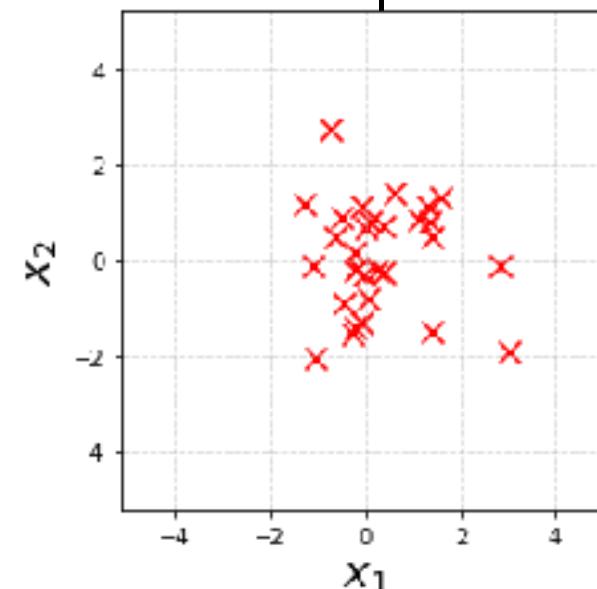
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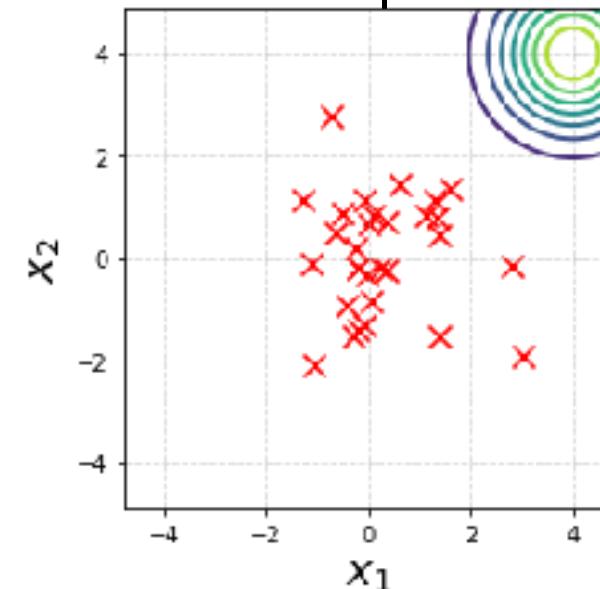
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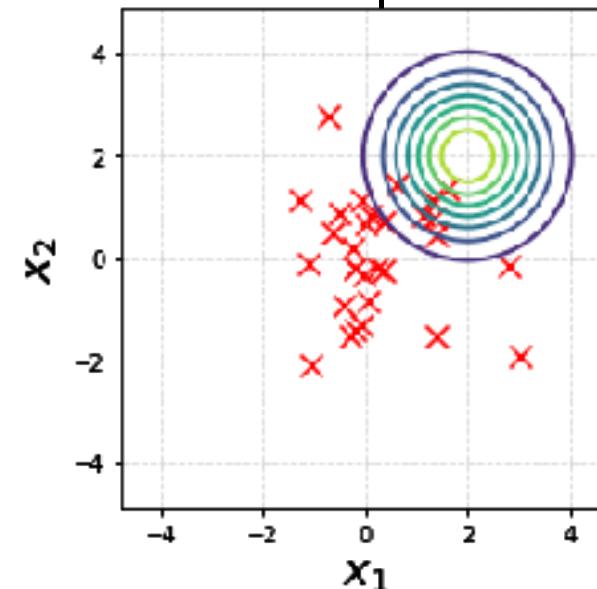
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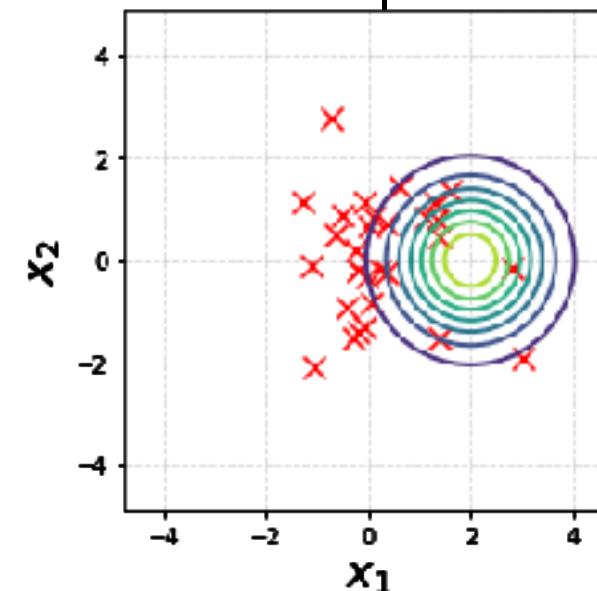
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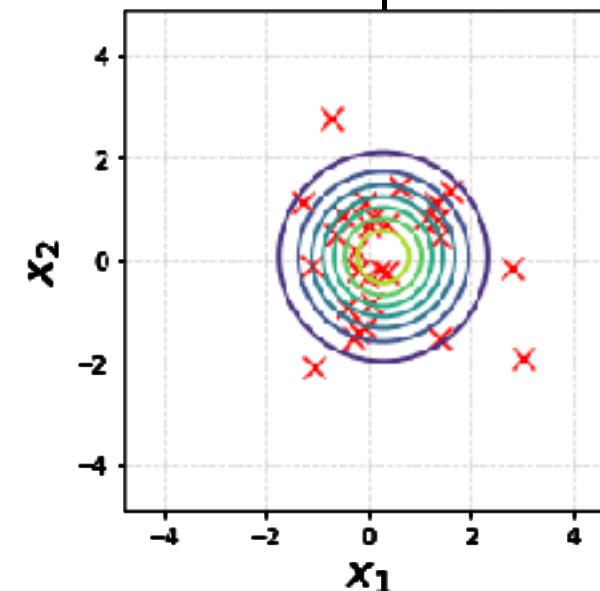
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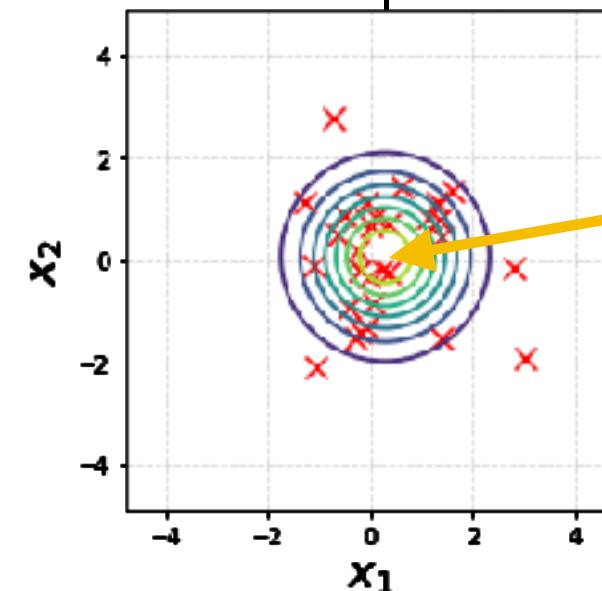
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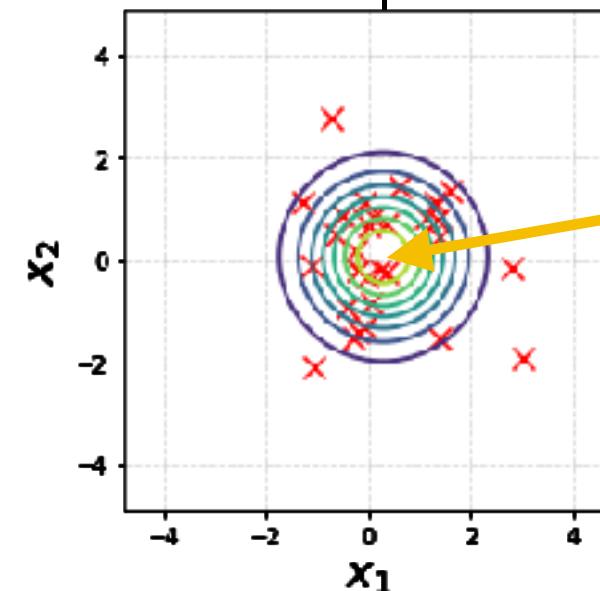
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Note: For more complex models, we may also want to compare higher moments...

The method of simulated moments

- **Problem:** We work with simulators, and so we can't necessarily compute the mean!

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The diagram consists of three question marks (???) arranged in a triangle. One arrow points from the top-left question mark to the expectation operator (\mathbb{E}). Another arrow points from the bottom-right question mark to the sample mean formula ($\frac{1}{n} \sum_{i=1}^n y_i$). A third arrow points from the bottom-left question mark to the distribution \mathbb{P}_θ .

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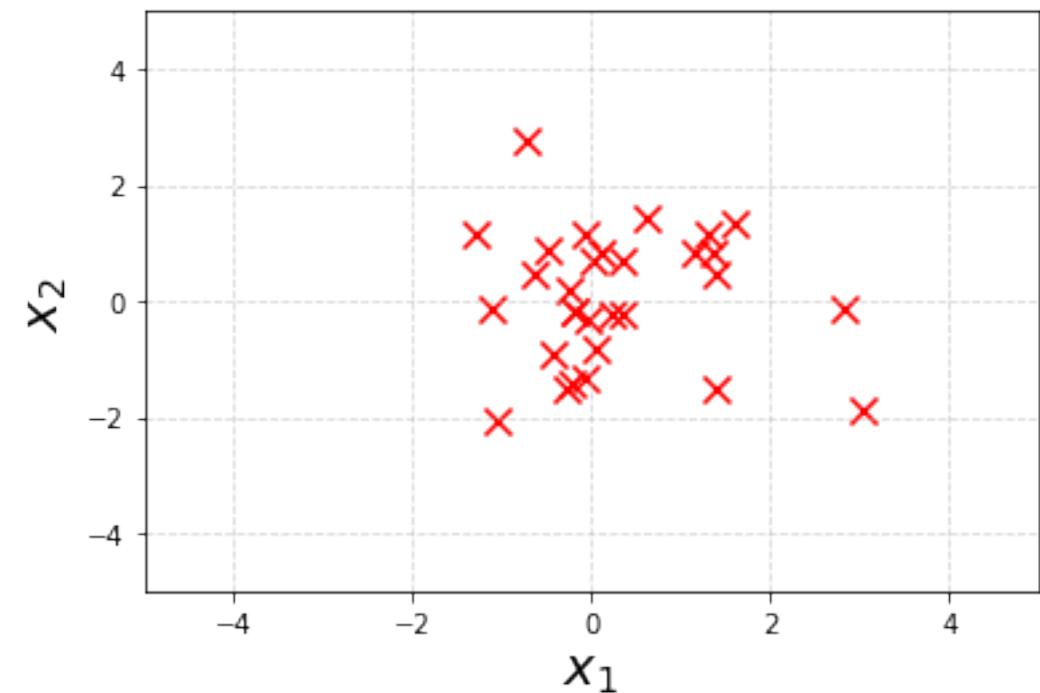
The diagram consists of three question marks (???) positioned above and to the left of the mathematical expression. Three yellow arrows point from these question marks to specific parts of the expression: one arrow points to the expectation operator \mathbb{E} , another points to the distribution \mathbb{P}_θ , and the third points to the sample mean formula $\frac{1}{n} \sum_{i=1}^n y_i$.

- **Method of simulated moments:** We repeat the method of moments, but we simulate at each iteration!

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The method of simulated moments

Fix grid $\theta_1, \dots, \theta_T \in \Theta$.

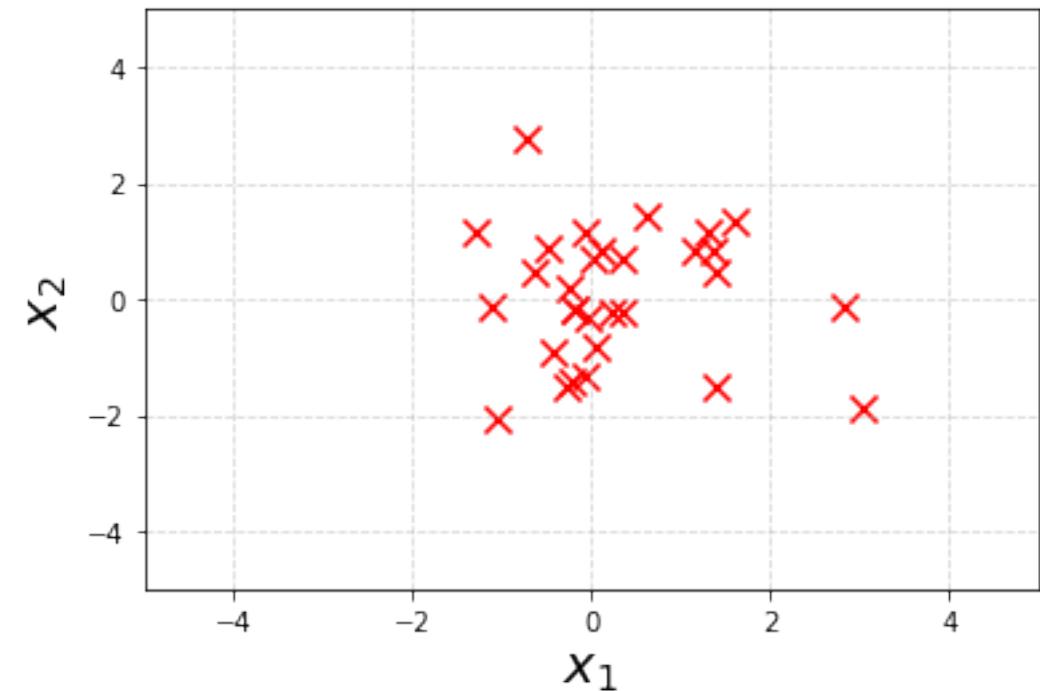


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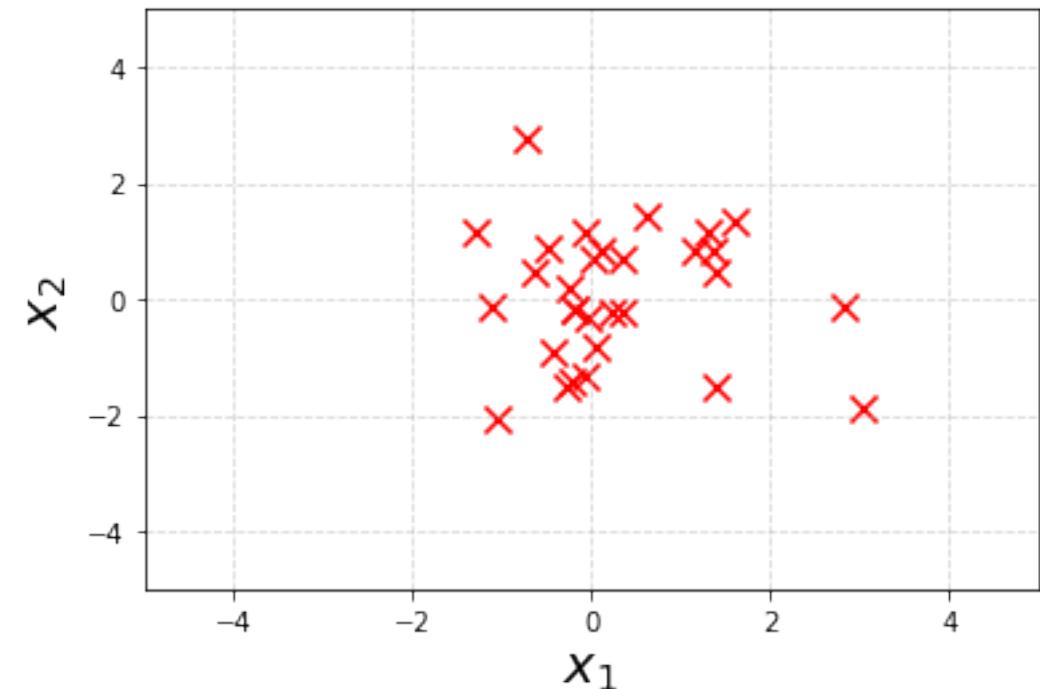
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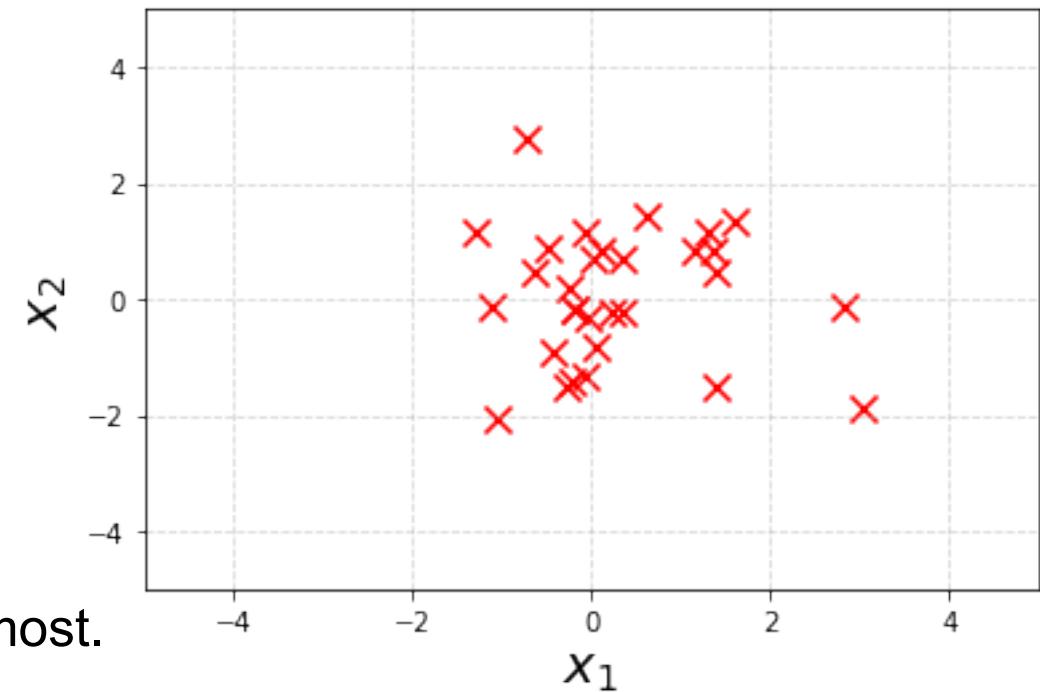
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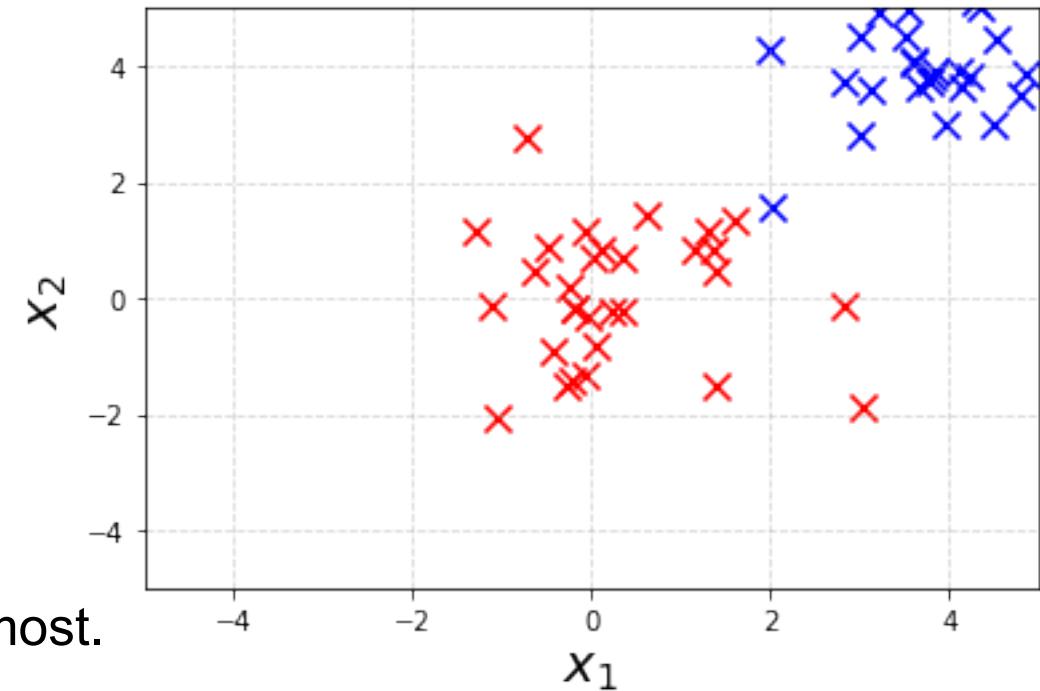
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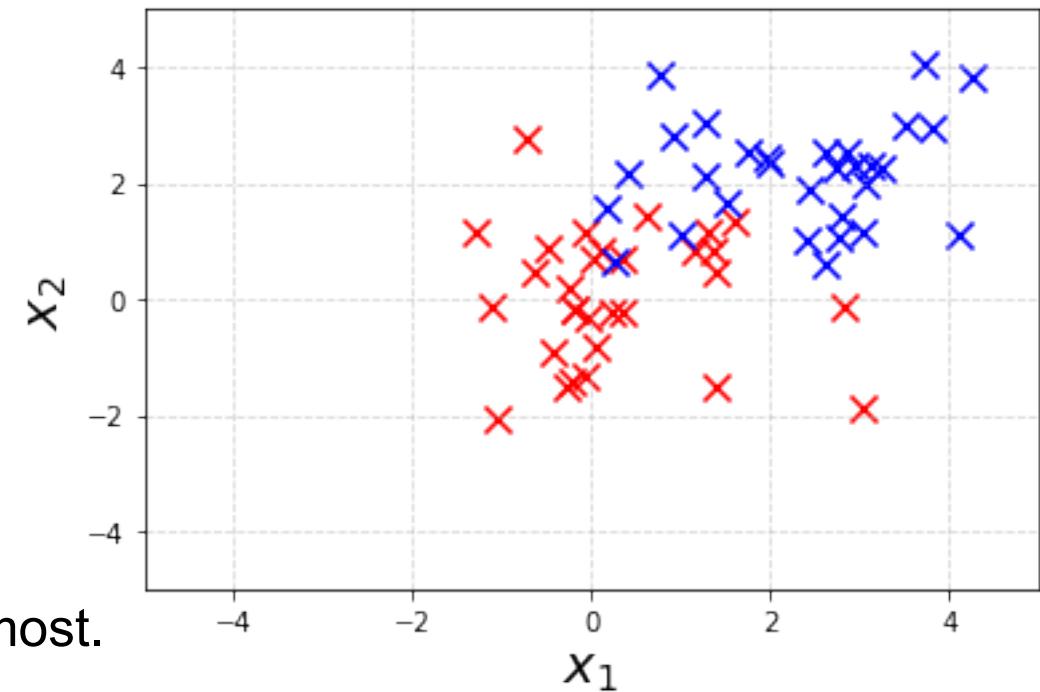
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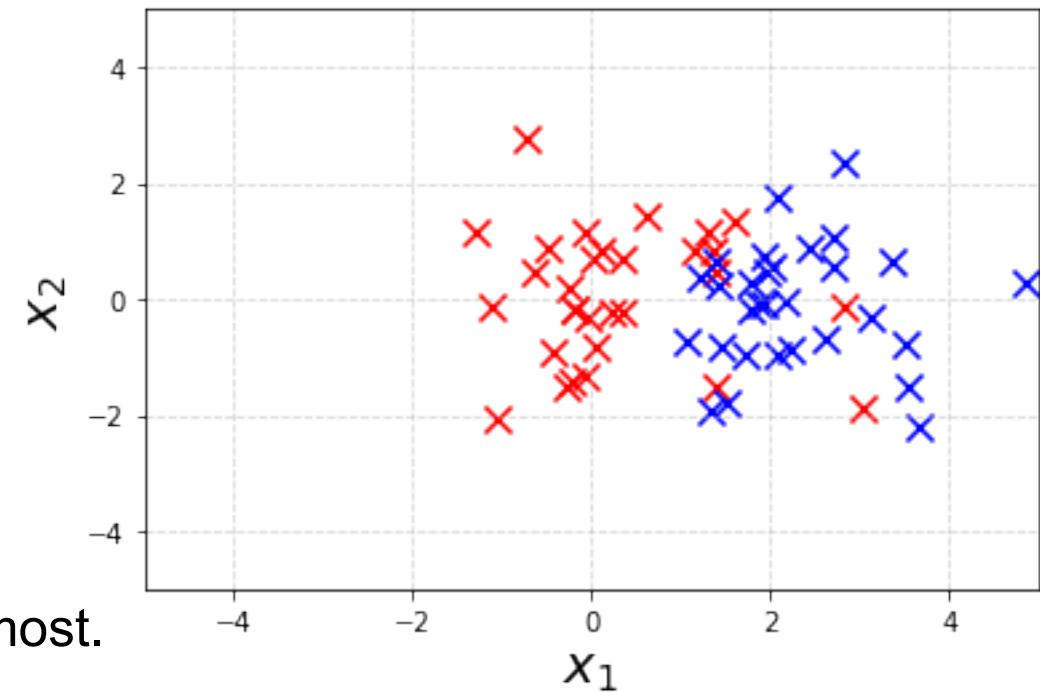
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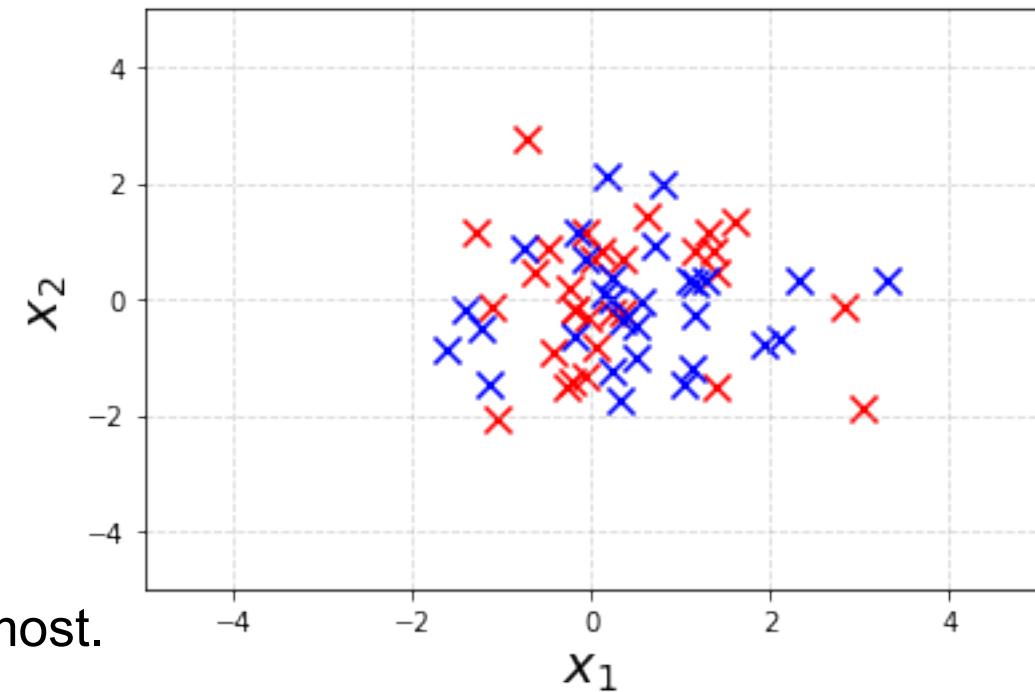
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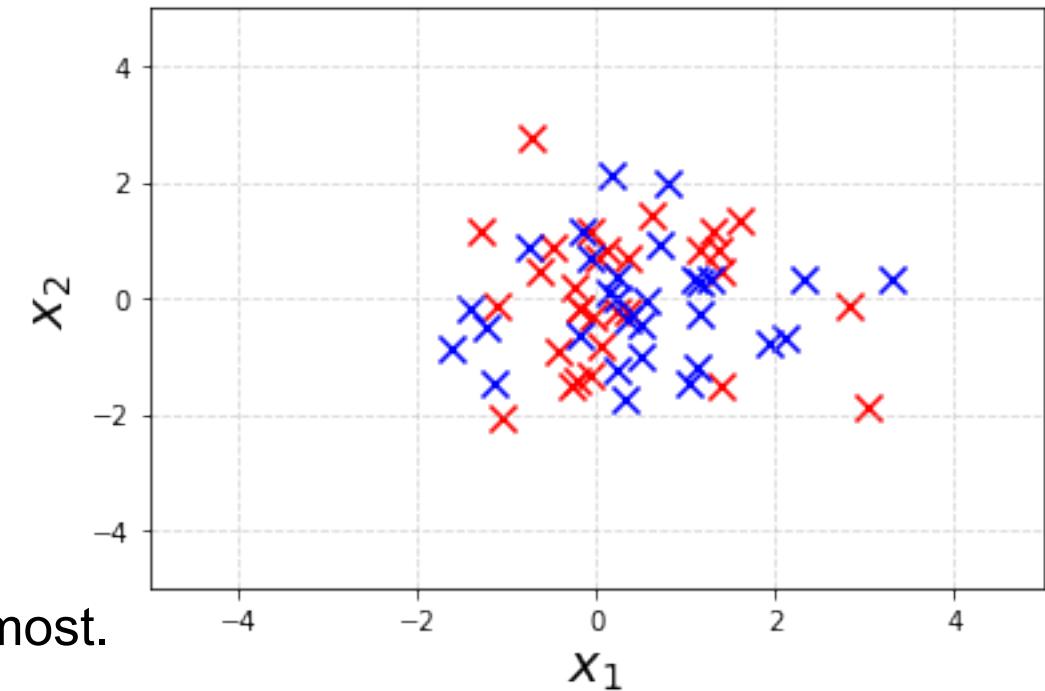
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- In practice, this is implemented much more efficiently than by grid search...

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- Can pick our favourite discrepancy/divergence/distance!

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This will typically make things intractable unless \mathcal{F} is picked carefully!

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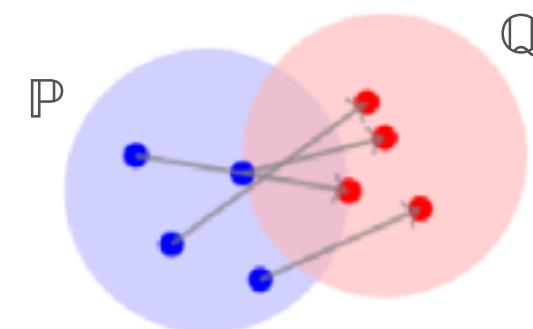
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Credit for figure:



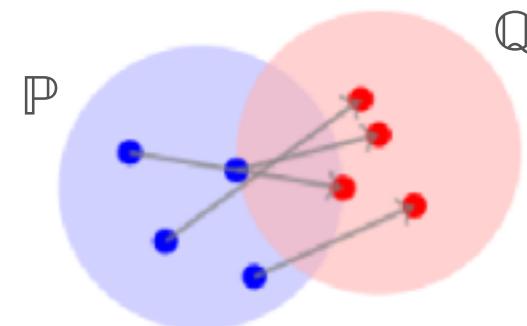
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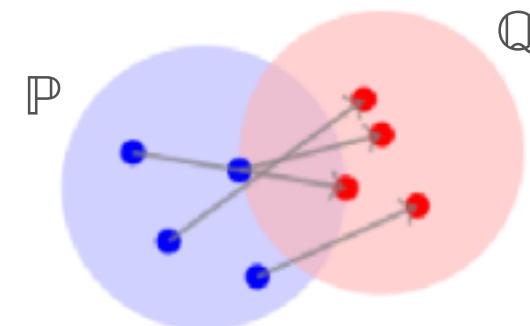
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The Wasserstein distance

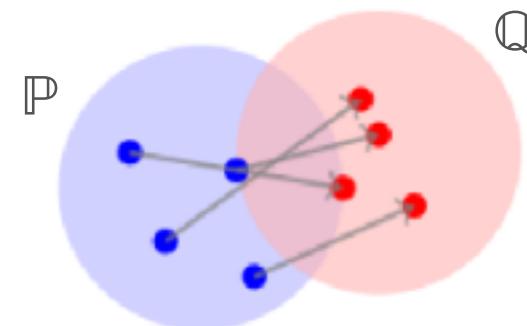
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(3) Interpretable ✓



Credit for figure:



The Wasserstein distance

$$W(\mathbb{P}, \mathbb{Q}) := \sup_{f \in \mathcal{F}_W} \left| \mathbb{E}_{X \sim \mathbb{P}}[f(X)] - \mathbb{E}_{X \sim \mathbb{Q}}[f(X)] \right|$$

$$\mathcal{F}_W := \{f: \mathcal{X} \rightarrow \mathbb{R} : |f(x) - f(y)| \leq \|x - y\|\}$$

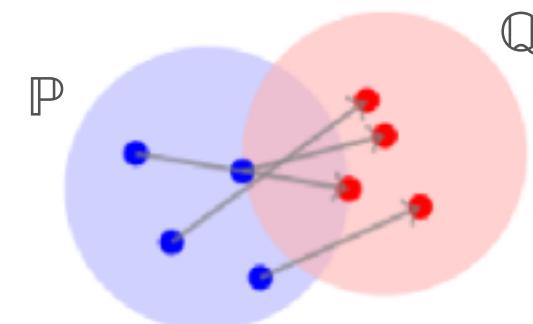
- Well-known interpretation as the **cost of moving mass** from \mathbb{P} to \mathbb{Q} !

(1) Divergence ✓

(2) Easy to estimate ~

(3) Interpretable ✓

(4) Robust ✗



Credit for figure:



Minimum Wasserstein estimators

- Once we have n samples from \mathbb{P}_θ and \mathbb{Q} , this turns into an optimal transport which can be solved in $O(n \log n)$ in $d = 1$ and $O(n^3)$ for $d > 1$.

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- This leads to the following estimator, usually approximated with stochastic optimisation:

$$\hat{\theta}_n := \arg \min_{\theta \in \Theta} W(\mathbb{P}_\theta, Q_n)$$

Bassetti, F., Bodini, A., & Regazzini, E. (2006). On minimum Kantorovich distance estimators. *Statistics & Probability Letters*, 76, 1298–1302.

Bernton, E., Jacob, P. E., Gerber, M., & Robert, C. P. (2017). Inference in generative models using the Wasserstein distance. *Information and Inference: A Journal of the IMA*, 8(4), 657–676.

The maximum mean discrepancy

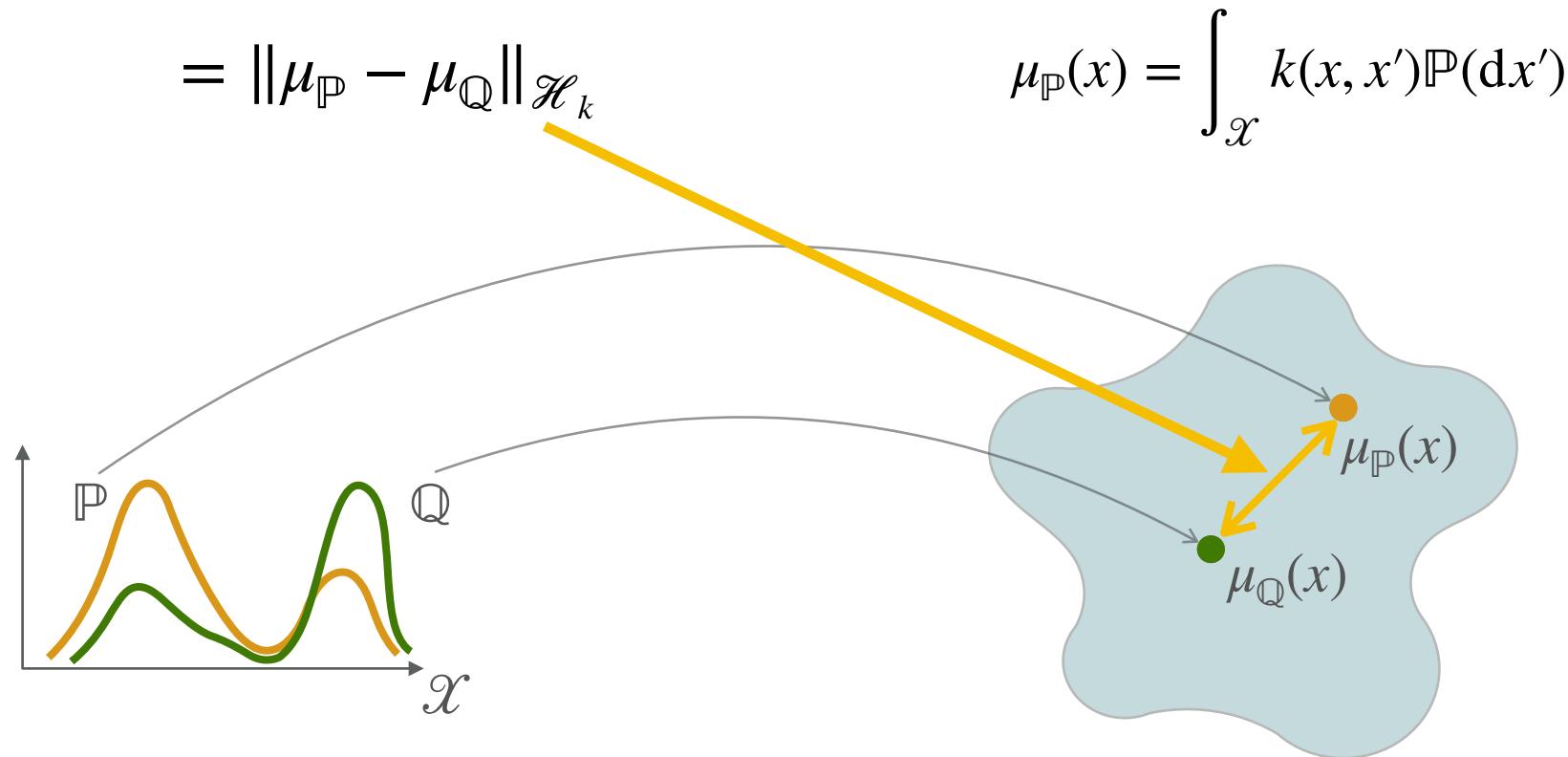
$$\text{MMD}(\mathbb{P}, \mathbb{Q}) := \sup_{f \in \mathcal{F}_{\text{MMD}}} \left| \mathbb{E}_{X \sim \mathbb{P}}[f(X)] - \mathbb{E}_{X \sim \mathbb{Q}}[f(X)] \right| \quad \mathcal{F}_{\text{MMD}} := \{f: \mathcal{X} \rightarrow \mathbb{R} : \|f\|_{\mathcal{H}_k} \leq 1\}$$

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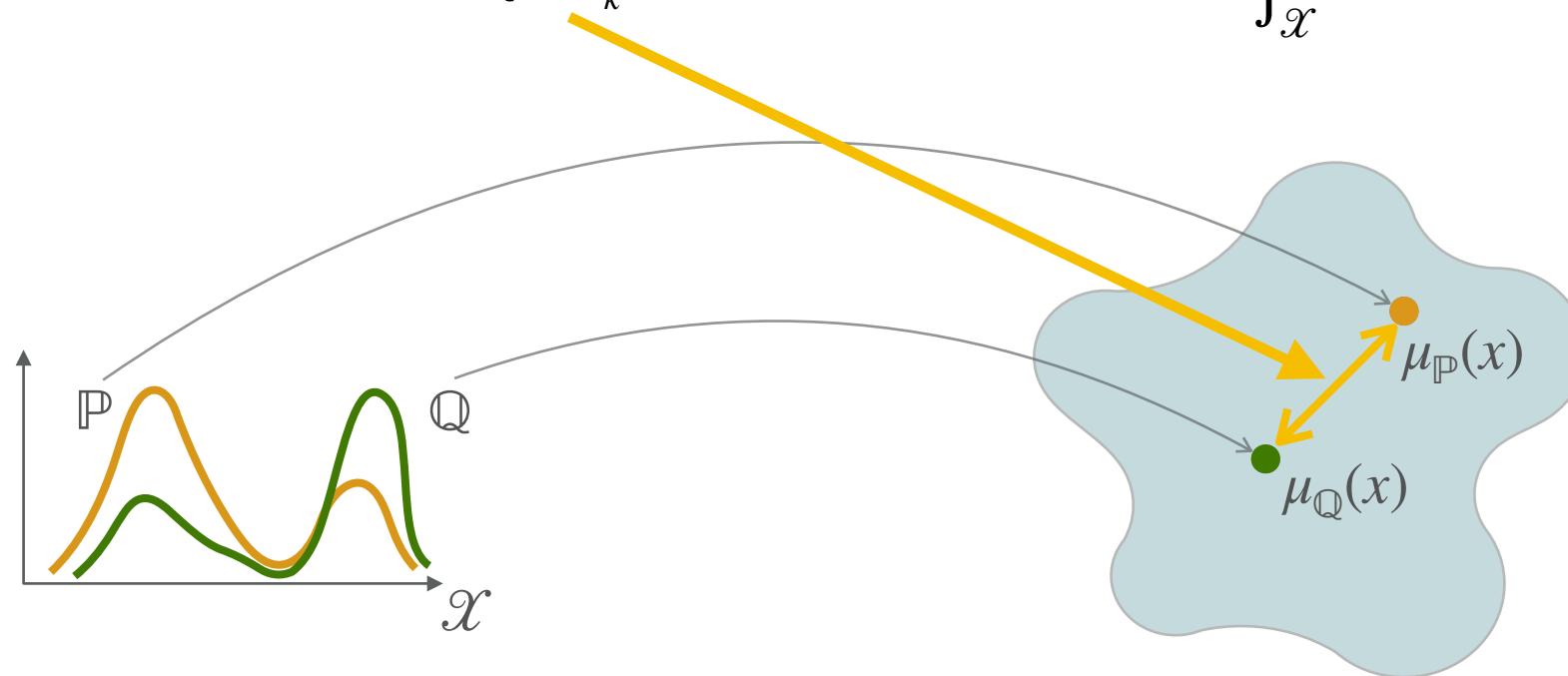


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$$\mu_{\mathbb{P}}(x) = \int_{\mathcal{X}} k(x, x') \mathbb{P}(\mathrm{d}x')$$

