

# Introductory Econometrics

## Lecture 11: Goodness of fit, estimation of $\sigma^2$

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## Fitted values

- ▶ Consider the multiple regression model with  $k$  regressors:  
 $Y_i = \beta_0 + \beta_1 X_{1,i} + \beta_2 X_{2,i} + \dots + \beta_k X_{k,i} + U_i.$
- ▶ Let  $\hat{\beta}_0, \hat{\beta}_1, \dots, \hat{\beta}_k$  be the OLS estimators.
- ▶ The fitted (or predicted) by the model value of  $Y$  is:  
 $\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_{1,i} + \hat{\beta}_2 X_{2,i} + \dots + \hat{\beta}_k X_{k,i}.$
- ▶ The residual is:  $\hat{U}_i = Y_i - \hat{Y}_i.$
- ▶ Consider the average of  $\hat{Y}$  :

$$\begin{aligned}\bar{\hat{Y}} &= \frac{1}{n} \sum_{i=1}^n \hat{Y}_i \\ &= \frac{1}{n} \sum_{i=1}^n (Y_i - \hat{U}_i) \\ &= \bar{Y} - \frac{1}{n} \sum_{i=1}^n \hat{U}_i = \bar{Y}\end{aligned}$$

because when there is an intercept,  $\sum_{i=1}^n \hat{U}_i = 0.$

# Sum-of-Squares

- The total variation of  $Y$  in the sample is:

$$SST = \sum_{i=1}^n (Y_i - \bar{Y})^2 \text{ (Total Sum-of-Squares).}$$

- The explained variation of  $Y$  in the sample is:

$$SSE = \sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2 \text{ (Explained or Model Sum-of-Squares).}$$

- The residual (unexplained or error) variation of  $Y$  in the sample is:

$$SSR = \sum_{i=1}^n \hat{U}_i^2 \text{ (Residual Sum-of-Squares).}$$

- If the regression contains an intercept:

$$SST = SSE + SSR.$$

# Proof of $SST=SSE+SSR$

► First,

$$\begin{aligned} SST &= \sum_{i=1}^n (Y_i - \bar{Y})^2 \\ &= \sum_{i=1}^n (\hat{Y}_i + \hat{U}_i - \bar{Y})^2 \\ &= \sum_{i=1}^n ((\hat{Y}_i - \bar{Y}) + \hat{U}_i)^2 \\ &= \sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2 + \sum_{i=1}^n \hat{U}_i^2 + 2 \sum_{i=1}^n (\hat{Y}_i - \bar{Y}) \hat{U}_i \end{aligned}$$

► Next, we will show that  $\sum_{i=1}^n (\hat{Y}_i - \bar{Y}) \hat{U}_i = 0$ .

# Proof of SST=SSE+SSR

- ▶ Since  $\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 X_{1,i} + \dots + \hat{\beta}_k X_{k,i}$ ,

$$\begin{aligned}\sum_{i=1}^n (\hat{Y}_i - \bar{Y}) \hat{U}_i &= \sum_{i=1}^n ((\hat{\beta}_0 + \hat{\beta}_1 X_{1,i} + \dots + \hat{\beta}_k X_{k,i}) - \bar{Y}) \hat{U}_i \\ &= \hat{\beta}_0 \sum_{i=1}^n \hat{U}_i + \hat{\beta}_1 \sum_{i=1}^n X_{1,i} \hat{U}_i + \dots + \hat{\beta}_k \sum_{i=1}^n X_{k,i} \hat{U}_i - \bar{Y} \sum_{i=1}^n \hat{U}_i.\end{aligned}$$

- ▶ The OLS normal equations for a model with an intercept:

$$\sum_{i=1}^n \hat{U}_i = \sum_{i=1}^n X_{1,i} \hat{U}_i = \dots = \sum_{i=1}^n X_{k,i} \hat{U}_i = 0.$$

- ▶ It follows that  $\sum_{i=1}^n (\hat{Y}_i - \bar{Y}) \hat{U}_i = 0$ .