(1) Divergence ✓



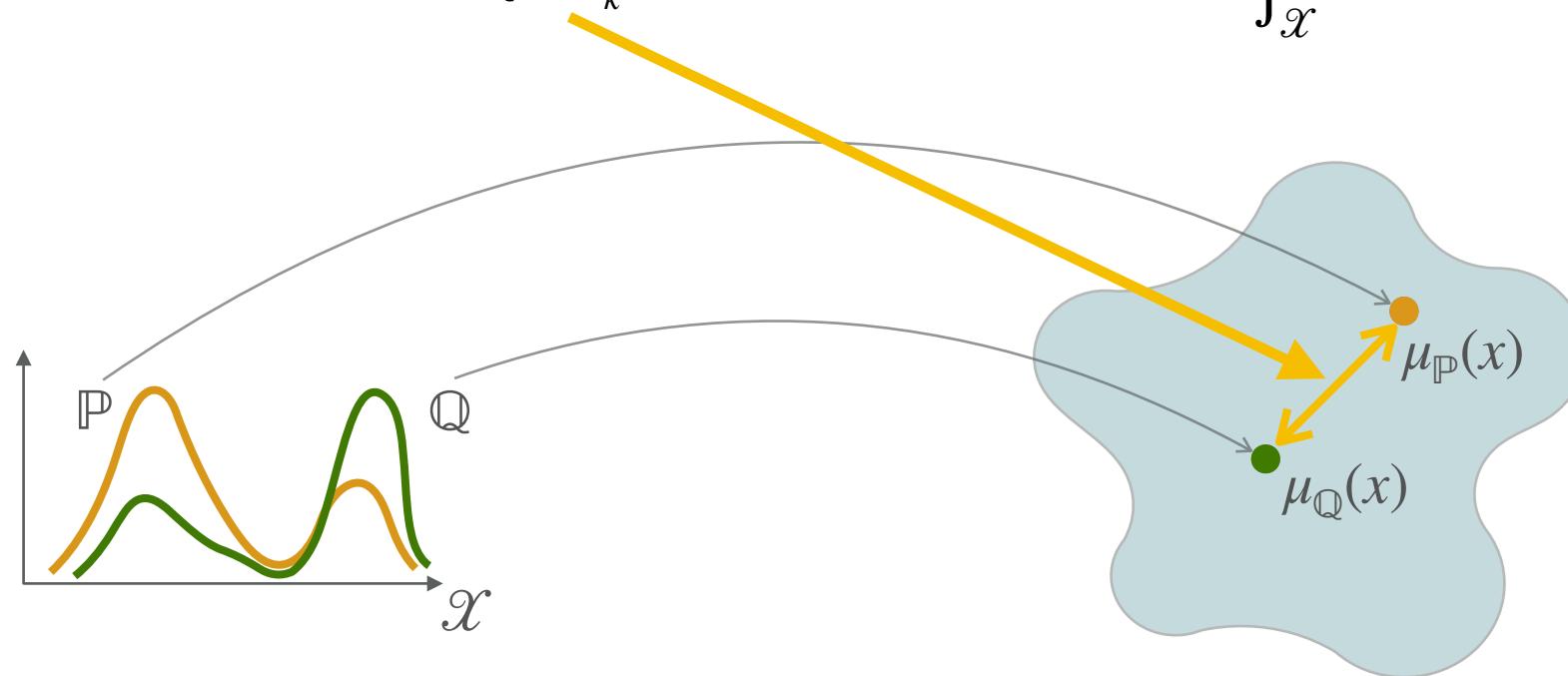
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- (2) Easy to estimate ✓

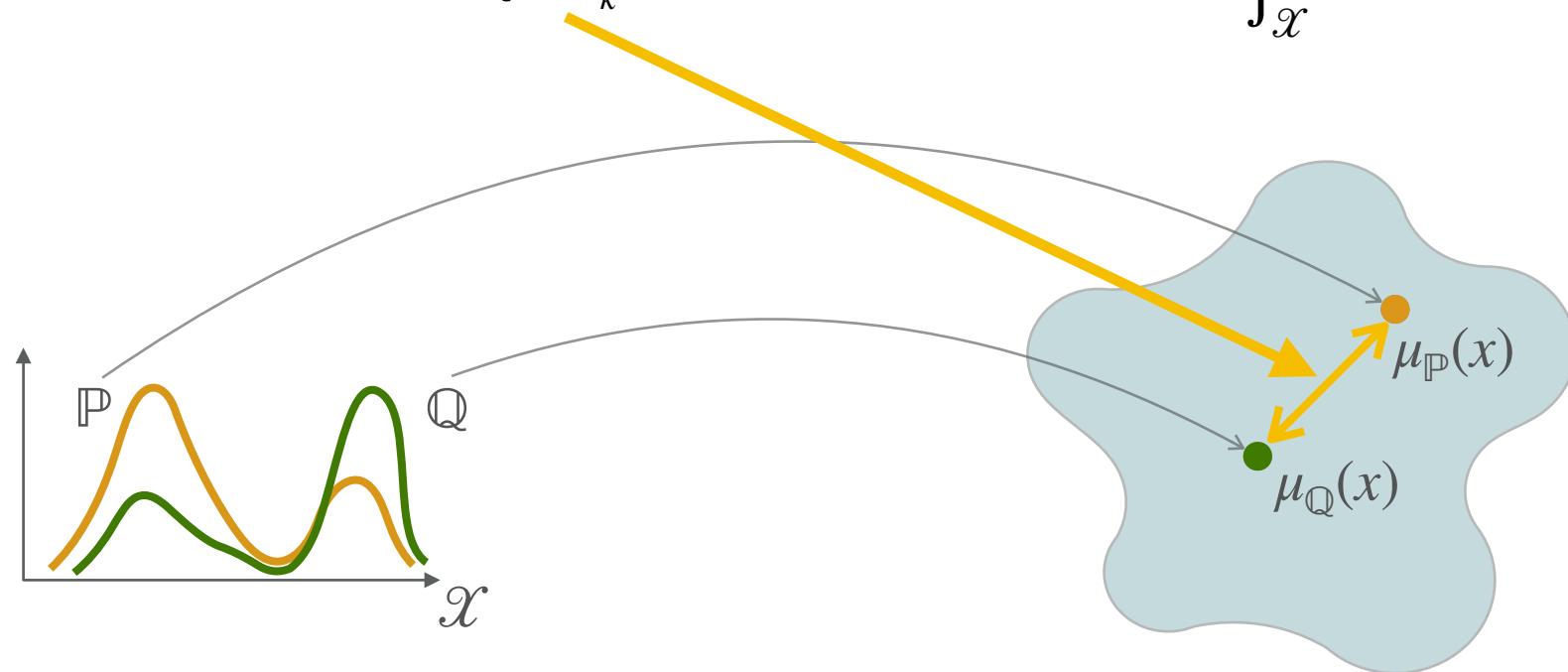
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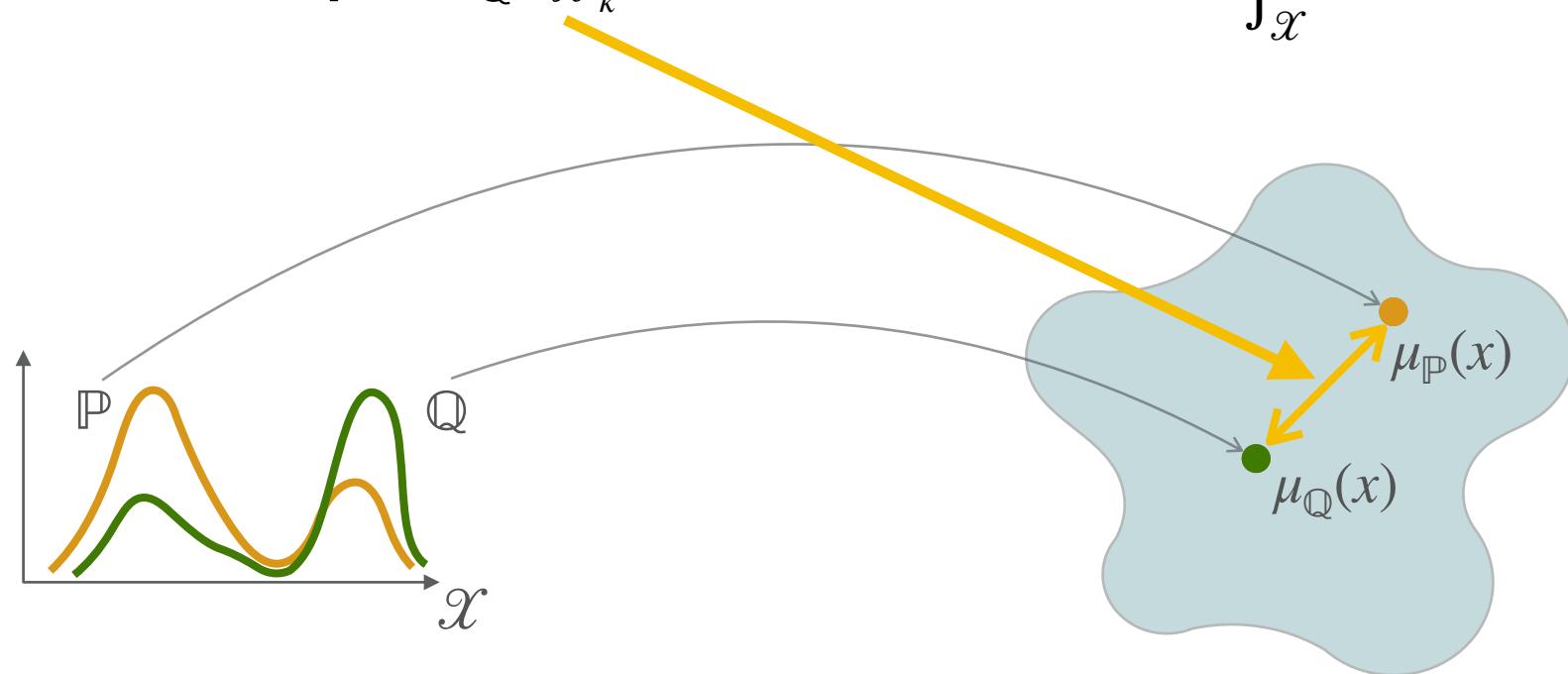
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Minimum MMD estimators

- Thanks to the ‘reproducing property’, we get:

$$\text{MMD}^2(\mathbb{P}, \mathbb{Q}) = \int_{\mathcal{X}} \int_{\mathcal{X}} k(x, y) \mathbb{P}(dx) \mathbb{P}(dy) - 2 \int_{\mathcal{X}} \int_{\mathcal{X}} k(x, y) \mathbb{P}(dx) \mathbb{Q}(dy) + \int_{\mathcal{X}} \int_{\mathcal{X}} k(x, y) \mathbb{Q}(dx) \mathbb{Q}(dy)$$

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- This leads to:

$$\hat{\theta}_n := \arg \min_{\theta \in \Theta} \text{MMD}^2(\mathbb{P}_\theta, Q_n)$$

Briol, F.-X., Barp, A., Duncan, A. B., & Girolami, M. (2019). Statistical inference for generative models with maximum mean discrepancy. *arXiv:1906.05944*.

Chérief-Abdellatif, B.-E., & Alquier, P. (2022). Finite sample properties of parametric MMD estimation: robustness to misspecification and dependence. *Bernoulli*, 28(1), 181–213.

Any Questions?

Approximate Bayesian Computation



(From now on we will mostly be Bayesian!)

Approximate Bayesian computation (ABC)

- Recall that we would like to approximate:

$$p(\theta | y_1, \dots, y_n) \propto \prod_{i=1}^n p(y_i | \theta) p(\theta)$$

Marin, J.-M., Pudlo, P., Robert, C. P., & Ryder, R. J. (2012). Approximate Bayesian computational methods. *Statistics and Computing*, 22, 1167–1180.

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$$q_{\text{ABC}}(\theta | y_1) \propto \left[\int_{\mathcal{X}} K_\epsilon(\|x_1 - y_1\|) p(x_1 | \theta) dx_1 \right] p(\theta)$$

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Surrogate likelihood!

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Uniform



Epanechnikov



Gaussian

Marin, J.-M., Pudlo, P., Robert, C. P., & Ryder, R. J. (2012). Approximate Bayesian computational methods. *Statistics and Computing*, 22, 1167–1180.

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A simple ABC sampler

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This is still intractable though!!

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Sampler for the ABC posterior

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We use the simulator:
 $x_{ti} = G_{\theta_t}(u_i), u_i \sim \mathbb{U}$



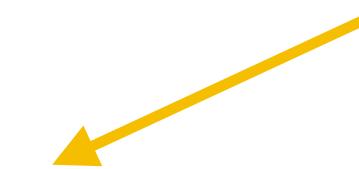
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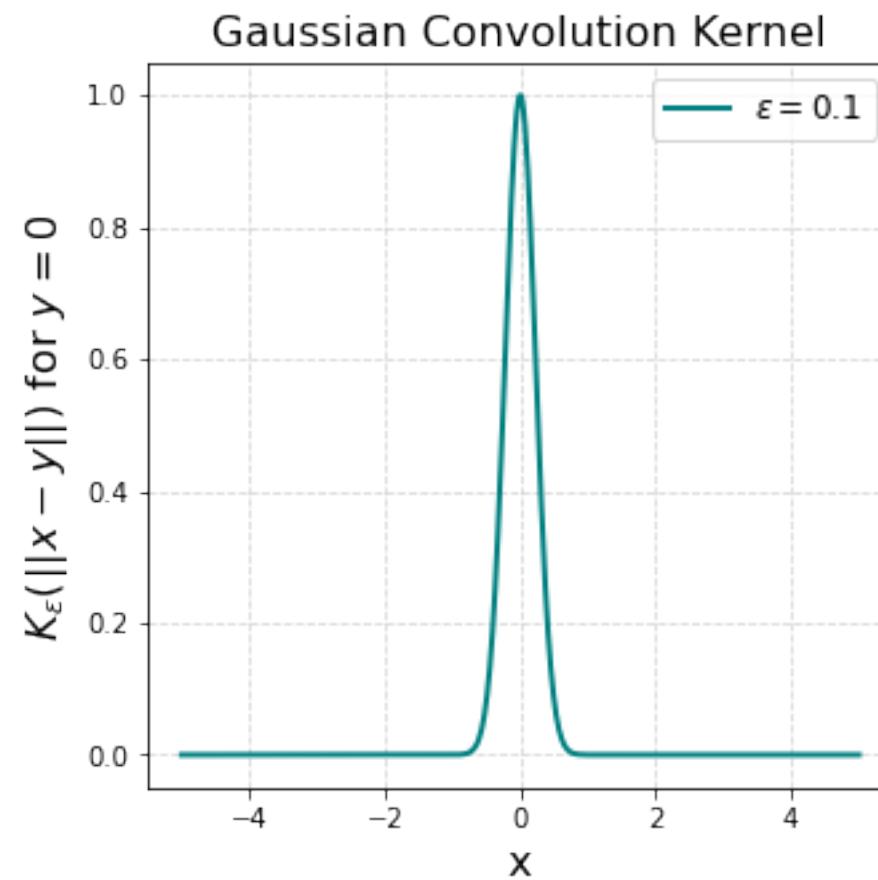
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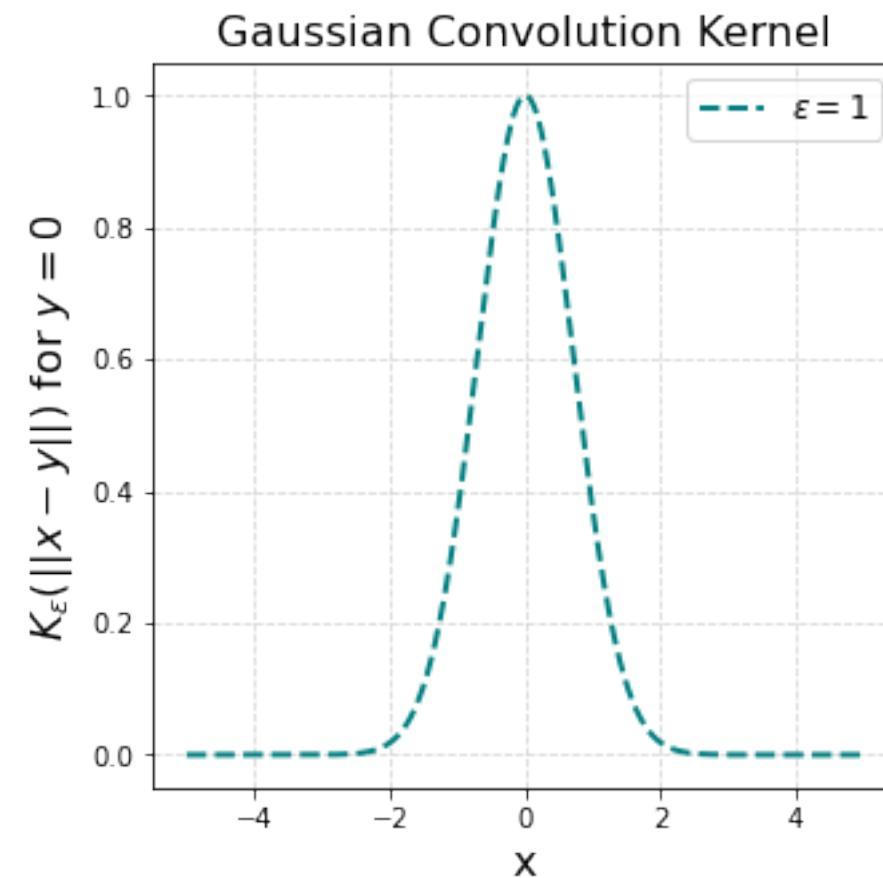


This is just a simple example; there are many more advanced sampling methods (e.g. SMC)

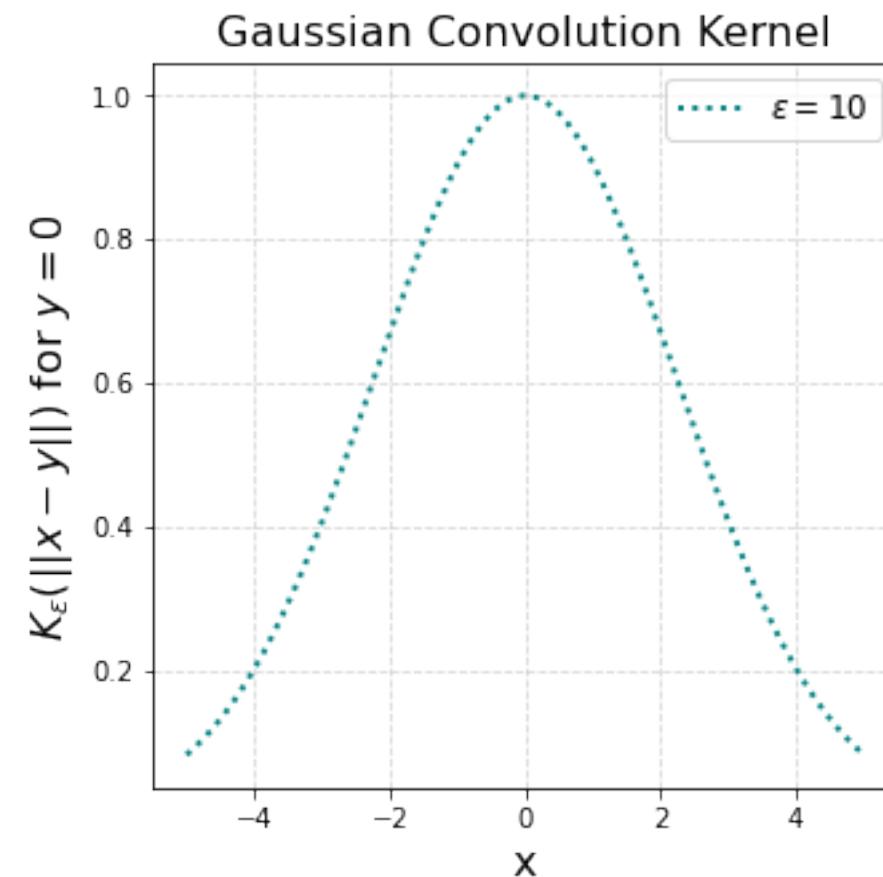
The impact of ϵ



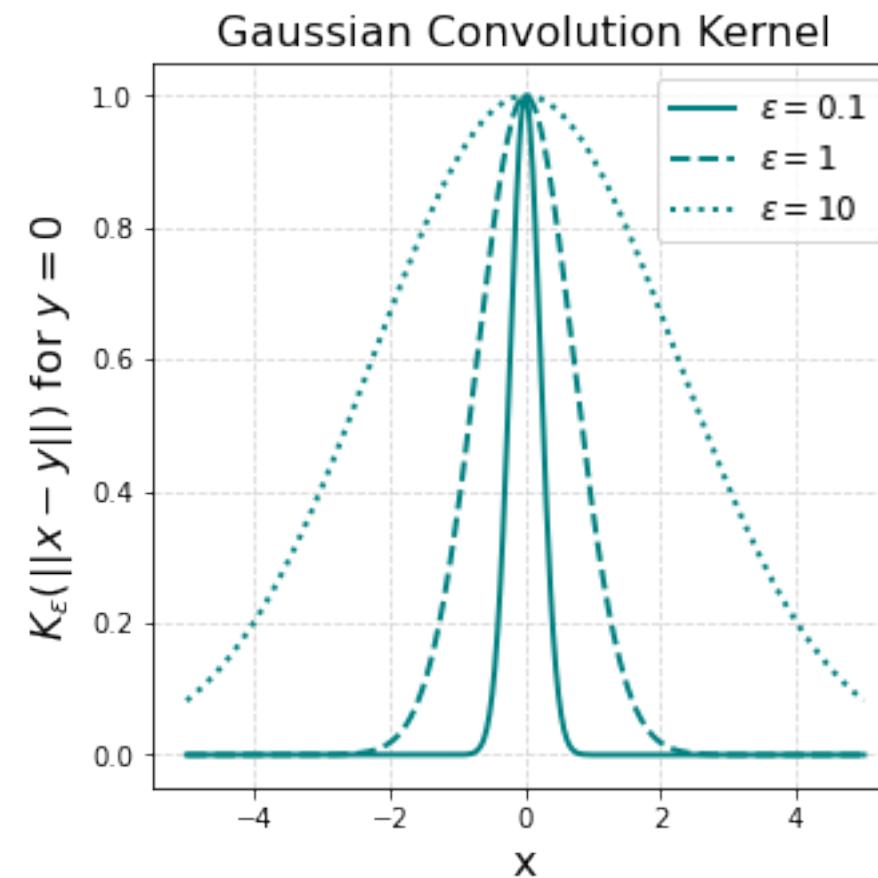
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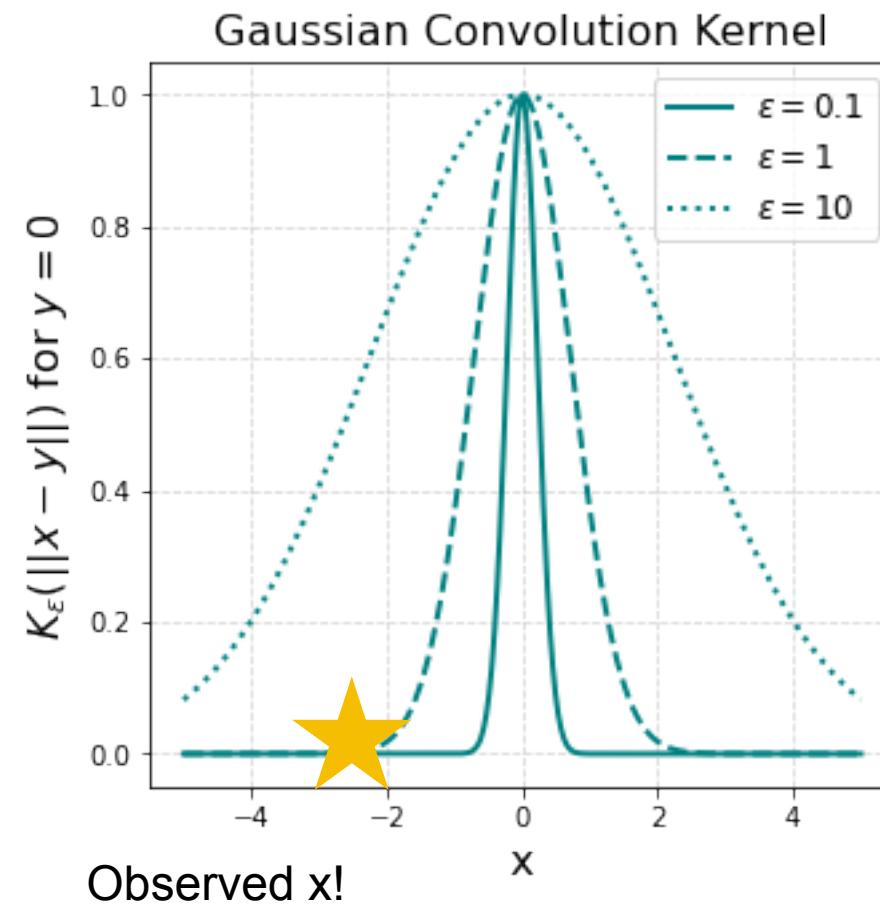
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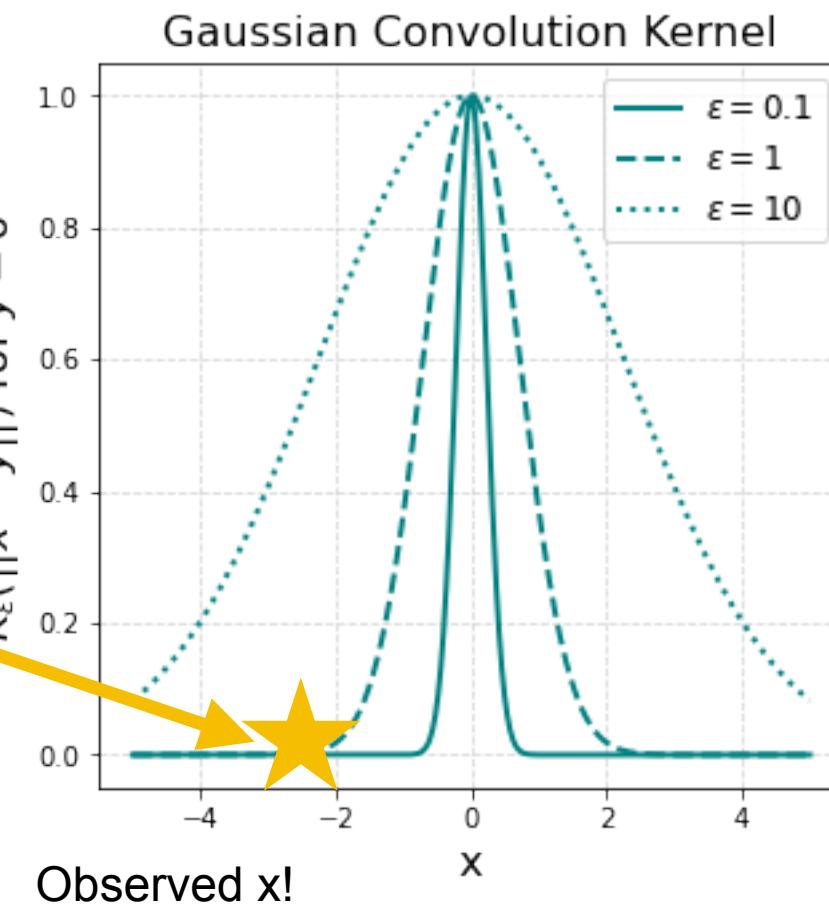
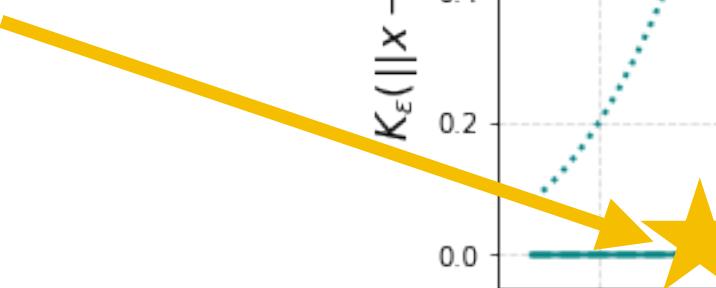


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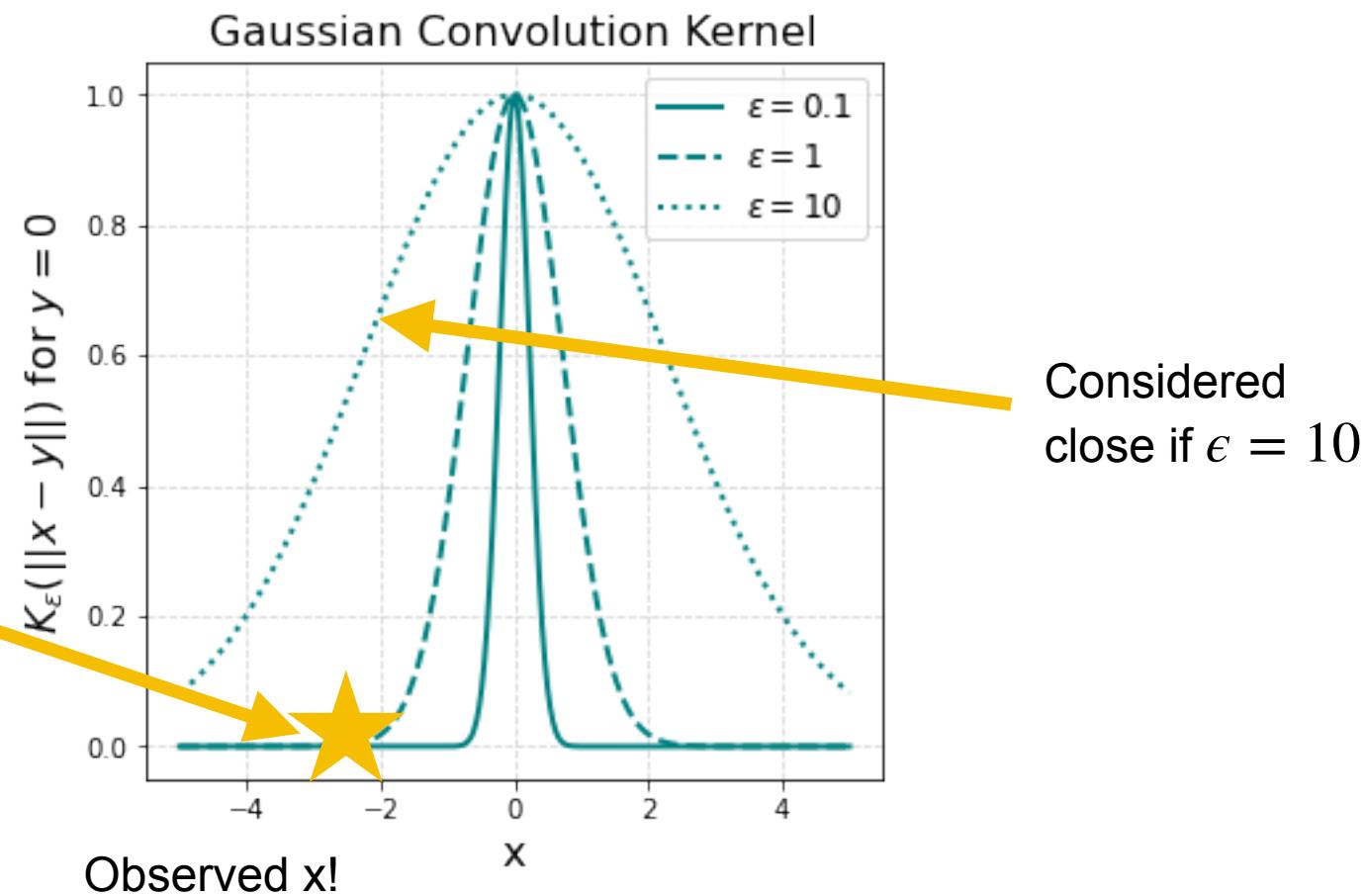
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Discrepancies-based ABC

$$q_{\text{ABC}}(\theta | y_1, \dots, y_n) \propto \int_{\mathcal{X}} \dots \int_{\mathcal{X}} K_\epsilon(\|x - y\|) \prod_{i=1}^n p(x_i | \theta) p(\theta) dx_1 \dots dx_n$$

~~$K_\epsilon(\|x - y\|)$~~

$K_\epsilon \left(D \left(\frac{1}{n} \sum_{i=1}^n \delta_{x_i}, \frac{1}{n} \sum_{i=1}^n \delta_{y_i} \right) \right) = K_\epsilon \left(D \left((\mathbb{P}_\theta)_n, \mathbb{Q}_n \right) \right)$

- Park, M., Jitkrittum, W., & Sejdinovic, D. (2016). K2-ABC: Approximate bayesian computation with kernel embeddings. *AISTATS*, 51, 398–407.
- Bernton, E., Jacob, P. E., Gerber, M., & Robert, C. P. (2019). Approximate Bayesian computation with the Wasserstein distance. *JRSSB*, 81(2), 235–269.
- Legramanti, S., Durante, D., & Alquier, P. (2025). Concentration and robustness of discrepancy-based ABC via Rademacher complexity. *The Annals of Statistics*, 53(1), 37–60.

Any Questions?

ML approaches to SBI



We have now already covered the state-of-the-art until 2020-ish!

SBI with conditional density estimators

- I probably don't need to convince you that machine learning methods are very good at emulation.... How can we use this for Bayes?

$$p(\theta | y_1, \dots, y_n) \propto \prod_{i=1}^n p(y_i | \theta) p(\theta)$$

Zammit-mangion, A., Sainsbury-Dale, M., & Huser, R. (2025). Neural methods for amortized parameter inference. *Annual Review of Statistics and Its Application*, 12, 311–335.

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- Both are conditional densities, and so we need to think about how we can use the 'power' of machine learning to emulate this type of quantity.

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- Both are conditional densities, and so we need to think about how we can use the 'power' of machine learning to emulate this type of quantity.
- We will start by emulating the likelihood; i.e. we want a flexible class: $\{q_\phi(x | \theta)\}_{\Phi \in \Phi}$

Zammit-mangion, A., Sainsbury-Dale, M., & Huser, R. (2025). Neural methods for amortized parameter inference. *Annual Review of Statistics and Its Application*, 12, 311–335.

Deistler, M., Boelts, J., Steinbach, P., Moss, G., Moreau, T., Gloeckler, M., Rodrigues, P. L. C., Linhart, J., Lappalainen, J. K., Miller, B. K., Gonçalves, P. J., Lueckmann, J.-M., Schröder, C., & Macke, J. H. (2025). Simulation-based inference: A practical guide. *arXiv:2508.12939*.

Some simpler models...

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- We can increase the flexibility:

$$q_\phi(x | \theta) = \sum_{c=1}^C w_c(\phi; \theta) \mathcal{N}(x | \mu_c(\phi; \theta), \Sigma_c(\phi; \theta))$$

Transformations and densities

- Consider some base distribution p_ν , and some transformation T such that

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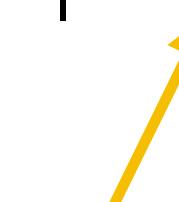
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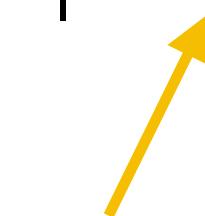
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Use neural networks!!

Normalising flows (I)

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We can also parametrise them!

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Normalising flows (I)

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Straightforward to create conditional density!

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Normalising flows (II)

$$q_{\phi}(x \mid \theta) = p_v(v) \left| \det J_{T_{\phi,\theta}}(x) \right|^{-1} \quad T_{\phi,\theta} = T_{\phi,\theta}^K \circ \dots \circ T_{\phi,\theta}^2 \circ T_{\phi,\theta}^1$$

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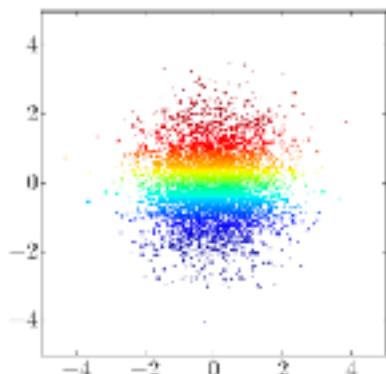
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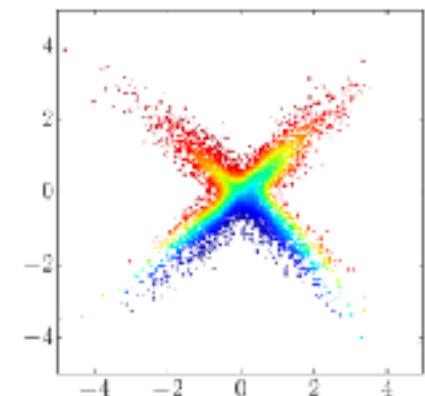
They can be, but (similarly to diffusion models) they do not typically encode any science, they are just constructed to be very flexible models!

Normalising flows (III)

$$p_v(v)$$



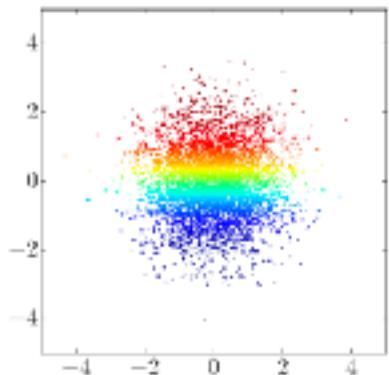
$$q_\phi(x)$$



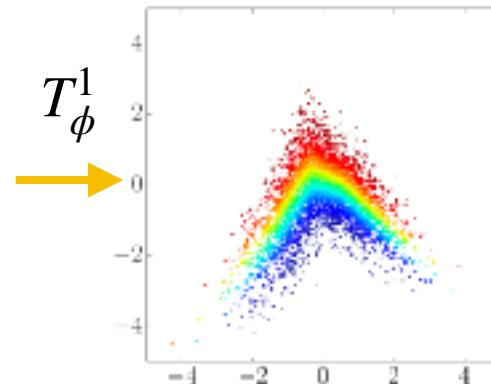
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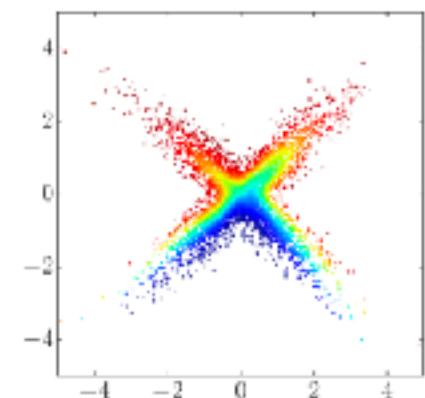
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$$T_\phi^1$$

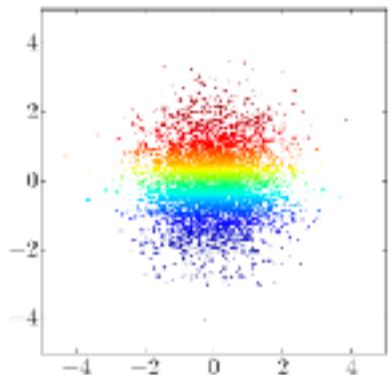
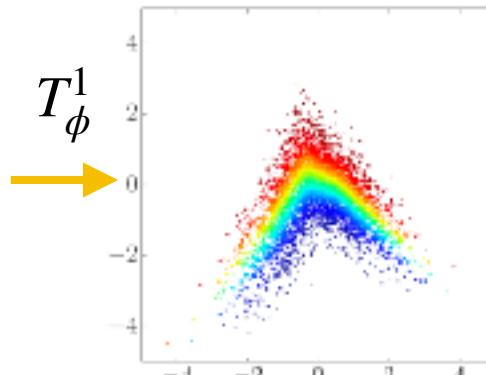
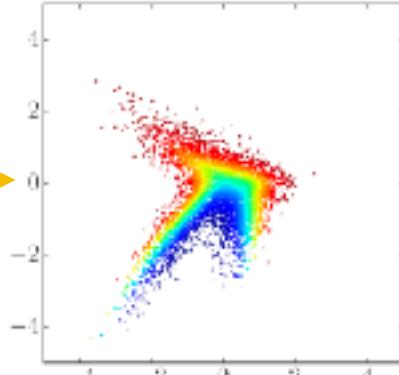
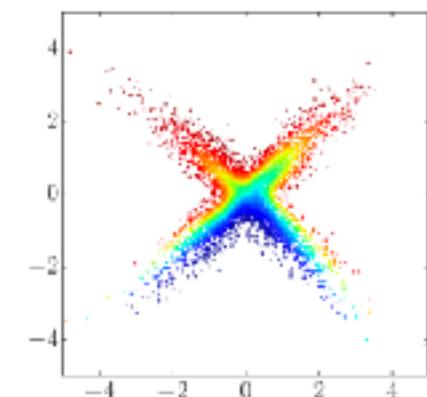


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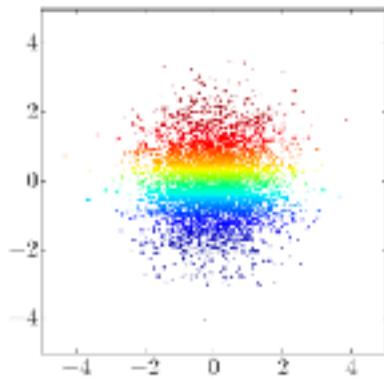
Normalising flows (III)

 $p_v(v)$  T_ϕ^1  T_ϕ^2  $q_\phi(x)$ 

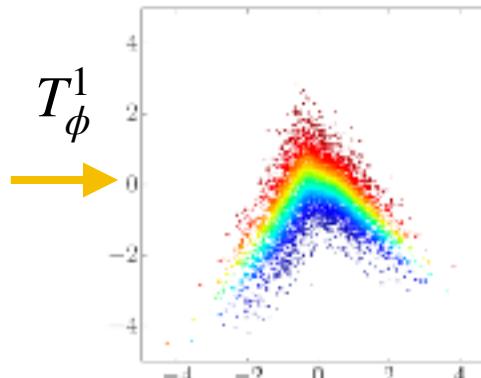
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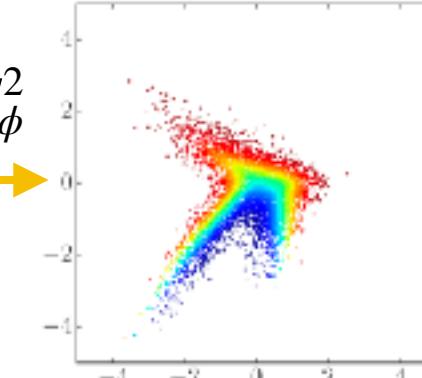
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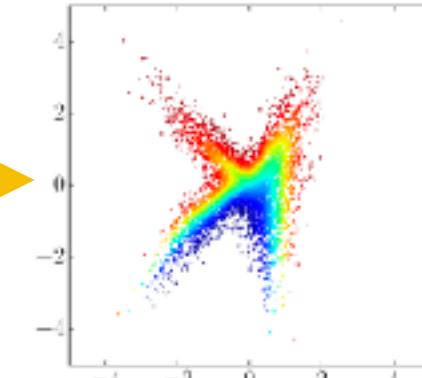
T_ϕ^1



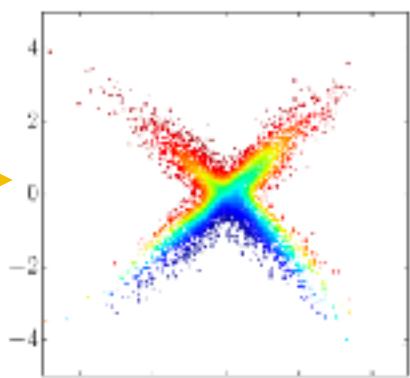
T_ϕ^2



T_ϕ^3



T_ϕ^4

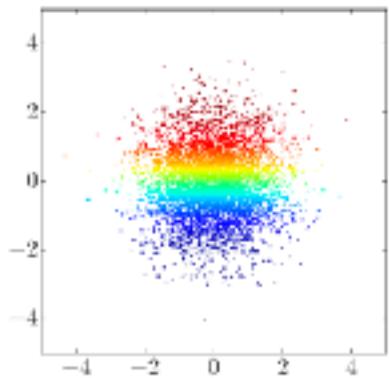


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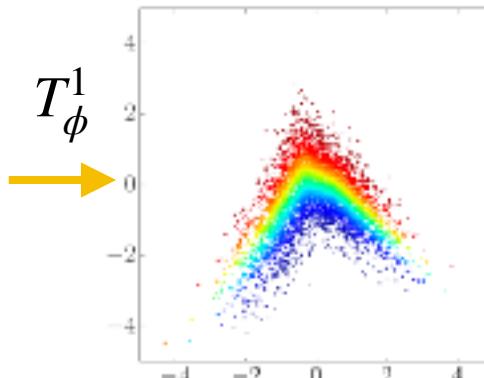
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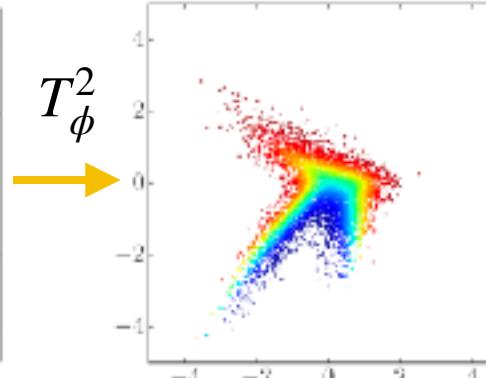
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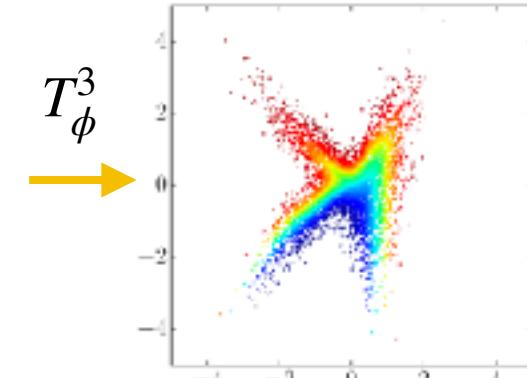
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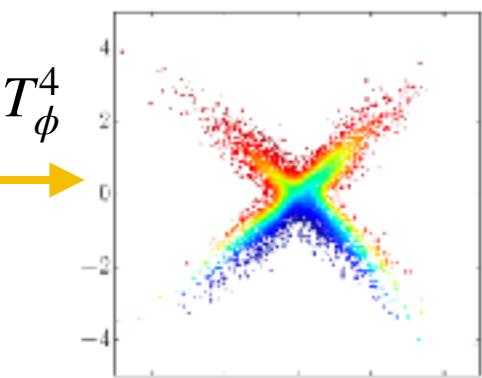
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$$q_\phi(x)$$

The composition of relatively simple transformations can give fairly complex maps!

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Neural likelihood estimation (NLE)

- **Step 1:** train $q_{\phi}(x | \theta)$ to approximate the likelihood using samples from the prior ($\theta_1, \dots, \theta_n \sim p(\theta)$) and simulator ($x_i \sim p(\cdot | \theta_i)$):

$$\hat{\phi}_n := \arg \min_{\phi \in \Phi} \ell_{\text{NLE}}(\phi), \quad \ell_{\text{NLE}}(\phi) = -\frac{1}{n} \sum_{i=1}^n \log q_{\phi}(x_i | \theta_i) \approx -\mathbb{E}_{\theta \sim p(\theta)}[\mathbb{E}_{x \sim p(\cdot | \theta)}[\log q_{\phi}(x | \theta)]]$$

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- **Step 2:** Approximate posterior (MCMC, VI) constructed with surrogate likelihood!

$$p_{\text{NLE}}(\theta | y_1, \dots, y_n) \propto \prod_{i=1}^n q_{\hat{\phi}_n}(y_i | \theta) p(\theta)$$

Amortisation for NLE

- Recall the NLE posterior:

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We still need to re-run MCMC/VI though... We are **partially amortised**.

Neural posterior estimation (NPE)

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- **Step 2:** Condition on the observed data:

$$p_{\text{NPE}}(\theta | y_1, \dots, y_n) = q_{\hat{\phi}_n}(\theta | y_1, \dots, y_n)$$

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- We have a direct handle on the new posterior; no need for MCMC/VI!



We are **fully amortised**.

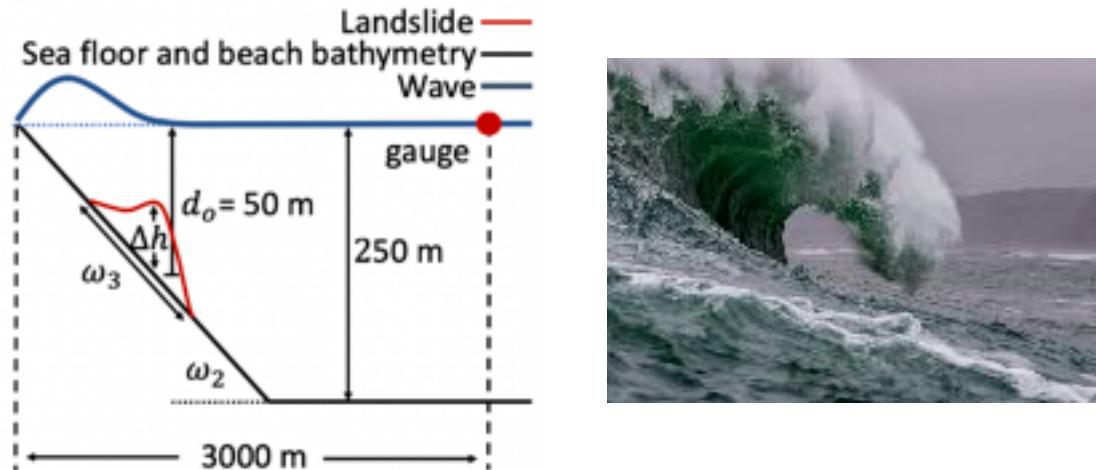
Any Questions?

Challenges with existing SBI methods



Challenge 1: Expensive simulators

Example 1:

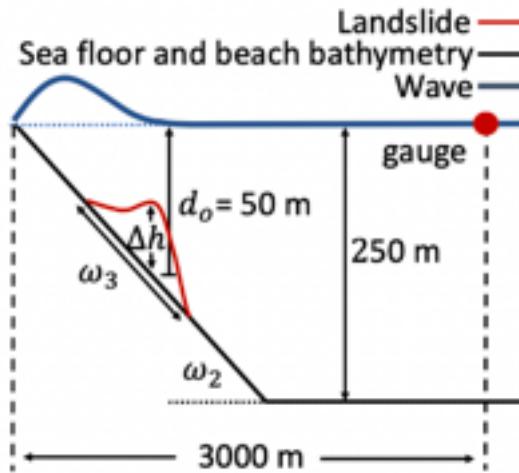


≈ 2 hours per sim on laptop

Li, K., Giles, D., Karvonen, T., Guillas, S., & Briol, F.-X. (2023).
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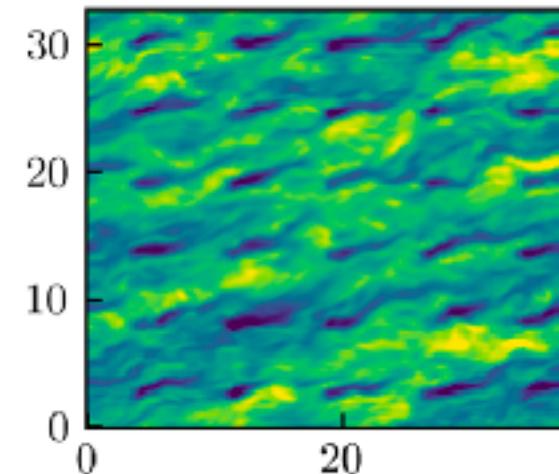
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Example 2:

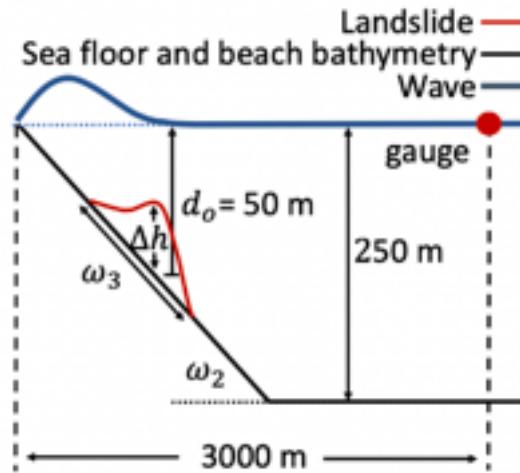


≈ 100 hours per sim on Met Office cluster

Kirby, A., Briol, F.-X., Dunstan, T. D., & Nishino, T. (2023). Data-driven modelling of turbine wake interactions and flow resistance in large wind farms. *Wind Energy*, 26(9), 875–1011.

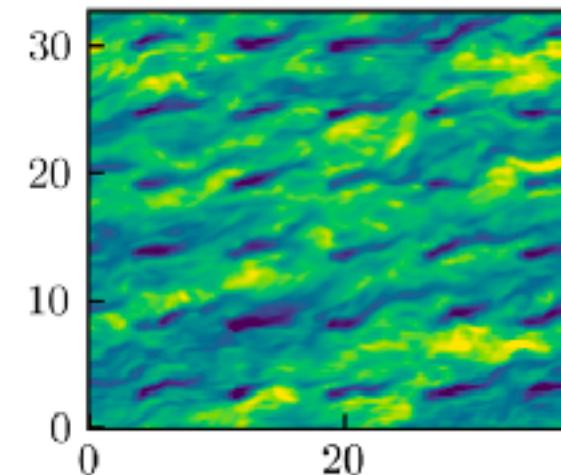
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≈ 2 hours per sim on laptop

Example 2:



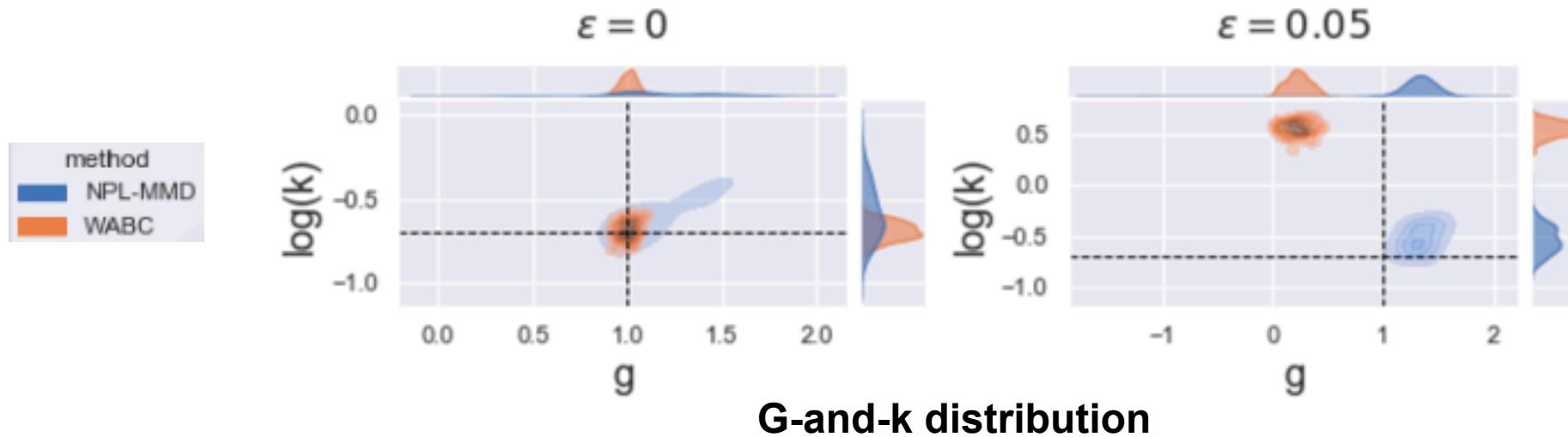
≈ 100 hours per sim on Met Office cluster



Currently out of reach of modern SBI methods!



Challenge 2: Model misspecification

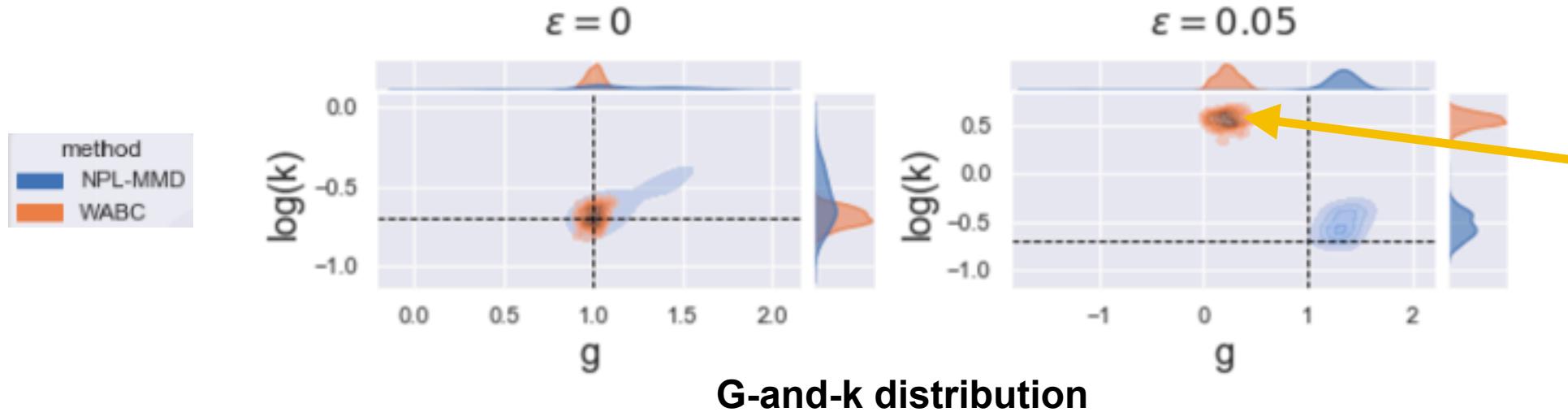


Dellaporta, C., Knoblauch, J., Damoulas, T. & **Briol, F-X** (2022). Robust Bayesian inference for simulator-based models via the MMD posterior bootstrap. AISTATS, 943-970. Best paper award.

Kelly, R. P., Warne, D. J., Frazier, D. T., Nott, D. J., Gutmann, M. U., & Drovandi, C. (2025). Simulation-based Bayesian inference under model misspecification. *arXiv:2503.12315*.



Challenge 2: Model misspecification



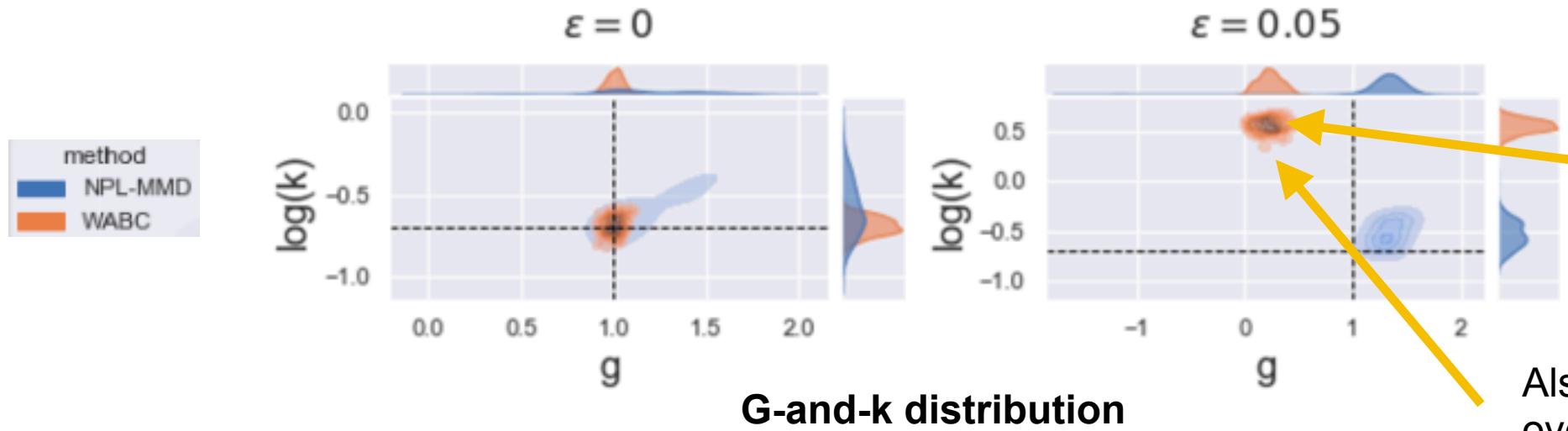
A tiny % of corrupted observations is enough to seriously affect the posterior.

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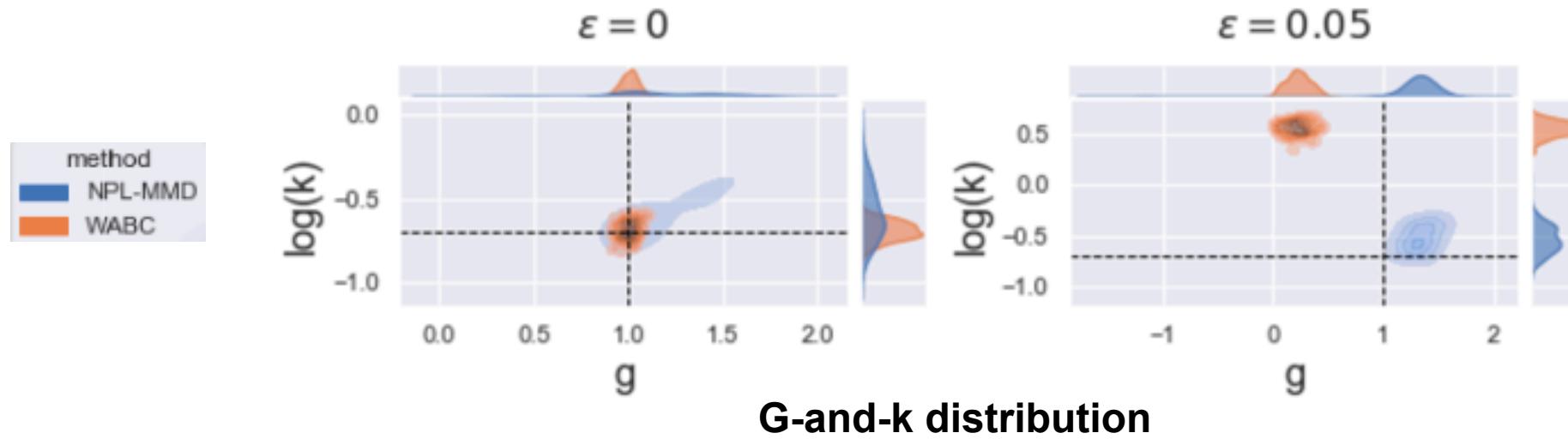
Also leads to serious overconfidence!

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Challenge 2: Model misspecification



Currently very few robust methods with theoretical guarantees

Challenge 3: Over-confidence

Published in Transactions on Machine Learning Research (11/2022)

A Trust Crisis In Simulation-Based Inference? Your Posterior Approximations Can Be Unfaithful

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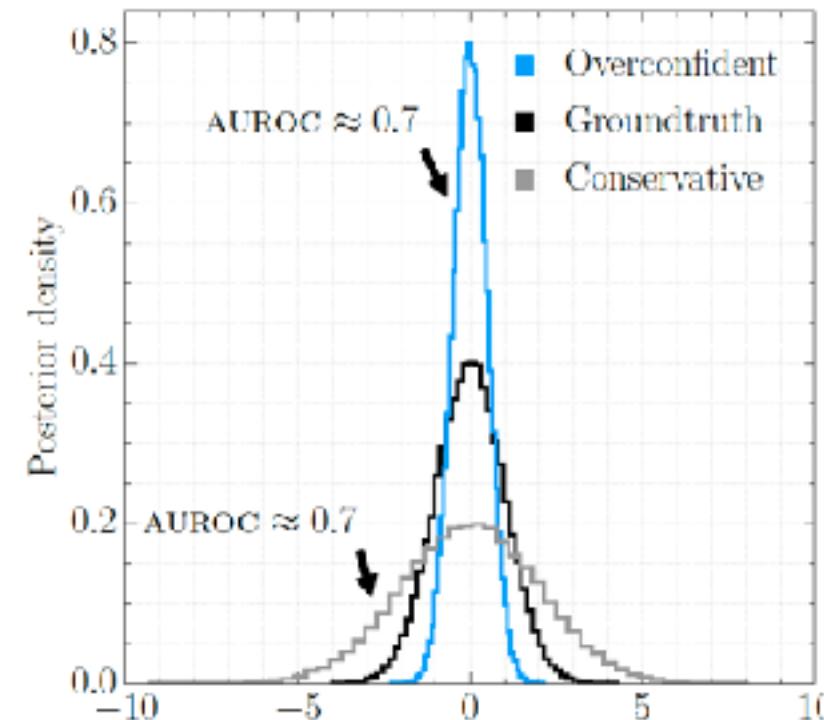
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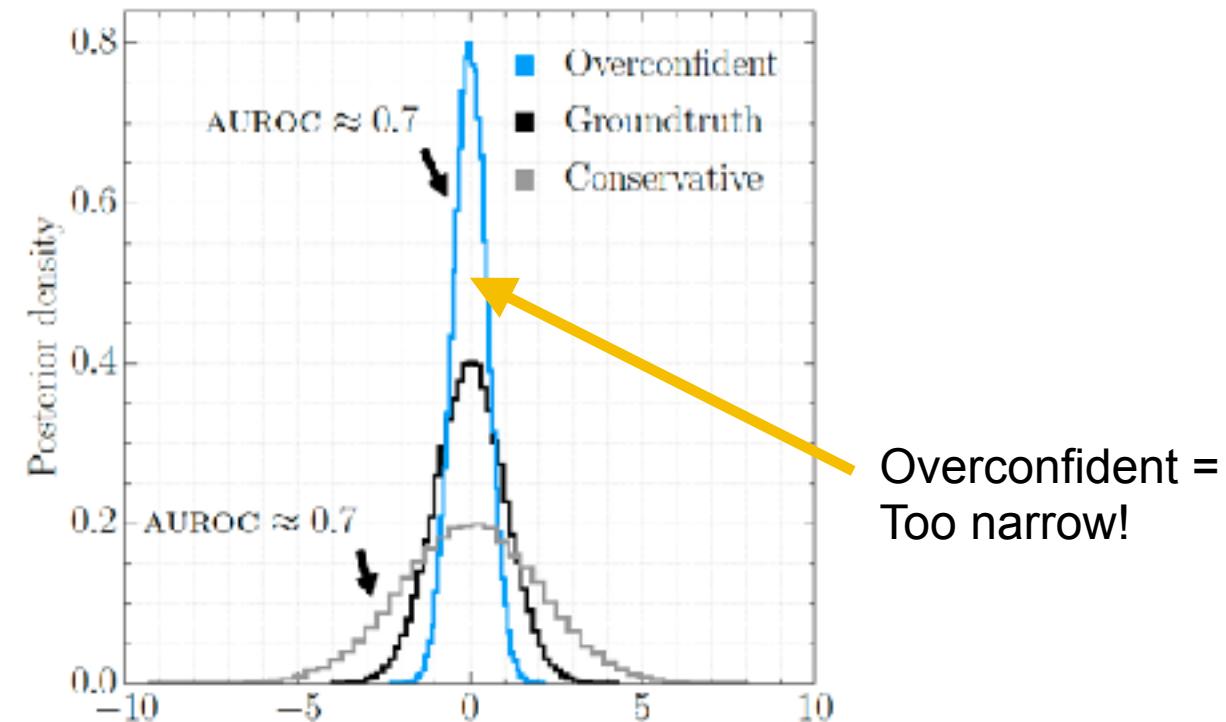
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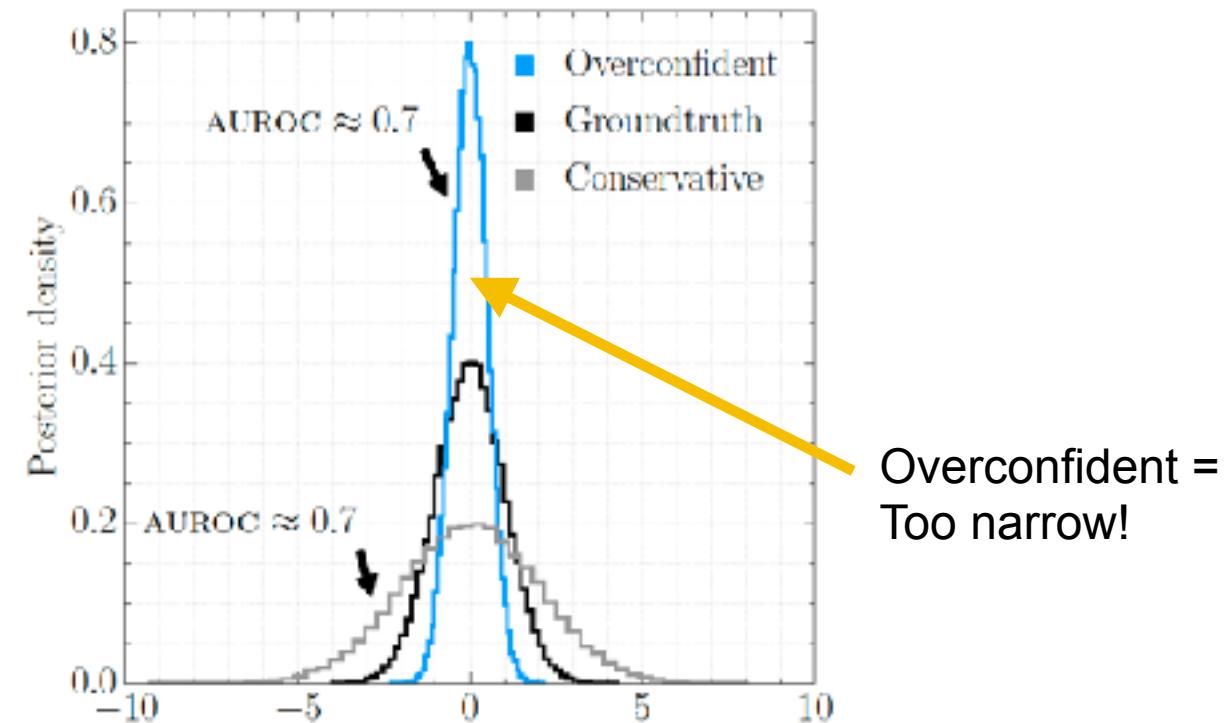
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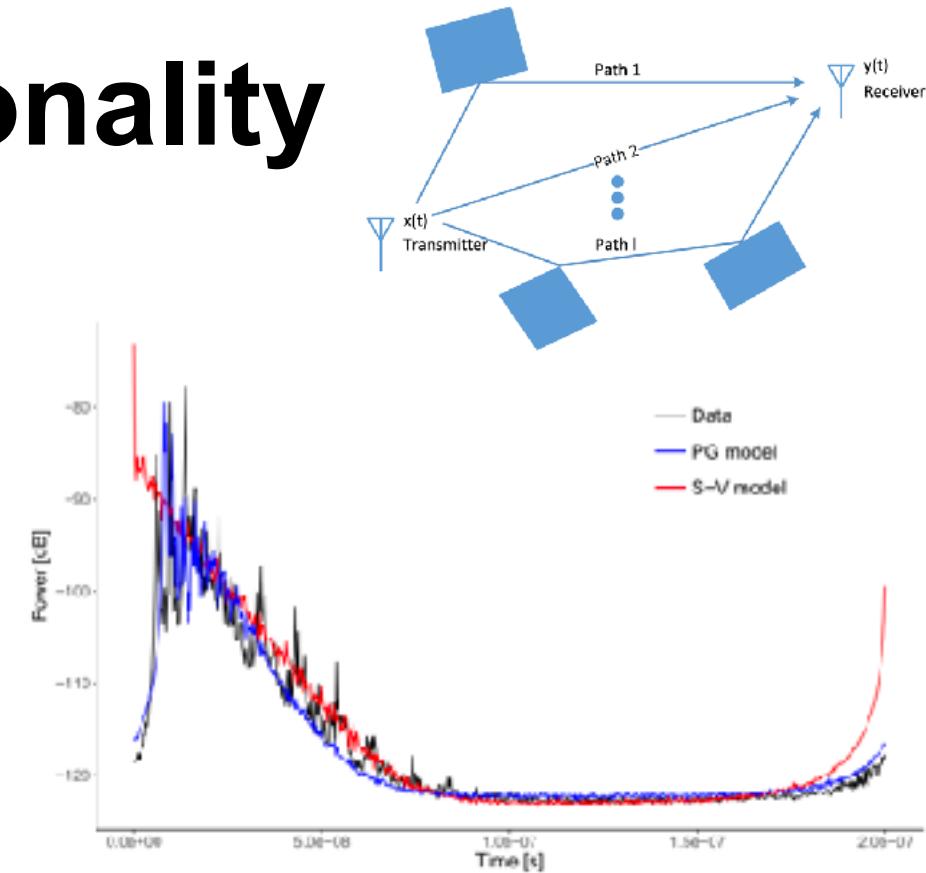
Observation 1 All benchmarked algorithms may produce non-conservative posterior approximations.

Challenge 4: High-dimensionality

- As with everything in stats/ML, the curse of dimensionality hurts us.... Computing distances or estimating densities is very tough!

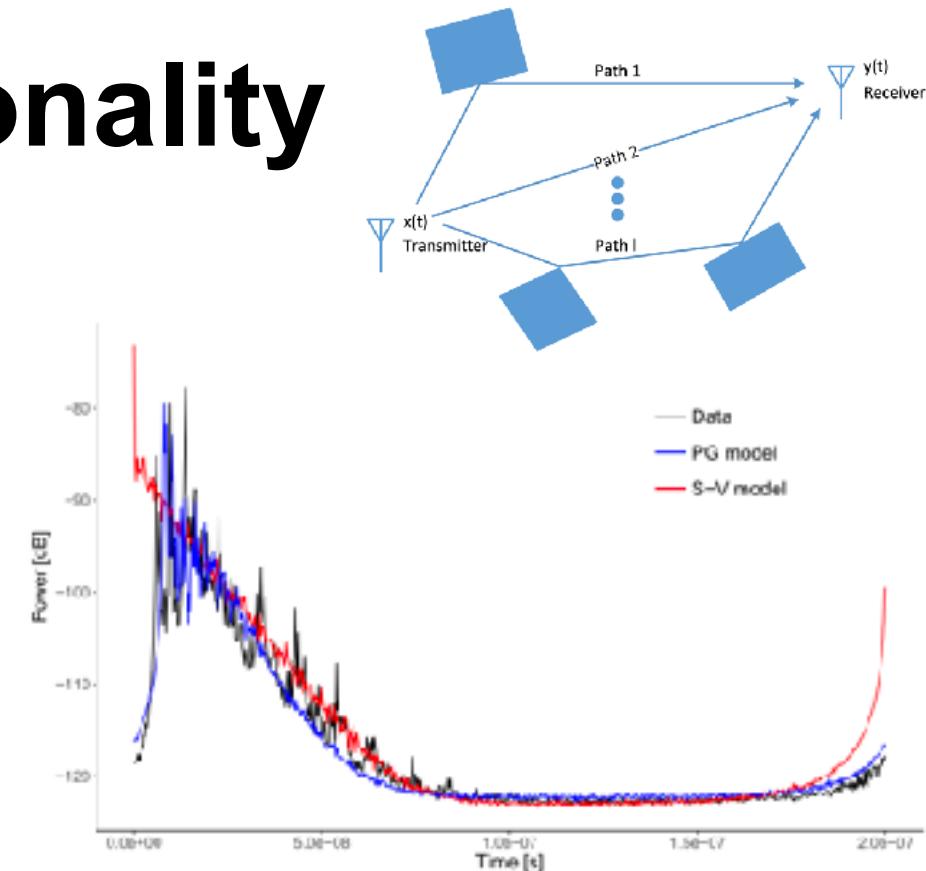
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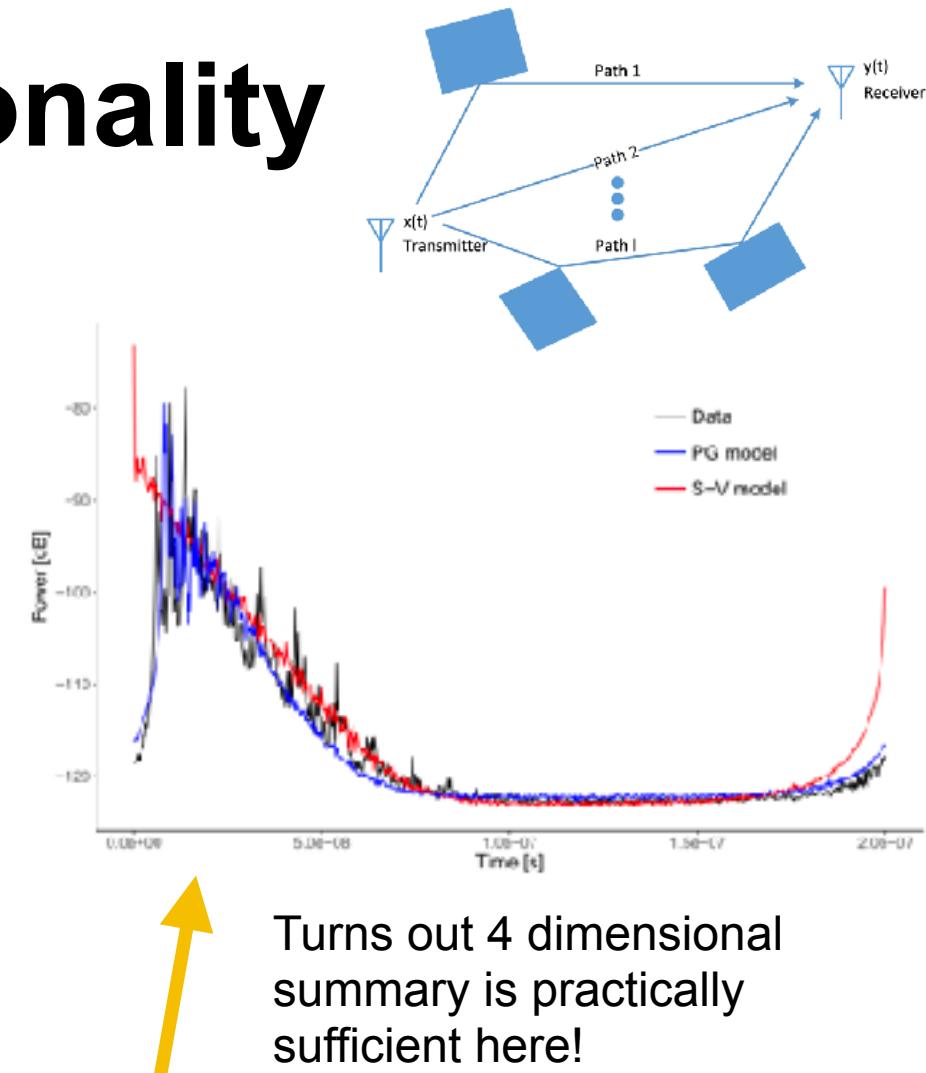


Turns out 4 dimensional summary is practically sufficient here!

Bharti, A., Briol, F.-X., & Pedersen, T. (2021). A general method for calibrating stochastic radio channel models with kernels. *IEEE Transactions on Antennas and Propagation*, 70(6), 3986–4001.

Challenge 4: High-dimensionality

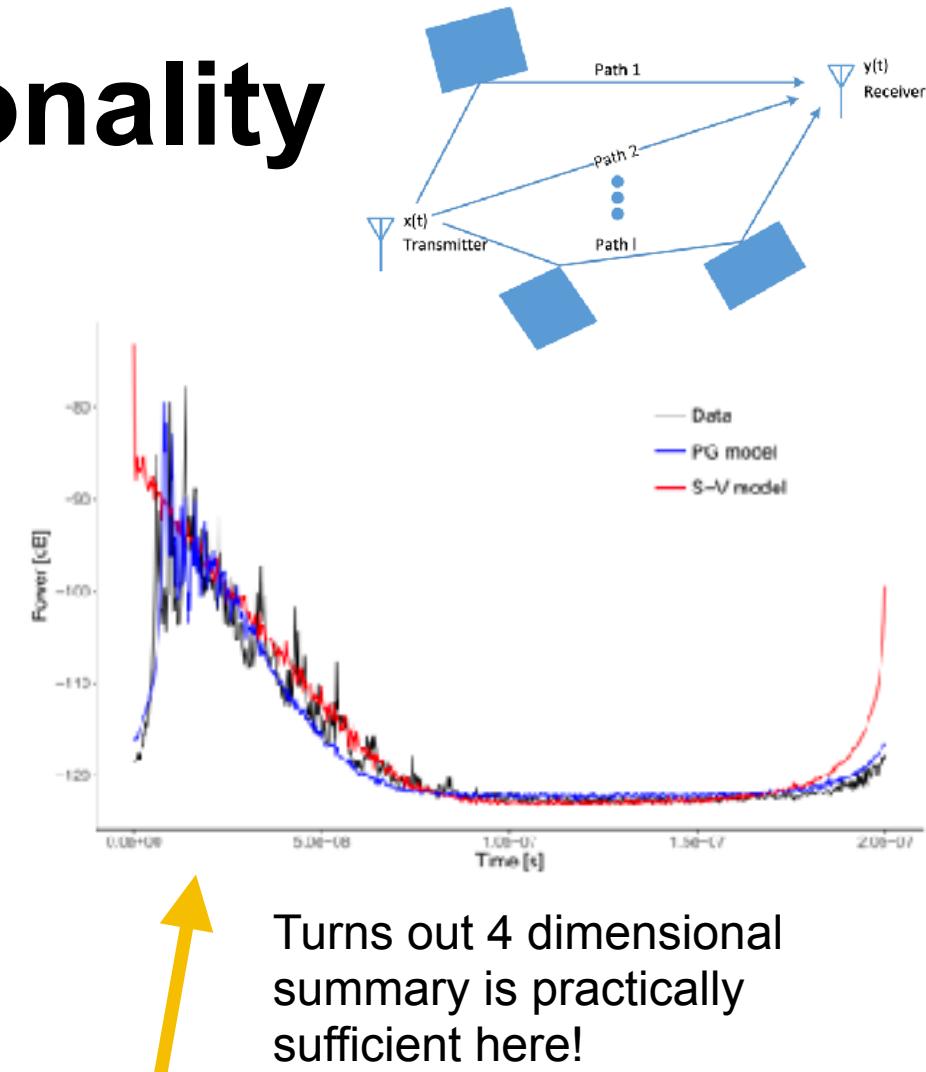
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- We therefore end up working with **summary statistics**, either hand-crafted or learnt via a neural network (i.e. a ‘summary network’).



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- We therefore end up working with **summary statistics**, either hand-crafted or learnt via a neural network (i.e. a ‘summary network’).
- Dimensionality of parameter space also a problem...



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Roadmap going ahead...

Background + challenges for SBI

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Snapshot 1:
Multi-fidelity methods for
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(NeurIPS?, 2025)

Roadmap going ahead...

Background + challenges for SBI

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Snapshot 2:
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Roadmap going ahead...

Background + challenges for SBI

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Snapshot 3:
Provably robust
generalisation of Bayes for
simulation-based inference
(AISTATS Best Paper
Award, 2022)

Any Questions?

Multilevel neural simulation-based inference



Paper: Hikida, Y., Bharti, A., Jeffrey, N. & **Briol, F-X** (2025). Multilevel neural simulation-based inference. arXiv:2506.06087. (to appear at NeurIPS?)

Code: <https://github.com/yugahikida/multilevel-sbi>

Challenge for SBI

Simulators can be really computationally expensive!

Challenge for SBI

Simulators can be really computationally expensive!

- Most simulators used in SBI papers take only a few seconds (or less) to run.
- Even if a simulator takes only a few minutes, we typically need thousands of simulations!
- Simulators that take more time are currently out of reach of existing methods.

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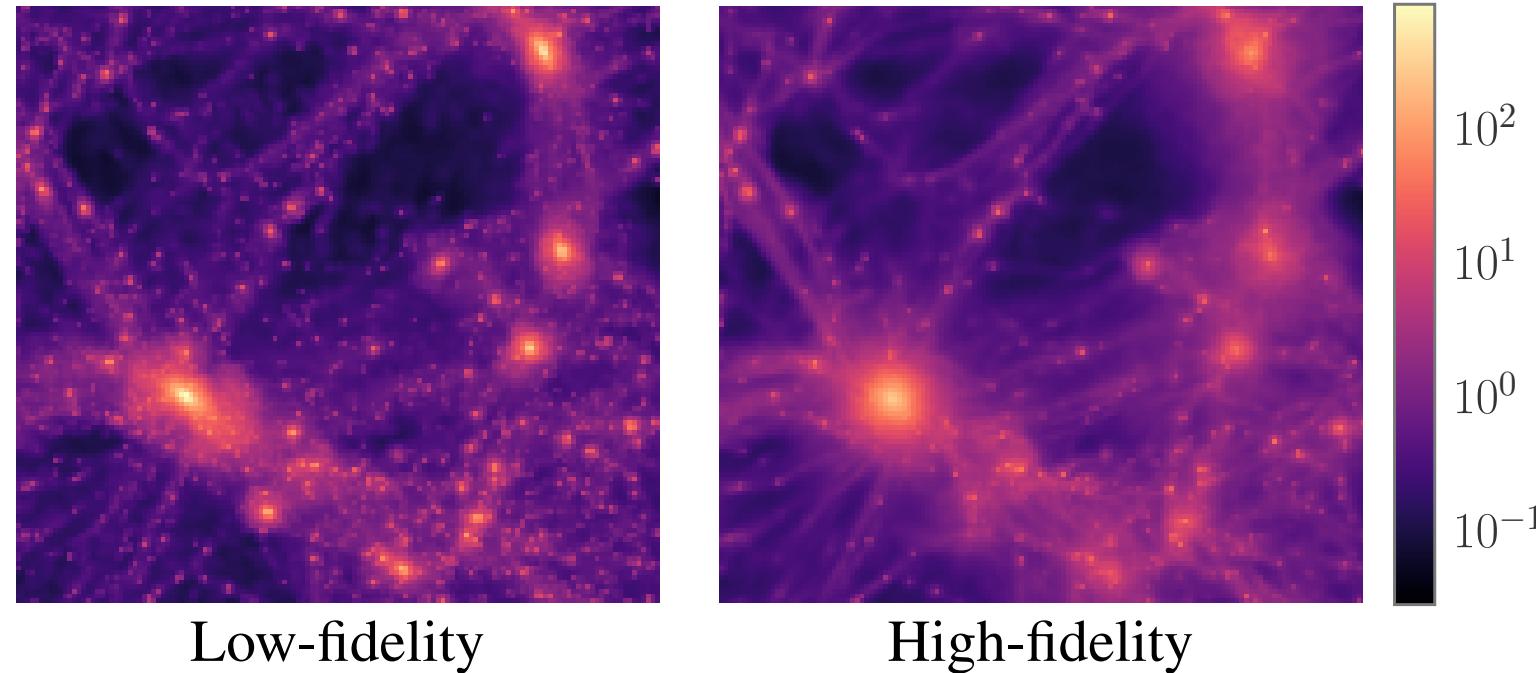
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 This leads to a form of model misspecification by design!