$$R^2$$

- Consider the following measure of goodness of fit:

$$\begin{aligned} R^2 &= \frac{\sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2}{\sum_{i=1}^n (Y_i - \bar{Y})^2} \\ &= \frac{SSE}{SST} \\ &= 1 - \frac{SSR}{SST} \\ &= 1 - \frac{\sum_{i=1}^n \hat{U}_i^2}{\sum_{i=1}^n (Y_i - \bar{Y})^2}. \end{aligned}$$

- $0 \leq R^2 \leq 1$ .
- $R^2$  measures the proportion of variation in  $Y$  in the sample explained by the  $X$ 's.

$R^2$  is a non-decreasing function of the number of the regressors

- Consider two models:

$$Y_i = \tilde{\beta}_0 + \tilde{\beta}_1 X_{1,i} + \tilde{U}_i,$$

$$Y_i = \hat{\beta}_0 + \hat{\beta}_1 X_{1,i} + \hat{\beta}_2 X_{2,i} + \hat{U}_i.$$

- We will show that

$$\sum_{i=1}^n \tilde{U}_i^2 \geq \sum_{i=1}^n \hat{U}_i^2$$

and therefore  $R^2$  that corresponds to the regression with one regressor is less or equal than  $R^2$  that corresponds to the regression with two regressors.

- This can be generalized to the case of  $k$  and  $k + 1$  regressors.

# Proof

- Consider

$$\sum_{i=1}^n (\tilde{U}_i - \hat{U}_i)^2 = \sum_{i=1}^n \tilde{U}_i^2 + \sum_{i=1}^n \hat{U}_i^2 - 2 \sum_{i=1}^n \tilde{U}_i \hat{U}_i.$$

- We will show that

$$\sum_{i=1}^n \tilde{U}_i \hat{U}_i = \sum_{i=1}^n \hat{U}_i^2.$$

- Then,

$$0 \leq \sum_{i=1}^n (\tilde{U}_i - \hat{U}_i)^2 = \sum_{i=1}^n \tilde{U}_i^2 - \sum_{i=1}^n \hat{U}_i^2,$$

or

$$\sum_{i=1}^n \tilde{U}_i^2 \geq \sum_{i=1}^n \hat{U}_i^2.$$



# Proof

$$\begin{aligned}\sum_{i=1}^n \tilde{U}_i \hat{U}_i &= \sum_{i=1}^n (Y_i - \tilde{\beta}_0 - \tilde{\beta}_1 X_{1,i}) \hat{U}_i \\&= \sum_{i=1}^n Y_i \hat{U}_i - \tilde{\beta}_0 \sum_{i=1}^n \hat{U}_i - \tilde{\beta}_1 \sum_{i=1}^n X_{1,i} \hat{U}_i \\&= \sum_{i=1}^n Y_i \hat{U}_i - \tilde{\beta}_0 \cdot 0 - \tilde{\beta}_1 \cdot 0 \\&= \sum_{i=1}^n (\hat{\beta}_0 + \hat{\beta}_1 X_{1,i} + \hat{\beta}_2 X_{2,i} + \hat{U}_i) \hat{U}_i \\&= \hat{\beta}_0 \cdot 0 + \hat{\beta}_1 \cdot 0 + \hat{\beta}_2 \cdot 0 + \sum_{i=1}^n \hat{U}_i^2 \\&= \sum_{i=1}^n \hat{U}_i^2\end{aligned}$$

# Adjusted $R^2$

- ▶ Since  $R^2$  cannot decrease when more regressors are added, even if the additional regressors are irrelevant, an alternative measure of goodness-of-fit has been developed.
- ▶ Adjusted  $R^2$ : the idea is to adjust  $SSR$  and  $SST$  for degrees of freedom:

$$\bar{R}^2 = 1 - \frac{SSR/(n - k - 1)}{SST/(n - 1)}.$$

- ▶  $\bar{R}^2 < R^2$ .
- ▶  $\bar{R}^2$  can decrease when more regressors are added.