SBI for cosmology



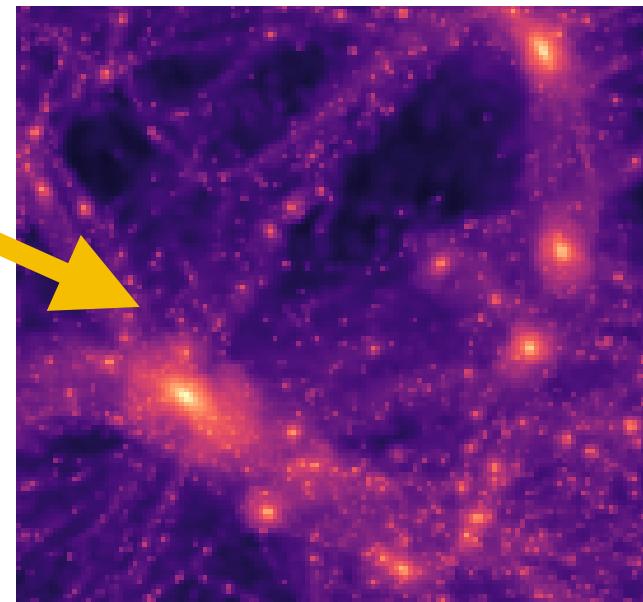
Jeffrey, N., et al. (2025). Dark energy survey year 3 results: likelihood-free, simulation-based w Λ CDM inference with neural compression of weak-lensing map statistics. *Monthly Notices of the Royal Astronomical Society*, 536(2), 1303–1322.

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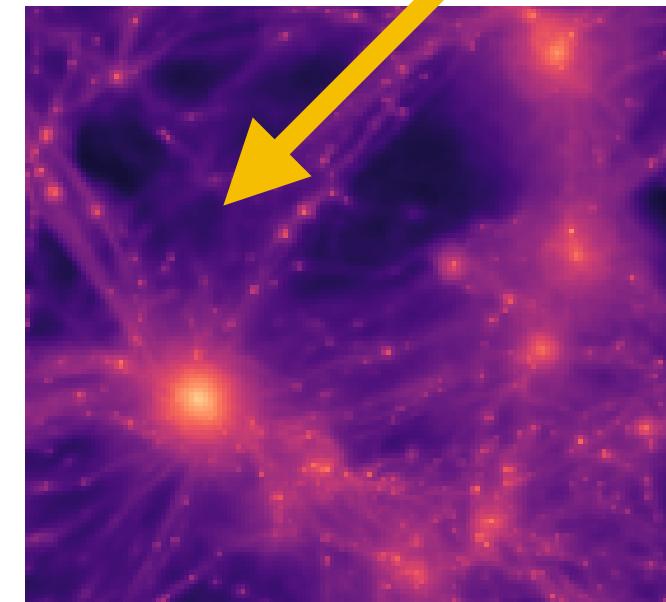
SBI for cosmology

Gravity-only N-body simulations

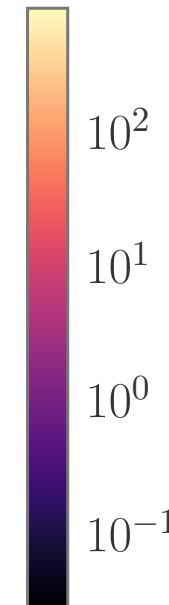


Low-fidelity

Hydrodynamic simulations



High-fidelity



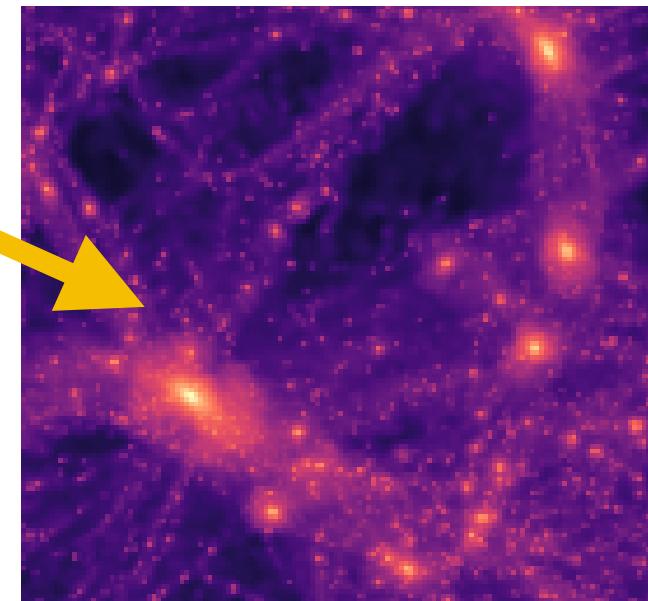
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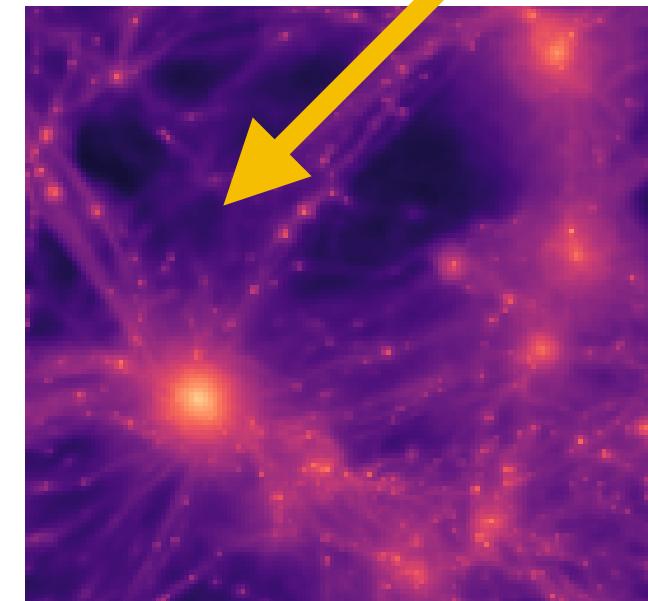
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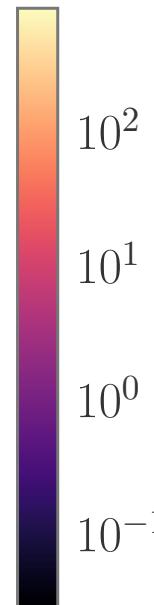


Low-fidelity

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High-fidelity



$\approx 100x$ more expensive!!

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Existing work on multi-fidelity in SBI

Many great works, but which are not specialised for neural-SBI:

Jasra, A., Jo, S., Nott, D., Shoemaker, C., & Tempone, R. (2019). Multilevel Monte Carlo in approximate Bayesian computation. *Stochastic Analysis and Applications*, 37(3), 346–360.

Prescott, T. P., & Baker, R. E. (2020). Multifidelity approximate Bayesian computation. *SIAM-ASA Journal on Uncertainty Quantification*, 8(1), 114–138.

Warne, D. J., Prescott, T. P., Baker, R. E., & Simpson, M. J. (2022). Multifidelity multilevel Monte Carlo to accelerate approximate Bayesian parameter inference for partially observed stochastic processes. *Journal of Computational Physics*, 469, 111543.

Existing work on multi-fidelity in SBI

One very recent attempt, but no theory and critical issue with hyper parameter selection:

Krouglova, A. N., Johnson, H. R., Confavreux, B., Deistler, M., & Gonçalves, P. J. (2025). Multifidelity simulation-based inference for computationally expensive simulators. *arXiv:2502.08416*.

Existing work on multi-fidelity in SBI

→ **Open problem:** Rigorous and theoretically-grounded multi-fidelity for neural SBI!

Neural likelihood estimation (NLE)

- **Step 1:** train $q_{\phi}(\cdot | \theta)$ to approximate the likelihood using samples from the prior ($\theta_1, \dots, \theta_n \sim p(\theta)$) and simulator ($x_i \sim p(\cdot | \theta_i)$):

$$\hat{\phi}_n := \arg \min_{\phi \in \Phi} \ell_{\text{NLE}}(\phi), \quad \ell_{\text{NLE}}(\phi) = -\frac{1}{n} \sum_{i=1}^n \log q_{\phi}(x_i | \theta_i) \approx -\mathbb{E}_{\theta \sim p(\theta)}[\mathbb{E}_{x \sim p(\cdot | \theta)}[\log q_{\phi}(x | \theta)]]$$

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- **Step 2:** Do Bayes with approximate likelihood!

$$p_{\text{NLE}}(\theta | y_1, \dots, y_n) \propto \prod_{i=1}^n q_{\hat{\phi}_n}(y_i | \theta) p(\theta)$$

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Typically the most **computationally expensive** step!!

- **Step 2:** Do Bayes with approximate likelihood!

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A better step 1?

$$\ell_{\text{NLE}}(\phi) = -\frac{1}{n} \sum_{i=1}^n \log q_\phi(x_i | \theta_i) \approx -\mathbb{E}_{\theta \sim p(\theta)}[\mathbb{E}_{x \sim p(\cdot | \theta)}[\log q_\phi(x | \theta)]]$$



Can we do this better/cheaper?!

A better step 1?

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Can we do this better/cheaper?!



Yes, Multilevel Monte Carlo!

Giles, M. B. (2015). Multilevel Monte Carlo methods. *Acta Numerica*, 24, 259–328.

Jasra, A., Law, K., & Suciu, C. (2020). Advanced Multilevel Monte Carlo Methods. *International Statistical Review*, 88(3), 548–579.

Multilevel Monte Carlo

Suppose we have a $f_0, f_1, \dots, f_L = f$ of increasing cost but also increasing accuracy. Then:

$$\mathbb{E}_{z \sim \mu}[f(z)]$$

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Very cheap - can
take n_0 large.

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 Very cheap - can take n_0 large.

 Very expensive - **cannot** take n_l large....
But low variance!

Multilevel NLE

$$-\mathbb{E}_{\theta \sim \pi} \left[\mathbb{E}_{x \sim \mathbb{P}_\theta} \left[\log q_\phi(x | \theta) \right] \right]$$

Multilevel NLE

$$-\mathbb{E}_{\theta \sim \pi} \left[\mathbb{E}_{x \sim \mathbb{P}_\theta} \left[\log q_\phi(x | \theta) \right] \right] = \mathbb{E}_{\theta \sim \pi, u \sim \mathbb{U}} \left[-\log q_\phi(G_\theta(u) | \theta) \right]$$

Change of measure



Multilevel NLE

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 This is now a joint expectation in the prior and \mathbb{U} !

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We can directly apply MLMC to it, where intermediate integrands are of the form:

$$f_\phi^l(\theta, u) = -\log q_\phi(G_\theta^l(u) | \theta)$$

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Multilevel neural SBI

Our ‘data’ is therefore:

$$\left\{ \theta_i^l, u_i^l, G_{\theta_i^l}^l(u_i^l), G_{\theta_i^l}^{l-1}(u_i^l) \right\} \quad \text{where} \quad \theta_i^l \sim \pi, u_i^l \sim \mathbb{U},$$

Multilevel neural SBI

Our ‘data’ is therefore:

$$\left\{ \theta_i^l, u_i^l, G_{\theta_i^l}^l(u_i^l), G_{\theta_i^l}^{l-1}(u_i^l) \right\} \quad \text{where} \quad \theta_i^l \sim \pi, u_i^l \sim \mathbb{U},$$


Seed-matched!

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Our objective for step 1 is:

$$\ell_{\text{ML-NLE}}(\phi) := \frac{1}{n_0} \sum_{i=1}^{n_0} f_\phi^0(u_i^0, \theta_i^0) + \sum_{l=1}^L \frac{1}{n_l} \sum_{i=1}^{n_l} \left(f_\phi^l(u_i^l, \theta_i^l) - f_\phi^{l-1}(u_i^l, \theta_i^l) \right)$$

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Note that we presented this for NLE, but the same could work for NPE, other scoring rules, etc...!

Challenges with training

$$\ell_{\text{ML-NLE}}(\phi) := \frac{1}{n_0} \sum_{i=1}^{n_0} f_\phi^0(u_i^0, \theta_i^0) + \frac{1}{n_1} \sum_{i=1}^{n_1} \left(f_\phi^1(u_i^1, \theta_i^1) - f_\phi^0(u_i^1, \theta_i^1) \right)$$

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$$\frac{1}{n_0} \sum_{i=1}^{n_0} \nabla f_\phi^0(u_i^0, \theta_i^0) \approx \mathbb{E}[\nabla f_\phi^0]$$
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Contradictory gradients! This is a problem when we are close to stationarity and n_0/n_1 are small... The variance of the negative term is always large!!

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Contradictory gradients! This is a problem when we are close to stationarity and n_0/n_1 are small... The variance of the negative term is always large!!

We fix the issue by normalising gradients so that these two terms have the same magnitude, and by projecting onto each other's normal planes, which stabilises training.

Bound on the variance

Under some mild assumptions, we get:

$$\text{Var} [\ell_{\text{ML-NLE}}(\phi)] \leq \frac{K_0(\phi)}{n_0} \left(\|G^0\|_{W^{1,4}(\pi \times \mathbb{U})}^4 + 1 \right) + \sum_{l=1}^L \frac{K_l(\phi)}{n_l} \left(\|G^l - G^{l-1}\|_{W^{1,4}(\pi \times \mathbb{U})}^2 + 1 \right)$$

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Large!

Small!

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The diagram consists of four yellow arrows pointing upwards from text labels to specific terms in the equation. The first arrow points from 'Large!' to the term $\frac{K_0(\phi)}{n_0}$. The second arrow points from 'Complexity of low-fidelity generator - large!' to the term $\|G^0\|_{W^{1,4}(\pi \times \mathbb{U})}^4$. The third arrow points from 'Small!' to the term $\frac{K_l(\phi)}{n_l}$. The fourth arrow points from 'Complexity of other integrands - small!' to the term $\|G^l - G^{l-1}\|_{W^{1,4}(\pi \times \mathbb{U})}^2$.

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Bound on the variance

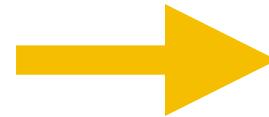
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- 3) The surrogate $q_\phi(\cdot | \theta)$ has a Lipschitz gradient locally, and does not blow up too fast.

Bound on the variance



Can use this to determine optimal samples per level!

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 Large!
 Complexity of low-fidelity generator - large!
 Small!
 Complexity of other integrands - small!

Assumptions:

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Simulations per level

Under some mild regularity conditions, we can find the optimal number of simulations per level assuming we have a maximum computational budget of C_{budget} :

$$n_0^\star \propto \frac{C_{\text{budget}}}{\sqrt{C_0}} \sqrt{\|G^0\|_{W^{1,4}(\pi \times \mathbb{U})}^4 + 1}, \quad n_l^\star \propto \frac{C_{\text{budget}}}{\sqrt{C_l + C_{l+1}}} \sqrt{\|G^l - G^{l-1}\|_{W^{1,4}(\pi \times \mathbb{U})}^2 + 1}.$$

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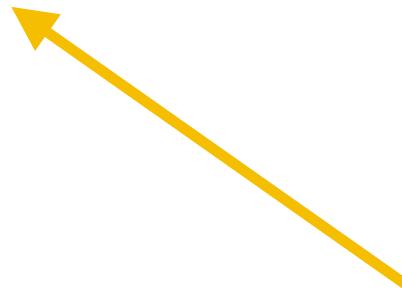
The more ‘complex’ the generator
(or the difference in generators),
the more simulations we need.

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The larger the cost of simulations at this level, the less simulations we can afford.



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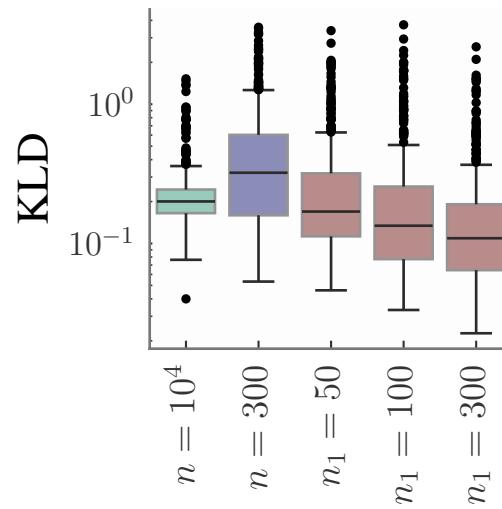
Note that these expressions contain a lot of quantities we may not know a-priori, but it is still indicative and helpful for selecting which simulations to run in practice.

G-and-k distribution

$$G_{\theta}^l(u) = \theta_1 + \theta_2 \left(1 + 0.8 \left(\frac{1 - \exp(-\theta_3 z_l(u))}{1 + \exp(-\theta_3 z_l(u))} \right) \right) \left(1 + z_l(u)^2 \right)^{\log(\theta_4)} z_l(u),$$

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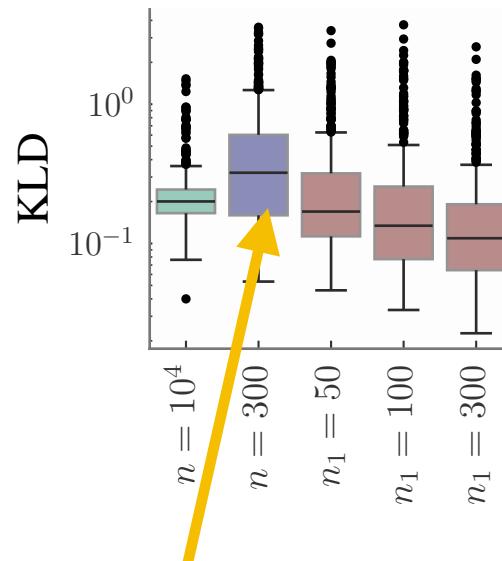


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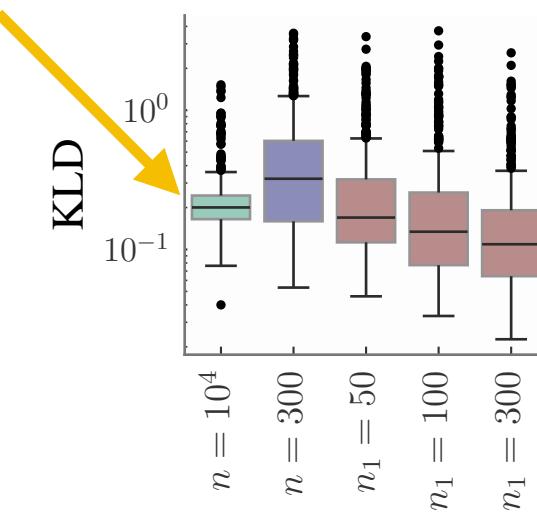
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High-fidelity only:
too few simulations!

G-and-k distribution

Low-fidelity only:
Many simulations,
but low quality!



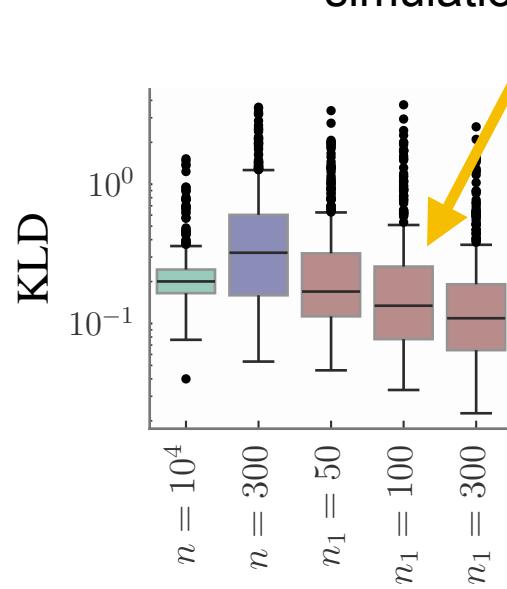
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G-and-k distribution

ML-NLE: both many simulations and high quality!

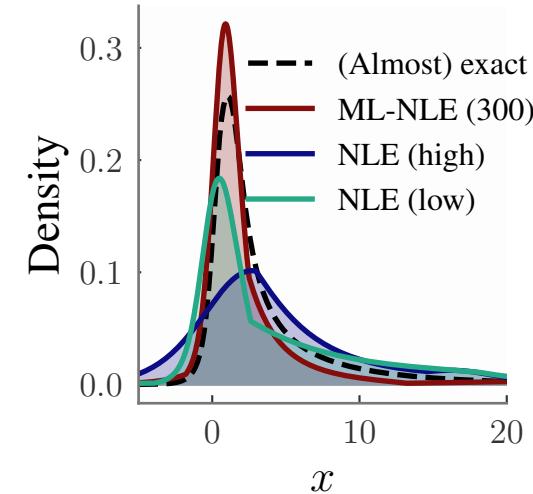
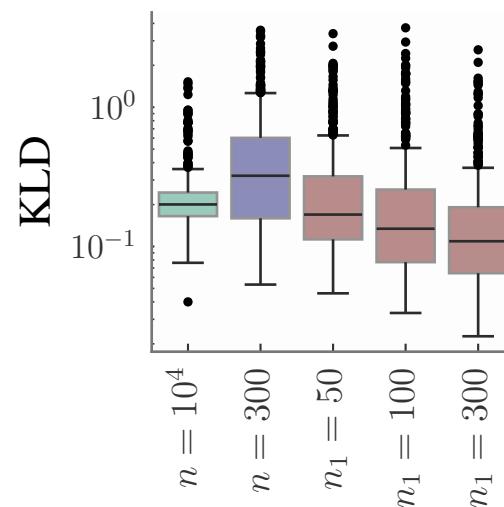


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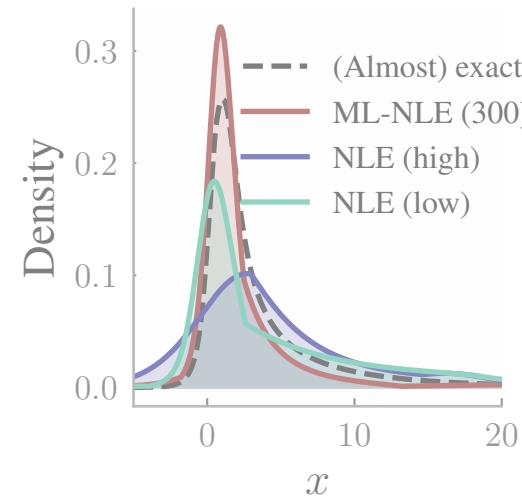
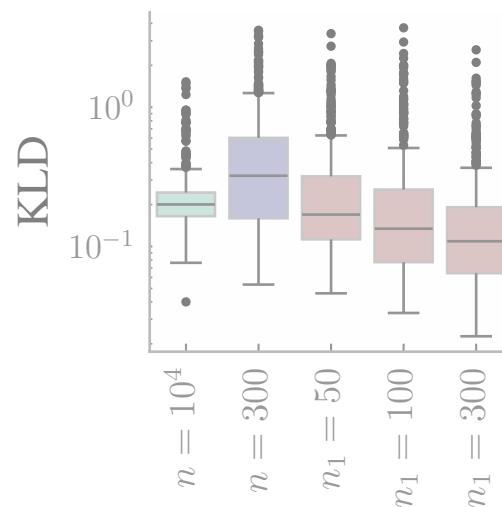
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G-and-k distribution



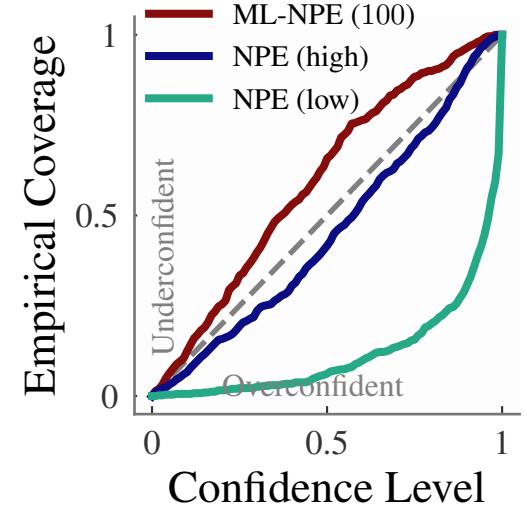
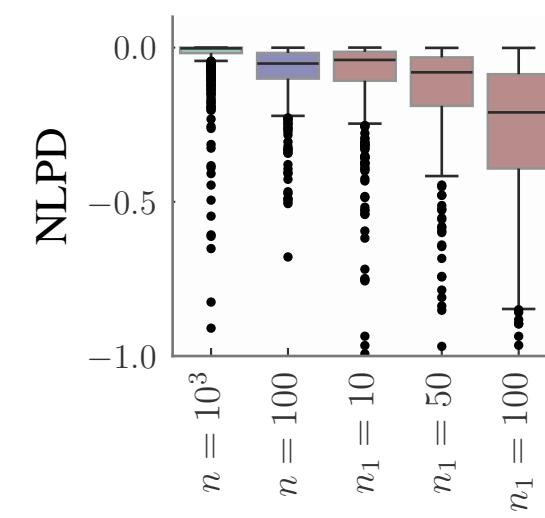
G-and-k distribution



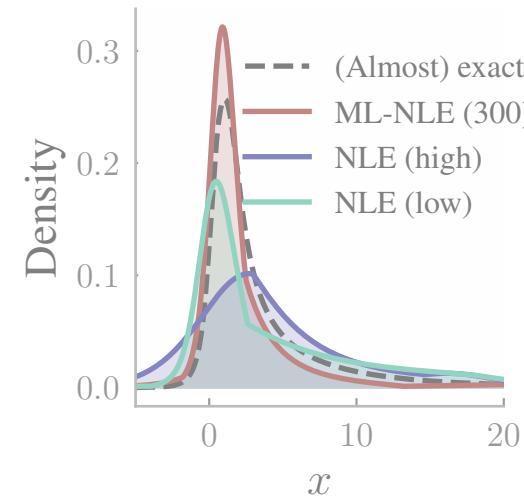
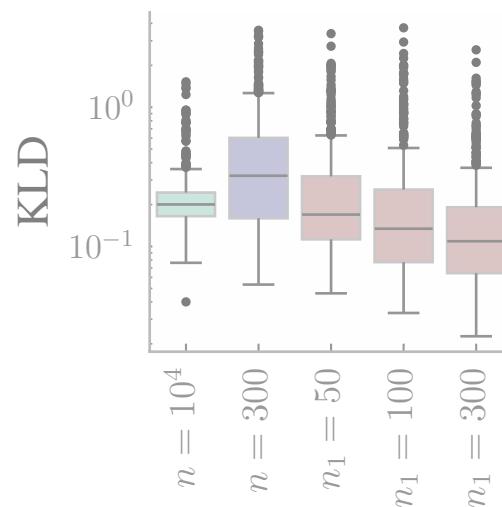
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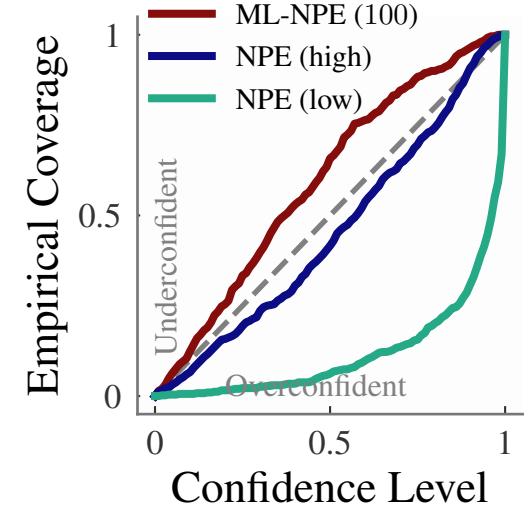
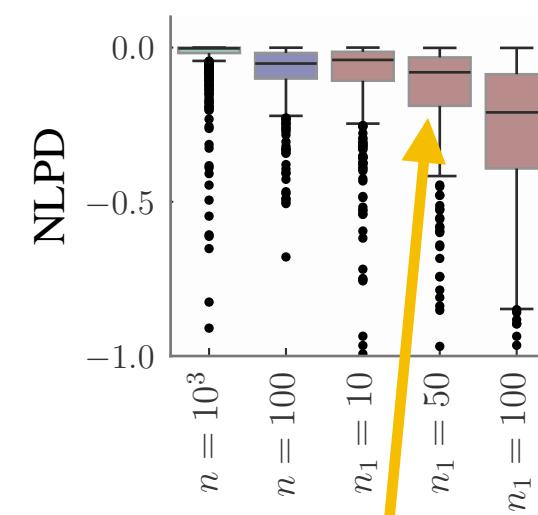
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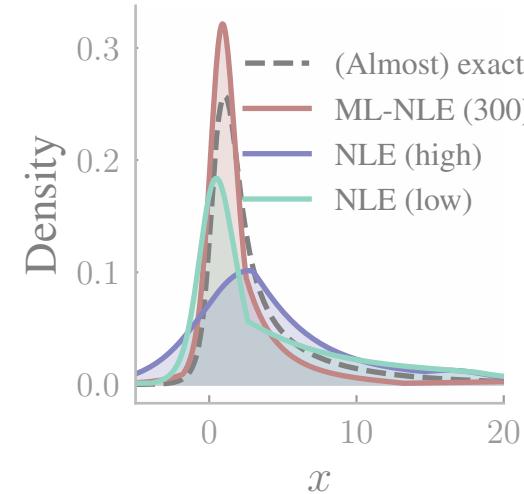
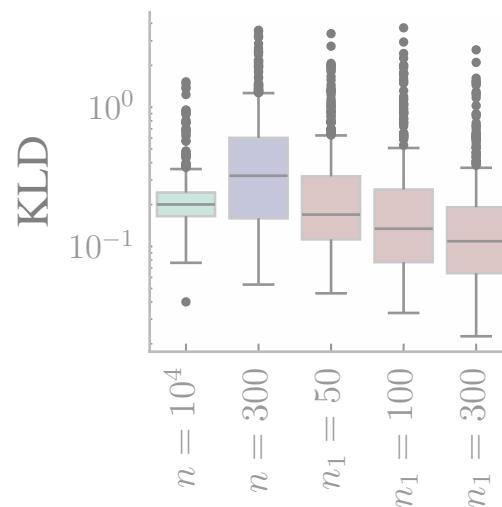
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ML-NPE: Similar conclusion!

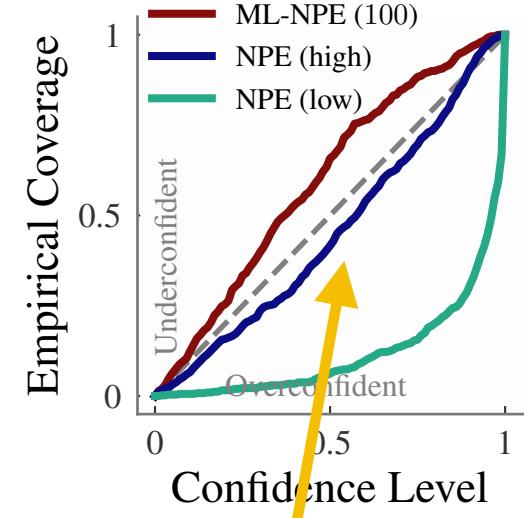
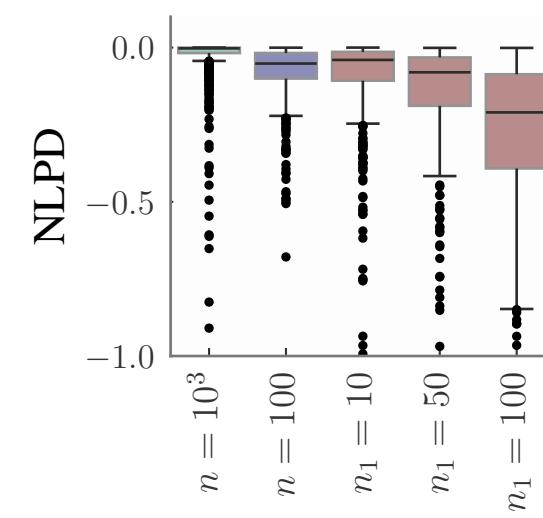
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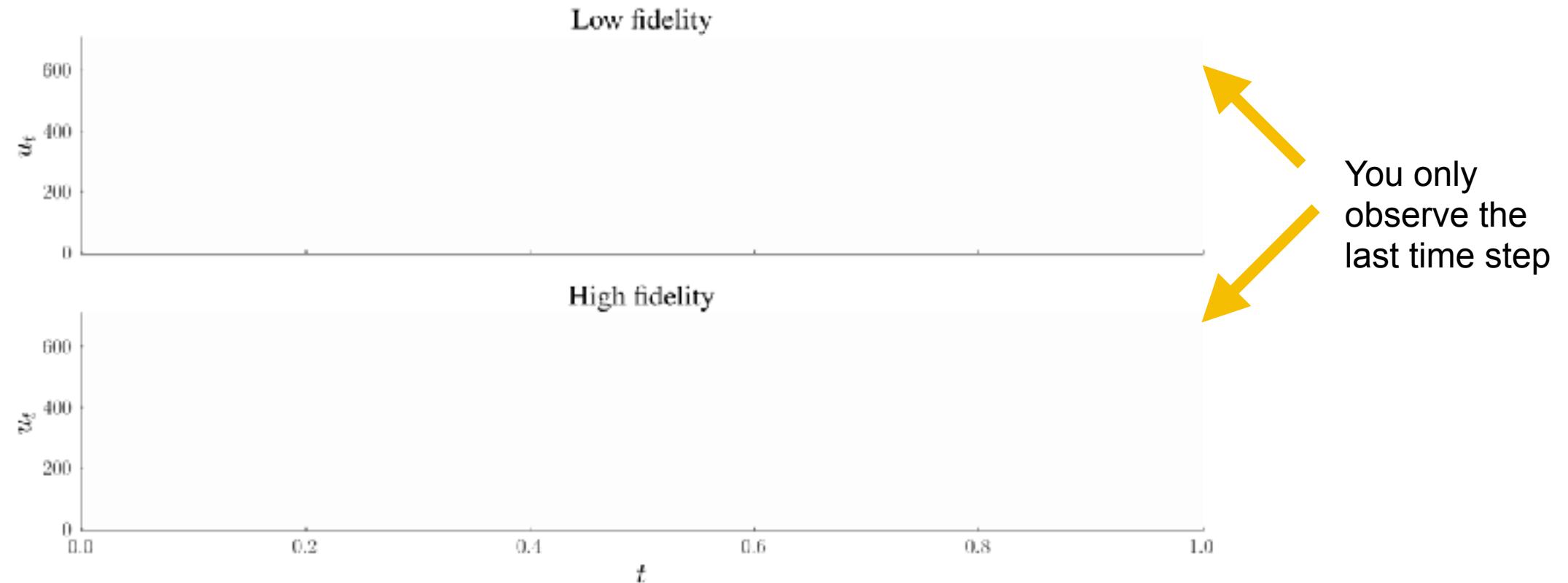
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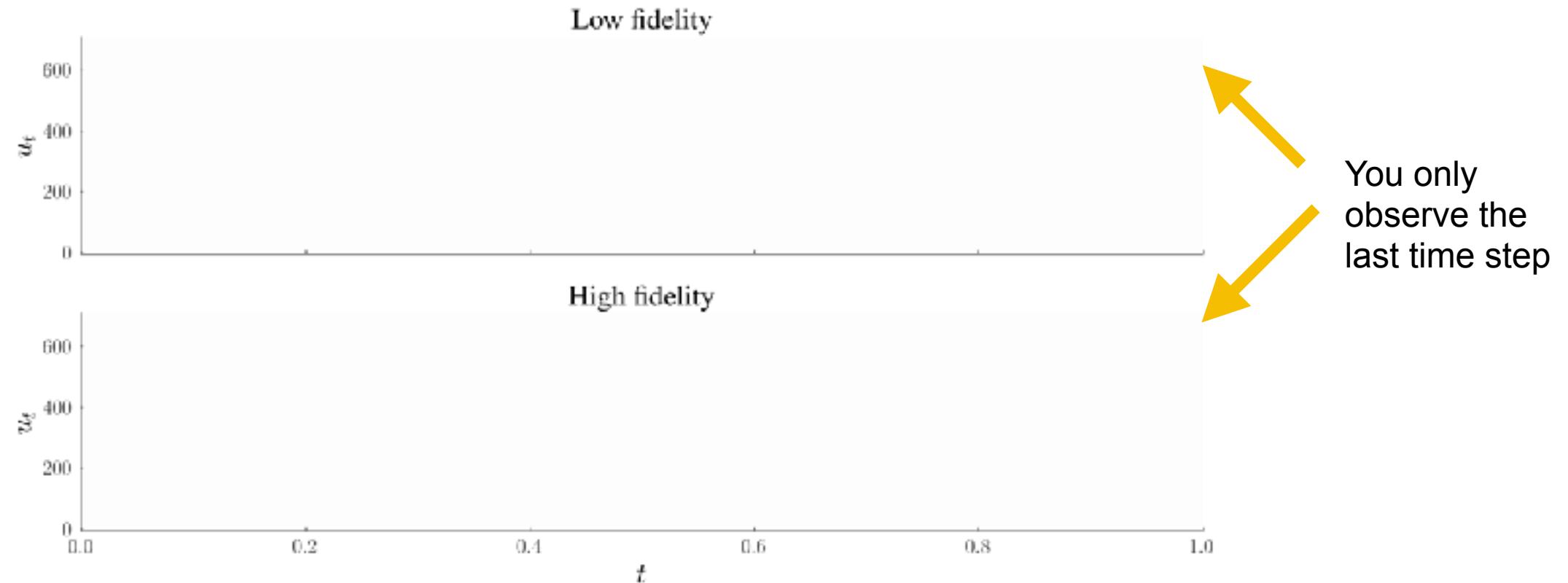
Coverage slightly cautious

Toggle-switch models for genes ($d=1$, $p=7$)



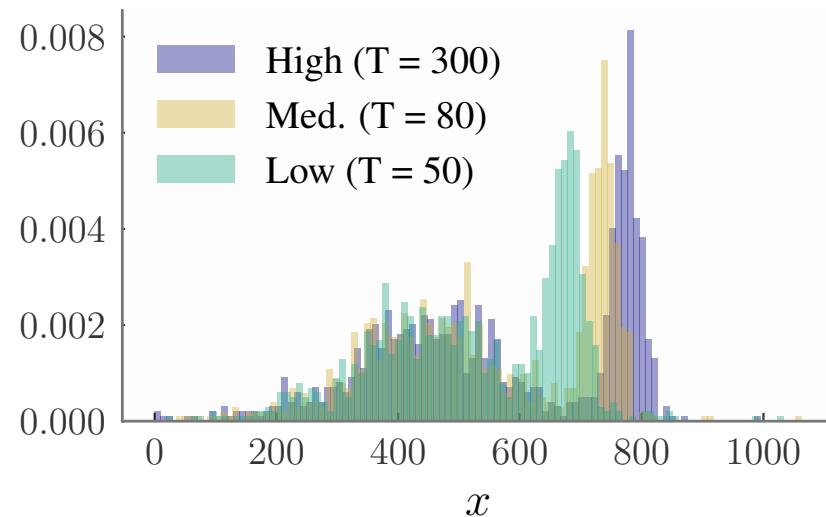
Bonassi, F. V., You, L., & West, M. (2011). Bayesian learning from marginal data in bionetwork models. *Statistical Applications in Genetics and Molecular Biology*, 10(1).

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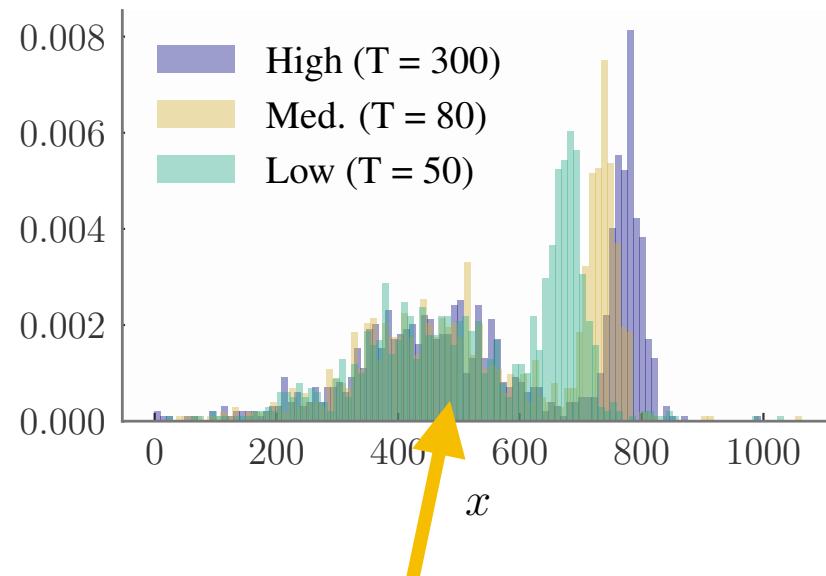
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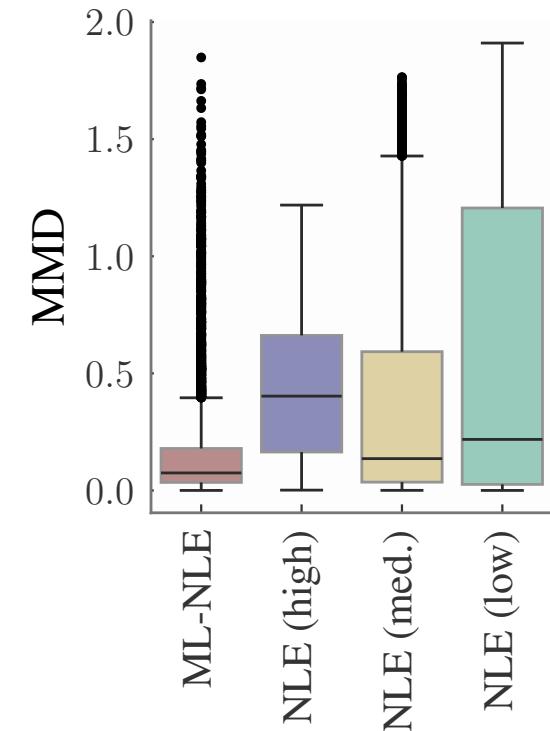
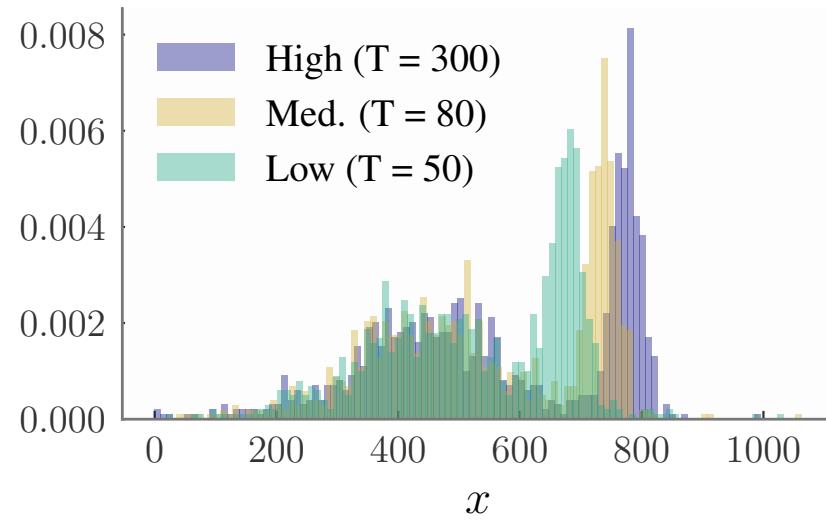
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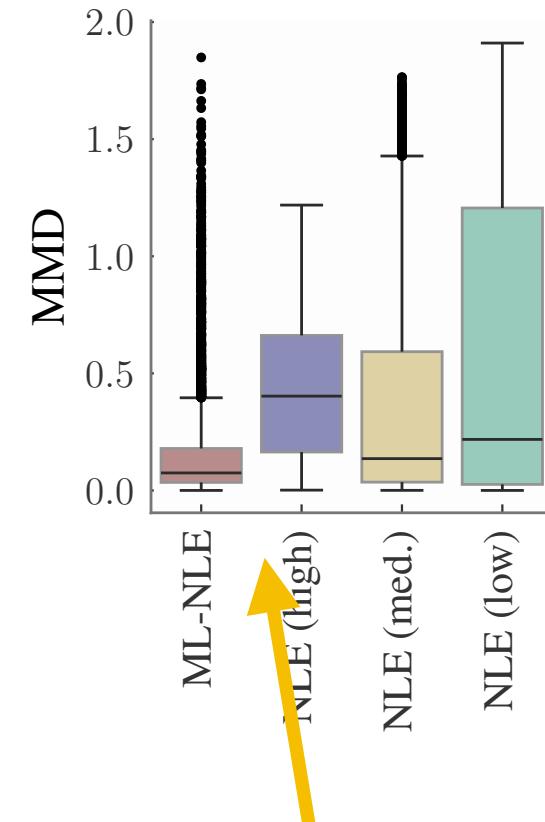
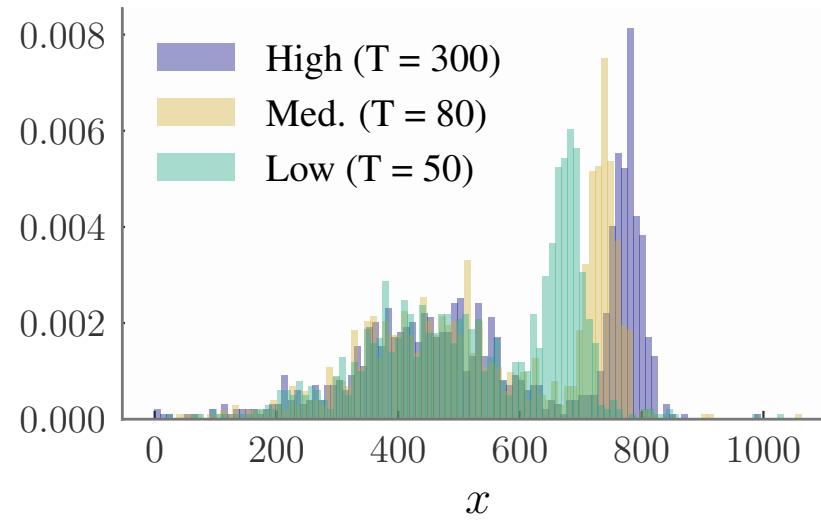
Observations bi-modal, with second mode only
well approximated for high-fidelity levels

Toggle-switch models for genes (d=1, p=7)



Bonassi, F. V., You, L., & West, M. (2011). Bayesian learning from marginal data in bionetwork models. *Statistical Applications in Genetics and Molecular Biology*, 10(1).

Toggle-switch models for genes (d=1, p=7)

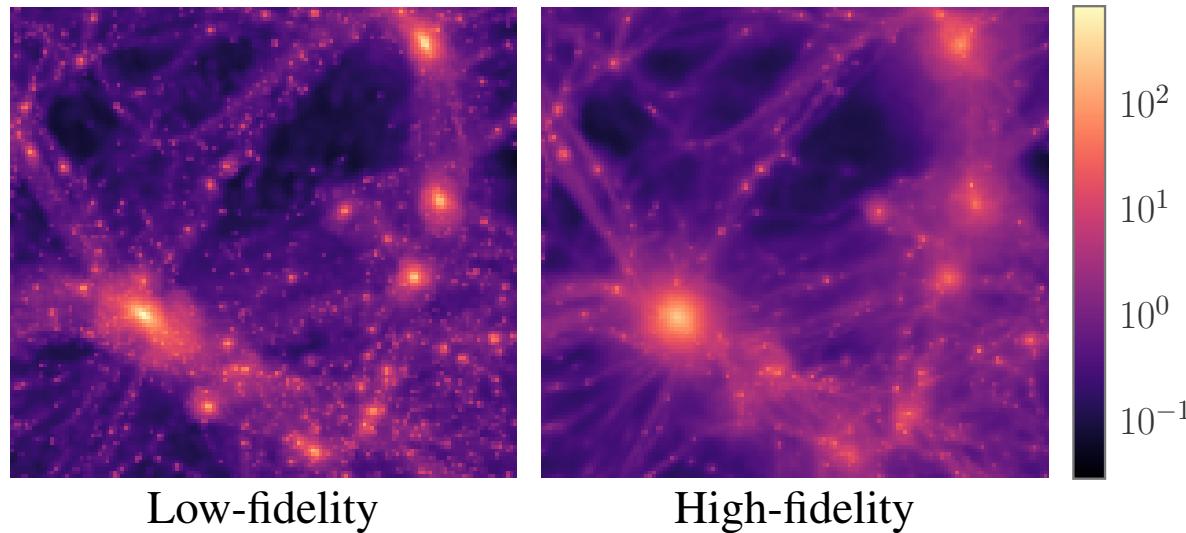


$$\begin{aligned} n_0 &= 10000 \\ n_1 &= 500 \\ n_2 &= 300 \end{aligned}$$

Bonassi, F. V., You, L., & West, M. (2011). Bayesian learning from marginal data in bionetwork models. *Statistical Applications in Genetics and Molecular Biology*, 10(1).

ML-NLE benefits from low-fidelity simulations for first mode but also from high-fidelity simulations for second mode

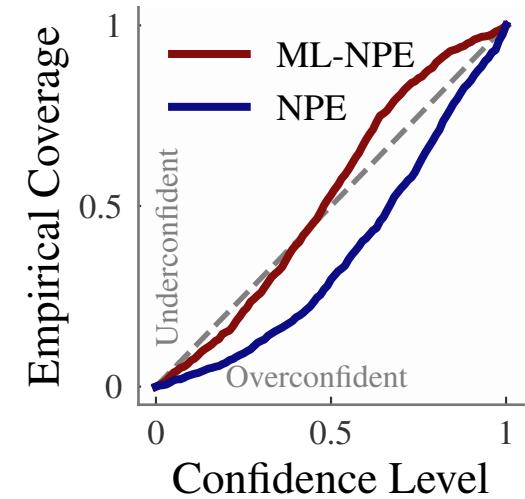
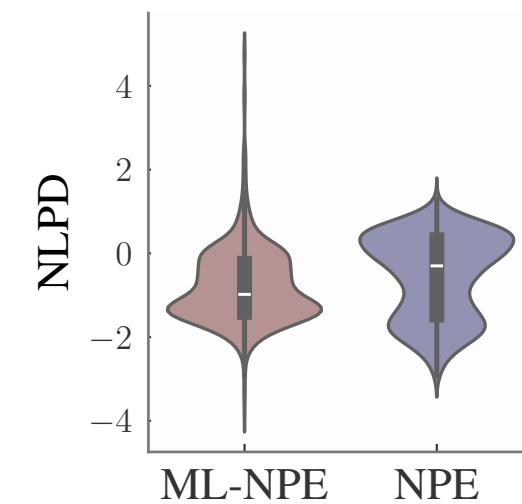
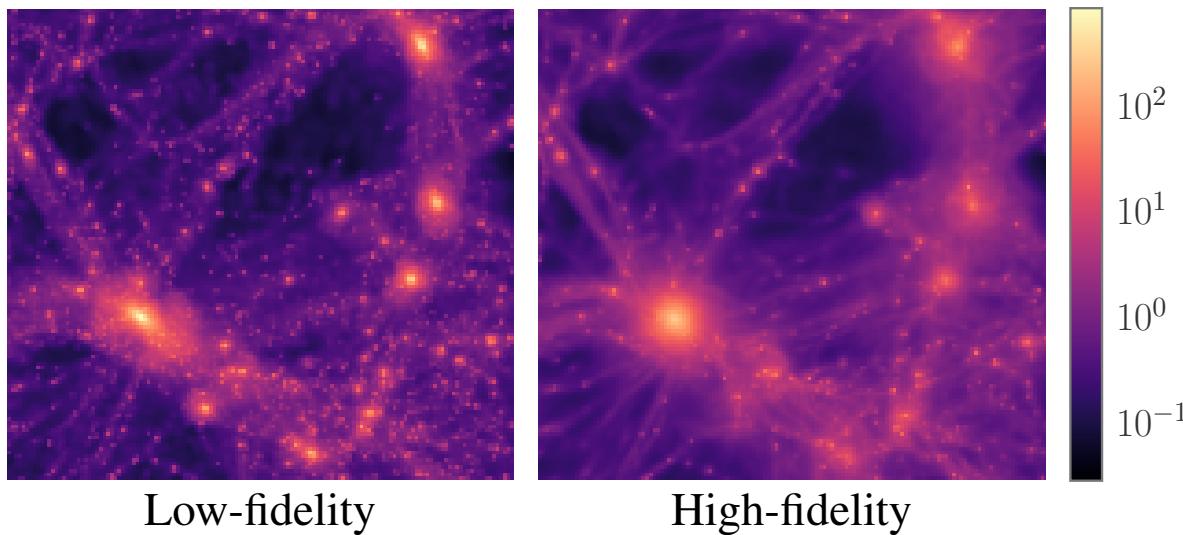
Back to cosmology.... (d=39, p=1)



NPE: $n = 20$ (all high fidelity!)

ML-NPE: $n_0 = 20, n_1 = 980$

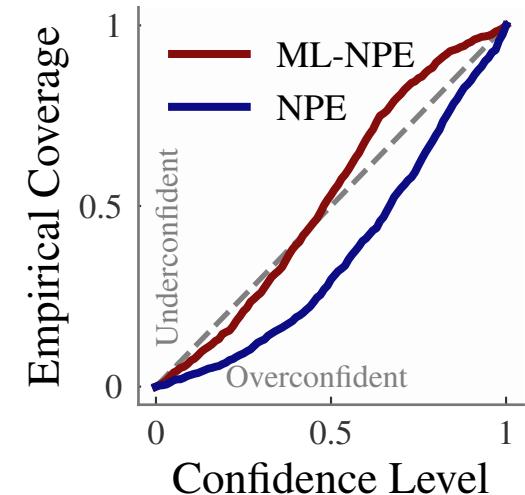
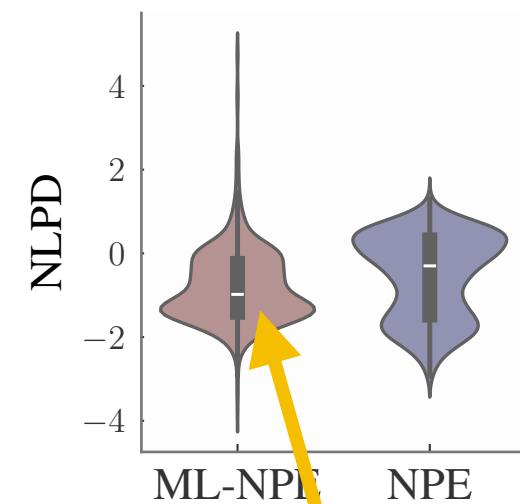
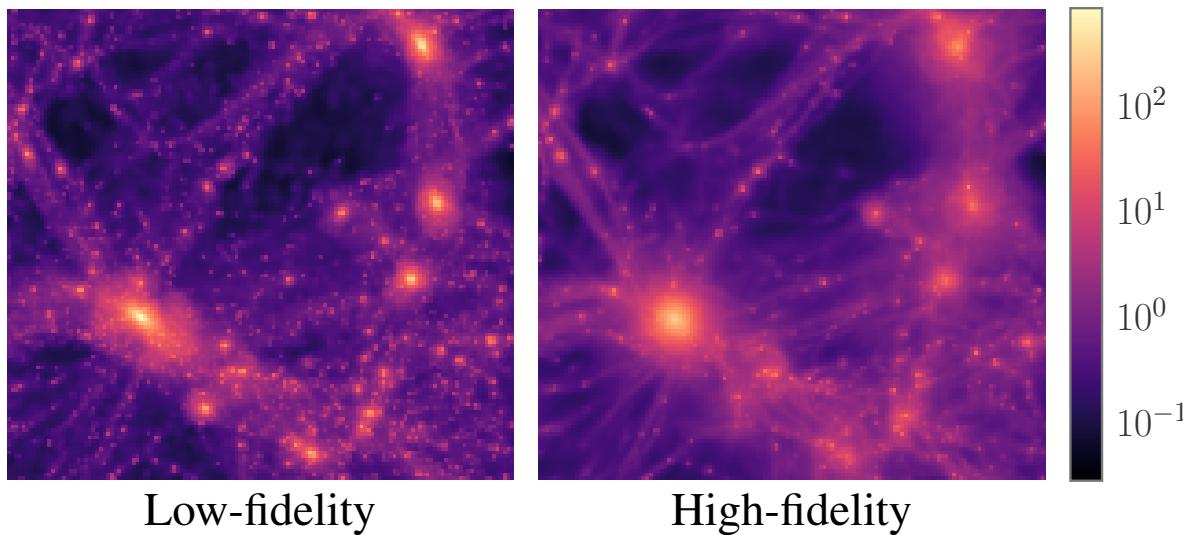
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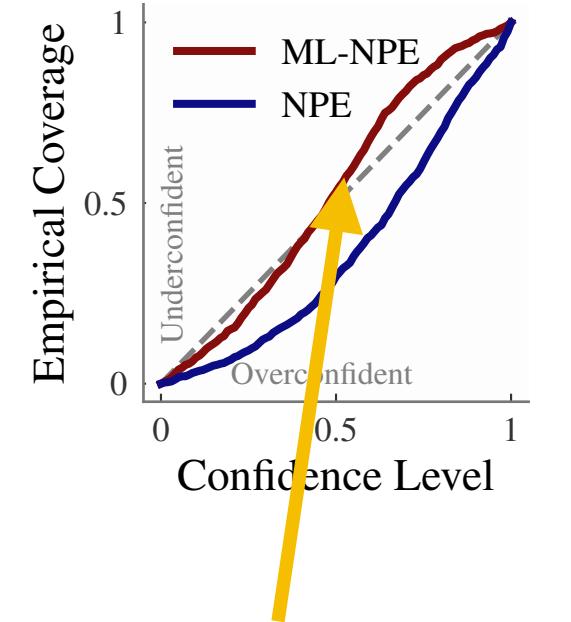
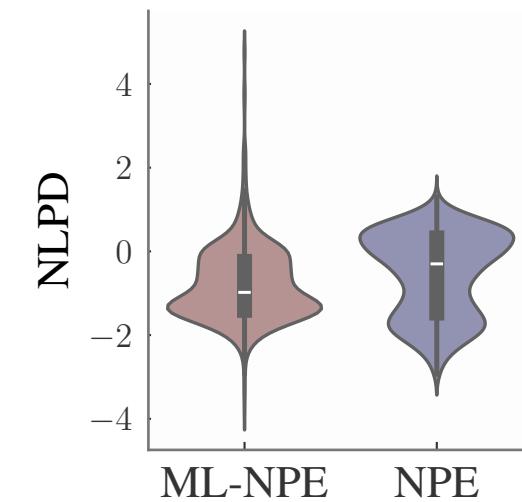
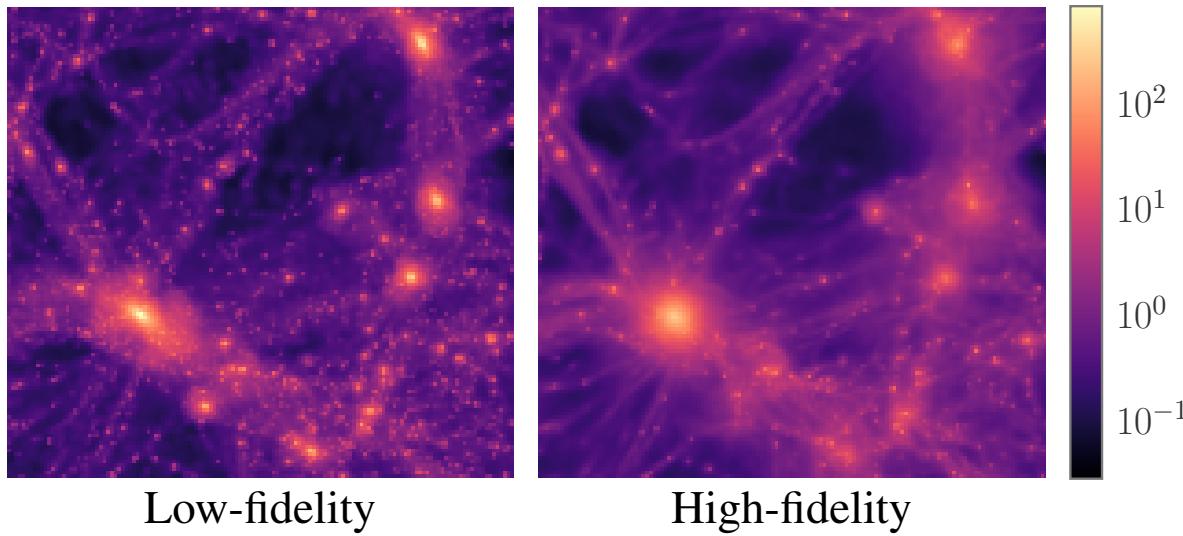
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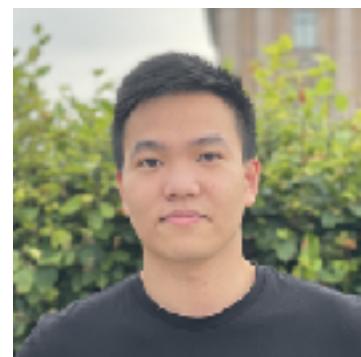
Improved calibration!

Any Questions?

Paper: Hikida, Y., Bharti, A., Jeffrey, N. & Briol, F-X (2025). Multilevel neural simulation-based inference. arXiv:2506.06087 (to appear at NeurIPS?).

Code: <https://github.com/yugahikida/multilevel-sbi>

Cost-aware simulation-based inference



Paper: Bharti, A., Huang, D., Kaski, S., & **Briol, F.-X.** (2025). Cost-aware simulation-based inference. International Conference on Artificial Intelligence and Statistics, 28–36.

Code: <https://github.com/huangdaolang/cost-aware-sbi>

Challenge for SBI

Simulators can be really computationally expensive!

Challenge for SBI

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However we may not have an easy way to obtain low-fidelity simulators....

Challenge for SBI

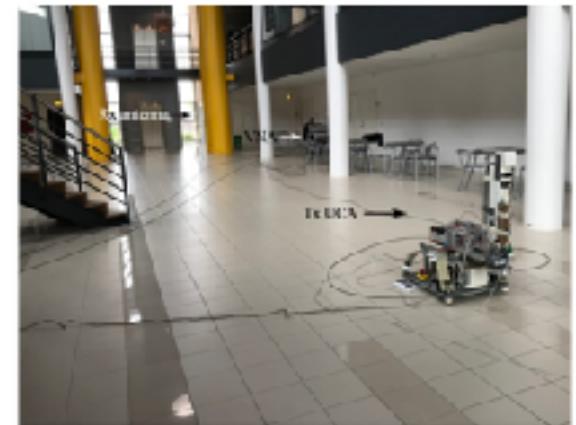
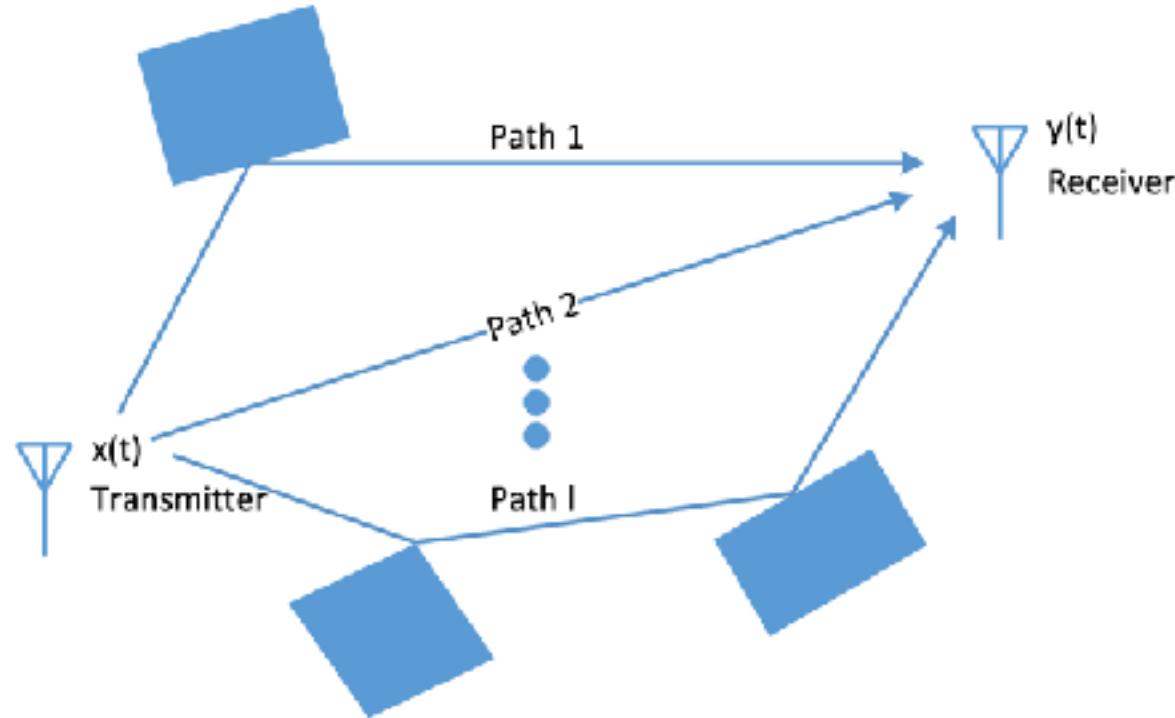
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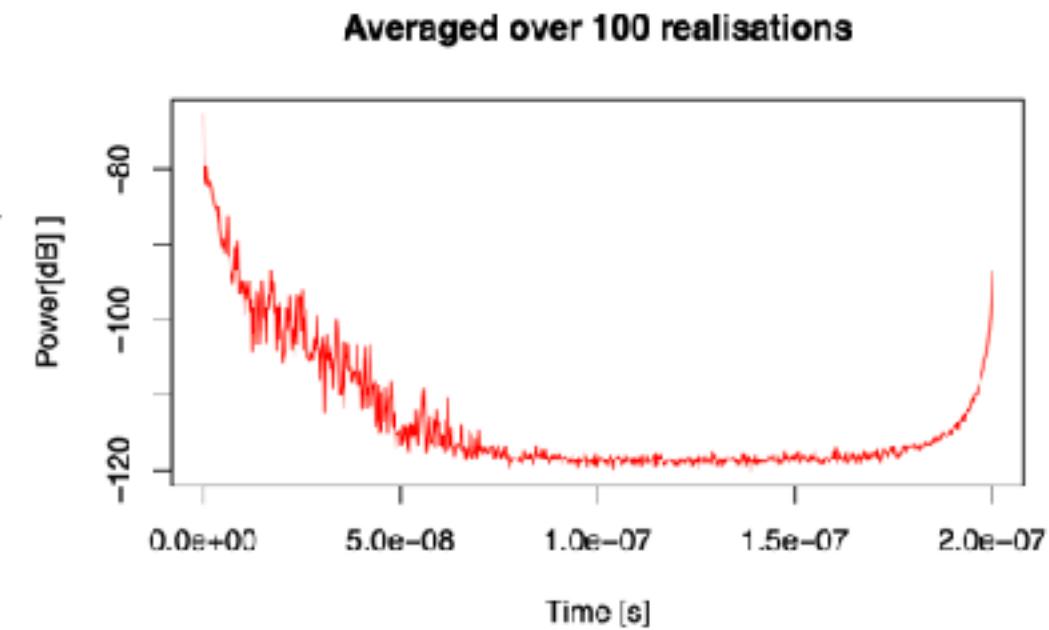
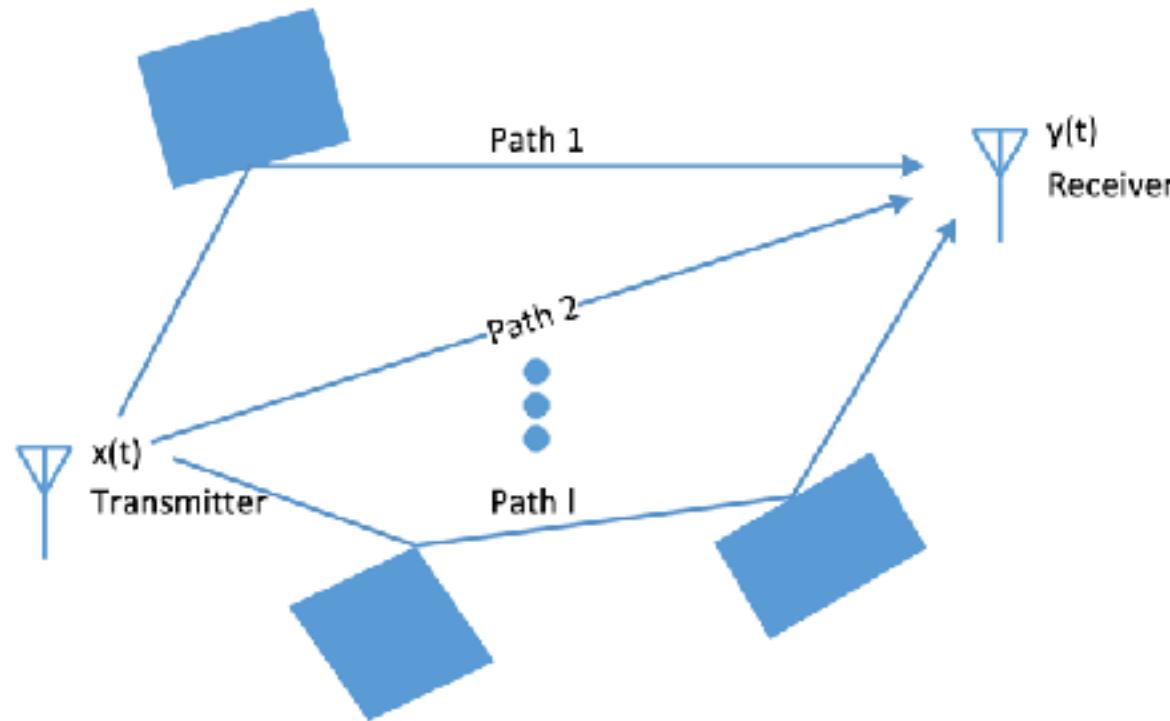
We can adjust our sampling to sample less often from expensive parameterisations!

SBI for radio-propagation



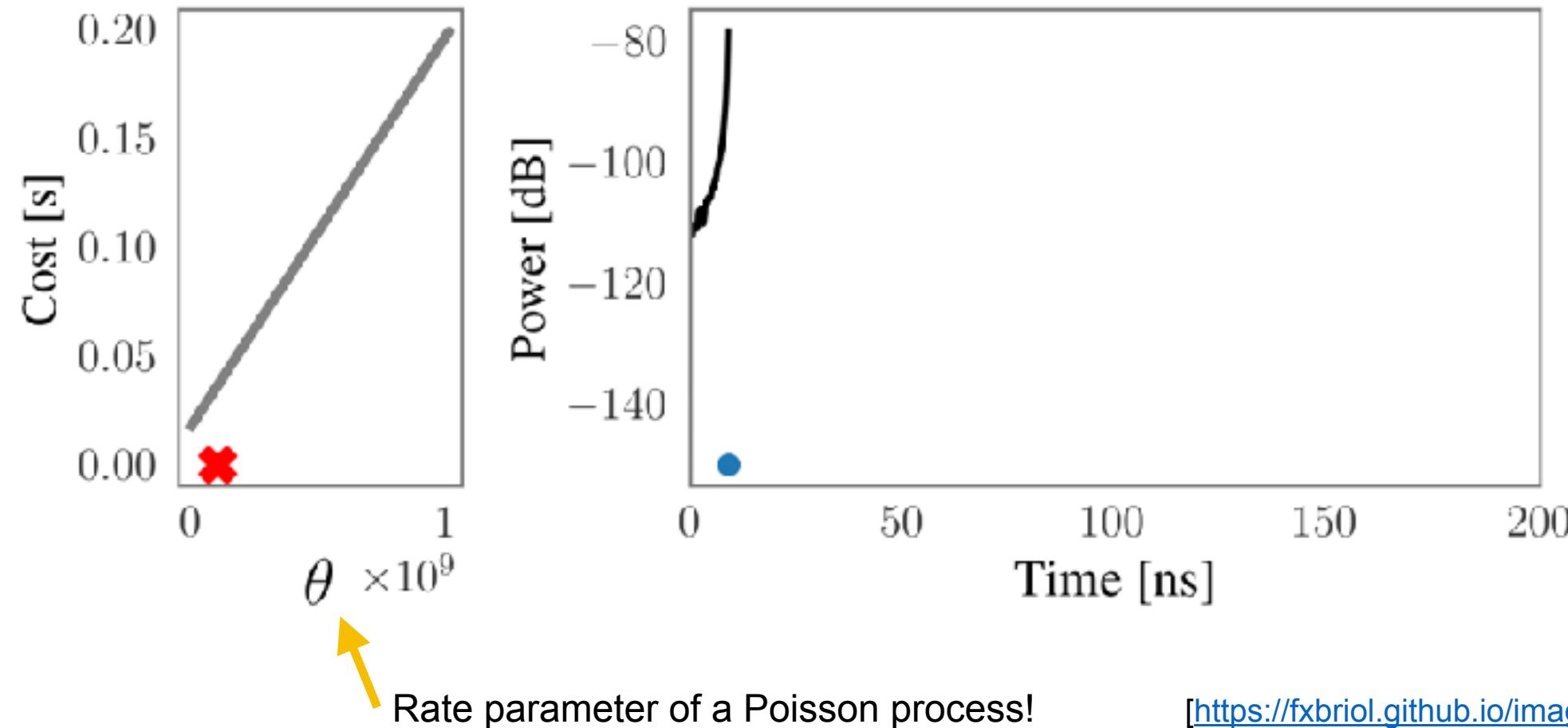
Bharti, A., **Briol, F-X.**, Pedersen, T. (2022). A general method for calibrating stochastic radio channel models with kernels. *IEEE Transactions on Antennas and Propagation*, vol. 70, no. 6, pp. 3986-4001, June 2022.

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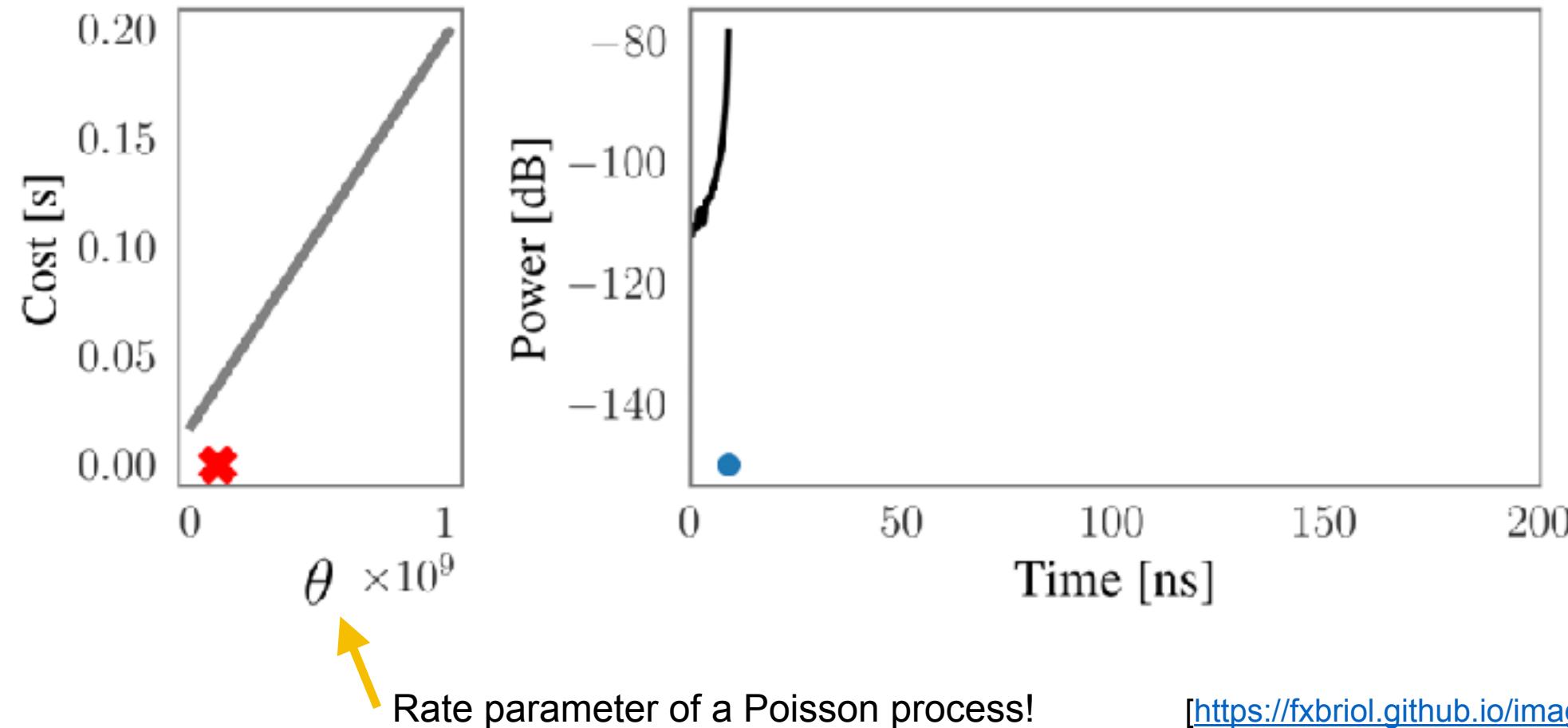


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The cost of simulations is not constant...



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Neural likelihood estimation (NLE)

- **Step 1:** train a conditional density model $q_\phi(\cdot | \theta)$ to approximate the likelihood using samples from the prior ($\theta_1, \dots, \theta_n \sim p(\theta)$) and simulator ($x_i \sim p(\cdot | \theta_i)$):

$$\hat{\phi}_n := \arg \min_{\phi \in \Phi} \ell_{\text{NLE}}(\phi), \quad \ell_{\text{NLE}}(\phi) = -\frac{1}{n} \sum_{i=1}^n \log q_\phi(x_i | \theta_i) \approx -\mathbb{E}_{\theta \sim p(\theta)}[\mathbb{E}_{x \sim p(x|\theta)}[\log q_\phi(x | \theta)]]$$

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- **Step 2:** Do Bayes with approximate likelihood!

$$p_{\text{NLE}}(\theta | y_1, \dots, y_n) \propto \prod_{i=1}^n q_{\hat{\phi}_n}(y_i | \theta) p(\theta)$$

A cheaper step 1?

$$\ell_{\text{NLE}}(\phi) = -\frac{1}{n} \sum_{i=1}^n \log q_\phi(x_i | \theta_i) \approx -\mathbb{E}_{\theta \sim p(\theta)}[\mathbb{E}_{x \sim p(\cdot | \theta)}[\log q_\phi(x | \theta)]]$$



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Can we do this better/cheaper?!

- Idea:**
- Let's make use of the cost function $c : \Theta \rightarrow \mathbb{R}$.
 - We can try to sample less often in expensive regions
 - but we still want to target the right objective.

Importance sampling

$$\mu = \int_{\Theta} f(\theta) \pi(\theta) d\theta$$

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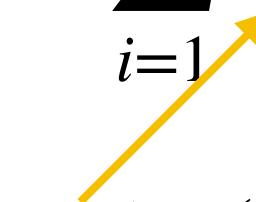
$$\approx \sum_{i=1}^N w(\theta_i) f(\theta_i) \quad \theta_1, \dots, \theta_N \sim \tilde{\pi}$$

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$$w_{\text{IS}}(\theta_i) = \frac{1}{N} \frac{\pi(\theta_i)}{\tilde{\pi}(\theta_i)}$$



Importance sampling

$$\begin{aligned}\mu &= \int_{\Theta} f(\theta) \pi(\theta) d\theta = \int_{\Theta} f(\theta) \frac{\pi(\theta)}{\tilde{\pi}(\theta)} \tilde{\pi}(\theta) d\theta \\ &\approx \sum_{i=1}^N w(\theta_i) f(\theta_i) \quad \theta_1, \dots, \theta_N \sim \tilde{\pi} \\ w_{\text{IS}}(\theta_i) &= \frac{1}{N} \frac{\pi(\theta_i)}{\tilde{\pi}(\theta_i)} \quad \text{w}_{\text{SNIS}}(\theta_i) = \frac{w_{\text{IS}}(\theta_i)}{\sum_{j=1}^N w_{\text{IS}}(\theta_j)}\end{aligned}$$

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Question: How do you pick the importance distribution?

Cost-aware importance sampling

$$\tilde{\pi}_g(\theta) \propto \frac{\pi(\theta)}{g(c(\theta))},$$

Cost-aware importance sampling

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We want a distribution similar to
our target π

Cost-aware importance sampling

$$\tilde{\pi}_g(\theta) \propto \frac{\pi(\theta)}{g(c(\theta))},$$

We do not want to sample often where the cost is large!

Cost-aware importance sampling

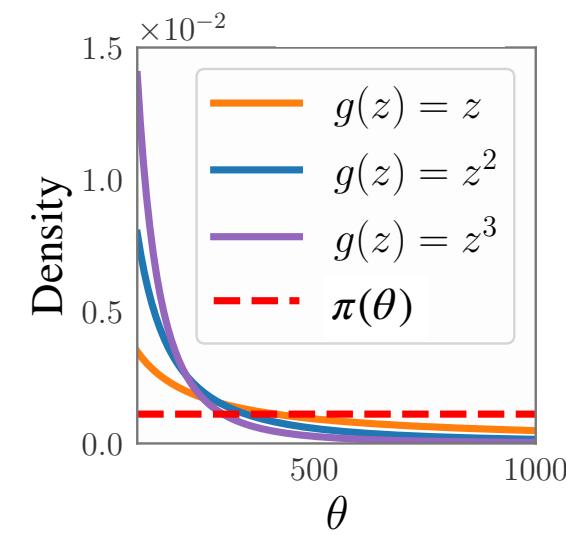
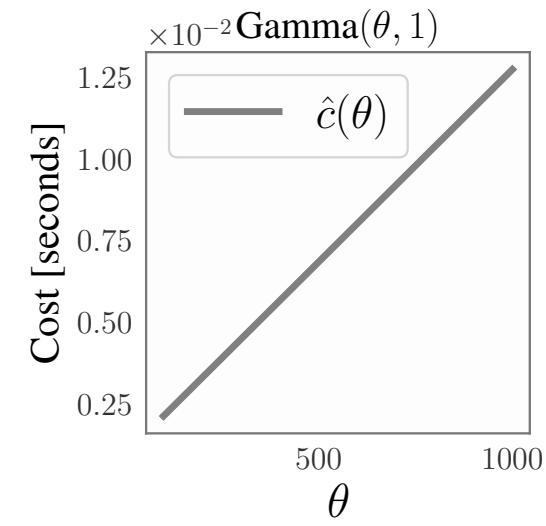
$$\tilde{\pi}_g(\theta) \propto \frac{\pi(\theta)}{g(c(\theta))},$$

$g : (0, \infty) \rightarrow (0, \infty)$ taken to
be non-decreasing.

Represents how much we
dislike ‘expensive’ parameters!