## Estimation of $\sigma^2$

- In the multiple linear regression model, we can estimate  $\sigma^2 = E[U_i^2]$  as follows:

Let

$$\hat{U}_i = Y_i - \hat{\beta}_0 - \hat{\beta}_1 X_{1,i} - \hat{\beta}_2 X_{2,i} - \dots - \hat{\beta}_k X_{k,i}.$$

An estimator for  $\sigma^2$  is

$$\begin{aligned} s^2 &= \frac{1}{n-k-1} \sum_{i=1}^n \hat{U}_i^2 \\ &= \frac{SSR}{n-k-1}. \end{aligned}$$

- The adjustment  $k+1$  is for the number of parameters we have to estimate in order to construct  $\hat{U}$ 's:

$$\hat{\beta}_0, \hat{\beta}_1, \dots, \hat{\beta}_k.$$

## Estimation of $\sigma^2$

$$s^2 = \frac{1}{n - k - 1} \sum_{i=1}^n \hat{U}_i^2.$$

- $s^2$  is an unbiased estimator of  $\sigma^2$  (i.e.,  $E[s^2] = \sigma^2$ ) when the following conditions hold:
1.  $Y_i = \beta_0 + \beta_1 X_{1,i} + \beta_2 X_{2,i} + \dots + \beta_k X_{k,i} + U_i$ .
  2. Conditional on  $X$ 's,  $E[U_i] = 0$  for all  $i$ 's.
  3. Conditional on  $X$ 's,  $E[U_i^2] = \sigma^2$  for all  $i$ 's (homoskedasticity).
  4. Conditional on  $X$ 's  $E[U_i U_j] = 0$  for all  $i \neq j$ .

# Stata

```
. regress rent avginc pop enroll
```

Source	SS	df	MS	Number of obs = 64		
Model	368241.042	3	122747.014	F( 3, 60)	=	29.05
Residual	253521.396	60	4225.35659	Prob > F	=	0.0000
Total	621762.438	63	9869.24504	R-squared	=	0.5923
				Adj R-squared	=	0.5719
				Root MSE	=	65.003

  

rent	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
avginc	.0119416	.001318	9.06	0.000	.0093052	.014578
pop	-.0003538	.0001621	-2.18	0.033	-.0006781	-.0000296
enroll	.0025595	.001264	2.02	0.047	.0000311	.0050879
_cons	120.772	34.53081	3.50	0.001	51.70009	189.8439

- ▶ We have 64 observations ( $n = 64$ ) and 3 regressors ( $k = 3$ ).
- ▶ SSE is displayed under Model SS (Sum of Squares): 368241.042.
- ▶ The Model df (degrees of freedom) is  $k = 3$ .
- ▶ The Model MS (Mean Squares) is  
 $SSE/k = 368241.042/3 = 122747.014$ .

# Stata

```
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- ▶ SSR is displayed under Residual SS: 253521.396.
- ▶ The Residual df is  $n - k - 1 = 64 - 3 - 1 = 60$ .
- ▶ The Residual MS is  $SSR / (n - k - 1) = s^2$ .
- ▶ The Residual MS is  $253521.396 / 60 = 4225.35659$ .

```
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```

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- SST is displayed under Total SS: 621762.438.
- The Total df is  $n - 1 = 64 - 1 = 63$ .
- The Total MS is  $SST/(n - 1) = 621762.438/63 = 9869.24504$ .

# Stata

```
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```

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- ▶  $R^2 = 1 - \frac{SSR}{SST} = 1 - \frac{253521.396}{621762.438} = 0.5923.$
- ▶  $\bar{R}^2 = 1 - \frac{SSR/(n-k-1)}{SST/(n-1)} = 1 - \frac{253521.396/60}{621762.438/63} = 0.5719.$
- ▶ Root MSE (Mean Squared Error) is  
 $s = \sqrt{s^2} = \sqrt{4225.35659} = 65.003.$