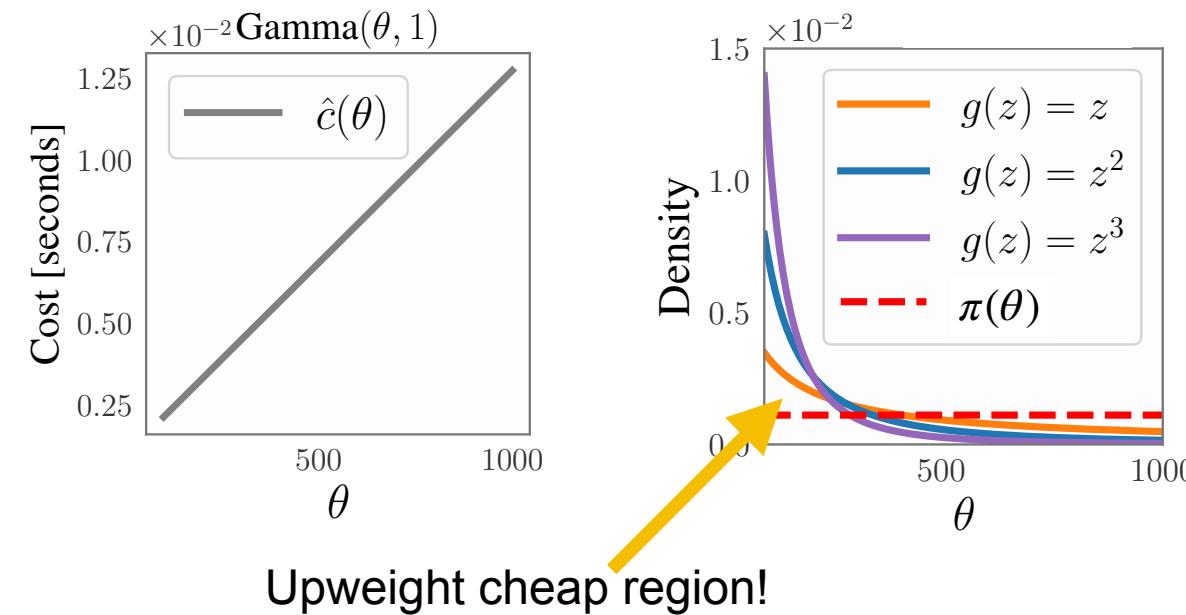
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Cost-aware importance sampling

$$w(\theta) = \frac{1}{N} \frac{\pi(\theta)}{\tilde{\pi}_g(\theta)} = \frac{B\pi(\theta)g(c(\theta))}{N\pi(\theta)} \propto g(c(\theta))$$



Through $\tilde{\pi}_g$, we sample less often from expensive regions, so we need to up-weight expensive samples.

Cost-aware importance sampling

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$$w_{\text{Ca}}(\theta_i) = \frac{w(\theta_i)}{\sum_{j=1}^n w(\theta_j)} = \frac{g(c(\theta_i))}{\sum_{j=1}^n g(c(\theta_j))}$$



We use SNIS weights

$$\mu = \int_{\Theta} f(\theta)\pi(\theta)d\theta \approx \sum_{i=1}^n w_{\text{Ca}}(\theta_i)f(\theta_i) = \hat{\mu}_n^{\text{Ca}}$$

Sampling from the cost-aware proposal

- We can use rejection sampling!

Sampling from the cost-aware proposal

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Repeat until n samples are accepted:

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2. Accept θ^* as a sample from $\tilde{\pi}_g$ with probability $A(\theta)$.

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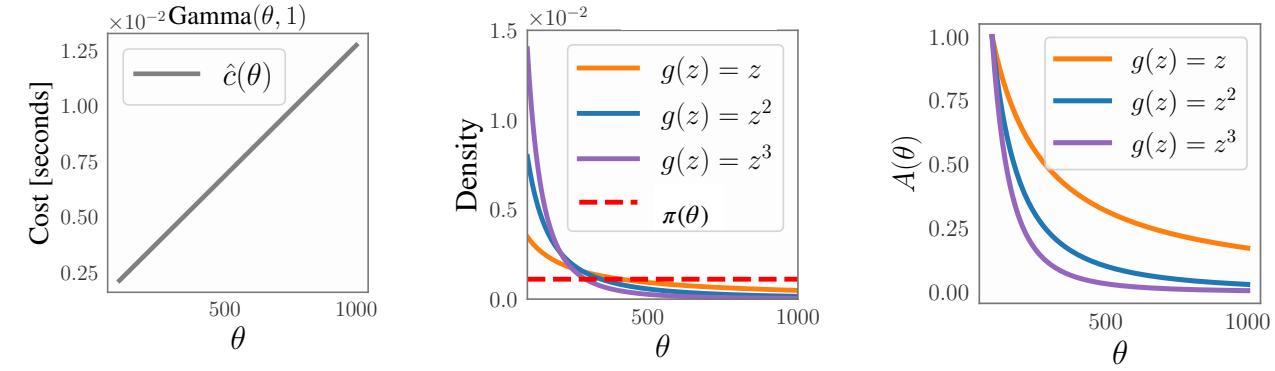
- $\tilde{\pi}_g$ is a density.
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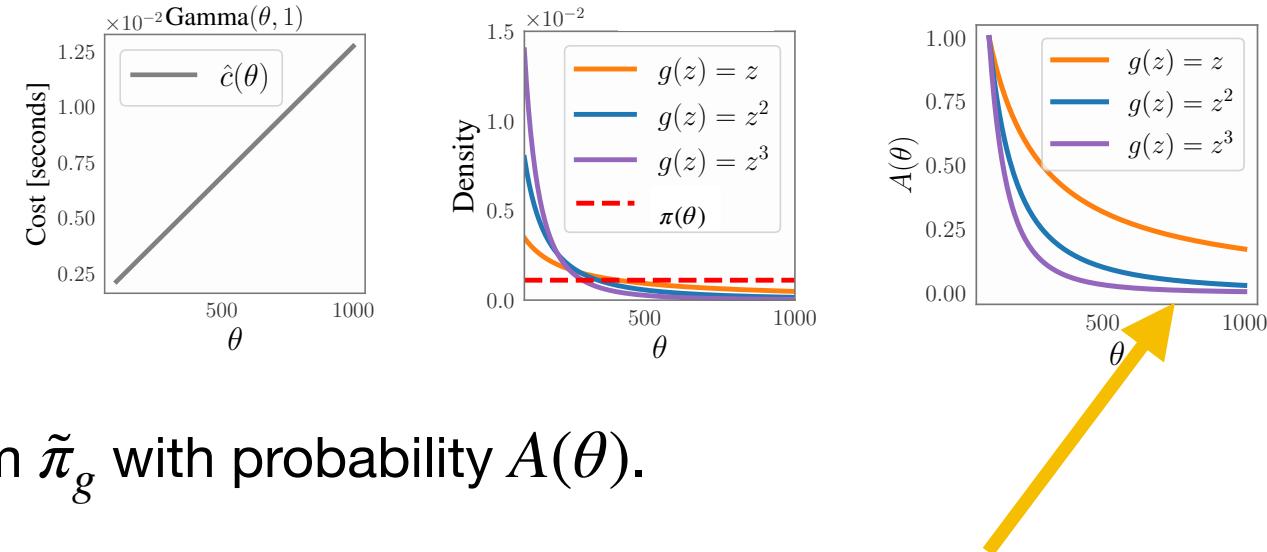
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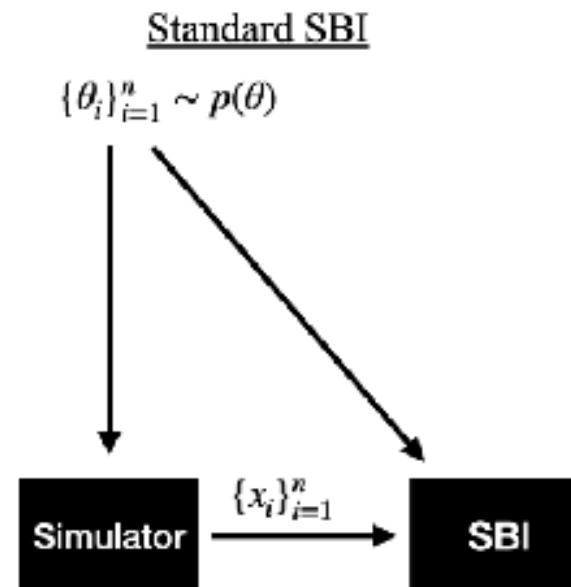
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Being cost-averse decreases acceptance prob!

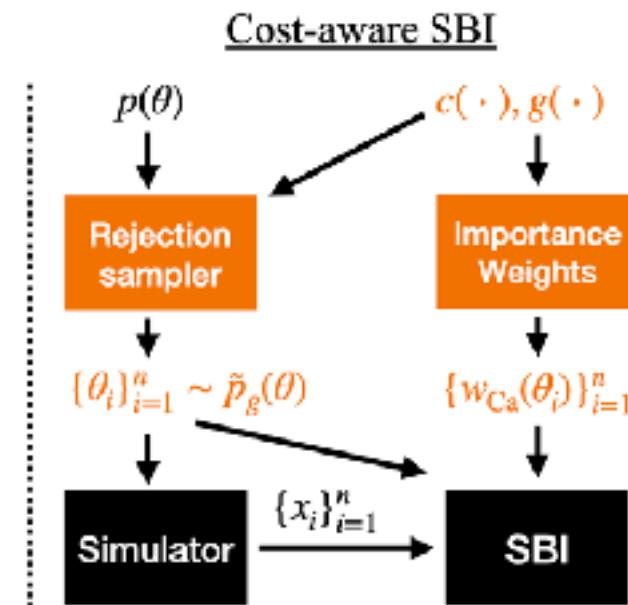
Putting it all together!

$$\ell_{\text{NLE}}(\phi) = -\frac{1}{n} \sum_{i=1}^n \log q_\phi(\mathbf{x}_i | \theta_i), \quad \theta_i \sim p(\theta), \mathbf{x}_i \sim p(\cdot | \theta)$$

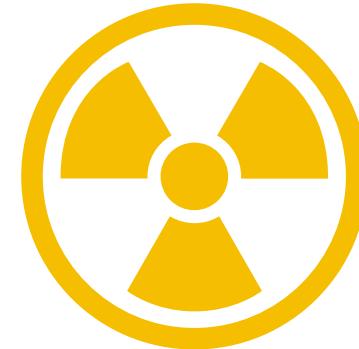


Putting it all together!

$$\ell_{\text{Ca-NLE}}(\phi) = -\frac{1}{n} \sum_{i=1}^n w_{\text{Ca}}(\theta_i) \log q_\phi(x_i | \theta_i), \quad \theta_i \sim \tilde{p}_g(\theta), x_i \sim p(\cdot | \theta)$$



Some reassuring results



Importance sampling can have infinite variance!!!

Some reassuring results

- Suppose that $g_{\max} = \sup_{\theta \in \Theta} g(c(\theta)) < \infty$. Then:

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- Suppose that $g_{\max} = \sup_{\theta \in \Theta} g(c(\theta)) < \infty$. Then:

- The weights are bounded: $\frac{g_{\min}}{ng_{\max}} \leq w_{\text{Ca}}(\theta_i) \leq \frac{g_{\max}}{ng_{\min}}$ $\forall i \in \{1, \dots, n\}$,

Some reassuring results

- Suppose that $g_{\max} = \sup_{\theta \in \Theta} g(c(\theta)) < \infty$. Then:
- 2. If f is square-integrable; i.e. $\int_{\Theta} f(\theta)^2 \pi(\theta) d\theta < \infty$, then $\text{Var}(\hat{\mu}_{\text{Ca}}) = \sigma_{\text{Ca}}^2$ where:

$$\frac{g_{\min}}{g_{\max}} \left(\sigma_{\text{MC}}^2 - \frac{\mu^2}{n} \right) \leq \sigma_{\text{Ca}}^2 \leq \frac{g_{\max}}{g_{\min}} \left(\sigma_{\text{MC}}^2 - \frac{\mu^2}{n} \right).$$

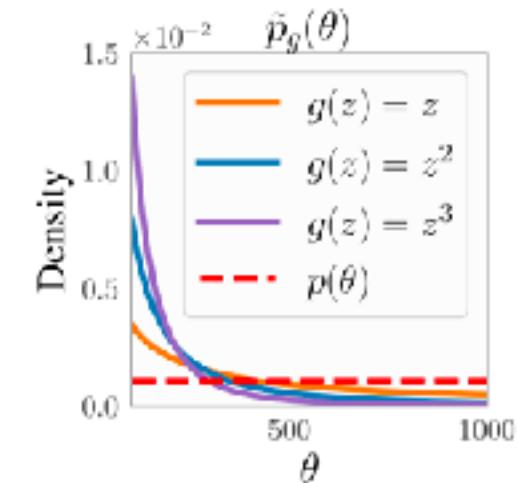
Some reassuring results

- Suppose that $g_{\max} = \sup_{\theta \in \Theta} g(c(\theta)) < \infty$. Then:

3. The ESS is bounded: $\left(\frac{g_{\min}}{g_{\max}} \right)^2 \leq \text{ESS} \leq \left(\frac{g_{\max}}{g_{\min}} \right)^2$.

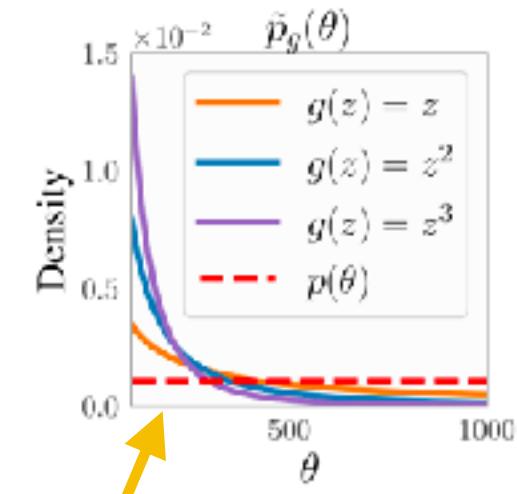
A Gamma simulator

- $\mathbb{P}_\theta = \text{Gamma}(\theta, 1)$,
- Simulator: Ahrens-Dieter acceptance-rejection method.
- Method: ABC!



A Gamma simulator

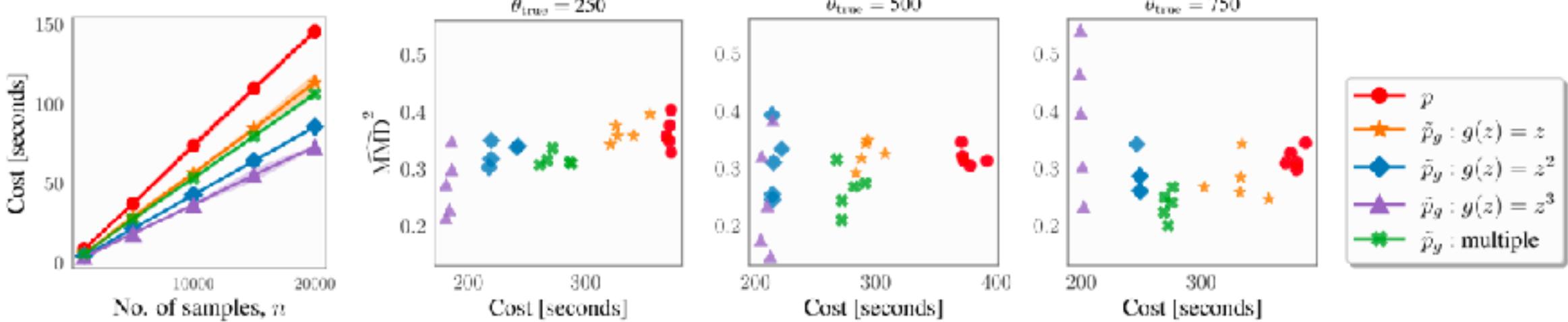
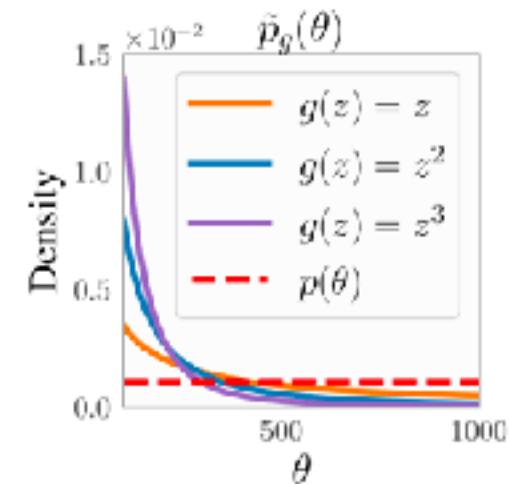
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Cost-aware pushes us to sample from small θ values!

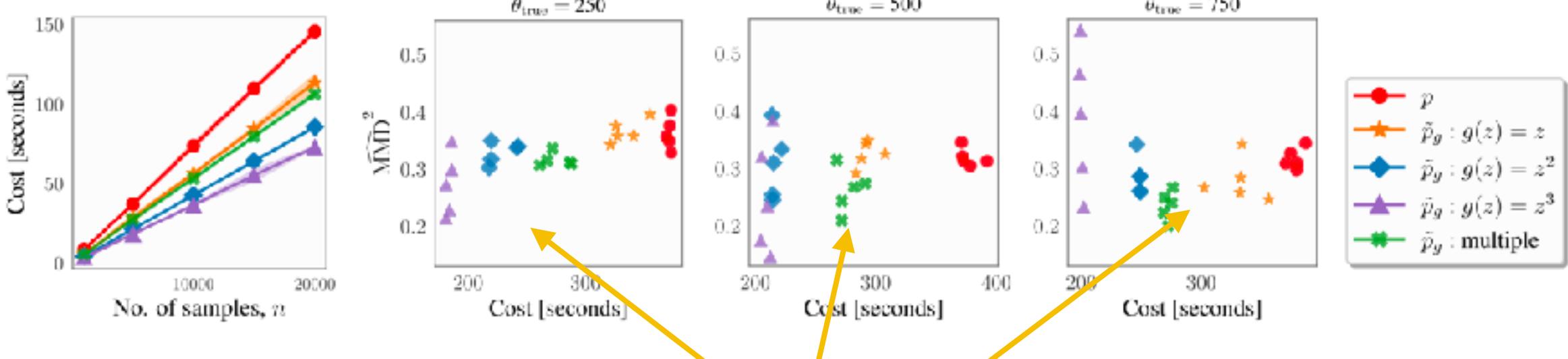
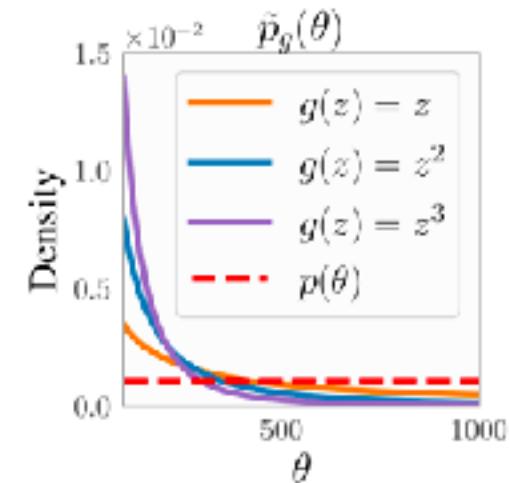
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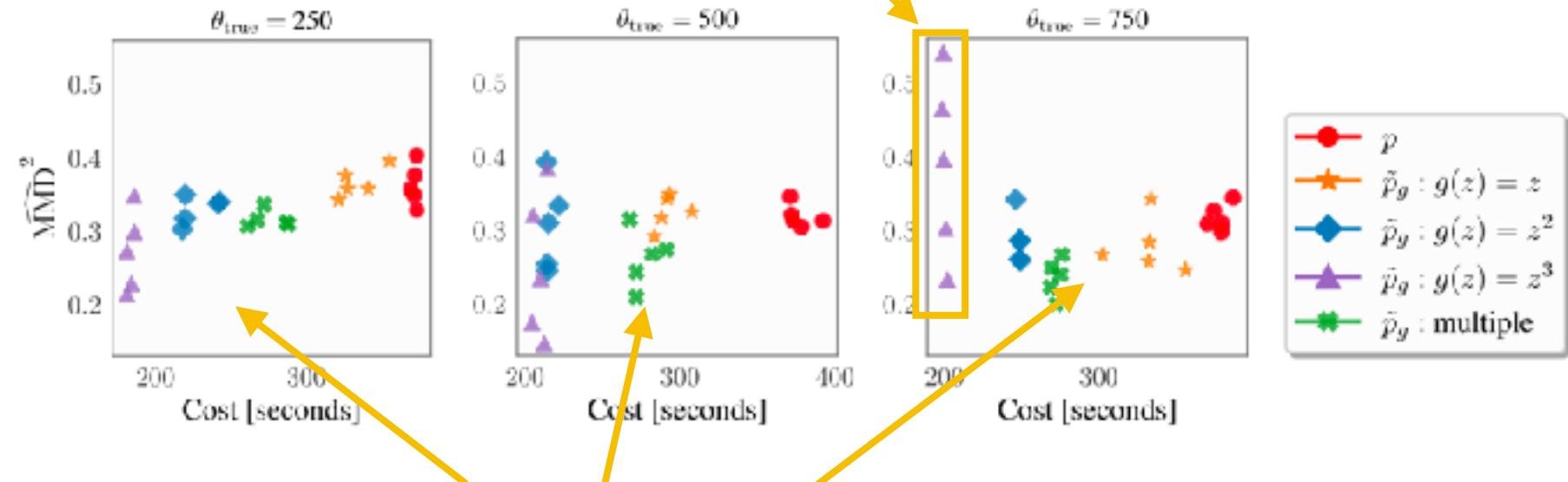
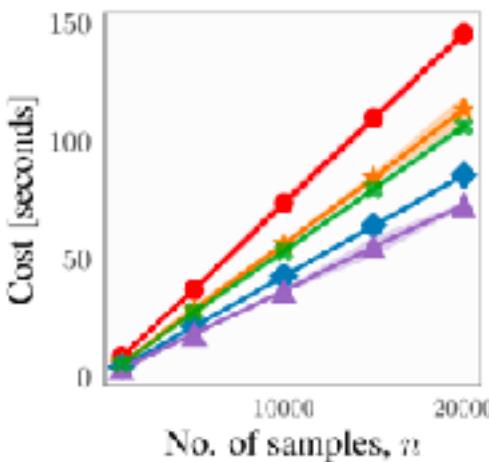
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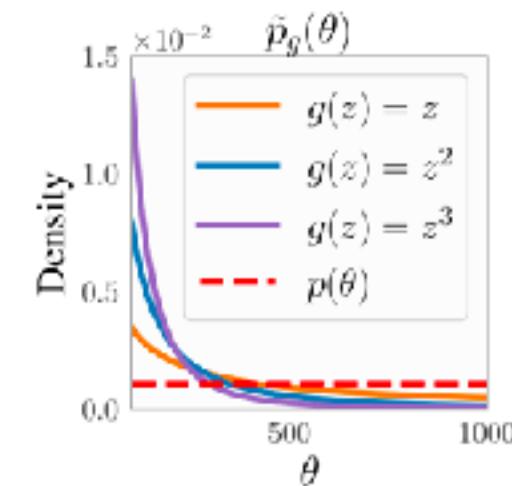
Being cost-aware tends to reduce your cost without a loss of accuracy!

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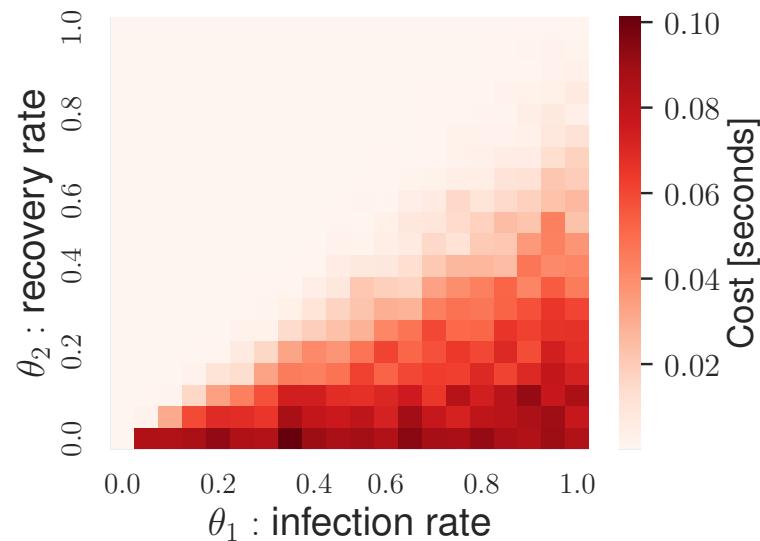
If truth in expensive region, being 'too' cost-aware won't be great!



Being cost-aware tends to reduce your cost without a loss of accuracy!

Some epidemiological models

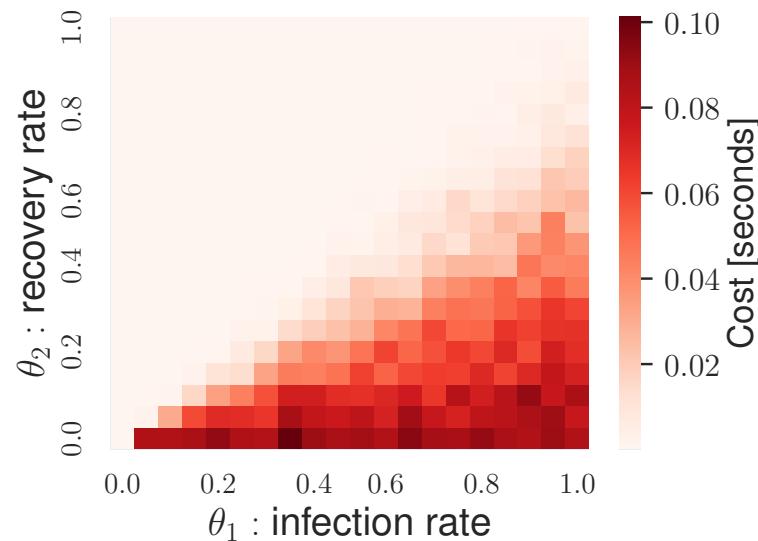
- We consider three different models with 1,2 and 3 parameters respectively, and use NPE.



Kypraios, T., Neal, P., and Prangle, D. (2017). A tutorial introduction to Bayesian inference for stochastic epidemic models using approximate Bayesian computation. Mathematical Biosciences, 287:42–53.

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$\widehat{\text{MMD}}^2 (\downarrow)$					Time saved (\uparrow)				
	NPE $g(z) = z^{0.5}$	Ca-NPE $g(z) = z$	Ca-NPE $g(z) = z^2$	Ca-NPE multiple		NPE $g(z) = z^{0.5}$	Ca-NPE $g(z) = z$	Ca-NPE $g(z) = z^2$	Ca-NPE multiple
Homogen.	0.02(0.02)	0.02(0.01)	0.02(0.02)	0.23(0.08)	0.05(0.04)	16%(2)	38%(2)	70%(2)	30%(5)
Temporal	0.03(0.03)	0.06(0.03)	0.07(0.03)	0.07(0.03)	0.05(0.04)	36%(4)	65%(2)	85%(1)	24%(5)
Bernoulli	0.02(0.00)	0.02(0.00)	0.02(0.01)	0.04(0.01)	0.02(0.00)	23%(4)	37%(4)	47%(3)	25%(6)

Kypraios, T., Neal, P., and Prangle, D. (2017). A tutorial introduction to Bayesian inference for stochastic epidemic models using approximate Bayesian computation. Mathematical Biosciences, 287:42–53.

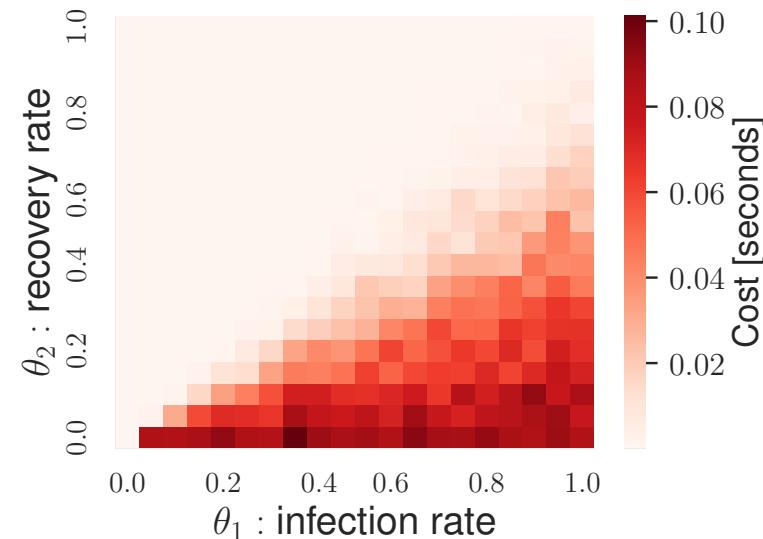
Some epidemiological models

- We consider three different models with 1, 2 and 3 parameters respectively, and use NPE.

$g(z) = z^{0.5}$: Same accuracy but modest improvement!

	$\widehat{\text{MMD}}^2 (\downarrow)$					Time saved (\uparrow)			
	NPE	Ca-NPE $g(z) = z^{0.5}$	Ca-NPE $g(z) = z$	Ca-NPE $g(z) = z^2$	Ca-NPE multiple	Ca-NPE $g(z) = z^{0.5}$	Ca-NPE $g(z) = z$	Ca-NPE $g(z) = z^2$	Ca-NPE multiple
Homogen.	0.02(0.02)	0.02(0.01)	0.02(0.02)	0.23(0.08)	0.05(0.04)	16%(2)	38%(2)	70%(2)	30%(5)
Temporal	0.03(0.03)	0.06(0.03)	0.07(0.03)	0.07(0.03)	0.05(0.04)	36%(4)	65%(2)	85%(1)	24%(5)
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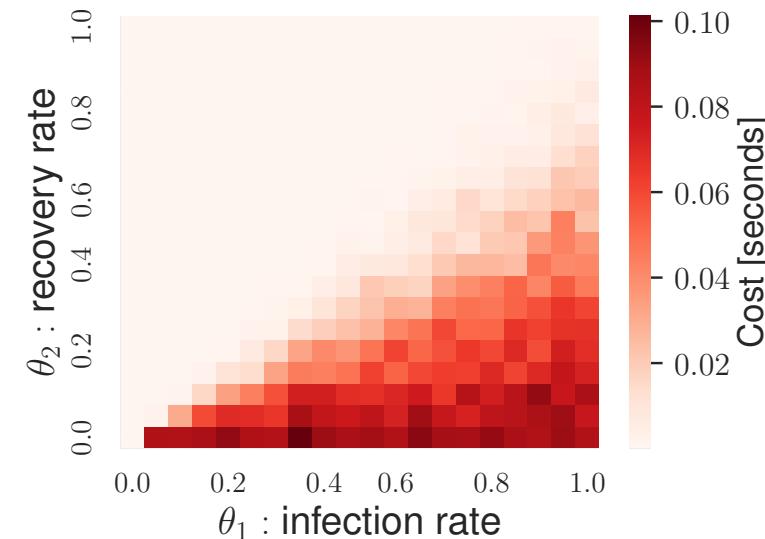


Some epidemiological models

- We consider three different models with 1, 2 and 3 parameters respectively, and use NPE.

$g(z) = z$: Still same accuracy but slightly better improvement!

	$\widehat{\text{MMD}}^2 (\downarrow)$				Time saved (\uparrow)				
	NPE	Ca-NPE $g(z) = z^{0.5}$	Ca-NPE $g(z) = z$	Ca-NPE $g(z) = z^2$	Ca-NPE multiple	Ca-NPE $g(z) = z^{0.5}$	Ca-NPE $g(z) = z$	Ca-NPE $g(z) = z^2$	Ca-NPE multiple
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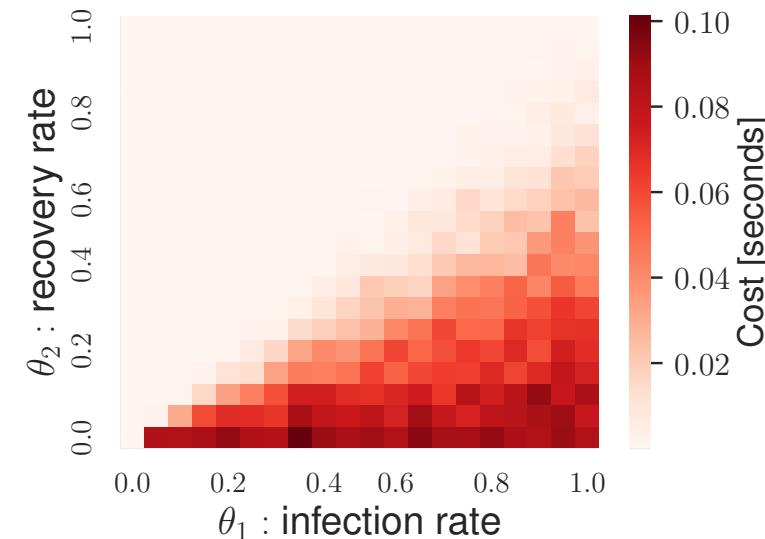


Some epidemiological models

- We consider three different models with 1, 2 and 3 parameters respectively, and use NPE.

$g(z) = z^2$: Worse accuracy but much cheaper

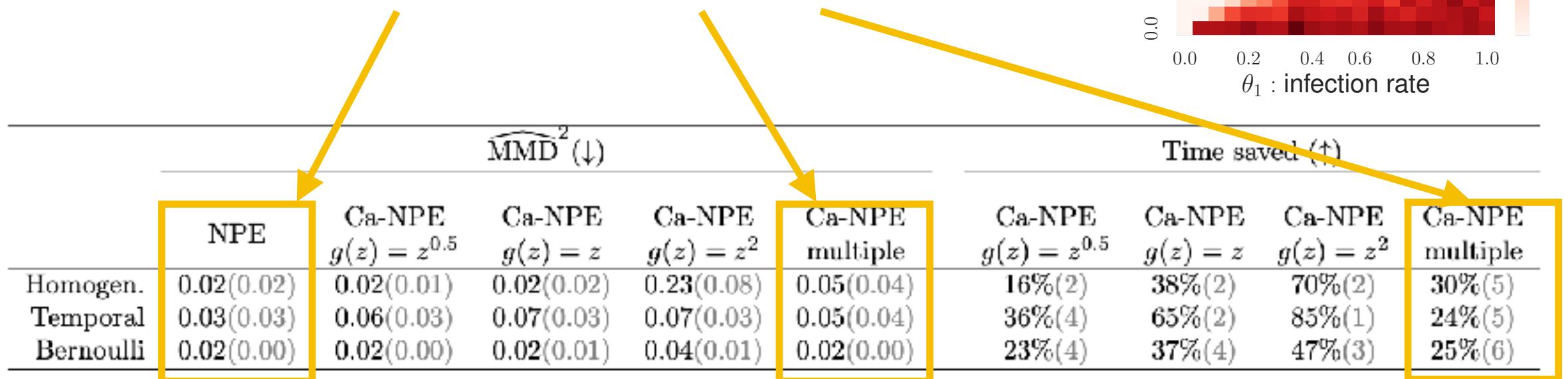
	$\widehat{\text{MMD}}^2 (\downarrow)$				Time saved (\uparrow)				
	NPE	Ca-NPE $g(z) = z^{0.5}$	Ca-NPE $g(z) = z$	Ca-NPE $g(z) = z^2$	Ca-NPE multiple	Ca-NPE $g(z) = z^{0.5}$	Ca-NPE $g(z) = z$	Ca-NPE $g(z) = z^2$	Ca-NPE multiple
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Some epidemiological models

- We consider three different models with 1, 2 and 3 parameters respectively, and use NPE.

Typically slight loss of accuracy but decent reduction in cost!



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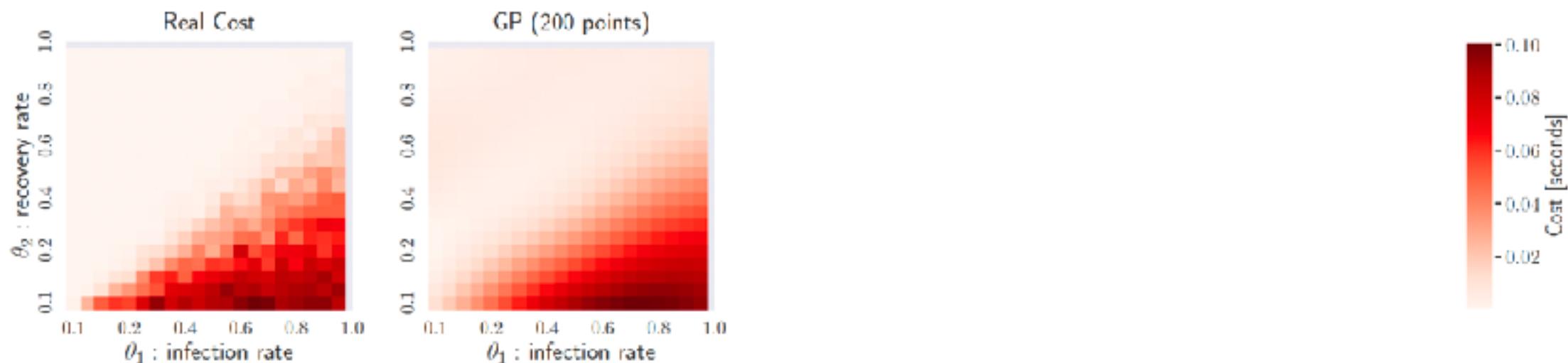
Estimating the cost function

When the cost function is unknown, it can be estimated through simulations+regression.
This is typically very cheap, and simulations can be re-used for inference!



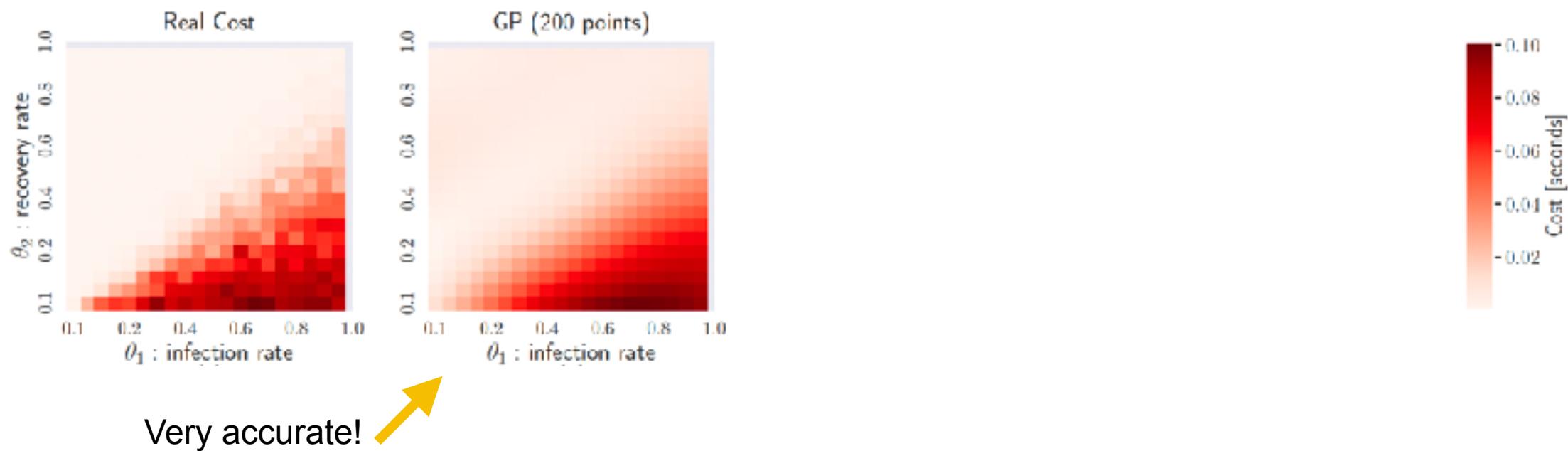
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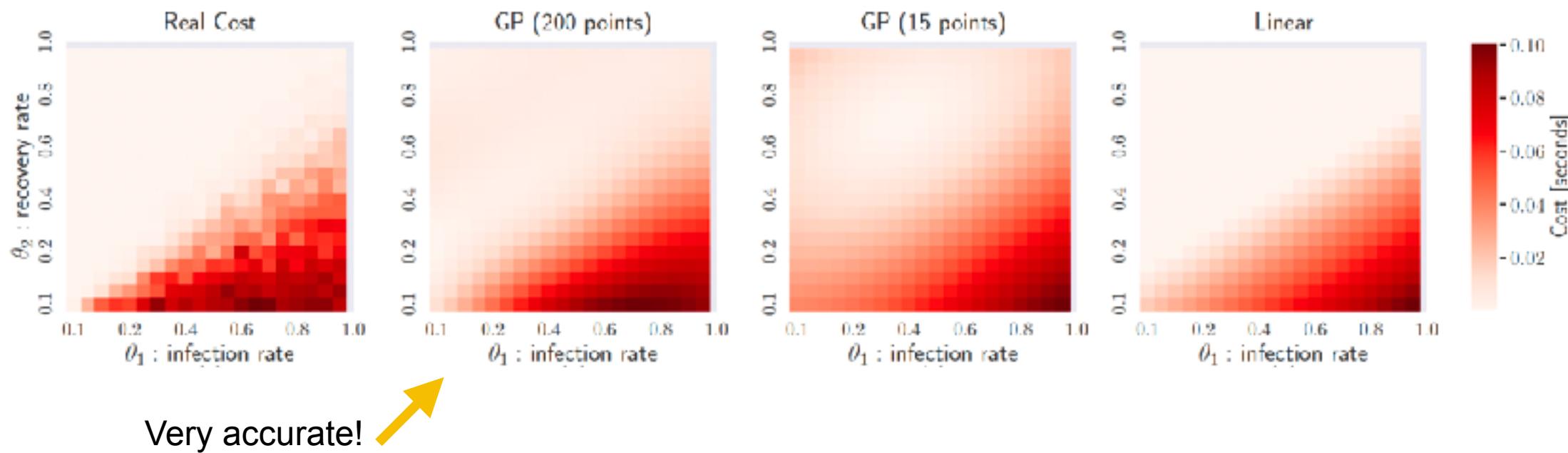
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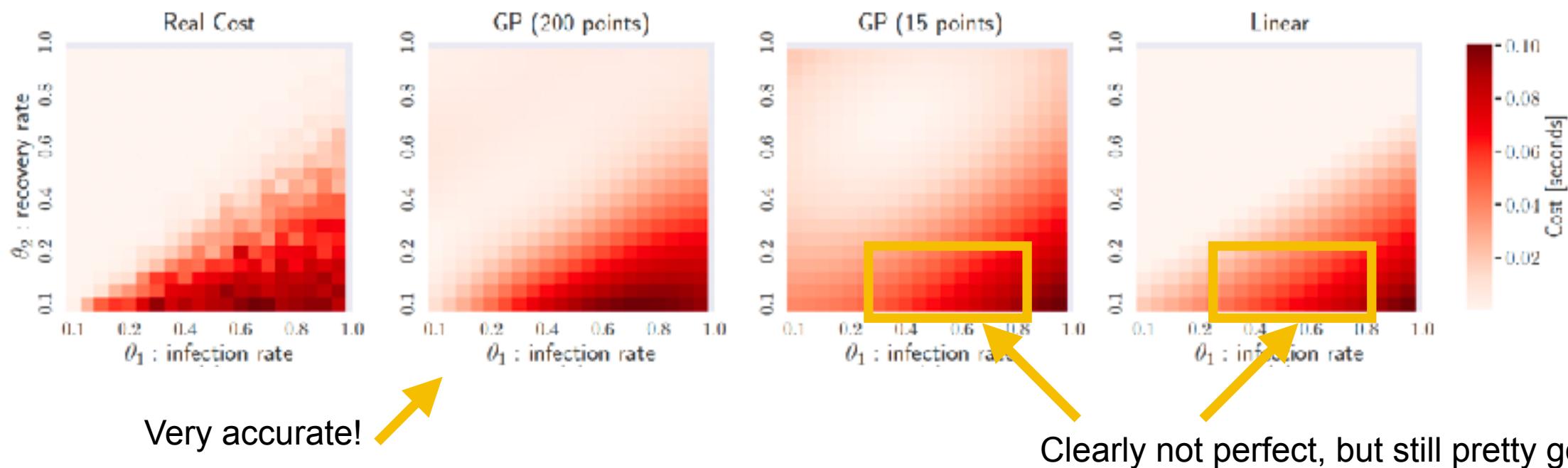
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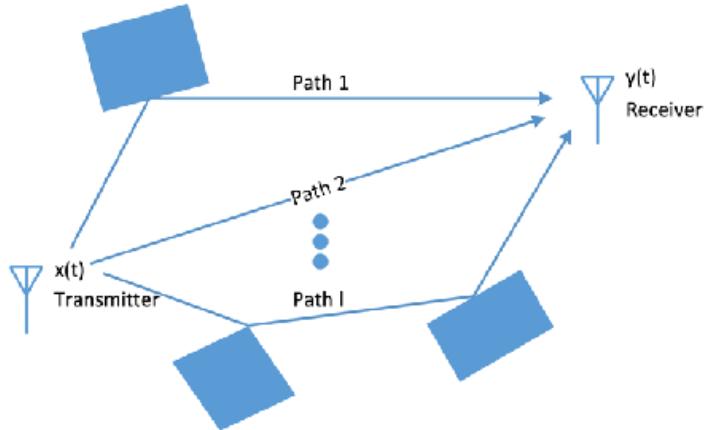


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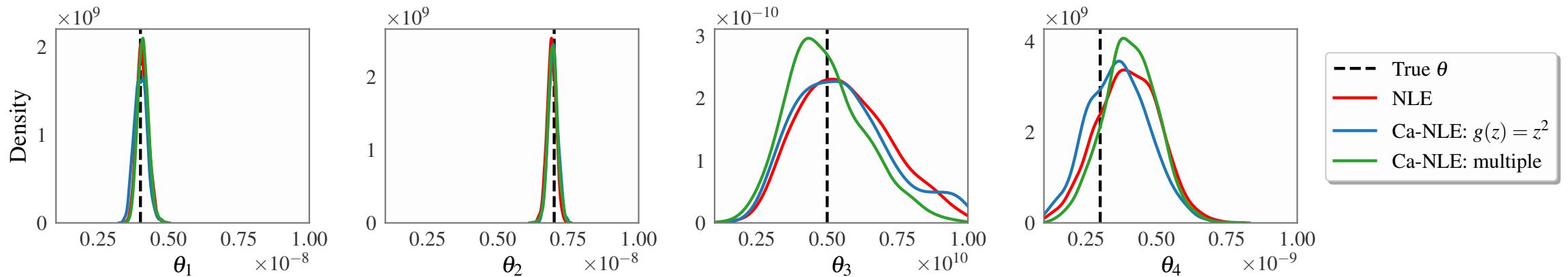


Back to radio-propagation



Computational Cost

- Standard NLE: 15.6h,
- Cost-aware NLE: 8.8h!!



Conclusion

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Conclusion

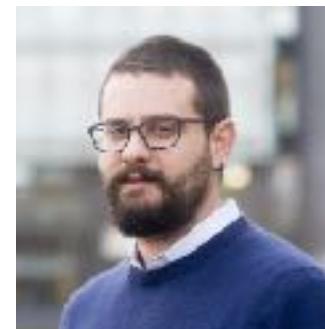
- We proposed a novel importance sampling algorithm which focuses on **down weighting sampling** in regions with a **large downstream cost**.
- Although I presented this for NLE/NPE, we also have experiments for ABC and it could be applied to any other sampling-based SBI method.
- Need more computational statisticians engaging with neural-based simulation inference!

Any Questions?

Paper: Bharti, A., Huang, D., Kaski, S., & Briol, F.-X. (2025). Cost-aware simulation-based inference. International Conference on Artificial Intelligence and Statistics, 28–36.

Code: <https://github.com/huangdaolang/cost-aware-sbi>

Robust Bayesian simulation-based inference



Paper: Dellaporta, C., Knoblauch, J., Damoulas, T. & **Briol, F-X** (2022). Robust Bayesian inference for simulator-based models via the MMD posterior bootstrap. AISTATS, 943-970. Best paper award.

Code: https://github.com/haritadell/npl_mmd_project



Connections with Jeremias' course

Optimisation-centric posteriors /
Generalised Variational Inference

$$q_n^*(\theta) = \arg \min_{q \in \mathcal{Q}} \left\{ \mathcal{L}(q, x_{1:n}) + D(q, \pi) \right\}$$

Gibbs/Generalised/

$$\pi_n^{\perp}(\theta | x_{1:n}) = \frac{\exp\{-\mathcal{L}(x_{1:n}, p_\theta)\} \cdot \pi(\theta)}{\int \exp\{-\mathcal{L}(x_{1:n}, p_\theta)\} \cdot \pi(\theta) d\theta}$$

Martingale posteriors &
resampling-based approaches

$$\begin{aligned} \text{For } i = 1, 2, \dots \\ X_{n+i+1} \sim p(X_{n+i} | x_{1:n}, X_{n+1:n+i}) \\ \theta^\infty = \operatorname{argmin}_{\theta \in \Theta} \mathcal{L}([x_{1:n}, X_{n+1:\infty}], \theta) \end{aligned}$$

[See Fong, Holmes, & Walker (2023)]

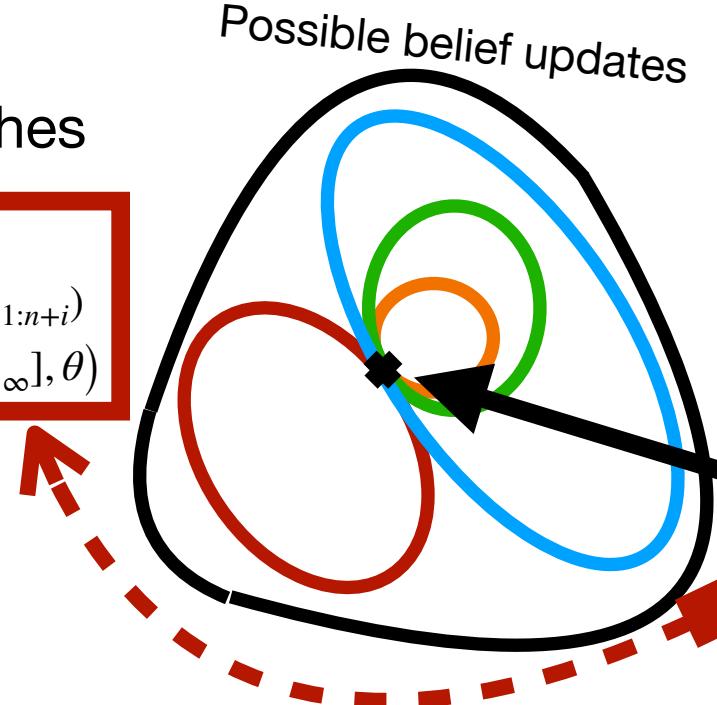
Possible belief updates

Power/Fractional/

$$\pi_n^{(\lambda)}(\theta | x_{1:n}) = \frac{p(x_{1:n} | \theta)^\lambda \cdot \pi(\theta)}{\int p(x_{1:n} | \theta)^\lambda \cdot \pi(\theta) d\theta}$$

Bayes'

$$\pi_n(\theta | x_{1:n}) = \frac{p(x_{1:n} | \theta) \cdot \pi(\theta)}{\int p(x_{1:n} | \theta) \cdot \pi(\theta) d\theta}$$





Connections with Jeremias' course

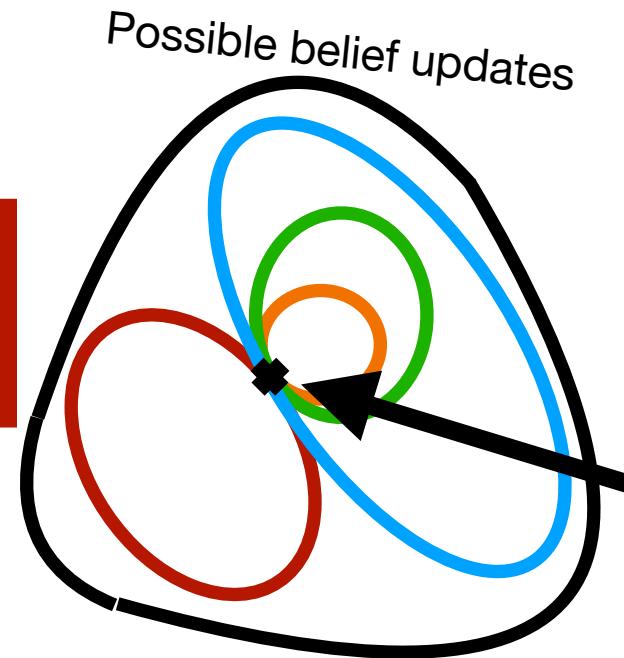
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[See Fong, Holmes, & Walker (2023)]



Non-parametric Learning

- Place a Dirichlet process $\text{DP}(\alpha; \mathbb{F})$ prior on \mathbb{Q}

Lyddon, S., Walker, S., & Holmes, C. (2018). Nonparametric learning from Bayesian models with randomized objective functions. *NeurIPS*, 2071–2081.

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$$\text{DP}(\alpha'; \mathbb{F}') \quad \alpha' = \alpha + n \quad \mathbb{F}' = \frac{\alpha}{\alpha + n} \mathbb{F} + \frac{n}{\alpha + n} \mathbb{Q}_n$$

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We still care about $\{\mathbb{P}_\theta\}_{\theta \in \Theta}$, so we map back to parameter space!

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The posterior bootstrap (i.e. NPL in practice)

(1) Sample $\mathbb{Q}^{(1)}, \mathbb{Q}^{(2)}, \dots$ from the DP posterior.

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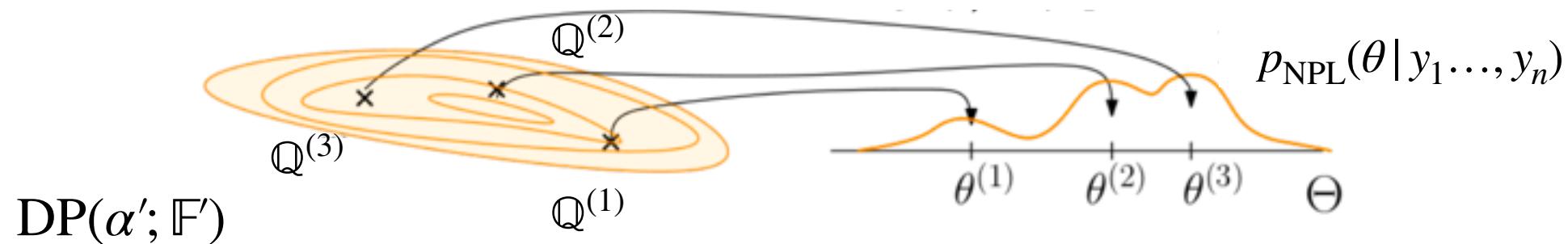
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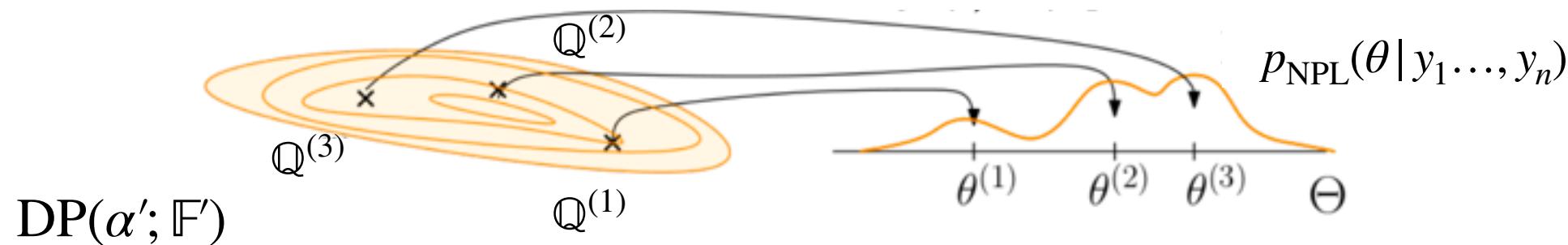
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Approximated with empirical loss



The MMD posterior bootstrap

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$$\text{MMD}^2(\mathbb{P}, \mathbb{Q}) = \int_{\mathcal{X}} \int_{\mathcal{X}} k(x, y) \mathbb{P}(dx) \mathbb{P}(dy) - 2 \int_{\mathcal{X}} \int_{\mathcal{X}} k(x, y) \mathbb{P}(dx) \mathbb{Q}(dy) + \int_{\mathcal{X}} \int_{\mathcal{X}} k(x, y) \mathbb{Q}(dx) \mathbb{Q}(dy)$$

Bounded!

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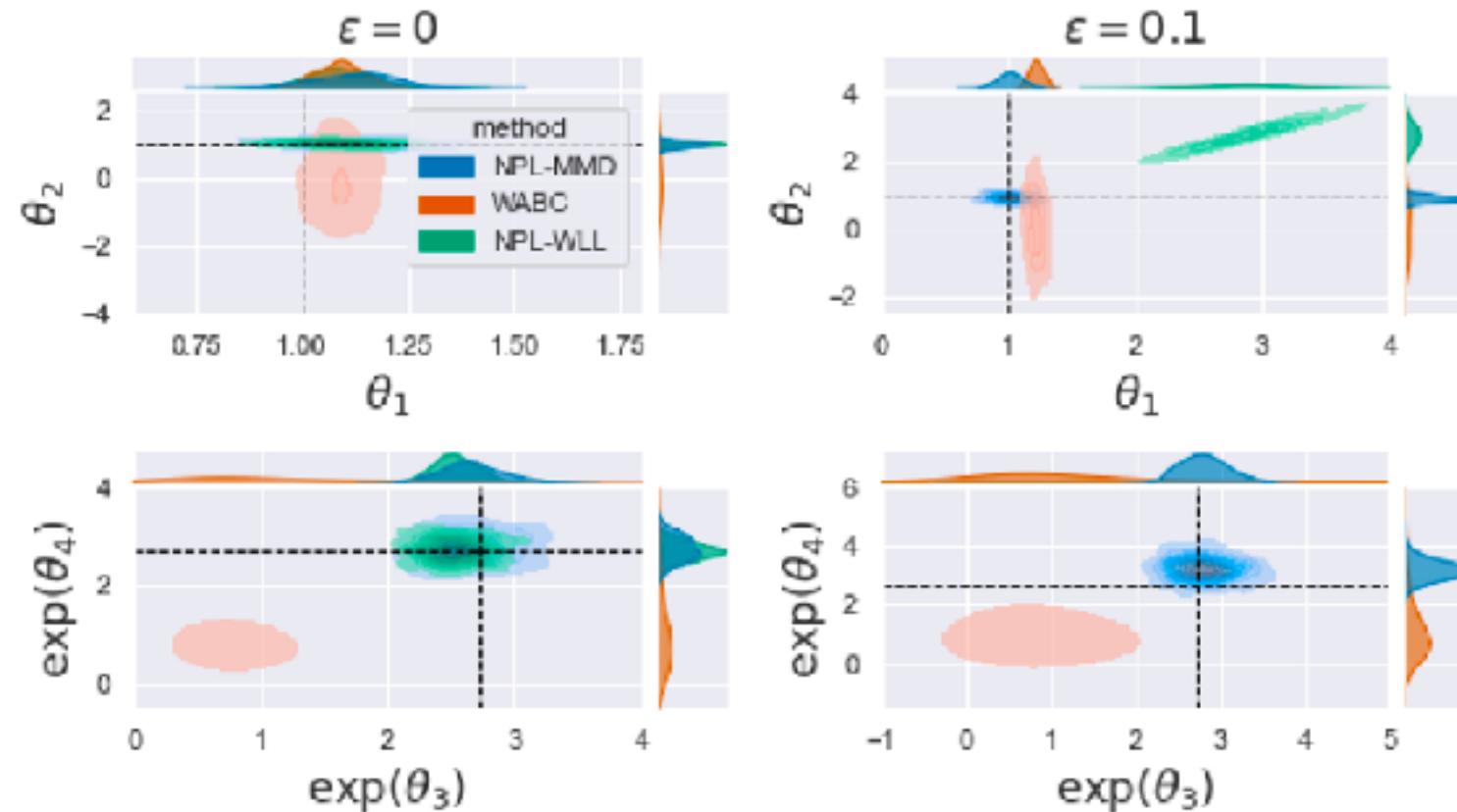
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Double robustness robust inference procedure and robust estimator!

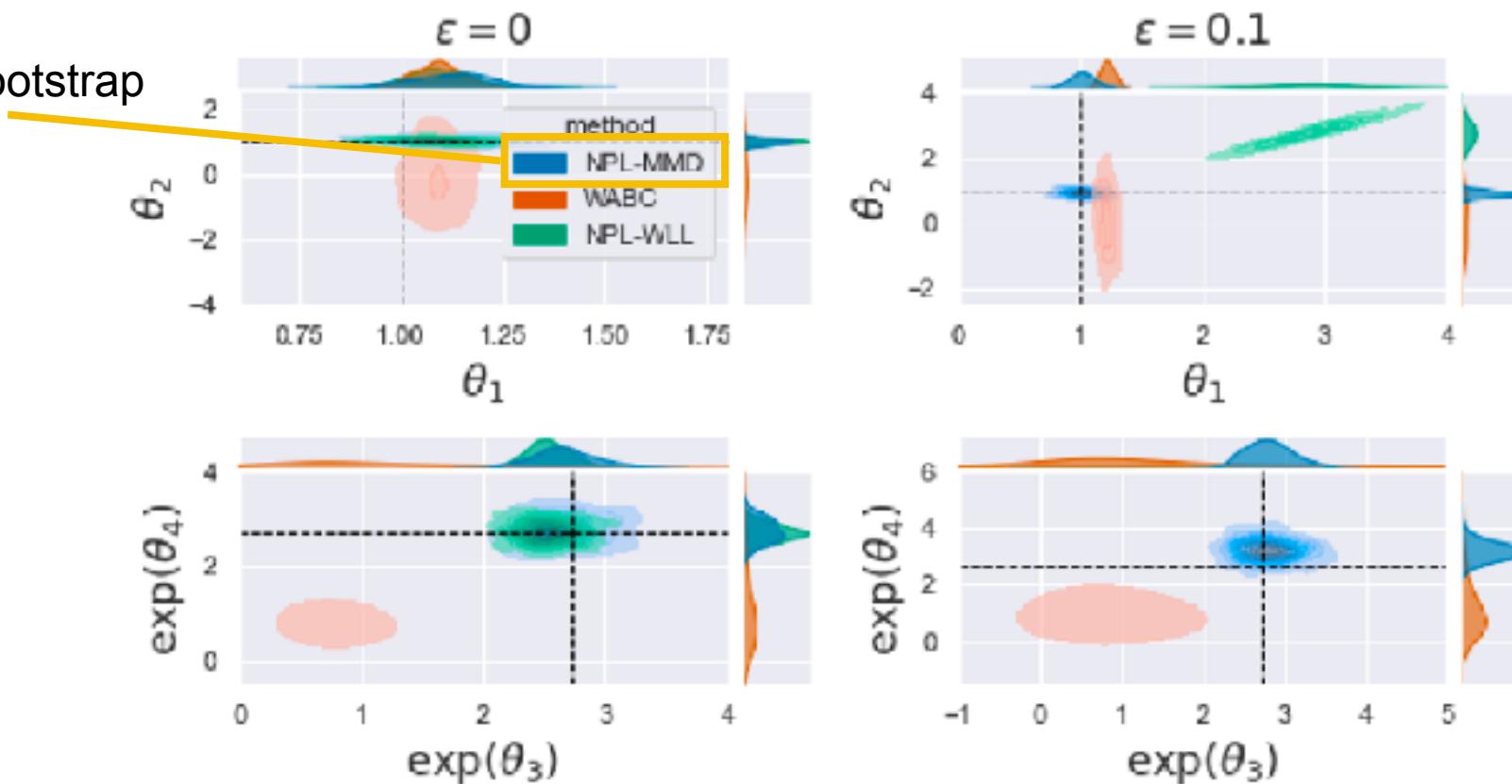
Example 1: Misspecified Gaussian



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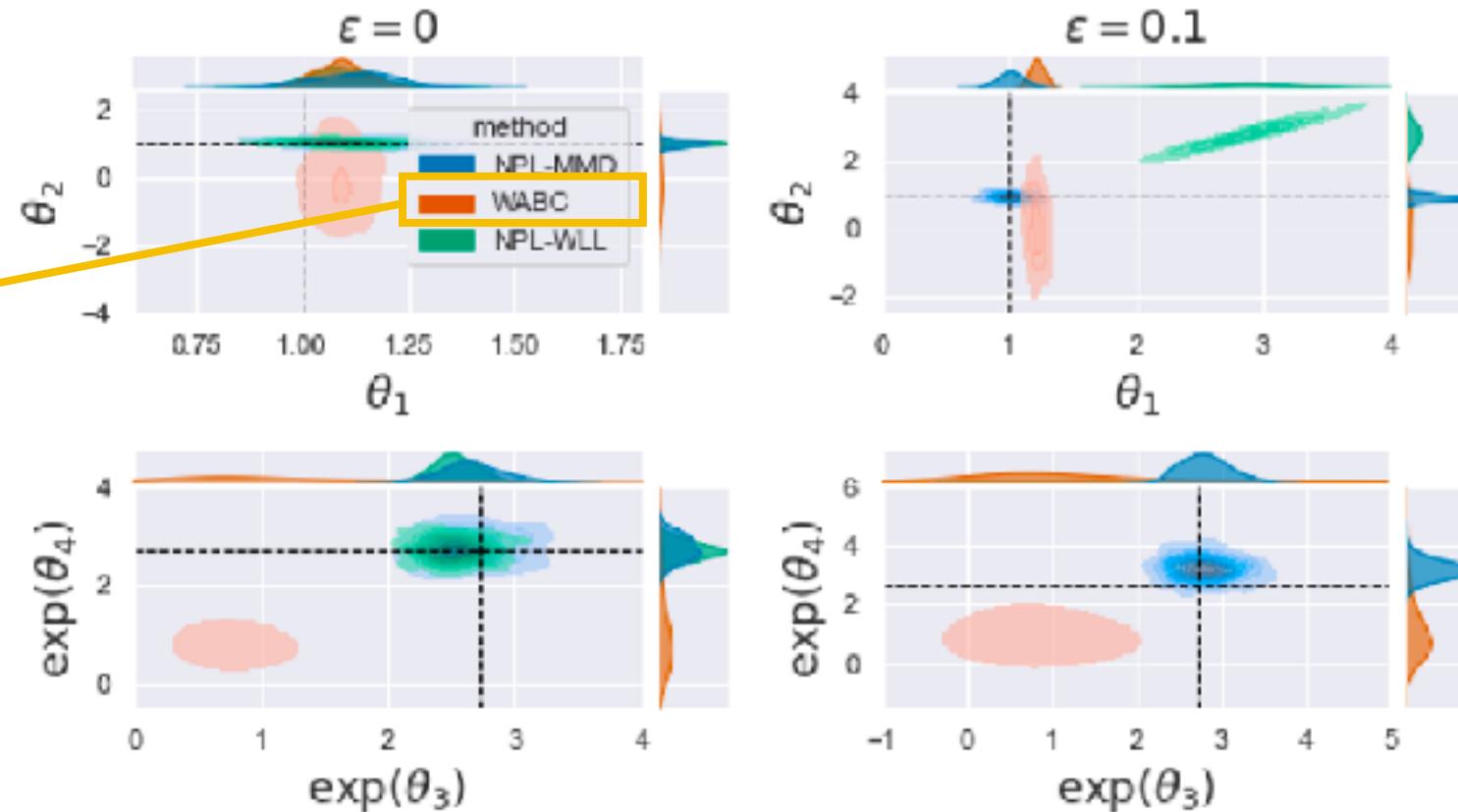
NPL-MMD:

MMD posterior bootstrap

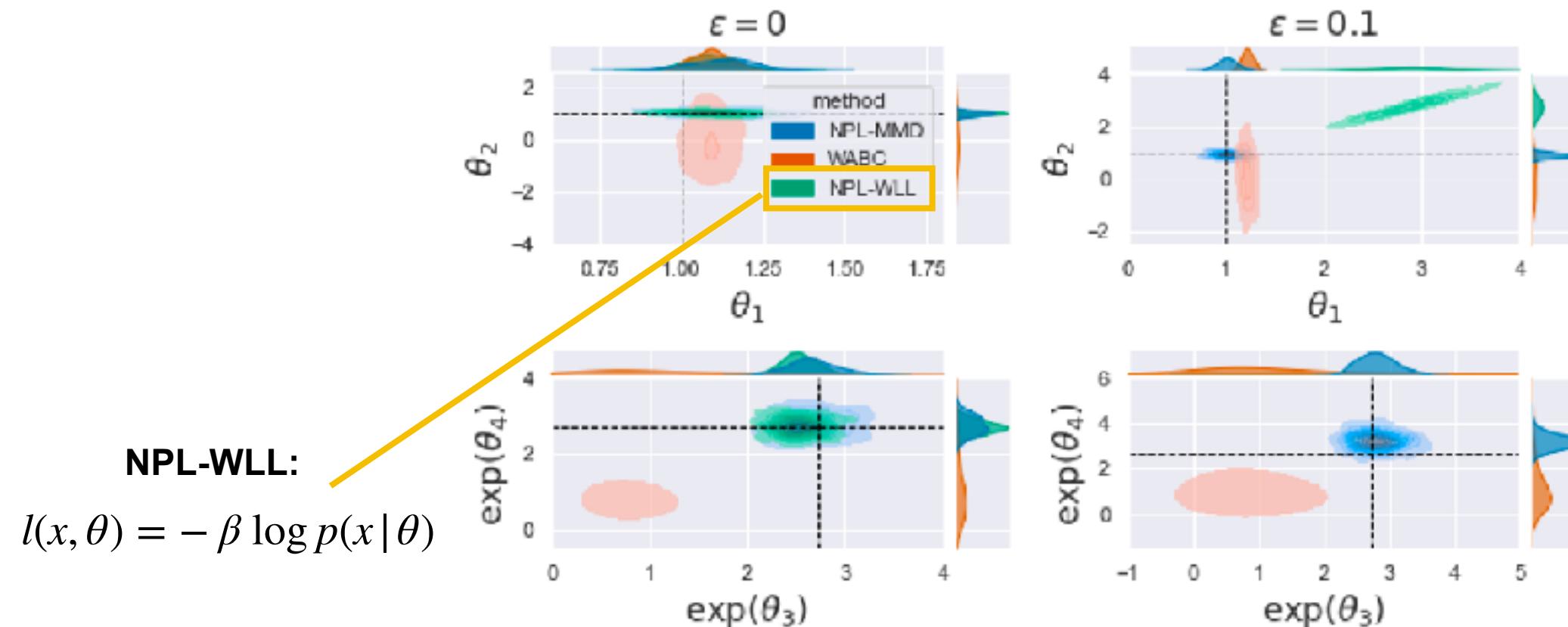


Example 1: Misspecified Gaussian

WABC:
ABC with Wasserstein
distance

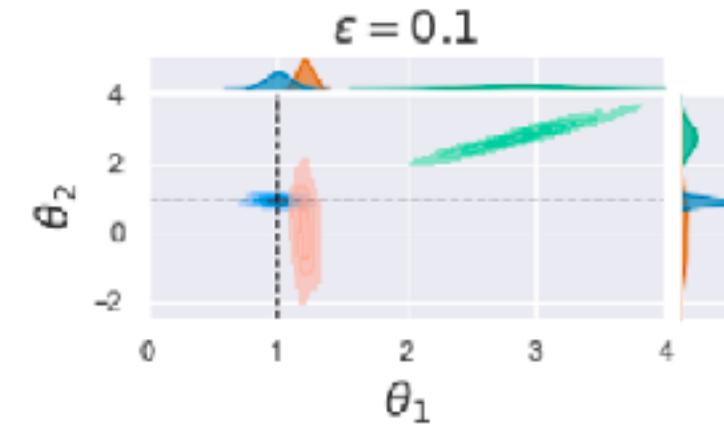
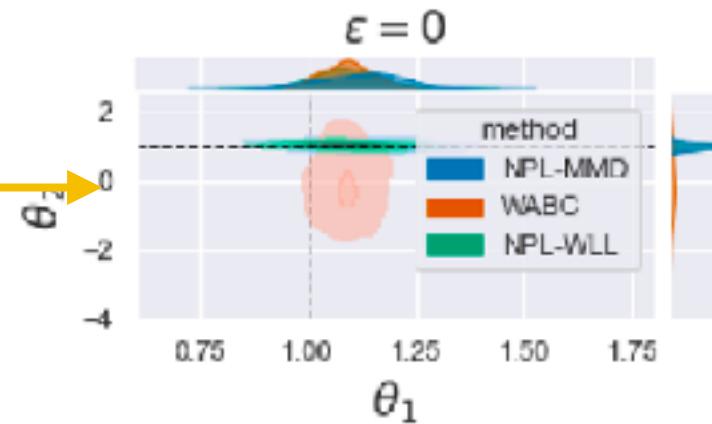


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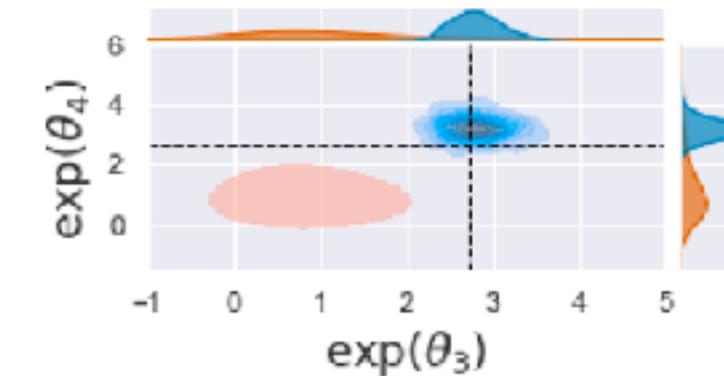
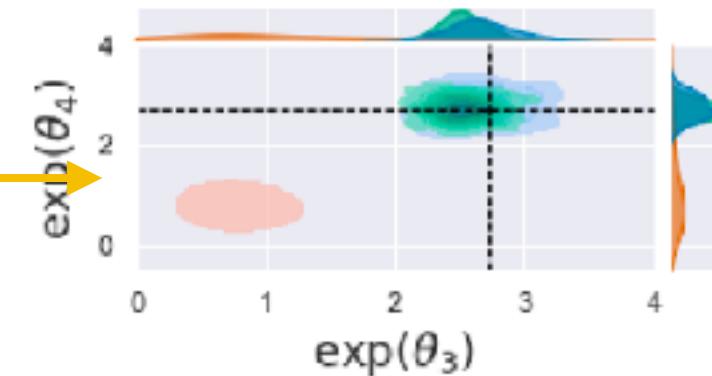


Example 1: Misspecified Gaussian

'Easy' parameters;
All do ok!

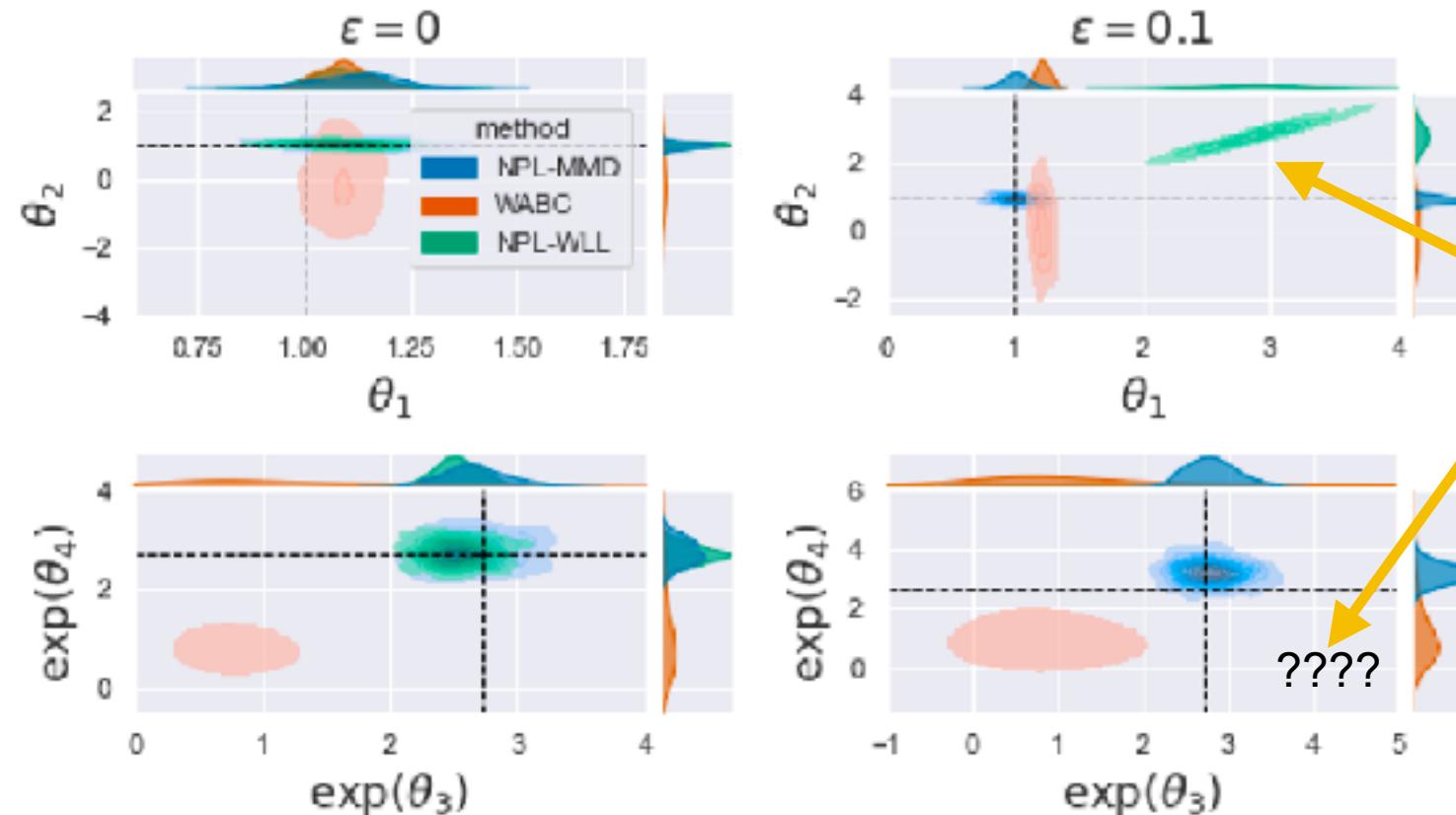


'Hard' parameters;
WABC already
struggles a bit



Well-specified case!

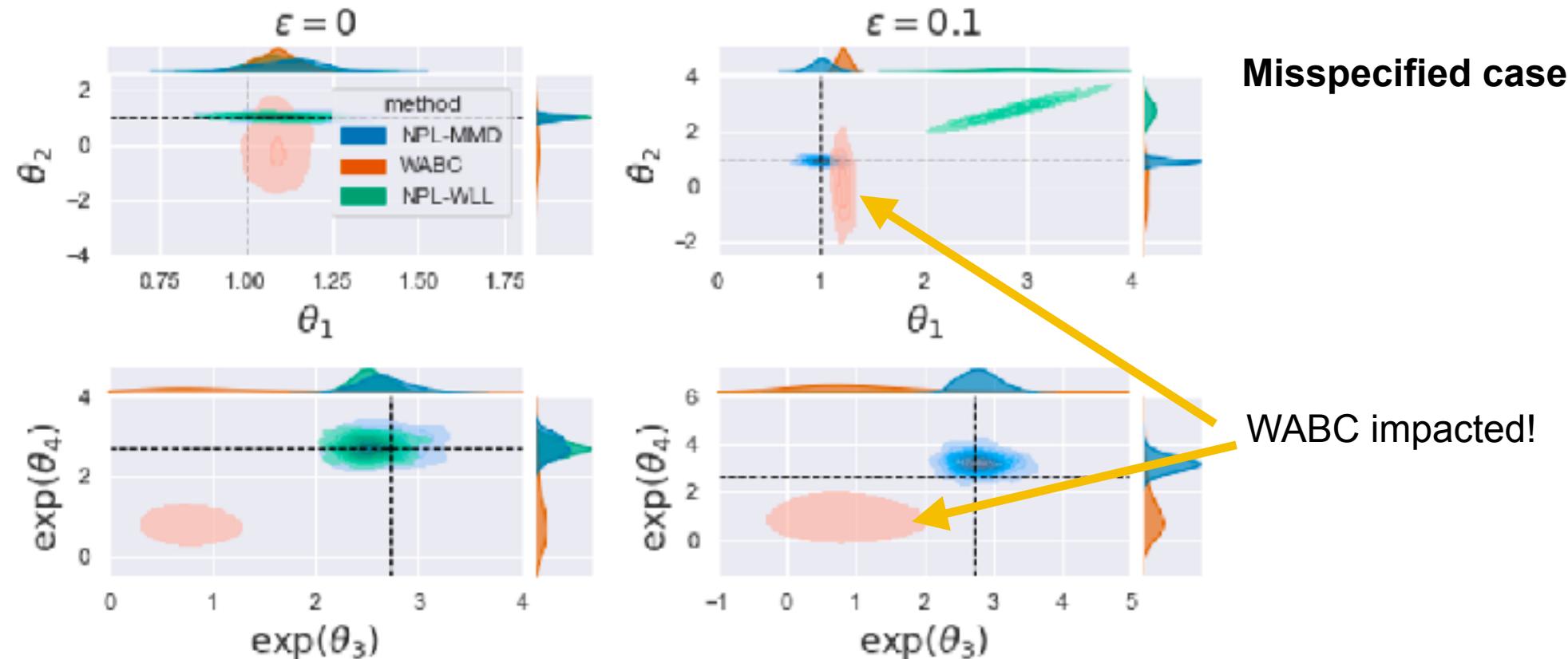
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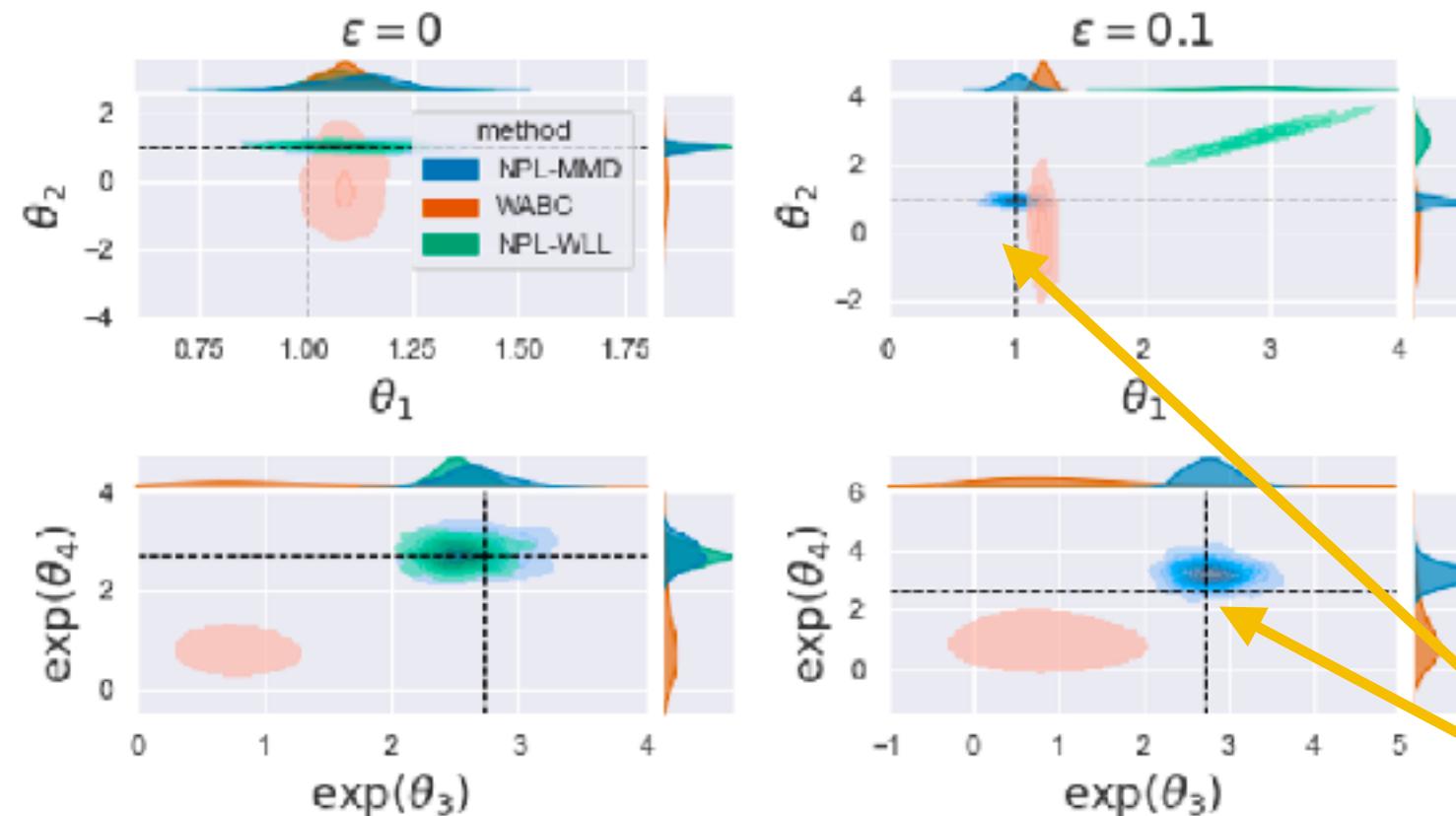
Misspecified case

NPL-WLL really
struggles

Example 1: Misspecified Gaussian



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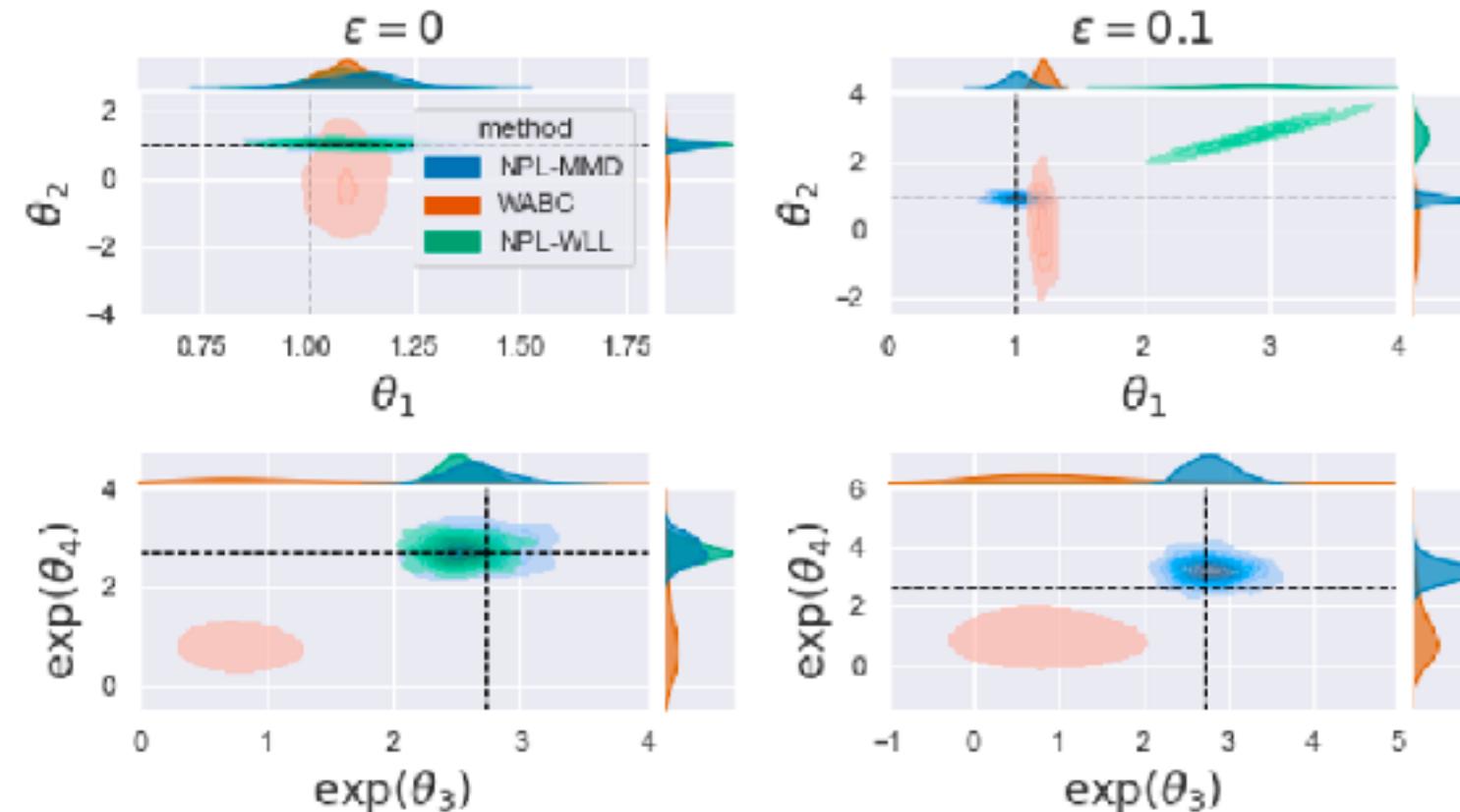


Misspecified case

MMD posterior bootstrap
barely impacted!!

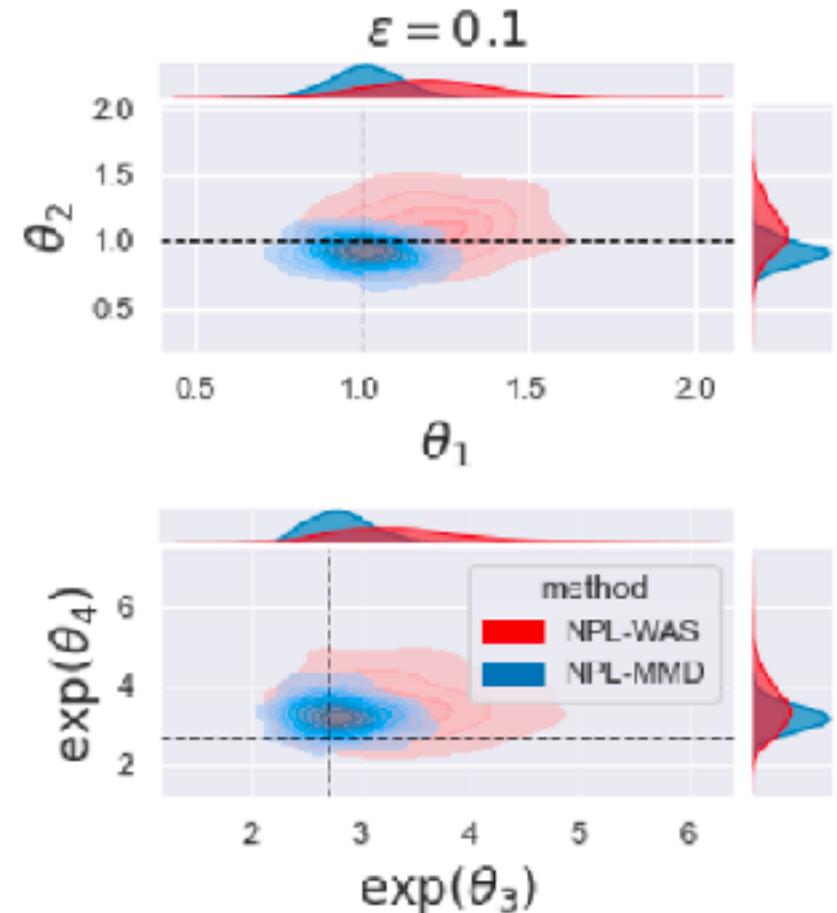
Example 1: Misspecified Gaussian

Time to run:
NPL-MMD: ≈ 2 mins
WABC: ≈ 1 hour



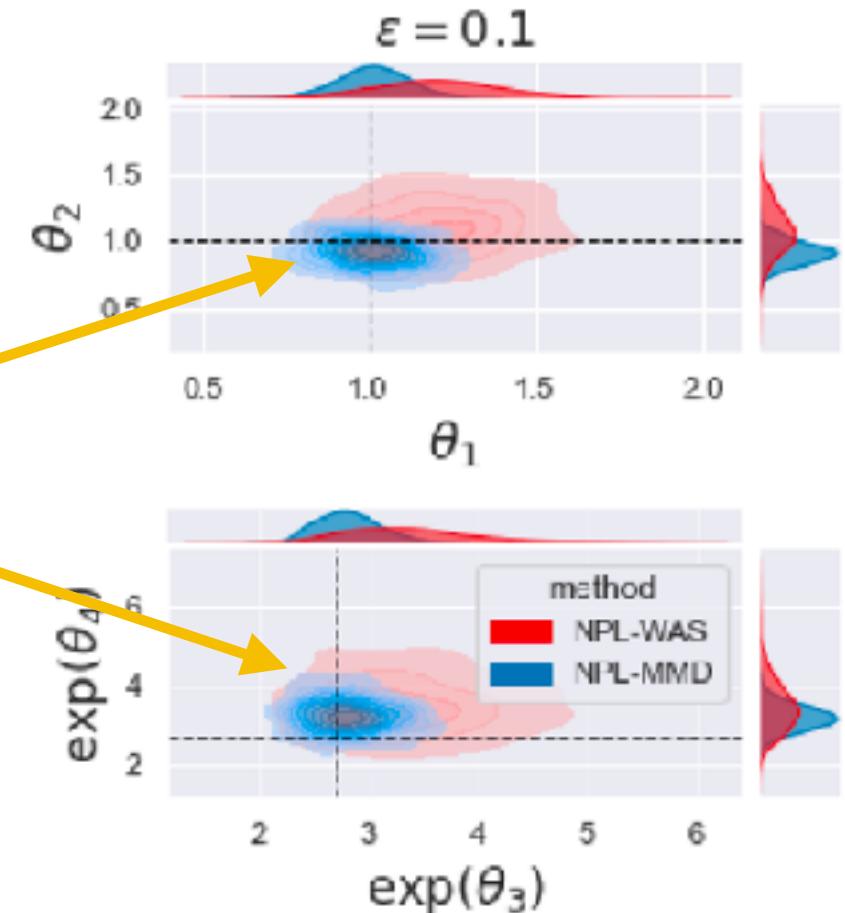
Example 1 continued: Wasserstein NPL

- In principle, nothing stops us from using the Wasserstein instead of MMD.



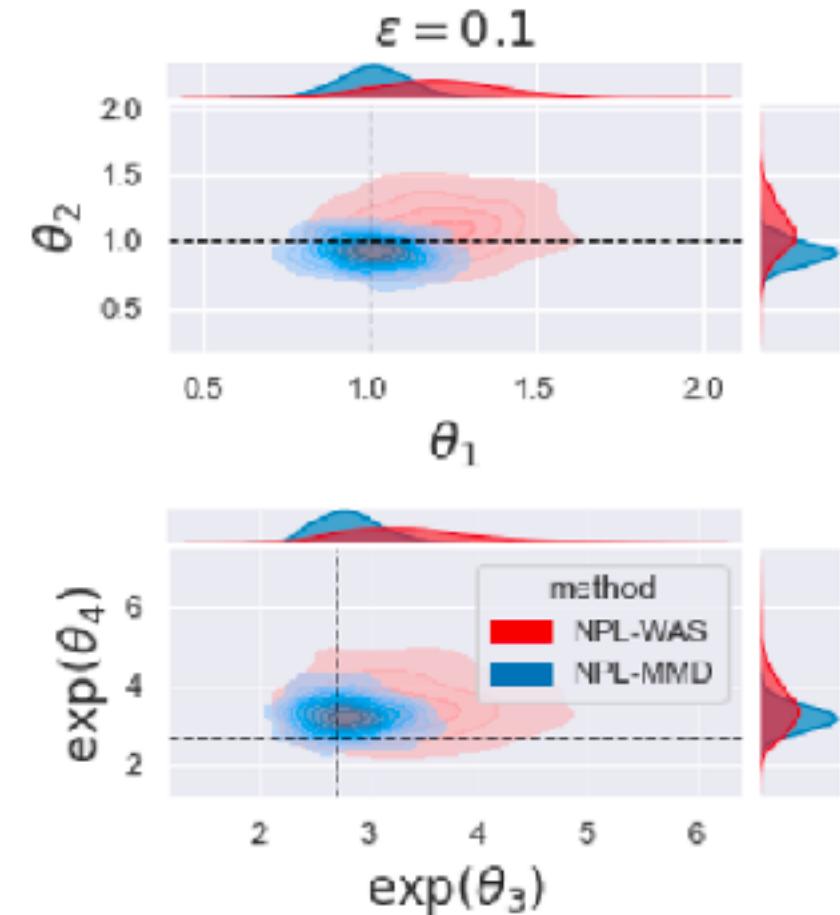
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- The results are still reasonable thanks to NPL framework, but much more diffuse



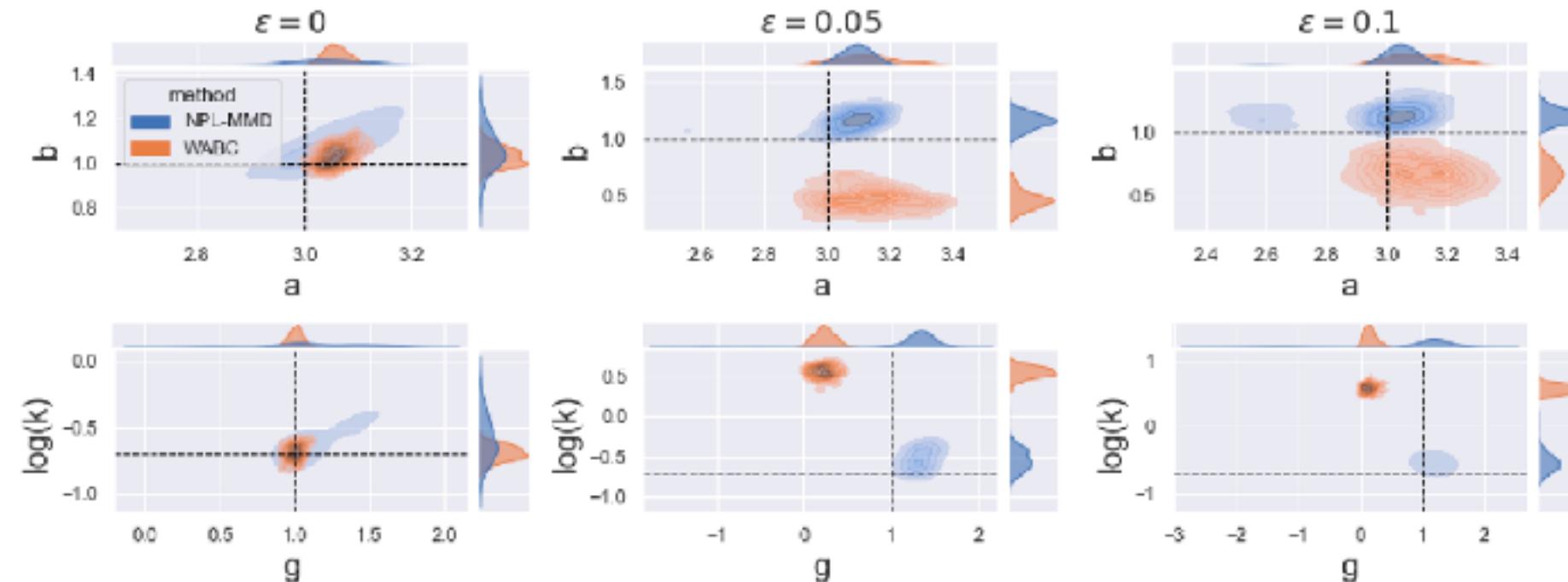
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→ We really do gain from having both a **robust inference** framework AND a **robust estimator**...

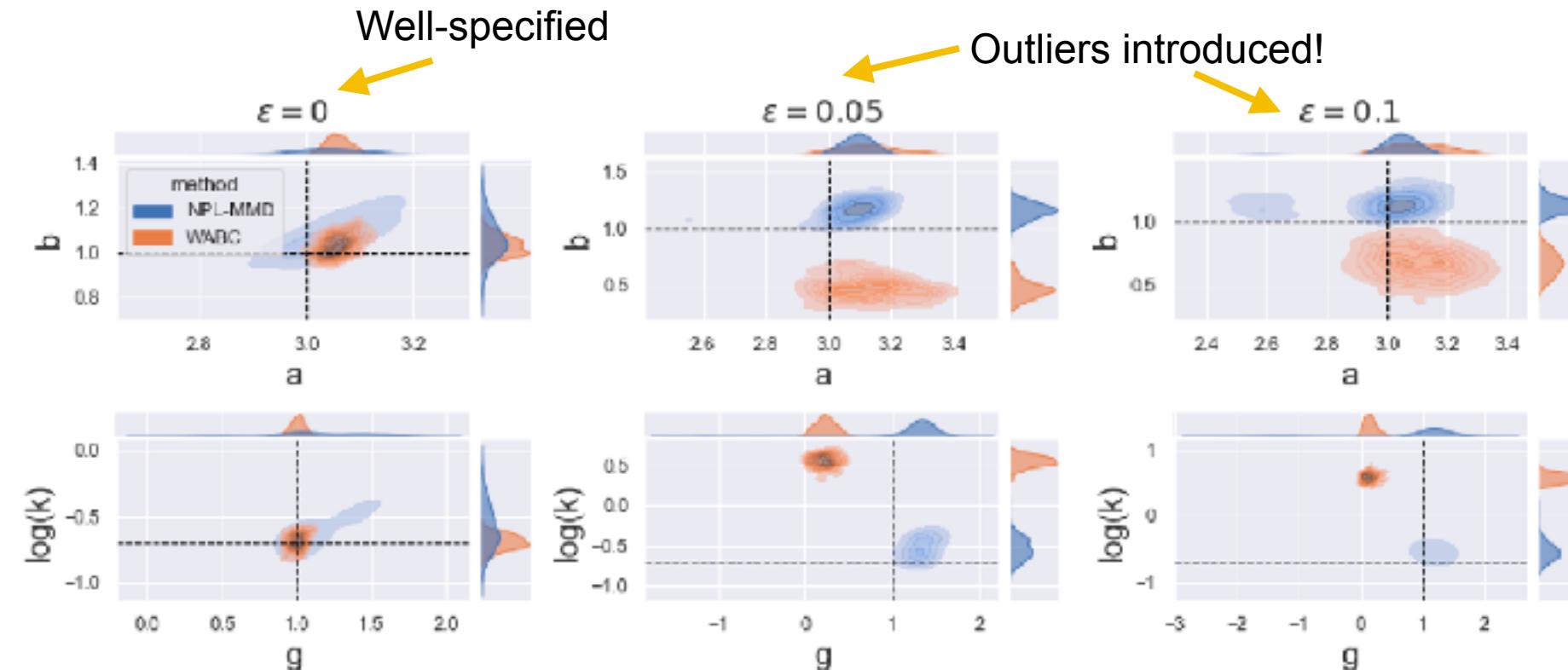
Misspecified g-and-k distribution



$$G_\theta(u) = \theta_1 + \theta_2 \left(1 + 0.8 \left(\frac{1 - \exp(-\theta_3 z(u))}{1 + \exp(-\theta_3 z(u))} \right) \right) (1 + z(u)^2)^{\log(\theta_4)} z(u),$$

$$z(u) = \Phi^{-1}(u) = \sqrt{2} \operatorname{erf}^{-1}(2u - 1), \quad u \sim \text{Unif}([0,1]),$$

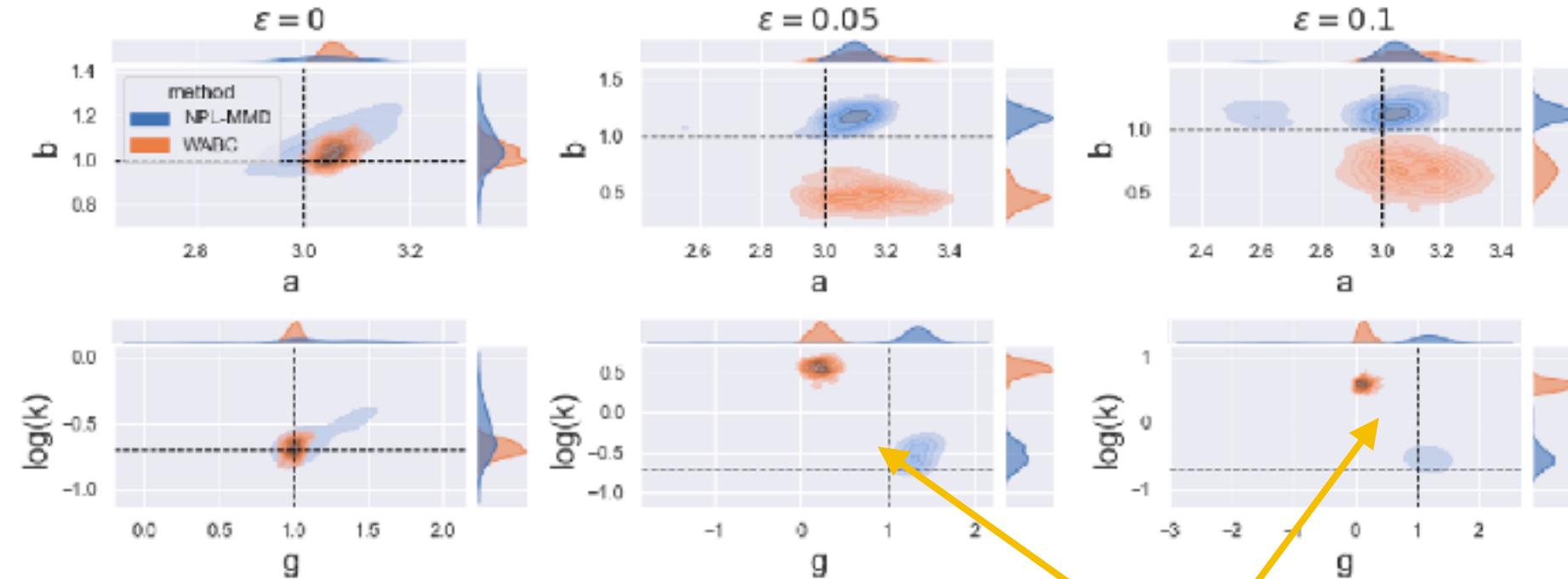
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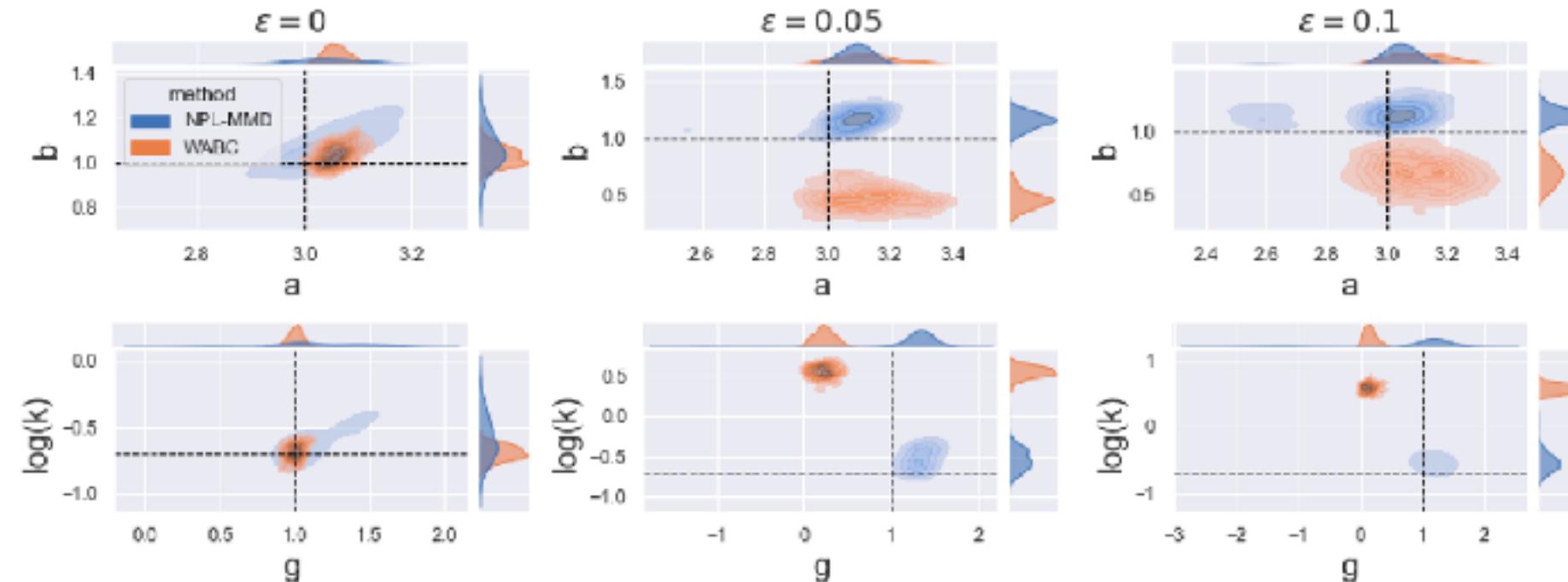
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Wasserstein ABC really struggles with outliers, but the MMD posterior bootstrap is not significantly impacted

Misspecified g-and-k distribution

Time to run:
 NPL-MMD: ≈ 30 sec
 WABC: ≈ 100 sec



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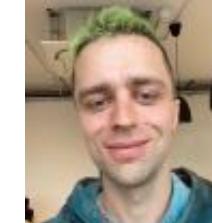
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Niu, Z., Meier, J., & **Briol, F.-X.** (2023). Discrepancy-based inference for intractable generative models using quasi-Monte Carlo. *Electronic Journal of Statistics*, 17(1), 1411–1456.

Bharti, A., Naslidnyk, M., Key, O., Kaski, S., & **Briol, F.-X.** (2023). Optimally-weighted estimators of the maximum mean discrepancy for likelihood-free inference. *International Conference on Machine Learning*, 2289–2312.

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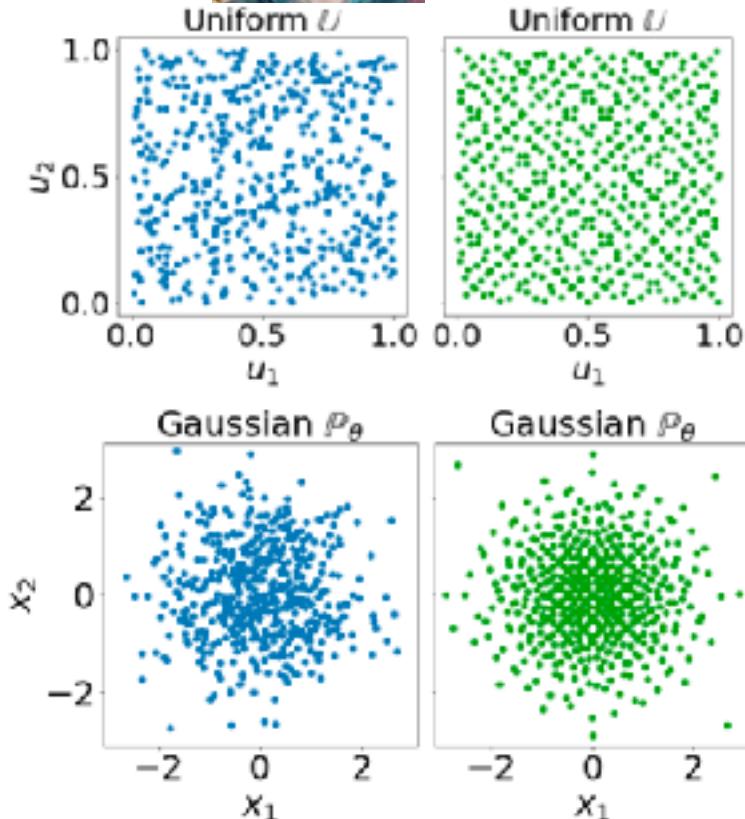
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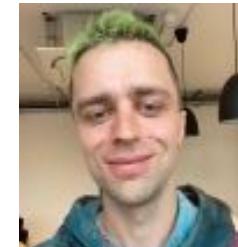
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Hypothesis testing for simulator misspecification



H_0 : Model/simulator is well-specified.

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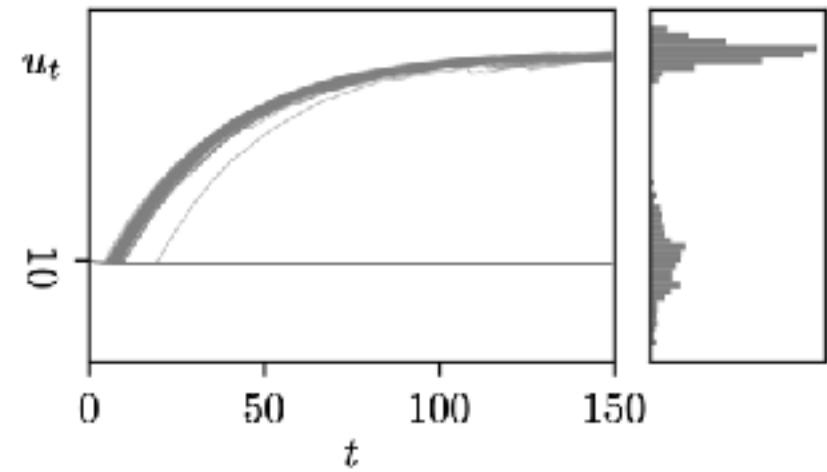
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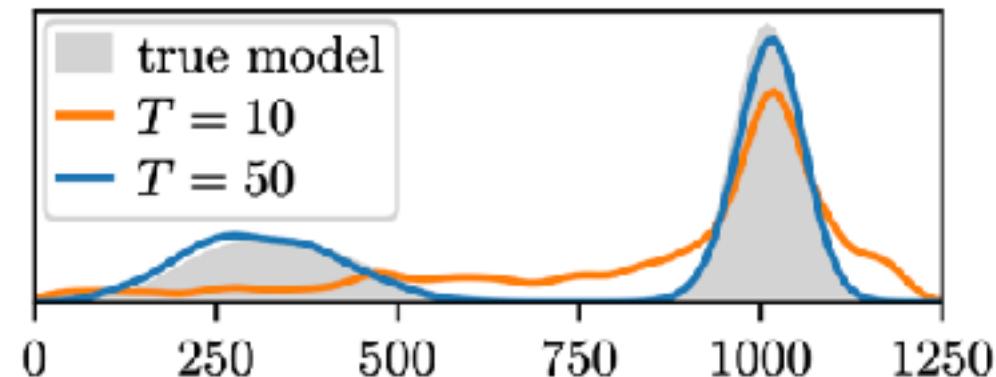
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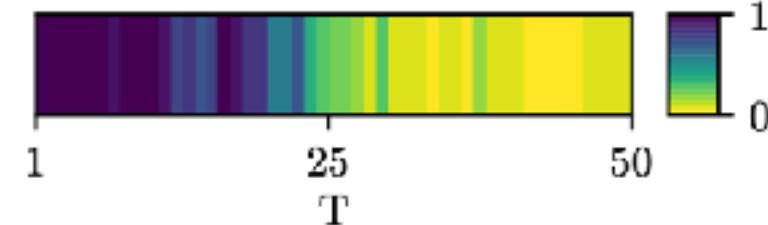
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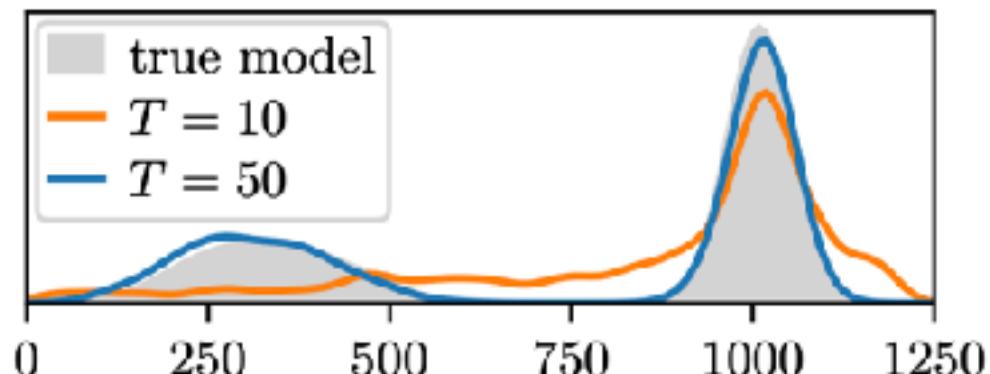
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% of rejects:



Toggle-switch model:



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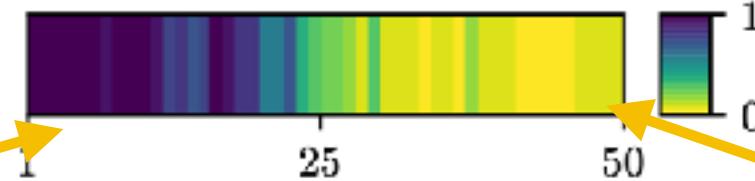
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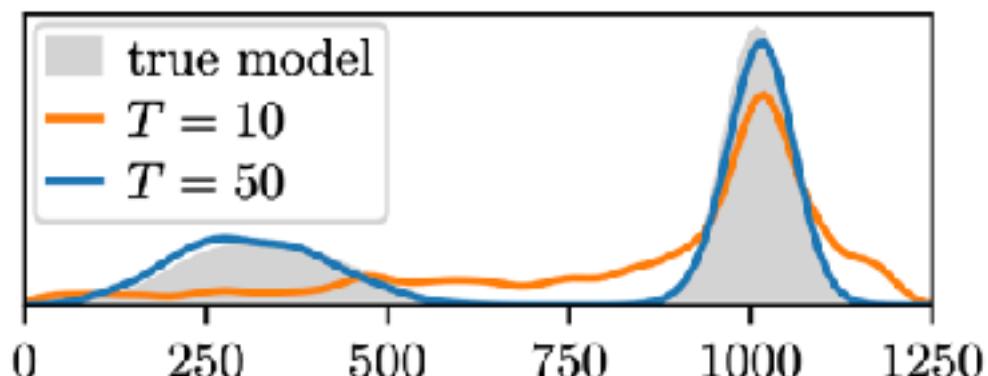
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Easy to distinguish

Toggle-switch model:



Looking ahead....

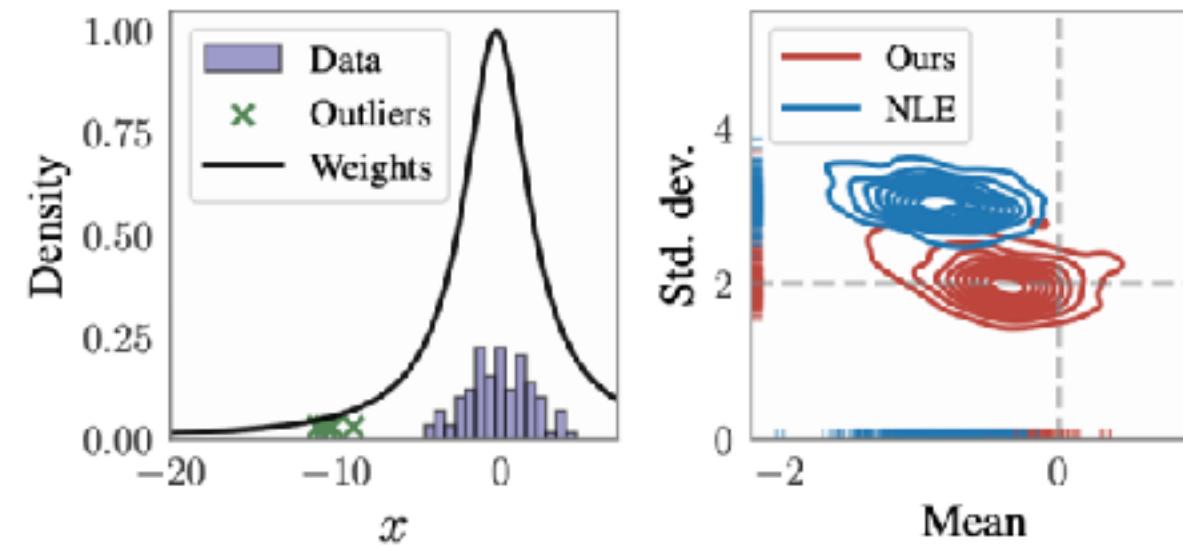


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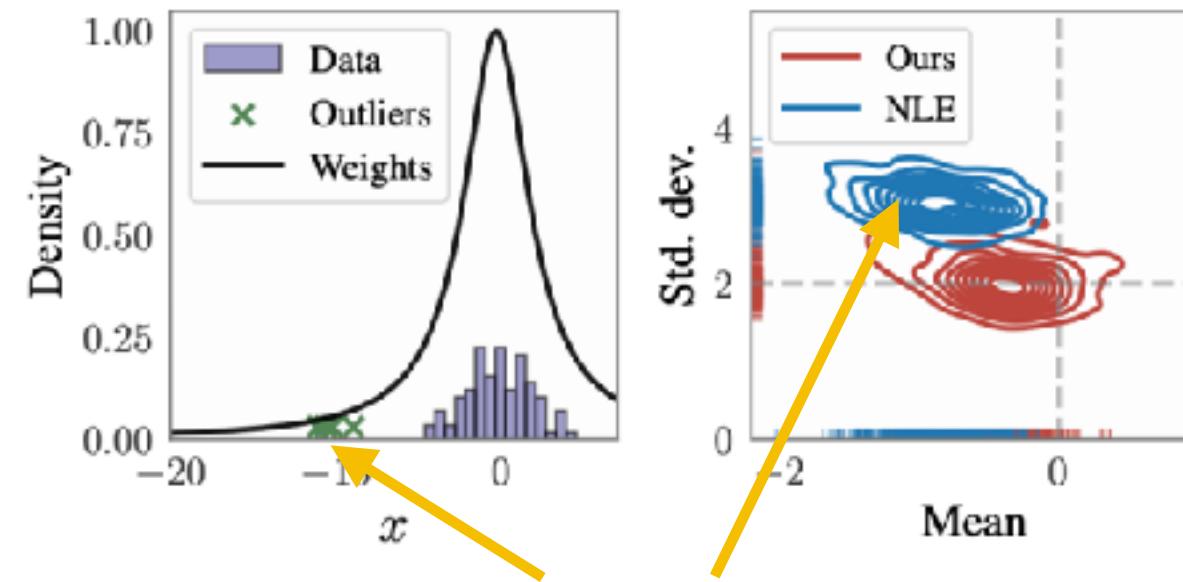
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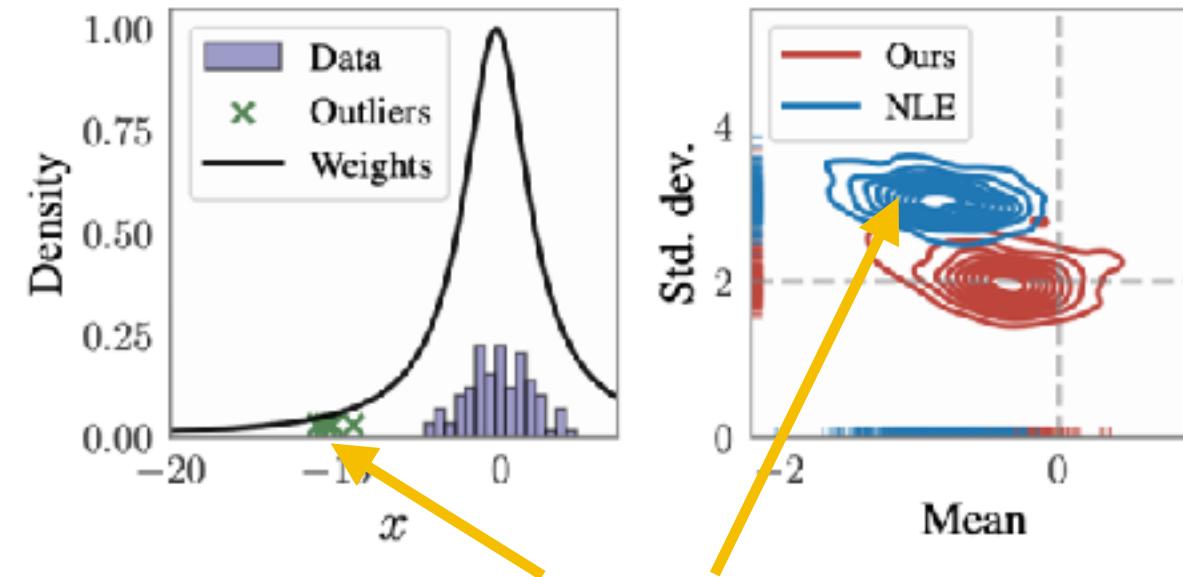


A few outliers can have a drastic impact on NLE!!

Looking ahead....



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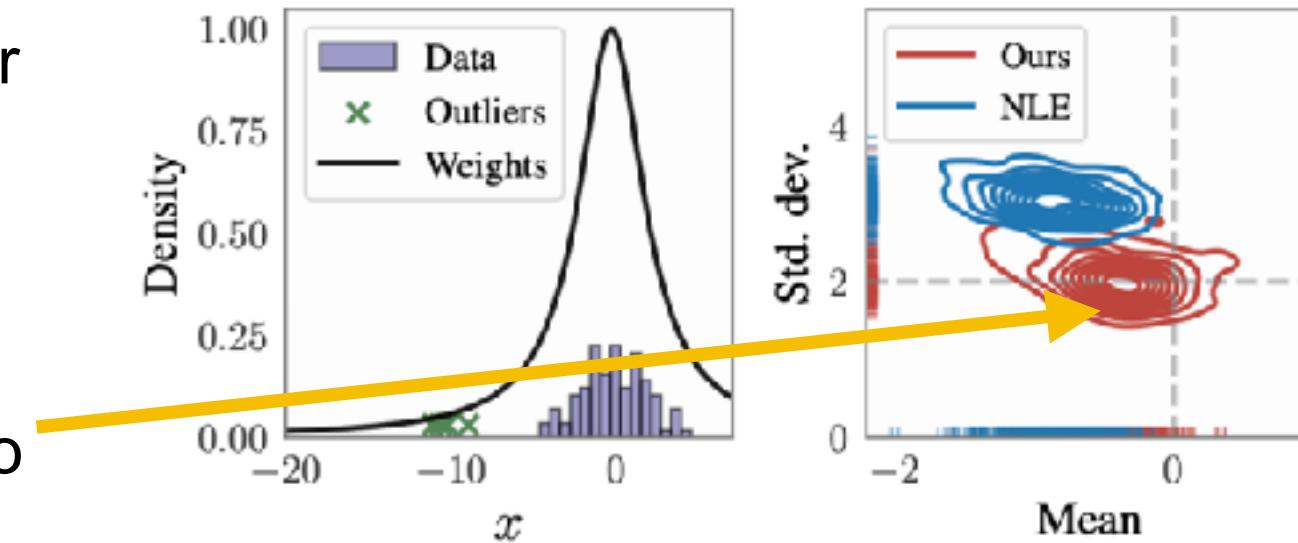


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Looking ahead....



- None of the methods in this section are well-suited for amortisation...
- Existing methods are either provably robust or amortised, but **not both**...!
- Currently working on a novel gen-Bayes method to resolve this.



Any Questions?

Paper: Dellaporta, C., Knoblauch, J., Damoulas, T. & **Briol, F-X** (2022). Robust Bayesian inference for simulator-based models via the MMD posterior bootstrap. AISTATS, 943-970. Best paper award.

Code: https://github.com/haritadell/npl_mmd_project

Summary of this course

- Basic methods:

Minimum distance
estimation

Approximate Bayesian
Computation

Neural simulation-
based inference

- Modern Challenges for SBI (expensive simulators, misspecification, calibration, high-dimensionality).
- Some illustrations of recent advances:

Hikida, Y., Bharti, A., Jeffrey, N. & **Briol, F-X** (2025). Multilevel neural simulation-based inference. arXiv:2506.06087 (to appear at NeurIPS?).

Bharti, A., Huang, D., Kaski, S., & **Briol, F.-X.** (2025). Cost-aware simulation-based inference. International Conference on Artificial Intelligence and Statistics, 28–36.

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- **Software:** sbi, bayesflow, etc...

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- Where should we go next?
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- Where are the computational statisticians (including me)?!
 - They were sleeping, but are slowly waking up to neural-based methods!

